

II. Stationary ARMA Processes.

1.) Suppose that Y_t is covariance stationary. As in Hamilton let γ_j denote the j th autocovariance of Y_t and let ρ_j denote its j th autocorrelation.

(a) Derive γ_j and ρ_j , $j = 0, 1, 2$, for the case where Y_t is MA(1) and for the case where Y_t is MA(2).

(b) Suppose instead that Y_t is AR(1), that is, $y_t = \phi y_{t-1} + \varepsilon_t$ where ε_t is white noise. Derive γ_j and ρ_j , $j = 0, 1, 2$. What can you conclude (or conjecture) about the relationship between the autoregressive parameter, ϕ , and the autocorrelation, ρ_j .

2.) Suppose that Y_t is a stationary AR(2) process. That is

$$y_t = c + \phi_1 y_{t-1} - \phi_2 y_{t-2} + \varepsilon_t \quad (1)$$

where c, ϕ_1 , and ϕ_2 are constants and ε_t is white noise. Furthermore, suppose that $\phi_1 = (a + b)$ and $\phi_2 = ab$ where $|a| < 1$ and $|b| < 1$. Derive the MA(∞) representation of Y_t and give explicit expressions for the MA coefficients in terms of the constants a and b .

3.) Consider the following stationary ARMA(p,q) process:

$$y_t = c + \phi_1 y_{t-1} + \dots + \phi_p y_{t-p} + \varepsilon_t + \theta_1 \varepsilon_{t-1} + \dots + \theta_q \varepsilon_{t-q} \quad (1)$$

Here, c, ϕ_j , and θ_j are constants and $\varepsilon_t \text{ i.i.d.}(0, \sigma^2)$

(a) Rewrite (1) as the first-order vector difference equation

$$\xi_t = \mathbf{F} \xi_{t-1} + \mathbf{G} \mathbf{v}_t \quad (2)$$

giving complete definitions of $\xi_t, \mathbf{F}, \mathbf{G}$, and \mathbf{v}_t in terms of the parameters of the original specification in (1).

4.) The present-value model of asset prices gives that the equilibrium price of an asset is the expected present discounted value of its dividend payments. Thus, if P_t denotes the current price of the asset and $r > 0$ denotes the known constant interest rate we have the equilibrium condition

Problem 1

Part (a):

① Suppose $Y_t \sim MA(1)$.

$$\text{Thus } Y_t = \mu + \varepsilon_t + \theta \varepsilon_{t-1} \quad (1)$$

where $\varepsilon_t \sim \text{iid } (0, \sigma^2)$

Since $E(Y_t) = \mu$

$$\gamma_0 = E[(Y_t - \mu)^2] = E[\varepsilon_t^2 + 2\theta \varepsilon_t \varepsilon_{t-1} + \theta^2 \varepsilon_{t-1}^2]$$

Since ε_t is white noise $E(\varepsilon_t \varepsilon_{t-1}) = 0$ and thus

$$\gamma_0 = E(\varepsilon_t^2) + \theta^2 E(\varepsilon_{t-1}^2) \text{ or}$$

$$\gamma_0 = \sigma^2 + \theta^2 \sigma^2 \quad \text{Thus}$$

$$\boxed{\gamma_0 = (1 + \theta^2) \sigma^2} \quad (2)$$

$$\begin{aligned} \gamma_1 &= E[(Y_t - \mu)(Y_{t-1} - \mu)] = E[(\varepsilon_t + \theta \varepsilon_{t-1})(\varepsilon_{t-1} + \theta \varepsilon_{t-2})] \\ &= E[\varepsilon_t \varepsilon_{t-1} + \theta \varepsilon_t \varepsilon_{t-2} + \theta \varepsilon_{t-1}^2 + \theta^2 \varepsilon_{t-1} \varepsilon_{t-2}] \end{aligned}$$

As ε_t is white noise $E(\varepsilon_t \varepsilon_{t-1}) = E(\varepsilon_t \varepsilon_{t-2}) = E(\varepsilon_{t-1} \varepsilon_{t-2}) = 0$

And thus

$$\gamma_1 = \cancel{\theta \sigma^2} \theta E(\varepsilon_{t-1}^2) \text{ or}$$

$$\boxed{\gamma_1 = \theta \sigma^2} \quad (3)$$

(2)

$$\begin{aligned}\gamma_2 &= E[(Y_t - \mu)(Y_{t-2} - \mu)] = E[(\varepsilon_t + \theta \varepsilon_{t-1})(\varepsilon_{t-2} + \theta \varepsilon_{t-3})] \\ &= E[\varepsilon_t \varepsilon_{t-2} + \theta \varepsilon_t \varepsilon_{t-3} + \theta \varepsilon_{t-1} \varepsilon_{t-2} + \theta^2 \varepsilon_{t-1} \varepsilon_{t-3}]\end{aligned}$$

