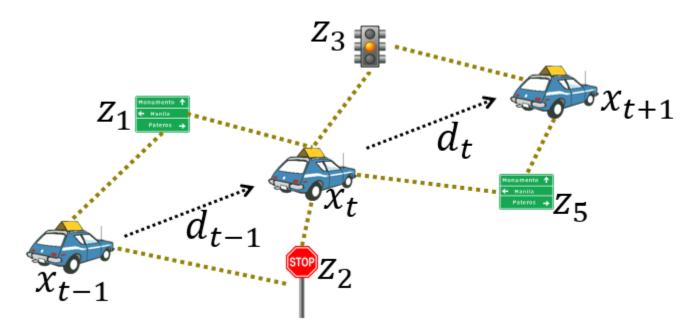
# A simple Localization Problem



#### → Given

- 1. Map: Global position of landmarks as  $z_i$
- 2. Motion model of the robot, e.g.  $x_{t+1} = x_t + d_t + \epsilon_t$  ,  $\epsilon_t \sim \mathcal{N}(0, R_t)$ .
- 3. Observations model as the bearing of the landmarks  $f_{i,t} = H(z_i x_t) + \eta_{i,t}$  ,  $\eta_{i,t} \sim \mathcal{N}(0,Q_t)$  .
- 4. Local observations at each time step for some of the landmarks  $f_{i,t}.$
- 5. Initial Location  $x_0 \sim \mathcal{N}(\mu_0, \Sigma_0)$ .

#### Find

The position and orientation of the robot,  $x_t$ 

!pip install -q requests scipy ipywidgets svgpath2mpl
from scipy.stats import multivariate\_normal
import requests
from io import BytesIO
import numpy as np
%matplotlib inline
import matplotlib.pyplot as plt
from IPython.display import display

from ipywidgets import interact, interactive, fixed, interact\_manual
import ipywidgets as widgets

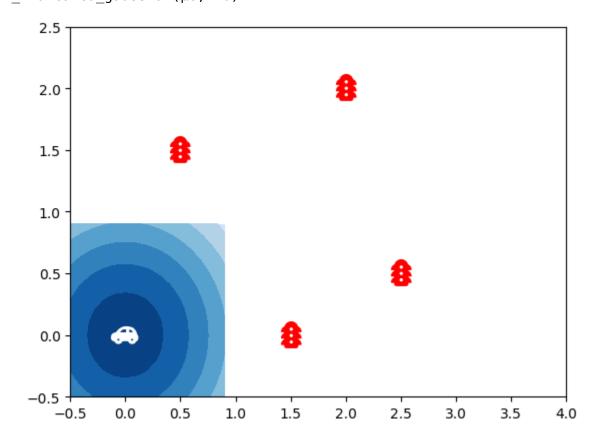
```
# 
mplcar = parse_path('M 9880,3580 C 8932,3481 8224,2669 8263,1725 8305,731 9137,-38
mplcar.vertices -= mplcar.vertices.mean(axis=0)
# 
mpltraffic = parse_path('M384 192h-64v-37.88c37.2-13.22 64-48.38 64-90.12h-64V32c0-
mpltraffic.vertices -= mpltraffic.vertices.mean(axis=0)

1.6/1.6 MB 11.8 MB/s eta 0:00:00
```

