

Association of short-term exposure to PM₁ with hospital admission from total and cause-specific respiratory diseases

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Abstract

Background and Objective: Evidence of short-term exposure to particulate matter with an aerodynamic diameter $\leq 1 \mu\text{m}$ (PM₁) on hospital admission for respiratory diseases (RDs) is limited. We aimed to estimate the associated risk of PM₁ on hospital admissions for RDs.

Methods: In this time-stratified case-crossover study, we assigned cases who had been admitted to hospital for RDs in Guangdong, China between 2016 and 2019. Exposure to PM₁ was assigned on the basis of the patient's residence for each case day and its control days. Conditional logistic regression models and distributed lag nonlinear models were used to quantify the association of PM₁ exposure with hospital admission for RDs at lag 0–1 days.

Results: A total of 408, 658 hospital admissions for total RDs were recorded in the study period. Each $10 \mu\text{g}/\text{m}^3$ increase in PM₁ was significantly associated with a 1.39% (95% confidence interval [CI]: 0.87%–1.91%), 1.97% (95% CI: 1.06%–2.87%) and 1.69% (95% CI: 0.67%–2.71%) increase in odds of hospital admissions for total RDs, chronic obstructive pulmonary disease (COPD) and pneumonia. The excess fraction of hospital admission for total RDs attributable to PM₁ exposure was 6.03%, while 6.59% for COPD and 7.48% for pneumonia. Besides, higher excess fractions were more pronounced for hospital admission of total RDs in older patients (>75 years).

Conclusion: Our results support that PM₁ is associated with increased risks of hospital admissions for RDs. It emphasizes the needs to pay attention to the effects of PM₁ on respiratory health, especially among elderly patients.

KEY WORDS

air pollution, case-crossover study, distributed lag nonlinear models, hospital admission, PM₁, respiratory disease

INTRODUCTION

Numerous epidemiological studies have shown that fine particles (PM_{2.5}) exposure is closely related to a variety of RDs such as COPD,^{1–3} pneumonia,^{4,5} and asthma.^{6,7} Smaller sized particulate matters like PM₁ is considered to be more toxic than PM_{2.5} due to its capability to carry more toxic carcinogens and deposit deeper in the lung.^{8–12} Although accumulating experimental studies have demonstrated the respiratory toxicity of PM₁ by activating

inflammatory responses and oxidative stress,^{13–15} the epidemiological evidence is limited.

To date, only a limited number of studies in Shenzhen, China and Vietnam, specifically reported the positive links between PM₁ exposure and increased hospitalizations for total RDs, while the effect estimates were inconsistent.^{16–18} These inconsistent findings might be due to differences in age and the categories of diseases among study participants. Additionally, the aforementioned studies used data from single monitoring stations or station-based measurements rather than individual exposures, potentially resulting in imprecise and incomparable estimates. Furthermore, it is still unknown whether PM₁ exposure is associated with an

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elevated risk of hospital admission for specific types of RDs such as acute exacerbation of chronic obstructive pulmonary disease (AECOPD), bacterial pneumonia, emphysema and chronic bronchitis, which is very important for early prevention, and standardized treatment of related sensitive population of specific RDs. In addition, all the above studies obtained data on PM₁ exposures from monitoring stations, which are crude for exposure assessment. In 2019, we developed an open-access dataset called ChinaHighAirPollutants (CHAP) using combination of artificial intelligence models with ground-based measurements, satellite remote sensing, model simulations and atmospheric reanalysis.¹⁹ This high spatial resolution dataset provides more accurate estimates of daily levels PM₁ in China.

Inhalation is the main route of ambient particulate matter entering the human body. Compared to external exposure, respiratory tract deposition dose represents the actual exposure level and therefore recommended to be considered when estimating the adverse effect of particles on respiratory health.²⁰ Current studies have supported the internal dose of exposures depends on the size distributions and population characteristics, including gender and age.^{21,22} However, few studies have estimated the deposition of PM₁ in human airways and further assessed the associated risk of hospital admission for respiratory diseases.^{21,22}

To fill the gap, we applied high-resolution grid air pollution dataset (CHAP), combined with multiple-path particle dosimetry (MPPD) model to assess the deposited PM₁. Besides, we conducted a large case-crossover study in Guangdong province to evaluate the risk of total and cause-specific respiratory hospitalizations attributable to PM₁, and to estimate the corresponding excess hospital admissions. Stratified analysis was further performed to explore the impact in hospitalization between age, gender and seasons.

METHODS

Study population

Based on the data from the Guangzhou Health Technology Identification and Human Resources Assessment Center, we included 408,658 patients who lived in Guangdong province, China and were admitted for RDs between 1 January 2016, and 31 December 2019. Besides, we extracted information including the date of admission, date of birth, gender, race and residential address of each patient from the database. The study was approved by the Ethics Committee of Sun Yat-sen University and informed consent was waived (ID: 2020-61).

Outcomes

The study outcome was hospital admissions for total RDs (International Statistical Classification of Diseases and Related Health Problems 10th Revision [ICD-10] codes: J00-J99) and 7 main types of RDs including asthma (ICD-10

SUMMARY AT A GLANCE

This study assessed the impact of PM₁ exposure on hospital admissions for respiratory diseases (RDs) in Guangdong, China. An increase in PM₁ was significantly associated with higher hospital admissions for RDs, especially among older patients, highlighting the need to focus on the effects of PM₁ on respiratory health.

codes: J45-J46), COPD (ICD-10 codes: J41-J44), acute exacerbation of chronic obstructive pulmonary disease (AECOPD) (ICD-10 codes: J44.0-J44.1), chronic bronchitis (ICD-10 codes: J41-J42), emphysema (ICD-10 codes: J43), pneumonia (ICD-10 codes: J12-J18) and bacterial pneumonia (ICD-10 codes: J15).

