

C&EE 110

Warren Kim

June 13, 2023

## Discussion 5

### Question 1

Given:

$$X \leq 31$$
$$\lambda = \frac{1}{\mu} = \frac{1}{44} = 0.0227$$

Then:

(a)

$$P(X \leq 31) = \int_0^{31} \lambda e^{-\lambda x} dx$$
$$= \int_0^{31} \frac{1}{44} e^{-\frac{x}{44}} dx$$
$$= 1 - e^{-\frac{31}{44}}$$
$$P(X \leq 31) = 0.5057$$

(b)  $V(X) = \frac{1}{\lambda^2} = 44^2 = 1936$

### Question 2

Given:

$$\alpha = 11.6$$
$$\beta = 2.2$$
$$\text{Range: } [3.5, 25]$$

**Weibull Distribution:**

$$E(X) = \alpha \Gamma \left( 1 + \frac{1}{\beta} \right)$$
$$V(X) = \alpha^2 \left\{ \Gamma \left( 1 + \frac{2}{\beta} \right) - \left[ \Gamma \left( 1 + \frac{1}{\beta} \right) \right]^2 \right\}$$
$$f(X, \alpha, \beta) = \left( \frac{\beta}{\alpha} \right) \left( \frac{x}{\alpha} \right)^{\beta-1} e^{-\left( \frac{x}{\alpha} \right)^\beta}$$
$$F(X, \alpha, \beta) = 1 - e^{-\left( \frac{x}{\alpha} \right)^\beta}$$

(a)  $E(X) = 11.6 \cdot \Gamma \left( 1 + \frac{1}{2.2} \right) = 10.27 \text{ m/s.}$

(b)  $V(X) = 11.6^2 \cdot \left\{ \Gamma \left( 1 + \frac{2}{2.2} \right) - \left[ \Gamma \left( 1 + \frac{1}{2.2} \right) \right]^2 \right\} = 23.91 \text{ m/s} \implies \sigma = \sqrt{V(X)} = 4.89 \text{ m/s.}$

(c)  $P(3.5 \leq X \leq 25) = F(X \leq 25) - F(X \leq 3.5) = \left( 1 - e^{-\left( \frac{25}{11.6} \right)^{2.22}} \right) - \left( 1 - e^{-\left( \frac{3.5}{11.6} \right)^{2.22}} \right) = 0.9284$

(d)  $\beta = 1 \implies \text{Exponential Distribution:}$

(i)  $E(X) = \frac{1}{\lambda} = \alpha = 11.6 \text{ m/s.}$

(ii)  $V(X) = \frac{1}{\lambda^2} = \alpha^2 = 134.56 \text{ (m/s)}^2$

(iii)

$$f(X, \lambda) = \lambda e^{-\lambda x}$$
$$F(X, \alpha, \beta) = 1 - e^{-\lambda x}$$
$$P(3.5 \leq X \leq 25) = \left( 1 - e^{-(11.6)(25)} \right) - \left( 1 - e^{-(11.6)(3.5)} \right) = 0.624$$

### Question 3

Given:

$$\mu = 0.35$$

$$\sigma = 0.2$$

$$T = 0.6$$

*Log-Normal Distribution:*  $\Phi$

$$\begin{aligned}\sigma_T^2 &= \ln \left( \frac{\sigma^2}{\mu^2} + 1 \right) \\ &= \ln \left( \frac{0.2^2}{0.35^2} + 1 \right)\end{aligned}$$

$$\sigma_T^2 = 0.283$$

$$\begin{aligned}\mu_T^2 &= \ln \mu - \frac{1}{2} \sigma_T^2 \\ &= \ln 0.35 - \frac{1}{2} (0.283) \\ \mu_T^2 &= -1.19\end{aligned}$$

$$(a) \quad P(T < 0.6) = \Phi \left( \frac{\ln 0.6 - \mu_T}{\sigma_T} \right) = \Phi(1.276) = 0.898.$$

(b)

$$\begin{aligned}P(T > t^*) &= 0.95 \iff 1 - P(T \leq t^*) = 0.95 \implies \Phi \left( \frac{\ln T^* - \mu_T}{\sigma_T} \right) = 0.05 \\ \implies -1.65 &= \frac{\ln T^* - \mu_T}{\sigma_T} \\ &= \exp(-1.65 \cdot \sqrt{0.283} - 1.19) \\ P(T > t^*) &= 0.126 \text{ s}\end{aligned}$$

