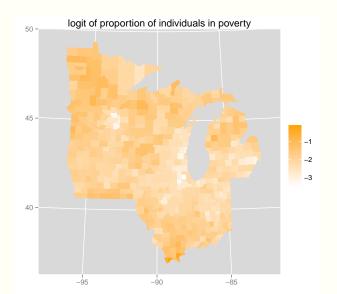
Local variable selection and parameter estimation for spatially varying coefficient models

Wesley Brooks

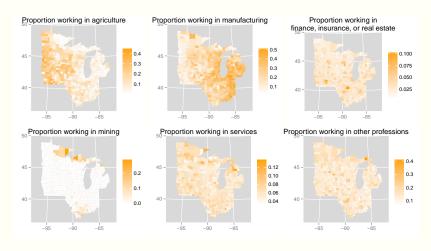
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Take a look at some data



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Sensible questions about the data

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- What are the sign and magnitude of that association?
- Is poverty rate associated with the same economic-structure variables across the entire region?
- Are the sign and magnitude of the associations constant across the region?

A review of existing methods

► Spatial regression

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- Varying coefficient regression

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 - Splines

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- Spatial regression
- Varying coefficient regression
 - Splines
 - Kernels
 - Wavelets
- Model selection via regularization

Some definitions

- ▶ Univariate spatial response process $\{Y(s) : s \in \mathcal{D}\}$
- ▶ Multivariate spatial covariate process $\{X(s) : s \in \mathcal{D}\}$
- ightharpoonup n = number of observations
- ▶ p = number of covariates
- ▶ Location (2-dimensional) s
- ► Spatial domain *D*

Further definitions

Geostatistical data:

- Observations are made at sampling locations s_i for i = 1,...,n
- E.g. elevation, temperature

► Areal data:

- Domain is partitioned into n regions $\{D_1, \ldots, D_n\}$
- The regions do not overlap, and they divide the domain completely: $\mathcal{D} = \bigcup_{i=1}^{n} D_i$
- Sampling locations s_i for i = 1, ..., n are the centroids of the regions
- E.g. poverty rate, population, spatial mean temperature

Existing approaches: spatial regression

► The typical spatial regression (?)

$$Y(\boldsymbol{s}) = \boldsymbol{X}(\boldsymbol{s})'\boldsymbol{\beta} + W(\boldsymbol{s}) + \varepsilon(\boldsymbol{s})$$

$$\operatorname{cov}(W(\boldsymbol{s}), W(\boldsymbol{t})) = \Gamma\left(\delta(\boldsymbol{s}, \boldsymbol{t})\right)$$

$$\delta(\boldsymbol{s}, \boldsymbol{t}) = \sqrt{\|\boldsymbol{t} - \boldsymbol{s}\|_2}$$
 E.g.
$$\Gamma(\delta(\boldsymbol{s}, \boldsymbol{t})) = \exp\{-\phi^{-1}\delta(\boldsymbol{s}, \boldsymbol{t})\}$$
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- ightharpoonup W(s) is a spatial random effect that accounts for autocorrelation in the response variable
- ▶ The coefficients β are constant
- ▶ Relies on a priori global variable selection

$$Y(s) = X(s)' \beta(s) + \varepsilon(s)$$

Existing approaches: varying coefficients regression

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- Assume an effect modifying variable S
- ► Coefficients are functions of S
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- This is a varying coefficient regression (VCR) model
 (?)
- If s is a spatial location then we have a spatially varying coefficient regression (SVCR) model

Existing approaches: spatially varying coefficient process

► Making model more flexible: coefficients in a spatial regression model can be allowed to vary (?)

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- ▶ $\{\beta_1(s): s \in \mathcal{D}\}, \dots, \{\beta_p(s): s \in \mathcal{D}\}$ are stationary spatial processes with Matèrn covariance functions
- ▶ Still relies on a priori global variable selection

Existing approaches: spline-based VCR and SVCR models

- Splines are a way to parameterize smooth functions
- Splines can be incorporated into a generalized additive model (GAM):

-
$$E{Y(t)} = f{X_1(t)} + \dots + f{X_p(t)}$$

▶ It is possible to parameterize a VCR model with splines for the coefficient functions:

-
$$E{Y(t)} = \beta_1(t)X_1(t) + \dots + \beta_p(t)X_p(t)$$

Existing approaches: Global selection in spline-based VCR models

Regularization methods for global variable selection in VCR models:

- ► The integral of a function squared (e.g. $\int \{f(t)\}^2 dt$) is zero if and only if the function is zero everywhere.
- Use regularization (maximize the likelihood plus a penalty) to encourage coefficient functions to be zero
- ► SCAD penalty (?) on the integral of the square of the coefficient function (?)
- ► Non-negative garrote penalty (?) on the integral of the square of the coefficient function (?)