Study design

We adopted a time-stratified case-crossover study, which is widely applied to evaluate the acute effect of ambient pollution on health.^{23,24} For each patient admitted to hospital, the admission date was defined as the case day, and the corresponding control days were the days on the same day, month and year as the case day. According to the approach, each case day could match for 3 or 4 control days.²⁴ For instance, if a case was admitted to the hospital on September 8, 2023 (Friday), the September 1, 2023 was the case day and other Fridays in September 2023 (i.e., September 15, 22 and 29) were the corresponding control days. As individual-level covariates were less likely to alter during case and corresponding control days, we did not consider them as confounders in our analysis.²⁵

Exposure assessment

Daily ambient PM₁ data used in the study was obtained from the CHAP dataset (spatial resolution: 10 km × 10 km).^{19,26,27} This dataset was produced by artificial intelligence model combination with ground-based measurements, satellite remote sensing, model simulations and atmospheric reanalysis of the big data, which had a full spatiotemporal coverage in China during 2015–2020 and good agreement with measured air pollutant concentrations in China. The cross-validated coefficient of determination (R^2) was 0.82 for PM₁. Based on the address for each patient, the daily 24-h average concentrations of PM₁ on the case and control days were extracted from the CHAP dataset.

The deposited fraction of PM₁ was further calculated by MPPD model based on Wang et al.²² Briefly, PM deposits in different regions of the respiratory system, including the head (mouth cavity and nasopharyngeal

passages), tracheobronchial and pulmonary regions. The deposited fractions in each lung region, along with physiological parameters such as age and sex, significantly impact the calculation of deposition amounts and can be effectively used to predict PM deposition in the human respiratory system.^{8,22,28,29} In this study, we defined deposited PM₁ as the sum of PM₁ concentrations in the tracheobronchial and pulmonary regions, excluding the head region. This is because PM₁ particles are less likely to deposit in the head region due to their small size and are typically cleared by breathing and coughing.³⁰ Therefore, to better capture the particles that are likely to remain and potentially cause health effects, we focused on the TB and pulmonary regions where PM₁ are more likely to deposit and persist. Because the study population was mostly elderly, we only considered notable differences in PM₁ deposition by gender. Therefore, we calculated the corresponding doses of PM₁ deposited in the different groups using the following equation:

$$\text{Male: Deposited PM}_1 (\mu\text{g}/\text{m}^3) = \text{PM}_1 \times 11.57\% \quad (1)$$

$$\text{Female: Deposited PM}_1 (\mu\text{g}/\text{m}^3) = \text{PM}_1 \times 12.83\% \quad (2)$$

Equations (1) and (2) represent the estimated amount of PM₁ deposited in the respiratory tract for males and females, respectively. The coefficients of 11.57% and 12.83% represent the deposition factors for males and females, which were derived based on physiological differences in lung structure and respiratory patterns between the gender.

Covariates

Daily concentrations of multiple gaseous pollutants, including sulfur dioxide (SO₂) (CV-R² = 0.84), nitrogen dioxide (NO₂) (CV-R² = 0.84), carbon monoxide (CO) (CV-R² = 0.80) and ozone (O₃) (CV-R² = 0.87), were collected from the CHAP dataset. Meteorological variables, included 24-h average temperature (°C) and relative humidity (%) were gained from China Meteorological Administration Land Data Assimilation System (CLDAS), with spatial resolution of 0.0625° × 0.0625° and temporal resolution of 1 day. The above covariates were extracted separately for case and control days according to the residential addresses of each case.

Statistical analysis

Exposure-response analysis

Spearman's correlation coefficient was used to evaluate the correlation between air pollutants. We applied conditional logistic regression model and distributed lag models (DLMs) to

investigate short-term effect of PM₁ including ambient PM₁ and deposited PM₁ on total and cause-specific RDs hospitalizations. Percent change in odds ($[\text{odds ratio} - 1] \times 100\%$) of hospital admissions and their 95% confidence interval associated with per 10 μg/m³ increase in ambient/deposited PM₁ on the day of hospitalization and within 1 day before hospitalization.³¹ We separately included ambient PM₁ and deposited PM₁ as continuous variables, and adjusted for temperature and humidity terms to assess exposure-response curves using natural spline functions with 6 degrees of freedom (df) and 3df, respectively.³² The cross-basis function for PM₁ exposure, established using DLM, consists of a linear function of the space of PM₁ exposure and the space of 2 d lag.³³ Besides that, distributed lag nonlinear models (DLNMs) were further applied. A likelihood ratio test was used to analyse the nonlinearity of their association. The lowest level of PM₁ concentration was used as a reference when assessing the association between PM₁ and hospitalizations for RDs.

TABLE 1 Baseline characteristics of the study population, 2016–2019.

