

$$1. \infty \quad E[X] = \frac{1+2}{2} = \frac{3}{2}$$

$$E[Y] = E[E[Y|X=x]] = E\left[\frac{1}{\sqrt{x}}\right] = E[X] = \frac{3}{2}$$

$$E[Z] = E[E[Z|X=x, Y=y]] = E[X] = \frac{3}{2}$$

$$\mu = \begin{bmatrix} \frac{3}{2} \\ \frac{3}{2} \\ \frac{3}{2} \end{bmatrix}$$

$$V[X] = \frac{(2-0)^2}{12} = \frac{1}{12}$$

$$V[Y] = E[V[Y|X]] + V[E[Y|X]]$$

$$= E\left[\left(\frac{1}{\sqrt{x}}\right)^2\right] + V[X] = E[X^{-1}] + V[X]$$

$$= \int_1^2 x^{-2} f_X(x) dx + V[X]$$

$$= \int_1^2 x^{-2} dx + \frac{1}{12}$$

$$= \left. \frac{x^{-1}}{-1} \right|_1^2 + \frac{1}{12}$$

$$= \frac{8}{3} - \frac{1}{3} + \frac{1}{12} = \frac{7}{3} + \frac{1}{12} = \frac{29}{12}$$

$$V[Z] = E[V[Z|X, Y]] + V[E[Z|X, Y]]$$

$$= E[V[Z|X]] + V[E[Z|X]]$$

$$= E[1] + V[X]$$

$$= \frac{13}{12}$$

$$\text{cov}(X, Y) = E[XY] - E[X]E[Y]$$

$$= E[E[XY|X=x]] - E[X]E[Y]$$

$$= E[X^{-1}] - E[X]E[Y]$$

$$= \frac{7}{3} - \frac{3}{2} \cdot \frac{3}{2} = \frac{7}{3} - \frac{9}{4} = \frac{1}{12}$$

$$\text{cov}(X, Z) = E[XZ] - E[X]E[Z]$$

$$= E[E[XZ|X=x]] - E[X]E[Z]$$

$$= E[X^2] - E[X]E[Z]$$

$$= \frac{7}{3} - \frac{3}{2} \cdot \frac{3}{2} = \frac{1}{12}$$

$$\text{cov}(Y, Z) = E[YZ] - E[Y]E[Z]$$

$$= E[E[YZ|X=x]] - E[Y]E[Z]$$

$$= \frac{7}{3} - \frac{3}{2} \cdot \frac{3}{2} = \frac{1}{12}$$

$$\Sigma = \begin{bmatrix} V[X] & \text{cov}(X,Y) & \text{cov}(X,Z) \\ \text{cov}(Y,X) & V[Y] & \text{cov}(Y,Z) \\ \text{cov}(Z,X) & \text{cov}(Z,Y) & V[Z] \end{bmatrix} = \begin{bmatrix} 1/2 & 1/2 & 1/2 \\ 1/2 & 29/12 & 1/2 \\ 1/2 & 1/2 & 13/12 \end{bmatrix}$$

b) See code

```
sigma = [1/12 1/12 1/12; 1/12 29/12 1/12; 1/12 1/12 13/12];  
[V, D] = eig(sigma);  
Q_t = transpose(V);  
disp(Q_t);
```

```
0.9963    -0.0326    -0.0796  
0.0772    -0.0670     0.9948  
0.0378     0.9972     0.0643
```

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$$2. \quad \gamma_1 = \sqrt{-2 \log X_1} \cos(2\pi X_2)$$

$$\gamma_2 = \sqrt{-2 \log X_1} \sin(2\pi X_2)$$

$$\gamma = (\gamma_1, \gamma_2)^T$$

$$f_\gamma(y) = f_X(H(y)) |\det J_H(y)|$$

$$f_X(x) = f_{X_1}(x_1) f_{X_2}(x_2) = \frac{1}{1-0} \cdot \frac{1}{1-0} = 1$$

$$f_\gamma(y) = |\det J_H(y)|$$

$$= \begin{vmatrix} \frac{\partial h_1}{\partial \gamma_1} & \frac{\partial h_1}{\partial \gamma_2} \\ \frac{\partial h_2}{\partial \gamma_1} & \frac{\partial h_2}{\partial \gamma_2} \end{vmatrix}$$

