

Limited Dependent Variable Models and Sample Selection Corrections.

• Limited dependent variables (LDV)

- LDV are substantively restricted
 - Binary variables take only two values, e.g. employed/not employed
 - Nonnegative variables, e.g. wages, prices, interest rates
 - Nonnegative variables with excess zeros, e.g. labor supply
 - Count variables, e.g. the number of arrests in a year

• Sample selection models

- The sample used to infer population relationships is endogenously selected, e.g. wage offer regression but data only about working women.

• Logit and Probit models for binary response

- Disadvantages of the LPM for binary dependent variables
 - Predictions sometimes lie outside the unit interval
 - Partial effects of explanatory variables are constant

• Nonlinear models for binary response

- Response probability is a nonlinear function of explanat. variables

$$P(y = 1|x) = G(\beta_0 + \beta_1 x_1 + \dots + \beta_k x_k) = G(x\beta)$$

↑
Probability of a "success" given explanatory variables


↑
A cumulative distribution function $0 < G(z) < 1$. The response probability is thus a function of the explanatory variables x .

↑
Shorthand vector notation: the vector of explanatory variables x also contains the constant of the model.

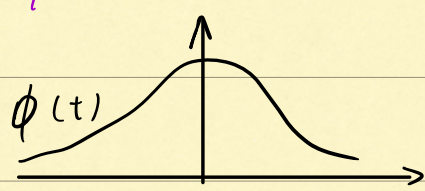
$$P(y=1|x) = G(\beta_0 + \beta_1 x_1 + \dots + \beta_k x_k) = G(x\beta).$$

$0 \leq G(z) \leq 1$

Candidates for Λ :

①. logit model: $\Lambda(z) = \frac{e^z}{1 + e^z} = \Lambda(z)$

C.d.f of standard logistic distribution

②. probit model: $\Lambda(z) = \Phi(z) = \int_{-\infty}^z \phi(z) dz$
C.d.f. of standard normal distribution.

 $\phi(z) = \frac{1}{\sqrt{2\pi}} e^{-\frac{z^2}{2}}$

代碼: `glm(y ~ x1 + x2 + ... + xk, family = binomial(link = probit))`
/ logit

Latent variable formulation of the logit and probit models.

$y^* = x\beta + e$ and $y = 1 \cdot [y^* > 0]$
{ standard logistic distribution $\Lambda(0,1)$
standard normal distribution $\mathcal{N}(0,1)$ }
indep of x

$\Rightarrow P(y=1|x) = P(y^* > 0|x)$

$= P(x\beta + e > 0|x)$

$= P(e > -x\beta | x)$

$= 1 - \underset{\text{or } \Phi}{\Lambda}(-x\beta | x)$

$= \underset{\text{or } \Phi}{\Lambda}(x\beta | x)$

• Interpretation of coefficients in Logit and Probit models

Continuous explanatory variables:

$$\frac{\partial P(y = 1|x)}{\partial x_j} = g(x\beta)\beta_j \quad \text{where} \quad g(z) \equiv \partial G(z)/\partial z > 0$$

How does the probability for $y = 1$ change if explanatory variable x_j changes by one unit?

Discrete explanatory variables:

$$G[\beta_0 + \beta_1 x_1 + \dots + \beta_k(c_k + 1)] - G[\beta_0 + \beta_1 x_1 + \dots + \beta_k c_k]$$

For example, explanatory variable x_k increases by one unit.

• Partial effects are nonlinear and depend on the level of x .

• Maximum likelihood estimation of Logit and Probit models

$$f(y_i|x_i; \beta) = [G(x_i\beta)]^{y_i} [1 - G(x_i\beta)]^{1-y_i}$$

The probability that individual i 's outcome is y_i , given that his/her characteristics are x_i

$$\log L(\beta) = \log \left(\prod_{i=1}^n f(y_i|x_i; \beta) \right) = \sum_{i=1}^n \log f(y_i|x_i; \beta)$$

Under random sampling

$$\max \log L(\beta) \rightarrow \hat{\beta}_0, \hat{\beta}_1, \dots, \hat{\beta}_k$$

Maximum likelihood estimates

• Properties of maximum likelihood estimators

- Maximum likelihood estimators are consistent, asymptotically normal, and asymptotically efficient if the distributional assumptions hold.

• Hypothesis testing after maximum likelihood estimation

- The usual t-tests and confidence intervals can be used.
- There are three alternatives to test multiple hypotheses:
 - Lagrange multiplier or score test (not discussed here)
 - Wald test (requires only estimation of unrestricted model)
 - Likelihood ratio test (restricted and unrestricted models needed)

$$LR = 2(\log L_{ur} - \log L_r) \sim \chi_q^2$$

Chi-square distribution with q degrees of freedom

The null hypothesis that the q hypotheses hold is rejected if the growth in maximized likelihood is too large when going from the restricted to the unrestricted model

• Goodness-of-fit measures for Logit and Probit models

- Percent correctly predicted = $\frac{\# \text{correctly predicted } \tilde{y}}{\# \tilde{y}}$

$$\tilde{y}_i = \begin{cases} 1 & \text{if } G(\mathbf{x}_i \hat{\beta}) \geq .5 \\ 0 & \text{otherwise} \end{cases}$$

\tilde{y}_i is predicted.

Individual i's outcome is predicted as one if the probability for this event is larger than .5, then percentage of correctly predicted $y = 1$ and $y = 0$ is counted

• Pseudo R-squared

$$\tilde{R}^2 = 1 - \log L_{ur} / \log L_0$$

Compare maximized log-likelihood of the model with that of a model that only contains a constant (and no explanatory variables)

• Correlation based measures

$$\text{Corr}(y_i, \tilde{y}_i), \text{Corr}(y_i, G(\mathbf{x}_i \hat{\beta}))$$

Look at correlation (or squared correlation) between predictions or predicted prob. and true values

• Reporting partial effects of explanatory variables

- The difficulty is that partial effects are not constant but depend on.
- Partial effects at the average:

$$\widehat{PEA}_j = g(\bar{\mathbf{x}} \hat{\beta}) \hat{\beta}_j$$

The partial effect of explanatory variable x_j is considered for an "average" individual (this is problematic in the case of explanatory variables such as gender)

- Average partial effects:

$$\widehat{APE}_j = n^{-1} \sum_{i=1}^n g(\mathbf{x}_i \hat{\beta}) \hat{\beta}_j$$

The partial effect of explanatory variable x_j is computed for each individual in the sample and then averaged across all sample members (makes more sense)

- Analogous formulas hold for discrete explanatory variables.

$glm(y \sim x_1 + x_2 + \dots + x_k, \text{family} = \text{binomial}(\text{link} = \text{probit}))$
logit

summary(A)

LogLik(A)

Likelihood ratio: $LR = 2(\text{LogLik}(u) - \text{LogLik}(r))$

$$\sim \chi^2_q$$

$$P(y=1 | x) = \int_{-\infty}^z \lambda(z) dz = \int_{-\infty}^z \frac{e^v}{(1+e^v)^2} dv.$$

$$= \frac{e^z}{(1+e^z)^2}$$

$$= \frac{e^{\beta_0 + \beta_1 x_1 + \dots + \beta_k x_k}}{1 + e^{\beta_0 + \beta_1 x_1 + \dots + \beta_k x_k}}$$

$$\widehat{P}(y=1 | x) = \frac{e^{\widehat{\beta}_0 + \widehat{\beta}_1 x_1 + \dots + \widehat{\beta}_k x_k}}{1 + e^{\widehat{\beta}_0 + \widehat{\beta}_1 x_1 + \dots + \widehat{\beta}_k x_k}}$$