PRACTICE FINAL

STAT 410 FALL 2020

Wednesday, December 9, 2020

Name & UIN	

Instructions:

- 1. You should finish this exam in 3 hours.
- 2. You can use any result shown in the lectures but you have to clearly state the result here.
- 3. Show all your supporting work. If you just write the answer you may not get a lot of credit.
- 4. There are 8 questions each worth 20 points. The maximum you can score is 160. This exam is worth 150 points and the remaining 10 points are for extra credit.

Problem 1: (6 + 6 + 8 points)

Let N(t) be a Poisson process with rate λ . Let $X = N(t_1), Y = N(t_2) - N(t_1)$ and $Z = N(t_3) - N(t_2)$ where $0 < t_1 < t_2 < t_3$ are some fixed times.

- a) What is the conditional distribution of X given X + Y + Z = 10.
- b) What is the conditional joint distribution of X,Y,Z given X+Y+Z=10? Is it a named distribution we have studied?
 - c) What is the conditional distribution of Y + Z given X + Y + Z = 10?

(a)
$$P(X=x|X+Y+Z=/0) = \frac{P(N(t_{i})=x) P(N(t_{3})-N(t_{i})=/o-x)}{P(N(t_{3})=/o)}$$

$$= \frac{(\lambda t_{i})^{x}}{x!} \frac{(\lambda (t_{3}-t_{i}))^{x}}{(/o-x)!} \frac{P(N(t_{3})=/o)}{e^{-\lambda (t_{3}-t_{i})}}$$

$$= \frac{(\lambda t_{3})^{x}}{(o!)} \frac{(\lambda t_{3})^{x}}{(/o-x)!} \frac{(\lambda t_{3}-t_{i})}{e^{-\lambda (t_{3}-t_{i})}}$$

$$= \frac{(\lambda t_{3})^{x}}{x!(/o-x)!} \frac{(\lambda t_{3})^{x}}{(\lambda t_{3})^{x}} \frac{(-\lambda t_{3})^{x}}{(-\lambda t_{3})^{x}}$$

$$= \frac{(\lambda t_{3})^{x}}{x!(/o-x)!} \frac{(\lambda t_{3})^{x}}{(\lambda t_{3})^{x}} \frac{(\lambda t_{3}-t_{i})^{x}}{(\lambda t_{3}-t_{i})^{x}}$$

$$= \frac{P(N_{i}(t_{1})=x) P(N(t_{3}-t_{1})=y) P(N(t_{3}-t_{2})=/o-xy)}{P(N(t_{3}-t_{2})=/o-xy)}$$

$$= \frac{(\lambda t_{3})^{x}}{x!(1o-x-y)!} \frac{(\lambda t_{3})^{x}}{(\lambda t_{3})^{x}} \frac{(\lambda t_{3}-t_{2})^{x}}{(\lambda t_{3})^{x}} \frac{(\lambda t_{3}-t_{2})^{x}}{(\lambda t_{3})^{x}}$$

$$= \frac{(\lambda t_{3})^{x}}{x!(1o-x-y)!} \frac{(\lambda t_{3})^{x}}{(\lambda t_{3})^{x}} \frac{(\lambda t_{3}-t_{3})^{x}}{(\lambda t_{3})^{x}} \frac{(\lambda t_{3}-t_{3})^{x}}{(\lambda t_{3})^{x}}$$

$$= \frac{(\lambda t_{3})^{x}}{x!(1o-x-y)!} \frac{(\lambda t_{3})^{x}}{(\lambda t_{3})^{x}} \frac{(\lambda t_{3}-t_{3})^{x}}{(\lambda t_{3})^{x}} \frac{(\lambda t_{3}-t_{3})^{x}}{(\lambda t_{3})^{x}}$$

$$= \frac{(\lambda t_{3})^{x}}{x!(1o-x)!} \frac{(\lambda t_{3})^{x}}{(\lambda t_{3})^{x}} \frac{(\lambda t_{3}-t_{3})^{x}}{(\lambda t_{3})^{x}} \frac{(\lambda t_{3}-t_{3})^{x}}{(\lambda t_{3})^{x}}$$

$$= \frac{(\lambda t_{3})^{x}}{x!(1o-x)!} \frac{(\lambda t_{3})^{x}}{(\lambda t_{3})^{x}} \frac{(\lambda t_{3}-t_{3})^{x}}{(\lambda t_{3})^{x}} \frac{(\lambda$$

