

Problem Set 3

Global Poverty and Economic Development

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1 Data Exercise - IV estimation

1.1

dta <- read.csv("./Data_IV.csv")
table(dta[dta\$assignment==0,]\$treated)
table(dta[dta\$assignment==1,]\$treated)</pre>

Table 1: IV Table

	treated = 0	treated = 1
assignment = 0	2499	0
assignment=1	1200	1301

$$\mathbb{E}[treated = 0 \mid assignment = 1] = \frac{1200}{1200 + 1301} = \frac{1200}{2501} \approx 0.480$$

1.2

model1 <- lm(y ~ treated, data=dta)</pre>

Results: see Table 2 column (1).

1.3

model2 <- lm(y ~ assignment, data=dta)</pre>

Results: see Table 2 column (2).

1.4

model3 <- ivreg(y ~ treated | assignment, data=dta)</pre>

Results: see Table 2 column (3).

Table 2: IV Regression

	Dependent variable:					
	OLS	y IIT	LATE	treated 1st Stage		
	(1)	(2)	(3)	(4)		
treated	6.583*** (0.120)		16.439*** (0.311)			
assignment	,	8.551*** (0.057)	,	0.520*** (0.010)		
Constant	2.536*** (0.061)	-0.028 (0.040)	-0.028 (0.114)	-0.000 (0.007)		
Observations R ²	5,000 0.374	5,000 0.820	5,000 -0.465	5,000 0.351		
Note:	*n<0.1: **n<0.05: ***n<0.01					

Note:

1.5

From 1.1, we see that the proportion of those assigned to treatment did not receive treatment is 0.480. We can further derive that the proportion of those assigned to treatment received treatment is 0.520 (1-0.480), which is also consistent with the first-stage regression (see Table 2 column (4)).

$$\begin{split} \delta^{Wald} &= \frac{\mathbb{E}[Y_i \mid Z=1] - \mathbb{E}[Y_i \mid Z=0]}{\mathbb{E}[D_i \mid Z=1] - \mathbb{E}[D_i \mid Z=0]} \\ &= \frac{\text{ITT}}{\text{1st stage}} \\ &= \frac{8.551}{0.520} \\ &\approx 16.44 \end{split}$$

1.6

Table 2 shows that the estimate in column (2) is different from the estimate in column (3). The OLS reduced-form regression captures the intent-to-treat effect while the IV regression captures the local average treatment effect. From 1.1 we see that the assignment of treatment is not strictly implemented. There are some units who are assigned to treatment but do not receive treatment. Therefore, the observed treatment indicator is not "clean". If we divide intent-to-treat effect by the proportion of units who were assigned to treatment and received treatment, we can derive the same result as in IV regression.

1.7

```
model5 \leftarrow lm(y ~assignment, data=dta, subset=dta$sex==1)
model6 <- lm(y ~ assignment, data=dta, subset=dta$sex==0)</pre>
```

Table 3: ITT and LATE Comparison

	Dependent variable:					
			у			
	ITT		LATE			
	sex=0	sex=1	sex=0	sex=1	both	
	(1)	(2)	(3)	(4)	(5)	
assignment	7.143*** (0.089)	9.959*** (0.043)				
treated	, ,	, ,	8.928*** (0.058)	41.392*** (2.012)	8.928*** (0.430)	
sex			,	, ,	-0.062 (0.344)	
treated:sex					32.464*** (1.493)	
Constant	0.003 (0.063)	$-0.059* \ (0.031)$	0.003 (0.033)	-0.059 (0.342)	0.003 (0.244)	
Observations R^2	2,498 0.721	2,502 0.955	2,498 0.925	2,502 -4.641	5,000 -2.318	
Note:	*p<0.1; **p<0.05; ***p<0.01					

1.8

Regression results: Table 3 column (1) and column (2). Two separate IV regressions capture the local average treatment effects across different gender groups.

1.9

$$LATE = \frac{ITT}{ITT_D}$$

ullet Note: ITT_D is the proportion of units who are assigned to treatment and receive treatment (i.e. the proportion of compliers).

In Table 3 from column (3) to column (4), we see that the difference between ITT and LATE among sex=1 is significantly larger than the difference between ITT and LATE among sex=0. We can infer that the proportion of compliers among sex=0 is much larger than the proportion of compliers among sex=1.

1.10

```
model9 <- ivreg(y ~ treated + sex | assignment + sex, data=dta)</pre>
```

In Table 3 from column (3) to column (5), the estimate from the IV regression after adding interaction term is quite close to the estimate from the IV regression for sex=0.

2 Theory - Centralized vs. Decentralized Corruption

Centralized Corruption

2.1

$$\begin{split} \Pi &= P \times (Q(P) - c) \\ &= P \times \left(\frac{\alpha}{2} - \beta \times P - c\right) \end{split}$$

$$\max_{P} \quad P \times \left(\frac{\alpha}{2} - \beta \times P - c\right)$$

FOC w.r.t. P:

$$\frac{\alpha}{2} - 2\beta \times P - c = 0 \implies P^* = \frac{1}{2\beta} \left(\frac{\alpha}{2} - c \right)$$
$$p^C = \frac{P^*}{N} = \frac{1}{2\beta N} \left(\frac{\alpha}{2} - c \right)$$

The monopolist sets the bribe level at $p^C=rac{1}{2eta N}\left(rac{lpha}{2}-c
ight)$ in each of the N identical checkpoints.

