

## 5.46

- a. The sampling distribution of  $\bar{X}$  is centered at  $E(\bar{X}) = \mu = 12$  cm, and the standard deviation of the  $\bar{X}$  distribution is  $\sigma_{\bar{X}} = \frac{\sigma_{\bar{X}}}{\sqrt{n}} = \frac{.04}{\sqrt{16}} = .01$  cm.
- b. With  $n = 64$ , the sampling distribution of  $\bar{X}$  is still centered at  $E(\bar{X}) = \mu = 12$  cm, but the standard deviation of the  $\bar{X}$  distribution is  $\sigma_{\bar{X}} = \frac{\sigma_{\bar{X}}}{\sqrt{n}} = \frac{.04}{\sqrt{64}} = .005$  cm.
- c.  $\bar{X}$  is more likely to be within .01 cm of the mean (12 cm) with the second, larger, sample. This is due to the decreased variability of  $\bar{X}$  that comes with a larger sample size.

## 5.50

- a.  $P(9,900 \leq \bar{X} \leq 10,200) \approx P\left(\frac{9,900 - 10,000}{500/\sqrt{40}} \leq Z \leq \frac{10,200 - 10,000}{500/\sqrt{40}}\right)$   
 $= P(-1.26 \leq Z \leq 2.53) = \Phi(2.53) - \Phi(-1.26) = .9943 - .1038 = .8905$ .
- b. According to the guideline given in Section 5.4,  $n$  should be greater than 30 in order to apply the CLT, thus using the same procedure for  $n = 15$  as was used for  $n = 40$  would not be appropriate.

## 5.53

- a. With the values provided,  
 $P(\bar{X} \geq 51) = P\left(Z \geq \frac{51 - 50}{1.2/\sqrt{9}}\right) = P(Z \geq 2.5) = 1 - .9938 = .0062$ .
- b. Replace  $n = 9$  by  $n = 40$ , and  
 $P(\bar{X} \geq 51) = P\left(Z \geq \frac{51 - 50}{1.2/\sqrt{40}}\right) = P(Z \geq 5.27) \approx 0$ .

## 5.60

$Y$  is normally distributed with  $\mu_Y = \frac{1}{2}(\mu_1 + \mu_2) - \frac{1}{3}(\mu_3 + \mu_4 + \mu_5) = -1$ , and

$$\sigma_Y^2 = \frac{1}{4}\sigma_1^2 + \frac{1}{4}\sigma_2^2 + \frac{1}{9}\sigma_3^2 + \frac{1}{9}\sigma_4^2 + \frac{1}{9}\sigma_5^2 = 3.167 \Rightarrow \sigma_Y = 1.7795.$$

Thus,  $P(0 \leq Y) = P\left(\frac{0 - (-1)}{1.7795} \leq Z\right) = P(.56 \leq Z) = .2877$  and

$$P(-1 \leq Y \leq 1) = P\left(0 \leq Z \leq \frac{2}{1.7795}\right) = P(0 \leq Z \leq 1.12) = .3686.$$

### 6.3

- a. We use the sample mean,  $\bar{x} = 1.3481$ .
- b. Because we assume normality, the mean = median, so we also use the sample mean  $\bar{x} = 1.3481$ . We could also easily use the sample median.
- c. We use the 90<sup>th</sup> percentile of the sample:  $\cancel{1.28} (1.28)\sigma = \bar{x} + 1.28s = 1.3481 + (1.28)(.3385) = 1.7814$ .
- d. Since we can assume normality,  

$$P(X < 1.5) \approx P\left(Z < \frac{1.5 - \bar{x}}{s}\right) = P\left(Z < \frac{1.5 - 1.3481}{.3385}\right) = P(Z < .45) = .6736.$$
- e. The estimated standard error of  $\bar{x} = \frac{\hat{\sigma}}{\sqrt{n}} = \frac{s}{\sqrt{n}} = \frac{.3385}{\sqrt{16}} = .0846$ .

### 6.5

Let  $\theta$  = the total audited value. Three potential estimators of  $\theta$  are  $\hat{\theta}_1 = N\bar{X}$ ,  $\hat{\theta}_2 = T - N\bar{D}$ , and  $\hat{\theta}_3 = T \cdot \frac{\bar{X}}{\bar{Y}}$ .

From the data,  $\bar{y} = 374.6$ ,  $\bar{x} = 340.6$ , and  $\bar{d} = 34.0$ . Knowing  $N = 5,000$  and  $T = 1,761,300$ , the three corresponding estimates are  $\hat{\theta}_1 = (5,000)(340.6) = 1,703,000$ ,  $\hat{\theta}_2 = 1,761,300 - (5,000)(34.0) = 1,591,300$ , and  $\hat{\theta}_3 = 1,761,300 \left( \frac{340.6}{374.6} \right) = 1,601,438.281$ .

### 6.28

- a.  $\left( \frac{x_1}{\theta} \exp[-x_1^2 / 2\theta] \right) \dots \left( \frac{x_n}{\theta} \exp[-x_n^2 / 2\theta] \right) = (x_1 \dots x_n) \frac{\exp[-\sum x_i^2 / 2\theta]}{\theta^n}$ . The natural log of the likelihood function is  $\ln(x_1 \dots x_n) - n \ln(\theta) - \frac{\sum x_i^2}{2\theta}$ . Taking the derivative with respect to  $\theta$  and equating to 0 gives  $-\frac{n}{\theta} + \frac{\sum x_i^2}{2\theta^2} = 0$ , so  $n\theta = \frac{\sum x_i^2}{2}$  and  $\theta = \frac{\sum x_i^2}{2n}$ . The mle is therefore  $\hat{\theta} = \frac{\sum X_i^2}{2n}$ , which is identical to the unbiased estimator suggested in Exercise 15.
- b. For  $x > 0$  the cdf of  $X$  is  $F(x; \theta) = P(X \leq x) = 1 - \exp\left[\frac{-x^2}{2\theta}\right]$ . Equating this to .5 and solving for  $x$  gives the median in terms of  $\theta$ :  $.5 = \exp\left[\frac{-x^2}{2\theta}\right] \Rightarrow x = \tilde{\mu} = \sqrt{-2\theta \ln(.5)} = \sqrt{1.3863\theta}$ . The mle of  $\tilde{\mu}$  is therefore  $\sqrt{1.3863\hat{\theta}}$ .

### 7.3

- a. A 90% confidence interval will be narrower. The  $z$  critical value for a 90% confidence level is 1.645, smaller than the  $z$  of 1.96 for the 95% confidence level, thus producing a narrower interval.
- b. Not a correct statement. Once an interval has been created from a sample, the mean  $\mu$  is either enclosed by it, or not. We have 95% confidence in the general procedure, under repeated and independent sampling.
- c. Not a correct statement. The interval is an estimate for the population mean, not a boundary for population values.
- d. Not a correct statement. In theory, if the process were repeated an infinite number of times, 95% of the intervals would contain the population mean  $\mu$ . We *expect* 95 out of 100 intervals will contain  $\mu$ , but we don't know this to be true.

### 7.6

- a.  $8439 \pm \frac{(1.645)(100)}{\sqrt{25}} = 8439 \pm 32.9 = (8406.1, 8471.9).$
- b.  $1 - \alpha = .92 \Rightarrow \alpha = .08 \Rightarrow \alpha / 2 = .04$  so  $z_{\alpha/2} = z_{.04} = 1.75.$