(C)
$$P(Y+Z=k|X+Y+Z=10) = P(X=10-k|X+Y+Z=10)$$

= $\frac{(0!)}{(10-k)!k!} (\frac{t_1}{t_3})^{(0-k)} (1-\frac{t_1}{t_3})^k$

$$\sim Bin^2 (lo, \frac{t_3-t_1}{t_3}).$$

Problem 2:(10 + 10 points)

Let W be double exponential random variable with pdf $f(x) = \exp(-|x|)/2$. A measuring device is used to observe W, but the device can only handle positive values, and gives a reading of 0 if W < 0; this is an example of censored data. So assume that $X = W\mathbb{I}(W > 0)$ is observed rather than W where $\mathbb{I}(W > 0)$ is the indicator of W > 0. Find E(X) and Var(X).

$$F_{X}(x) = \begin{cases} \frac{1}{2}, & x = 0 \\ \frac{1}{2} + \int_{0}^{1} \frac{1}{2} e^{-t} dt = |-\frac{1}{2}e^{-x}, & x > 0 \end{cases}$$

$$f_{X}(x) = \begin{cases} 0, & x < 0 \\ \frac{1}{2} + \int_{0}^{1} \frac{1}{2} e^{-t} dt = |-\frac{1}{2}e^{-x}, & x > 0 \end{cases}$$

$$E(X) = \int_{0}^{\infty} \frac{1}{2} e^{-x} x dx = -\int_{0}^{\infty} \frac{1}{2} x de^{-x}$$

$$= -\frac{1}{2} x e^{-x} \Big|_{0}^{\infty} + \int_{0}^{\infty} \frac{1}{2} e^{-x} dx.$$

$$= -\frac{1}{2} e^{-x} \Big|_{0}^{\infty} = \frac{1}{2}$$

$$Var(X) = E(X^{2}) - E^{2}(X)$$

$$= \int_{0}^{\infty} \frac{1}{2} e^{-X} x^{2} dx - \frac{1}{4} = -\int_{0}^{\infty} \frac{1}{2} x^{2} de^{-X} - \frac{1}{4}$$

$$= -\frac{1}{2} x^{2} e^{-X} \Big|_{0}^{\infty} + \int_{0}^{\infty} x e^{-X} dx - \frac{1}{4}$$

$$= \frac{3}{4}$$

Problem 3:(6 + 6 + 8 points)

A drunken man wanders around randomly in a large space. At each step, he flips a fair coin and chooses to move either horizontally or vertically. Then he moves W_n steps horizontally or vertically (depending on the coin toss) where $W_1, W_2, \ldots, W_n, \ldots$ is a sequence of i.i.d discrete random variables supported on the non negative integers. Let (X_n, Y_n) and R_n be his position and distance from the origin after n steps, respectively.

(a) Determine whether or not X_n is independent of Y_n . Give reasoning.

= $n E(w_1^2) + (n^2 - n^4)(Ew_1)^2 - 2[-\frac{n}{4}(Ew_1)^2 + \frac{n^2}{4}(Ew_1)^2]$

Problem 4: (10 + 10 points)

We have observed random variables Y_1, Y_2, \ldots, Y_n i.i.d N(0, 1) random variables.

- (a) For fixed j and k with $1 \le j < k \le n$ find $P(Y_{new} \in [Y_{(j)}, Y_{(k)}])$.
- (b) Does your answer change if the Y_i are i.i.d Unif(0,1) random variables? Does your answer change if the Y_i are i.i.d Bernoulli(0.5) random variables?

(a)
$$Y_{(1)} Y_{(2)} Y_{(2)} Y_{(3)}$$
, $Y_{(4)} Y_{(4)} Y_{(4)}$

$$P(Y_{new} \in [Y_{(j)}, Y_{(k)}] = \frac{k-j}{n+1}$$
(b) Doesn't charge when $Y_i \sim U_{nif}(0,1)$.

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<u>Problem 5:</u> (5 + 5 + 5 + 5 points)

Let X_1, X_2, \dots, X_n and Y be i.i.d continuous random variables with CDF F. Let $V_i = \mathbb{I}(X_i < Y)$.

