

FIN 591: Homework #3

Due on Wednesday, April 11, 2018

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Problem 1

- a. Since the final payoff of P is 1, using continuous-time version stochastic discount factor, $P_t(\tau)$ is derived as follows.

$$\begin{aligned} P_t(\tau) &= E_t \left[\frac{U_c(C_{t+\tau, t+\tau})}{U_c(C_t, t)} \times 1 \right] \\ &= E_t \left[\frac{e^{-\phi(t+\tau)} C_{t+\tau}^{\gamma-1}}{e^{\phi t} C_t^{\gamma-1}} \right] \\ &= E_t \left[e^{-\phi \tau} \frac{C_{t+\tau}^{\gamma-1}}{C_t^{\gamma-1}} \right] \end{aligned} \quad (1)$$

- b. From $P_t(\tau) = E_t \left[\frac{e^{-\phi(t+\tau)} C_{t+\tau}^{\gamma-1}}{e^{\phi t} C_t^{\gamma-1}} \right]$, we can find that process M_t is equal to $e^{-\phi t} C_t^{\gamma-1}$. Therefore, using Ito's lemma, dynamics of M_t can be derived as equation (2).

$$\begin{aligned} dM_t &= -\phi e^{-\phi t} C_t^{\gamma-1} dt + e^{-\phi t} (\gamma-1) C_t^{\gamma-2} C [(\mu_c - \lambda k) dt + \sigma_c dZ_c] \\ &\quad + \frac{1}{2} e^{-\phi t} (\gamma-1)(\gamma-2) C^2 C_t^{\gamma-3} \sigma_c^2 dt + [e^{-\phi t} (Y C)^{\gamma-1} - e^{-\phi t} C^{\gamma-1}] dq \\ &= [-\phi + (\gamma-1)(\mu_c - \lambda k) + \frac{1}{2} (\gamma-1)(\gamma-2) \sigma_c^2] M_t dt + (\gamma-1) \sigma_c M_t dZ_c + (Y^{\gamma-1} - 1) M_t dq \end{aligned} \quad (2)$$

- c. Since $E \left[\frac{dM}{M} \right] = -r dt$, the following equation holds.

$$\begin{aligned} r &= -E \left[\frac{dM}{M} \right] \\ &= \phi - (\gamma-1)(\mu_c - \lambda k) - \frac{1}{2} (\gamma-1)(\gamma-2) \sigma_c^2 - \lambda E[e^{(\gamma-1) \log Y} - 1] \\ &= \phi - (\gamma-1)(\mu_c - \lambda k) - \frac{1}{2} (\gamma-1)(\gamma-2) \sigma_c^2 - \lambda (e^{(\gamma-1)\alpha + \frac{1}{2}(\gamma-1)^2 \delta^2} - 1) \end{aligned} \quad (3)$$

Since μ_c, k, λ are constant, instantaneous risk free rate is constant.

- d. Since an asset price is equal to sum of discounted future payoffs, assuming some regularity conditions hold, S_t is represented as follows.

$$\begin{aligned} S_t &= E_t \left[\int_t^\infty \frac{M_s}{M_t} D_s ds \right] \\ &= E_t \left[\int_t^\infty \frac{e^{-\phi s} C_s^{\gamma-1}}{e^{-\phi t} C_t^{\gamma-1}} D_s ds \right] \\ \Rightarrow \frac{S_t}{D_t} &= E_t \left[\int_t^\infty e^{-\phi(s-t)} \left(\frac{C_s}{C_t} \right)^{\gamma-1} \left(\frac{D_s}{D_t} \right) ds \right] \\ &= E_t \left[\int_t^\infty e^{-\phi(s-t) + (\gamma-1) \log(C_s/C_t) + \log(D_s/D_t)} ds \right] \\ &= \int_t^\infty E_t \left[e^{-\phi(s-t) + (\gamma-1) \log(C_s/C_t) + \log(D_s/D_t)} ds \right] \end{aligned} \quad (4)$$

Considering the process of C_t , $E_t[e^{(\gamma-1) \log(C_s/C_t)}]$ is calculated as follows.

