

MAT 653: Statistical Simulation

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Classical Monte Carlo Integration

The general problem is to evaluate

$$E_f[h(X)] = \int_{\mathcal{X}} h(x)f(x) dx,$$

where h is some known function, X is r.v. $\sim f$ and $\mathcal{X} = \text{supp}(f)$. In the following discussion, the underlying \mathcal{X} is implicitly assumed.

The principle of Monte Carlo method for evaluating $E_f[h(X)]$ is to generate an i.i.d. random sample sequence (X_1, \dots, X_n) from f and propose to estimate it by empirical average

$$\bar{h}_n = \frac{1}{n} \sum_{j=1}^n h(X_j)$$

this is called the **Monte Carlo estimator** for $E_f[h(X)]$. This works because by Law of Large numbers, $\bar{h}_n \rightarrow E_f[h(X)]$.

Properties of estimator \bar{h}_n

1. $E_f(\bar{h}_n) = \mathbb{E}_f[h(X)]$

Proof: $\mathbb{E}_f(\bar{h}_n) = \frac{1}{n} \sum_{j=1}^n \mathbb{E}_f[h(X_j)] = \frac{1}{n} \sum_{j=1}^n \mathbb{E}_f[h(X)] = E_f[h(X)]$

2. $\text{Var}_f(\bar{h}_n) = \frac{1}{n} \text{Var}_f[h(X)]$

Proof:

$$\begin{aligned} \text{Var}_f(\bar{h}_n) &= \text{Var}_f\left[\frac{1}{n} \sum_{j=1}^n h(X_j)\right] = \frac{1}{n^2} \text{Var}_f\left[\sum_{j=1}^n h(X_j)\right] \\ &\stackrel{(*)}{=} \frac{1}{n^2} \sum_{j=1}^n \text{Var}_f[h(X_j)] = \frac{1}{n^2} \sum_{j=1}^n \text{Var}_f[h(X)] = \frac{1}{n} \text{Var}_f[h(X)] \end{aligned}$$

where $(*)$ holds since $X_j \stackrel{iid}{\sim} f$, $\text{Cov}(X_i, X_j) = 0$ ($i \neq j$)

And $\text{Var}_f(\bar{h}_n)$ can also be estimated from (X_1, \dots, X_n) by $\frac{1}{n^2} \sum_{j=1}^n \left[h(X_j) - \frac{1}{n} \sum_{i=1}^n h(X_i)\right]^2$.

Proof:

$$\begin{aligned}
 \text{Var}_f(\bar{h}_n) &= \frac{1}{n} \text{Var}_f[h(X)] \\
 &= \frac{1}{n} \mathbb{E}_f\{[h(X) - \mathbb{E}_f(h(X))]^2\} \quad (\text{where } \mathbb{E}_f(h(X)) \text{ can be estimated by } \bar{h}_n) \\
 &\leftarrow \frac{1}{n} \mathbb{E}_f\{[h(X) - \bar{h}_n]^2\} \quad (\text{now estimate } \mathbb{E}_f[h(X) - \bar{h}_n]^2) \\
 &\leftarrow \frac{1}{n^2} \sum_{j=1}^n [h(X_j) - \bar{h}_n]^2 \\
 &= \frac{1}{n^2} \sum_{j=1}^n \left[h(X_j) - \frac{1}{n} \sum_{i=1}^n h(X_i) \right]^2
 \end{aligned}$$

Example: For $\mathcal{N}(0, 1)$, want to estimate c.d.f $\phi(t) = \int_{-\infty}^t \frac{1}{\sqrt{2\pi}} e^{-x} dx$.

Say $f(x) = \frac{1}{\sqrt{2\pi}} e^{-x}$, we can rewrite $\phi(t) = \int_{-\infty}^t \frac{1}{\sqrt{2\pi}} e^{-x} dx = \int_{-\infty}^{\infty} \mathbf{1}_{(-\infty, t)}(x) f(x) dx$.

Therefore, we can generate a sample of size n , $\{X_i\}_{i=1}^n \stackrel{iid}{\sim} f$, then use $\bar{h}_n = \frac{1}{n} \sum_{i=1}^n \mathbf{1}_{(-\infty, t)}(X_i)$ to estimate $\phi(t)$.