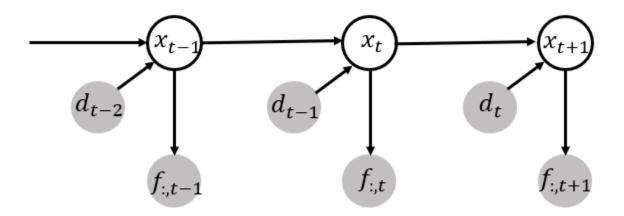
### , Input data

```
Scale Model noise by a scalar
#@title Input data
url = 'http://www-personal.umich.edu/~dh
                                           value scale_q
response = requests.get( url)
data = np.load(BytesIO(response.content)
                                             scale q:
f i t = [data['obs_vals_%d' % k] for k i
visible all t = [data['obs ids %d' % k]
d = data['d']
                                           Scale Observation noise by a
z = data['z']
                                           scalar value scale r
#@markdown ### Scale Model noise by a sc
scale q = 1 #@param {type: "slider", min
Q = data['Q'] * scale q
                                             scale r:
#@markdown ### Scale Observation noise b
scale r = 2 #@param {type: "slider", min
R = data['R'] * scale r
H = data['H']
\mu\theta = data['mu\theta']
\Sigma 0 = data['sigma0']
def plot truncated gaussian(mean, cov, s=1,n=10):
  xsigma = np.sgrt(cov[0, 0])
  ysigma = np.sqrt(cov[1, 1])
  xslice = slice(mean[0]-s*xsigma, mean[0]+s*xsigma, xsigma/n)
  yslice = slice(mean[0]-s*ysigma, mean[0]+s*ysigma, ysigma/n)
  x, y = np.mgrid[xslice, yslice]
  pos = np.empty(x.shape + (2,))
  pos[:, :, 0] = x; pos[:, :, 1] = y
  rv = multivariate normal(mean, cov)
  plt.contourf(x, y, rv.pdf(pos), cmap='Blues')
  plt.plot(mean[0], mean[1], color='w', marker=mplcar, markersize=20, linestyle='
def plot particles(w particles, x particles):
  plt.scatter(x particles[:, 0], x particles[:, 1], c=w particles.flatten(), cmap
  mean = (w particles * x particles).sum(axis=0)
  plt.plot(mean[0], mean[1], color='y', marker=mplcar, markersize=20, linestyle='
```

```
def plot_landmarks(z):  
# Let's plot our global markers  
plt.plot(z[:, 0], z[:, 1], color='r', marker=mpltraffic, markersize=20, linesty  
plt.xlim(-0.5, 4.0)  
plt.ylim(-0.5, 2.5)  
plot_landmarks(z)  
plot_truncated_gaussian(\mu0, \Sigma0)
```



# Bayes filter



1

Let all the observation upto time t be represented by  $f_{1:t}$  and all the odometry inputs be denoted by  $d_{1:t-1}$ . The localization problem is the estimation of the probability  $p(x_t|f_{1:t},d_{1:t-1})$ . We assume the process generating the data to be a Markovian process with state  $x_t$ . This means that given  $x_{t-1}$ , future states  $x_t, x_{t+1}, \ldots$  are independent of all past states  $x_{t-2}, x_{t-3}, \ldots$ . This allows us to write the localization problem as a recursive equation

$$p(x_t|f_{1:t},d_{1:t-1}) = \int_{\mathcal{X}_{t-1}} p(x_t|f_t,d_{t-1},x_{t-1}) p(x_{t-1}|f_{1:t-1},d_{1:t-2}) dx_{t-1}$$

Note that the second term inside the integral is the solution of previous localization problem. The first term is broken down into two steps, (1) the motion update or prediction step (2) measurement update step. Due to Bayes rule, we can write

$$p(x_t|f_t,d_{t-1},x_{t-1}) = rac{p(f_t,d_{t-1}|x_t,x_{t-1})p(x_t|x_{t-1})}{p(f_t,d_{t-1}|x_{t-1})}$$

Using the independence of  $f_t \perp d_{t-1}|x_t$  and  $f_t \perp x_{t-1}|x_t$ , and  $p(f_t,d_{t-1}|x_t,x_{t-1})=p(f_t|x_t)p(d_{t-1}|x_t,x_{t-1})$ . Overall with some rearrangement, we get

$$p(x_t|f_t,d_{t-1},x_{t-1}) = rac{p(f_t|x_t)p(x_t|d_{t-1},x_{t-1})}{p(f_t|x_{t-1})}$$

This is convinient since we know the measurement model  $p(f_t|x_t)$  and the motion model  $p(x_t|d_{t-1},x_{t-1})$ . We assume that all observations are equally likely and hence only act as a normalizing factor:

$$p(x_t|f_t, d_{t-1}, x_{t-1}) \propto p(f_t|x_t)p(x_t|d_{t-1}, x_{t-1})$$

Putting this back in the Bayes recursive formulation, we get

$$p(x_t|f_{1:t},d_{1:t-1}) \propto p(f_t|x_t) \int_{\mathcal{X}_{t-1}} p(x_t|d_{t-1},x_{t-1}) p(x_{t-1}|f_{1:t-1},d_{1:t-2}) dx_{t-1}$$