$$\begin{aligned} E_t[e^{(\gamma-1) \log(C_s/C_t)}] &= E_t \left[e^{(\gamma-1)(\mu_c - \frac{1}{2} \sigma_c^2 - \lambda k)(s-t) + \frac{1}{2} (\gamma-1)^2 \sigma_c^2 (s-t) + (\gamma-1) \log y(s,t)} \right] \\ y(s, t) &= \prod_{i=s}^t Y_i \end{aligned} \quad (5)$$

Since $\log y(s, t) = \sum_{i=s}^t \log Y_i$, and $\log Y_i$'s are independently and identically distributed as $N(\alpha, \delta^2)$, $\log(C_s/C_t)$ is also normally distributed, and its expected value from equation (5) is calculated as follows.

$$E_t[e^{(\gamma-1)\log(C_s/C_t)}] = e^{(\gamma-1)(\mu_c - \frac{1}{2}\sigma_c^2 - \lambda k)(s-t) + \frac{1}{2}(\gamma-1)^2\sigma_c^2(s-t) + (\gamma-1)(\alpha + \frac{1}{2}\delta^2)(s-t)} \quad (6)$$

Applying the result from equation (6) and considering the correlation between z_d and z_c is ρ , $\frac{S_t}{D_t}$ from equation (4) is solved as follows.

$$\begin{aligned} \frac{S_t}{D_t} &= \int_t^\infty e^{-(s-t)[\phi + (1-\gamma)(\mu_c - \frac{1}{2}\sigma_c^2 - \lambda k) + \frac{1}{2}(1-\gamma)^2\sigma_c^2 + (1-\gamma)(\alpha + \frac{1}{2}\delta^2) - \mu_d + (1-\gamma)\rho\sigma_c\sigma_d]} ds \\ &= -\frac{1}{A} e^{-(s-t)A} \Big|_t^\infty = \frac{1}{A} \\ A &= \phi + (1-\gamma)(\mu_c - \frac{1}{2}\sigma_c^2 - \lambda k) + \frac{1}{2}(1-\gamma)^2\sigma_c^2 + (1-\gamma)(\alpha + \frac{1}{2}\delta^2) - \mu_d + (1-\gamma)\rho\sigma_c\sigma_d \end{aligned} \quad (7)$$

Problem 2

a. Considering the process of risky asset price, intertemporal budget constraint is derived as follows.

$$\begin{aligned} dW &= \omega \frac{dS}{S} + (1-\omega)r dt - C dt \\ &= (\omega(\mu - \lambda k - r)W + rW - C)dt + \sigma W dZ + \omega(Y-1)W dq \end{aligned} \quad (8)$$

b. Investors maximize $E_0[\int_0^T e^{-\phi t} u(C_t) dt]$, subject to $dW = (\omega(\mu - \lambda k - r)W + rW)dt + \sigma W dZ + \omega(Y-1)W dq$.

Let $J(W, s) = \max_{C, \omega} E_s[\int_0^T e^{-\phi t} u(C_t) dt]$. Then the following equation follows.

$$\begin{aligned} J(W, 0) &= \max_{C, \omega} E_0 \left[\int_0^{\Delta t} e^{-\phi t} u(C_t) dt + J(W, \Delta t) \right] \\ &= \max_{C, \omega} E_0 [u(C_0)\Delta t + J(W, 0) + J_W(\omega(\mu - \lambda k - r)W + rW - C)\Delta t \\ &\quad + \frac{1}{2}\omega^2\sigma^2 J_{WW}\Delta t + (J(\omega(Y-1)W, 0) - J(W, 0))dq] \end{aligned} \quad (9)$$

Letting $\Delta t \rightarrow 0$, equation (9) becomes equation (10), and it is Bellman equation.

$$0 = \max_{C, \omega} [u(c_0) + J_W(\omega(\mu - \lambda k - r)W + rW - C) + \frac{1}{2}\omega^2\sigma^2 J_{WW} + \lambda(E[J(\omega(Y-1)W, 0)] - J(W, 0))] \quad (10)$$

c. Applying first order condition to equation (10), the following equation holds.

$$\begin{aligned} u_c &= J_W \\ J_W(\mu - \lambda k - r)W + \omega\sigma^2 J_{WW} + \frac{\partial}{\partial \omega}(\lambda(E[J(\omega(Y-1)W, 0)] - J(W, 0))) &= 0 \end{aligned} \quad (11)$$