

Notations (common)

Each row is the recording for neuron i , $\mathbf{y}_i = (y_{i1}, \dots, y_{iT})'$, $i = 1, \dots, N$. Denote the cluster index for neuron i as $z_i \in \{1, \dots\}$. The number of neurons in cluster j is $n_j = \sum_{i=1}^N I(z_i = j)$, and $\sum_{j=1,2,\dots} n_j = N$.

Model

Denote the latent vector in cluster j as $\mathbf{x}_t^{(j)} \in R^{p_j}$. For simplicity, assume all $p_j = p$. Each observation follows a Poisson distribution as follows:

$$\log \lambda_{it} = d_i + \mathbf{c}_i' \mathbf{x}_t^{(z_i)}$$

$$y_{it} \sim \text{Poisson}(\lambda_{it})$$

Where $\mathbf{c}_i \in R^p$ and $\mathbf{x}_t^{(z_i)} \in R^p$.

Denote all latent states as $\mathbf{x}_t = (\mathbf{x}_t^{(1)'}, \mathbf{x}_t^{(2)'}, \dots)'$ and they evolve linearly with Gaussian noise:

$$\mathbf{x}_1 \sim N(\mathbf{x}_0, \mathbf{Q}_0)$$

$$\mathbf{x}_{t+1} | \mathbf{x}_t \sim N(\mathbf{A}\mathbf{x}_t + \mathbf{b}, \mathbf{Q})$$

To simplify, assume \mathbf{Q}_0 is known (e.g. $\mathbf{Q}_0 = \mathbf{I}$).

If we assume block diagonal (as in Joshua et al., 2020) for process noise covariance, we can write things as:

$$\mathbf{x}_{t+1}^{(j)} | \mathbf{x}_t^{(1)}, \mathbf{x}_t^{(2)}, \dots \sim N\left(\sum_{l=1,\dots} \mathbf{A}_{j \leftarrow l} \mathbf{x}_t^{(l)} + \mathbf{b}_j, \mathbf{Q}^{(j)}\right)$$

Notice $\{\mathbf{A}_{j \leftarrow l}\}$ forms the full transition matrix as:

$$\mathbf{A} = \begin{pmatrix} \mathbf{A}_{1 \leftarrow 1} & \mathbf{A}_{1 \leftarrow 2} & \dots \\ \mathbf{A}_{2 \leftarrow 1} & \mathbf{A}_{2 \leftarrow 2} & \dots \\ \dots & \dots & \dots \end{pmatrix}$$

If the j^{th} row block of \mathbf{A} is $\mathbf{A}_j = (\mathbf{A}_{j \leftarrow 1} \quad \mathbf{A}_{j \leftarrow 2} \quad \dots)$. Then, $\sum_{l=1,\dots} \mathbf{A}_{j \leftarrow l} \mathbf{x}_t^{(l)} + \mathbf{b}_j = \mathbf{A}_j \mathbf{x}_t + \mathbf{b}_j$.

If we further let \mathbf{Q} be diagonal: denote the k^{th} row of \mathbf{x}_t , \mathbf{A} , \mathbf{b} as x_{kt} , \mathbf{a}_k , b_k . The corresponding process noise variance as q_k . Then:

$$x_{k,t+1} | x_{kt} \sim N(\mathbf{a}_k' \mathbf{x}_t + b_k, q_k)$$

The parameters need to estimate:

- (1) Latent vectors: \mathbf{x}_t
- (2) Initials: \mathbf{x}_0
- (3) Linear mapping for latent vectors: $\{d_i\}_{i=1}^N$ and $\{\mathbf{c}_i\}_{i=1}^N$
- (4) Linear dynamics for latent vectors: \mathbf{A} and \mathbf{b}
- (5) Process noise: \mathbf{Q}

Since the progress noise is independent in the model, $f(\mathbf{y}_i|\boldsymbol{\Theta}_j) = \prod_{t=1}^T P(y_{it}|\boldsymbol{\Theta}_j)$, where $P(\cdot)$ is the Poisson density and $\boldsymbol{\Theta}_j$ is the parameters in cluster j .