$$\gamma_1^2 + \gamma_2^2 = (-2 \log X_1) \cos^2(2\pi X_2) + (-2 \log X_1) \sin^2(2\pi X_2)$$

$$= (-2 \log X_1) (\cos^2(2\pi X_2) + \sin^2(2\pi X_2))$$

$$= -2 \log X_1$$

$$-\frac{\gamma_1^2 + \gamma_2^2}{2} = \log X_1 \Rightarrow X_1 = e^{-\frac{\gamma_1^2 + \gamma_2^2}{2}}$$

$$\frac{\gamma_2}{\gamma_1} = \frac{\sin(2\pi X_2)}{\cos(2\pi X_2)} = \tan(2\pi X_2)$$

$$X_2 = \frac{1}{2\pi} \tan^{-1}(\gamma_2/\gamma_1)$$

$$\text{let } h_1 = X_1, h_2 = X_2$$

$$\frac{\partial h_1}{\partial \gamma_1} = e^{-\frac{\gamma_1^2 + \gamma_2^2}{2}} (-\gamma_1)$$

$$\frac{\partial h_1}{\partial \gamma_2} = e^{-\frac{\gamma_1^2 + \gamma_2^2}{2}} (-\gamma_2)$$

$$\frac{\partial h_2}{\partial \gamma_1} = \frac{1}{2\pi} \frac{1}{1 + (\gamma_2/\gamma_1)^2} \gamma_2 (-\gamma_1^{-2}) = \frac{1}{2\pi} \frac{-\gamma_2}{(1 + \gamma_2^2/\gamma_1^2) \gamma_1^2} = -\frac{\gamma_2}{2\pi (\gamma_1^2 + \gamma_2^2)}$$

$$\frac{\partial h_2}{\partial \gamma_2} = \frac{1}{2\pi} \frac{1}{1 + (\gamma_2/\gamma_1)^2} \frac{1}{\gamma_1} = \frac{1}{2\pi} \frac{1}{1 + \gamma_2^2/\gamma_1^2} \cdot \frac{1}{\gamma_1} = \frac{1}{2\pi} \frac{\gamma_1^2}{\gamma_1^2 + \gamma_2^2} \frac{1}{\gamma_1}$$

$$= \frac{\gamma_1}{2\pi (\gamma_1^2 + \gamma_2^2)}$$

$$f_\gamma(y) = \left| -\gamma_1 e^{-\frac{\gamma_1^2 + \gamma_2^2}{2}} \frac{\gamma_1}{2\pi (\gamma_1^2 + \gamma_2^2)} - (-\gamma_2) e^{-\frac{\gamma_1^2 + \gamma_2^2}{2}} \frac{-\gamma_2}{2\pi (\gamma_1^2 + \gamma_2^2)} \right|$$

$$= \left| \frac{e^{-\frac{\gamma_1^2 + \gamma_2^2}{2}}}{2\pi (\gamma_1^2 + \gamma_2^2)} (-\gamma_1^2 - \gamma_2^2) \right| = \left| -\frac{1}{2\pi} e^{-\frac{\gamma_1^2 + \gamma_2^2}{2}} \right|$$

$$= \frac{1}{2\pi} e^{-\frac{\gamma_1^2 + \gamma_2^2}{2}}$$

$$f_{Y_1}(y) = \int_{-\infty}^{\infty} f_Y(y) dY_2 = \int_{-\infty}^{\infty} \frac{1}{2\pi} e^{-\frac{Y_1^2 + Y_2^2}{2}} dY_2 = \frac{1}{\sqrt{2\pi}} e^{-\frac{Y_1^2}{2}}$$

$$f_{Y_2}(y) = \int_{-\infty}^{\infty} f_Y(y) dY_1 = \int_{-\infty}^{\infty} \frac{1}{2\pi} e^{-\frac{Y_1^2 + Y_2^2}{2}} dY_1 = \frac{1}{\sqrt{2\pi}} e^{-\frac{Y_2^2}{2}}$$

$$f_Y(y) = \frac{1}{2\pi} e^{-\frac{Y_1^2 + Y_2^2}{2}} = \frac{1}{\sqrt{2\pi}} e^{-\frac{Y_1^2}{2}} \frac{1}{\sqrt{2\pi}} e^{-\frac{Y_2^2}{2}} = f_{Y_1}(y) f_{Y_2}(y)$$

