## **Population proportion**

- Let *p* denote the proportion of individuals or objects in a population who possess a specified property (probability of success). In order to make inference about p, naturally we would look at the sample proportion, which is X/n. X is the number of S's in the sample. In practice, X normally should follow a hyper-geometric distribution, which can often be approximated by a binomial distribution, and when n is large, it can further be approximated by a normal distribution.
- We first consider large sample tests.

# Large-sample tests

Thanks to the Central Limit Theorem, we have

$$Z = \frac{\hat{p} - p_0}{\sqrt{p_0(1 - p_0)/n}} \sim N(0, 1)$$

under the null hypothesis.

- Thus the rejection region is determined by
- 1.  $H_a$ :  $p > p_0$ :  $Z > z_\alpha$
- 2.  $H_a$ :  $p < p_0$ :  $Z < -z_\alpha$
- 3.  $H_a$ :  $p \neq p_0$ :  $Z > z_{\alpha/2}$  or  $Z_0 < -z_{\alpha/2}$
- The test procedures are valid provided that  $np_0 \ge 10$  and  $n(1-p_0) \ge 10$ .

# **Example**

Ex. (Defective rate cont.) A factory claims that less than 10% of the components they produce are defective. A consumer group is skeptical of the claim and checks a random sample of 300 components and finds that 39 are defective. Is there evidence that 10% of all components made at the factory are defective?

$$H_0: p = 0.10$$
  $H_a: p > 0.10$ 

$$\hat{p} = \frac{39}{300} = 0.13$$
  $Z = \frac{0.13 - 0.1}{\sqrt{0.1(1 - 0.1)/300}} = 1.72$ 

 $z_{0.05}$  = 1.645. Z >  $z_{0.05}$ , thus we would reject  $H_0$  at level  $\alpha$ =0.05.

# Type II error

• We consider the upper tail test with rejection region  $Z > z_{\alpha}$  and try to calculate the type II error associated with this test.

$$\beta(p') = P(H_0 \text{ is not rejected when } p = p')$$

$$= P\left(\frac{\hat{p} - p_0}{\sqrt{p_0(1 - p_0)/n}} \le z_\alpha | p = p'\right)$$

$$= P\left(\frac{\hat{p} - p'}{\sqrt{p_0(1 - p_0)/n}} \le z_\alpha + \frac{p_0 - p'}{\sqrt{p_0(1 - p_0)/n}} | p = p'\right)$$

$$= P\left(\frac{\hat{p} - p'}{\sqrt{p'(1 - p')/n}} \le \frac{z_\alpha \sqrt{p_0(1 - p_0)/n}}{\sqrt{p'(1 - p')/n}} + \frac{(p_0 - p')}{\sqrt{p'(1 - p')/n}} | p = p'\right)$$

$$= \Phi\left(\frac{p_0 - p' + z_\alpha \sqrt{p_0(1 - p_0)/n}}{\sqrt{p'(1 - p')/n}}\right)$$

Formulas for different alternatives can be derived in a similar way. (Book p.308)

# Determining sample size

• If we specify a particular alternative p' and specify a  $\beta$  value that can be tolerated (e.g. 0.1). Then from

$$\beta = \Phi\left(\frac{p_0 - p' + z_\alpha \sqrt{p_0(1 - p_0)/n}}{\sqrt{p'(1 - p')/n}}\right) \Longrightarrow -z_\beta = \frac{p_0 - p' + z_\alpha \sqrt{p_0(1 - p_0)/n}}{\sqrt{p'(1 - p')/n}}$$

 Therefore, in order to achieve the specified type I and type II error, one has to have a sample size of at least

$$n = \left(\frac{z_{\alpha}\sqrt{p_0(1-p_0)} + z_{\beta}\sqrt{p'(1-p')}}{p' - p_0}\right)^2$$

- For two sided test, we have to change  $z_{\alpha}$  to  $z_{\alpha/2}$  in the above formula.
- Difference between the sample size calculation formula in chapter 7 and the one above.

# **Example**

Ex. A package-delivery service advertises that at least 90% of all packages brought to its office by 9 a.m. for delivery in the same city are delivered by noon that day. Let p denote the true proportion of such packages that are delivered as advertised and consider the hypothesis  $H_0$ : p = 0.9 versus  $H_a$ : p < 0.9. If only 80% of the packages are delivered, how likely is it that a level .01 test based on n=225 packages will detect such departure from  $H_0$ ? What should the sample size be to ensure that  $\beta(0.8) = 0.01$ ? With  $\alpha = .01$ ,  $p_0 = .9$ , p' = .8, and n = 225.