Conditional Priors

- (1) Latent vectors $\mathbf{x}_t^{(j)}$: the conditional prior is defined by
(Assume there are J clusters)

$$\mathbf{x}_1 \sim N(\mathbf{x}_0, \mathbf{Q}_0)$$

$$\mathbf{x}_{t+1}|\mathbf{x}_t \sim N(\mathbf{A}\mathbf{x}_t + \mathbf{b}, \mathbf{Q})$$

- (2) Initials \mathbf{x}_0 :
(Assume there are J clusters)

$$\mathbf{x}_0 \sim N(\boldsymbol{\mu}_{\mathbf{x}_{00}}, \boldsymbol{\Sigma}_{\mathbf{x}_{00}})$$

$$\boldsymbol{\mu}_{\mathbf{x}_{00}} = \mathbf{0}_{Jp} \text{ and } \boldsymbol{\Sigma}_{\mathbf{x}_{00}} = 10^2 \mathbf{I}_{Jp}$$

- (3) Linear mapping for latent vectors $\{d_i\}_{i=1}^N$ and $\{c_i\}_{i=1}^N$:

$$(d_i, c_i)' \sim N(\boldsymbol{\mu}_{dc0}, \boldsymbol{\Sigma}_{dc0})$$

$$\text{Where } \boldsymbol{\mu}_{dc0} = \mathbf{0}_{p+1} \text{ and } \boldsymbol{\Sigma}_{dc0} = \mathbf{I}_{p+1} \times 10^{-2}$$

- (4) Linear dynamics for latent vectors \mathbf{A} and \mathbf{b} :
(if the number of cluster is J)

- a. Assume \mathbf{Q} be diagonal:

$$(b_k, a'_k)' \sim N(\boldsymbol{\mu}_{ba_k0}, \boldsymbol{\Sigma}_{ba0})$$

$$\text{Where } \boldsymbol{\mu}_{ba_k0} = (0, e'_k)' \text{ and } \boldsymbol{\Sigma}_{ba0} = 0.25 \mathbf{I}_{Jp+1}$$

- b. Assume \mathbf{Q} be block-diagonal:

$$\text{Denote } \tilde{\mathbf{A}}_j = \text{vec}(\mathbf{A}_j)$$

$$(b'_j, \tilde{\mathbf{A}}_j)' \sim N(\boldsymbol{\mu}_{bA_0}, \boldsymbol{\Sigma}_{bA_0})$$

$$\text{Where } \boldsymbol{\mu}_{bA_0} = (\mathbf{0}'_{p+(j-1)p^2}, \text{vec}(\mathbf{I}_p)', \mathbf{0}'_{(J-j)p^2})' \text{ and } \boldsymbol{\Sigma}_{bA_0} = 0.25 \mathbf{I}_{p(1+pJ)}$$

- c. No constraints on \mathbf{Q} :

$$\text{Denote } \tilde{\mathbf{A}} = \text{vec}(\mathbf{A})$$

$$(b', \tilde{\mathbf{A}})' \sim N(\boldsymbol{\mu}_{BA_0}, \boldsymbol{\Sigma}_{BA_0})$$

$$\text{Where } \boldsymbol{\mu}_{BA_0} = (\mathbf{0}'_{Jp}, \text{vec}(\mathbf{I}_{Jp})')' \text{ and } \boldsymbol{\Sigma}_{BA_0} = 0.25 \mathbf{I}_{Jp(1+Jp)}$$

- (5) Process noise \mathbf{Q} :

- a. Assume \mathbf{Q} be diagonal:

$$q_k \sim IG(\frac{\nu_0}{2}, \frac{\nu_0 \sigma_0^2}{2})$$

$$\text{Where } \nu_0 = 4 \text{ and } \sigma_0^2 = 10^{-4}$$

- b. Assume \mathbf{Q} be block-diagonal:

$$\mathbf{Q}^{(j)} \sim W^{-1}(\Psi_0, \nu_0)$$

$$\text{Where } \nu_0 = p + 2 \text{ and } \Psi_0 = \mathbf{I}_p \times 10^{-4}.$$

(To make the mean of $\mathbf{Q}^{(j)}$ loosely centered around $\mathbf{I}_p \times 10^{-4}$)

c. No constraints on \mathbf{Q} :

(if the number of cluster is J)

$$\mathbf{Q} \sim W^{-1}(\Psi_{\mathbf{Q}_0}, \nu_{\mathbf{Q}_0})$$

Where $\nu_{\mathbf{Q}_0} = Jp + 2$ and $\Psi_{\mathbf{Q}_0} = \mathbf{I}_{Jp} \times 10^{-4}$.