Characteristics	Value
Case days, no.	408,658
Control days, no.	1387,734
Cause-specific, no. (%)	
Pneumonia (J12–J18)	121,979 (29.8)
Bacterial pneumonia (J15)	40,729 (10.0)
COPD (J41–J44)	130,266 (31.9)
Chronic bronchitis (J41–J42)	12,614 (3.1)
Emphysema (J43)	1063 (0.3)
AECOPD (J44.0–J44.1)	102,497 (25.1)
Asthma (J45–J46)	8628 (2.1)
Other respiratory diseases	147,785 (36.2)
Gender, no. (%)	
Male	247,517 (60.6)
Female	161,141 (39.4)
Age, year	
Mean (SD)	76.4 (9.33)
≤75 years (no. [%])	184,084 (45.0)
>75 years (no. [%])	224,574 (55.0)
Race, no. (%)	
Han	406,718 (99.5)
Other	994 (0.2)
Unknown	946 (0.2)
Season, no. (%)	
Warm	191,471 (46.8)
Cool	217,187 (53.2)

Abbreviations: AECOPD, acute exacerbation of chronic obstructive pulmonary disease; COPD, chronic obstructive pulmonary diseases.

Excess hospital admission

We further assessed excess hospital admission by estimating excess fraction and number of excess hospital admission. Based on a previously proposed method,^{34,35} excess fraction was calculated using the following equation:

$$\text{Excess fraction}_{x,t} = 1 - e^{\left(-\sum_{l=0}^L \beta_{x_{t-l}}\right)}$$

$$\begin{aligned} \text{Number of excess hospital admission}_{x,t} \\ = N_t \times \text{excess fraction}_{x,t} \end{aligned}$$

In above equation, $\sum \beta_{x_{t-l}}$ corresponds to the logarithm of the cumulative odds ratio of pollutant exposure in DLMs or DLNMs at time t ; N_t is the number of hospital admission for RDs at time t . The excess fraction_{x,t} and number of excess hospital admission_{x,t} on behalf of the amount of time during time $t-l_0, \dots, t-L$ that can be attributed to exposure to pollutant, compared to the reference level of exposure to pollutant, that is, the minimum pollutant exposure level. The empirical confidence intervals (eCIs) were calculated for estimated excess fraction and number of excess hospital admission by Monte Carlo simulation.

Stratified analysis

Gender (male, female), age (≤ 75 years old, > 75 years old), season (Cool season, October to March of next year; Warm season, April–September) were conducted for subgroup analyses, to investigate possible susceptible modification.

Sensitivity analysis

To evaluate the robustness of this finding, we fitted single- and multi-pollutant model to control for co-exposure to air pollutants (NO₂, SO₂, CO and O₃). Principal component analysis (PCA) was used for dimensionality reduction, and several correlated pollutants were transformed into a set of linearly uncorrelated mixed variables, called principal components. Principal components (PC) choosed by Kaiser's eigenvalue greater than 1.0 rule and Cattell's Scree test³⁶ were selected as covariate for the multi-pollutant models, and likelihood ratio test was used to examine the heterogeneity between the single- and multi-pollutant. In this study, all analyses were performed using R version 4.1.1. All p -values were two-sided, and a p -value < 0.05 was considered as statistically significant.

T A B L E 2 Distribution of exposure to ambient air pollutants and meteorological conditions on case days and control days.

	Mean	SD	Percentile						
			5th	25th	50th	75th	95th		
On case days ($n = 408,658$)									
Air pollutant									
PM ₁ , $\mu\text{g}/\text{m}^3$	22.2	11.9	7.9	13.6	19.3	28.1	45.3		
Deposited PM ₁ , $\mu\text{g}/\text{m}^3$	2.7	1.5	0.9	1.6	2.3	3.4	5.5		
CO, mg/m^3	0.89	0.22	0.61	0.74	0.85	0.99	1.33		
SO ₂ , $\mu\text{g}/\text{m}^3$	10.7	4.0	6.0	7.8	9.9	12.7	18.5		
NO ₂ , $\mu\text{g}/\text{m}^3$	46.9	21.5	20.6	32.4	42.9	56.1	87.1		
O ₃ , $\mu\text{g}/\text{m}^3$	93.1	50.6	20.0	55.0	88.0	124.1	185.1		
Meteorological condition									
Temperature, $^\circ\text{C}$	23.2	6.0	12.5	18.9	24.0	28.3	31.1		
Relative humidity, %	76.0	13.9	48.4	68.5	79.1	86.7	92.7		
On control days ($n = 1,387,734$)									
Air pollutant									
PM ₁ , $\mu\text{g}/\text{m}^3$	22.0	11.8	7.8	13.5	19.1	27.9	44.9		
Deposited PM ₁ , $\mu\text{g}/\text{m}^3$	2.7	1.4	0.9	1.6	2.3	3.4	5.4		
CO, mg/m^3	0.89	0.22	0.61	0.74	0.85	0.99	1.32		
SO ₂ , $\mu\text{g}/\text{m}^3$	10.7	4.0	5.9	7.7	9.9	12.6	18.4		
NO ₂ , $\mu\text{g}/\text{m}^3$	46.3	21.1	20.4	32.1	42.5	55.3	85.7		
O ₃ , $\mu\text{g}/\text{m}^3$	93.1	50.4	19.7	55.1	87.9	123.8	184.4		
Meteorological condition									
Temperature, $^\circ\text{C}$	23.2	6.1	12.3	18.8	24.1	28.3	31.2		
Relative humidity, %	76.1	13.7	49.0	68.5	79.0	86.7	92.7		

Abbreviations: CO, carbon monoxide; NO₂, nitrogen dioxide; O₃, ozone; PM₁, particulate matter with aerodynamic diameter $\leq 1 \mu\text{m}$; SO₂, sulfur dioxide.