Example (Drawback of classical M.C. integration): If $Z \sim \mathcal{N}(0, 1)$, and we are asked to evaluate $P(Z > 4.5)$.

$$P(Z > 4.5) = \mathbb{E}_f[\mathbf{1}_{(4.5, \infty)}(Z)]$$

Even though we can use classical M.C. integration: generate a sample of size n , $\{Z_i\}_{i=1}^n \stackrel{iid}{\sim} \mathcal{N}(0, 1)$, then use $\bar{h}_n = \frac{1}{n} \sum_{i=1}^n \mathbf{1}_{(4.5, \infty)}(Z_i)$ to estimate $P(Z > 4.5)$. But actually, simulating $\{Z_i\}_{i=1}^n \stackrel{iid}{\sim} \mathcal{N}(0, 1)$ probably only produce a hit $Z_i \in (4.5, \infty)$ once in large amount of iterations. However, using importance sampling can solve this issue.

Importance Sampling

Given another p.d.f g s.t. $\text{supp}(g) \supseteq \text{supp}(f)$.

$$\mathbb{E}_f[h(X)] = \int h(x) f(x) dx = \int h(x) \frac{f(x)}{g(x)} g(x) dx = \mathbb{E}_g \left[h(X) \frac{f(X)}{g(X)} \right]$$

Now, we can generate a sample of size n , $\{X_i\}_{i=1}^n \stackrel{iid}{\sim} g$, then use $\bar{h}_n = \frac{1}{n} \sum_{i=1}^n h(X_i) \frac{f(X_i)}{g(X_i)}$ to estimate $\mathbb{E}_f[h(X)] = \int h(x) f(x) dx$.

Where g is called **importance function**, $\frac{f(X_i)}{g(X_i)}$ is called **importance weight** for X_i and $\left(X_i, \frac{f(X_i)}{g(X_i)}\right)$ is called **importance sample**. $\frac{1}{n} \sum_{i=1}^n h(X_i) \frac{f(X_i)}{g(X_i)}$ is called **importance sampling estimator** for $\mathbb{E}_f[h(X)]$.

Example (continue the previous example): recall that $Z \sim \mathcal{N}(0, 1)$, and we are asked to evaluate $P(Z > 4.5)$.

Now we can take g to be density function of $\exp(1)$ (right) truncated at 4.5:

$$\begin{aligned} g(y) &= P(\exp(1) = y | \exp(1) > 4.5) = \frac{P(\exp(1) = y, \exp(1) > 4.5)}{P(\exp(1) > 4.5)} = \frac{P(\exp(1) = y)}{P(\exp(1) > 4.5)} \\ &= \frac{e^{-y}}{1 - \int_0^{4.5} e^{-x} dx} = \frac{e^{-y}}{1 - (1 - e^{-4.5})} = e^{-(y-4.5)} \quad (y > 4.5) \end{aligned}$$

Now we can generate a sample of size n , $\{Y_i\}_{i=1}^n \stackrel{iid}{\sim} g$, recall $f(x) = \frac{1}{\sqrt{2\pi}} e^{-x}$, then the importance sampling estimator for $P(Z > 4.5) = \mathbb{E}_f[\mathbf{1}_{(4.5, \infty)}(Z)]$ becomes

$$\frac{1}{n} \sum_{i=1}^n \mathbf{1}_{(4.5, \infty)}(Y_i) \frac{f(Y_i)}{g(Y_i)} = \frac{1}{n} \sum_{i=1}^n \mathbf{1}_{(4.5, \infty)}(Y_i) \frac{e^{-\frac{Y_i^2}{2} + Y_i - 4.5}}{\sqrt{2\pi}} = \frac{1}{n} \sum_{i=1}^n \frac{e^{-\frac{Y_i^2}{2} + Y_i - 4.5}}{\sqrt{2\pi}}$$

where here $\mathbf{1}_{(4.5, \infty)}(Y_i) = 1$ since $\{Y_i\}_{i=1}^n \stackrel{iid}{\sim} g$.

Remark: we need $\text{supp}(g) \supseteq \text{supp}(f)$ to make sure the value of importance weight $\frac{f(X_i)}{g(X_i)}$ is meaningful. Actually here having a weaker condition $\text{supp}(g) \supseteq \text{supp}(h \times f)$ suffices since we have the fraction $\frac{h(X_i)f(X_i)}{g(X_i)}$.