Existing approaches: wavelet methods for VCR models

Wavelet methods involve decomposing a function into local frequency components. Wavelet methods for VCR models include using Bayesian variable selection or the Lasso to estimate which local frequency components have nonzero coefficients (??).

These methods achieve sparsity in the local frequency components but not in the local covariates, and so are not suitable for local variable selection.

Existing approaches: Local regression

Local regression uses a kernel function at each sampling location to weight observations based on their distance from the sampling location. An example is the bisquare kernel:

$$w_{ii'} = \begin{cases} \left[1 - (\phi^{-1}\delta_{ii'})^2\right]^2 & \text{if } \delta_{ii'} < \phi, \\ 0 & \text{if } \delta_{ii'} \ge \phi. \end{cases}$$
 (2)

Where ϕ is a bandwidth parameter. Given the weights, a local model is

Given the weights, a local model is fit at each sampling location using the local likelihood (?)

Existing approaches: Local likelihood

Calibrate the model by doing the following at each sampling location:

- Weight each observation's likelihood
- Weights are given by the kernel

$$\begin{split} L &= \prod_{i'=1}^{n} \left(L_{i'} \right)^{w_{ii'}} \\ \ell &= \sum_{i'=1}^{n} w_{ii'} \left\{ \log \sigma_i^2 + \sigma_i^{-2} \left(y_{i'} - \boldsymbol{x}_{i'}' \boldsymbol{\beta}_i \right)^2 \right\} \end{split}$$

Where $\beta_i = \beta(s_i)$.

Existing approaches: geographically weighted regression

When the effect modifying variable *s* refers to spatial location, the method of local regression is called geographically weighted regression (GWR) (??)

Existing approaches: bandwidth estimation for GWR

- Smaller bandwidth: less bias, more flexible coefficient surface
- Large bandwidth: less variance, less flexible coefficient surface
- ► Estimate the degrees of freedom used in estmating the coefficient surface (?):
 - $\hat{y} = Hy$ $\nu = tr(H)$
- ► Then the corrected AIC for bandwidth selection is:
- $\blacktriangleright \ \ \mathrm{AIC_c} = 2n\log\sigma + n\left\{\frac{n+\nu}{n-2-\nu}\right\}$

Introduction

Existing approaches: geographically weighted Lasso

Within a GWR model, using the Lasso (?) for local variable selection is called the geographically weighted Lasso (GWL) (?).

- The GWL requires estimating a Lasso tuning parameter for each local model
- ? estimates the local Lasso tuning parameter at location s_i by minimizing a jacknife criterion: $|y_i \hat{y}_i|$
- ► The jacknife criterion can only be calculated where data are observed, making it impossible to use the GWL to impute missing data or to estimate the value of the coefficient surface at new locations
- Also, the Lasso is known to be biased in variable selection and suboptimal for coefficient estimation

Geographically weighted elastic net (GWEN)

- Local variable selection in a GWR model using the adaptive elastic net (AEN) (?)
- Under suitable conditions, the AEN has an oracle property for selection

$$S(\beta_i) = -2\ell_i(\beta_i) + \mathcal{J}_2(\beta_i)$$

$$= \sum_{i'=1}^n w_{ii'} \left\{ \log \sigma_i^2 + \left(\sigma_i^2\right)^{-1} (y_{i'} - \boldsymbol{x}'_{i'}\beta_i)^2 \right\}$$

$$+ \alpha_i \lambda_i \sum_{j=1}^p |\beta_{ij}| / \gamma_{ij}$$

$$+ (1 - \alpha_i) \lambda_i^* \sum_{i=1}^p (\beta_{ij} / \gamma_{ij})^2$$

Geographically weighted elastic net (GWEN)

where $\sum_{i'=1}^n w_{ii'} \left(y_{i'} - x_{i'}' \beta_i \right)^2$ is the weighted sum of squares minimized by traditional GWR, and $\mathcal{J}_2(\beta_i) = \alpha_i \lambda_i^* \sum_{j=1}^p |\beta_{ij}|/\gamma_{ij} + (1-\alpha_i)\lambda_i^* \sum_{j=1}^p \left(\beta_{ij}/\gamma_{ij} \right)^2$ is the AEN penalty.

