

# **FIN 591: Homework #3**

Due on Wednesday, April 11, 2018

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## Problem 1

- a. Since the final payoff of  $P$  is 1, using continuous-time version stochastic discount factor,  $P_t(\tau)$  is derived as follows.

$$\begin{aligned} P_t(\tau) &= E_t \left[ \frac{U_c(C_{t+\tau, t+\tau})}{U_c(C_t, t)} \times 1 \right] \\ &= E_t \left[ \frac{e^{-\phi(t+\tau)} C_{t+\tau}^{\gamma-1}}{e^{\phi t} C_t^{\gamma-1}} \right] \\ &= E_t \left[ e^{-\phi \tau} \frac{C_{t+\tau}^{\gamma-1}}{C_t^{\gamma-1}} \right] \end{aligned} \quad (1)$$

- b. From  $P_t(\tau) = E_t \left[ \frac{e^{-\phi(t+\tau)} C_{t+\tau}^{\gamma-1}}{e^{\phi t} C_t^{\gamma-1}} \right]$ , we can find that process  $M_t$  is equal to  $e^{-\phi t} C_t^{\gamma-1}$ . Therefore, using Ito's lemma, dynamics of  $M_t$  can be derived as equation (2).

$$\begin{aligned} dM_t &= -\phi e^{-\phi t} C_t^{\gamma-1} dt + e^{-\phi t} (\gamma-1) C_t^{\gamma-2} C [(\mu_c - \lambda k) dt + \sigma_c dZ_c] \\ &\quad + \frac{1}{2} e^{-\phi t} (\gamma-1)(\gamma-2) C^2 C_t^{\gamma-3} \sigma_c^2 dt + [e^{-\phi t} (Y C)^{\gamma-1} - e^{-\phi t} C^{\gamma-1}] dq \\ &= [-\phi + (\gamma-1)(\mu_c - \lambda k) + \frac{1}{2} (\gamma-1)(\gamma-2) \sigma_c^2] M_t dt + (\gamma-1) \sigma_c M_t dZ_c + (Y^{\gamma-1} - 1) M_t dq \end{aligned} \quad (2)$$

- c. Since  $E \left[ \frac{dM}{M} \right] = -r dt$ , the following equation holds.

$$\begin{aligned} r &= -E \left[ \frac{dM}{M} \right] \\ &= \phi - (\gamma-1)(\mu_c - \lambda k) - \frac{1}{2} (\gamma-1)(\gamma-2) \sigma_c^2 - \lambda E[e^{(\gamma-1) \log Y} - 1] \\ &= \phi - (\gamma-1)(\mu_c - \lambda k) - \frac{1}{2} (\gamma-1)(\gamma-2) \sigma_c^2 - \lambda (e^{(\gamma-1)\alpha + \frac{1}{2}(\gamma-1)^2 \delta^2} - 1) \end{aligned} \quad (3)$$

Since  $\mu_c, k, \lambda$  are constant, instantaneous risk free rate is constant.

- d. Since an asset price is equal to sum of discounted future payoffs, assuming some regularity conditions hold,  $S_t$  is represented as follows.

$$\begin{aligned} S_t &= E_t \left[ \int_t^\infty \frac{M_s}{M_t} D_s ds \right] \\ &= E_t \left[ \int_t^\infty \frac{e^{-\phi s} C_s^{\gamma-1}}{e^{-\phi t} C_t^{\gamma-1}} D_s ds \right] \\ \Rightarrow \frac{S_t}{D_t} &= E_t \left[ \int_t^\infty e^{-\phi(s-t)} \left( \frac{C_s}{C_t} \right)^{\gamma-1} \left( \frac{D_s}{D_t} \right) ds \right] \\ &= E_t \left[ \int_t^\infty e^{-\phi(s-t) + (\gamma-1) \log(C_s/C_t) + \log(D_s/D_t)} ds \right] \\ &= \int_t^\infty E_t \left[ e^{-\phi(s-t) + (\gamma-1) \log(C_s/C_t) + \log(D_s/D_t)} ds \right] \end{aligned} \quad (4)$$

Considering the process of  $C_t$ ,  $E_t[e^{(\gamma-1) \log(C_s/C_t)}]$  is calculated as follows.

$$\begin{aligned} E_t[e^{(\gamma-1) \log(C_s/C_t)}] &= E_t \left[ e^{(\gamma-1)(\mu_c - \frac{1}{2} \sigma_c^2 - \lambda k)(s-t) + \frac{1}{2} (\gamma-1)^2 \sigma_c^2 (s-t) + (\gamma-1) \log y(s,t)} \right] \\ y(s, t) &= \prod_{i=s}^t Y_i \end{aligned} \quad (5)$$

