

HOMEWORK 8

Chapter 6

28.

- a. $\left(\frac{x_1}{\theta} \exp[-x_1^2 / 2\theta]\right) \dots \left(\frac{x_n}{\theta} \exp[-x_n^2 / 2\theta]\right) = (x_1 \dots x_n) \frac{\exp[-\sum x_i^2 / 2\theta]}{\theta^n}$. The natural log of the likelihood function is $\ln(x_1 \dots x_n) - n \ln(\theta) - \frac{\sum x_i^2}{2\theta}$. Taking the derivative with respect to θ and equating to 0 gives $-\frac{n}{\theta} + \frac{\sum x_i^2}{2\theta^2} = 0$, so $n\theta = \frac{\sum x_i^2}{2}$ and $\theta = \frac{\sum x_i^2}{2n}$. The mle is therefore $\hat{\theta} = \frac{\sum X_i^2}{2n}$, which is identical to the unbiased estimator suggested in Exercise 15.

Chapter 7

3.

- a. A 90% confidence interval will be narrower. The z critical value for a 90% confidence level is 1.645, smaller than the z of 1.96 for the 95% confidence level, thus producing a narrower interval.
- b. Not a correct statement. Once an interval has been created from a sample, the mean μ is either enclosed by it, or not. We have 95% confidence in the general procedure, under repeated and independent sampling.
- c. Not a correct statement. The interval is an estimate for the population mean, not a boundary for population values.
- d. Not a correct statement. In theory, if the process were repeated an infinite number of times, 95% of the intervals would contain the population mean μ . We *expect* 95 out of 100 intervals will contain μ , but we don't know this to be true.

4.

- a. $58.3 \pm \frac{1.96(3)}{\sqrt{25}} = 58.3 \pm 1.18 = (57.1, 59.5)$.
- c. $58.3 \pm \frac{2.58(3)}{\sqrt{100}} = 58.3 \pm .77 = (57.5, 59.1)$.
- e. $n = \left[\frac{2(2.58)3}{1} \right]^2 = 239.62 \approx 240$.

6.

- a. $8439 \pm \frac{(1.645)(100)}{\sqrt{25}} = 8439 \pm 32.9 = (8406.1, 8471.9)$.
- b. $1 - \alpha = .92 \Rightarrow \alpha = .08 \Rightarrow \alpha / 2 = .04$ so $z_{\alpha/2} = z_{.04} = 1.75$.