

~~Answer Handwritten~~

Answer Key

**Econ 6700: Math Methods II, Midterm Exam, Fall 2020.**

You may consult the book, your notebook, or any other materials, but you may not work with another student or consult any other individual; all of the work must be yours and yours alone.

Answer any three of the following four questions. All questions receive equal weight.

**Question 1.** Let  $\mathbf{A}$  be a  $(2 \times 2)$  matrix whose elements are labelled as follows,

$\mathbf{A}_{(2 \times 2)} = \begin{bmatrix} a_{11} & a_{12} \\ a_{21} & a_{22} \end{bmatrix}$ . Let  $r_1$  and  $r_2$  denote the eigenvalues of  $\mathbf{A}$ . Consider the eigenvectors of  $\mathbf{A}$  normalized so that the first element of each is 1. Thus, denote the eigenvectors of  $\mathbf{A}$  that are associated with  $r_1$  and  $r_2$ , respectively, by

$$\mathbf{v}_1 = \begin{bmatrix} 1 \\ v_{12} \end{bmatrix} \text{ and } \mathbf{v}_2 = \begin{bmatrix} 1 \\ v_{22} \end{bmatrix}.$$

A.) Derive expressions for  $v_{12}$  and  $v_{22}$ , each in terms of  $a_{11}, a_{12}, a_{21}, a_{22}, r_1$ , and  $r_2$ . Derive expressions for  $r_1$  and  $r_2$ , each in terms of  $a_{11}, a_{12}, a_{21}$ , and  $a_{22}$ .

B.) Prove that the sum of the eigenvalues of  $\mathbf{A}_{(2 \times 2)}$  equals the sum of its diagonal elements. Prove that the product of the eigenvalues of  $\mathbf{A}_{(2 \times 2)}$  is equal to its determinant.

C.) Suppose that  $\mathbf{P}^{-1}\mathbf{A}\mathbf{P} = \mathbf{D}$  where  $\mathbf{D}_{(k \times k)}$  is a diagonal matrix. Prove that the diagonal elements of  $\mathbf{D}_{(k \times k)}$  are the eigenvalues of  $\mathbf{A}_{(k \times k)}$  and that the columns of  $\mathbf{P}_{(k \times k)}$  are the corresponding eigenvectors.

(Exam continued next page.)

**Question 2.** Consider the following model:

$$y_t = y_{t+1} - (i_t - \pi_{t+1}) \quad (1)$$

$$\pi_t = \beta \pi_{t+1} + \alpha y_t + v_t \quad (2)$$

$$i_t = \gamma \pi_t + \varepsilon_t \quad (3)$$

where  $y_t$ ,  $i_t$ , and  $\pi_t$  denote output, the nominal interest rate, and inflation, respectively.

The parameters  $\beta$ ,  $\alpha$ , and  $\gamma$  are positive. The shocks  $v_t$  and  $\varepsilon_t$  are both i.i.d.  $(0, \sigma^2)$ .

A.) Rewrite the system described by equations (1)-(3) as  $\mathbf{x}_{t+1} = \mathbf{Ax}_t + \mathbf{Bu}_t$  where

$\mathbf{x}_t = \begin{pmatrix} y_t \\ \pi_t \end{pmatrix}$ . Give each of the elements of the matrices  $\mathbf{A}$  and  $\mathbf{B}$  in terms of the

parameters,  $\beta$ ,  $\alpha$ , and  $\gamma$ . Specify the elements of  $\mathbf{u}_t$  in terms of the variables and parameters of the original system, equations (1)-(3).

B.) Now, instead, use the system described by equations (1)-(3) to derive a 2<sup>nd</sup> order difference equation in  $\pi_t$ . Show how the coefficients and shocks in this 2<sup>nd</sup> order difference equation depend on the parameters and variables in the original system, equations (1)-(3).

C.) Suppose now that, instead of both shocks being i.i.d.  $(0, \sigma^2)$ , each is a constant. That is, let  $v_t = v_s$  and  $\varepsilon_t = \varepsilon_s \forall t$  where  $v_s$  and  $\varepsilon_s$  are nonzero constants. Using the 2<sup>nd</sup> order difference equation that you derived in part B, derive the stationary-state value of  $\pi_t$ .