Therefore, Y_1 and Y_2 are independent

$$b) f_{Y_1}(y) = \int_{-\infty}^{\infty} f_Y(y) dY_2 = \int_{-\infty}^{\infty} \frac{1}{2\pi} e^{-\frac{Y_1^2 + Y_2^2}{2}} dY_2 = \frac{1}{\sqrt{2\pi}} e^{-\frac{Y_1^2}{2}}$$

$$f_{Y_2}(y) = \int_{-\infty}^{\infty} f_Y(y) dY_1 = \int_{-\infty}^{\infty} \frac{1}{2\pi} e^{-\frac{Y_1^2 + Y_2^2}{2}} dY_1 = \frac{1}{\sqrt{2\pi}} e^{-\frac{Y_2^2}{2}}$$

$Y_1 \sim N(0, 1)$ and $Y_2 \sim N(0, 1)$

$$3. a) g(\gamma) = \sum_{x,r} \sum_r^{-1} (\gamma - \mu_r) + \mu_x$$

$$\begin{aligned} \sum_{x,r} &= \sum_x Gx + w = \sum_x Gx + \sum_x w = \sum_x Gx \\ &= E[X(GX)^T] - E[X]E[GX]^T \\ &= (E[XX^T] - E[X]E[X]^T)G^T \\ &= \sum_x G^T \end{aligned}$$

$$\sum_r = \sum_{GX+w} = \sum_{GX} + \sum_{GX,w} + \sum_{w,GX} + \sum_w$$

$$\sum_{GX} = G \sum_x G^T$$

$$\begin{aligned} \sum_{GX,w} &= E[GXW^T] - E[GX]E[W]^T \\ &= (E[XW^T] - E[X]E[W]^T)G \\ &= \sum_{xw} G = 0 \end{aligned}$$

$$\sum_{w,GX} = \sum_{GX,w}^T = 0$$

$$\sum_r = G \sum_x G^T + \sum_w$$

$$\begin{aligned} \mu_r &= E[r] = E[GX + w] \\ &= E[GX] = G E[X] \\ &= G \mu_x \end{aligned}$$

$$g(\gamma) = \sum_x G^T (G \sum_x G^T + \sum_w)^{-1} (\gamma - G \mu_x) + \mu_x$$

b) See code

```
n = 10;
G = [1 2; 3, 4];
G_T = G';
e = 0.03;
X = zeros(0, 2);
X_hat = zeros(0, 2);

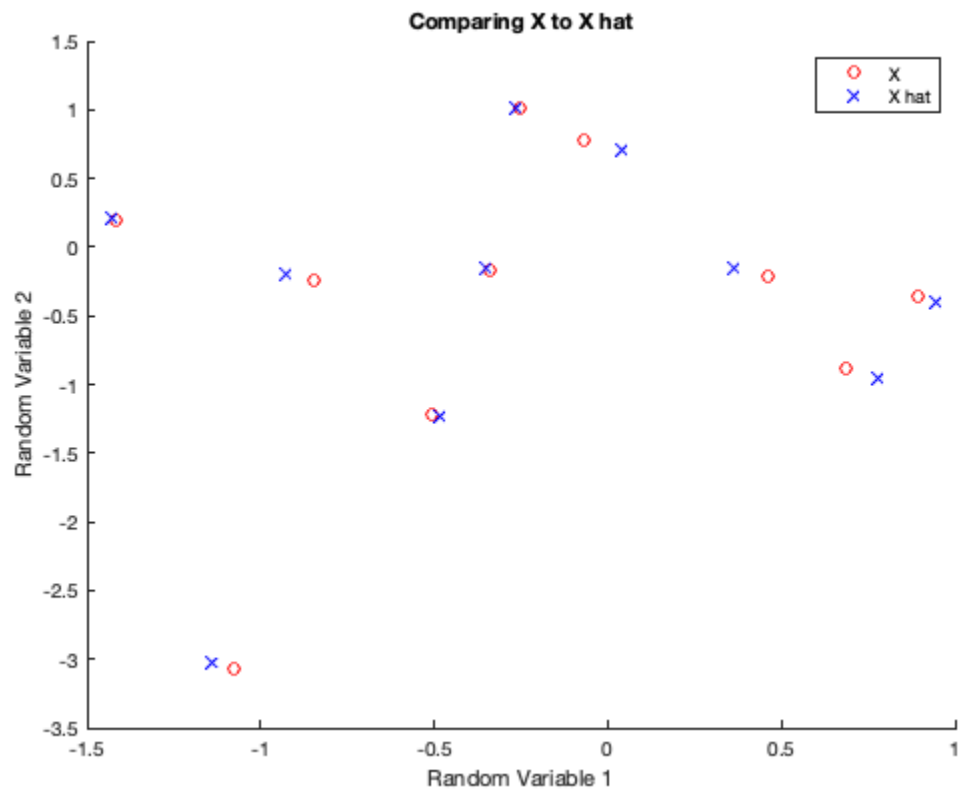
sigma_X = [1 0; 0 1];
sigma_W = [e^2 0; 0 e^2];
miu_X = [0; 0];

for i = 1:n
    x = normrnd(0, 1, 2, 1);
    w = mvnrnd([0; 0], sigma_W)';
    y = G * x + w;

    x_hat = sigma_X * G_T * (G * sigma_X * G_T + sigma_W)^(-1) * (y - G *
miu_X) + miu_X;

    X_hat = [X_hat; x_hat'];
    X = [X; x'];
end

hold on
scatter(X(:, 1), X(:, 2), 'red');
scatter(X_hat(:, 1), X_hat(:, 2), 'blue', 'x');
legend('X', 'X hat');
title('Comparing X to X hat');
xlabel('Random Variable 1');
ylabel('Random Variable 2');
hold off
```



