

# A note on gaussian AR(1) process [WiP!!!]

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July 18, 2021

## Abstract

This note presents elementary properties of autoregressive AR(1) gaussian process. It is intended as a future reference for the author, hence not too much effort is made to ensure that propounding of all steps is sufficient to satisfy the tastes of more fastidious readers. Also, notation is at times quite sloppy and is assumed to be self-explanatory.

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# 1 AR(1) process - overview

Let  $\{Y_t\}_{t=-\infty}^{\infty}$  be a weakly stationary stochastic process that obeys the equation:

$$Y_t = \phi Y_{t-1} + c + \epsilon_t \quad (1)$$

with gaussian error term, i.e.:

$$\epsilon_t \sim N(0, \sigma^2) \quad (2)$$

and all the error terms being uncorrelated. Furthermore, suppose that  $T+1$  consecutive realizations of the process  $\{Y_t\}$  are observed for times  $t = 0, 1, \dots, T$  and that they are denoted  $y_0, y_1, y_2, \dots, y_T$ .

## 2 Likelihood function

I start by looking at the conditional distribution of  $Y_t$  with respect to  $Y_{t-1}$ . To this end, suppose that the value of  $Y_{t-1}$  is known and equal to  $y_{t-1}$ . Then:

$$Y_t = c + \phi y_{t-1} + \epsilon_t$$

Unsurprisingly, then:

$$E(Y_t | Y_{t-1} = y_{t-1}) = c + \phi y_{t-1}$$

$$Var(Y_t | Y_{t-1} = y_{t-1}) = \sigma^2$$

As  $\{\epsilon_t\}$  has been assumed to be gaussian, it follows that the conditional probability density function of  $Y_t$  with respect to  $Y_{t-1} = y_{t-1}$  is given by:

$$f_{Y_t | Y_{t-1}}(y_t | y_{t-1}) = \frac{1}{\sigma \sqrt{2\pi}} \exp \left( -\frac{1}{2} \frac{(y_t - c - \phi y_{t-1})^2}{\sigma^2} \right) \quad (3)$$

This formula will be of great use in the next step which is rewriting of joint probability density function of  $(Y_T, Y_{T-1}, \dots, Y_1, Y_0)$ .

Consequently, let's deal with the joint density. Using the simple fact presented in [1], I have:

$$\begin{aligned} f_{Y_T, Y_{T-1}, \dots, Y_1, Y_0}(y_T, y_{T-1}, \dots, y_1, y_0) &= f_{Y_T | Y_{T-1}, \dots, Y_1, Y_0}(y_T | y_{T-1}, \dots, y_1, y_0) \cdot \\ &\quad f_{Y_{T-1} | Y_{T-2}, \dots, Y_1, Y_0}(y_{T-1} | y_{T-2}, \dots, y_1, y_0) \cdot \dots \cdot \\ &\quad f_{Y_1 | Y_0}(y_1 | y_0) \cdot f_{Y_0}(y_0) \end{aligned}$$

Written more succintly:

$$f_{Y_T, Y_{T-1}, \dots, Y_1, Y_0}(y_T, y_{T-1}, \dots, y_1, y_0) = f_{Y_0}(y_0) \cdot \prod_{t=1}^T f_{Y_t|Y_{t-1}, \dots, y_0}(y_t|y_{t-1}, \dots, y_0)$$

Since I aim for *conditional* likelihood, I divide both sided by  $f_{Y_0}(y_0)$  and obtain:

$$f_{Y_T, Y_{T-1}, \dots, Y_1|Y_0}(y_T, y_{T-1}, \dots, y_1|y_0) = \prod_{t=1}^T f_{Y_t|Y_{t-1}, \dots, y_0}(y_t|y_{t-1}, \dots, y_0)$$

Now, by equation (3) one can easily see that since the value of the process at time  $t$  is conditional only on its value at time  $t - 1$  the conditional densities in the product in the equation above reduce as follows:

$$f_{Y_t|Y_{t-1}, Y_{t-2}, \dots, Y_0}(y_t|y_{t-1}, y_{t-2}, \dots, y_0) = f_{Y_t|Y_{t-1}}(y_t|y_{t-1}) \quad (4)$$

The intuitive explanation of the above is that by (3) the density of  $Y_t$  is conditional on  $Y_{t-1}$  only and not on  $Y_{t-2}, Y_{t-3}, \dots$ . More formal approach to this topic is presented in one of the later sections.