Give the stationary state value of  $\pi_t$  as a function of  $\beta$ ,  $\alpha$ ,  $\gamma$ ,  $v_s$  and  $\varepsilon_s$ .

**Question 3.** Using Hamilton's notation, consider a Markov chain. The transition probability matrix is  $\mathbf{P}_{(N \times N)}$  and  $\xi_t$  is an  $(N \times 1)$  column vector whose  $i$ th element is 1 when  $s_t = i$  and whose other elements are zero.

You may take the following propositions as given

(i) A matrix and its transpose have the same eigenvalues.

(ii)  $E(\xi_{t+M} | \xi_t) = \mathbf{P}^M \xi_t$

(iii) For the decomposition  $\mathbf{P} = \mathbf{T}\Lambda\mathbf{T}^{-1}$ , the first row of  $\mathbf{T}^{-1}$  is the (transpose of) the eigenvector of  $\mathbf{P}'$  associated with the (1,1) element of  $\Lambda$ .

A.) Prove that 1 is an eigenvalue of  $\mathbf{P}$ . Show that there exists a probability vector  $\pi$  such that  $\pi = \mathbf{P}\pi$ .

B.) Assuming that all other eigenvalues are real and less than 1 in absolute value, prove that  $\lim_{M \rightarrow \infty} \mathbf{P}^M = \pi\psi'$  where, here,  $\psi_{(N \times 1)}$  is a vector of ones. *That is,  $\psi_{(N \times 1)} = [1, 1, \dots, 1]$ .*

(Exam continued next page.)

**Question 4.** Throughout this question you should assume that  $y_t$  is covariance stationary. You should also assume that  $|\phi| < 1$  and that  $\varepsilon_t$  is white noise.

A.) Derive the MA( $\infty$ ) representation of the following two stochastic processes:

$$y_t = c + \phi y_{t-1} + \varepsilon_t \quad (1)$$

and

$$y_t = c + \phi y_{t-1} + \varepsilon_t + \theta \varepsilon_{t-1}. \quad (2)$$

B.) Using the lag operator, derive the autoregressive, AR, representation of the following MA(1) process:

$$y_t = \mu + \varepsilon_t - \theta \varepsilon_{t-1}, \quad |\theta| < 1. \quad (3)$$

What is the order of the AR representation?

C.) Let  $\gamma_j$  denote the  $j$ th autocovariance of  $y_t$  and let  $\rho_j$  denote its  $j$ th autocorrelation. Derive  $\gamma_j$  and  $\rho_j$ ,  $j = 0, 1$ , for the case where  $y_t$  is MA(1).

D.) Let  $P_t$  denote the expected present discounted value of  $y_t$ . Thus  $P_t = E_t \sum_{i=0}^{\infty} \beta^i y_{t+i}$  where  $0 < \beta < 1$ . Here  $E_t[\cdot]$  denotes  $E[\cdot | I_t]$  where  $I_t$ , the current information set, is  $I_t = \{y_t, y_{t-1}, y_{t-2}, \dots; P_{t-1}, P_{t-2}, P_{t-3}, \dots; \Theta\}$ , and where  $\Theta$  denotes the structure of the relevant stochastic process including its parameters (and including  $\beta$ ).

Derive the value of  $P_t$  (the REE solution) for the ARMA(1,1) process in equation (2) above, where  $c = 0$ .

Question 1

1.1

A.) If  $v_j$  is Re-e-vector of A Associated w/  
Re e-value  $r_j$  Then

$$(A - r_j I) v_j = \begin{bmatrix} 0 \\ 0 \end{bmatrix} \text{ or}$$

$$\begin{bmatrix} a_{11} - r_j & a_{12} \\ a_{21} & a_{22} - r_j \end{bmatrix} \begin{bmatrix} 1 \\ v_{j2} \end{bmatrix} = \begin{bmatrix} 0 \\ 0 \end{bmatrix}$$

First row:  $(a_{11} - r_j) + a_{12}v_{j2} = 0 \text{ or}$

$$v_{j2} = \frac{r_j - a_{11}}{a_{12}} \quad (1)$$

for  $j=1, 2.$

Second row:  $a_{21} + (a_{22} - r_j)v_{j2} = 0 \text{ or}$

$$v_{j2} = \frac{-a_{21}}{a_{22} - r_j} \quad (2)$$

for  $j=1, 2.$

Either (1) or (2) Answers This part of PART A.

