# **Bayesian Statistics**

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Spring Semester 2022

# Basics in Bayesian statistics

• Likelihood:  $f(x \mid \theta)$ 

• Prior:  $\pi(\theta)$ 

• Posterior:  $\pi(\theta \mid x) = \frac{\pi(\theta) f(x \mid \theta)}{f(x)} \propto \pi(\theta) f(x \mid \theta)$ 

• Prior predictive density:  $f(x) = \int f(x \mid \theta) \pi(\theta) d\theta$ 

• Posterior predictive density:  $f(y \mid x) = \int f(y \mid x, \theta) \pi(\theta \mid x) d\theta$ 

# Bayesian point estimates

• Posterior mean:

$$\mathbb{E}((\theta - T)^2 \mid x) = \mathbb{E}(\theta^2 \mid x) - 2\mathbb{E}(\theta \mid x)T + T^2$$

This is minimized for  $T = T(X) = \mathbb{E}(\theta \mid x)$ 

• Posterior median:

$$\mathbb{E}(|\theta - T| \mid x) = \int_{-\infty}^{T} (T - \theta)\pi(\theta \mid x)d\theta + \int_{T}^{\infty} (\theta - T)\pi(\theta \mid x)d\theta$$

Using the Leibniz integral rule it follows that

$$\frac{\partial}{\partial T} \mathbb{E}(|\theta - T| \mid x) = \int_{-\infty}^{T} \pi(\theta \mid x) d\theta - \int_{T}^{\infty} \pi(\theta \mid x) d\theta$$

This equals zero if  $T = T(X) = \text{median}\pi(\theta \mid x)$ 

Posterior mode:

$$\mathbb{E}(1_{[-\varepsilon,\varepsilon]^c}(T-\theta)\mid x) = 1 - \int_{T-\varepsilon}^{T+\varepsilon} \pi(\theta\mid x) d\theta$$

For small  $\varepsilon$ , we have

$$\int_{T-\varepsilon}^{T+\varepsilon} \pi(\theta \mid x) d\theta \approx 2\varepsilon \pi(\theta \mid x)$$

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This is maximized, i.e.,  $\mathbb{E}(1_{[-\varepsilon,\varepsilon]^c}(T-\theta)\mid x)$  is minimized, for  $T=T(X)=\mathsf{mode}\pi(\theta\mid x)$ 

# Bayesian decision theory

• Posterior risk:

$$\rho(T(x), \pi) = \mathbb{E}(L(T(X), \theta) \mid x) = \int_{\Theta} L(T(x), \theta) \pi(\theta \mid x) d\theta$$

• Frequentist risk:

$$R(T, \theta) = \mathbb{E}_{\theta}(L(T(X), \theta), \theta) = \int_{\mathbf{X}} L(T(X), \theta) f(X \mid \theta) dX$$

• Bayes factor:

$$B_{01}(x) = \frac{f(x \mid \theta_0)}{f(x \mid \theta_1)} = \frac{\pi(\theta_0 \mid x)\pi(\theta_1)}{\pi(\theta_1 \mid x)\pi(\theta_0)} = \frac{\frac{\pi(\theta_0 \mid x)}{\pi(\theta_1 \mid x)}}{\frac{\pi(\theta_0}{\pi(\theta_1)}} = \frac{\mathsf{Posterior} \ \mathsf{odds}}{\mathsf{Prior} \ \mathsf{odds}}$$

### **Bayesian asymptotics**

• Frequentist asymptotocs:

$$\begin{split} \widehat{\theta}_n &\overset{\mathsf{approx}}{\sim} \mathcal{N}\left(\theta_0, \frac{1}{n} I(\theta_n)^{-1}\right) \\ 2\left(\log L_n(\widehat{\theta}_n) - \log L_n(\theta_n)\right) \overset{\mathsf{d}}{\longrightarrow} \chi_p^2 \end{split}$$