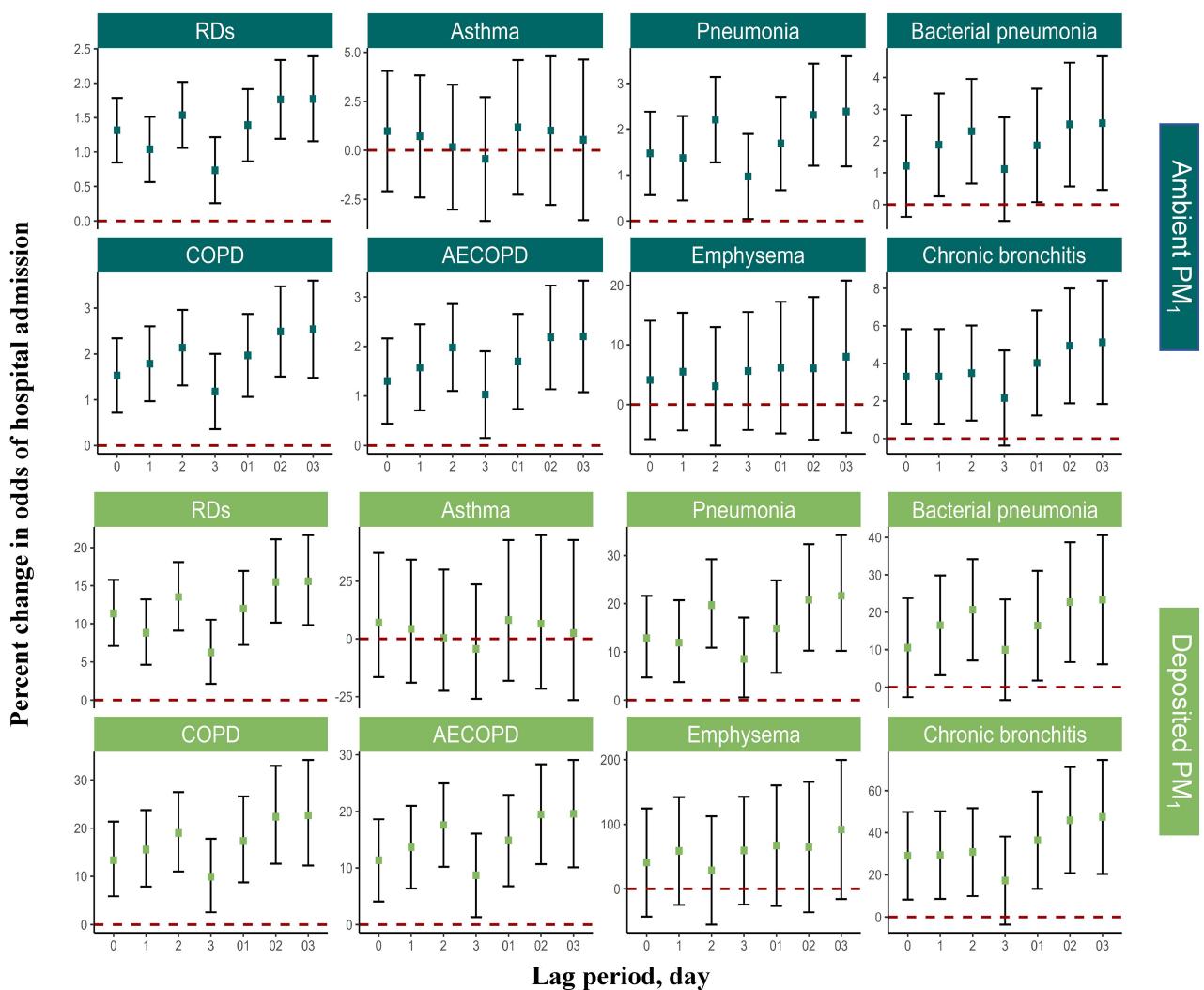


FIGURE 1 Percent change in odds of hospital admission associated with each $10 \mu\text{g}/\text{m}^3$ increase of exposure to ambient/deposited PM₁ in different lag days. AECOPD, acute exacerbation of chronic obstructive pulmonary disease; COPD, chronic obstructive pulmonary diseases; PM₁, particulate matter with aerodynamic diameter $\leq 1 \mu\text{m}$.

RESULTS

Descriptive outcomes

We collected a total of 408,658 hospitalized records for total RDs from 1 January 2016, to 31 December 2019, including 130,266 COPD (31.8%), 121,979 pneumonia (29.8%), 8628 asthma cases (2.1%) and other respiratory disease cases (36.3%). Over half of the patients were male (60.6%), 45.2% were aged ≤ 75 years and 46.8% of admissions were in warm season (Table 1). The distribution of air pollutants and meteorological conditions in case and control days are shown in Table 2. The mean concentrations of PM₁, deposited PM₁, CO, SO₂, NO₂ and O₃ on the case day were $22.2 \mu\text{g}/\text{m}^3$, $2.7 \mu\text{g}/\text{m}^3$, $0.89 \text{ mg}/\text{m}^3$, $10.7 \mu\text{g}/\text{m}^3$, $46.9 \mu\text{g}/\text{m}^3$ and $93.1 \mu\text{g}/\text{m}^3$ respectively, which was similar to those in control days. PM₁ was strongly correlated with deposited PM₁ ($r = 0.99$), while moderately

correlated with SO₂, NO₂ and CO ($r = 0.58$, 0.53 and 0.51 , respectively; all p -values <0.001 ; shown in Figure S1 in the Supporting Information).