I Show below that  $v_{j2}$  from (1) is the same as  
 $v_{j2}$  from (2) But Such a proof is NOT  
part of The answer.

A.) continued: If  $r$  is an e-value of A Then

$$\text{Det}(A - rI) = 0 \quad \text{or}$$

$$\text{Det} \begin{bmatrix} (a_{11}-r) & a_{12} \\ a_{21} & (a_{22}-r) \end{bmatrix} = 0 \quad \text{or}$$

$$(a_{11}-r)(a_{22}-r) - a_{21}a_{12} = 0 \quad \text{or}$$

$$a_{11}a_{22} - (a_{11}+a_{22})r + r^2 - a_{21}a_{12} = 0 \quad \text{or}$$

$$r^2 - (a_{11}+a_{22})r + (a_{11}a_{22} - a_{21}a_{12}) = 0 \quad (3)$$

Thus

$$r_{1,2} = \frac{1}{2} \left\{ (a_{11}+a_{22}) \pm \sqrt{(a_{11}+a_{22})^2 - 4(a_{11}a_{22} - a_{21}a_{12})} \right\} \quad (4)$$

Eqn (4) Answers the Second part of Part A.

Digress to Show That (1) and (2) give the same  $r_{1,2}$ .

Must Show  $\left( \frac{r_j - a_{11}}{a_{12}} \right) = \left( \frac{-a_{21}}{a_{22} - r_j} \right) \quad \text{or}$

$$(r_j - a_{11})(a_{22} - r_j) = -a_{21}a_{12} \quad \text{or}$$

$$-a_{11}a_{22} - r_j^2 + (a_{11} + a_{22})r_j = -a_{21}a_{12} \quad \text{or}$$

$$0 = r_j^2 - (a_{11} + a_{22})r_j + a_{11}a_{22} - a_{21}a_{12} \quad \text{which is (3)}$$

QED

B.) Let  $Q_1 \equiv (a_{11} + a_{22})$  and  $Q_2 \equiv [(a_{11} + a_{22})^2 - 4(a_{11}a_{22} - a_{12}a_{21})]$   
 So Part (4) gives

$$P_1 = \frac{1}{2} [Q_1 - Q_2^{\frac{1}{2}}] \text{ and } P_2 = \frac{1}{2} [Q_1 + Q_2^{\frac{1}{2}}] \quad (5)$$

Thus  ~~$P_1 + P_2$~~   $P_1 + P_2 = \frac{1}{2} [Q_1 - Q_2^{\frac{1}{2}}] + \frac{1}{2} [Q_1 + Q_2^{\frac{1}{2}}]$

$$P_1 + P_2 = Q_1 \text{ so}$$

$$P_1 + P_2 = a_{11} + a_{22} \quad \underline{\text{QED}} \begin{array}{l} \text{Sum of Diagonals} \\ = \text{Sum of Roots} \end{array}$$

$$\begin{aligned} \text{Also } P_1 \cdot P_2 &= \frac{1}{2} [Q_1 - Q_2^{\frac{1}{2}}] \frac{1}{2} [Q_1 + Q_2^{\frac{1}{2}}] \\ &= \frac{1}{4} [Q_1^2 - Q_2^{\frac{1}{2}} Q_1 + Q_1 Q_2^{\frac{1}{2}} - (Q_2^{\frac{1}{2}})^2] \\ &= \frac{1}{4} [(a_{11} + a_{22})^2 - (a_{11} + a_{22})^2 + 4(a_{11}a_{22} - a_{12}a_{21})] \end{aligned}$$

$$\text{Thus } P_1 \cdot P_2 = a_{11}a_{22} - a_{12}a_{21} = \text{Det}(A). \quad \underline{\text{QED}} \begin{array}{l} \text{Prod of Roots} \\ = \text{Det}(A) \end{array}$$

There is an alternate proof on Pg 99 of Simon + Blume.

C.) We have  $P^{-1}AP=D$  where  $D$  is a diagonal matrix.

