

Bayesian Statistics

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Basics in Bayesian statistics

- Likelihood: $f(x | \theta)$
- Prior: $\pi(\theta)$
- Posterior: $\pi(\theta | x) = \frac{\pi(\theta)f(x | \theta)}{f(x)} \propto \pi(\theta)f(x | \theta)$
- Prior predictive density: $f(x) = \int f(x | \theta)\pi(\theta)d\theta$
- Posterior predictive density: $f(y | x) = \int f(y | x, \theta)\pi(\theta | x)d\theta$

Bayesian point estimates

- Posterior mean:

$$\mathbb{E}((\theta - T)^2 | x) = \mathbb{E}(\theta^2 | x) - 2\mathbb{E}(\theta | x)T + T^2$$

This is minimized for $T = T(X) = \mathbb{E}(\theta | x)$

- Posterior median:

$$\mathbb{E}(|\theta - T| | x) = \int_{-\infty}^T (T - \theta)\pi(\theta | x)d\theta + \int_T^{\infty} (\theta - T)\pi(\theta | x)d\theta$$

Using the Leibniz integral rule it follows that

$$\frac{\partial}{\partial T} \mathbb{E}(|\theta - T| | x) = \int_{-\infty}^T \pi(\theta | x)d\theta - \int_T^{\infty} \pi(\theta | x)d\theta$$

This equals zero if $T = T(X) = \text{median}\pi(\theta | x)$

- Posterior mode:

$$\mathbb{E}(1_{[-\varepsilon, \varepsilon]^c}(T - \theta) | x) = 1 - \int_{T-\varepsilon}^{T+\varepsilon} \pi(\theta | x)d\theta$$

For small ε , we have

$$\int_{T-\varepsilon}^{T+\varepsilon} \pi(\theta | x)d\theta \approx 2\varepsilon\pi(\theta | x)$$

This is maximized, i.e., $\mathbb{E}(1_{[-\varepsilon, \varepsilon]^c}(T - \theta) | x)$ is minimized, for $T = T(X) = \text{mode}\pi(\theta | x)$

Bayesian decision theory

- Posterior risk:

$$\rho(T(x), \pi) = \mathbb{E}(L(T(X), \theta) | x) = \int_{\Theta} L(T(x), \theta) \pi(\theta | x) d\theta$$

- Frequentist risk:

$$R(T, \theta) = \mathbb{E}_{\theta}(L(T(X), \theta), \theta) = \int_{\mathbf{X}} L(T(x), \theta) f(x | \theta) dx$$

- Bayes factor:

$$B_{01}(x) = \frac{f(x | \theta_0)}{f(x | \theta_1)} = \frac{\pi(\theta_0 | x) \pi(\theta_1)}{\pi(\theta_1 | x) \pi(\theta_0)} = \frac{\frac{\pi(\theta_0 | x)}{\pi(\theta_1 | x)}}{\frac{\pi(\theta_0)}{\pi(\theta_1)}} = \frac{\text{Posterior odds}}{\text{Prior odds}}$$

Bayesian asymptotics

- Frequentist asymptotics:

$$\hat{\theta}_n \overset{\text{approx}}{\sim} \mathcal{N}\left(\theta_0, \frac{1}{n} I(\theta_0)^{-1}\right)$$

$$2 \left(\log L_n(\hat{\theta}_n) - \log L_n(\theta_0) \right) \xrightarrow{d} \chi_p^2$$

- Bayesian asymptotics:

$$\theta | (x_1, \dots, x_n) \overset{\text{approx}}{\sim} \mathcal{N}\left(\hat{\theta}_n, \frac{1}{n} I(\hat{\theta}_n)^{-1}\right)$$

Likelihood principle

- Repeat the trial a fixed number n times and observe the random number X of trials where the event occurred:

$$P(X = x) = \binom{n}{x} p^x (1-p)^{n-x}$$

- Binomial distribution
- Reject H_0 if $x > c_1$

- Repeat the experiment a random number of N times until the event has occurred a fixed number of times x :

$$P(N = n) = \binom{n-1}{x-1} p^x (1-p)^{n-x}$$

- Negative binomial distribution
- Reject H_0 if $n < c_2$

Priors

- Conjugate priors:

$$\pi_{\xi}(\theta) f(x | \theta) \propto \pi_{\xi'}(\theta)$$

- Improper priors:

$$\int_{\theta} \pi(\theta) = \infty$$

- Jeffreys prior:

$$\pi(\theta) \propto \det(I(\theta))^{1/2}$$

- Reference prior:

$$\begin{aligned}
I(X, \theta) &= \int_X f(x) \int_{\Theta} \pi(\theta | x) \log \frac{\pi(\theta | x)}{\pi(\theta)} d\theta dx \\
&= \int_X f(x) \int_{\Theta} \frac{\pi(x, \theta)}{f(x)} \log \frac{\pi(x, \theta)}{\pi(\theta)f(x)} d\theta dx \\
&= \int_{X \times \Theta} \pi(x, \theta) \log \frac{\pi(x, \theta)}{\pi(\theta)f(x)} dx d\theta \\
&= KL(\pi(x, \theta), \pi(\theta)f(x))
\end{aligned}$$