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$$4. a) X = \begin{bmatrix} x_1 \\ x_2 \end{bmatrix} = \begin{bmatrix} x_1 \\ 3x_1 \end{bmatrix} = \begin{bmatrix} 1 \\ 3 \end{bmatrix} [x_1] + \begin{bmatrix} 0 \\ 0 \end{bmatrix}$$

Since X can be expressed as $X = AZ + \mu$ where $Z \sim N(0, I)$,

X is Gaussian since its components are jointly distributed

$$b) \text{cov}(X) = CC^T = \begin{bmatrix} 1 \\ 3 \end{bmatrix} \begin{bmatrix} 1 & 3 \end{bmatrix} = \begin{bmatrix} 1 & 3 \\ 3 & 9 \end{bmatrix}$$

$$c) \det(\Sigma_X) = 9 - 3 \cdot 3 = 0$$

Σ_X is singular, so X does not have a density

5. a) If X is Gaussian, it can be expressed as

$$X = BZ + \mu, \text{ where } Z = (Z_1, \dots, Z_m)^T \text{ and } Z_1, \dots, Z_m \stackrel{\text{iid}}{\sim} N(0, 1)$$

Then,

$$Y = A(BZ + \mu) = ABZ + A\mu$$

$$\text{Let } C = AB \text{ and } \mu_Y = A\mu$$

Then,

$$Y = CZ + \mu_Y$$

Since $Z = (Z_1, \dots, Z_m)^T$ and $Z_1, \dots, Z_m \stackrel{\text{iid}}{\sim} N(0, 1)$, Y 's components are jointly normally distributed, and Y is normal, proving the claim.

b) We can disprove the claim by providing a counterexample:

Let X be a non-Gaussian vector:

$$X = \begin{bmatrix} X_1 \\ X_2 \end{bmatrix} \text{ where } X_1 \sim N(0, 1) \text{ and } X_2 \not\sim N(0, 1)$$

$$\text{Let } A = \begin{bmatrix} 1 & 0 \\ 1 & 0 \end{bmatrix}.$$

Then,

$$Y = AX = \begin{bmatrix} 1 & 0 \\ 1 & 0 \end{bmatrix} \begin{bmatrix} X_1 \\ X_2 \end{bmatrix} = \begin{bmatrix} X_1 \\ X_1 \end{bmatrix}$$

Since Y 's components follow $\sim N(0, 1)$, they are jointly normally distributed, so Y is Gaussian. However, X is not Gaussian, so the claim is disproved.