Let  $P = \begin{bmatrix} V_1 & V_2 & \cdots & V_K \end{bmatrix}_{(K \times K)}$  and

let  $D = \begin{bmatrix} d_1 & 0 & \cdots & 0 \\ 0 & d_2 & & \\ \vdots & & \ddots & \\ 0 & \cdots & & d_K \end{bmatrix}_{(K \times K)}$

Pre multiply both sides of  $P^{-1}AP=D$  by  $P$  to get

$$AP = PD$$

or

$$A[V_1 \ V_2 \ \cdots \ V_K] = [V_1 \ V_2 \ \cdots \ V_K] \begin{bmatrix} d_1 & 0 & \cdots & 0 \\ 0 & d_2 & & \\ \vdots & & \ddots & \\ 0 & \cdots & & d_K \end{bmatrix}$$

Equating, column by column we have

$$Av_1 = d_1 v_1 \text{ or } (A - d_1 I) v_1 = 0_{(K \times 1)}$$

$$Av_2 = d_2 v_2 \text{ or } (A - d_2 I) v_2 = 0_{(K \times 1)}$$

:

$$Av_K = d_K v_K \text{ or } (A - d_K I) v_K = 0_{(K \times 1)}$$

So  $d_j$  and  $v_j$  satisfy the defn of an eigenvalue and of an eigenvector, respectively, for  $j=1, 2, \dots, K$ .

QED

Question 2

2.1

Begin From

$$y_t = y_{t+1} - (i_t - \pi_{t+1}) \quad (1)$$

$$\pi_t = \beta \pi_{t+1} + \alpha y_t + v_t \quad (2)$$

$$i_t = \gamma \pi_t + \varepsilon_t \quad (3)$$

A.) Use (3) in (1) to get

$$y_t = y_{t+1} - \gamma \pi_t + \pi_{t+1} - \varepsilon_t \text{ or}$$

$$y_{t+1} = y_t + \gamma \pi_t - \pi_{t+1} + \varepsilon_t \quad (4)$$

Re-write (2) as  $\beta \pi_{t+1} = \pi_t - \alpha y_t - v_t$  or

$$\pi_{t+1} = \frac{1}{\beta} \pi_t - \frac{\alpha}{\beta} y_t - \frac{1}{\beta} v_t \quad (5)$$

use (5) in (4) to get

$$y_{t+1} = y_t + \gamma \pi_t - \left[ \frac{1}{\beta} \pi_t - \frac{\alpha}{\beta} y_t - \frac{1}{\beta} v_t \right] + \varepsilon_t \text{ or}$$

$$y_{t+1} = \left(1 + \frac{\alpha}{\beta}\right) y_t + \left(\gamma - \frac{1}{\beta}\right) \pi_t + \left(\varepsilon_t + \frac{1}{\beta} v_t\right) \quad (6)$$

(2.2)

Taken together eqns (5) and (6) can be written as

$$\begin{bmatrix} \gamma_{t+1} \\ \pi_{t+1} \end{bmatrix} = \begin{bmatrix} (1 + \frac{\alpha}{\beta}) & (\gamma - \frac{1}{\beta}) \\ -\frac{\alpha}{\beta} & \frac{1}{\beta} \end{bmatrix} \begin{bmatrix} \gamma_t \\ \pi_t \end{bmatrix} + \begin{bmatrix} 1 & \frac{1}{\beta} \\ 0 & -\frac{1}{\beta} \end{bmatrix} \begin{bmatrix} \varepsilon_t \\ v_t \end{bmatrix} \quad (?)$$

cf.

$$x_{t+1} = A x_t + B u_t \quad (8)$$

B. Once again, (3) in (1) gives (4)

$$\gamma_{t+1} = \gamma_t + \gamma \pi_t - \pi_{t+1} + \varepsilon_t \quad (4)$$

Invert (2) to write

$$\cancel{\beta \pi_t} \ \& \ \gamma_t = \pi_t - \beta \pi_{t+1} - \gamma_t \quad \text{or}$$

$$\gamma_t = \frac{1}{\alpha} \pi_t - \left( \frac{\beta}{\alpha} \right) \pi_{t+1} - \frac{1}{\alpha} \gamma_t \quad (9)$$