MCMC iteration

(1) Update $\mathbf{x}_t^{(j)}$:

use local Normal approximation at prior, i.e. update by adaptive smoothing.

Notice: adaptive filtering is not the “exact Laplace approximation”. In Laplace approximation, we are evaluating at the mode/ maximum of posterior, but in adaptive filtering, it is evaluated at the prior...

An IMPROTANT drawback for adaptive filtering/ smoothing is that it don't give the covariance between different time steps. This will ruin the sampling and estimations for other parameters. But the mean estimations for adaptive filtering/ smoothing is perfect.

solution: Use the posterior mean from adaptive smoothing as a warm start and then pass to NR (the Exact Laplace Approximation) to refine the estimation if necessary. If too slow, just use adaptive smoother estimation.

THE EXACT LAPLACE APPROXIMATION

Denote t^{th} column of mean firing rate and observation as $\tilde{\lambda}_t = (\lambda_{1t}, \dots, \lambda_{Nt})'$ and $\tilde{\mathbf{y}}_t = (y_{1t}, \dots, y_{Nt})'$. The linear mapping matrix for all observations is \mathbf{C} , such that $\log \tilde{\lambda}_t = \mathbf{d} + \mathbf{C}\mathbf{x}_t$. Let $\mathbf{x} = (\mathbf{x}'_1, \dots, \mathbf{x}'_T)'$ and $f(\mathbf{x}) = \log P(\mathbf{x}|\{\mathbf{y}_i\}_{i=1}^N, \mathbf{C}, \mathbf{Q}_0, \mathbf{A}, \mathbf{b}, \mathbf{Q}, \dots)$

The first and second derivative with respect to \mathbf{x}

$$\begin{aligned} \frac{\partial f}{\partial \mathbf{x}_1} &= \mathbf{C}'(\tilde{\mathbf{y}}_1 - \tilde{\lambda}_1) - \mathbf{Q}_0^{-1}(\mathbf{x}_1 - \mathbf{x}_0) + \mathbf{A}'\mathbf{Q}^{-1}(\mathbf{x}_2 - \mathbf{A}\mathbf{x}_1 - \mathbf{b}) \\ \frac{\partial f}{\partial \mathbf{x}_t} &= \mathbf{C}'(\tilde{\mathbf{y}}_t - \tilde{\lambda}_t) - \mathbf{Q}^{-1}(\mathbf{x}_t - \mathbf{A}\mathbf{x}_{t-1} - \mathbf{b}) + \mathbf{A}'\mathbf{Q}^{-1}(\mathbf{x}_{t+1} - \mathbf{A}\mathbf{x}_t - \mathbf{b}) \\ \frac{\partial f}{\partial \mathbf{x}_T} &= \mathbf{C}'(\tilde{\mathbf{y}}_T - \tilde{\lambda}_T) - \mathbf{Q}^{-1}(\mathbf{x}_T - \mathbf{A}\mathbf{x}_{T-1} - \mathbf{b}) \\ \frac{\partial^2 f}{\partial \mathbf{x}_1 \partial \mathbf{x}'_1} &= -\mathbf{C}'\text{Diag}(\tilde{\lambda}_1)\mathbf{C} - \mathbf{Q}_0^{-1} - \mathbf{A}'\mathbf{Q}^{-1}\mathbf{A} & \frac{\partial^2 f}{\partial \mathbf{x}_1 \partial \mathbf{x}'_2} &= \mathbf{A}'\mathbf{Q}^{-1} \\ \frac{\partial^2 f}{\partial \mathbf{x}_t \partial \mathbf{x}'_t} &= -\mathbf{C}'\text{Diag}(\tilde{\lambda}_t)\mathbf{C} - \mathbf{Q}^{-1} - \mathbf{A}'\mathbf{Q}^{-1}\mathbf{A} & \frac{\partial^2 f}{\partial \mathbf{x}_t \partial \mathbf{x}'_{t-1}} &= \mathbf{Q}^{-1}\mathbf{A} & \frac{\partial^2 f}{\partial \mathbf{x}_t \partial \mathbf{x}'_{t+1}} &= \mathbf{A}'\mathbf{Q}^{-1} \\ \frac{\partial^2 f}{\partial \mathbf{x}_T \partial \mathbf{x}'_T} &= -\mathbf{C}'\text{Diag}(\tilde{\lambda}_T)\mathbf{C} - \mathbf{Q}^{-1} \end{aligned}$$

