Metropolis Algorithm (Continuous State Space)

Instructor: Dr. Wei Li Scribe: Kyle Beiter

Nov 11th, 2021

Metropolis Algorithm in Continuous Space

We consider a continuous state space, from which we would like to sample a target density $f(\cdot)$ or a kernel \tilde{f} . Our goal is to find a transition function $p(\cdot|\cdot)$ to satisfy the 'detailed balance conditions', i.e: for a proposed x^* ,

$$p(x^*|x)f(x) = p(x|x^*)f(x^*).$$

If we can find such a p, then the long run distribution of the ergodic markov chain with transition distribution $p(\cdot|\cdot)$ is $f(\cdot)$.

Method

Choose an arbitrary transition density $q(\cdot|\cdot)$, without requiring that the detailed balance conditions hold. Then use an acceptance probability on proposed samples as

$$\alpha(x^*|x) = \min\Big(1, \frac{q(x|x^*)f(x^*)}{q(x^*|x)f(x)}\Big).$$

The goal of this is to regulate the flow of probabilities between the current state and the proposed state. So given that our current state is x then the probability of a particular state being proposed and accepted is $\alpha(x^*|x)q(x^*|x)$. Similarly, we can calculate the probability that any proposal is accepted as $a(x) := \int \alpha(x^*|x)p(x^*|x)dx^*$. A move is not accepted with probability r(x) := 1 - a(x), in which case we set $x^* = x$. The (effective) transition density is then

$$p(x^*|x) = \alpha(x^*|x)q(x^*|x) + r(x)\delta_x(x^*)$$

where $\delta_x(x^*) = 1$ if $x^* = x$ and 0 otherwise. Then since we are in continuous state space, this can be generalized to

$$\int_A p(x^*|x)dx^* = p(A|x) = \int_A \alpha(x^*|x)q(x^*|x)dx^* + r(x)\delta_A(x)$$

where $\delta_A(x) = 1$ if $x \in A$ and 0 otherwise.

We can show that the (effective) transition density indeed satisfy the detailed balance condition.

Proof:

Algorithm

- 1. Given a target $f(\cdot)$ and a chosen $q(\cdot|\cdot)$, generate x^* from $q(\cdot|x^{(s)})$
- 2. Compute $\alpha(x^*|x^{(s)}) = \min(1, \frac{q(x^{(s)}|x^*)f(x^*)}{q(x^*|x^{(s)})f(x^{(s)})})$
- 3. Generate $u \sim \text{unif}(0,1)$
 - i. if $u < \alpha(x^*|x^{(s)})$, set $x^{(s+1)} = x^*$
 - ii. if $u > \alpha(x^*|x^{(s)})$, set $x^{(s+1)} = x^{(s)}$

Remark

A necessary condition for the algorithm to work is $\operatorname{supp}(f) \subseteq \operatorname{supp}(q(\cdot|x))$ for all $x \in \operatorname{supp}(f)$. A minimal necessary condition would be $\operatorname{supp}(f) \subseteq \bigcup_{x \in \operatorname{Supp}(f)} [\operatorname{supp}(q(\cdot|x))]$

Choice of Transition Density

- 1. Classical Choice (Random Walk): Choose $q(x^*|x) = \tilde{q}(x^* x)$ where \tilde{q} is symmetric about 0. For example $x^* = x^{(s)} + \epsilon^{(s)}$, where $\epsilon^{(s)} \sim \text{unif}(-1,1)$ or $\epsilon^{(s)} \sim N(0,\tau^2)$. Then we would have $x^* \sim \text{unif}(x^{(s)} 1, x^{(s)} + 1)$, or $x^* \sim N(x^{(s)}, \tau^2)$ respectively. Then due to the symmetry about 0, $q(x^*|x) = q(x|x^*)$, so the acceptance probability simplifies to $\alpha(x^*|x) = \min(1, \frac{f(x^*)}{f(x)})$.
- 2. Independent Metropolis Hastings: Choose $x^* \sim q(\cdot) = q(\cdot|x)$ independent of the current value $x^{(s)}$. For example $q \sim N(0,1)$ throughout.

Practicalities

1. Burn-in

- i. Run algorithm until some iteration B for which Markov Chain appears to reach stationarity.
- ii. Run algorithm B' more times and keep $\{x^{B+1}, \dots, x^{B+B'}\}\$
- iii. Use empirical distribution of $\{x^{B+1}, \dots, x^{B+B'}\}$ to approximate the density f.
- 2. **Thinning:** After burning to discard first B observations, then thin the chain by including every kth draw.
- 3. **Trace Plot**: The plot of the values generated from the metropolis algorithm against the iteration number. Want to check:
 - a. The chain doesn't get stuck in a certain area of space.
 - b. The chain can move relatively fast. Between 30%-60% acceptance rate is ideal.
 - c. The chain moves at an appropriate speed. Proposals are not too bold (standard deviation of p is too high), however they are large enough to be consequential (standard deviation of p is not too small)
- 4. "ACF" Autocorrelation Function: Calculate the "lag j sample covariance" as

$$\hat{\gamma}_j = \frac{1}{T} \sum_{t=j+1}^T (x^{(t)} - \bar{x})(x^{(t-j)} - \bar{x})$$

and the "lag j sample autocorrelation" as $\hat{\rho}_j = \frac{\hat{\gamma}_j}{\hat{\gamma}_0}$. Then plotting autocorrelation $\hat{\rho}_j$ v.s. the lag index (j), relatively high values would indicate a high degree of correlation between draws and therefore, slow mixing.