Type II error: 
$$\beta(p') = 1 - \Phi\left(\frac{p_0 - p' - z_\alpha\sqrt{p_0(1-p_0)/n}}{\sqrt{p'(1-p')/n}}\right)$$

$$= 1 - \Phi\left(\frac{.9 - .8 - 2.33\sqrt{(.9)(.1)/225}}{\sqrt{(.8)(.2)/225}}\right)$$

$$= 1 - \Phi(2.00) = .0228$$

# **Example cont.**

• Using  $z_{01}$ =2.33, the sample size can then be calculated from

$$n = \left(\frac{z_{\alpha}\sqrt{p_{0}(1-p_{0})/n} + z_{\beta}\sqrt{p'(1-p')/n}}{p'-p_{0}}\right)^{2}$$
$$= \left(\frac{2.33\sqrt{(.9)(.1)} + 2.33\sqrt{(.8)(.2)}}{.8-.9}\right)^{2} \approx 266$$

• 1- $\beta$  is often referred to as the power of a test. It is the probability that the test can actually detect the alternative given the alternative is true! For  $\alpha$ -level tests, the bigger the power the better!

# **Small sample tests**

- For testing population proportions, when the sample size is small, the normal approximation is no longer appropriate. Thus a more accurate test should be used.
- As mentioned before, the sample proportion is X/n. X is the number of S's in the sample and can be treated as a binomial random variable. Thus a rejection region can be constructed using binomial cdf/pmf.
- Can we get an exact  $\alpha$ -level test using binomial?

#### P-Value

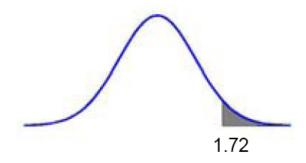
- To report the result of a hypothesis-testing analysis is to simply say whether the null hypothesis was rejected at a specified level of significance. This type of statement is somewhat inadequate because it says nothing about whether the conclusion was a very close call or quite clear cut.
- P-value is a quantity that conveys much information about the strength of evidence against  $H_0$  and allows an individual decision maker to draw a conclusion at any specified level  $\alpha$ .
- The P-value (observed significance level) is the probability, under the null hypothesis, that the test statistic is more **extreme** than the observed statistic.

## **Example cont.**

Ex. (Defective rate cont.) A factory claims that less than 10% of the components they produce are defective. A consumer group is skeptical of the claim and checks a random sample of 300 components and finds that 39 are defective. Is there evidence that 10% of all components made at the factory are defective?

If 
$$H_0$$
 is true,  $Z = \frac{\hat{p} - p_0}{\sqrt{p_0(1 - p_0)/n}} \sim N(0, 1)$ 

$$\hat{p} = \frac{39}{300} = 0.13$$
  $Z = \frac{0.13 - 0.1}{\sqrt{0.1(1 - 0.1)/300}} = 1.72$ 



#### Remarks

- P-value is corresponding to the smallest level of significance at which  $H_0$  would be rejected when a specified test procedure is used on a given data set. The smaller the P-value, the more contradictory is the data to  $H_0$ .
- Once the P-value has been determined, the conclusion at any particular level α results from comparing the P-value to α:
  - 1. P-value  $\leq \alpha \rightarrow \text{reject } H_0 \text{ at level } \alpha$ .
  - 2. P-value >  $\alpha \rightarrow$  do not reject  $H_0$  at level  $\alpha$ .
- To calculate P-value:
  - 1. Calculate the test statistic as before.
  - 2. Compute probability that we will reject the null if the threshold is the test statistic obtained from 1.
- Question: what is the relationship of P-value of the one-sided test and the P-value of the two-sided test?

# Two sample tests

- A new drug is claimed to significantly reduce the blood pressure for high blood pressure patients. What kind of tests can we use to verify the claim?
- A new drug is claimed to perform much better in terms of reducing blood pressure than an old drug. What kind of tests can we use to verify the claim?

# Things to cover

- As in the one sample testing problem, we will cover the following cases:
  - Two normal populations with known variance.
  - 2. Two populations with unknown distribution and large sample size.
  - Two normal populations with unknown variance.
  - 4. Two population proportions with large sample size.
  - 5. Tests about variances. (NOT required.)
- Basic assumptions for comparing population means:
  - 1.  $X_1, X_2, ..., X_m$  is a random sample (i.i.d.) from a population with mean  $\mu_1$  and variance  $\sigma_1^2$ .
  - Y<sub>1</sub>, Y<sub>2</sub>, ..., Y<sub>n</sub> is a random sample (i.i.d.) from a population with mean  $\mu_2$  and variance  $\sigma_2^2$ .
  - 3. The X and Y samples are independent of one another.

