

Bayesian Data Analysis CHAPTER 17: Metric Predicted Variable with One Metric Predictor

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Overview of Chapter 17

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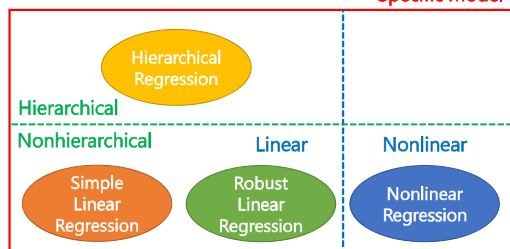
■ Purpose

- ✓ Predict one metric variable from one metric predictor

ex.) y x
weight height

■ Outline

Specific Model



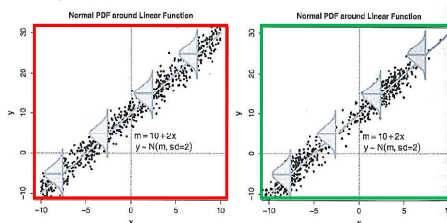
17.1 Simple Linear Regression

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■ Step

- ✓ Generate any random x
- ✓ Compute the mean predicted value of y by $\mu = \theta_0 + \theta_1 x$
- ✓ Generate random variable for datum y from a normal distribution (μ : mean, σ : standard deviation)

■ Figure



Generate x from
Left :
a uniform distribution
Right :
a bimodal distribution
Both shows data from
the same model

17.1 Simple Linear Regression

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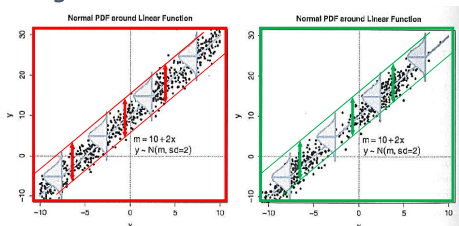
■ Assumption

- ✓ *Homogeneity of variance*
At every value of x , the variance of y is the same

■ Note

- ✓ Simple Linear Regression describes **tendencies**, not causality

■ Figure



Both assumes
Homogeneity of variance

But hard to see it
on the right figure
because there is an area
which x is sparse

17.2 Robust Linear Regression

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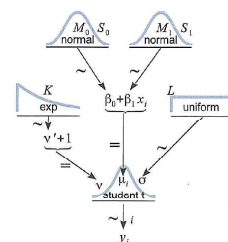
■ Object

- ✓ Data which have outliers

■ Assumption

- ✓ The datum y_i is a **t-distributed** random value around the central tendency $\mu_i = \theta_0 + \theta_1 x_i$

■ Diagram



17.2 Robust Linear Regression

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■ Goal

- ✓ Determine what combinations of $\theta_0, \theta_1, \sigma, \nu$ are credible, given the data
The answer (from Bayes' rule):

$$p(\beta_0, \beta_1, \sigma, \nu | D) = \frac{p(D | \beta_0, \beta_1, \sigma, \nu) p(\beta_0, \beta_1, \sigma, \nu)}{\int \int \int \int d\beta_0 d\beta_1 d\sigma d\nu p(D | \beta_0, \beta_1, \sigma, \nu) p(\beta_0, \beta_1, \sigma, \nu)}$$

→ Complicated...

Use JAGS or Stan!

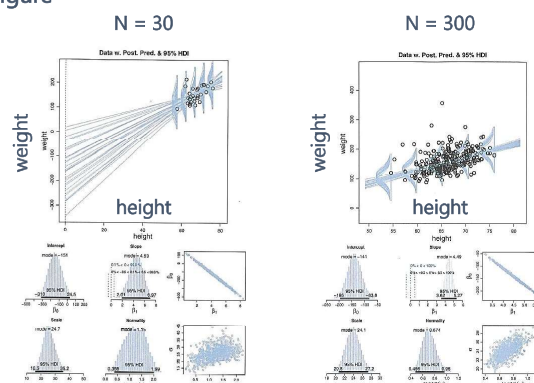
■ What we have to do

- ✓ Specify sensible priors
- ✓ Make sure that the MCMC process generates a trustworthy sample that is converged and well mixed
→ Talk about it later

17.2 Robust Linear Regression

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■ Figure



■ Problem with using raw data

- ✓ Parameter-correlation problem

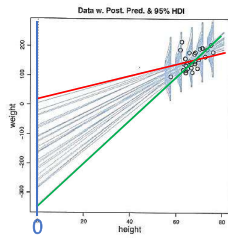
The credible slopes and intercepts trade off