So, the gradient is:

$$\nabla = \frac{\partial f}{\partial \mathbf{x}} = \left(\left(\frac{\partial f}{\partial \mathbf{x}_1} \right)', \dots, \left(\frac{\partial f}{\partial \mathbf{x}_T} \right)' \right)'$$

And the block tri-diagonal Hessian:

$$H = \frac{\partial^2 f}{\partial \mathbf{x} \partial \mathbf{x}'} = \begin{pmatrix} \frac{\partial^2 f}{\partial \mathbf{x}_1 \partial \mathbf{x}_1'} & \mathbf{A}' \mathbf{Q}^{-1} & 0 & \dots & 0 \\ \mathbf{Q}^{-1} \mathbf{A} & \frac{\partial^2 f}{\partial \mathbf{x}_2 \partial \mathbf{x}_2'} & \mathbf{A}' \mathbf{Q}^{-1} & \dots & \vdots \\ 0 & \mathbf{Q}^{-1} \mathbf{A} & \frac{\partial^2 f}{\partial \mathbf{x}_3 \partial \mathbf{x}_3'} & \dots & \vdots \\ \vdots & \vdots & \vdots & \ddots & \vdots \\ 0 & \dots & \dots & \dots & \frac{\partial^2 f}{\partial \mathbf{x}_T \partial \mathbf{x}_T'} \end{pmatrix}$$

Use Newton-Raphson to find $\boldsymbol{\mu}_x = \operatorname{argmax}_x (f(\mathbf{x}))$ and $\boldsymbol{\Sigma}_x = - \left[\frac{\partial^2 f}{\partial \mathbf{x} \partial \mathbf{x}'} \Big|_{\mathbf{x}=\boldsymbol{\mu}_x} \right]^{-1}$

When using Newton-Raphson, $H \setminus \nabla$ in MATLAB will make use of block tri-diagonal structure automatically.

See two examples in newtonTest.m.

To sample efficiently, use Cholesky decomposition of $\boldsymbol{\Sigma}_x^{-1} = \mathbf{R}\mathbf{R}'$: sample $\mathbf{Z} \sim N(\mathbf{R}'\boldsymbol{\mu}_x, \mathbf{I})$, then $\mathbf{x} = (\mathbf{R}')^{-1}\mathbf{Z} \sim N(\boldsymbol{\mu}_x, \boldsymbol{\Sigma}_x)$.

(2) Update \mathbf{x}_0 :

$$P(\mathbf{x}_0 | \mathbf{x}_1, \mathbf{Q}_0 \dots) \propto N(\mathbf{x}_1 | \mathbf{x}_0, \mathbf{Q}_0) N(\mathbf{x}_0 | \boldsymbol{\mu}_{x_0}, \boldsymbol{\Sigma}_{x_0})$$

Because of independence, we can update element by element. To write things easily, just write in matrix form. By conjugacy, $\mathbf{x}_0 | \mathbf{x}_1, \mathbf{Q}_0 \dots \sim N(\boldsymbol{\mu}_{x_0}, \boldsymbol{\Sigma}_{x_0})$

$$\begin{aligned} \boldsymbol{\Sigma}_{x_0} &= [\boldsymbol{\Sigma}_{x_0}^{-1} + \mathbf{Q}_0^{-1}]^{-1} \\ \boldsymbol{\mu}_{x_0} &= \boldsymbol{\Sigma}_{x_0} (\boldsymbol{\Sigma}_{x_0}^{-1} \boldsymbol{\mu}_{x_0} + \mathbf{Q}_0^{-1} \mathbf{x}_1) \end{aligned}$$

(3) Update $\{d_i\}_{i=1}^N$ and $\{\mathbf{c}_i\}_{i=1}^N$:

$$P(d_i, \mathbf{c}_i | \mathbf{y}_i, \{\mathbf{x}_t^{(z_i)}\}_{t=1}^T, \dots) \propto \left[\prod_{t=1}^T P(y_{it} | \mathbf{x}_t^{(z_i)}, d_i, \mathbf{c}_i) \right] P((d_i, \mathbf{c}_i)')$$

