Continuous RV

- Recall the definition of pmf for a discrete rv. P(X=x). Can we extend this definition to continuous rv's?
- Uniform random variable: X is equally likely to be any number on [0,1], what is the probability P(X=0.5)?
- The probability model for a continuous random variable assigns probabilities to intervals of outcomes rather than to individual outcomes.
- The probability model of X is often described by a smooth curve, which is the probability density function (pdf) of X.

The CDF

- Although the idea of pmd does not extend to the continuous rv's, the idea of cdf still works.
- The cumulative distribution function (cdf) F(x) for a continuous rv X is defined for every number x by

$$F(x) = P(X \le x) = \int_{-\infty}^{x} f(y)dy$$

- F(x) is in fact the probability that a rv X is smaller than x. F(x) increases smoothly as x increases. $F(-\infty) = 0$, $F(+\infty) = 1$.
- It is easy to compute probabilities using F(x).
 - P(X > a) = 1 F(a)
 - $P(a \le X \le b) = F(b) F(a)$

pdf from cdf

- If X is a continuous rv with pdf f(x) and cdf F(x), then at every x at which the derivative F'(x) exists, F'(x) = f(x). f(x) is often a smooth curve, which is the probability density function (pdf) of X.
- Let p be a number between 0 and 1. The (100p)th percentile (quantile) of the distribution of a continuous rv X, denoted by $\eta(p)$, is defined by

$$p = F(\eta(p)) = \int_{-\infty}^{\eta(p)} f(y)dy$$

• The median of a continuous distribution, denoted by $\tilde{\mu}$, is the 50th percentile, so $\tilde{\mu}$ satisfies .5 = F($\tilde{\mu}$). That is, half the area under the density curve is to the left of $\tilde{\mu}$ and half is to the right of $\tilde{\mu}$.

PDF

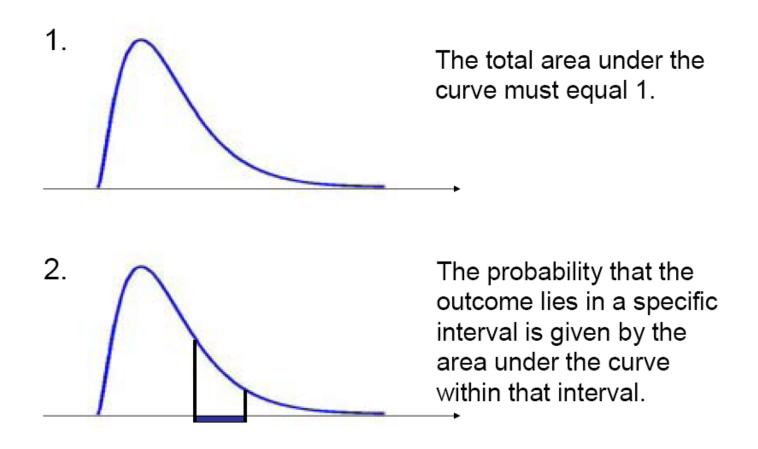
• The probability density function (pdf) of a continuous rv X is a function f(x) such that for any two numbers a and b with $a \le b$,

$$P(a \le X \le b) = \int_{a}^{b} f(x)dx$$

The graph of f(x) is often referred to as the *density curve*.

- This means the area under the density curve represents probability!
- Note that $0 \le f(x)$ for all x.
- f(x)dx can be treated as P(X=x)!

Properties of PDF



Uniform Distribution

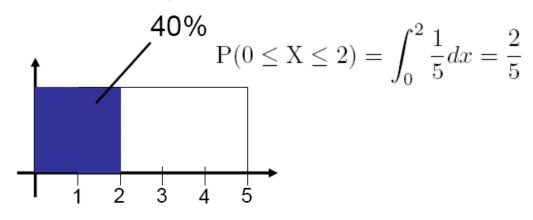
 A continuous rv X is said to have a uniform distribution on the interval [A, B] if the pdf of X is

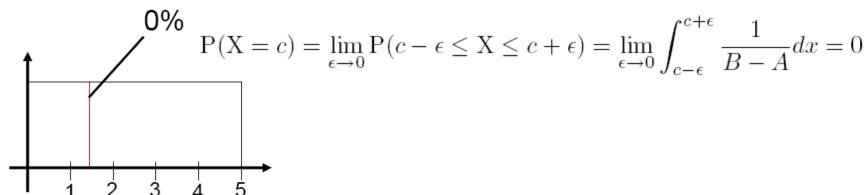
$$f(x; A, B) = \begin{cases} \frac{1}{B-A} & A \le x \le B\\ 0 & \text{otherwise} \end{cases}$$

- Verify that this is a proper pdf.
 - 1. $f(x) \ge 0$ for all x.
 - 2. Area under f(x) should be equal to 1.

