

Writing Exercise

Bochen Wang
Tianyuan Liu

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1 Exercise 4.1

1.1

$$\begin{aligned} \max_a \quad & a^T B a \\ \text{s.t.} \quad & a^T W a = 1 \end{aligned}$$

corresponding to the lagrangian

$$\begin{aligned} L_P &= a^T B a - \lambda(a^T W a - 1) \\ \text{Let } \frac{\partial L_P}{\partial a} &= 2Ba - 2\lambda W a = 0 \\ W^{-1} B a &= \lambda a \end{aligned}$$

$\therefore a$ is a eigenvector of $W^{-1}B$

$$\begin{aligned} \therefore B a &= \lambda W a \\ \therefore a^T B a &= \lambda a^T W a \\ \therefore a^T W a &= 1 \\ \therefore a^T B a &= \lambda \end{aligned}$$

$\therefore a$ is the eigenvector corresponding to the biggest eigenvalue of $W^{-1}B$

2 Exercise 4.2

2.1

From equation 4.10 we know the linear discriminant function is:

$$\delta_k(x) = x^T \Sigma^{-1} \mu_k - \frac{1}{2} \mu_k^T \Sigma^{-1} \mu_k + \log \pi_k$$

so LDA classifies to class 2 if

$$\begin{aligned}
\delta_2(x) &> \delta_1(x) \\
x^T \Sigma^{-1} \mu_2 - \frac{1}{2} \mu_2^T \Sigma^{-1} \mu_2 + \log \pi_2 &> x^T \Sigma^{-1} \mu_1 - \frac{1}{2} \mu_1^T \Sigma^{-1} \mu_1 + \log \pi_1 \\
x^T \Sigma^{-1} (\mu_2 - \mu_1) &> \frac{1}{2} \mu_2^T \Sigma^{-1} \mu_2 - \frac{1}{2} \mu_1^T \Sigma^{-1} \mu_1 + \frac{1}{2} \mu_1^T \Sigma^{-1} \mu_2 \\
&- \left(\frac{1}{2} \mu_1^T \Sigma^{-1} \mu_2 \right)^T + \log(N_1) - \log(N_2) \\
x^T \Sigma^{-1} (\mu_2 - \mu_1) &> \frac{1}{2} (\mu_2 + \mu_1)^T \Sigma^{-1} (\mu_2 - \mu_1) - \log\left(\frac{N_2}{N_1}\right)
\end{aligned}$$

2.2

Let U_i be the class indicator vector of class i , so $U = U_1 + U_2$ is a vector with all entries 1. When we encode the target as $a_1 = -\frac{N}{N_1}$ and $a_2 = \frac{N}{N_2}$, we have $Y = a_1 U_1 + a_2 U_2$. Then, $RSS = \sum_{i=1}^N (y_i - \beta_0 - \beta^T x_i)^2 = (Y - \beta_0 U - X\beta)^T (Y - \beta_0 U - X\beta)$. To minimize the RSS,

$$\begin{aligned}
\frac{\partial RSS}{\partial \beta} &= 2X^T X\beta - 2X^T Y + 2\beta_0 X^T U = 0 \\
\frac{\partial RSS}{\partial \beta_0} &= 2U^T U\beta_0 - 2U^T (Y - X\beta) = 2N\beta_0 - 2U^T (Y - X\beta) = 0
\end{aligned}$$

From the second equation, we get:

$$\beta_0 = \frac{U^T (Y - X\beta)}{N} \quad (1)$$

Plug it back to first equation, we get:

$$(X^T X - \frac{X^T U U^T X}{N})\beta = X^T Y - \frac{X^T U U^T Y}{N} \quad (2)$$

For l.h.s:

Notice that $X^T U = X^T (U_1 + U_2) = N_1 \mu_1 + N_2 \mu_2$. plug it to l.h.s. We get:

$$l.h.s = X^T X - \frac{1}{N} (N_1^2 \mu_1 \mu_1^T + N_2^2 \mu_2 \mu_2^T + N_1 N_2 \mu_1 \mu_2^T + N_1 N_2 \mu_2 \mu_1^T) \quad (3)$$