Since ε_t is white noise

$$E(\varepsilon_t \varepsilon_{t-2}) = E(\varepsilon_t \varepsilon_{t-3}) = E(\varepsilon_{t-1} \varepsilon_{t-2}) = E(\varepsilon_{t-1} \varepsilon_{t-3}) = 0$$

And thus

$$\boxed{\gamma_2 = 0} \quad (4)$$

~~Applying the same Arg~~

Since $\rho_j \equiv \frac{\gamma_j}{\gamma_0}$ we have that

$$\rho_0 = 1 \quad \text{of course}$$

$$\rho_1 = \frac{\gamma_1}{\gamma_0} = \frac{\theta \sigma^2}{(1 + \theta^2) \sigma} \quad \text{or}$$

$$\boxed{\rho_1 = \left(\frac{\theta}{1 + \theta^2} \right)} \quad (5)$$

$$\text{and } \rho_2 = \frac{\gamma_2}{\gamma_0} = \frac{0}{(1 + \theta^2) \sigma} \quad \text{or}$$

$$\boxed{\rho_2 = 0} \quad (6)$$

$$(ii) Y_t \sim MA(2)$$

Thus

$$Y_t = \mu + \varepsilon_t + \theta_1 \varepsilon_{t-1} + \theta_2 \varepsilon_{t-2} \quad (7)$$

where $\varepsilon_t \sim iid(0, \sigma^2)$

~~No = E[Y_t]~~ Since $E(Y_t) = \mu$

$$\begin{aligned} \gamma_0 &= E[(Y_t - \mu)^2] = E[(\varepsilon_t + \theta_1 \varepsilon_{t-1} + \theta_2 \varepsilon_{t-2})^2] \\ &= E[\varepsilon_t^2 + 2\theta_1 \varepsilon_t \varepsilon_{t-1} + 2\theta_2 \varepsilon_t \varepsilon_{t-2} + \theta_1^2 \varepsilon_{t-1}^2 + 2\theta_1 \theta_2 \varepsilon_{t-1} \varepsilon_{t-2} \\ &\quad + \theta_2^2 \varepsilon_{t-2}^2] \end{aligned}$$

$$\text{Since } E(\varepsilon_t^2) = E(\varepsilon_{t-1}^2) = E(\varepsilon_{t-2}^2) = \sigma^2$$

~~and E(\varepsilon_t \varepsilon_{t-1}) = 0 for~~

$$\text{and } E(\varepsilon_t \varepsilon_{t'}) = 0 \text{ for } t \neq t'$$

$$\gamma_0 = \sigma^2 + \theta_1^2 \sigma^2 + \theta_2^2 \sigma^2 \text{ or}$$

$$\boxed{\gamma_0 = [1 + \theta_1^2 + \theta_2^2] \sigma^2 \quad (8)}$$

(4)

$$\gamma_1 = E[(Y_t - \mu)(Y_{t-1} - \mu)] = E[(\varepsilon_t + \theta_1 \varepsilon_{t-1} + \theta_2 \varepsilon_{t-2})(\varepsilon_{t-1} + \theta_1 \varepsilon_{t-2} + \theta_2 \varepsilon_{t-3})]$$

Where, ~~Noting that~~ $E(\varepsilon_t \varepsilon_{t-k}) = 0$ for $k \neq 0$, so

$$\gamma_1 = E[\theta_1 \varepsilon_{t-1}^2 + \theta_2 \theta_1 \varepsilon_{t-2}^2] \quad \text{or}$$

$$\gamma_1 = \theta_1 \sigma^2 + \theta_2 \theta_1 \sigma^2 \quad \text{or}$$

$$\boxed{\gamma_1 = [\theta_1 + \theta_1 \theta_2] \sigma^2 \quad (6)}$$

$$\begin{aligned} \gamma_2 &= E[(Y_t - \mu)(Y_{t-2} - \mu)] = \\ &E[(\varepsilon_t + \theta_1 \varepsilon_{t-1} + \theta_2 \varepsilon_{t-2})(\varepsilon_{t-2} + \theta_1 \varepsilon_{t-3} + \theta_2 \varepsilon_{t-4})] \\ &= E(\theta_2 \varepsilon_{t-2}^2) \quad \text{so} \end{aligned}$$

$$\boxed{\gamma_2 = \theta_2 \sigma^2 \quad (7)}$$

Further $\rho_0 = 1$ of course

$$\rho_1 = \left[\frac{\theta_1 + \theta_1 \theta_2}{1 + \theta_1^2 + \theta_2^2} \right] \quad (8)$$

$$\text{and } \rho_2 = \left[\frac{\theta_2}{1 + \theta_1^2 + \theta_2^2} \right] \quad (9)$$