Thus

$$\gamma_{t+1} = \frac{1}{\alpha} \pi_{t+1} - \left( \frac{\beta}{\alpha} \right) \pi_{t+2} - \frac{1}{\alpha} \gamma_{t+1} \quad (10)$$

Use (9) and (10) in (4) to get

$$\frac{1}{\alpha} \Pi_{t+1} - \frac{\beta}{\alpha} \Pi_{t+2} - \frac{1}{\alpha} V_{t+1} = \frac{1}{\alpha} \Pi_t - \frac{\beta}{\alpha} \Pi_{t+1} - \frac{1}{\alpha} V_t \\ + \gamma \Pi_t - \Pi_{t+1} + \varepsilon_t$$

or

$$\frac{1}{\alpha} \Pi_{t+1} - \frac{\beta}{\alpha} \Pi_{t+2} - \frac{1}{\alpha} V_{t+1} = \left( \frac{1}{\alpha} + \gamma \right) \Pi_t - \left( \frac{\beta}{\alpha} + 1 \right) \Pi_{t+1} + \varepsilon_t - \frac{1}{\alpha} V_t$$

or, collecting

$$- \frac{\beta}{\alpha} \Pi_{t+2} + \left( \frac{1}{\alpha} + \frac{\beta}{\alpha} + 1 \right) \Pi_{t+1} - \left( \frac{1}{\alpha} + \gamma \right) \Pi_t = \varepsilon_t + \frac{1}{\alpha} (V_{t+1} - V_t)$$

Multiply through by  $-\frac{\alpha}{\beta}$  to get

$$\Pi_{t+2} - \left( \frac{1}{\beta} + 1 + \frac{\gamma}{\beta} \right) \Pi_{t+1} + \left( \frac{1}{\beta} + \frac{\gamma}{\beta} \right) \Pi_t = \frac{1}{\beta} (V_t - V_{t+1}) - \frac{\alpha}{\beta} \varepsilon_t$$

or

$$\boxed{\Pi_{t+2} - \left( 1 + \frac{1+\gamma}{\beta} \right) \Pi_{t+1} + \left( \frac{1+\gamma}{\beta} \right) \Pi_t = \frac{1}{\beta} (V_t - V_{t+1}) - \frac{\alpha}{\beta} \varepsilon_t}$$

(11)

Eqn (11) answers part (B).

c) In (11) set  $\gamma_t = \gamma_{t+1} = \gamma_s$  and  $\varepsilon_t = \varepsilon_s$ .

Also Set  $\pi_{t+2} = \pi_{t+1} = \pi_t$ . The ~~the~~ result is

$$\pi_t = \left(1 + \frac{1+\alpha}{\beta}\right) + \frac{1+\alpha\gamma}{\beta}$$

$$\pi_s - \left(1 + \frac{1+\alpha}{\beta}\right) \pi_s + \left(\frac{1+\alpha\gamma}{\beta}\right) \pi_s = -\frac{\alpha}{\beta} \varepsilon_s.$$

so

$$\left[1 - 1 - \frac{1+\alpha}{\beta} + \frac{1}{\beta} + \frac{\alpha\gamma}{\beta}\right] \pi_s = -\frac{\alpha}{\beta} \varepsilon_s \quad \text{or}$$

$$\left[\frac{-1}{\beta} - \frac{\alpha}{\beta} + \frac{1}{\beta} + \frac{\alpha\gamma}{\beta}\right] \pi_s = -\frac{\alpha}{\beta} \varepsilon_s \quad \text{or}$$

$$\frac{\alpha(\gamma-1)}{\beta} \pi_s = -\frac{\alpha}{\beta} \varepsilon_s \quad \text{or}$$

$$(\gamma-1) \pi_s = -\varepsilon_s \quad \text{or}$$

$$\boxed{\pi_s = (1-\gamma) \varepsilon_s \quad (12)}$$

Eqn (12) answers Part C

A.) Since  $P$  is a TRANSITION PROBABILITY MATRIX, The elements of each of its Columns sum to 1. Thus,

for  $\Psi = \begin{pmatrix} 1 & 1 & \dots & 1 \end{pmatrix}'_{(N \times 1)}$

$$P' \Psi = \Psi \quad (1)$$

Rewrite (1) as

$$(P' - 1 \cdot I_N) \Psi = \vec{q}_{(N \times 1)} \quad (2)$$

And we see that 1 is an eigenvalue of  $P'$ , and  $\Psi$  is its associated eigenvector. Since  $P$  and  $P'$  have the same eigenvalues it follows that  
1 is an eigenvalue of  $P$ .