### Question 4

Given:

$$\mu_X = 100 \text{ kips}$$

$$\sigma_X = 10 \text{ kips}$$

$$\mu_Y = 40 \text{ kips}$$

$$\sigma_Y = 10 \text{ kips}$$

*Assumed Normal Distribution:*

(a)

$$\begin{aligned}\mu_Z &= \mu_X + \mu_Y = 100 + 40 = 140 \\ \sigma_Z &= \sqrt{\sigma_X^2 + \sigma_Y^2} = \sqrt{10^2 + 10^2} = 14.14 \text{ kips}\end{aligned}$$

(b)

$$\begin{aligned}P(Z \geq z^*) &= 0.05 \iff 1 - P(Z < z^*) = 0.05 \\&\implies 1 - P(Z < z^*) = 1 - \Phi\left(\frac{z^* - \mu_Z}{\sigma_Z}\right) \\&= 0.05 \\&\implies \Phi\left(\frac{z^* - \mu_Z}{\sigma_Z}\right) = 0.95 \\&\implies z^* = (1.65)(14.14) + 140 = 163.33 \text{ kips}\end{aligned}$$

(c)  $P(Z < 170) = \Phi\left(\frac{170-140}{14.14}\right) = \Phi(2.12) = 0.983.$

## Discussion 6

### Question 1

Given:

$$d = \sqrt{r^2 - 40^2}$$

We have that  $P(R \leq r) = F(r) = \frac{\text{length of the fault within distance } r}{\text{total length of the fault}} = \frac{2\sqrt{r^2 - 40^2}}{60}$  Then, we have

$$F(r) = \begin{cases} \frac{2\sqrt{r^2 - 40^2}}{60} & 40 \leq r \leq 50 \\ 1 & r > 50 \\ 0 & r < 40 \end{cases}$$

and

$$f(r) = \begin{cases} \frac{r}{30\sqrt{r^2 - 40^2}} & 40 \leq r \leq 50 \\ 1 & r > 50 \\ 0 & r < 40 \end{cases}$$

### Question 2

Given:

$$f(y_1, y_2) = \begin{cases} k(y_1 + y_2) & 1 \leq y_1 \leq 2, 4 \leq y_2 \leq 5 \\ 0 & \text{else} \end{cases}$$

**PDF:**

$$\int_{-\infty}^{\infty} \int_{-\infty}^{\infty} f(y_1, y_2) dy_1 dy_2 = 1$$

(a)

$$\begin{aligned} \int_1^2 \int_4^5 f(y_1, y_2) dy_1 dy_2 &= \int_1^2 \int_4^5 k(y_1 + y_2) dy_1 dy_2 = 6k \\ \implies k &= \frac{1}{6} \end{aligned}$$

(b)

$$\begin{aligned} P(Y_1 \leq y_1, Y_2 \leq y_2) &= \int_1^{y_1} \int_4^{y_2} \frac{1}{6} (t_1 + t_2) dt_1 dt_2 \\ &= \frac{1}{6} \int_1^{y_1} \left[ t_1 t_2 + \frac{1}{2} t_2^2 \right] \Big|_4^{y_2} dy_1 \\ &= \frac{1}{6} \int_1^{y_1} \left[ t_1 (y_2 - 4) + \frac{1}{2} y_2^2 - \frac{16}{2} \right] dy_1 \\ &= \frac{1}{6} \left[ \frac{1}{2} t_1^2 (y_2 - 4) + \frac{1}{2} y_2^2 - 8 \right] \Big|_1^{y_1} \\ P(Y_1 \leq y_1, Y_2 \leq y_2) &= \frac{1}{12} (y_1 - 1)(y_2 - 4)(y_1 + y_2 + 5) \quad 1 \leq y_1 \leq 2, 4 \leq y_2 \leq 5 \end{aligned}$$

$$(c) P(Y_1 \leq 1.5, Y_2 \leq 4.5) = (1.5 - 1)(4.5 - 4)(1.5 + 4.5 + 5) = 0.229$$