When the slope is small, the intercept is big

When the slope is big, the intercept is small

→ MCMC sampling is difficult

Two parameter values change slowly



■ Ways to make the sampling more efficient

- ✓ Change the sampling algorithm → Stan : HMC
- ✓ Transform the data → JAGS : Standardization

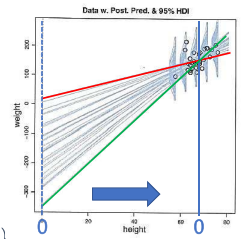
■ Mean centering

- ✓ Slide the axis

so that zero falls under the mean

→ The Slope changes without any big changes on the intercept

→ Solve parameter-correlation problem(???)



■ Standardize data

- ✓ Re-scaling the data relative to their mean (M)

and standard deviation (SD): $z_x = \frac{(x - M_x)}{SD_x}$ and $z_y = \frac{(y - M_y)}{SD_y}$

- ✓ Linear Regression using standardized data

$$\begin{aligned} z_y &= \beta_0 + \beta_1 z_x \\ \frac{(\hat{y} - M_y)}{SD_y} &= \beta_0 + \beta_1 \frac{(x - M_x)}{SD_x} \end{aligned} \quad \text{by definition of the model}$$

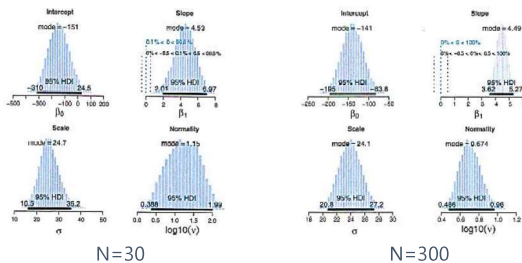
from Equation 17.1

$$\hat{y} = \underbrace{\beta_0}_{\beta_0} + \underbrace{\beta_1}_{\beta_1} \frac{(x - M_x)}{SD_x} \quad (17.2)$$

β_0 : the slope with the data
 β_1 : the intercept with the data

■ Interpreting the posterior distribution

- ✓ Compare N=30 regression and N=300 one
- ✓ The slope, intercept and scale are about the same
- ✓ The certainty of the estimate for N=300 is tighter than for N=30
- ✓ The normality parameter for N=300 is bigger than for N=30



■ Object

- ✓ Data that each individual contributes multiple observations ex.) Reading-ability scores of children across several years
Family income for different size of the family, for different regions

■ Assumption

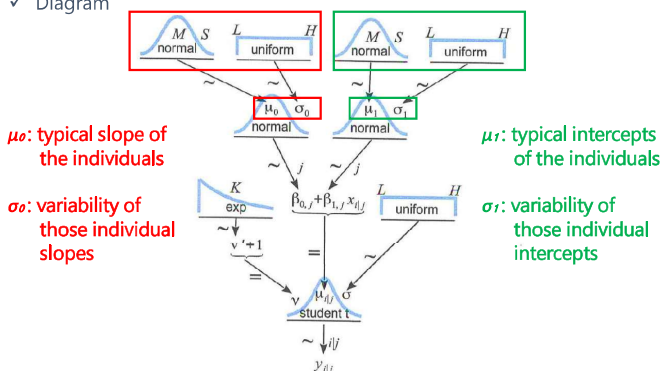
- ✓ Each individual is representative of the group
→ Every individual informs the estimate of the group slope and intercept
Get sharing of information across individuals

■ Goal

- ✓ describe each individual with a linear regression
- ✓ Estimate the typical slope and intercept of the group overall

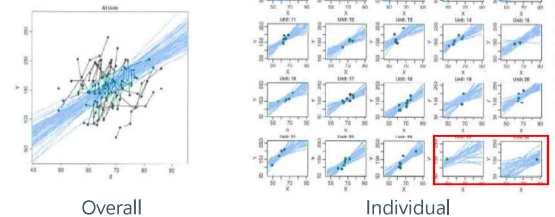
■ The model and implementation in JAGS

- ✓ Diagram



■ The posterior distribution : Shrinkage and prediction

- ✓ Overall : Clearly positive by integrating each individual slope
- ✓ Individual : Notable shrinkage of the estimates of the individuals
- ✓ The estimates are tightly constrained by each data



■ Object

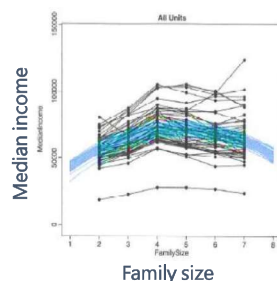
- ✓ Data that y appears to have a nonlinear trend as x increases