The estimate of the covariance matrix used in LDA is:

$$\begin{aligned}
(N-2)\Sigma &= \sum_{i:y_i=a_1} (x_i - \mu_1)(x_i - \mu_1)^T + \sum_{i:y_i=a_2} (x_i - \mu_2)(x_i - \mu_2)^T \\
&= X^T X - N_1 \mu_1 \mu_1^T - N_2 \mu_2 \mu_2^T
\end{aligned}$$

Also $\Sigma_B = (\mu_1 - \mu_2)(\mu_1 - \mu_2)^T$. So

$$\begin{aligned}
(N-2)\Sigma + \frac{N_1 N_2}{N} \Sigma_B &= X^T X + \left(\frac{N_1 N_2}{N} - N_1\right) \mu_1 \mu_1^T + \left(\frac{N_1 N_2}{N} - N_2\right) \mu_2 \mu_2^T \\
&- \frac{N_1 N_2}{N} \mu_1 \mu_2^T - \frac{N_1 N_2}{N} \mu_2 \mu_1^T \\
&= X^T X - \frac{N_1^2}{N} \mu_1 \mu_1^T - \frac{N_2^2}{N} \mu_2 \mu_2^T - \frac{N_1 N_2}{N} \mu_1 \mu_2^T - \frac{N_1 N_2}{N} \mu_2 \mu_1^T
\end{aligned}$$

Which is the same as equation(3).

For r.h.s: The first term is equal to :

$$X^T Y = X^T (a_1 U_1 + a_2 U_2) = a_1 N_1 \mu_1 + a_2 N_2 \mu_2$$

The second term is equal to:

$$\begin{aligned} \frac{1}{N} X^T U U^T Y &= \frac{1}{N} (N_1 \mu_1 + N_2 \mu_2) \\ &= \frac{a_1 N_1^2 + a_2 N_1 N_2}{N} \mu_1 + \frac{a_2 N_2^2 + a_1 N_1 N_2}{N} \mu_2 \end{aligned}$$

Combing two terms together, and with l.h.s we get:

$$[(N-2)\Sigma + \frac{N_1 N_2}{N} \Sigma_B] \beta = \frac{N_1 N_2}{N} (a_1 - a_2) (\mu_1 - \mu_2) \quad (4)$$

plug $a_1 = -\frac{N}{N_1}$ and $a_2 = \frac{N}{N_2}$ back to the equation:

$$r.h.s = \frac{N_1 N_2}{N} \left(-\frac{N^2}{N_1 N_2} \right) (\mu_1 - \mu_2) = N(\mu_2 - \mu_1) \quad (5)$$

So we get:

$$[(N-2)\Sigma + \frac{N_1 N_2}{N} \Sigma_B] \beta = N(\mu_2 - \mu_1) \quad (6)$$

2.3

Let $c = (\mu_2 - \mu_1)^T \beta$

$$\Sigma_B \beta = (\mu_2 - \mu_1)(\mu_2 - \mu_1)^T \beta = c(\mu_2 - \mu_1)$$

$\Sigma_B \beta$ is int the direction of $(\mu_2 - \mu_1)$. Combined with result of last question, $\beta \propto \Sigma^{-1}(\mu_2 - \mu_1)$. Therefore the least-squares regression coefficient is identical to the LDA coefficient, up to a scalar multiple.

2.4

Since (5) holds for any encoding a_1 and a_2 , the result also holds for any encoding.

2.5

Notice that $U^T Y = a_1 N_1 + a_2 N_2 = -N + N = 0$. So from equation(1), we get: $\beta_0 = -\frac{1}{N} (U^T X \beta) = -\left(\frac{N_1}{N} \mu_1^T + \frac{N_2}{N} \mu_2^T \right) \beta$. so the decision function is:

$$f(x) = x^T \beta + \beta_0 = \left(X^T - \frac{N_1}{N} \mu_1^T - \frac{N_2}{N} \mu_2^T \right) \beta$$

Notice from the result of 1.3. We have $\beta = c \Sigma^{-1}(\mu_2 - \mu_1)$ for some value c . The decision is:

$$\begin{aligned} f(x) &> 0 \\ x^T \Sigma^{-1}(\mu_2 - \mu_1) &> \left(\frac{N_1}{N} \mu_1^T + \frac{N_2}{N} \mu_2^T \right) \Sigma^{-1}(\mu_2 - \mu_1) \end{aligned}$$

So this is not the same as the LDA rule unless the classes have equal numbers of observations.