Part(b) Suppose Y_t is $AR(1)$, That is,

$$Y_t = \phi Y_{t-1} + \varepsilon_t \quad (9)$$

$$\varepsilon_t \sim iid(0, \sigma^2)$$

TAKING EXPECTATIONS

$$E(Y_t) = \phi E(Y_{t-1}) + E(\varepsilon_t) \quad \nearrow 0$$

Since Y_t is STATIONARY $E(Y_t) = E(Y_{t-1}) = \mu$ so

$$\mu = \phi \mu \text{ or}$$

$$(1 - \phi) \mu = 0. \text{ Stationarity requires } |\phi| < 1$$

$$\text{so } (1 - \phi) \neq 0 \text{ Therefore } \boxed{\mu = 0} \quad (10)$$

~~Not~~ To Derive Autocovariances it is useful to work with the $MA(\infty)$ representation. Write (9)

as

$$(1 - \phi L) Y_t = \varepsilon_t \quad \text{or}$$

$$Y_t = \left(\frac{1}{1 - \phi L} \right) \varepsilon_t \quad \text{or}$$

$$Y_t = \varepsilon_t + \phi \varepsilon_{t-1} + \phi^2 \varepsilon_{t-2} + \dots \quad (11)$$

Since $\mu = 0$ we have

$$\gamma_0 = E(Y_t^2) = E[(\varepsilon_t + \phi \varepsilon_{t-1} + \phi^2 \varepsilon_{t-2} + \dots)^2]$$

Since $E(\varepsilon_t \varepsilon_{t-j}) = 0$ for $t \neq \tau$ this gives

$$\gamma_0 = E[(\varepsilon_t^2 + \phi^2 \varepsilon_{t-1}^2 + (\phi^2)^2 \varepsilon_{t-2}^2 + \dots)]$$

$$= \sigma^2 + \phi^2 \sigma^2 + \phi^4 \sigma^2 + \dots \quad \text{or}$$

$$\gamma_0 = (1 + \phi^2 + \phi^4 + \phi^6 + \dots) \sigma^2$$

or Since $|\phi^2| < 1$

$$\gamma_0 = \left[\frac{1}{1 - \phi^2} \right] \sigma^2 \quad (12)$$

Next, consider

$$\gamma_1 = E(Y_t Y_{t-1}) = E[(\varepsilon_t + \phi \varepsilon_{t-1} + \phi^2 \varepsilon_{t-2} + \dots)(\varepsilon_{t-1} + \phi \varepsilon_{t-2} + \dots)]$$

Since $E(\varepsilon_t \varepsilon_{t-j}) = 0$ for $j = 1, 2, 3, \dots$ This gives

$$\gamma_1 = E[\phi(\varepsilon_{t-1} + \phi \varepsilon_{t-2} + \phi^2 \varepsilon_{t-3} + \dots)(\varepsilon_{t-1} + \phi \varepsilon_{t-2} + \dots)]$$



$$\text{or } \gamma_1 = \phi E[\varepsilon_{t-1}^2 + \phi^2 \varepsilon_{t-2}^2 + \phi^4 \varepsilon_{t-3}^2 + \dots]$$

$$\gamma_1 = \phi [1 + \phi^2 + \phi^4 + \dots] \sigma^2 \quad \text{and Thus}$$

$$\boxed{\gamma_1 = \left[\frac{\phi}{1 - \phi^2} \right] \sigma^2 \quad (13)}$$

Next, Consider

$$\gamma_2 = E(Y_t Y_{t-2}) = E[(\varepsilon_t + \phi \varepsilon_{t-1} + \phi^2 \varepsilon_{t-2} + \dots)(\varepsilon_{t-2} + \phi \varepsilon_{t-3} + \dots)]$$

$$\text{Since } E(\varepsilon_t \varepsilon_{t-j}) = 0 \text{ for } j = 1, 2, 3, \dots$$

$$\text{and } E(\varepsilon_{t-1} \varepsilon_{t-j}) = 0 \text{ for } j = 2, 3, 4, \dots$$

$$\gamma_2 = \phi^2 E[(\varepsilon_{t-2} + \phi \varepsilon_{t-3} + \phi^2 \varepsilon_{t-4} + \dots)^2] \quad \text{or}$$

$$\gamma_2 = \phi^2 [1 + \phi^2 + \phi^4 + \dots] \sigma^2 \quad \text{or}$$

$$\boxed{\gamma_2 = \left[\frac{\phi^2}{1 - \phi^2} \right] \sigma^2 \quad (14)}$$