Since  $\log y(s, t) = \sum_{i=s}^t \log Y_i$ , and  $\log Y_i$ 's are independently and identically distributed as  $N(\alpha, \delta^2)$ ,  $\log(C_s/C_t)$  is also normally distributed, and its expected value from equation (5) is calculated as follows.

$$\mathbb{E}_t[e^{(\gamma-1)\log(C_s/C_t)}] = e^{(\gamma-1)(\mu_c - \frac{1}{2}\sigma_c^2 - \lambda k)(s-t) + \frac{1}{2}(\gamma-1)^2\sigma_c^2(s-t) + (\gamma-1)\alpha(s-t) + \frac{1}{2}(\gamma-1)^2\delta^2(s-t)} \quad (6)$$

Applying the result from equation (6) and considering the correlation between  $z_d$  and  $z_c$  is  $\rho$ ,  $\frac{S_t}{D_t}$  from equation (4) is solved as follows.

$$\begin{aligned} \frac{S_t}{D_t} &= \int_t^\infty e^{-(s-t)[\phi + (1-\gamma)(\mu_c - \frac{1}{2}\sigma_c^2 - \lambda k) + \frac{1}{2}(1-\gamma)^2\sigma_c^2 + (1-\gamma)\alpha + \frac{1}{2}(1-\gamma)^2\delta^2 - \mu_d + (1-\gamma)\rho\sigma_c\sigma_d]} ds \\ &= -\frac{1}{A} e^{-(s-t)A} \Big|_t^\infty = \frac{1}{A} \\ A &= \phi + (1-\gamma)(\mu_c - \frac{1}{2}\sigma_c^2 - \lambda k) + \frac{1}{2}(1-\gamma)^2\sigma_c^2 + (1-\gamma)\alpha + \frac{1}{2}(1-\gamma)^2\delta^2 - \mu_d + (1-\gamma)\rho\sigma_c\sigma_d \end{aligned} \quad (7)$$

## Problem 2

a. Considering the process of risky asset price, intertemporal budget constraint is derived as follows.

$$\begin{aligned} dW &= \omega_t \frac{dS}{S} + (1 - \omega_t)rdt - C_t dt \\ &= (\omega_t(\mu - \lambda k - r)W_t + rW_t - C_t)dt + \sigma\omega_t W_t dz + \omega_t W_t (Y_t - 1)dq \end{aligned} \quad (8)$$

b. Investors maximize  $\mathbb{E}_0[\int_0^T e^{-\phi t} u(C_t)dt]$ , subject to the equation (8).

Let  $J(W_t, t) = \max_{C_t, \omega_t} \mathbb{E}_t[\int_t^T e^{-\phi s} u(C_s)ds]$ . Then the following equation follows.

$$\begin{aligned} J(W_t, t) &= \max_{C_t, \omega_t} \mathbb{E}_t \left[ \int_t^{t+\Delta t} e^{-\phi s} u(C_s)ds + J(W_{t+\Delta t}, t + \Delta t) \right] \\ &= \max_{C_t, \omega_t} \mathbb{E}_t [u(C_t)\Delta t + J(W_t, t) + J_t\Delta t + J_W(\omega_t(\mu - \lambda k - r)W_t + rW_t - C_t)\Delta t \\ &\quad + \frac{1}{2}\omega_t^2 W_t^2 \sigma^2 J_{WW}\Delta t + (J(1 + \omega_t W_t(Y_t - 1), t) - J(W_t, t))dq], \quad (t \in [0, T]) \end{aligned} \quad (9)$$

Letting  $\Delta t \rightarrow 0$ , equation (9) becomes equation (10), and it is the Bellman equation.

$$\begin{aligned} 0 &= \max_{C, \omega} [u(C_t) + L(J)] \\ L(J) &= J_t + J_W(\omega_t W_t(\mu - \lambda k - r) + rW_t - C_t) + \frac{1}{2}\omega_t^2 W_t^2 \sigma^2 J_{WW} + \lambda \mathbb{E}_t [J(1 + \omega_t W_t(Y_t - 1), t) - J(W_t, t)] \end{aligned} \quad (10)$$

c. Applying first order condition to equation (10), the following equation holds.

$$\begin{aligned} u_C &= J_W \\ J_W W_t(\mu - \lambda k - r) + \omega_t W_t^2 \sigma^2 J_{WW} + \lambda \mathbb{E}_t [W_t(Y_t - 1)J_W(1 + \omega_t W_t(Y_t - 1), t)] &= 0 \end{aligned} \quad (11)$$