Reference:

<http://cbio.ensmp.fr/~jvert/svn/tutorials/course/1102Senegal/solution.pdf>

3 RLU Exercise 11.3.1

3.1

$$\begin{aligned}
 M^T M &= \begin{bmatrix} 1 & 3 & 5 & 0 & 1 \\ 2 & 4 & 4 & 2 & 3 \\ 3 & 5 & 3 & 4 & 5 \end{bmatrix} \begin{bmatrix} 1 & 2 & 3 \\ 3 & 4 & 5 \\ 5 & 4 & 3 \\ 0 & 2 & 4 \\ 1 & 3 & 5 \end{bmatrix} \\
 &= \begin{bmatrix} 36 & 37 & 38 \\ 37 & 49 & 61 \\ 38 & 61 & 84 \end{bmatrix} \\
 MM^T &= \begin{bmatrix} 1 & 2 & 3 \\ 3 & 4 & 5 \\ 5 & 4 & 3 \\ 0 & 2 & 4 \\ 1 & 3 & 5 \end{bmatrix} \begin{bmatrix} 1 & 3 & 5 & 0 & 1 \\ 2 & 4 & 4 & 2 & 3 \\ 3 & 5 & 3 & 4 & 5 \end{bmatrix} \\
 &= \begin{bmatrix} 14 & 26 & 22 & 16 & 22 \\ 26 & 50 & 46 & 28 & 40 \\ 22 & 46 & 50 & 20 & 32 \\ 16 & 28 & 20 & 20 & 26 \\ 22 & 40 & 32 & 26 & 35 \end{bmatrix}
 \end{aligned}$$

3.2

Two eigenvalues for $M^T M$ and MM^T are: **15.43** and **153.57**.

3.3

Two eigenvectors for $M^T M$ are:

$$\begin{bmatrix} -0.816 & 0.409 \\ -0.126 & 0.564 \\ 0.564 & 0.718 \end{bmatrix}$$

Two eigenvectors for MM^T are:

$$\begin{bmatrix} 0.159 & 0.298 \\ -0.033 & 0.571 \\ -0.736 & 0.521 \\ 0.510 & 0.323 \\ 0.414 & 0.459 \end{bmatrix}$$

3.4

$$\begin{aligned}
 MM^T &= (UDV^T)(UDV^T)^T \\
 &= UDV^TVD^TU^T \\
 &= UD^2U^T
 \end{aligned}$$

This is the eigendecomposition of MM^T .
 So U should be a N * N matrix where ith column is the ith eigenvector.
 U is:

$$\begin{bmatrix} 0.298 & 0.159 \\ 0.571 & -0.033 \\ 0.521 & -0.736 \\ 0.323 & 0.510 \\ 0.459 & 0.414 \end{bmatrix}$$

Similarly, V is:

$$\begin{bmatrix} 0.409 & -0.816 \\ 0.564 & -0.126 \\ 0.718 & 0.564 \end{bmatrix}$$

Σ is the diagonal matrix which the ith diagonal value is the square of ith eigenvalue.

$$\begin{bmatrix} 12.39 & 0 \\ 0 & 3.93 \end{bmatrix}$$

3.5

$$\begin{bmatrix} 0.298 \\ 0.571 \\ 0.521 \\ 0.323 \\ 0.459 \end{bmatrix} * 12.39 * \begin{bmatrix} 0.409 \\ 0.564 \\ 0.718 \end{bmatrix}^T = \begin{bmatrix} 1.51 & 2.08 & 2.65 \\ 2.89 & 3.98 & 5.07 \\ 2.64 & 3.64 & 4.64 \\ 1.64 & 2.25 & 2.87 \\ 2.33 & 3.20 & 4.08 \end{bmatrix}$$

3.6

The energy retained is $\frac{12.39^2}{12.39^2+3.93^2} = 0.91$.