Practice Exam 2 Econ471 Fall 2020 Solution

Part I. Multiple Choice Questions

- 1. The normality assumption implies that:
- a. the population error u is dependent on the explanatory variables and is normally distributed with mean equal to one and variance σ^2 .
- b. the population error u is independent of the explanatory variables and is normally distributed with mean equal to one and variance σ .
- c. the population error u is dependent on the explanatory variables and is normally distributed with mean zero and variance σ .
- d. the population error u is independent of the explanatory variables and is normally distributed with mean zero and variance σ^2 .

Answer: d

- 2. Consider the equation, $Y = \beta_1 + \beta_2 X_2 + u$. A null hypothesis, H_0 : $\beta_2 = 0$ states that:
- a. X_2 has no effect on the expected value of β_2 .
- b. X₂ has no effect on the expected value of Y.
- c. β_2 has no effect on the expected value of Y.
- d. Y has no effect on the expected value of X₂.

Answer: b

- 3. Which of the following statements is true?
- a. If the calculated value of F statistic is higher than the critical value, we reject the alternative hypothesis in favor of the null hypothesis.
- b. The F statistic is always nonnegative as SSR_r is never smaller than SSR_{ur} .
- c. Degrees of freedom of a restricted model is always less than the degrees of freedom of an unrestricted model.
- d. The F statistic is more flexible than the t statistic to test a hypothesis with a single restriction.

Answer: b

- 4. Which of the following statements is true?
- a. The standard error of a regression, $\hat{\sigma}$, is not an unbiased estimator for σ , the standard deviation of the error, u, in a multiple regression model.
- b. All estimators that are unbiased are also consistent.
- c. Almost all economists agree that unbiasedness is a minimal requirement for an estimator in regression analysis.
- d. All estimators in a regression model that are consistent are also unbiased.

Answer: a

- 5. In a multiple regression model, the OLS estimator is consistent if:
- a. there is no correlation between the dependent variables and the error term.
- b. there is a perfect correlation between the dependent variables and the error term.
- c. the sample size is less than the number of parameters in the model.
- d. there is no correlation between the independent variables and the error term.

Answer: d

- 6. If $\delta_1 = \text{Cov}(x_1, x_2) / \text{Var}(x_1)$ where x_1 and x_2 are two independent variables in a regression equation, which of the following statements is true?
- a. If x_2 has a positive partial effect on the dependent variable, and $\delta_1 > 0$, then the inconsistency in the simple regression slope estimator associated with x_1 is negative.
- b. If x_2 has a positive partial effect on the dependent variable, and $\delta_1 > 0$, then the inconsistency in the simple regression slope estimator associated with x_1 is positive.
- c. If x_1 has a positive partial effect on the dependent variable, and $\delta_1 > 0$, then the inconsistency in the simple regression slope estimator associated with x_1 is negative.
- d. If x_1 has a positive partial effect on the dependent variable, and $\delta_1 > 0$, then the inconsistency in the simple regression slope estimator associated with x_1 is positive.

Answer: b

- 7. Which of the following statements is true under the Gauss-Markov assumptions?
- a. Among a certain class of estimators, OLS estimators are best linear unbiased, but are asymptotically inefficient.
- b. Among a certain class of estimators, OLS estimators are biased but asymptotically efficient.
- c. Among a certain class of estimators, OLS estimators are best linear unbiased and asymptotically efficient.
- d. The LM test is independent of the Gauss-Markov assumptions.

Answer: c

- 8. Changing the unit of measurement of any independent variable, where log of the independent variable appears in the regression:
- a. affects only the intercept coefficient.
- b. affects only the slope coefficient.
- c. affects both the slope and intercept coefficients.
- d. affects neither the slope nor the intercept coefficient.

Answer: a

- 9. If a regression equation has only one explanatory variable, say x_1 , its standardized coefficient must lie in the range:
- a. -2 to 0.
- b. -1 to 1.
- c. 0 to 1.
- d. 0 to 2.

Answer: b

- 10. Which of the following correctly represents the equation for adjusted R²?
- a. $\bar{R}^2 = 1 [SSR/(n-1)]/[SST/(n+1)]$
- b. $\bar{R}^2 = 1 [SSR/(n-k-1)]/[SST/(n+1)]$
- c. $\bar{R}^2 = 1 [SSR/(n-k-1)]/[SST/(n-1)]$
- d. $\bar{R}^2 = 1 [SSR]/[SST/(n-1)]$

Answer: c

11. The following simple model is used to determine the annual savings of an individual on the basis of his annual income and education.

Savings = $\beta_0 + \delta_0$ Edu + β_1 Inc+u

The variable 'Edu' takes a value of 1 if the person is educated and the variable 'Inc' measures the income of the individual. The inclusion of another binary variable in this model that takes a value of 1 if a person is uneducated, will give rise to the problem of _____.

- a. omitted variable bias
- b. self-selection
- c. dummy variable trap
- d. heteroskedastcity

Answer: c

Part 2. Quantitative Questions

1. The F-statistic for testing a set of linear hypotheses is given by the formula

$$F = \frac{(SSR_r - SSR_{ur})/q}{SSR_{ur}/(n-k-1)}$$

where SSR_r is the sum of squared residuals from the restricted regression, SSR_{ur} is the sum of squared residuals from the unrestricted regression, q is the number of restrictions under the null hypothesis, and k is the number of regressors in the unrestricted regression. Prove that this formula is the same as the following formula based on the regression R^2 of the restricted and unrestricted regression:

$$F = \frac{(R_{ur}^2 - R_r^2)/q}{(1 - R_{ur}^2)/(n - k - 1)}$$

Solution: Since $R^2=1-\frac{SSR}{SST}$, SSR=SST(1- R^2). Substitution into the first equation then results in the second equation, once the "1" in the numerator is canceled, and the *SST* is factored out in the numerator.

