HOMEWORK 5

4.

a.
$$\int_{-\infty}^{\infty} f(x;\theta) dx = \int_{0}^{\infty} \frac{x}{\theta^{2}} e^{-x^{2}/2\theta^{2}} dx = -e^{-x^{2}/2\theta^{2}} \int_{0}^{\infty} e^{-(-1)\theta} dx = 0$$

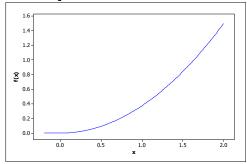
b.
$$P(X \le 200) = \int_{-\infty}^{200} f(x;\theta) dx = \int_{0}^{200} \frac{x}{\theta^2} e^{-x^2/2\theta^2} dx = -e^{-x^2/2\theta^2} \Big]_{0}^{200} \approx -.1353 + 1 = .8647$$
. $P(X < 200) = P(X \le 200) \approx .8647$, since *X* is continuous. $P(X \ge 200) = 1 - P(X < 200) \approx .1353$.

c.
$$P(100 \le X \le 200) = \int_{100}^{200} f(x;\theta) dx = -e^{-x^2/20,000} \Big|_{100}^{200} \approx .4712$$
.

d. For
$$x > 0$$
, $P(X \le x) = \int_{-\infty}^{x} f(y;\theta) dy = \int_{0}^{x} \frac{y}{\theta^{2}} e^{-y^{2}/2\theta^{2}} dx = -e^{-y^{2}/2\theta^{2}} \Big|_{0}^{x} = 1 - e^{-x^{2}/2\theta^{2}}$.

5.

a.
$$1 = \int_{-\infty}^{\infty} f(x)dx = \int_{0}^{2} kx^{2}dx = \frac{kx^{3}}{3} \Big|_{0}^{2} = \frac{8k}{3} \Rightarrow k = \frac{3}{8}.$$



b.
$$P(0 \le X \le 1) = \int_0^1 \frac{3}{8} x^2 dx = \frac{1}{8} x^3 \Big]_0^1 = \frac{1}{8} = .125.$$

c.
$$P(1 \le X \le 1.5) = \int_{1}^{1.5} \frac{3}{8} x^2 dx = \frac{1}{8} x^3 \Big|_{1}^{1.5} = \frac{1}{8} (\frac{3}{2})^3 - \frac{1}{8} (1)^3 = \frac{19}{64} = .296875$$

d.
$$P(X \ge 1.5) = 1 - \int_{1.5}^{2} \frac{3}{8} x^2 dx = \frac{1}{8} x^3 \Big|_{1.5}^{2} = \frac{1}{8} (2)^3 - \frac{1}{8} (1.5)^3 = .578125.$$

12.

a.
$$P(X < 0) = F(0) = .5$$
.

b.
$$P(-1 \le X \le 1) = F(1) - F(-1) = .6875.$$

c.
$$P(X > .5) = 1 - P(X \le .5) = 1 - F(.5) = 1 - .6836 = .3164$$
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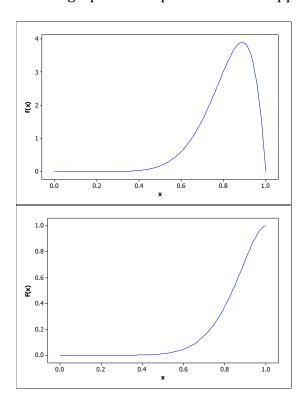
d.
$$f(x) = F'(x) = \frac{d}{dx} \left(\frac{1}{2} + \frac{3}{32} \left(4x - \frac{x^3}{3} \right) \right) = 0 + \frac{3}{32} \left(4 - \frac{3x^2}{3} \right) = .09375 \left(4 - x^2 \right).$$

e. By definition, $F(\beta b) = .5 \cdot F(0) = .5$ from **a** above, which is as desired.

- 15.
- **a.** Since X is limited to the interval (0, 1), F(x) = 0 for $x \le 0$ and F(x) = 1 for $x \ge 1$. For 0 < x < 1.

$$F(x) = \int_{-\infty}^{x} f(y) dy = \int_{0}^{x} 90y^{8} (1-y) dy = \int_{0}^{x} (90y^{8} - 90y^{9}) dy = 10y^{9} - 9y^{10} \Big|_{0}^{x} = 10x^{9} - 9x^{10}.$$

The graphs of the pdf and cdf of *X* appear below.



- **b.** $F(.5) = 10(.5)^9 9(.5)^{10} = .0107.$
- c. $P(.25 < X \le .5) = F(.5) F(.25) = .0107 [10(.25)^9 9(.25)^{10}] = .0107 .0000 = .0107.$

Since *X* is continuous, $P(.25 \le X \le .5) = P(.25 < X \le .5) = .0107$.