Exposure-response analysis

Each $10 \mu\text{g}/\text{m}^3$ increase of ambient PM₁ was significantly associated with 1.39% (95% CI: 0.87%–1.91%), 1.69% (0.67%–2.71%), 1.97% (1.06%–2.87%), 1.90% (0.88%–2.92%), 1.86% (0.08%–3.64%) and 4.02% (1.22%–6.82%) in total RDs, pneumonia, COPD, AECOPD, bacterial pneumonia and chronic bronchitis at lag 01 days, respectively. No significant associations were observed between PM₁ and hospital admission on asthma and emphysema with any different lag days. We also observed that deposited PM₁ had similar short-term effects as did ambient PM₁ (Figure 1). J-shaped associations were observed between ambient PM₁

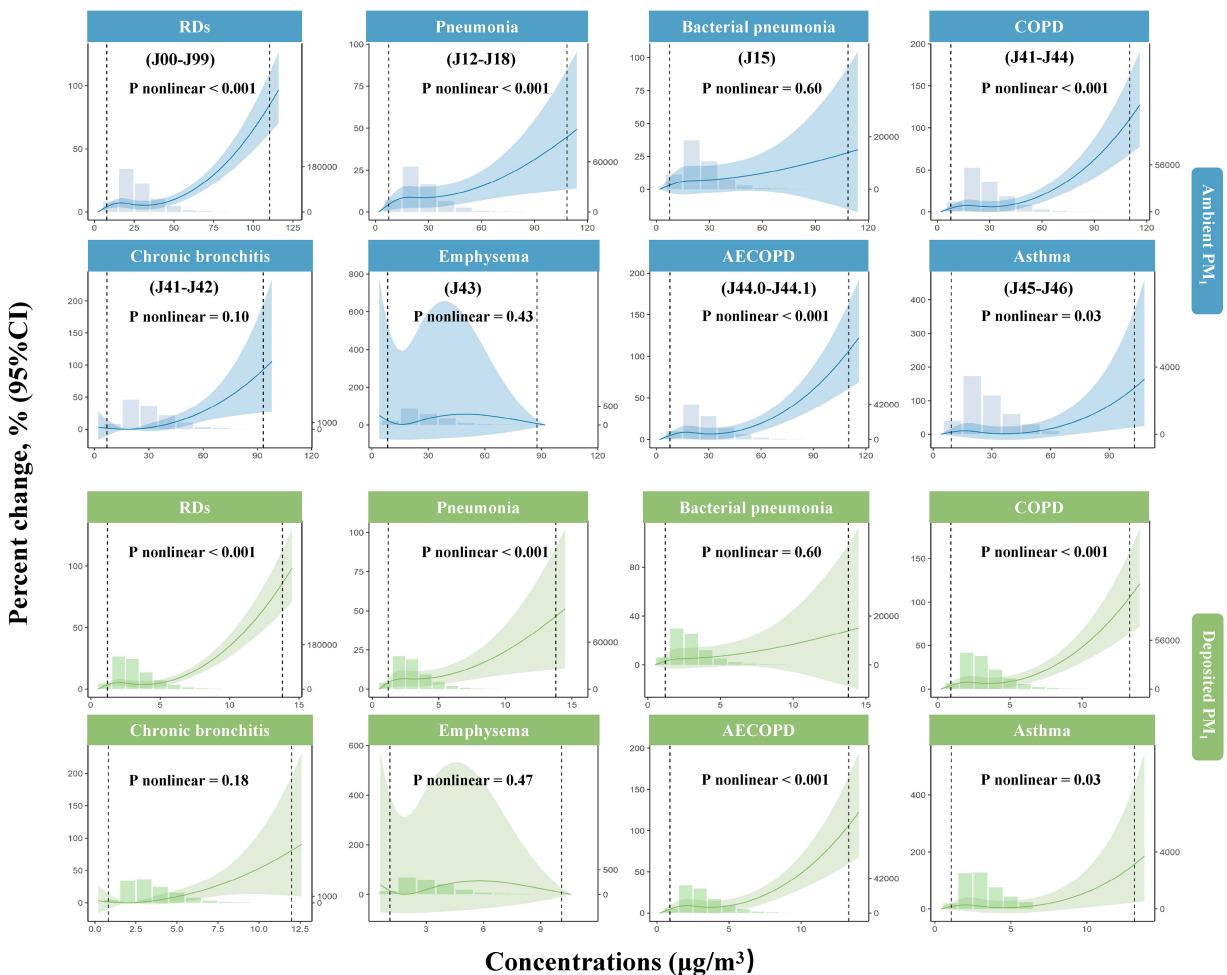


FIGURE 2 Cumulative exposure–response curves for associations of lag 0–1 day exposure to ambient/deposited PM₁ with hospital admission from total and cause-specific respiratory diseases. The solid lines with shaded regions indicate percent changes in odds of hospital admission and their corresponding 95% confidence intervals, while the two dotted lines represent 5% and 95% of the ambient PM₁ exposure, respectively. AECOPD, acute exacerbation of chronic obstructive pulmonary disease; CI, confidence interval; COPD, chronic obstructive pulmonary diseases; PM₁, particulate matter with aerodynamic diameter $\leq 1 \mu\text{m}$.

and deposited PM₁ with hospital admissions for total RDs, COPD, AECOPD, pneumonia and asthma (p for nonlinear <0.05); while the associations of PM₁ with hospital admissions for emphysema, bacterial pneumonia and chronic bronchitis were linear (Figure 2). In multi-pollutant models, we observed the lower estimates of ambient PM₁ and deposited PM₁ on hospital admissions on total RDs, COPD, AECOPD and pneumonia when exposure to gaseous pollutants (NO₂, SO₂, CO and O₃) were adjusted, though the exposure-response curves still represented as J-shapes (Figure S2 in the Supporting Information).