### Motion update or prediction step

Computing the term  $\operatorname{ar{bel}}(x_{t-1}) = \int_{\mathcal{X}_{t-1}} p(x_t|d_{t-1},x_{t-1}) p(x_{t-1}|f_{1:t-1},d_{1:t-2}) dx_{t-1}$ 

### Measurement update step

Computing the term  $p(x_t|f_{1:t},d_{1:t-1}) \propto p(f_t|x_t) \mathrm{bel}(x_{t-1})$ 

The bayes filter can be implemented in many ways, depending upon how the probability distribution is represented:

1. Particle filter

- 2. Histogram fitler
- 3. Kalman filter

# Particle filter or Sequential Importance resampling

A probability distribution can be computationally represented as samples and weights.

Given a bunch of samples  $\{x_i\}_{i=1}^N$  samples sampled from a proposal distribution q(x)>0, if we want to compute the expected value of a function g(x) over target distribution p(x) 2

$$\mathbb{E}_{p(x)}[g(x)] = \sum_{i=1}^N rac{g(x)p(x)}{q(x)}$$

# Representing $p(x_0)$ as particles

First step is to convert  $x_0 \sim \mathcal{N}(\mu_0, \Sigma_0)$  as particles,  $(w_t^{(j)}, x_t^{(j)}): w_t^{(j)} \approxeq p(x_t^{(j)}|f_{1:t}, d_{1:t-1}).$ 

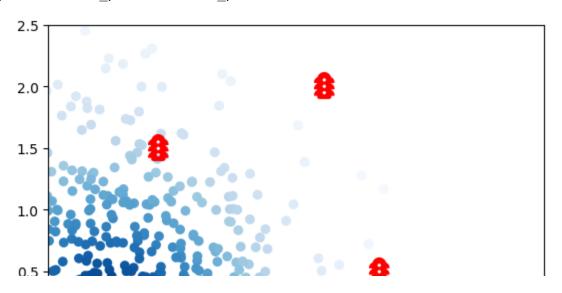
#@markdown Enter number of particles as
logM = 3 #@param {type: "slider", min: 1
M = 10\*\*logM
def initialize\_particles(M, μ0, Σ0):
 x0\_particles = np.random.multivariat
 w\_particles\_unnormalized = multivari
 w0\_particles = w\_particles\_unnormali
 return w0 particles, x0 particles

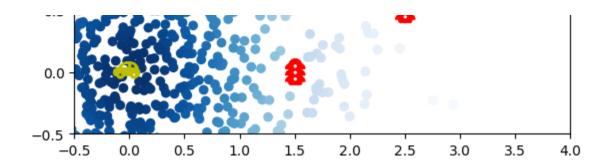
w0\_particles, x0\_particles = initialize\_
plot\_landmarks(z)
plot\_particles(w0\_particles, x0\_particle

Enter number of particles as  $10^{logM}$ 

logM:

3



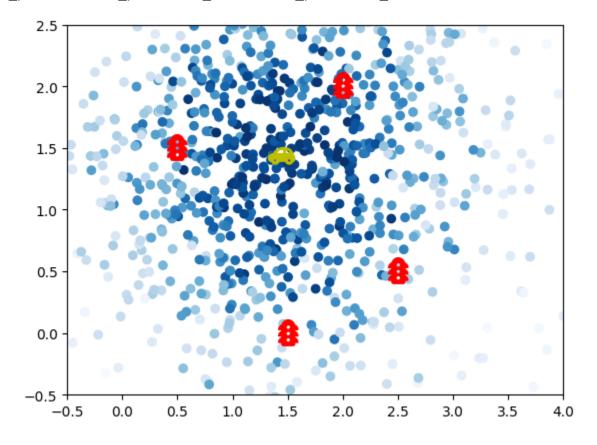


### Particle filter motion update or prediction step

$$x_{t+1} \sim \mathcal{N}(x_t + d_t, R_t)$$

```
mean_0 = (w0_particles * x0_particles).sum(axis=0)
def motion_update(w0_particles, x0_particles, d_t):
    x1_particles = np.empty_like(x0_particles)
    for i, x0 in enumerate(x0_particles):
        x1_particles[i, :] = np.random.multivariate_normal(x0 + d_t, R)
    return w0_particles, x1_particles
```