(8)

Since $\rho_j = \frac{\gamma_j}{\gamma_0}$ from (12) (13) and (14)
we have

$$\rho_0 = 1 \text{ of course,}$$

$$\rho_1 = \phi$$

$$\rho_2 = \phi^2$$

we can infer (or conjecture) that

For the AR(1) Process

$$\boxed{\rho_j = \phi^j}$$

Problem 2. Consider

$$y_t = c + \phi_1 y_{t-1} + \phi_2 y_{t-2} + \varepsilon_t \quad (1)$$

write as

$$y_t - \phi_1 y_{t-1} + \phi_2 y_{t-2} = c + \varepsilon_t \quad (2)$$

Since $\phi_1 = (a+b)$ and $\phi_2 = ab$ we have

$$y_t - (a+b)y_{t-1} + ab y_{t-2} = c + \varepsilon_t \quad \text{or}$$

$$(1 - (a+b)L + abL^2)y_t = c + \varepsilon_t \quad \text{or}$$

$$(1 - aL)(1 - bL)y_t = c + \varepsilon_t \quad \text{or}$$

$$y_t = \frac{c}{(1-a)(1-b)} + \frac{\varepsilon_t}{(1-aL)(1-bL)} \quad (3)$$

Note that, since $|b| < 1$ we can write

$$\frac{1}{(1-aL)(1-bL)} \varepsilon_t = \left(\frac{1}{1-aL} \right) [1 + bL + b^2L^2 + \dots] \varepsilon_t \quad \text{or}$$

$$\frac{1}{(1-aL)(1-bL)} \varepsilon_t = \left(\frac{1}{1-aL} \right) [\varepsilon_t + b\varepsilon_{t-1} + b^2\varepsilon_{t-2} + \dots]$$

And, since $|a| < 1$ we have

$$\frac{1}{(1-aL)(1-bL)} \varepsilon_t = [1 + aL + a^2L^2 + a^3L^3 + \dots] [\varepsilon_t + b\varepsilon_{t-1} + b^2\varepsilon_{t-2} + \dots]$$

$$= \varepsilon_t + b\varepsilon_{t-1} + b^2\varepsilon_{t-2} + b^3\varepsilon_{t-3} + \dots + a\varepsilon_{t-1} + ab\varepsilon_{t-2} + ab^2\varepsilon_{t-3} + \dots + a^2\varepsilon_{t-2} + a^2b\varepsilon_{t-3} + \dots$$

$$= \varepsilon_t + (a+b)\varepsilon_{t-1} + (b^2+ab+a^2)\varepsilon_{t-2} + (b^3+ab^2+a^2b+a^3)\varepsilon_{t-3} + (b^4+ab^3+a^2b^2+a^3b+a^4)\varepsilon_{t-4} + (b^5+ab^4+a^2b^3+a^3b^2+a^4b+a^5)\varepsilon_{t-5} + \dots$$

There are some issues relating to the output of this model which should be noted and discussed further. The most notable concern is the possibility of multicollinearity between the dependent variable and my variable capturing institutional quality. In order to investigate the validity of these endogeneity concerns I will first examine the relationship between the possible multiplicity between the two variables. The possible multiplicity between the two variables is a concern because if the two variables are highly correlated, the model may suffer from multicollinearity, which would lead to unreliable estimates of the coefficients. This is a common issue in econometric models, especially when dealing with time series data. To address this, I will use various diagnostic tests such as the Variance Inflation Factor (VIF) to assess the degree of multicollinearity. If the VIF values are high, it indicates a strong correlation between the variables, and I may need to consider alternative specifications or data transformations to improve the model's reliability.

Collecting we have

$$\frac{y}{\sqrt{t}} = \frac{c}{(1-a)(1-b)} + \sum_{j=0}^{\infty} \psi_j \varepsilon_{t-j} \quad (4)$$

where $\psi_0 = 1$

$$\psi_1 = (a+b)$$

$$\psi_2 = (b^2 + ab + a^2)$$

$$\psi_3 = (b^3 + ab^2 + a^2b + a^3)$$

$$\psi_j = (b^j + ab^{j-1} + a^2b^{j-2} + \dots + a^{j-1}b + a^j)$$

Problem 3 y_t follows a stationary ARMA(p, q) process:

$$y_t = c + \phi_1 y_{t-1} + \dots + \phi_p y_{t-p} + \varepsilon_t + \theta_1 \varepsilon_{t-1} + \dots + \theta_q \varepsilon_{t-q} \quad (1)$$

First, take expectations and note that

$$E(y_t) = E(y_{t-1}) = \dots = E(y_{t-p}) = \mu$$

so that

$$\mu = c + \phi_1 \mu + \dots + \phi_p \mu \text{ or}$$

$$c = (1 - \phi_1 - \phi_2 - \dots - \phi_p) \mu \quad (2)$$

This in (1) gives

$$z_t = \phi_1 z_{t-1} + \dots + \phi_p z_{t-p} + \varepsilon_t + \theta_1 \varepsilon_{t-1} + \dots + \theta_q \varepsilon_{t-q} \quad (3)$$

where $z_{t-j} = (y_{t-j} - \mu)$ for $j = 0, 1, 2, \dots, p$.