Excess hospital admission

The excess fraction of hospital admission for total RDs attributed to ambient PM₁ was 6.03% (95% eCI: 2.95%–8.7%), which corresponds to 24,631 (95% eCI: 12,044, 35,573) (Table 3). The excess fraction computed using

estimates from the DLNMs was 7.48%, 6.59% and 7.2%, of hospital admission for pneumonia, COPD and AECOPD, respectively, and the estimates from the DLMs was 9.08% of hospital admission for chronic bronchitis. Nevertheless, no significant excess fraction was observed of hospital admission of asthma, bacterial pneumonia and emphysema. The excess fraction remained or slightly changed with adjustment for deposited PM₁.

Stratified analysis

Figure 3 show excess fraction of hospital admission for RDs stratified by age, season and sex. We observed higher excess fractions of hospital admission for total RDs in patients aged >75 , female and in warm season. In addition, the association of PM₁ with hospital admissions for total RDs were J-shaped in subtype except warm season (Figures S3 and S4 in the Supporting Information).

TABLE 3 Excess fraction and number of excess hospital admissions from respiratory disease associated with lag 0–1 day exposure to ambient/deposited PM₁, using DLM or DLNM.

Disease	N	Ambient PM ₁		<i>p</i> for nonlinear ^b	Deposited PM ₁		<i>p</i> for nonlinear ^b
		Excess fraction, % (95% eCI) ^a	Number of excess hospital admissions, N (95% eCI) ^a		Excess fraction, % (95% eCI) ^a	Number of excess hospital admissions, N (95% eCI) ^a	
Total respiratory disease (J00–J99)	408,658	6.03 (2.95, 8.7)	24,631 (12,044, 35,573)	<0.001	4.45 (2.07, 6.62)	18,197 (8450, 27,049)	<0.001
Pneumonia (J12–J18)	121,979	7.48 (2.17, 12.03)	9130 (2650, 14,676)	<0.001	5.75 (1.70, 9.28)	7015 (2072, 11,315)	<0.001
Bacterial pneumonia (J15)	40,728	2.49 (−0.70, 5.58)	1014 (−284, 2274)	0.60	2.2 (−0.61, 4.93)	896 (−248, 2009)	0.60
COPD (J41–J44)	130,266	6.59 (0.94, 11.48)	8589 (1226, 14,952)	<0.001	6.83 (1.00, 11.81)	8899 (1304, 15,389)	<0.001
AECOPD (J44.0–J44.1)	102,496	7.2 (0.87, 12.59)	7332 (892, 12,909)	<0.001	8.12 (2.17, 14.12)	8321 (2229, 14,468)	<0.001
Emphysema (J43)	1062	13.78 (−5.44, 28.63)	146 (−58, 304)	0.43	13.00 (−4.92, 26.97)	138 (−52, 286)	0.47
Chronic bronchitis (J41–J42)	12,613	9.08 (3.21, 14.61)	1146 (405, 1843)	0.10	8.83 (2.86, 14.44)	1114 (361, 1822)	0.18
Asthma (J45–J46)	8628	6.23 (−11.28, 20.28)	537 (−974, 1749)	0.03	8.96 (−9.9, 23.36)	773 (−854, 2015)	0.03

Abbreviations: AECOPD, acute exacerbation of chronic obstructive pulmonary disease; CI, confidence interval; COPD, chronic obstructive pulmonary diseases; eCI: empirical confidence interval.

^aEstimated using the DLNM if the *p* value for nonlinear was <0.05 or using the DLM if the *p* value for nonlinear was >0.05.

^bEstimated using the likelihood ratio test.

DISCUSSION

In this large case-crossover study, short-term effect of PM₁ was significantly associated with an increased risk of RDs hospital admissions. Furthermore, our findings suggest J-shaped associations between PM₁ exposure and hospitalization for total RDs, as well as COPD and pneumonia. The excess fraction of hospital admission for total RDs was 6.03%, which equivalents to 24,631. And the excess fraction was 7.48% and 6.59% of hospital admission for pneumonia and COPD, respectively. The higher excess fraction of admissions for total RDs was observed in patients older than 75 years, females, and in the warmer seasons.

Similarly, three previous case-crossover studies (one conducted in Shenzhen, China and two in Hanoi, Vietnam) also estimated acute impacts of PM₁ on hospital admissions of RDs and reported positive associations,^{16–18} but the estimates are inconsistent. Each 10 µg/m³ increase in PM₁ was associated with a 2.5% elevated risk of total RDs hospitalization among 8934 Vietnamese children,¹⁷ while a 9% increase in 6078 Chinese adults.¹⁶ Another study in Vietnam among 302,345 children, reported a 1.4% increase in risk.¹⁸ All of the results are higher than that in our study, with a 1.39% increase risk of total RDs. Although the average concentration of PM₁ in our study (22.2 µg/m³) was close to that in Shenzhen city (19 µg/m³), the sample size in our study was

larger than that study conducted in Shenzhen, which may have enhanced the robustness of our results. Furthermore, the study in Shenzhen incorporated a substantial proportion of patients with respiratory infections. Nevertheless, it neglected to target chronic non-communicable respiratory diseases such as chronic obstructive pulmonary disease (COPD), which constitute a considerable proportion of respiratory diseases in China. In addition, the possible reasons for the inconsistent estimates between Vietnam and our study could be related to heterogeneity in the characteristics of subjects and PM₁ concentration.

