Financial Econometrics Econ 40357 ARIMA (Auto Regressive Integrated Moving Average) Models Part 1.

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Overview: Univariate, parametric time-series models

- Time series are interesting when there is dependence over time.
- Strategy is to develop **model** that describes the time-series data.
 - Familiar story in econometrics. Can characterize properties of theoretical model.
 - If data conforms to model, use properties of model to generate infererence, say something about the real world.
- Time-series models: Describe how the current state depends on past states. Then use the current state to predict future states.
- Estimation and prediction. Prediction ⇔ forecast.

What we learn in this segment

- Parametric models, often found to be useful for modeling stationary time series. So called ARIMA models.
- AutoRegressive Integrated Moving Average
- Key features are knowing about
 - First two moments (mean and variance)
 - Autocovariance, autocorrelation: Characterizing dependence over time
 - Conditional expectation. To be used as forecasting model.
 - Estimation, using estimated models to forecast
 - How to evaluate the forecasts.

Covariance stationarity (again)

The time series $\{y_t\}_{t=1}^T$ is covariance (weakly) stationary if the mean, variance, and autocovariances of the process are constant.

$$E(y_t) = \mu$$

$$E(y_t - \mu)^2 = \sigma^2$$

$$E(y_t - \mu)(y_{t-k} - \mu) = \gamma_k$$

Conditional expectation

Q: What function *F* minimizes the mean square prediction error,

$$E[y_{t+1} - F(y_{t+1}|I_t)]^2$$

A:

$$E(y_{t+1}|I_t) = \int y_{t+1} \rho(y_{t+1}|I_t) dy_{t+1}$$

where $p(y_{t+1}|I_t)$ is conditional pdf of y_{t+1} , I_t is the available information at t.

- Important result: Conditional expectation is minimum mean-square error predictor. It's the best!
- Think of fitted value of regression as conditional expectation.
 Systematic part of regression also called projection.
- Notational Convention $E_t(X_{t+k}) \equiv E(X_{t+k}|I_t)$

The white noise process (again)

- Stochastic (random) nature of the world
- White noise is the basic building block of all time series

$$y_t = \sigma \epsilon_t, \quad \epsilon_t \stackrel{iid}{\sim} (0, 1)$$

- These are random shocks, no dependence over time, representing purely unpredictable events. It's a model of news.
- We didn't say they are normally distributed. In time-series, it doesn't matter because all inference is asymptotic
- By itself, it is uninteresting, because there is no dependence over time.
- Next, I show you how we build in dependence.

Moving Average models

- An MA(k) process. y_t is correlated with y_{t-k} and possibly $y_{t-1}, ..., y_{t-k+1}$.
- The MA(1). Example might be daily returns with slow moving capital. News occurs today. High frequency traders pounce, institutional investors, move later in the day. Retail investors don't know until they see the nightly Bloomberg report.

The MA(1) model

• Let y_t be the observations

$$y_t = \mu + \epsilon_t + \theta \epsilon_{t-1}$$

where $\epsilon_t \stackrel{iid}{\sim} (0, \sigma_{\epsilon}^2)$. Shift time index back one period,

$$y_{t-1} = \mu + \epsilon_{t-1} + \theta \epsilon_{t-2}$$

Calculate the mean of y_t

$$E(y_t) = E(\mu + \epsilon_t + \theta \epsilon_{t-1}) = \mu$$

Calculate the variance of y_t

$$\sigma_y^2 = Var(y_t) = E(y_t - \mu)^2 = E(\varepsilon_t + \theta \varepsilon_{t-1})^2$$

$$= E(\varepsilon_t^2 + \theta^2 \varepsilon_{t-1}^2 + 2\theta \varepsilon_t \varepsilon_{t-1})$$
(2)

$$= E\left(\epsilon_t^2\right) + E\left(\theta^2 \epsilon_{t-1}^2\right) + E\left(2\theta \epsilon_t \epsilon_{t-1}\right) \tag{3}$$

$$= \sigma_{\epsilon}^{2} + \theta^{2} \sigma_{\epsilon}^{2} + 2\theta \underbrace{E\left(\epsilon_{t} \epsilon_{t-1}\right)}_{0} \tag{4}$$

$$= \left(1 + \theta^2\right) \sigma_{\epsilon}^2 \tag{5}$$

Calculate the auto covariance function

$$\gamma_{1} = Cov (y_{t}, y_{t-1}) = E (\tilde{y}_{t}, \tilde{y}_{t-1})
= E (\epsilon_{t} + \theta \epsilon_{t-1}) (\epsilon_{t-1} + \theta \epsilon_{t-2})
= E (\epsilon_{t} \epsilon_{t-1} + \theta \epsilon_{t-1}^{2} + \theta \epsilon_{t} \epsilon_{t-2} + \theta^{2} \epsilon_{t-1} \epsilon_{t-2})
= \theta \sigma_{\epsilon}^{2}$$
(7)