Notice $d_i + \mathbf{c}_i' \mathbf{x}_t^{(z_i)} = (1, \mathbf{x}_t^{(z_i)})' (d_i, \mathbf{c}_i)'$. Then the problem is reduced to Bayesian Poisson regression. Denote $(d_i, \mathbf{c}_i)' = \boldsymbol{\zeta}_i \in R^{p+1}$ and $(1, \mathbf{x}_t^{(z_i)})' = \tilde{\mathbf{x}}_t^{(z_i)}$. To get an efficient update, use the Laplace approximation: $P(\boldsymbol{\zeta}_i | \mathbf{y}_i, \{\mathbf{x}_t^{(z_i)}\}_{t=1}^T, \dots) = \exp f(\boldsymbol{\zeta}_i) \approx N(\boldsymbol{\zeta}_i | \boldsymbol{\mu}_{\zeta_i}, \boldsymbol{\Sigma}_{\zeta_i})$

$$\frac{\partial f}{\partial \boldsymbol{\zeta}_i} = \frac{\partial l}{\partial \boldsymbol{\zeta}_i} - \boldsymbol{\Sigma}_{dc0}^{-1} (\boldsymbol{\zeta}_i - \boldsymbol{\mu}_{dc0}) = \left[\sum_{t=1}^T \tilde{\mathbf{x}}_t^{(z_i)} (y_{it} - \lambda_{it}) \right] - \boldsymbol{\Sigma}_{dc0}^{-1} (\boldsymbol{\zeta}_i - \boldsymbol{\mu}_{dc0})$$

$$\frac{\partial^2 f}{\partial \zeta_i \partial \zeta_i'} = \frac{\partial^2 l}{\partial \zeta_i \partial \zeta_i'} - \Sigma_{dc0}^{-1} = - \left[\sum_{t=1}^T \lambda_{it} \tilde{\mathbf{x}}_t^{(z_i)} \tilde{\mathbf{x}}_t^{(z_i)'} \right] - \Sigma_{dc0}^{-1}$$

Where l is Poisson log-likelihood.

Use Newton-Raphson to find $\mu_{\zeta_j} = \operatorname{argmax}_{\zeta_j} (f(\zeta_j))$ and $\Sigma_{\zeta_j} = - \left[\frac{\partial^2 f}{\partial \zeta_j \partial \zeta_j'} \Big|_{\zeta_j = \mu_{\zeta_j}} \right]^{-1}$

(4) Update \mathbf{A} and \mathbf{b}

a. Assume \mathbf{Q} be diagonal:

$$P((b_k, \mathbf{a}_k')' | \{x_{kt}\}_{t=1}^T, q_k, \dots) \propto \left[\prod_{t=2}^T N(x_{kt} | \mathbf{a}_k' \mathbf{x}_{t-1} + b_k, q_k) \right] P((b_k, \mathbf{a}_k')')$$

Again, rewrite $\mathbf{a}_k' \mathbf{x}_{t-1} + b_k = (1, \mathbf{x}_{t-1}') (b_k, \mathbf{a}_k')'$. Then the problem is reduced to Bayesian linear regression. Denote $(b_k, \mathbf{a}_k')' = \boldsymbol{\gamma}_k \in R^{p+1}$ and $\tilde{\mathbf{X}} = ((1, \mathbf{x}_1')', \dots, (1, \mathbf{x}_{T-1}')')$. Then by conjugacy,

$$\begin{aligned} P(\boldsymbol{\gamma}_k | \{x_{kt}\}_{t=1}^T, q_k, \dots) &= N(\boldsymbol{\gamma}_k | \boldsymbol{\mu}_{\boldsymbol{\gamma}_k}, \boldsymbol{\Sigma}_{\boldsymbol{\gamma}_k}) \\ \boldsymbol{\Sigma}_{\boldsymbol{\gamma}_k} &= \left[\boldsymbol{\Sigma}_{ba0}^{-1} + \frac{\tilde{\mathbf{X}}' \tilde{\mathbf{X}}}{q_k} \right]^{-1} \\ \boldsymbol{\mu}_{\boldsymbol{\gamma}_k} &= \boldsymbol{\Sigma}_{\boldsymbol{\gamma}_k} (\boldsymbol{\Sigma}_{ba0}^{-1} \boldsymbol{\mu}_{ba0} + \tilde{\mathbf{X}}' (x_{k2}, \dots, x_{kT})' / q_k) \end{aligned}$$