Consider now the eigenvector of  $P$  associated with the eigenvalue  $\lambda_1 = 1$ . Denote this e-vector by  $X_{(N \times 1)}$ . Thus

~~$PX =$~~   $(P - 1 \cdot I_N) X = \vec{q}_{(N \times 1)} \quad (3)$

$$X \neq 0$$



And, Thus  $X = P X$  (4)

If we Normalize the elements of  $X$  so that they sum to 1, This Normalized vector,  $\overline{\pi}$ , is a PROBABILITY Vector. So

$$\overline{\pi} = P \overline{\pi} \quad (5)$$

and  $\psi' \overline{\pi} = 1 \quad (6)$

B. We seek to Establish That

$$\lim_{M \rightarrow \infty} P^M = \overline{\pi}_{(N \times 1)} \psi'_{(1 \times N)} \quad (7)$$

Note That  $\overline{\pi} \psi' = [\overline{\pi}_{(N \times 1)} \overline{\pi}_{(N \times 1)} \dots \overline{\pi}_{(N \times 1)}]_{(N \times N)} \quad (8)$

Use The Decomposition

$$P = T \Lambda T^{-1} \quad (9)$$

Where  $\Lambda_{(N \times N)}$  is a diagonal matrix with the eigenvalues of  $P$  along its principle diagonal and where the columns of  $T_{(N \times N)}$  are the corresponding eigenvectors of  $P$ . Order the eigenvalues in  $\Lambda$  so that  $\lambda_1 = 1$  is in the  $(1, 1)$  position.

~~Since the first row, 2, 3, ..., N satisfy~~

Since  $|\lambda_i| < 1$  for  $i = 2, 3, \dots, N$ , it follows that

$$\lim_{M \rightarrow \infty} \Lambda^M = \begin{bmatrix} 1 & 0 & \cdots & 0 \\ 0 & 0 & & \vdots \\ \vdots & & \ddots & \vdots \\ 0 & \cdots & 0 & 0 \end{bmatrix}_{(N \times N)} \quad \text{and thus}$$

$$\lim_{M \rightarrow \infty} P^M = T \begin{bmatrix} 1 & 0 & \cdots & 0 \\ 0 & 0 & & \vdots \\ \vdots & & \ddots & \vdots \\ 0 & \cdots & 0 & 0 \end{bmatrix} T^{-1} = X \cdot Y'(q)$$

where  $X_{(N \times 1)}$  is the first column of  $P$  and  $Y'_{(1 \times N)}$  is the first row of  $P$ .

Since  $X$  is 1<sup>st</sup> col of  $T$  it is the e-vector  
of  $P$  associated with the ~~e-e~~ e-value  
 $\lambda_1 = 1$ . It is thus proportional to  $\pi$ :

$$X = \beta \pi \quad (10)$$

Where  $\beta \neq 0$  is a scalar.

(Transpose of) Since  $Y'$  is the first row of  $T^{-1}$  it is the  
e-vector of  $P'$  associated w/ the e-value  $\lambda_1 = 1$ .

Thus

$$P'Y = Y \quad (11)$$

Now, since  $P'\Psi = \Psi \quad (1)$

$Y$  must be proportional to  $\Psi$ :

$$Y = \alpha \Psi \quad (12)$$

Where  $\alpha \neq 0$  is a scalar.

Use (10) and (12) in RHS(9) to get



3.5

~~3.5~~

$$\lim_{M \rightarrow \infty} P^M = (\beta \alpha) \pi \psi' \quad (13)$$

$$\text{From } E(\xi_{n+1} | \xi_n) = P^n \xi_n \quad (14)$$

Note that  $P^M$  is itself a TRANSITION PROBABILITY MATRIX.