Geographically weighted elastic net (GWEN)

It is necessary to estimate an AEN tuning parameter for each local model. Using the local BIC allows fitting a local model at any location within the domain

$$\begin{split} \mathsf{BIC}_{\mathsf{loc},i} &= -2\sum_{i'=1}^n \ell_{ii'} + \log\left(\sum_{i'=1}^n w_{ii'}\right) \mathsf{df}_i \\ &= -2\sum_{i'=1}^n \log\left[\left(2\pi \hat{\sigma}_i^2\right)^{-1/2} \exp\left\{-\frac{1}{2}\hat{\sigma}_i^{-2} \left(y_{i'} - \boldsymbol{x}_{i'}' \hat{\boldsymbol{\beta}}_{i'}\right)^2\right\}\right]^w \\ &+ \log\left(\sum_{i'=1}^n w_{ii'}\right) \mathsf{df}_i \end{split}$$

(3)

Geographically weighted elastic net (GWEN)

$$= \sum_{i'=1}^n w_{ii'} \left\{ \log \left(2\pi\right) + \log \hat{\sigma}_i^2 + \hat{\sigma}_i^{-2} \left(y_{i'} - \boldsymbol{x}_{i'}' \hat{\boldsymbol{\beta}}_{i'}\right)^2 \right\}$$

$$+ \log \left(\sum_{i'=1}^n w_{ii'}\right) \mathsf{df}_i$$

Simulation study

Simulating covariates

Five covariates $\tilde{X}_1,\ldots,\tilde{X}_5$ were simulated by Gaussian random fields on the domain $[0,1]\times[0,1]$ on a 30×30 grid of sampling locations:

$$\begin{split} \tilde{X}_j \sim N(0, \Sigma) \text{ for } j = 1, \dots, 5 \\ \{\Sigma\}_{i,i'} = \exp\{-\tau^{-1}\delta_{ii'}\} \text{ for } i,i' = 1, \dots, n \end{split}$$

Where the covariates were simulated with colinearity, the colinearity was induced by multiplying the design matrix by the square root of the colinearity matrix:

$$\begin{aligned} \operatorname{diag}(\Omega_{5\times 5}) &= 1\\ \operatorname{off-diag}(\Omega_{5\times 5}) &= \rho\\ X &= \tilde{X}R \end{aligned} \tag{4}$$

Where $\Omega_{5\times 5}=R'R$ is the Cholesky decomposition.

Simulation study

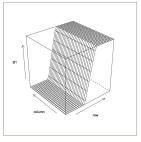
Simulating the response

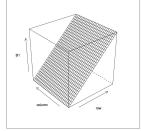
- ► $Y(s) = X(s)'\beta(s) = \sum_{j=1}^{5} \beta_j(s)X_j(s) + \varepsilon(s)$
- $\triangleright \varepsilon \sim iid \ N(0, \sigma^2)$
- ▶ $\beta_1(s)$, the coefficient function for X_1 , is nonzero in part of the domain.
- ▶ Coefficients for $X_2, ..., X_5$ are zero everywhere

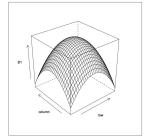
Simulation study

Coefficient functions

Call these functions step, gradient, and parabola:







Selection

			st	ер		gradient				
		е	net	la	asso	е	enet		lasso	
	location	β_1	β_2 - β_5	β						
		0.99	0.00	0.99	0.00	1.00	0.00	1.00	0.00	0.3
	1	0.99	0.02	0.99	0.02	1.00	0.01	1.00	0.01	0.
	ı	0.99	0.00	1.00	0.00	1.00	0.00	1.00	0.00	0.2
		0.96	0.05	0.91	0.04	0.99	0.03	0.99	0.01	0.
		1.00	0.00	1.00	0.00	1.00	0.01	1.00	0.00	1.0
	2	1.00	0.03	1.00	0.03	1.00	0.02	1.00	0.02	1.0
	2	1.00	0.01	1.00	0.00	1.00	0.01	1.00	0.01	1.0
		0.99	0.05	0.97	0.04	1.00	0.02	0.99	0.01	0.9
		0.91	0.01	0.91	0.00	1.00	0.01	1.00	0.01	1.0
	0	0.96	0.05	0.96	0.05	1.00	0.01	1.00	0.01	1.0
	3	0.92	0.05	0.95	0.02	1.00	0.02	1.00	0.01	1.0
		0.92	0.08	0.87	0.05	1.00	0.02	0.98	0.02	0.9
		0.48	0.01	0.43	0.01	1.00	0.01	1.00	0.01	1.0