Autocorrelation

$$\rho\left(y_{t}, y_{t-1}\right) = \operatorname{Corr}(y_{t}, y_{t-1}) = \frac{\gamma_{1}}{\sigma_{y}\sigma_{y}} = \frac{\gamma_{1}}{\left(1 + \theta\right)\sigma_{\varepsilon}^{2}}$$

and for any k > 1,

$$\gamma_k = Cov(y_t, y_{t-k}) = 0.$$

MA(1) process is covariance stationary and displays one period dependence (memory).

MA(1) Forecasting formula

 Use the fact that the conditional expectation (projection), the fitted value of model (regression), is the optimal forecast. One period ahead forecast

$$E_{t}(y_{t+1}) = E_{t}(\mu + \epsilon_{t+1} + \theta \epsilon_{t}) = \mu + \theta \epsilon_{t}$$
$$E_{t}(y_{t+2}) = E_{t}(\mu + \epsilon_{t+1} + \theta \epsilon_{t+1}) = \mu$$

And for any $k \ge 2$, the model has no forecasting power

$$E_t(y_{t+k}) = \mu$$

The MA(2) model

 Observations correlated with (exhibit dependence) at most 2 lags of itself.

$$y_t = \mu + \epsilon_t + \theta_1 \epsilon_{t-1} + \theta_2 \epsilon_{t-2}$$

I assign you to verify the following

$$E(y_t) = \mu$$

$$Var(y_t) = \left(1 + \theta_1^2 + \theta_2^2\right) \sigma_{\epsilon}^2,$$

$$Cov(y_t, y_{t-1}) = (\theta_1 + \theta_1\theta_2) \sigma_{\epsilon}^2$$

$$Cov(y_t, y_{t-2}) = \theta_2,$$

$$Cov(y_t, y_{t-k}) = 0 \text{ for } k > 2$$

- MA(2) Forecasts
- One-step ahead forecast

$$E_t(y_{t+1}) = E_t(\mu + \epsilon_{t+1} + \theta_1 \epsilon_t + \theta_2 \epsilon_{t-1})$$

= $\mu + \theta_1 \epsilon_t + \theta_2 \epsilon_{t-1}$

Two-step ahead forecast

$$E_{t}(y_{t+2}) = E_{t}(\mu + \epsilon_{t+2} + \theta_{1}\epsilon_{t+1} + \theta_{2}\epsilon_{t}) = \mu + \theta_{2}\epsilon_{t}$$

Three-step ahead forecast

$$E_t(y_{t+3}) = E(\mu + \epsilon_{t+3} + \theta_1 \epsilon_{t+2} + \theta_2 \epsilon_{t+1}) = \mu$$

Hence for any $k \geq 3$, $E_t(y_{t+k}) = \mu$.

How to Estimate MA models?

- There are no independent variables, so you can't run least squares regression.
- We do something called maximum likelihood estimation.
- Illustrate the idea with the MA(1) model.

Maximum Likelihood Estimation of MA(1)

• The ϵ_t are random variables. Let's assume they are drawn from a normal distribution, $N\left(0,\sigma_{\epsilon}^2\right)$. The marginal probability density function (pdf) for ϵ_t is

$$f_1\left(\epsilon_t\right) = rac{1}{\sigma_\epsilon \sqrt{2\pi}} e^{-rac{\epsilon_t^2}{2\sigma_\epsilon^2}}$$

The joint pdf for $\epsilon_1, \epsilon_2, ..., \epsilon_t, \epsilon_{t+1}, ..., \epsilon_T$ is the product of the f_1 (), because the $\epsilon's$ are independent.

$$f(\epsilon_T, \epsilon_{T-1}, ..., \epsilon_1) = \left(\frac{2}{\sigma_{\epsilon} \sqrt{2\pi}}\right)^T e^{-\frac{1}{2\sigma_{\epsilon}^2} \sum_{l=1}^T \epsilon_t^2}$$

Maximum Likelihood Estimation of MA(1)