For example, is

$$P^M = \begin{bmatrix} \phi_{11} & \phi_{12} & \cdots & \phi_{1N} \\ \phi_{21} & \ddots & \ddots & \vdots \\ \vdots & & & \vdots \\ \phi_{N1} & \cdots & \phi_{NN} \end{bmatrix}$$

$\phi_{12} = \text{Prob}\{S_{n+1}=2 | S_n=1\}$   
 $\phi_{13} = \text{Prob}\{S_{n+1}=3 | S_n=1\}$   
 etc

elements of each of the  
Thus The Columns of  $P^M$  must sum TO 1.

Note from (8) That The elements of each of  
The Columns of  $\pi \psi'$  must also sum TO 1.

Thus, in (13),  $\beta \alpha = 1$  and (7)

is established.

QED (7)

A.)

$$\hat{y}_t = c + \phi \hat{y}_{t-1} + \varepsilon_t \quad (1)$$

$$(1 - \phi L) \hat{y}_t = c + \varepsilon_t$$

$$\hat{y}_t = \left( \frac{c}{1 - \phi L} \right) + \left( \frac{\varepsilon_t}{1 - \phi L} \right)$$

$$\boxed{\hat{y}_t = \left( \frac{c}{1 - \phi} \right) + \varepsilon_t + \phi \varepsilon_{t-1} + \phi^2 \varepsilon_{t-2} + \dots}$$

or

$$\boxed{\hat{y}_t = \left( \frac{c}{1 - \phi} \right) + \sum_{j=0}^{\infty} \phi^j \varepsilon_{t-j} \quad (4)}$$

$$\hat{y}_t = c + \phi \hat{y}_{t-1} + \varepsilon_t + \theta \varepsilon_{t-1} \quad (2)$$

$$(1 - \phi L) \hat{y}_t = c + \varepsilon_t + \theta \varepsilon_{t-1}$$

$$\hat{y}_t = \left( \frac{c}{1 - \phi L} \right) + \left( \frac{1}{1 - \phi L} \right) \varepsilon_t + \left( \frac{\theta}{1 - \phi L} \right) \varepsilon_{t-1}$$

$$\begin{aligned} \hat{y}_t &= \left( \frac{c}{1 - \phi} \right) + (\varepsilon_t + \phi \varepsilon_{t-1} + \phi^2 \varepsilon_{t-2} + \dots) \\ &\quad + \theta (\varepsilon_{t-1} + \phi \varepsilon_{t-2} + \phi^2 \varepsilon_{t-3} + \dots) \end{aligned}$$

$$\begin{aligned} \hat{y}_t &= \left( \frac{c}{1 - \phi} \right) + \varepsilon_t + (\phi + \theta) \varepsilon_{t-1} + (\phi^2 + \theta \phi) \varepsilon_{t-2} \\ &\quad + (\phi^3 + \theta^2 \phi) \varepsilon_{t-3} + \dots \end{aligned}$$

$$\boxed{\hat{y}_t = \left( \frac{c}{1 - \phi} \right) + \varepsilon_t + (\phi + \theta) \sum_{j=0}^{\infty} \phi^j \varepsilon_{t-1-j} \quad (5)}$$

(4.2) ~~(3)~~

B.)  $y_t = \mu + \varepsilon_t - \Theta \varepsilon_{t-1} \quad (\Theta | < 1) \quad (3)$

$$y_t = \mu + (1 - \Theta L) \varepsilon_t$$

$$\left(\frac{1}{1-\Theta L}\right) y_t = \left(\frac{\mu}{1-\Theta L}\right) + \varepsilon_t \quad \text{or}$$

$$y_t + \Theta y_{t-1} + \Theta^2 y_{t-2} + \dots = \left(\frac{\mu}{1-\Theta}\right) + \varepsilon_t \quad \text{or}$$

$$\boxed{y_t = -\Theta y_{t-1} - \Theta^2 y_{t-2} - \Theta^3 y_{t-3} - \dots + \left(\frac{\mu}{1-\Theta}\right) + \varepsilon_t}$$

$$\boxed{\text{or} \quad y_t = \left(\frac{\mu}{1-\Theta}\right) - \sum_{j=1}^{\infty} \Theta^j y_{t-j} + \varepsilon_t \quad (6)}$$

Eqn (6) is an infinite order AR process, that is AR( $\infty$ ).