• Bayesian asymptotics:

$$\theta \mid (x_1, \cdots, x_n) \stackrel{\mathsf{approx}}{\sim} \mathcal{N}\left(\widehat{\theta}_n, \frac{1}{n}I(\widehat{\theta}_n)^{-1}\right)$$

# Likelihood principle

ullet Repeat the trial a fixed number n times and observe the random number X of trials where the event occurred:

$$P(X = x) = \binom{n}{x} p^x (1-p)^{n-x}$$

- Binomial distribution
- Reject  $H_0$  if  $x > c_1$

ullet Repeat the experiment a random number of N times until the event has occurred a fixed number of times x:

$$P(N = n) = \binom{n-1}{x-1} p^x (1-p)^{n-x}$$

- Negative binomial distribution
- Reject  $H_0$  if  $n < c_2$

#### **Priors**

• Conjugate priors:

$$\pi_{\xi}(\theta) f(x \mid \theta) \propto \pi_{\xi'}(\theta)$$

• Improper priors:

$$\int_{\theta} \pi(\theta) = \infty$$

• Jeffreys prior:

$$\pi(\theta) \propto \det(I(\theta))^{1/2}$$

• Reference prior:

$$I(X,\theta) = \int_X f(x) \int_{\Theta} \pi(\theta \mid x) \log \frac{\pi(\theta \mid x)}{\pi(\theta)} d\theta dx$$

$$= \int_X f(x) \int_{\Theta} \frac{\pi(x,\theta)}{f(x)} \log \frac{\pi(x,\theta)}{\pi(\theta)f(x)} d\theta dx$$

$$= \int_{X \times \Theta} \pi(x,\theta) \log \frac{\pi(x,\theta)}{\pi(\theta)f(x)} dx d\theta$$

$$= KL(\pi(x,\theta),\pi(\theta)f(x))$$

### **Hierarchical Bayes models**

• Hierarchical Bayes models:

$$\pi(\xi)\pi(\theta \mid \xi)f(x \mid \theta)$$

• Marginal posterior: approach 1

- Compute the marginal prior:

$$\pi(\theta) = \int \pi(\theta \mid \xi) \pi(\xi) d\xi$$

- Then use Bayes formula:

$$\pi(\theta \mid x) \propto \pi(\theta) f(x \mid \theta)$$

- Marginal posterior: approach 2
  - Law of total probability (computationally easier):

$$\pi(\theta \mid x) = \int \pi(\theta \mid x, \xi) \pi(\xi \mid x) d\xi \propto \int \pi(\theta \mid x, \xi) \pi(\xi) f(x \mid \xi) d\xi$$

\* 
$$\pi(\xi \mid x) = \int \pi(\theta, \xi \mid x) d\theta = \int \pi(\theta \mid x) \pi(\theta \mid \xi) d\theta$$

\* 
$$\pi(\xi \mid x) = \frac{\pi(\theta, \xi \mid x)}{(\theta \mid \xi)}$$

\* 
$$\pi(\xi \mid x) = \int \pi(\theta, \xi \mid x) d\theta = \int \pi(\theta \mid x) \pi(\theta \mid \xi) d\theta$$
  
\*  $\pi(\xi \mid x) = \frac{\pi(\theta, \xi \mid x)}{\pi(\theta \mid x, \xi)}$   
\*  $\int f(x \mid \xi) = \int \pi(x, \theta \mid \xi) d\theta = \int f(x \mid \theta) \pi(\theta \mid \xi) d\theta$ 

# **Empirical Bayes method**

• Marginal posterior can be computed as

$$\pi(\theta \mid x) \propto \int \pi(\theta \mid x, \xi) f(x \mid \xi) \pi(\xi) d\xi$$

• Instead of approximating this integral, the empirical Bayes method uses

$$\pi(\theta \mid x) \approx \pi(\theta \mid x, \hat{\xi}(x)) = \frac{f(x \mid \theta)\pi(\theta \mid \hat{\xi}(x))}{f(x \mid \hat{\xi}(x))}$$

where

$$\hat{\xi}(x) = \arg\max_{\xi} f(x \mid \xi) = \arg\max_{\xi} \int f(x \mid \theta) \pi(\theta \mid \xi) d\theta$$