2. Consider the following multiple regression model

$$Y_i = \beta_0 + \beta_1 X_{1i} + \beta_2 X_{2i} + \beta_3 X_{3i} + u_i$$

You want to consider certain hypotheses involving more than one parameter. You decide to test the joint hypotheses using the F-statistics. For each of the cases below specify a restricted model and indicate how you would compute the F-statistic to test for the validity of the restrictions.

- (a) $\beta_1 = -\beta_2; \ \beta_3 = 0$
- (b) $\beta_1 + \beta_2 + \beta_3 = 1$
- (c) $\beta_1 = \beta_2/\beta_3$

Solution: (a) The restricted model is $Y_i=eta_0+eta_1(X_{1i}-X_{2i})+u_i$ and the F-statistic would be

$$F = \frac{(SSR_r - SSR_{ur})/2}{SSR_{ur}/(n-3-1)}$$

(b) The restricted model is $Y_i-X_{3i}=\beta_0+\beta_1(X_{1i}-X_{3i})+\beta_2(X_{2i}-X_{3i})+u_i$ and the F-statistic would be

$$F = \frac{(SSR_r - SSR_{ur})/1}{SSR_{ur}/(n-3-1)}$$

(c) This is not a linear restriction. Hence you cannot use the F-test to test for its validity.

3.

- (i) With df = n 2 = 86, we obtain the 5% critical value from Table G.2 with df = 90. Because each test is two-tailed, the critical value is 1.987. The t statistic for H₀: $\beta_0 = 0$ is about -.89, which is much less than 1.987 in absolute value. Therefore, we fail to reject $\beta_0 = 0$. The t statistic for H₀: $\beta_1 = 1$ is (.976 1)/.049 \approx -.49, which is even less significant. (Remember, we reject H₀ in favor of H₁ in this case only if |t| > 1.987.)
- (ii) We use the SSR form of the F statistic. We are testing q = 2 restrictions and the df in the unrestricted model is 86. We are given SSR_r = 209,448.99 and SSR_{ur} = 165,644.51. Therefore,

$$F = \frac{(209,448.99 - 165,644.51)}{165,644.51} \cdot \left(\frac{86}{2}\right) \approx 11.37,$$

which is a strong rejection of H₀: from Table G.3c, the 1% critical value with 2 and 90 df is 4.85.

- (iii) We use the *R*-squared form of the *F* statistic. We are testing q = 3 restrictions and there are 88 5 = 83 df in the unrestricted model. The *F* statistic is $[(.829 .820)/(1 .829)](83/3) \approx 1.46$. The 10% critical value (again using 90 denominator df in Table G.3a) is 2.15, so we fail to reject H₀ at even the 10% level. In fact, the *p*-value is about .23.
- (iv) If heteroskedasticity were present, Assumption MLR.5 would be violated, and the *F* statistic would not have an *F* distribution under the null hypothesis. Therefore, comparing the *F* statistic against the usual critical values, or obtaining the *p*-value from the *F* distribution, would not be especially meaningful.
- **4.** Write $y = \beta_0 + \beta_1 x + u$, and take the expected value: $E(y) = \beta_0 + \beta_1 E(x) + E(u)$, or $\mu_y = \beta_0 + \beta_1 \mu_x$ since E(u) = 0, where $\mu_y = E(y)$ and $\mu_x = E(x)$. We can rewrite this as $\beta_0 = \mu_y \beta_1 \mu_x$. Now, $\hat{\beta}_0 = \overline{y} \hat{\beta}_1 \overline{x}$. Taking the plim of this we have P(x) = P(x) = P(x) = P(x). Taking the plim of this we have P(x) = P(x) = P(x) = P(x). We can rewrite this as P(x) = P(x) = P(x) = P(x). Now, P(x) = P(x) = P(x) = P(x). Now, P(x) = P(x) = P(x). Taking the plim of this we have P(x) = P(x) = P(x). We can rewrite this as P(x) = P(x) = P(x). Now, P(x) = P(x) = P(x). Now, P(x) = P(x) = P(x). Now, P(x) = P(x) = P(x). Taking the plim of this we have P(x) = P(x). Now, P(x) = P(x). Taking the plim of this we have P(x) = P(x).
- 5. (i) If $\triangle cigs = 10$ then $\triangle \log \widehat{(bwght)} = -.0044(10) = -.044$, which means about a 4.4% lower birth weight.

- (ii) A white child is estimated to weigh about 5.5% more, other factors in the first equation fixed. Further, $t_{white} \approx 4.23$, which is well above any commonly used critical value. Thus, the difference between white and nonwhite babies is also statistically significant.
- (iii) If the mother has one more year of education, the child's birth weight is estimated to be .3% lower. This is not a huge effect, and the *t* statistic is only one, so it is not statistically significant.
- (iv) The two regressions use different sets of observations. The second regression uses fewer observations because *motheduc* or *fatheduc* are missing for some observations. We would have to reestimate the first equation (and obtain the *R*-squared) using the same observations used to estimate the second equation.