

$$1. \quad f(x, y) = \frac{y^2 - x^2}{8} e^{-y} \quad 0 < y < \infty, \quad -y \leq x \leq y$$

$$\begin{aligned}
 E[X|Y=y] &= \int_{-y}^y x f_{X|Y}(x|y) dx \\
 &= \int_{-y}^y x \frac{f_{X,Y}(x,y)}{f_Y(y)} dx = \int_{-y}^y x \frac{y^2 - x^2}{8f_Y(y)} e^{-y} dx \\
 &= \frac{e^{-y}}{8f_Y(y)} \int_{-y}^y x(y^2 - x^2) dx \\
 &= \frac{e^{-y}}{8f_Y(y)} \left[\frac{x^2 y^2}{2} - \frac{x^4}{4} \right]_{-y}^y \\
 &= \frac{e^{-y}}{8f_Y(y)} \left[\left(\frac{y^4}{2} - \frac{y^4}{4} \right) - \left(\frac{y^4}{2} - \frac{y^4}{4} \right) \right] \\
 &= 0
 \end{aligned}$$

2. a) Let A be the event that a student has influenza.

The indicator of A is $I_A = \begin{cases} 1 & \text{if student has influenza} \\ 0 & \text{if student doesn't have influenza} \end{cases}$

We want to find $f_{X|I_A}(x | I_A = 1)$

$$f_{X|I_A}(x | I_A = 1) = P(X=x | I_A=1) = \frac{P(I_A=1 | X=x) P(X=x)}{P(I_A=1)}$$

Denominator:

$$\begin{aligned} P(I_A=1) &= P(A) = \int_0^1 q(x) f_X(x) dx \\ &\propto \int_0^1 x^\gamma x^{\alpha-1} (1-x)^{\beta-1} dx = \int_0^1 x^{\gamma+\alpha-1} (1-x)^{\beta-1} dx \end{aligned}$$

Numerator:

$$\begin{aligned} P(I_A=1 | X=x) P(X=x) &= q(x) P(X=x) \\ &\propto q(x) f_X(x) \\ &\propto x^\gamma x^{\alpha-1} (1-x)^{\beta-1} = x^{\gamma+\alpha-1} (1-x)^{\beta-1} \\ f_{X|I_A}(x | I_A=1) &= \frac{x^{\gamma+\alpha-1} (1-x)^{\beta-1}}{\int_0^1 x^{\gamma+\alpha-1} (1-x)^{\beta-1} dx} \\ &= \frac{x^{\gamma+\alpha-1} (1-x)^{\beta-1}}{B(\gamma+\alpha, \beta)} \end{aligned}$$

The conditional density follows a Beta distribution: $\text{Be}(\gamma+\alpha, \beta)$

b) Using the expected value of Beta distribution on $\text{Be}(\gamma+\alpha, \beta)$:

$$E[X | \text{student has influenza}] = \frac{\gamma+\alpha}{\gamma+\alpha+\beta}$$

c) See code

```
n = 10^4;
alpha = 2;
beta = 6;
gamma = 2;

X = betarnd(alpha, beta, n, 1);
exposure = 0;
counter = 0;
for i=1:n
    x = X(i);
    q = x^gamma;
    if (rand() < q)
        exposure = exposure + x;
        counter = counter + 1;
    end
end

sample = exposure / counter
expected = (gamma + alpha) / (gamma + alpha + beta)

sample =

    0.3955

expected =

    0.4000
```

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3. I guess that there will be 2 loops.

Let X denote the number of loops by the end of the process, and let n be the number of shoelaces. That means there are $2n$ free ends. Let $X = X_1 + X_2 + \dots + X_n$ where X_i is the number of loops obtained at iteration i .

On the first iteration, you have to choose 2 free ends out of the $2n$ available. After choosing the first free end, the only way to make a loop is by choosing the other end of the strand, and this happens w.p. $\frac{1}{2n-1}$ since there are $2n-1$ available strands to choose from. If you do not choose a free end on the same strand, you make no loop, which happens w.p. $\frac{2n-2}{2n-1}$. Therefore,
$$E[X_1] = \frac{1}{2n-1}(1) + \frac{2n-2}{2n-1}(0) = \frac{1}{2n-1}$$

After the first iteration, the number of free ends becomes $2n-2$ regardless of whether or not you made a loop. Therefore, for the second iteration,
$$E[X_2] = \frac{1}{2n-3}(1) + \frac{2n-4}{2n-3}(0) = \frac{1}{2n-3}$$

