

1. Selection

Theorem 1.1. *If $h\sqrt{n}b_n \xrightarrow{p} \infty$ then $P\left\{\hat{\beta}_{(b)}(\mathbf{s}) = 0\right\} \rightarrow 1$.*

Proof. The proof is by contradiction.

Recall that the objective to be minimized by $\hat{\beta}_{(p)}(\mathbf{s})$ is

$$Q\{\beta(\mathbf{s})\} = (1/2)\{\mathbf{Y} - \mathbf{Z}(\mathbf{s})\beta(\mathbf{s})\}^T \mathbf{W}(\mathbf{s})\{\mathbf{Y} - \mathbf{Z}(\mathbf{s})\beta(\mathbf{s})\} + n \sum_{j=1}^p \lambda_j \|\beta_p(\mathbf{s})\| \quad (1)$$

Let $\hat{\beta}_p(\mathbf{s}) \neq 0$. Then (1) is differentiable w.r.t. $\beta_p(\mathbf{s})$ and Q is maximized at

$$\begin{aligned} 0 &= \mathbf{Z}_{(p)}^T(\mathbf{s})\mathbf{W}(\mathbf{s})\left\{\mathbf{Y} - \mathbf{Z}_{(-p)}(\mathbf{s})\hat{\beta}_{(-p)}(\mathbf{s}) - \mathbf{Z}_{(p)}(\mathbf{s})\hat{\beta}_{(p)}(\mathbf{s})\right\} + n\lambda_p \frac{\hat{\beta}_{(p)}(\mathbf{s})}{\|\beta_{(p)}(\mathbf{s})\|} \\ &= \mathbf{Z}_{(p)}^T(\mathbf{s})\mathbf{W}(\mathbf{s})\left[\mathbf{Y} - \mathbf{Z}(\mathbf{s})\beta(\mathbf{s}) - \frac{h^2\kappa_2}{2\kappa_0}\{\beta_{uu}(\mathbf{s}) + \beta_{vv}(\mathbf{s})\}\right] \\ &\quad + \mathbf{Z}_{(p)}^T(\mathbf{s})\mathbf{W}(\mathbf{s})\mathbf{Z}_{(-p)}(\mathbf{s})\left[\beta_{(-p)}(\mathbf{s}) + \frac{h^2\kappa_2}{2\kappa_0}\{\beta_{(-p),uu}(\mathbf{s}) + \beta_{(-p),vv}(\mathbf{s})\} - \hat{\beta}_{(-p)}(\mathbf{s})\right] \\ &\quad + \mathbf{Z}_{(p)}^T(\mathbf{s})\mathbf{W}(\mathbf{s})\mathbf{Z}_{(p)}(\mathbf{s})\left[\beta_{(p)}(\mathbf{s}) + \frac{h^2\kappa_2}{2\kappa_0}\{\beta_{(p),uu}(\mathbf{s}) + \beta_{(p),vv}(\mathbf{s})\} - \hat{\beta}_{(p)}(\mathbf{s})\right] \\ &\quad + n\lambda_p \frac{\hat{\beta}_{(p)}(\mathbf{s})}{\|\beta_{(p)}(\mathbf{s})\|} \\ &= hn^{-1/2}\mathbf{Z}_{(p)}^T(\mathbf{s})\mathbf{W}(\mathbf{s})\left\{\mathbf{Y} - \mathbf{Z}(\mathbf{s})\beta(\mathbf{s}) - \frac{h^2\kappa_2}{2\kappa_0}\{\beta_{uu}(\mathbf{s}) + \beta_{vv}(\mathbf{s})\}\right\} \\ &\quad + \left\{n^{-1}\mathbf{Z}_{(p)}^T(\mathbf{s})\mathbf{W}(\mathbf{s})\mathbf{Z}_{(-p)}(\mathbf{s})\right\}h\sqrt{n}\left[\beta_{(-p)}(\mathbf{s}) + \frac{h^2\kappa_2}{2\kappa_0}\{\beta_{(-p),uu}(\mathbf{s}) + \beta_{(-p),vv}(\mathbf{s})\} - \hat{\beta}_{(-p)}(\mathbf{s})\right] \\ &\quad + \left\{n^{-1}\mathbf{Z}_{(p)}^T(\mathbf{s})\mathbf{W}(\mathbf{s})\mathbf{Z}_{(p)}(\mathbf{s})\right\}h\sqrt{n}\left[\beta_{(p)}(\mathbf{s}) + \frac{h^2\kappa_2}{2\kappa_0}\{\beta_{(p),uu}(\mathbf{s}) + \beta_{(p),vv}(\mathbf{s})\} - \hat{\beta}_{(p)}(\mathbf{s})\right] \\ &\quad + h\sqrt{n}\lambda_p \frac{\hat{\beta}_{(p)}(\mathbf{s})}{\|\beta_{(p)}(\mathbf{s})\|} \quad (2) \end{aligned}$$

From Lemma 2 of ?, $n^{-1}\left\{\mathbf{Z}_{(p)}^T(\mathbf{s})\mathbf{W}(\mathbf{s})\mathbf{Z}_{(-p)}(\mathbf{s})\right\} = O_p(1)$ and $n^{-1}\left\{\mathbf{Z}_{(p)}^T(\mathbf{s})\mathbf{W}(\mathbf{s})\mathbf{Z}_{(p)}(\mathbf{s})\right\} =$