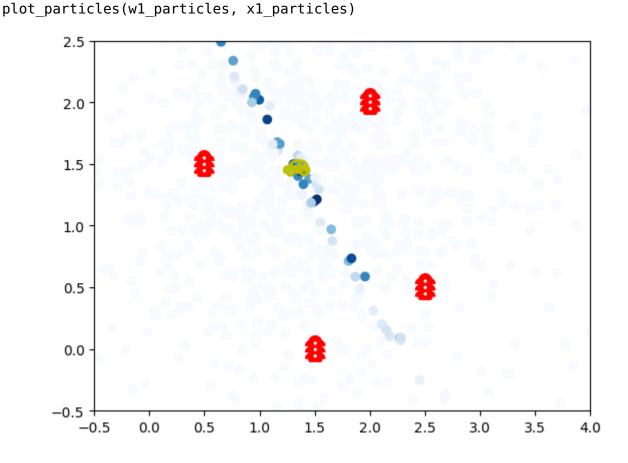
w1\_particles\_motion, x1\_particles\_motion = motion\_update(w0\_particles, x0\_particle
plot\_landmarks(z)
plot particles(w1 particles motion, x1 particles motion)



### Particle filter measurement update

$$egin{aligned} f_{i,t}|x_t &\sim \mathcal{N}(H(z_i-x_t),Q) \ p(f_t|x_t) &= \prod_i p(f_t|x_t) \ w_{t+1}^{(j)} &\propto w_t^{(j)} p(f_t|x_t) \end{aligned}$$

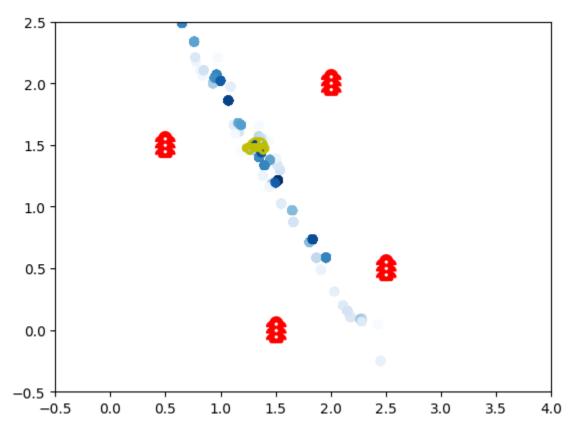
```
def measurement_update(w1_particles_motion, x1_particles_motion, visible_t, ft):
    diff_i_t = z[visible_t, np.newaxis, :] - x1_particles_motion[np.newaxis, :, :
    VMD = diff_i_t.shape
    obs_mean = (diff_i_t.reshape(-1, VMD[-1]) @ H.T).reshape(*VMD[:2],-1)
    p_f_t = 1
    for i, fit in enumerate(ft):
        p_f_t = p_f_t * multivariate_normal(fit, Q).pdf(obs_mean[i])
        assert p_f_t.shape[0] == VMD[1]
    w1_particles_unnormalized = w0_particles * p_f_t.reshape(-1,1)
    w1_particles = w1_particles_unnormalized / w1_particles_unnormalized.sum()
    return w1_particles, x1_particles_motion
w1_particles, x1_particles = measurement_update(w1_particles_motion, x1_particles_plot landmarks(z))
```