C.)  $y_t \sim MA(1)$  so

$$y_t = \mu + \varepsilon_t + \Theta \varepsilon_{t-1} \quad (7)$$

$$\varepsilon_t \sim iid(0, \sigma^2)$$



$$\gamma_0 = E[(y_t - \mu)^2] = E[(\varepsilon_t + \theta \varepsilon_{t-1})^2]$$

$$= E[\varepsilon_t^2 + 2\theta \varepsilon_t \varepsilon_{t-1} + \theta^2 \varepsilon_{t-1}^2] = E(\varepsilon_t^2) + 2\theta E(\varepsilon_t \varepsilon_{t-1}) + \theta^2 E(\varepsilon_{t-1}^2)$$

$$\gamma_0 = \sigma^2 + \theta^2 \sigma^2$$

or  $\boxed{\gamma_0 = (1+\theta^2)\sigma^2 \quad (8)}$

$$\gamma_1 = E[(y_t - \mu)(y_{t-1} - \mu)] = E[(\varepsilon_t + \theta \varepsilon_{t-1})(\varepsilon_{t-1} + \theta \varepsilon_{t-2})]$$

$$= E[\varepsilon_t \varepsilon_{t-1} + \theta \varepsilon_{t-1}^2 + \theta \varepsilon_t \varepsilon_{t-2} + \theta^2 \varepsilon_{t-1} \varepsilon_{t-2}]$$

$$= E(\varepsilon_t \varepsilon_{t-1}) + \theta E(\varepsilon_{t-1}^2) + \theta E(\varepsilon_t \varepsilon_{t-2}) + \theta^2 E(\varepsilon_{t-1} \varepsilon_{t-2})$$

$\boxed{\gamma_1 = \theta \sigma^2 \quad (9)}$

Next  $p_j \equiv \frac{\gamma_j}{\gamma_0}$  so

$p_1 = 1$  and  $p_2 = \frac{\theta \sigma^2}{(1+\theta^2)\sigma^2} = \left(\frac{\theta}{1+\theta^2}\right) \quad (10)$

(4.4)

(3.4)

D.) We have

$$P_t = E \sum_{i=0}^{\infty} \beta^i y_{t+i} \quad (11)$$

where

$$y_t = \phi y_{t-1} + \varepsilon_t + \theta \varepsilon_{t-1} \quad (2')$$

From (2')

$$\boxed{E y_t = y_t} \quad (12.1)$$

$$E y_{t+1} = E[\phi y_t + \varepsilon_{t+1} + \theta \varepsilon_t] \text{ so}$$

$$\boxed{E y_{t+1} = \phi y_t + \theta \varepsilon_t} \quad (12.2)$$

$$E y_{t+2} = E[\phi y_{t+1} + \varepsilon_{t+2} + \theta \varepsilon_{t+1}] = \phi E y_{t+1} \text{ so}$$

$$\boxed{E y_{t+2} = \phi^2 y_t + \phi \theta \varepsilon_t} \quad (12.3)$$

$$E y_{t+3} = E[\phi y_{t+2} + \varepsilon_{t+3} + \theta \varepsilon_{t+2}] = \phi E y_{t+2} \text{ so}$$

$$\boxed{E y_{t+3} = \phi^3 y_t + \phi^2 \theta \varepsilon_t} \quad (12.4)$$



4.5 3.5

Thus, in general

$$\mathbb{E}_t Y_{t+j} = \phi^j Y_t + \theta \phi^{j-1} \varepsilon_t \text{ for } j=1, 2, \dots \quad (12)$$

So (11) Becomes

~~Probabilistic~~

$$P_t = Y_t + \beta \phi Y_t + \beta \phi^2 Y_t + \beta \phi^3 Y_t + \dots \\ + \beta \theta \varepsilon_t + \beta^2 \theta \phi \varepsilon_t + \beta^3 \theta \phi^2 \varepsilon_t + \dots \quad \text{or}$$

$$P_t = \sum_{j=0}^{\infty} (\beta \phi)^j Y_t + \beta \theta \sum_{j=0}^{\infty} (\beta \phi)^j \varepsilon_t \quad \text{or}$$

$$P_t = \left[ \frac{1}{1 - \beta \phi} \right] Y_t + \left[ \frac{\beta \theta}{1 - \beta \phi} \right] \varepsilon_t \quad (13)$$