Now (4) can be written

$$\bar{z}_t = F_{(p \times p)} \bar{z}_{t-1} + (G_{(p \times q)} V_{t(q \times 1)}) \quad (5)$$

where

$$\bar{z}_t = \begin{bmatrix} y_t - \mu \\ y_{t-1} - \mu \\ \vdots \\ y_{t-(p-1)} - \mu \end{bmatrix} \quad (p \times 1)$$

$$F = \begin{bmatrix} \phi_1 & \phi_2 & \phi_3 & \dots & \phi_p \\ 1 & 0 & 0 & \dots & 0 \\ 0 & 1 & 0 & \dots & 0 \\ 0 & 0 & 1 & \dots & 0 \\ \vdots & \vdots & \vdots & \ddots & \vdots \\ 0 & \dots & 0 & 0 & 1 \end{bmatrix} \quad (p \times p)$$

$$G = \begin{bmatrix} 1 & \theta_1 & \theta_2 & \dots & \theta_q \\ 0 & 0 & 0 & \dots & 0 \\ \vdots & \vdots & \vdots & \ddots & \vdots \\ 0 & 0 & 0 & \dots & 0 \end{bmatrix} \quad [p \times (q+1)]$$

$$V_t = [\varepsilon_t \ \varepsilon_{t-1} \ \varepsilon_{t-2} \ \dots \ \varepsilon_{t-q}]^T \quad [(q+1) \times 1]$$

Define $\xi_t = \begin{bmatrix} z_t \\ z_{t-1} \\ \vdots \\ z_{t-p+1} \end{bmatrix}_p$ so that $\xi_{t-1} = \begin{bmatrix} z_{t-1} \\ z_{t-2} \\ \vdots \\ z_{t-p} \end{bmatrix}_{(p \times 1)}$

and (3) can be written as

$$\xi_t = \begin{bmatrix} \phi_1 & \phi_2 & \dots & \phi_p \\ 1 & 0 & \dots & 0 \\ 0 & 1 & \dots & 0 \\ \vdots & \vdots & \ddots & \vdots \\ 0 & 0 & \dots & 1 \end{bmatrix} \xi_{t-1} + (\xi_t + \theta_1 \xi_{t-1} + \dots + \theta_q \xi_{t-q}) \quad (4)$$

Table 3: Specification Test

	Test Statistic	p-value
Linear Model	0.041E-20	0.01
Nonlinear Model	0.67E-28	0.00

let $V_t = \begin{bmatrix} \xi_t \\ \xi_{t-1} \\ \vdots \\ \xi_{t-q} \end{bmatrix}_{(q+1) \times 1}$

and $G = \begin{bmatrix} 1 & \theta_1 & \theta_2 & \dots & \theta_q \\ 0 & 1 & 0 & \dots & 0 \\ \vdots & \vdots & \vdots & \ddots & \vdots \\ 0 & 0 & 0 & \dots & 1 \end{bmatrix}_{(p \times (q+1))}$

Problem 4

Let $R \equiv 1+r$ hence $(\frac{1}{1+r}) = R^{-1}$ and write (1) as

$$P_t = E_t \sum_{j=1}^{\infty} R^{-j} d_{t+j} \quad (2)$$

or

$$P_t = E_t [R^{-1} d_{t+1} + R^{-2} d_{t+2} + R^{-3} d_{t+3} + \dots] \quad (2')$$

$$(a) d_t = \mu + \varepsilon_t + \theta \varepsilon_{t-1}$$

$$\text{Here } E_t d_{t+1} = \mu + \theta \varepsilon_t$$

$$\text{and } E_t d_{t+j} = \mu \quad \text{for } j = 2, 3, 4, \dots$$

Thus (2') gives

$$P_t = R^{-1}(\mu + \theta \varepsilon_t) + R^{-2}\mu + R^{-3}\mu + \dots$$

$$P_t = R^{-1}\theta \varepsilon_t + R^{-1}[\mu + R^{-1}\mu + R^{-2}\mu + \dots] \quad (3)$$