Apart from the significant associations between PM₁ and hospital admission for total RDs, we have new findings that short-term exposure to PM₁ is significantly associated with increased risk of hospital admission for specific RDs (including COPD, pneumonia, AECOPD, chronic bronchitis and bacterial pneumonia), which may be attributed to shared underlying pathological mechanisms associated with PM₁-related RDs. Inhaled PM can be deposited in the trachea, bronchus and distal lungs through oral and nasopharyngeal passages, causing lung injury.³⁷ Fine particles mainly deposit in the surrounding small airways and alveoli. When the particle size is less than 1 µm, the particles can remain in the alveolar region for several years.³⁸ Hence, when assessing the health effects of PM₁, it is crucial to consider the internal exposure dose alongside the potential

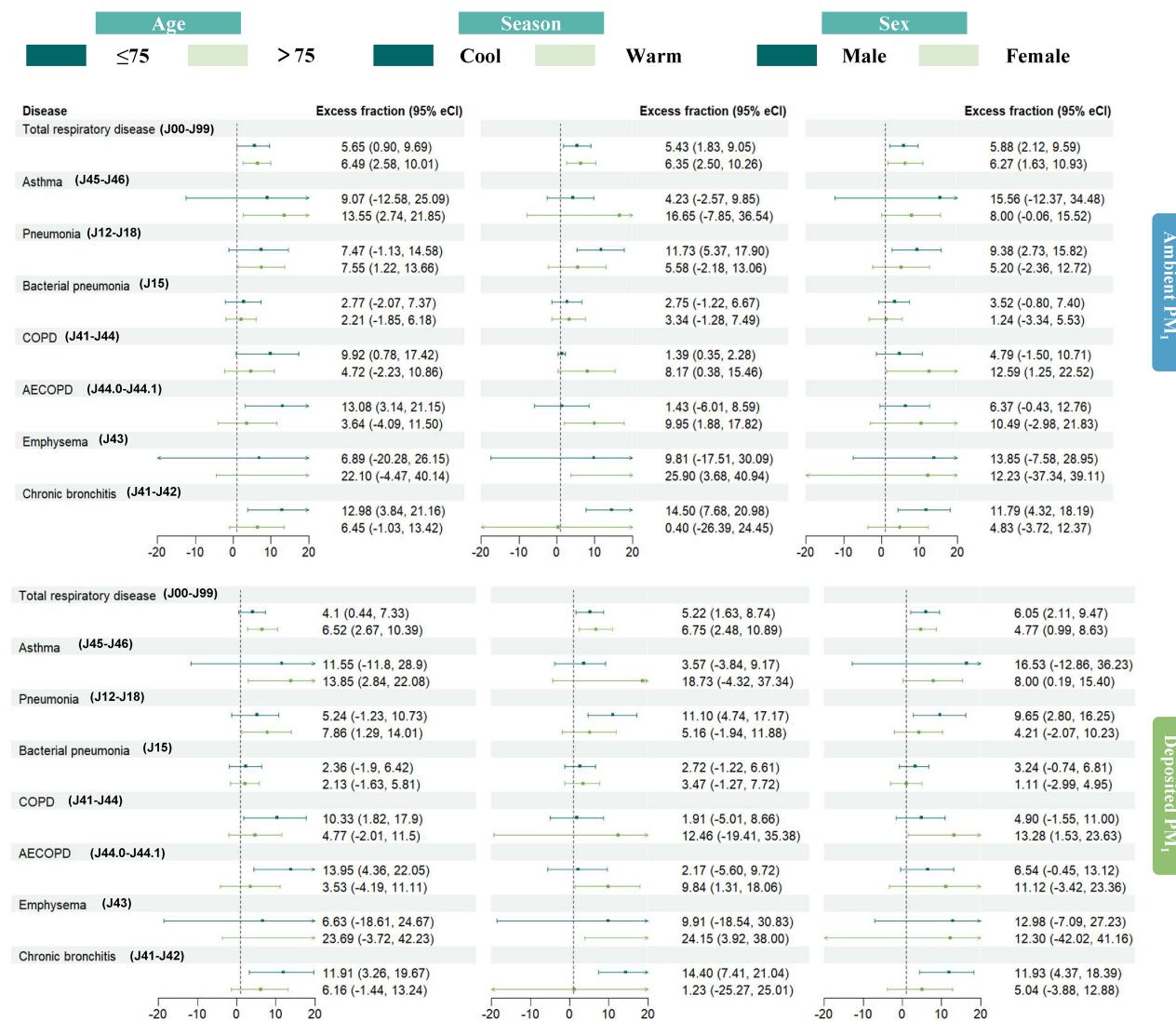


FIGURE 3 Excess fraction of hospital admission from respiratory diseases associated with lag 0–1 day exposure to ambient/deposited PM₁, stratified by age, season and sex. The dotted black lines indicate the estimates of 0. PM₁, particulate matter with aerodynamic diameter $\leq 1 \mu\text{m}$; AECOPD, acute exacerbation of chronic obstructive pulmonary disease; COPD, chronic obstructive pulmonary diseases; eCI: empirical confidence interval; PM₁, particulate matter with aerodynamic diameter $\leq 1 \mu\text{m}$.

toxicity of the chemicals contained in PM₁ itself. Thus, we applied MPPD model to calculate deposited PM₁, combined with individual-level on PM₁ concentrations, to obtain more accurate effect estimate. In addition, few studies have specifically focused on the chemical characteristics of PM₁ (e.g., heavy metals, polycyclic aromatic hydrocarbons, organic compounds and black carbon) and their impact on respiratory disease hospitalizations. We acknowledged that future studies should aim to identify and quantify these toxic components to better understand which chemicals within PM₁ contribute most to the adverse respiratory health effects observed.