# **Bayesian linear regression**

• Model:

$$y = \alpha \mathbf{1} + X_{\gamma} \beta_{\gamma} + \varepsilon$$

• *g*-prior of Zellner:

$$\beta_{\gamma} \mid \sigma^2 \sim \mathcal{N}\left(\beta_{\gamma}^0, g\sigma^2(X_{\gamma}^T X_{\gamma}^{-1})\right)$$

–  $\beta_{\gamma}^{0}$  is the prior mean. Often  $\beta_{\gamma}^{0}=0$ 

# Laplace approximation

• Laplace approximations are used to approximate integrals of the form

$$\int h(\theta)q(\theta)d\theta$$

where

- q is a possibly unnormalized smooth density which is concentrated around its mode  $\theta_0 = rg \max \log q( heta)$
- $-\ h$  is an arbitrary smooth function
- Expanding  $\log q(\theta)$  into a second-order Taylor series at  $\theta_0$ :

$$\log q(\theta) \approx \log q(\theta_0) - \frac{1}{2} (\theta - \theta_0)^T J(\theta_0) (\theta - \theta_0)$$

$$q(\theta) \approx q(\theta_0) \exp\left(-\frac{1}{2}(\theta - \theta_0)^T J(\theta_0)(\theta - \theta_0)\right)$$

where  $J(\theta)$  is the negative Hessian:

$$J(\theta)_{ij} = -\frac{\partial^2}{\partial \theta_i \partial \theta_j} \log q(\theta)$$

• Expanding  $h(\theta)$  into a first-order Taylor series at  $\theta_0$ :

$$h(\theta) \approx h(\theta_0) + \frac{\partial h}{\partial \theta} (\theta_0)^T (\theta - \theta_0)$$

• Laplace approximation is given by

$$\int h(\theta)q(\theta)d\theta \approx \int \left(h(\theta_0) + \frac{\partial h}{\partial \theta}(\theta_0)^T(\theta - \theta_0)\right) q(\theta_0) \exp\left(-\frac{1}{2}(\theta - \theta_0)^T J(\theta_0)(\theta - \theta_0)\right) d\theta$$

$$= h(\theta_0)q(\theta_0) \int \exp\left(-\frac{1}{2}(\theta - \theta_0)^T J(\theta_0)(\theta - \theta_0)\right) d\theta$$

$$= \int h(\theta)q(\theta)d\theta \approx h(\theta_0)q(\theta_0)(\det J(\theta_0))^{-1/2}(2\pi)^{p/2}$$

- $-\theta_0 = \arg\max\log q(\theta)$
- $J(\theta)$  is the negative Hessian

$$J(\theta)_{ij} = -\frac{\partial^2}{\partial \theta_i \partial \theta_j} \log q(\theta)$$

# Importance sampling

• Monte Carlo sampling:

$$\mathbb{E}(f(x)) = \int f(x)p(x)dx \approx \frac{1}{n} \sum_{i} f(x_i)$$

• Importance sampling:

$$\mathbb{E}(f(x)) = \int f(x)p(x)dx = \int f(x)\frac{p(x)}{q(x)}q(x)dx \approx \frac{1}{n}\sum_{i} f(x_i)\frac{p(x_i)}{q(x_i)}$$

# **Hamiltonian Monte Carlo**

• Hamiltonian H(x, u):

$$H(x, u) = -\log \pi(x) + \sum_{i=1}^{p} \frac{u_i^2}{2m_i}$$
$$\tilde{\pi}(x, u) \propto \exp(-H(x, u))$$

- Physical interpretation:
  - $-\ x$  is the position
  - $\boldsymbol{u}$  is the momentum
  - $-\log \pi(x)$  is the potential energy
  - $\sum_{i=1}^p rac{u_i^2}{2m_i}$  is the kinetic energy H(x,u) is the total energy in the system