- (a) Are V_i independent? Are V_i conditionally independent given Y? Explain.
- (b) Let $N = \sum_{i=1}^{n} V_i$. Find EN.
- (c) Find the conditional distribution of N, given Y = y
- (d) Find Var(N).

(a) No Vi depend on Y., Yes
$$P(V_1=1 | V_2=1) > P(V_1=1 | V_2=0)$$
.

(b)
$$EN = nEV_i = \frac{n}{2}$$

(C)
$$V_{i}|Y=y=I(X_{i}< y)\sim Bern(F(y)).$$

(d)
$$Var(N) = E(Var(N|Y)) + Var(E(N|Y))$$
.

$$= E(nF(y)(I-F(y))) + Var(nF(y)).$$

$$= \frac{n}{2} - \frac{n}{3} + \frac{n^2}{12} = \frac{n^2}{12} + \frac{n}{6}$$

Problem 6: (6 + 6 + 8 points)

Let Z_1, Z_2 be N(0,1) i.i.d random variables.

- (a) As a function of Z_1, Z_2 , create an Gamma(2, 1) r.v. X (your answer can also involve the standard Normal CDF).
- (b) Let $Y = \exp(-R)$, where $R = \sqrt{Z_1^2 + Z_2^2}$. Write down (but do not evaluate) an integral for E(Y).
- (c) Let $X_1 = Z_1 + Z_2$ and $X_2 = Z_1 Z_2$. Determine whether X_1 and X_2 are independent.

(a) C.D. F of exp(1)
$$F(x) = 1 - e^{-x} \times 20$$
.

$$= \sum_{i=1}^{n} F^{i}(y) = -\log(1-y), y \in [0,1]$$

$$F^{i}(\Phi(x)) \sim \exp(1). \quad i = 1, L$$

$$\Phi(Z_{i}) \sim Unif(0,1) = \sum_{i=1}^{n} 1 - \Phi(Z_{L}) \sim Unif(0,1)$$

$$= \sum_{i=1}^{n} -(\log(\Phi(Z_{i})) \sim \exp(1))$$

$$= \sum_{i=1}^{n} -(\log(\Phi(Z_{i})) -(\log(\Phi(Z_{L}))) \sim Camma(2,1)$$

$$= \sum_{i=1}^{n} \exp(\frac{1}{2}(Z_{L})) -(\log(\Phi(Z_{L}))) \sim Camma(2,1)$$

(C)

Problem 7: (6 + 6 + 8 points)

A Markov chain X_0, X_1, \ldots with state space $\{-2, -1, 0, 1, 2\}$ proceeds as follows. The chain starts at $X_0 = 0$. If X_n is not an endpoint 2 or -2, then X_{n+1} is $X_n + 1$ or $X_n - 1$, each with probability 1/2. Otherwise, the chain gets reflected of the endpoint, i.e., from 2 it always goes to 1 and from -2 it always goes to -1.

- (a) Is $|X_0|, |X_1|, |X_2|, \ldots$ also a Markov chain? Explain.
- (b) (b) Let sgn be the sign function: sgn(x) = 1 if x > 0, sgn(x) = -1 if x < 0, and sgn(0) = 0. Is $sgn(X_0), sgn(X_1), sgn(X_2), \ldots$ a Markov chain? Explain.
 - (b) Find the stationary distribution of the chain X_0, X_1, X_2, \ldots

 $=) S = \left(\frac{1}{8}, \frac{2}{8}, \frac{2}{8}, \frac{2}{5}, \frac{1}{5}\right).$

Problem 8: (10 + 10 points)

(a) Let $X \sim Exp(\lambda)$ and $Y|X = x \sim Poi(x)$. What is the distribution of Y?

Let N_t be a Poisson process with rate λ . Define τ to be the time of first arrival in the process N_t , so that $\tau = \min\{t : N_t = 1\}$.

(b) What is the distribution of $N_{2\tau} - N_{\tau}$?

(a). Consider two possion process I: rate / II: rate λ begin at a same time.

Let X be the time of first event belong to II.

Y be the number of events belong to I before
the time of the first event belongs to II.

Y~ Geom (7t1).

(b) N2T-NT T ~ Pois(T).

 $T \sim E \times po(\lambda)$ => $N \times T - N + \sim Cleom(\frac{\lambda}{\lambda + 1})$.