For now, let's focus on obtaining an alaytical formula for the conditional likelihood function. Substituting (4) into (2) yields:

$$f_{Y_T, Y_{T-1}, \dots, Y_1|Y_0}(y_T, y_{T-1}, \dots, y_1|y_0) = \prod_{t=1}^T \frac{1}{\sigma\sqrt{2\pi}} \exp\left(-\frac{1}{2} \frac{(y_t - c - y_{t-1})^2}{\sigma^2}\right) \quad (5)$$

Since we consider the above as a function of the model parameters  $\phi, c, \sigma$  conditional on the sample  $y_0, y_1, \dots, y_T$  I introduce the conditional liekelihood  $L(\phi, c, \sigma|y_T, t_{T-1}, \dots, y_1, y_0)$  as defined using the equation (5) :

$$= \prod_{t=1}^T \frac{1}{\sigma\sqrt{2\pi}} \exp\left(-\frac{1}{2} \frac{(y_t - c - y_{t-1})^2}{\sigma^2}\right) \quad (6)$$

### 3 Derivation of condntional maximum likelihood estimators of the parameters

This section contains transformations of the formula...

## 4 Rewriting AR(1) as MA( $\infty$ )

Let's go back to square one and aim for another interesting feature of AR(1) process. Start with the equation (1):

$$Y_t = c + \phi y_{t-1} + \epsilon_t$$

Writing this equation for time  $t - 1$ :

$$Y_{t-1} = c + \phi y_{t-2} + \epsilon_{t-1}$$

Substitution of the formula for  $Y_{t-1}$  into the formula for  $Y_t$  yields:

$$\begin{aligned} Y_t &= c + \phi(c + \phi y_{t-2} + \epsilon_{t-1}) + \epsilon_t \\ &= c + \phi c + \epsilon_t + \phi \epsilon_{t-1} + \phi^2 y_{t-2} \end{aligned}$$

Iterating one more time I obtain:

$$Y_t = c(1 + \phi + \phi^2) + (\epsilon_t + \phi \epsilon_{t-1} + \phi^2 \epsilon_{t-2}) + \phi^3 Y_{t-3}$$

Generalizing:

$$Y_t = c \sum_{k=0}^n \phi^k + \sum_{k=0}^n \phi^k \epsilon_{t-k} + \phi^{n+1} Y_{t-n-1}$$

Since  $|\phi| < 1$ , it follows that  $\lim_{n \rightarrow \infty} \phi^{n+1} = 0$ . Therefore<sup>1</sup>:

$$Y_t = \frac{c}{1 - \phi} + \sum_{k=0}^{\infty} \phi^k \epsilon_{t-k}$$

## 5 Unconditional distribution of $Y_t$

I continue the analysis in the vein of (4). Since all  $\epsilon_t$  are normally, independently distributed with  $\epsilon_t \sim N(0, \sigma^2)$ , then the  $Y_t$  is also normally distributed with mean:

$$E(Y_t) = \frac{c}{1 - \phi}$$

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<sup>1</sup>I deliberately restrain from intruding the polynomial notation here to keep focus of what is of greatest interest here.

and variance:

$$\begin{aligned} \text{Var}(Y_t) &= \text{Var}(\sum_{n=0}^{\infty} \phi^n \epsilon_{t-n}) \\ &= \sum_{n=0}^{\infty} \phi^{2n} \sigma^2 \\ &= \frac{\sigma^2}{1-\phi^2} \end{aligned}$$

This allows me to write down the unconditional density of  $Y_t$ :

$$f_{Y_t}(y_t) = \frac{1}{\sqrt{2\pi}(\sigma/\sqrt{1-\phi^2})} \exp\left(-\frac{1}{2} \frac{(y_t - c/(1-\phi))^2}{\sigma^2/(1-\phi^2)}\right)$$

## 6 Joint density of $(Y_{t+2}, Y_{t+1}, Y_t)$

Building on top of the results of two previous sections I aim at deriving the joint density of  $\{Y_t\}$  at three consecutive time instances, i.e.  $Y_{t+2}, Y_{t+1}, Y_t$ .

## 7 Unconditional maximum likelihood

## 8 References

The exposition in this note follows closely the one in [2], chapter 5. Another useful reference is [3].

## References

- [1] A. Wegrzyn, “A note on conditional probabilities of multivariate continuous variables with applications to time series analysis.” [https://github.com/wegar-2/latex\\_files/blob/master/probability\\_calculus/conditional\\_probabilities\\_for\\_time\\_series.pdf](https://github.com/wegar-2/latex_files/blob/master/probability_calculus/conditional_probabilities_for_time_series.pdf), 2021.
- [2] J. D. Hamilton, *Time Series Analysis*. Princeton University Press, 1994.
- [3] W. A. Fuller, *Introduction to Statistical Time Series, 2nd Edition*. Wiley, 1996.