MSE of $\beta_1(s)$

, - (,						
function	location	GWEN	GWAL	u.enet	u.lasso	oracle	GWR
		0.026	0.025	0.057	0.057	0.062	0.008
	4	0.042	0.040	0.193	0.180	0.102	0.016
	1	0.036	0.014	0.055	0.067	0.080	0.016
		0.093	0.130	0.230	0.285	0.144	0.030
		0.063	0.058	0.043	0.043	0.038	0.055
	2	0.087	0.084	0.064	0.064	0.073	0.084
	2	0.068	0.049	0.045	0.040	0.036	0.052
		0.140	0.128	0.082	0.093	0.074	0.096
		0.025	0.025	0.019	0.019	0.004	0.010
eton	3	0.021	0.021	0.015	0.015	0.007	0.011
step		0.027	0.021	0.018	0.014	0.006	0.019
		0.027	0.038	0.020	0.031	0.007	0.016
		0.026	0.026	0.028	0.025	0.034	0.054
	4	0.046	0.050	0.054	0.057	0.073	0.081
	4	0.025	0.030	0.030	0.027	0.036	0.063
		0.035	0.036	0.043	0.046	0.072	0.083
		0.000	0.000	0.000	0.000	0.000	0.008
	5	0.002	0.002	0.001	0.000	0.000	0.014
	5	0.000	0.000	0.000	0.000	0.000	0.021

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Variance of $\beta_1(s)$

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			0.036	0.036	0.195	0.180	0.098	0.016
			0.027	0.013	0.056	0.068	0.080	0.016
			0.059	0.100	0.226	0.276	0.145	0.030
			0.014	0.013	0.006	0.006	0.006	0.008
		2	0.017	0.017	0.011	0.011	0.008	0.013
		2	0.012	0.010	0.005	0.004	0.006	0.010
			0.021	0.033	0.021	0.037	0.008	0.014
			0.022	0.023	0.019	0.019	0.004	0.009
	eton	3	0.021	0.021	0.014	0.014	0.005	0.008
	step		0.024	0.021	0.018	0.014	0.005	0.016
			0.024	0.036	0.020	0.032	0.005	0.014
			0.022	0.023	0.023	0.022	0.006	0.007
		4	0.025	0.024	0.025	0.022	0.006	0.008
		4	0.021	0.025	0.024	0.023	0.005	0.013
			0.026	0.027	0.029	0.032	0.009	0.015
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		4	0.021	0.025	0.024	0.023	0.005	0.013
			0.026	0.027	0.029	0.032	0.009	0.015
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			0.027	0.013	0.056	0.068	0.080	0.016
			0.059	0.100	0.226	0.276	0.145	0.030
			0.014	0.013	0.006	0.006	0.006	0.008
		2	0.017	0.017	0.011	0.011	0.008	0.013
		2	0.012	0.010	0.005	0.004	0.006	0.010
			0.021	0.033	0.021	0.037	0.008	0.014
			0.022	0.023	0.019	0.019	0.004	0.009
	eton	3	0.021	0.021	0.014	0.014	0.005	0.008
	step		0.024	0.021	0.018	0.014	0.005	0.016
			0.024	0.036	0.020	0.032	0.005	0.014
			0.022	0.023	0.023	0.022	0.006	0.007
		4	0.025	0.024	0.025	0.022	0.006	0.008
		4	0.021	0.025	0.024	0.023	0.005	0.013
			0.026	0.027	0.029	0.032	0.009	0.015
		5	0.000	0.000	0.000	0.000	0.000	0.007
			0.002	0.002	0.001	0.000	0.000	0.014
		3	0.000	0.000	0.000	0.000	0.000	0.021

Bias of $\beta_1(s)$

	function	location	GWEN	GWAL	u.enet	u.lasso	oracle	GWR
_			-0.056	-0.049	0.001	0.005	0.015	-0.007
		1	-0.080	-0.069	0.020	0.040	0.072	0.002
		ı	-0.093	-0.037	-0.010	-0.009	-0.005	0.003
			-0.185	-0.177	-0.075	-0.110	0.032	-0.009
			-0.222	-0.213	-0.193	-0.191	-0.178	-0.217
		2	-0.264	-0.259	-0.231	-0.232	-0.256	-0.268
		2	-0.236	-0.197	-0.199	-0.188	-0.176	-0.204
			-0.345	-0.309	-0.248	-0.236	-0.257	-0.286
		3	-0.057	-0.047	-0.006	-0.006	0.025	0.024
	ston		-0.009	0.004	0.022	0.024	0.047	0.051
	step		-0.052	-0.007	0.003	0.020	0.039	0.055
			-0.062	-0.046	-0.011	-0.014	0.046	0.046
			0.066	0.058	0.071	0.057	0.168	0.218
		4	0.147	0.165	0.170	0.188	0.260	0.272
		4	0.062	0.071	0.077	0.067	0.174	0.223
			0.098	0.098	0.121	0.119	0.250	0.262
			0.000	0.000	0.000	0.000	0.000	-0.022
		_	0.003	0.001	-0.005	-0.003	0.000	-0.018