$O_p(1)$. From Theorem 3 of ?, we have that $hn^{1/2} \left[\hat{\beta}_{(-p)}(\mathbf{s}) - \beta_{(-p)}(\mathbf{s}) - \frac{h^2 \kappa_2}{2\kappa_0} \{ \beta_{(p),uu}(\mathbf{s}) + \beta_{(p),vv}(\mathbf{s}) \} \right] = O_p(1)$ and $hn^{1/2} \left[\hat{\beta}_{(p)}(\mathbf{s}) - \beta_{(p)}(\mathbf{s}) - \frac{h^2 \kappa_2}{2\kappa_0} \{ \beta_{(p),uu}(\mathbf{s}) + \beta_{(p),vv}(\mathbf{s}) \} \right] = O_p(1)$. So the second and third terms of the sum in (2) are $O_p(1)$. We showed in the proof of ?? that $hn^{-1/2} \mathbf{Z}_{(p)}^T(\mathbf{s}) \mathbf{W}(\mathbf{s}) \{ \mathbf{Y} - \mathbf{Z}(\mathbf{s}) \beta(\mathbf{s}) \} = O_p(1)$.

Because the first three terms of the sum in 2 are $O_p(1)$, for $\hat{\beta}_{(p)}(\mathbf{s})$ to be a solution, we must have that $hn^{1/2} \lambda_p \frac{\hat{\beta}_{(p)}(\mathbf{s})}{\|\hat{\beta}_{(p)}(\mathbf{s})\|} = O_p(1)$.

But since by assumption $\hat{\beta}_{(p)}(\mathbf{s}_i) \neq 0$, there must be some $k \in \{1, \dots, d_p\}$ such that $|\hat{\beta}_{(p),k}(\mathbf{s})| = \max\{|\hat{\beta}_{(p),k'}(\mathbf{s})| : 1 \leq k' \leq d_p\}$. And for this k , we have that $|\hat{\beta}_{(p),k}(\mathbf{s})|/\|\hat{\beta}_{(p)}(\mathbf{s})\| \geq 1/\sqrt{d_p} > 0$.

Now since $\sqrt{n}b_n \rightarrow \infty$, we have that $hn^{1/2} \lambda_p \frac{\hat{\beta}_{(p)}(\mathbf{s})}{\|\hat{\beta}_{(p)}(\mathbf{s})\|}$ is unbounded and therefore dominates the $O_p(1)$ terms of the sum in (2). So for large enough n , $\hat{\beta}_{(p)}(\mathbf{s}) \neq 0$ cannot maximize Q . \square

Theorem 1.2. *If $h\sqrt{n}a_n \xrightarrow{p} 0$ then $P\left\{\hat{\beta}_{(a)}(\mathbf{s}) \neq 0\right\} \rightarrow 1$.*

Proof. Again, the proof is by contradiction.

Assume that $\hat{\beta}_{(k)} = 0$ for some $k < p_0$. For the adaptive group lasso, the covariate group k is shrunk to zero if

$$\left\| \left\{ \mathbf{Z}_{(k)}^T(\mathbf{s}) \mathbf{W}(\mathbf{s}) \mathbf{Z}_{(k)}^T(\mathbf{s}) \right\}^{-1} \mathbf{Z}_{(k)}^T(\mathbf{s}) \mathbf{r}_{(k)}(\mathbf{s}) \right\|^2 \leq \frac{h^2 n \lambda^2}{\|\tilde{\beta}_{(k)}\|^2}$$

where $\mathbf{r}_{(k)}(\mathbf{s})$ is the residual after accounting for all covariate groups except group k . That is,

$\mathbf{r}_{(k)}(\mathbf{s}) = \mathbf{Y} - \mathbf{Z}_{(-k)}(\mathbf{s}) \beta_{(-k)}(\mathbf{s})$. But $\|\tilde{\beta}_{(k)}\| > 0$ implies that $\left\| \left\{ \mathbf{Z}_{(k)}^T(\mathbf{s}) \mathbf{W}(\mathbf{s}) \mathbf{Z}_{(k)}^T(\mathbf{s}) \right\}^{-1} \mathbf{Z}_{(k)}^T(\mathbf{s}) \mathbf{r}_{(k)}(\mathbf{s}) \right\|^2 > 0$ and $\frac{h^2 n \lambda^2}{\|\tilde{\beta}_{(k)}\|^2} \leq h^2 n a_n^2 \rightarrow 0$. So

$$P\left\{\hat{\boldsymbol{\beta}}_{(k)}(\mathbf{s}) \neq 0\right\} \leq P\left[\left\|\left\{\mathbf{Z}_{(k)}^T(\mathbf{s})\mathbf{W}(\mathbf{s})\mathbf{Z}_{(k)}^T(\mathbf{s})\right\}^{-1}\mathbf{Z}_{(k)}^T(\mathbf{s})\mathbf{r}_{(k)}(\mathbf{s})\right\|^2 \leq h^2na_n^2\right] \rightarrow 0. \quad (3)$$

□