# Particle resampling step

```
def resample_particles(w1_particles, x1_particles):
    effective_number = 1 / (w1_particles**2).sum()
    if effective_number > w1_particles.shape[0]/2:
        return w1_particles, x1_particles
    x1_indices = np.random.choice(w1_particles.shape[0], w1_particles.shape[0], particles_resampled = x1_particles[x1_indices]
    w1_particles_resampled = w1_particles[x1_indices] / w1_particles[x1_indices].
    return w1_particles_resampled, x1_particles_resampled

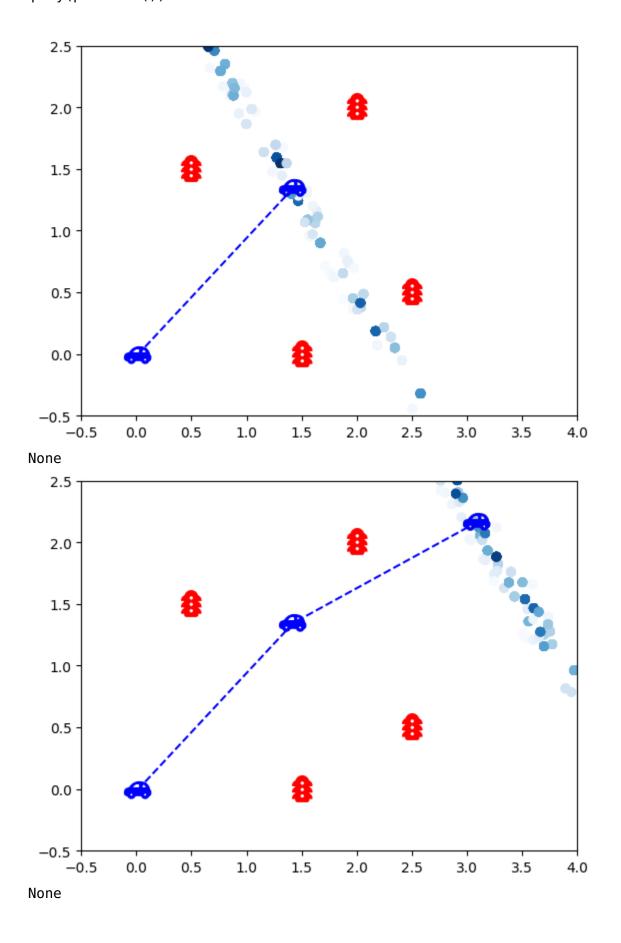
w1_particles_resampled, x1_particles_resampled = resample_particles(w1_particles, plot_landmarks(z))
plot_particles(w1_particles_resampled, x1_particles_resampled)
```



### Full particle particle filter algorithm

```
wt_particles, xt_particles = initialize_particles(M, μ0, Σ0)
means = [(wt_particles * xt_particles).sum(axis=0)]
for t in range(d.shape[0]):
    wlm, xlm = motion_update(wt_particles, xt_particles, d[t])
    wl, xl = measurement_update(wlm, xlm, visible_all_t[t+1], f_i_t[t+1])
    wt_particles, xt_particles = resample_particles(wl, xl)
    plot_landmarks(z)
    plot_particles(wt_particles, xt_particles)
    means.append((wt_particles * xt_particles).sum(axis=0))
    mean np = np.array(means)
```

plt.plot(mean\_np[:, 0], mean\_np[:, 1], markersize=20, marker=mplcar, linestyle= display(plt.show())



#### Kalman Filter

Kalman filter assumes Gaussian distributions and linear measurement and motion models.

Assume motion model

$$x_{t+1} = A_t x_t + B_t d_t + \epsilon_t, \quad \epsilon_t \sim \mathcal{N}(0, R_t)$$

and measurement model

$$f_t = H_t x_t + \eta_t, \quad \eta_t \sim \mathcal{N}(0,Q_t)$$

# Kalman filter model for the simple Localization problem

The measurement model in simple localization problem is linear when written as  $f_{i,t} = H(z_i - x_t) + \eta_t$ . It is affine. Moreover, it needs to be stacked to compute  $f_t$  for all i in a single matrix multiplication. Let the new state be defined as  $\bar{x}_t = [x_t^\top, 1]^\top$ . Then the motion model becomes,

$$ar{x}_{t+1} = ar{x}_t + egin{bmatrix} 1 & 0 \ 0 & 1 \ 0 & 0 \end{bmatrix} d_t + ar{\epsilon}_t, \quad ar{\epsilon}_t \sim \mathcal{N}(0, ar{R}_t)$$

where

$$ar{R}_t = egin{bmatrix} R_t & 0 \ 0 & \delta \end{bmatrix}$$

. 
$$\delta=10^{-6}$$
 .