(d)

$$\begin{aligned}f_1(y_1) &= \int_4^5 \frac{1}{6}(y_1 + y_2)dy_2 \\&= \frac{1}{6} \int_4^5 (y_1 + y_2)dy_2 \\&= \frac{1}{6} \left[ y_1 + \frac{25}{2} - \frac{16}{2} \right] \\f_1(y_1) &= \frac{1}{6} \left( y_2 + \frac{3}{2} \right) \qquad 4 \leq y_2 \leq 5\end{aligned}$$

(e)

$$\begin{aligned}P(Y_1 \leq 1.5 | Y_2 \geq 4.5) &= \frac{P(Y_1 \leq 1.5 \cap Y_2 \geq 4.5)}{P(Y_1 \geq 4.5)} \\&= \frac{\int_1^{1.5} \int_{4.5}^5 \frac{1}{6}(y_1 + y_2)dy_2dy_1}{\int_{4.5}^5 \frac{1}{6} \left( y_2 + \frac{3}{2} \right) dy_2} \\P(Y_1 \leq 1.5 | Y_2 \geq 4.5) &= \frac{0.25}{0.52} = 0.48\end{aligned}$$

## Discussion 7

### Question 1

Given:

$$\begin{aligned}n &= 50 \\ \mu &= 12.68 \\ \sigma &= 6.83\end{aligned}$$

**Assumed Normal Distribution:**

$$P(|\bar{X} - \mu| \leq 0) = 0.95$$

$$P(\mu - z_{\frac{\alpha}{2}} \cdot \sigma_{\bar{X}} \leq \mu_{\bar{X}} \leq \mu + z_{\frac{\alpha}{2}} \cdot \sigma_{\bar{X}}) = 1 - \alpha$$

where the confidence interval is  $1 - \alpha = 0.95$  with significance level 0.05.

(a)

$$\begin{aligned}z_{\frac{\alpha}{2}} &= z_{\frac{0.05}{2}} = z_{0.025} = 1.96 && \text{from z-table} \\ \sigma_{\bar{X}} &= \frac{\sigma}{\sqrt{n}} = \frac{6.83}{\sqrt{50}} = 0.9695\end{aligned}$$

Then we get

$$\begin{aligned}P(\mu - z_{\frac{\alpha}{2}} \cdot \sigma_{\bar{X}} \leq \mu_{\bar{X}} \leq \mu + z_{\frac{\alpha}{2}} \cdot \sigma_{\bar{X}}) &= P(12.68 - (1.96)(0.9695) \leq \mu_{\bar{X}} \leq 12.68 + (1.96)(0.9695)) = 0.95 \\ \implies 95\% \text{ confidence interval is } 12.68 \pm 1.89 &\iff (10.79, 14.57)\end{aligned}$$

(b)

$$\begin{aligned}P(11.09 \leq \mu_{\bar{X}} \leq 14.27) &= 1 - \alpha \\ \implies 14.27 &= \mu + z_{\frac{\alpha}{2}} \cdot \frac{\sigma}{\sqrt{n}} \\ &= 12.68 + z_{\frac{\alpha}{2}} \cdot 0.9695 \\ z_{\frac{\alpha}{2}} &= 1.646 \\ \implies \frac{\alpha}{2} &= 0.05 \implies \alpha = 0.1, \text{ confidence interval is } 100(1 - \alpha) = 90\%\end{aligned}$$

**Question 2-4:** See Canvas

**Discussion 8/9:** See Canvas

## MLE

**Note:**  $\ln \iff \log$ ,  $\exp x = e^x$

Given

$$L = \prod_{i=1}^N p(x_i) = \prod_{i=1}^N \frac{\lambda^{x_i} \exp(-\lambda)}{x_i!}$$

$$\begin{aligned}LL &= \log \left\{ \prod_{i=1}^N \frac{\lambda^{x_i} \exp(-\lambda)}{x_i!} \right\} \\ &= \sum_{i=1}^N \log(\lambda^{x_i} \exp(-\lambda)) - \log(x_i!) && \log\left(\frac{x}{y}\right) = \log(x) - \log(y) \\ &= \sum_{i=1}^N -\lambda + (x_i) \log(\lambda) - \log(x_i!) && \log(xy) = \log(x) + \log(y), \log(x^a) = a \log(x), \log(e^x) = x \\ LL &= -N\lambda + \sum_{i=1}^N [(x_i) \log(\lambda) - \log(x_i!)]\end{aligned}$$

Then  $\frac{dLL}{d\lambda}$  is

$$\frac{dLL}{d\lambda} = -N\lambda$$