Since $|R^{-1}| < 1$

$$[\mu + R^{-1}\mu + R^{-2}\mu + \dots] = \left[\frac{1}{1-R^{-1}} \right] \mu = \left(\frac{1+r}{r} \right) \mu \quad (4)$$

Subst from (4) into (3) to get

$$P_t = R^{-1} \Theta \varepsilon_t + \frac{1}{1+r} \left[\frac{1+r}{r} \right] \mu \quad \text{or}$$

$$\boxed{P_t = \frac{\mu}{r} + \Theta R^{-1} \varepsilon_t} \quad (5)$$

$$(b) d_t = \mu + \varepsilon_t + \Theta_1 \varepsilon_{t-1} + \Theta_2 \varepsilon_{t-2}$$

$$\text{Here } E_t d_{t+1} = \mu + \Theta_1 \varepsilon_t + \Theta_2 \varepsilon_{t-1}$$

$$E_t d_{t+2} = \mu + \Theta_2 \varepsilon_t$$

$$\text{AND } E_t d_{t+j} = \mu \quad \text{for } j=3, 4, 5, \dots$$

Substituting into (2') gives

$$P_t = R^{-1} [\mu + \Theta_1 \varepsilon_t + \Theta_2 \varepsilon_{t-1}] + R^{-2} [\mu + \Theta_2 \varepsilon_t] \\ + R^{-3} \mu + R^{-4} \mu + R^{-5} \mu + \dots \quad \text{or}$$

$$P_t = R^{-1} [\Theta_1 \varepsilon_t + \Theta_2 \varepsilon_{t-1}] + R^{-2} \Theta_2 \varepsilon_t + \\ R^{-1} [\mu + R^{-1} \mu + R^{-2} \mu + \dots] \quad (6)$$

Using (4) in (6) gives

$$P_t = R^{-1} [\Theta_1 + R^{-1} \Theta_2] \varepsilon_t + R^{-1} \Theta_2 \varepsilon_{t-1} \\ + \frac{1}{1+r} \left[\frac{1+r}{r} \right] \mu \quad \text{or}$$

$$P_t = \frac{\mu}{r} + R^{-1} [\Theta_1 + R^{-1} \Theta_2 \varepsilon_t + R^{-1} \Theta_2 \varepsilon_{t-1}] \quad (7)$$

IN (5) and (7) ~~NOTE~~ Note That if Dividends ARE MA(1) The Asset Price is MA(1) and if Dividends are MA(2) The Asset Price is MA(2).

~~$$(C) \quad d_t = c + \phi d_{t-1} + \varepsilon_t \quad (8)$$~~

write

~~$$X_t \equiv \begin{bmatrix} c \\ d_t \end{bmatrix} \text{ and } (8) \text{ becomes}$$~~

~~$$X_t = F X_{t-1} + V_t$$~~

~~$$\text{where } F \equiv \begin{bmatrix} 1 & 0 \\ 1 & \phi \end{bmatrix} \text{ and } V_t = \begin{bmatrix} 0 \\ \varepsilon_t \end{bmatrix} \quad (9)$$~~

~~ASB~~

$$c) d_t = c + \phi d_{t-1} + \varepsilon_t \quad (8)$$

"Brute-Force" method.

Work From (2')

$$P_t = E_t [R^{-1} d_{t+1} + R^{-2} d_{t+2} + R^{-3} d_{t+3} + \dots] \quad (2')$$

From (8) it follows that

$$E_t d_{t+1} = c + \phi d_t \quad (9.1)$$

$$E_t d_{t+2} = c + \phi E_t d_{t+1} \quad \text{so}$$

$$E_t d_{t+2} = c + \phi c + \phi^2 d_t \quad (9.2)$$

$$E_t d_{t+3} = c + \phi E_t d_{t+2} \quad \text{so}$$

$$E_t d_{t+3} = c + \phi c + \phi^2 c + \phi^3 d_t \quad (9.3)$$

$$E_t d_{t+4} = c + \phi c + \phi^2 c + \phi^3 c + \phi^4 d_t \quad (9.4)$$

etc.