Example

Ex. Suppose a bus arrives equally likely at any time between 7:00 – 7:05 AM. What is the probability it arrives sometime between 7:00 – 7:02 AM?





Expected Values

- Notice that the pdf f(x) of a continuous distribution is actually playing the role of pmf p(x) of a discrete distribution.
- Recall that the expected value of a discrete distribution is calculated by

$$\mu_X = \mathcal{E}(\mathcal{X}) = \sum_{x \in D} x \cdot p(x)$$

 Therefore, similarly we can define the expected value of a continuous distribution by

$$\mu_X = E(X) = \int_{-\infty}^{\infty} x \cdot f(x) dx$$

 Take advantage of the symmetry of particular distributions, when calculating expectations.

Variance

- With a similar argument as in the discrete case, we can also define the expectation of a function of a continuous rv as well as the variance of a continuous rv.
- Proposition: if X is a continuous rv with pdf f(x) and h(X) is any function of X, then

$$E[h(X)] = \int_{-\infty}^{\infty} h(x) \cdot f(x) dx$$

As a special case of the above proposition, the variance of X is defined by

$$\sigma_X^2 = \operatorname{Var}(X) = \operatorname{E}(X - \operatorname{E}(X))^2 = \int_{-\infty}^{\infty} (x - \mu_X)^2 \cdot f(x) dx$$

The standard deviation (SD) of X is $\sigma_X = \sqrt{\mathrm{Var}(\mathrm{X})}$.

Examples

<u>Ex.</u> Prove for continuous rv X, as in the discrete case, that $Var(X) = E(X^2) - [E(X)]^2$.

Ex. If a stick of length 1 is broken at random into two pieces. What is the expected length of the longer piece?

Properties

- Some properties of mean and variance hold in the continuous case in a similar way as in the discrete case.
- For example, under linear transformation of X, we have
- 1. E(aX+b) = aE(X) + b
- 2. $Var(aX+b) = a^2Var(X)$
- Exercise: prove the above formulas rigorously!

Uniform RV

- We call a uniform rv U a standard uniform, if and only if U ~ uniform on [0,1]
- For a standard uniform rv U, we can easily calculate,

$$E(U) = \int_0^1 x \cdot 1 dx = \frac{1}{2}$$

$$E(U^2) = \int_0^1 x^2 \cdot 1 dx = \frac{1}{3}$$

$$Var(U) = E(U^2) - [E(U)]^2 = \frac{1}{12}$$

General Uniform

- Note that a general case of uniform distribution X on [A, B] can be treated as a linear transform of a standard uniform, i.e., X = (B A)U + A.
- Proposition:

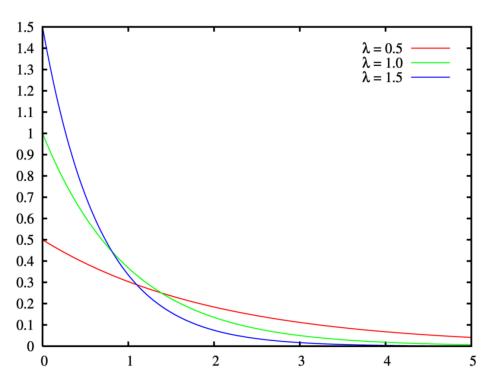
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If X is a continuous uniform rv on [A, B], then E(X) = (B + A)/2, Var(X) = (B - A)^2/12
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• R command: dunif(x, min=0, max=1), punif(q, min=0, max=1), qunif(p, min=0, max=1).

The Exponential Distribution

• X is said to have an exponential distribution with parameter λ (λ >0) if the pdf of X is

$$f(x;\lambda) = \begin{cases} \lambda e^{-\lambda x} & x \ge 0\\ 0 & \text{otherwise} \end{cases}$$



More on Exponential

Note that an exponential rv X can only take positive values. And the cdf of X is

$$F(x;\lambda) = \begin{cases} \int_0^x \lambda e^{-\lambda y} dy = 1 - e^{-\lambda x} & x \ge 0\\ 0 & x < 0 \end{cases}$$

- Thus $P(X>x) = 1 F(x; \lambda) = e^{-\lambda x}$
- Proposition: (proof?)

If X is an exponential rv with parameter λ , then $E(X) = 1/\lambda$, $Var(X) = 1/\lambda^2$

• R command: dexp(x, lamda=1), pexp(q, lamda=1), qexp(p, lamda=1).