Hierarchical Bayes models

- Hierarchical Bayes models:

$$\pi(\xi)\pi(\theta | \xi)f(x | \theta)$$

- Marginal posterior: approach 1

- Compute the marginal prior:

$$\pi(\theta) = \int \pi(\theta | \xi)\pi(\xi)d\xi$$

- Then use Bayes formula:

$$\pi(\theta | x) \propto \pi(\theta)f(x | \theta)$$

- Marginal posterior: approach 2

- Law of total probability (computationally easier):

$$\pi(\theta | x) = \int \pi(\theta | x, \xi)\pi(\xi | x)d\xi \propto \int \pi(\theta | x, \xi)\pi(\xi)f(x | \xi)d\xi$$

$$* \pi(\xi | x) = \int \pi(\theta, \xi | x)d\theta = \int \pi(\theta | x)\pi(\theta | \xi)d\theta$$

$$* \pi(\xi | x) = \frac{\pi(\theta, \xi | x)}{\pi(\theta | x, \xi)}$$

$$* \int f(x | \xi) = \int \pi(x, \theta | \xi)d\theta = \int f(x | \theta)\pi(\theta | \xi)d\theta$$

Empirical Bayes method

- Marginal posterior can be computed as

$$\pi(\theta | x) \propto \int \pi(\theta | x, \xi)f(x | \xi)\pi(\xi)d\xi$$

- Instead of approximating this integral, the empirical Bayes method uses

$$\pi(\theta | x) \approx \pi(\theta | x, \hat{\xi}(x)) = \frac{f(x | \theta)\pi(\theta | \hat{\xi}(x))}{f(x | \hat{\xi}(x))}$$

where

$$\hat{\xi}(x) = \arg \max_{\xi} f(x | \xi) = \arg \max_{\xi} \int f(x | \theta)\pi(\theta | \xi)d\theta$$

Bayesian linear regression

- Model:

$$y = \alpha \mathbf{1} + X_{\gamma} \beta_{\gamma} + \varepsilon$$

- g -prior of Zellner:

$$\beta_{\gamma} | \sigma^2 \sim \mathcal{N}(\beta_{\gamma}^0, g\sigma^2(X_{\gamma}^T X_{\gamma}^{-1}))$$

- β_{γ}^0 is the prior mean. Often $\beta_{\gamma}^0 = 0$

Laplace approximation

- Laplace approximations are used to approximate integrals of the form

$$\int h(\theta)q(\theta)d\theta$$

where

- q is a possibly unnormalized smooth density which is concentrated around its mode $\theta_0 = \arg \max \log q(\theta)$
- h is an arbitrary smooth function

- Expanding $\log q(\theta)$ into a second-order Taylor series at θ_0 :

$$\log q(\theta) \approx \log q(\theta_0) - \frac{1}{2}(\theta - \theta_0)^T J(\theta_0)(\theta - \theta_0)$$

$$q(\theta) \approx q(\theta_0) \exp \left(-\frac{1}{2}(\theta - \theta_0)^T J(\theta_0)(\theta - \theta_0) \right)$$

where $J(\theta)$ is the negative Hessian:

$$J(\theta)_{ij} = -\frac{\partial^2}{\partial \theta_i \partial \theta_j} \log q(\theta)$$

- Expanding $h(\theta)$ into a first-order Taylor series at θ_0 :

$$h(\theta) \approx h(\theta_0) + \frac{\partial h}{\partial \theta}(\theta_0)^T (\theta - \theta_0)$$

- Laplace approximation is given by

$$\begin{aligned} \int h(\theta)q(\theta)d\theta &\approx \int \left(h(\theta_0) + \frac{\partial h}{\partial \theta}(\theta_0)^T (\theta - \theta_0) \right) q(\theta_0) \exp \left(-\frac{1}{2}(\theta - \theta_0)^T J(\theta_0)(\theta - \theta_0) \right) d\theta \\ &= h(\theta_0)q(\theta_0) \int \exp \left(-\frac{1}{2}(\theta - \theta_0)^T J(\theta_0)(\theta - \theta_0) \right) d\theta \\ &= \int h(\theta)q(\theta)d\theta \approx h(\theta_0)q(\theta_0)(\det J(\theta_0))^{-1/2}(2\pi)^{p/2} \end{aligned}$$

- $\theta_0 = \arg \max \log q(\theta)$
- $J(\theta)$ is the negative Hessian

$$J(\theta)_{ij} = -\frac{\partial^2}{\partial \theta_i \partial \theta_j} \log q(\theta)$$

Importance sampling

- Monte Carlo sampling:

$$\mathbb{E}(f(x)) = \int f(x)p(x)dx \approx \frac{1}{n} \sum_i f(x_i)$$

- Importance sampling:

$$\mathbb{E}(f(x)) = \int f(x)p(x)dx = \int f(x)\frac{p(x)}{q(x)}q(x)dx \approx \frac{1}{n} \sum_i f(x_i)\frac{p(x_i)}{q(x_i)}$$