Bias of $\beta_1(s)$

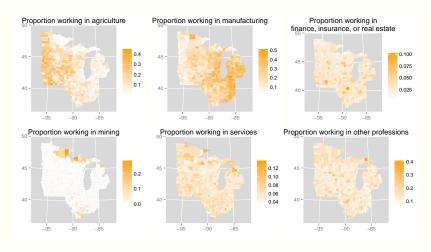
	function	location	GWEN	GWAL	u.enet	u.lasso	oracle	GWR
			-0.056	-0.049	0.001	0.005	0.015	-0.007
		4	-0.080	-0.069	0.020	0.040	0.072	0.002
		1	-0.093	-0.037	-0.010	-0.009	-0.005	0.003
			-0.185	-0.177	-0.075	-0.110	0.032	-0.009
		2	-0.222	-0.213	-0.193	-0.191	-0.178	-0.217
			-0.264	-0.259	-0.231	-0.232	-0.256	-0.268
			-0.236	-0.197	-0.199	-0.188	-0.176	-0.204
			-0.345	-0.309	-0.248	-0.236	-0.257	-0.286
		3	-0.057	-0.047	-0.006	-0.006	0.025	0.024
	-4		-0.009	0.004	0.022	0.024	0.047	0.051
	step		-0.052	-0.007	0.003	0.020	0.039	0.055
			-0.062	-0.046	-0.011	-0.014	0.046	0.046
			0.066	0.058	0.071	0.057	0.168	0.218
		1	0.147	0.165	0.170	0.188	0.260	0.272
		4	0.062	0.071	0.077	0.067	0.174	0.223
			0.098	0.098	0.121	0.119	0.250	0.262
			0.000	0.000	0.000	0.000	0.000	-0.022
			0.003	0.001	-0.005	-0.003	0.000	-0.018

Bias of $\beta_1(s)$

	function	location	GWEN	GWAL	u.enet	u.lasso	oracle	GWR
			-0.056	-0.049	0.001	0.005	0.015	-0.007
		4	-0.080	-0.069	0.020	0.040	0.072	0.002
		1	-0.093	-0.037	-0.010	-0.009	-0.005	0.003
			-0.185	-0.177	-0.075	-0.110	0.032	-0.009
		2	-0.222	-0.213	-0.193	-0.191	-0.178	-0.217
			-0.264	-0.259	-0.231	-0.232	-0.256	-0.268
			-0.236	-0.197	-0.199	-0.188	-0.176	-0.204
			-0.345	-0.309	-0.248	-0.236	-0.257	-0.286
		3	-0.057	-0.047	-0.006	-0.006	0.025	0.024
	-4		-0.009	0.004	0.022	0.024	0.047	0.051
	step		-0.052	-0.007	0.003	0.020	0.039	0.055
			-0.062	-0.046	-0.011	-0.014	0.046	0.046
			0.066	0.058	0.071	0.057	0.168	0.218
		1	0.147	0.165	0.170	0.188	0.260	0.272
		4	0.062	0.071	0.077	0.067	0.174	0.223
			0.098	0.098	0.121	0.119	0.250	0.262
			0.000	0.000	0.000	0.000	0.000	-0.022
			0.003	0.001	-0.005	-0.003	0.000	-0.018

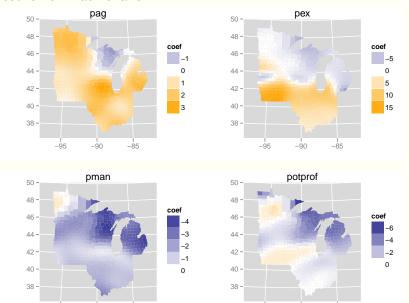
Data analysis

Revisiting the introductory example

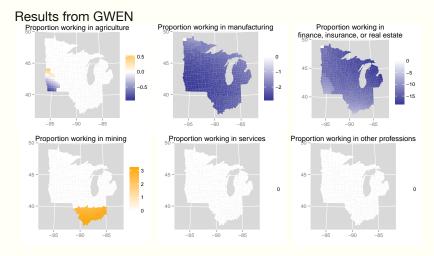


Data analysis

Results from traditional GWR



Data analysis



Future work

- Apply the GWEN to data with non-Gaussian response variable
- Incorporate spatial autocorrelation in the model (simulated errors were iid)

References I