Substitute From (9.1) - (9.4) into (2') to get

$$P_t = R^{-1}(c + \phi d_t) + R^{-2}(c + \phi c + \phi^2 d_t) + \\ R^{-3}(c + \phi c + \phi^2 c + \phi^3 d_t) + R^{-4}(c + \phi c + \phi^2 c + \phi^3 c + \phi^4 d_t) \\ + \dots$$

or 

$$P_z = R^{-1}c + R^{-2}(1+\phi)c + R^{-3}(1+\phi+\phi^2)c + R^{-4}(1+\phi+\phi^2+\phi^3)c + \dots \\ + R^{-1}\phi[1 + R^{-1}\phi + R^{-2}\phi^2 + \dots]dz \quad (10)$$

Now, Note That

$$R^{-1}\phi[1 + R^{-1}\phi + R^{-2}\phi^2 + \dots] = \frac{\phi}{1+r} \left[\frac{1}{1 - R^{-1}\phi} \right] = \frac{\phi}{1+r} \left[\frac{1}{1 - (\frac{\phi}{1+r})} \right] \\ = \frac{\phi}{1+r} \left[\frac{1}{(\frac{1+r-\phi}{1+r})} \right] = \frac{\phi}{1+r} \cdot \frac{1+r}{1+r-\phi} = \frac{\phi}{1+r-\phi}$$

This in (10) gives

$$P_z = R^{-1}c \left[1 + R^{-1}(1+\phi) + R^{-2}(1+\phi+\phi^2) + R^{-3}(1+\phi+\phi^2+\phi^3) + \dots \right] \\ + \left(\frac{\phi}{1+r-\phi} \right) dz \quad (11)$$

Now, Notice That

$$1+\phi = (1 + \phi + \phi^2 + \phi^3 + \dots) \\ - (\phi^2 + \phi^3 + \phi^4 + \dots)$$

$$= \left(\frac{1}{1-\phi} \right) - \left(\frac{\phi^2}{1-\phi} \right) \quad \text{so}$$

$$(1+\phi) = \left(\frac{1-\phi^2}{1-\phi} \right) \quad (12.1)$$

A similar argument gives

$$1 + \phi + \phi^2 = \frac{1 - \phi^3}{1 - \phi} \quad (12.2)$$

$$1 + \phi + \phi^2 + \phi^3 = \frac{1 - \phi^4}{1 - \phi} \quad (12.3)$$

etc.

Using $1 = \frac{1 - \phi}{1 - \phi}$ and (12.1) - (12.3) we have that

$$R^{-1}C [1 + R^{-1}(1 + \phi) + R^{-2}(1 + \phi + \phi^2) + \dots] =$$

$$R^{-1}C \left[\frac{1 - \phi}{1 - \phi} + R^{-1} \left(\frac{1 - \phi^2}{1 - \phi} \right) + R^{-2} \left(\frac{1 - \phi^3}{1 - \phi} \right) + \dots \right] =$$

$$CR^{-1} \left[\frac{1}{1 - \phi} + R^{-1} \left(\frac{1}{1 - \phi} \right) + R^{-2} \left(\frac{1}{1 - \phi} \right) + \dots \right]$$

$$- CR^{-1} \left[\frac{\phi}{1 - \phi} + \frac{R^{-1}\phi^2}{1 - \phi} + \frac{R^{-2}\phi^3}{1 - \phi} + \dots \right]$$

$$= CR^{-1} \left(\frac{1}{1 - \phi} \right) \left(\frac{1}{1 - R^{-1}} \right) - CR^{-1} \left(\frac{\phi}{1 - \phi} \right) \left(\frac{1}{1 - R^{-1}\phi} \right)$$

$$= \left(\frac{C}{1 + r} \right) \left(\frac{1}{1 - \phi} \right) \left(\frac{1 + r}{r} \right) - \left(\frac{C}{1 + r} \right) \left(\frac{\phi}{1 - \phi} \right) \left(\frac{1 + r}{1 + r - \phi} \right)$$

$$= \left(\frac{C}{1 - \phi} \right) \left[\frac{1}{r} - \frac{\phi}{1 + r - \phi} \right]$$



$$= \left(\frac{C}{1-\phi} \right) \left[\frac{1+r-\phi-\phi r}{r(1+r-\phi)} \right] = \frac{C}{1-\phi} \left[\frac{(1+r)(1-\phi)}{r(1+r-\phi)} \right]$$

$$= \frac{1+r}{(1+r-\phi)r} C$$

~~using this~~ Thus

$$R^+ C \left[1 + R^{-1}(1+\phi) + R^{-2}(1+\phi+\phi^2) + \dots \right]$$

$$= \frac{1+r}{(1+r-\phi)r} C \quad (13)$$

use (13) in (11) to get

$$P_z = \frac{1+r}{(1+r-\phi)r} C + \left(\frac{\phi}{1+r-\phi} \right) d \quad (14)$$

Method using Selector Matrix

Define

$$X_t = \begin{bmatrix} c \\ d_t \end{bmatrix} \quad F = \begin{bmatrix} 1 & 0 \\ 1 & \phi \end{bmatrix} \quad V_t = \begin{bmatrix} 0 \\ \varepsilon_t \end{bmatrix}$$