Generalizing this to the i th iteration,
$$E[X_i] = \frac{1}{2n-2i+1}$$

$$\begin{aligned} E[X] &= E[X_1 + X_2 + \dots + X_n] = E[X_1] + E[X_2] + \dots + E[X_n] \\ &= \sum_{i=1}^n \frac{1}{2n-2i+1} = \sum_{i=1}^n \frac{1}{2(n-i)+1} \\ &= \sum_{i=0}^{n-1} \frac{1}{2i+1} \\ &= \sum_{i=1}^n \frac{1}{2i-1} \end{aligned}$$

For $n=100$, this summation evaluates to $\sum_{i=1}^{100} \frac{1}{2i-1} = 3.284$ loops, which is quite close to my initial guess.

4. a) We claim that $E[N_k] = 1 + m + \dots + m^{k-1}$. We go by induction on k .

Base case: $k = 1$

After trial 1, no matter the outcome, we have fulfilled having 1 consecutive outcome, so $E[N_1] = 1$, which is exactly the claim for $k = 1$.

Induction Hypothesis

Suppose that $E[N_k] = 1 + m + \dots + m^{k-1}$ holds for some arbitrary $k \in \mathbb{Z}^+$.

Induction step

We want to prove that $E[N_{k+1}] = 1 + m + \dots + m^k$.

$$E[N_{k+1}] = E[E[N_{k+1} | N_k]]$$

Given N_k , there are two cases. If you select the consecutive outcome, $N_{k+1} = N_k + 1$ w.p. $\frac{1}{m}$. If you select the non-repeating outcome, $N_{k+1} = N_k + E[N_{k+1}]$ w.p. $\frac{m-1}{m}$.

$$\begin{aligned} E[E[N_{k+1} | N_k]] &= E\left[(N_k + 1) \frac{1}{m} + (N_k + E[N_{k+1}]) \frac{m-1}{m}\right] \\ &= E\left[\frac{N_k}{m} + \frac{1}{m} + \frac{N_k(m-1)}{m} + E[N_{k+1}] \frac{m-1}{m}\right] \\ &= E\left[N_k + \frac{1}{m} + E[N_{k+1}] \frac{m-1}{m}\right] \end{aligned}$$

$$E[N_{k+1}] = E[N_k] + \frac{1}{m} + E[N_{k+1}] \frac{m-1}{m}$$

$$\frac{1}{m} E[N_{k+1}] = E[N_k] + \frac{1}{m}$$

$$E[N_{k+1}] = m E[N_k] + 1$$

$$\begin{aligned} &= m(m^{k-1} + m^{k-2} + \dots + 1) + 1 \\ &= m^k + m^{k-1} + \dots + 1 \end{aligned}$$

by Induction Hypothesis

which proves our claim

$$\begin{aligned}
 b) \sigma_q &= \sqrt{V[N_q]} \approx \sqrt{E[N_q]^2} = E[N_q] \\
 &= 1 + m + \dots + m^8 \quad m = 10 \\
 &= 1 + 10 + \dots + 10^8 \\
 &= 11111111
 \end{aligned}$$

$$\begin{aligned}
 E[N_q] - n_q &= 11111111 - 24658620 \\
 &= 86452502 < \sigma_q
 \end{aligned}$$

since n_q is within one standard deviation of the mean, the evidence is weak

$$\begin{aligned}
 5. \quad E[X] &= E[E[X|Q]] \\
 &= \int_0^1 E[X|Q=q] f_Q(q) dq \\
 &= \int_0^1 E[X|Q=q] dq \\
 &= \int_0^1 nq dq = \frac{nq^2}{2} \Big|_0^1 = \frac{n}{2}
 \end{aligned}$$

$$\begin{aligned}
 V[X] &= E[V[X|Q]] + V[E[X|Q]] \\
 &= E[nq(1-q)] + V[nq] \\
 &= \int_0^1 nq - nq^2 dq + n^2 V[q] \\
 &= \frac{nq^2}{2} - \frac{nq^3}{3} \Big|_0^1 + n^2 \frac{(1-0)^2}{12} \\
 &= \frac{n}{2} - \frac{n}{3} + \frac{n^2}{12} \\
 &= \frac{n^2}{12} + \frac{n}{6} \\
 &= \frac{n^2 + 2n}{12}
 \end{aligned}$$