2.2

$$\begin{split} \pi^C &= P^* \times (Q(P^*) - c) \\ &= P^* \times \left(\frac{\alpha}{2} - \beta \times P^* - c\right) \\ &= \frac{1}{2\beta} \left(\frac{\alpha}{2} - c\right) \times \left(\frac{\alpha}{2} - \frac{1}{2} \left(\frac{\alpha}{2} - c\right) - c\right) \\ &= \frac{1}{2\beta} \left(\frac{\alpha}{2} - c\right) \times \left(\frac{\alpha}{4} - \frac{c}{2}\right) \\ &= \frac{1}{4\beta} \left(\frac{\alpha}{2} - c\right)^2 \end{split}$$

The total profits from centralized corruption is $\pi^C = \frac{1}{4\beta} \left(\frac{\alpha}{2} - c\right)^2$

Decentralized Corruption

2.3

Setup:

- 2 checkpoints set bribe level at 0
- ullet N-3 checkpoints set bribe level at p^C

$$\pi_{i} = p_{i} \times (Q(P) - c)$$

$$= p_{i} \times \left(\frac{\alpha}{2} - \beta \times P - c\right)$$

$$= p_{i} \times \left(\frac{\alpha}{2} - \beta \times (p_{i} + (N - 3) \cdot p^{C}) - c\right)$$

$$= p_{i} \times \left(\frac{N + 3}{2N} \left(\frac{\alpha}{2} - c\right) - \beta p_{i}\right)$$

$$\max_{p_i} \quad p_i \times \left(\frac{N+3}{2N} \left(\frac{\alpha}{2} - c\right) - \beta p_i\right)$$

FOC w.r.t. p_i :

$$\frac{N+3}{2N}\left(\frac{\alpha}{2}-c\right)-2\beta p_i=0 \implies p_i=\frac{N+3}{4\beta N}\left(\frac{\alpha}{2}-c\right)$$

$$\frac{p_i}{p^C} = \frac{N+3}{2} > 1 \quad (N \geq 2 \text{ by assumption}) \implies p_i > p^C$$

The checkpoint i would set $p_i=\frac{N+3}{4\beta N}\left(\frac{\alpha}{2}-c\right)$ and it would be larger than $p^C.$

2.4

Denote bribe level under decentralized corruption by p^D . In a symmetric equilibrium, $P = N \times p^D$.

$$\begin{split} \pi_i &= p_i \times (Q(P) - c) \\ &= p_i \times \left(\frac{\alpha}{2} - \beta \times P - c\right) \\ &= p_i \times \left(\frac{\alpha}{2} - \beta \times \left(p_i + \sum_{j \neq i} p_j\right) - c\right) \end{split}$$

$$\max_{p_i} \quad p_i \times \left(\frac{\alpha}{2} - \beta \times \left(p_i + \sum_{j \neq i} p_j\right) - c\right)$$

FOC w.r.t. p_i :

$$\frac{\alpha}{2} - 2\beta p_i - \beta \sum_{j \neq i} p_j - c = 0$$

$$\frac{\alpha}{2} - 2\beta p^D - \beta (N - 1) \times p^D - c = 0$$

$$p^D = \frac{1}{\beta (N + 1)} \left(\frac{\alpha}{2} - c\right)$$

$$\frac{p^D}{p^C} = \frac{2N}{N+1} > 1 \quad (N \ge 2 \text{ by assumption}) \implies p^D > p^C$$

The equilibrium bribe is $\frac{1}{\beta(N+1)}\left(\frac{\alpha}{2}-c\right)$ and it is larger than $p^C.$

2.5

Denote the total profits from decentralized corruption by π^D .

$$\begin{split} \pi^D &= N \times p^D \times (Q(P) - c) \\ &= N \times p^D \times \left(\frac{\alpha}{2} - \beta \times P - c\right) \\ &= N \times p^D \times \left(\frac{\alpha}{2} - \beta N \times p^D - c\right) \\ &= \frac{N}{\beta(N+1)^2} \left(\frac{\alpha}{2} - c\right)^2 \end{split}$$

$$\frac{\pi^D}{\pi^C} = \frac{4N}{(N+1)^2} < 1 \quad (N \geq 2 \text{ by assumption}) \implies \pi^D < \pi^C$$

The total profits from decentralized corruption are $\frac{N}{\beta(N+1)^2}\left(\frac{\alpha}{2}-c\right)^2$ and they are smaller than π^C .

2.6

I would ensure corruption to be centralized. From **2.4**, we see that $p^D > p^C$ (i.e. the bribe level at each checkpoint is higher in decentralized corruption than in centralized corruption). In decentralized bribing setting, independent price-setters lead to higher corruption per trip. The decentralized price is so high that it leads to lower total profits from corruption since each price-setter does not take into account the externality of his price decision on demand at other checkpoints.

Centralized is better. Higher corruption but also higher GDP (i.e. higher number of trips) and lower corruption per GDP (i.e. lower p). This implies that the marginal lorry driver is more likely to stay in the market.