Self-normalized version of importance sampling

Again, we want to evaluate $\mathbb{E}_f[h(X)]$. But now $f \propto \tilde{f}$, $g \propto \tilde{g}$. Say $f = c_0 f_0$ where f_0 is the unnormalized p.d.f, c_0 is the normalizing constant. And $g = c_1 g_1$, where g_0 is the unnormalized p.d.f, c_1 is the normalizing constant.

Now, suppose we know f_0 and g_0 , but possibly not c_0 and c_1 . And suppose we know how to generate random sample from g .

$$\begin{aligned} \mathbb{E}_f[h(X)] &= \int h(x)f(x) dx \\ &= \frac{\int \frac{h(x)f(x)}{g(x)} g(x) dx}{\int \frac{f(x)}{g(x)} g(x) dx} \\ &= \frac{\int \frac{f_0(x)}{g_0(x)} h(x) g(x) dx}{\int \frac{f_0(x)}{g_0(x)} g(x) dx} \quad (\text{now } w(x) := \frac{f_0(x)}{g_0(x)}) \\ &= \frac{\int w(x) h(x) g(x) dx}{\int w(x) g(x) dx} = \frac{\mathbb{E}_g[w(x)h(x)]}{\mathbb{E}_g[w(x)]} \end{aligned}$$

Then we can generate a sample of size n , $\{X_i\}_{i=1}^n \stackrel{iid}{\sim} g$, and use $\hat{\mu} = \frac{\frac{1}{n} \sum_{i=1}^n w(X_i)h(X_i)}{\frac{1}{n} \sum_{i=1}^n w(X_i)}$ to estimate $\mu = \mathbb{E}_f[h(X)]$.

Here $\hat{\mu}$ is called **Self-normalized importance sampling estimator** for μ .

The reason why $\hat{\mu} \rightarrow \mu$.

$$\hat{\mu} = \frac{\frac{1}{n} \sum_{i=1}^n \frac{f_0(X_i)}{g_0(X_i)} h(X_i)}{\frac{1}{n} \sum_{i=1}^n \frac{f_0(X_i)}{g_0(X_i)}} = \frac{\frac{1}{n} \sum_{i=1}^n \frac{f(X_i)}{g(X_i)} h(X_i)}{\frac{1}{n} \sum_{i=1}^n \frac{f(X_i)}{g(X_i)}} \xrightarrow{LLN} \frac{\mathbb{E}_g\left[\frac{f(X)}{g(X)} h(X)\right]}{\mathbb{E}_g\left[\frac{f(X)}{g(X)}\right]}$$

The last term equals to $\int \frac{h(x)f(x)}{g(x)} g(x) dx / \int \frac{f(x)}{g(x)} g(x) dx = \int h(x)f(x) dx = \mathbb{E}_f[h(X)] = \mu$.

Self-normalize Importance Sampling

Target: $E_f(h(x))$

$f \propto f_0$, g is importance pdf, and $\text{supp}(f) \subset \text{supp}(g)$.

Self-normalized importance sampling $\implies (x_i, \frac{f_0(x_i)}{g_0(x_i)})$ importance sample. It turns out this can be recycled by multinomial resampling into a sample that is from f .

Step 1: Sample $x_i \sim g$, obtain $(x_i, \frac{f_0(x_i)}{g_0(x_i)})_{i=1}^n$

Step 2: importance weights $w_i = \frac{f_0(x_i)}{g_0(x_i)}$. Let $\hat{w}_i = \frac{\frac{1}{n} w_i}{\sum_j \frac{1}{n} w_j}$. Notice that $\hat{w}_i \in (0, 1)$, $\sum \hat{w}_i = 1$

Step 3: draw a random sample of size m with replacement from x_1, \dots, x_n with weighted probabilities by $\hat{w}_1, \dots, \hat{w}_n$: that is for $k = 1, 2, \dots, m$,

$$X_k^* = \begin{cases} x_1 & \text{with prob} = \hat{w}_1 \\ x_2 & \text{with prob} = \hat{w}_2 \\ \vdots & \\ x_n & \text{with prob} = \hat{w}_n \end{cases}$$

It turns out that $(X_1^*, X_2^*, \dots, X_m^*)$ is a random sample from f .

Remark: for this sampling to work, the target sample size m should be no more than 10% of the original sample size n .

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