The measurement model becomes.

$$egin{bmatrix} f_{1,t} \ f_{2,t} \ dots \ f_{N,t} \end{bmatrix} = egin{bmatrix} -H & Hz_1 \ -H & Hz_2 \ dots & dots \ -H & Hz_N \end{bmatrix} ar{x}_t + egin{bmatrix} \eta_1 \ \eta_2 \ dots \ \eta_N \end{bmatrix}$$

,

$$ar{\eta} \sim \mathcal{N} \left( 0, egin{bmatrix} Q & 0 & \dots & 0 \ 0 & Q & \dots & 0 \ dots & dots & \dots & dots \ 0 & 0 & \dots & Q \end{bmatrix} 
ight)$$

Hence, the Kalman filter constants for the simple localization problem are:

$$\lceil -H \quad Hz_1 \rceil \qquad \qquad \lceil Q \quad 0 \quad \dots \quad 0 \rceil$$

$$ar{A}_t = I_{3 imes 3}, ar{B}_t = I_{3 imes 2}, ar{H}_t = egin{bmatrix} -H & Hz_2 \ dots & dots \ -H & Hz_N \end{bmatrix}, ar{Q}_t = egin{bmatrix} dots & Q & \dots & 0 \ dots & dots & \dots & dots \ 0 & 0 & \dots & Q \end{bmatrix}, ar{R}_t = egin{bmatrix} R & 0 \ 0 & \delta \end{bmatrix}$$

```
mu \theta = \text{np.hstack}((\mu \theta, 1))
sigma 0 = np.eye(mu 0.shape[0])
sigma \ \Theta[:-1, :-1] = \Sigma \Theta
D = \mu 0.shape[0]
Abart = np.eye(D+1)
Bbart = np.zeros((D+1,D))
Bbart[:D,:D] = np.eye(D)
def Hbart create(H, z, visible t):
  Hstacked = np.repeat(H, len(visible t), axis=0)
  Hzis = []
  for i, lid in enumerate(visible t):
    Hzis.append(H @ z[lid])
  return np.hstack((-Hstacked, np.vstack(Hzis)))
Rbart = np.eye(D+1) * 1e-6
Rbart[:D, :D] = R
def Qbart create(Q, visible t):
  from scipy.linalg import block diag
  return block diag(*([Q]*len(visible t)))
```

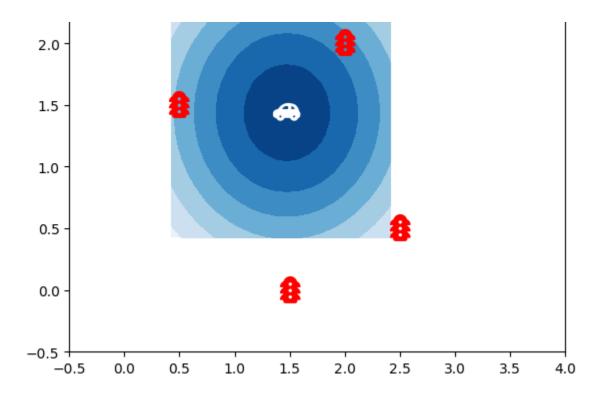
### Motion or prediction step

Thrun et al. 2005 marginal of a gaussian marginal of a gaussian without schur complement kalman filter using marginal

$$ar{\mu}_{t+1} = A\mu_t + Bd_t \ ar{\Sigma}_{t+1} = A_t \Sigma_t A_t^{ op} + R_t$$

```
def motion_update_kf(mu_t, sigma_t, d_t, At, Bt, Rt):
    mu_tp1 = At @ mu_t + Bt @ d_t
    sigma_tp1 = At @ sigma_t @ At.T + Rt
    return mu_tp1, sigma_tp1

mu_1_motion, sigma_1_motion = motion_update_kf(mu_0, sigma_0, d[0], Abart, Bbart,
plot_truncated_gaussian(mu_1_motion[:D], sigma_1_motion[:D, :D])
plot_landmarks(z)
```