b. Assume \mathbf{Q} be block-diagonal:

$$P((b_j', \tilde{\mathbf{A}}_j')' | \{\mathbf{x}_t^{(j)}\}_{t=1}^T, \mathbf{Q}^{(j)}, \dots) \propto \left[\prod_{t=2}^T N(\mathbf{x}_t^{(j)} | \mathbf{A}_j \mathbf{x}_{t-1} + b_j, \mathbf{Q}^{(j)}) \right] P((b_j', \tilde{\mathbf{A}}_j')')$$

Again, notice $\mathbf{A}_j \mathbf{x}_{t-1} + b_j = ((1, \mathbf{x}_{t-1}')' \otimes \mathbf{I}_p) (b_j', \tilde{\mathbf{A}}_j')'$, then the problem is reduced to Bayesian linear regression. Denote $(1, \mathbf{x}_{t-1}')' \otimes \mathbf{I}_p = \tilde{\mathbf{X}}_{t-1} \in R^{p \times p(1+pJ)}$, $\tilde{\mathbf{X}} = (\tilde{\mathbf{X}}_1', \dots, \tilde{\mathbf{X}}_{T-1}')'$ and $(b_j', \tilde{\mathbf{A}}_j')' = \boldsymbol{\gamma}_j \in R^{p(1+pJ)}$. Then

$$\begin{aligned} \prod_{t=2}^T N(\mathbf{x}_t^{(j)} | \mathbf{A}_j \mathbf{x}_{t-1} + b_j, \mathbf{Q}^{(j)}) &= \prod_{t=2}^T N(\mathbf{x}_t^{(j)} | \tilde{\mathbf{X}}_{t-1} \boldsymbol{\gamma}_j, \mathbf{Q}^{(j)}) \\ &= N((\mathbf{x}_2^{(j)'}', \dots, \mathbf{x}_T^{(j)'}')' | \tilde{\mathbf{X}} \boldsymbol{\gamma}_j, \mathbf{I}_{T-1} \otimes \mathbf{Q}^{(j)}) \end{aligned}$$

By conjugacy:

$$\begin{aligned} P(\boldsymbol{\gamma}_j | \{\mathbf{x}_t^{(j)}\}_{t=1}^T, \mathbf{Q}^{(j)}, \dots) &= N(\boldsymbol{\gamma}_j | \boldsymbol{\mu}_{\boldsymbol{\gamma}_j}, \boldsymbol{\Sigma}_{\boldsymbol{\gamma}_j}) \\ \boldsymbol{\Sigma}_{\boldsymbol{\gamma}_j} &= \left[\boldsymbol{\Sigma}_{ba0}^{-1} + \tilde{\mathbf{X}}' \mathbf{I}_{T-1} \otimes (\mathbf{Q}^{(j)})^{-1} \tilde{\mathbf{X}} \right]^{-1} \\ \boldsymbol{\mu}_{\boldsymbol{\gamma}_j} &= \boldsymbol{\Sigma}_{\boldsymbol{\gamma}_j} \left(\boldsymbol{\Sigma}_{ba0}^{-1} \boldsymbol{\mu}_{ba0} + \tilde{\mathbf{X}}' \mathbf{I}_{T-1} \otimes (\mathbf{Q}^{(j)})^{-1} (\mathbf{x}_2^{(j)'}', \dots, \mathbf{x}_T^{(j)'}')' \right) \end{aligned}$$

c. No constraints on \mathbf{Q} :

$$P((b', \tilde{\mathbf{A}}')' | \{\mathbf{x}_t\}_{t=1}^T, \mathbf{Q}, \dots) \propto \left[\prod_{t=2}^T N(\mathbf{x}_t | \mathbf{A} \mathbf{x}_{t-1} + b, \mathbf{Q}) \right] P((b', \tilde{\mathbf{A}}')')$$