Besides that, our study suggested that ambient PM₁ exposure may have a greater adverse impact on hospital admission of RDs in older adults. The elderly may generally have an elevated risk of particulate air pollution related effects due to the degradation of biological functions and

immune systems, as well as underlying chronic diseases.³⁹ Simultaneously, exposure to PM may exacerbate lower respiratory tract inflammation, especially in susceptible populations such as the elderly.⁴⁰ Monitoring ambient PM₁ could be a valuable tool in protecting the elderly, as short-term exposure to elevated PM₁ levels has been shown to have significant adverse effects on their health. It is therefore crucial to minimize outdoor activities for older adults during periods of high PM₁ exposure. Practical interventions, such as improving indoor air quality through the use of air purifiers, issuing health advisories on days with elevated PM₁ levels, and encouraging healthcare providers to adjust treatment plans for high-risk individuals, could help mitigate these effects.

Most existing studies focus on PM_{2.5}, whose relationship with the occurrence and mortality of respiratory diseases is

well-established.^{3,23,41} However, research on PM₁ is severely lacking. Although PM₁ is an important component of PM_{2.5}, the correlation between PM₁ and PM_{2.5} varies depending on environmental factors, such as emission sources and seasonality. In areas dominated by combustion processes (e.g., vehicle emissions or industrial activities), PM₁ can form a larger proportion of PM_{2.5}, resulting in stronger correlations. Resuspended road dust contributes primarily to PM_{2.5} but is largely absent in PM₁.⁴² A European study found that PM₁'s contribution to PM_{2.5} varied significantly between summer (58%) and winter (29%).⁴³ Additionally, the smaller size of PM₁ allows it to penetrate deeper into the lungs, reaching the alveolar region, where it may cause more severe damage than larger particles like PM_{2.5}.^{16,44}

This study has several strengths. First, we investigated associations between PM₁ and hospital admissions for RDs with considerably large sample size, ensuring the adequate statistical power of our results. Second, the PM₁ exposure data we evaluated were extracted from a high spatial grid dataset based on individual addresses, rather than directly using pollution data from monitoring stations, which facilitates the assessment of more accurate individual level exposures. Third, we applied the MPPD model to simulate the deposition of PM₁, and the quantification of PM₁ deposition helps refine the risk assessment. At the same time, the effect of deposited PM₁ on respiratory hospitalization was taken into account in this study, which has not been considered in any study to date. In addition, by applying a time-stratified case-crossover design, we could effectively control potential confounders in all analyses, including individual-level covariates (e.g., age and sex) and behaviour like smoking.

This study has several limitations that should be acknowledged. First, the lack of information on actual inhalation may have led to an underestimation of the true measure of effect. Second, we focused on the elderly population (aged >60) in our study limits the generalizability of the findings. Furthermore, this restricted age range prevented us from accounting for the effects of age when calculating PM₁ deposition using the MPPD model. Third, as this study was conducted in a single province (Guangdong) in China, the generalizability of the results to other regions should be considered with caution.

In conclusion, short-term exposure to PM₁ was significantly associated with an increased risk of hospitalization for RDs, with older adults potentially being a particularly sensitive population. Our finding highlights the importance of reducing ambient PM₁ exposure as a possible strategy to mitigate RDs-related hospitalizations. Furthermore, the exposure-response relationship between PM₁ and hospitalizations for RDs could provide valuable data for the refinement of air quality standards for PM₁.

AUTHOR CONTRIBUTIONS

Chenghui Zhong: Formal analysis (lead); writing – original draft (lead). **Qi Tian:** Formal analysis (equal); writing – original draft (equal). **Jing Wei:** Data curation (lead); formal analysis (equal); writing – review and editing (equal).

Wenfeng Lu: Formal analysis (equal); writing – review and editing (equal). **Ruijun Xu:** Formal analysis (equal); writing – review and editing (equal). **Meiqi Lan:** Formal analysis (equal); writing – review and editing (equal). **Nan Hu:** Formal analysis (equal); writing – review and editing (equal). **Lan Qiu:** Formal analysis (equal); writing – review and editing (equal). **Han Zhang:** Formal analysis (equal); writing – review and editing (equal). **SaiFeng Li:** Formal analysis (equal); writing – review and editing (equal). **Chunxiang Shi:** Formal analysis (equal); writing – review and editing (equal). **Yuewei Liu:** Conceptualization (equal); writing – review and editing (equal). **Yun Zhou:** Conceptualization (equal); writing – review and editing (equal).

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CONFLICT OF INTEREST STATEMENT

None declared.

DATA AVAILABILITY STATEMENT

Data are available in a public, open access repository. The air pollution data are available at (weijing-rs.github.io/product.html). The clinical and meteorological data are not publicly available.

HUMAN ETHICS APPROVAL DECLARATION

The study was approved by the Ethical Committee of School of Public Health, Sun Yat-sen University with a waiver of informed consent.

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