#### **Test statistics**

 Since we are comparing the population means, a natural test statistic to use would be the difference of two sample means. Because of independence we have,

$$E(\bar{X} - \bar{Y}) = \mu_1 - \mu_2$$

$$Var(\bar{X} - \bar{Y}) = \frac{\sigma_1^2}{m} + \frac{\sigma_2^2}{n}$$

## Case I: normal, known variance

$$H_0: \mu_1 - \mu_2 = \Delta_0$$

Test statistic: 
$$\frac{\overline{X} - \overline{Y} - \Delta_0}{\sqrt{\frac{{\sigma_1}^2}{m} + \frac{{\sigma_2}^2}{n}}} \sim N(0, 1)$$

vs Alternative Hypothesis:

$$H_a: \mu_1 - \mu_2 > \Delta_0$$
, reject if  $\frac{\overline{X} - \overline{Y} - \Delta_0}{\sqrt{\frac{\sigma_1^2}{m} + \frac{\sigma_2^2}{n}}} > Z_{\alpha}$ 

$$H_a: \mu_1 - \mu_2 < \Delta_0$$
 , reject if  $\dfrac{\overline{X} - \overline{Y} - \Delta_0}{\sqrt{\dfrac{{\sigma_1}^2}{m} + \dfrac{{\sigma_2}^2}{n}}} < -Z_{\alpha}$ 

$$H_a: \mu_1 - \mu_2 \neq \Delta_0 \text{, reject if } \frac{\overline{X} - \overline{Y} - \Delta_0}{\sqrt{\frac{\sigma_1^2}{m} + \frac{\sigma_2^2}{n}}} < -Z_{\alpha/2} \text{ or } \frac{\overline{X} - \overline{Y} - \Delta_0}{\sqrt{\frac{\sigma_1^2}{m} + \frac{\sigma_2^2}{n}}} > Z_{\alpha/2}$$

### **Questions**

- How to compute P-value for case I?
- How to compute type II errors for case I?
- In a balanced design, derive the sample size calculation formula (for alternative ">"):

$$m = n = \frac{(\sigma_1^2 + \sigma_2^2)(Z_{\alpha} + Z_{\beta})^2}{(\Delta' - \Delta_0)^2}$$

# Case II: large sample

$$H_0: \mu_1 - \mu_2 = \Delta_0$$

Test statistic: 
$$\frac{\overline{X} - \overline{Y} - \Delta_0}{\sqrt{\frac{s_1^2}{m} + \frac{s_2^2}{n}}} \sim AN(0,1)$$

vs Alternative Hypothesis:

$$H_a: \mu_1 - \mu_2 > \Delta_0$$
 , reject if  $\frac{\overline{X} - \overline{Y} - \Delta_0}{\sqrt{\frac{s_1^2}{m} + \frac{s_2^2}{n}}} > Z_{\alpha}$ 

$$H_a: \mu_1 - \mu_2 < \Delta_0$$
, reject if  $\frac{\overline{X} - \overline{Y} - \Delta_0}{\sqrt{\frac{s_1^2}{m} + \frac{s_2^2}{n}}} < -Z_{\alpha}$ 

$$H_a: \mu_1 - \mu_2 \neq \Delta_0 \text{ , reject if } \frac{\overline{X} - \overline{Y} - \Delta_0}{\sqrt{\frac{s_1^2}{m} + \frac{s_2^2}{n}}} < -Z_{\alpha/2} \text{ or } \frac{\overline{X} - \overline{Y} - \Delta_0}{\sqrt{\frac{s_1^2}{m} + \frac{s_2^2}{n}}} > Z_{\alpha/2}$$

## **Questions**

• How to construct confidence interval for  $\mu_1 - \mu_2$  in case II?

# Case III: normal, unknown variance

$$H_0: \mu_1 - \mu_2 = \Delta_0$$

Test statistic:  $\frac{\overline{X} - \overline{Y} - \Delta_0}{\sqrt{\frac{s_1^2}{m} + \frac{s_2^2}{n}}} \sim t_v$ , v is the df of the t-distribution and it's approximately estimated

by the sampled data: 
$$v = \frac{\left(\frac{S_1^2}{m} + \frac{S_2^2}{n}\right)^2}{\frac{\left(S_1^2/m\right)^2}{m-1} + \frac{\left(S_2^2/n\right)^2}{n-1}}$$
, and round  $v$  town to the nearest integer.