(2x1) (2x2) (2x1)

and $S = \begin{bmatrix} 0 & 1 \end{bmatrix}$ so that (8)

(1x2)

(Can be written as

$$X_t = F X_{t-1} + V_t \quad (15)$$

and

$$d_t = S \cdot X_t \quad (16)$$

Begin From (2)

$$P_t = E_t \sum_{j=1}^{\infty} R^{-j} E_t d_{t+j} \quad (2)$$

Which we can now write as



$$P_t = E_t \sum_{j=1}^{\infty} R^{-j} S X_{t+j} \quad \text{or, Since } R^{-1} \text{ is a scalar,}$$

$$P_t = S \sum_{j=1}^{\infty} R^{-j} E_t X_{t+j} \quad (17)$$

$$\text{Since } E_t V_{t+1} = E_t V_{t+2} = E_t V_{t+3} = \dots = 0$$

We can use (15) in (17) to get

$$P_t = S \sum_{j=1}^{\infty} R^{-j} F^j X_t \quad \text{or}$$

$$P_t = S R^{-1} F \left[\sum_{j=0}^{\infty} R^{-j} F^j \right] X_t \quad (18)$$

The Eigenvalues of F are 1 and ϕ Both of which are less than $(R^{-1})^{-1} = R = 1+r$ in Absolute Value.

Therefore we can invoke HAMILTON'S PROPOSITION 1.3

(page 20) and write (18) as

$$P_t = S R^{-1} F [I - R^{-1} F]^{-1} X_t \quad (19)$$

Since $[I - R^{-1}F] = \begin{bmatrix} (1 - R^{-1}) & 0 \\ -R^{-1} & (1 - R^{-1}\phi) \end{bmatrix}$ it follows that

$$[I - R^{-1}F]^{-1} = \frac{1}{(1 - R^{-1})(1 - R^{-1}\phi)} \begin{bmatrix} (1 - R^{-1}\phi) & 0 \\ +R^{-1} & (1 - R^{-1}) \end{bmatrix} \quad (20)$$

Use (20) in (19) to write

$$P_z = \begin{bmatrix} 0 & I \end{bmatrix} R^{-1} \begin{bmatrix} 1 & 0 \\ 1 & \phi \end{bmatrix} \frac{1}{(1 - R^{-1})(1 - R^{-1}\phi)} \begin{bmatrix} 1 - R^{-1}\phi & 0 \\ +R^{-1} & (1 - R^{-1}) \end{bmatrix} \begin{bmatrix} C \\ d_z \end{bmatrix}$$

or

$$P_z = \begin{bmatrix} 0 & I \end{bmatrix} \frac{R^{-1}}{(1 - R^{-1})(1 - R^{-1}\phi)} \begin{bmatrix} (1 - R^{-1}\phi) & 0 \\ (1 - R^{-1}\phi + \phi R^{-1}) & \phi(1 - R^{-1}) \end{bmatrix} \begin{bmatrix} C \\ d_z \end{bmatrix}$$

$(1 \times 2) \quad (1 \times 1) \quad (2 \times 2) \quad (2 \times 1)$

or

$$P_z = \frac{R^{-1}}{(1 - R^{-1})(1 - R^{-1}\phi)} \begin{bmatrix} 0 & I \end{bmatrix} \begin{bmatrix} (1 - R^{-1}\phi)C \\ C + \phi(1 - R^{-1}\phi)d_z \end{bmatrix}$$

or

$$P_z = \frac{R^{-1}}{(1 - R^{-1})(1 - R^{-1}\phi)} \begin{bmatrix} C + \phi(1 - R^{-1}\phi)d_z \end{bmatrix} \quad (21)$$

or



Note That $\left[\frac{R^{-1}\phi}{1-R^{-1}\phi} \right] = \frac{R}{\bar{R}} \left[\frac{R^{-1}\phi}{1-R^{-1}\phi} \right] = \frac{\phi}{R-\phi} = \frac{\phi}{1+r-\phi}$

This in (21) gives

$$P_z = \frac{R^{-1}}{(1-R^{-1})(1-R^{-1}\phi)} C + \left(\frac{\phi}{1+r-\phi} \right) d_z \quad (22)$$

$$\frac{R^{-1}}{(1-R^{-1})(1-R^{-1}\phi)} = \frac{\left(\frac{1}{1+r} \right)}{\left(\frac{r}{1+r} \right) \left(\frac{1+r-\phi}{1+r} \right)} = \frac{(1+r)}{(1+r-\phi)r}$$

This in (22) gives

$$P_z = \frac{(1+r)}{r(1+r-\phi)} C + \left(\frac{\phi}{1+r-\phi} \right) d_z \quad (23)$$

cf (14)