■ Example

- ✓ Family size and family income for each state in the U.S

■ Model

- ✓ $\mu_i = b_0 + b_1 x + b_2 x^2$
- ✓ If $b_2 = 0$ → linear model
- ✓ If $|b_2|$ is big
→ nonlinear model is reasonable



■ Nonlinear Regression using standardized data

$$\begin{aligned} z_y &= \beta_0 + \beta_1 z_x + \beta_2 z_x^2 \\ \frac{(\hat{y} - M_y)}{SD_y} &= \beta_0 + \beta_1 \frac{(x - M_x)}{SD_x} + \beta_2 \frac{(x - M_x)^2}{SD_x^2} \end{aligned} \quad \text{by definition of the model}$$

from Equation 17.1

$$\hat{y} = \underbrace{\beta_0}_{\beta_0} + \underbrace{\beta_1}_{\beta_1} \frac{(x - M_x)}{SD_x} + \underbrace{\beta_2}_{\beta_2} \frac{(x - M_x)^2}{SD_x^2} \quad (17.3)$$

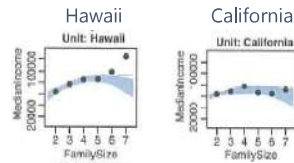
β_0 : the slope with the data
 β_1 : the intercept with the data
 β_2 : the quadratic coefficient with the data

■ Weighting data

- ✓ The data report the median income based on different numbers of families at each size
→ every median has a different amount of sampling noise
- ✓ Consider "margin of error"
If margin of error is **high**, noise parameter should be **increased**
If margin of error is **small**, noise parameter should be **decreased**

Results and interpretation

- ✓ The quadratic coefficient is $-2200 \sim -1700$
 - Nonlinear model is reasonable
- ✓ Hawaii (the amount of data is not big)
 - The trend is **upward**, but the curve is **downward curvature**
 - Shrinkage from the group
- ✓ California
 - a **narrow** spread at family size 2
 - a **large** spread at family size 7
 - the most of the data for large family sizes have **large standard errors**



Further extensions

	An example of family income and family size	Extensions
Trend	linear quadratic	higher-order polynomial Sinusoidal exponential
Noise distribution	a single lying noise for all individuals	vary among individuals
Distribution for parameters	normal distribution	t distribution
Considering Covariation	No	Use a multivariate normal prior on the intercept and slope

17. 5 Procedure and Perils for Expanding a Model

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Posterior predictive check

- ✓ Visualize the data and the posterior predictions
 - If the prediction doesn't seem to fit the data, change the model
 - New model should be both **meaningful** and **computationally tractable**
- ✓ Create a posterior predictive sampling distribution
 - measure of discrepancy between the predictions and the data

Ways to extend a model

- ✓ Add a parameter
 - ex.) $\mu_i = \beta_0 + \beta_1 X_i \rightarrow \mu_i = \beta_0 + \beta_1 X_i + \beta_2 X_i^2$
 - You can check the validity of the model by considering β_2
- ✓ Try a completely different model
 - Compare models by **Bayesian model comparison**
- ✓ "double dipping": data are used to change the prior distribution

17. 5 Procedure and Perils for Expanding a Model

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Steps to extend a JAGS or Stan model

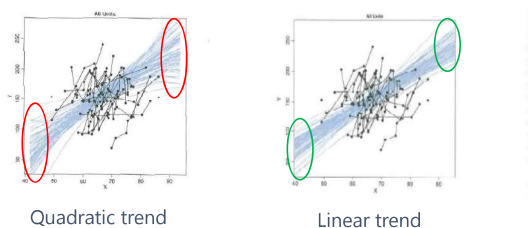
- ✓ Carefully specify the model with its new parameters
 - Draw a diagram
- ✓ Be sure all the new parameters have sensible priors
- ✓ Define initial values for all the new parameters
 - You can let JAGS initialize parameters automatically
- ✓ Tell JAGS to track the new parameters
 - Stan automatically tracks
- ✓ Modify the summary and graphics output to properly display the extended model
 - You should write **R code** because graphics are displayed by R

17. 5 Procedure and Perils for Expanding a Model

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Perils of adding parameters

- ✓ Increase in uncertainty of a parameter estimate
 - The curvature and slope trade-off strongly
 - Even if curvature is 0, the certainty of slope decreases
- ✓ The one of the ways to solve the problem is standardizing the data



Quadratic trend

Linear trend