### Case III cont.

vs Alternative Hypothesis:

$$H_a: \mu_1 - \mu_2 > \Delta_0$$
, reject if  $\frac{\overline{X} - \overline{Y} - \Delta_0}{\sqrt{\frac{s_1^2}{m} + \frac{s_2^2}{n}}} > t_{\alpha,\nu}$ 

$$H_a: \mu_1 - \mu_2 < \Delta_0 \text{ , reject if } \frac{\overline{X} - \overline{Y} - \Delta_0}{\sqrt{\frac{s_1^2}{m} + \frac{s_2^2}{n}}} < -t_{\alpha, \nu}$$

$$H_a: \mu_1 - \mu_2 \neq \Delta_0 \text{, reject if } \frac{\overline{X} - \overline{Y} - \Delta_0}{\sqrt{\frac{s_1^2}{m} + \frac{s_2^2}{n}}} < -t_{\alpha/2, \nu} \text{ or } \frac{\overline{X} - \overline{Y} - \Delta_0}{\sqrt{\frac{s_1^2}{m} + \frac{s_2^2}{n}}} > t_{\alpha/2, \nu}$$

### **Questions**

- How to compute P-values of the test?
- How to construct confidence interval for  $\mu_1 \mu_2$  in case III?
- What if we know that  $\sigma_1^2 = \sigma_2^2$ ?

The *pooled estimator* of  $\sigma^2 = \sigma_1^2 = \sigma_2^2$  is given by

$$S_p^2 = \frac{m-1}{m+n-2} \cdot S_1^2 + \frac{n-1}{m+n-2} \cdot S_2^2$$

### Case IV

$$H_0: p_1 - p_2 = 0$$

Test statistic: 
$$\frac{\hat{p}_1 - \hat{p}_2}{\sqrt{\hat{p}(1-\hat{p})\left(\frac{1}{m} + \frac{1}{n}\right)}}, \quad \hat{p} = \frac{m}{m+n} \hat{p}_1 + \frac{n}{m+n} \hat{p}_2 \quad \text{(the weighted average of } \hat{p}_1$$

and  $\hat{p}_2$ )

### Case IV cont.

vs Alternative Hypothesis:

$$H_a: p_1 - p_2 > 0$$
, reject if  $\frac{\hat{p}_1 - \hat{p}_2}{\sqrt{\hat{p}(1-\hat{p})\left(\frac{1}{m} + \frac{1}{n}\right)}} > Z_{\alpha}$ 

$$H_a: p_1 - p_2 < 0$$
, reject if  $\frac{\hat{p}_1 - \hat{p}_2}{\sqrt{\hat{p}(1-\hat{p})\left(\frac{1}{m} + \frac{1}{n}\right)}} < -Z_{\alpha}$ 

$$H_a: p_1 - p_2 \neq 0 \text{, reject if } \frac{\hat{p}_1 - \hat{p}_2}{\sqrt{\hat{p}(1-\hat{p})\left(\frac{1}{m} + \frac{1}{n}\right)}} > Z_{\alpha/2} \text{ or } \frac{\hat{p}_1 - \hat{p}_2}{\sqrt{\hat{p}(1-\hat{p})\left(\frac{1}{m} + \frac{1}{n}\right)}} < -Z_{\alpha/2}$$

### **Remarks and Questions**

- Type II error and sample size calculation on book p.356.
- Sample size calculation formula (balanced design):

$$m = n = \frac{\left[Z_{\alpha}\sqrt{(p_1 + p_2)(q_1 + q_2)/2} + Z_{\beta}\sqrt{p_1q_1 + p_2q_2}\right]^2}{(p_1 - p_2)^2}$$

 Compare the sample size calculation formula with case I. What is the difference between the two?

#### **Tests on variances**

- How to do hypothesis testing when the null is:  $H_0$ :  $\sigma^2 = \sigma_0^2$ .
- > Chi-square test.
- In the pooled t-test, we assume the two populations have the same variance. How to test if the two variances are indeed the same?
- > F test.

## **Example**

Ex. Suppose you wish to test the effect of Prozac (antidepressants) on the well-being of depressed individuals, using a standardized "well-being scale" that sums Likert-type items to obtain a score that could range from 0 to 20. Higher scores indicate greater well-being (that is, Prozac is having a positive effect). 9 individuals participate in the experiment, the scores before and after using Prozac are recorded. Is the effect of Prozac significant?

	moodpre	moodpost
1	3.00	5.00
2	.00	1.00
3	6.00	5.00
4	7.00	7.00
5	4.00	10.00
6	3.00	9.00
7	2.00	7.00
8	1.00	11.00
9	4.00	8.00

#### Paired t-test

- As in the previous example, the data is paired, the two scores (before and after) recorded for each individual are dependent, but the between individuals the pairs are independent.
- Thus in order to test  $H_0$ :  $\mu_1 \mu_2 = 0$ , one has to look at the difference of each pair. The problem eventually becomes a one sample t-test problem.