### Measurement update step

#### Thrun et al. 2005

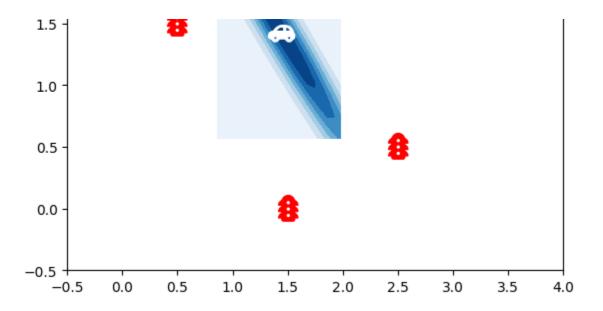
$$K_{t+1} = ar{\Sigma}_{t+1} H_t^ op (H_t ar{\Sigma}_{t+1} H_t^ op + Q_t)^{-1} \ \mu_{t+1} = ar{\mu}_{t+1} + K_{t+1} (f_t - H_t \mu_{t+1}) \ \Sigma_{t+1} = (I - K_{t+1} H_t) ar{\Sigma}_{t+1}$$

```
def measurement_update_kf(mu_t, sigma_t, f_is, Hbart, Qbart):
    diff = f_is - Hbart @ mu_t
    HSigmaH_minus_Q = Hbart @ sigma_t @ Hbart.T + Qbart
    mu_tp1 = mu_t + sigma_t @ Hbart.T @ np.linalg.solve(HSigmaH_minus_Q, diff)
    I = np.eye(mu_t.shape[0])
    sigma_tp1 = (I - sigma_t @ Hbart.T @ np.linalg.solve(HSigmaH_minus_Q, Hbart)) @
    return mu_tp1, sigma_tp1

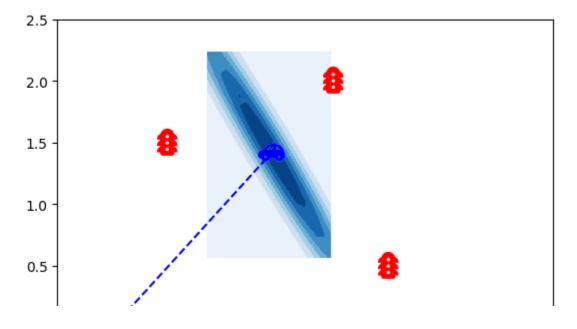
f_is = f_i_t[1].flatten()
mu_1, sigma_1 = measurement_update_kf(mu_1_motion, sigma_1_motion, f_is, Hbart_cr

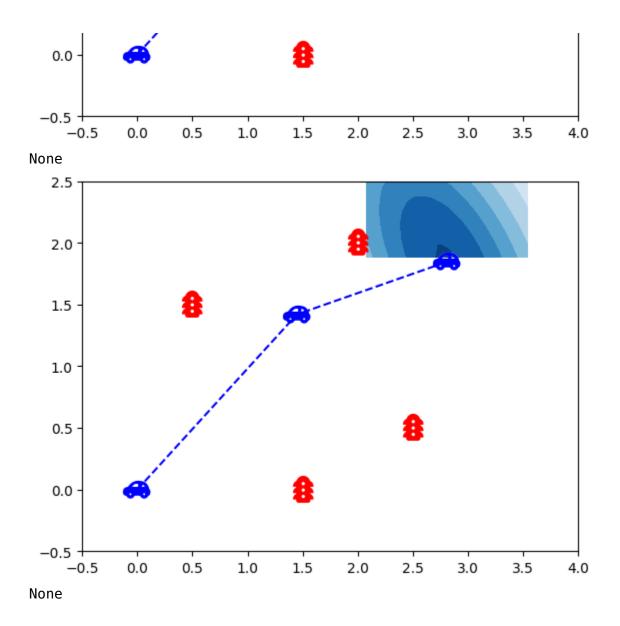
plot_truncated_gaussian(mu_1[:D], sigma_1[:D, :D])
plot_landmarks(z)
```





### Full Kalman Filter algorithm





[MacKay1998]: MacKay, David JC. "Introduction to monte carlo methods." Learning in graphical models. Springer, Dordrecht, 1998. 175-204

[Thrun2005]: Thrun, Sebastian, Wolfram Burgard, and Dieter Fox. Probabilistic robotics. MIT press, 2005