• Notice that $\epsilon_t = y_t - \mu - \theta \epsilon_{t-1}$, $\epsilon_{t-1} = y_{t-1} - \mu - \theta \epsilon_{t-2}$, $\epsilon_{t-2} = y_{t-2} - \mu - \theta \epsilon_{t-3}$, ... This means $\epsilon_t = y_t - \mu - \theta \left(y_{t-1} - \mu - \theta \left(y_{t-2} - \mu - \theta \left(\dots \right) \right) \right)$ $\epsilon_{t-1} = y_{t-1} - \mu - \theta \left(y_{t-2} - \mu - \theta \left(y_{t-3} - \mu - \theta \left(\dots \right) \right) \right)$ \vdots

 Substitute these back into the joint pdf, and we get a function of the y_t, which I won't write out specifically.

$$f\left(y_T,y_{T-1},...,y_1\big|\mu,\theta,\sigma_\epsilon^2\right)$$

This is now a function of the **data**. By substitution of the MA(1) model into the joint pdf, we've transformed the pdf into a function of the data. This is called a **likelihood** function. PDFs are for random variables. Likelihood functions are for data.

• Maximum likelihood estimation is done by asking the computer to search and those $\mu, \eta, \sigma_{\epsilon}^2$ that maximizes f().

Let's apply MA(1) to daily stock returns

Eviews/ARIMA Models.wf1

Code: equation eqma1.ls(optmethod=opg) djiaret c ma(1)

Dependent Variable: DJIARET

Method: ARMA Maximum Likelihood (OPG - BHHH)

Date: 09/12/19 Time: 10:50 Sample: 8/25/2014 8/22/2019 Included observations: 1208

Convergence 28 iterations

Variable C

MA(1) SIGMASQ R-squared Adjusted R-squared S.E. of regression

Sum squared resid Log likelihood F-statistic

Prob(F-statistic)
Inverted MA Roots

Coefficient 0.000385 -0.008067 7.25E-05 0.000058 -0.001601 0.008523 0.087529 4043.555 0.035126 0.965485 .01

Std. Error
0.000260
0.019399
1.90E-06
Mean dependent var
S.D. dependent var
Akaike info criterion
Schwarz criterion
Hannan-Quinn criter.
Durbin-Watson stat

Prob. 0.1390 0.6776 0.0000 0.000384 0.008516 -6.689660 -6.677003 -6.684894 1.965005

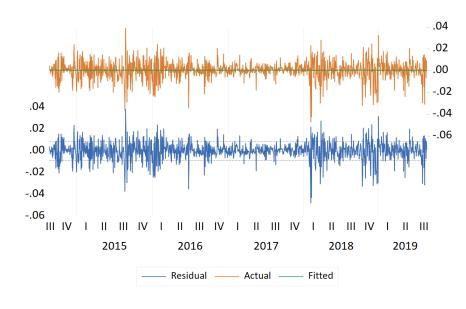
t-Statistic

1.480358

-0.415869

38 04024

Let's apply MA(1) to daily stock returns



Let's apply MA(5) to daily stock returns

equation eqma5.ls(optmethod=opg) dijaret c ma(1) ma(2) ma(3) ma(4) ma(5)

Dependent Variable: DJIARET

Method: ARMA Maximum Likelihood (OPG - BHHH)

Date: 09/12/19 Time: 10:50 Sample: 8/25/2014 8/22/2019 Included observations: 1208

Convergence achieved after 36 iterations

Variable C MA(1) MA(2) MA(3) MA(4) MA(5) SIGMASO R-squared Adjusted R-squared S.E. of regression Sum squared resid Log likelihood F-statistic Prob(F-statistic) Inverted MA Roots

Coefficient
0.000378
-0.005196
-0.037908
0.067974
-0.019959
-0.050492
7.18E-05
0.008933
0.003981
0.008499
0.086752
4048.525
1.804113
0.094952
.54
46+.25i

Std. Error 0.000266 0.019996 0.021856 0.022404 0.022182 0.025925 1.94E-06 Mean dependent var S.D. dependent var Schwarz criterion Schwarz criterion Hannan-Quinn criter. Durbin-Watson stat	t-Statistic 1.421888 -0.259843 -1.734453 3.331478 -0.899776 -1.947578 36.98844
.1955i	.19+.55i

0.1553 0.7950 0.0831 0.0009 0.3684 0.0517 0.0000 0.000384 0.008516 -6.691267

-6.661733

-6.680146

1 971303

-.46-.25i

Proh

Let's apply MA(5) to daily stock returns