- **d.** The 75th percentile is the value of *x* for which F(x) = .75: $10x^9 9x^{10} = .75$ $\Rightarrow x = .9036$ using software.
- e. $E(X) = \int_{-\infty}^{\infty} x \cdot f(x) dx = \int_{0}^{1} x \cdot 90x^{8} (1-x) dx = \int_{0}^{1} (90x^{9} 90x^{10}) dx = 9x^{10} \frac{90}{11}x^{11} \Big]_{0}^{1} = 9 \frac{90}{11}$ $= \frac{9}{11} = .8182$. Similarly, $E(X^{2}) = \int_{-\infty}^{\infty} x^{2} \cdot f(x) dx = \int_{0}^{1} x^{2} \cdot 90x^{8} (1-x) dx = \dots = .6818$, from which $V(X) = .6818 - (.8182)^{2} = .0124$ and $\sigma_{X} = .11134$.
- f. $\mu \pm \sigma = (.7068, .9295)$. Thus, $P(\mu \sigma \le X \le \mu + \sigma) = F(.9295) F(.7068) = .8465 .1602 = .6863$, and the probability X is more than 1 standard deviation from its mean value equals 1 .6863 = 3137.

e.
$$\Phi(1.37) = .9147$$
.

f.
$$P(-1.75 < Z) + [1 - P(Z < -1.75)] = 1 - \Phi(-1.75) = .9599$$
.

g.
$$\Phi(2) - \Phi(-1.50) = .9104$$
.

29.

e.
$$P(c \le Z) = .121 \Rightarrow 1 - P(Z \le c) = .121 \Rightarrow 1 - \Phi(c) = .121 \Rightarrow \Phi(c) = .879 \Rightarrow c = 1.17$$
.

d.
$$P(-c \le Z \le c) = \Phi(c) - \Phi(-c) = \Phi(c) - (1 - \Phi(c)) = 2\Phi(c) - 1 = .668 \Rightarrow \Phi(c)$$

= .834 \Rightarrow
 $c = 0.97$.

e.
$$P(c \le |Z|) = 1 - P(|Z| < c) = 1 - [\Phi(c) - \Phi(-c)] = 1 - [2\Phi(c) - 1] = 2 - 2\Phi(c)$$

= .016 $\Rightarrow \Phi(c) = .992 \Rightarrow c = 2.41$.

42. The probability *X* is within .1 of its mean is given by $P(\mu - .1 \le X \le \mu + .1) = P\left(\frac{(\mu - .1) - \mu}{\sigma} < Z < \frac{(\mu + .1) - \mu}{\sigma}\right) = \Phi\left(\frac{.1}{\sigma}\right) - \Phi\left(\frac{.1}{\sigma}\right) = 2\Phi\left(\frac{.1}{\sigma}\right) - 1$. If we require this to equal 95%, we find $2\Phi\left(\frac{.1}{\sigma}\right) - 1 = .95 \Rightarrow \Phi\left(\frac{.1}{\sigma}\right) = .975 \Rightarrow \frac{.1}{\sigma} = 1.96$ from the standard normal table. Thus, $\sigma = \frac{.1}{1.96} = .0510$.

Alternatively, use the empirical rule: 95% of all values lie within 2 standard deviations of the mean, so we want $2\sigma = .1$, or $\sigma = .05$. (This is not quite as precise as the first answer.)

With μ = .500 inches, the acceptable range for the diameter is between .496 and .504 inches, so unacceptable bearings will have diameters smaller than .496 or larger than .504.

The new distribution has μ = .499 and σ =.002.

$$P(X < .496 \text{ or } X > .504) = P\left(Z < \frac{.496 - .499}{.002}\right) + P\left(Z > \frac{.504 - .499}{.002}\right) = P(Z < -1.5) + P(Z > 2.5) =$$

$$\Phi(-1.5) + [1 - \Phi(2.5)] = .073.7.3\%$$
 of the bearings will be unacceptable.

60.

a.
$$P(X \le 100) = 1 - e^{-(100)(.01386)} = 1 - e^{-1.386} = .7499.$$

 $P(X \le 200) = 1 - e^{-(200)(.01386)} = 1 - e^{-2.772} = .9375.$
 $P(100 \le X \le 200) = P(X \le 200) - P(X \le 100) = .9375 - .7499 = .1876.$

b. First, since *X* is exponential,
$$\mu = \frac{1}{\lambda} = \frac{1}{.01386} = 72.15$$
, $\sigma = 72.15$. Then

$$P(X > \mu + 2\sigma) = P(X > 72.15 + 2(72.15)) = P(X > 216.45) = 1 - (1 - e^{-0.01386(216.45)}) = e^{-3} = .0498.$$

- c. Remember the median is the solution to F(x) = .5. Use the formula for the exponential cdf and solve for x: $F(x) = 1 e^{-.01386x} = .5 \Rightarrow e^{-.01386x} = .5 \Rightarrow .01386x = \ln(.5) \Rightarrow x = -\frac{\ln(.5)}{01386} = 50.01 \text{ m}.$
- **70.** To find the (100p)th percentile, set F(x) = p and solve for x: $p = F(x) = 1 e^{-\lambda x}$ $\Rightarrow e^{-\lambda x} = 1 p \Rightarrow$ $-\lambda x = \ln(1-p) \Rightarrow x = -\frac{\ln(1-p)}{\lambda}$.

To find the median, set p = .5 to get $\beta = -\frac{\ln(1 - .5)}{\lambda} = \frac{.693}{\lambda}$.

88. The data values and *z* percentiles provided result in the probability plot below. The plot shows some non-trivial departures from linearity, especially in the lower tail of the distribution. This indicates a normal distribution might not be a good fit to the population distribution of clubhead velocities for female golfers.