Let $(b', \tilde{\mathbf{A}}')' = \boldsymbol{\gamma} \in R^{p(1+pJ)}$, $(1, \mathbf{x}_{t-1}')' \otimes \mathbf{I}_{pJ} = \tilde{\mathbf{G}}_{t-1} \in R^{Jp \times Jp(1+pJ)}$ and $\tilde{\mathbf{G}} = (\tilde{\mathbf{G}}_1', \dots, \tilde{\mathbf{G}}_{T-1}')'$. By conjugacy, $\boldsymbol{\gamma} | \{\mathbf{x}_t\}_{t=1}^T, \mathbf{Q}, \dots \sim N(\boldsymbol{\mu}_{\boldsymbol{\gamma}}, \boldsymbol{\Sigma}_{\boldsymbol{\gamma}})$.

$$\begin{aligned}\Sigma_{\gamma} &= [\Sigma_{BA_0}^{-1} + \tilde{\mathbf{G}}' \mathbf{I}_{T-1} \otimes (\mathbf{Q})^{-1} \tilde{\mathbf{G}}]^{-1} \\ \mu_{\gamma} &= \Sigma_{\gamma} \left(\Sigma_{BA_0}^{-1} \mu_{BA_0} + \tilde{\mathbf{G}}' \mathbf{I}_{T-1} \otimes (\mathbf{Q})^{-1} (\mathbf{x}'_2, \dots, \mathbf{x}'_T) \right)\end{aligned}$$

(5) Update \mathbf{Q} :

a. Assume \mathbf{Q} be diagonal:

Let $\mathbf{a}'_k \mathbf{x}_{t-1} + b_k = \mu_{x_{kt}}$

$$P(q_k | \{x_{kt}\}_{t=1}^T, b_k, \mathbf{a}'_k, \dots) \propto \left[\prod_{t=2}^T N(x_{kt} | \mu_{x_{kt}}, q_k) \right] IG\left(q_k \left| \frac{\nu_0}{2}, \frac{\nu_0 \sigma_0^2}{2} \right.\right)$$

By conjugacy,

$$P(q_k | \{x_{kt}\}_{t=1}^T, b_k, \mathbf{a}'_k, \dots) = IG\left(q_k \left| \frac{\nu_0 + T - 1}{2}, \frac{\nu_0 \sigma_0^2 + \sum_{t=2}^T (x_{kt} - \mu_{x_{kt}})^2}{2} \right.\right)$$

b. Assume \mathbf{Q} be block-diagonal:

Let $\mathbf{A}_j \mathbf{x}_{t-1} + \mathbf{b}_j = \mu_{x_t^{(j)}}$

$$P(\mathbf{Q}^{(j)} | \{\mathbf{x}_t^{(j)}\}_{t=1}^T, \mathbf{A}_j, \mathbf{b}_j, \dots) \propto \left[\prod_{t=2}^T N(\mathbf{x}_t^{(j)} | \mu_{x_t^{(j)}}, \mathbf{Q}^{(j)}) \right] W^{-1}(\mathbf{Q}^{(j)} | \Psi_0, \nu_0)$$

By conjugacy,

$$\begin{aligned}P(\mathbf{Q}^{(j)} | \{\mathbf{x}_t^{(j)}\}_{t=1}^T, \mathbf{A}_j, \mathbf{b}_j, \dots) \\ = W^{-1}\left(\mathbf{Q}^{(j)} | \Psi_0 + \sum_{t=2}^T (\mathbf{x}_t^{(j)} - \mu_{x_t^{(j)}})(\mathbf{x}_t^{(j)} - \mu_{x_t^{(j)}})', T - 1 + \nu_0\right)\end{aligned}$$

c. No constraints on \mathbf{Q} :

Let $\mathbf{A} \mathbf{x}_{t-1} + \mathbf{b} = \mu_{x_t}$

$$P(\mathbf{Q} | \{\mathbf{x}_t\}_{t=1}^T, \mathbf{A}, \mathbf{b}, \dots) \propto \left[\prod_{t=2}^T N(\mathbf{x}_t | \mu_{x_t}, \mathbf{Q}) \right] W^{-1}(\mathbf{Q} | \Psi_{\mathbf{Q}_0}, \nu_{\mathbf{Q}_0})$$

By conjugacy,

$$\mathbf{Q} | \{\mathbf{x}_t\}_{t=1}^T, \mathbf{A}, \mathbf{b}, \dots \sim W^{-1}\left(\Psi_0 + \sum_{t=2}^T (\mathbf{x}_t - \mu_{x_t})(\mathbf{x}_t - \mu_{x_t})', T - 1 + \nu_0\right)$$