
Cardinality constrained submodular maximization for random streams

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Abstract

We consider the problem of maximizing submodular functions in single-pass streaming and secretaries-with-shortlists models, both with random arrival order. For cardinality constrained monotone functions, Agrawal, Shadravan, and Stein [ASS19] gave a single-pass $(1 - 1/e - \varepsilon)$ -approximation algorithm using only linear memory, but their exponential dependence on ε makes it impractical even for $\varepsilon = 0.1$. We simplify both the algorithm and the analysis, obtaining an exponential improvement in the ε -dependence (in particular, $O(k/\varepsilon)$ memory). Extending these techniques, we also give a simple $(1/e - \varepsilon)$ -approximation for non-monotone functions in $O(k/\varepsilon)$ memory. For the monotone case, we also give a corresponding unconditional hardness barrier of $1 - 1/e + \varepsilon$ for single-pass algorithms in randomly ordered streams, even assuming unlimited computation.

Finally, we show that the algorithms are simple to implement and work well on real world datasets.

1 Introduction

Over the past few decades, submodularity has become recognized as a useful property occurring in a wide variety of discrete optimization problems. Submodular functions model the property of diminishing returns, whereby the gain in a utility function decreases as the set of items considered increases. This property occurs naturally in machine learning, information retrieval, and influence maximization, to name a few (see [IKBA20] and the references within).

In many settings, the data is not available in a random-access model; either for external reasons (customers arriving online) or because of the massive amount of data. When the data becomes too big to store in memory, we look for *streaming algorithms* that pass (once) through the data, and efficiently decide for each element whether to discard it or keep it in a small buffer in memory.

In this work, we consider algorithms that process elements arriving in a random order. Note that the classical greedy algorithm which iteratively adds the best element [NWF78] cannot be used here, and hence we must look for new algorithmic techniques. The motivation for considering random element arrival comes from the prevalence of submodularity in big-data applications, in which data is often logged in batches that can be modelled as random samples from an underlying

distribution.¹ The problem of streaming submodular maximization has recently received significant attention both for random arrival order and the more pessimistic worst-case order arrival [ASS19, AEF⁺20, AF19, BMKK14, BV14, CGQ15, FKK18, FNSZ20, HKMY20, IV19, KMZ⁺19, MV19, NTM⁺18, HTW20, Sha20].

1.1 Submodular functions and the streaming model

Let $f : 2^E \rightarrow \mathbb{R}_+$ be a non-negative set function satisfying $f(S \cup \{e\}) - f(S) \geq f(T \cup \{e\}) - f(T)$ for all $S \subseteq T \subseteq E \setminus \{e\}$. Such a function is called *submodular*. For simplicity, we assume $f(\emptyset) = 0$.² We use the shorthand $f(e|S) := f(S \cup \{e\}) - f(S)$ to denote the *marginal* of e on top of S . When $f(S) \leq f(T)$ for all $S \subseteq T$, f is called *monotone*. We consider the following optimization problem:

$$OPT := \max_{|S| \leq k} f(S)$$

where k is a cardinality constraint for the solution size.

Our focus on submodular maximization in the streaming setting. In this setting, an algorithm is given a single pass over a dataset in a streaming fashion, where the stream is a some permutation of the input dataset and each element is seen once. The stream is in *random order* when the permutation is uniformly random. When there are no constraints on the stream order, we call the stream *adversarial*.

At each step of the stream, our algorithm is allowed to maintain a buffer of input elements. When a new element is streamed in, the algorithm can choose to add the element to its buffer. To be as general as possible, we assume an oracle model — that is, we assume there is an oracle that returns the value of $f(S)$ for any set S . The decision of the algorithm to add an element is based only on queries to the oracle on subsets of the buffer elements. The algorithm may also choose to throw away buffered elements at any given time. The goal in the streaming model is to minimize memory use, and the complexity of the algorithm is the maximum number of input elements the algorithm stores in the buffer at any given time.

For the oracle model, an important distinction is between *weak oracle* access or and *strong oracle* access. In the *weak oracle* setting, the algorithm is only allowed to query sets of feasible elements (sets that have cardinality less than k). In the *strong oracle* setting however, the algorithm is allowed to query any set of elements. All our results apply to both the weak and strong oracle models.

Our aim will be to develop algorithms that only make one pass over the data stream, using $\tilde{O}_\varepsilon(k)$ memory, where k is the maximum size of a solution set in \mathcal{I} . We assume that k is small relative to $n := |E|$, the size of the ground set.

1.2 Our contributions

On the algorithmic side, we give $(1 - 1/e - \varepsilon)$ and $(1/e - \varepsilon)$ approximation for cardinality constrained monotone and non-monotone submodular maximization respectively, both using $O(k/\varepsilon)$ memory. The monotone result has an exponential improvement in memory requirements compared to Agrawal et al. [ASS19] (in terms of the dependence on ε), while the non-monotone result is the first to appear in the random-order streaming model, and improves upon the best known polynomial-time approximations under adversarial orders [AEF⁺20]. The algorithms are extremely simple to implement, and perform well on real world data (see Section 5), even compared to *offline* greedy algorithms.

On the hardness side, we prove that a $(1 - 1/e + \varepsilon)$ -approximation for monotone submodular maximization would require $\Omega(n)$ memory (even with unlimited queries and computational power). This improves the hardness bound of $7/8$ from [ASS19].

1.3 Related work

Prior work on this problem has focused on both the adversarial and random-order streaming setting. Algorithmic and hardness results further depend on whether the function f is monotone or non-

¹Note that the random order assumption generalizes the typical assumption of i.i.d. sampling stream elements from a known distribution, as the random order assumption applies equally well to unknown distributions.

²Submodular functions f with $f(\emptyset) > 0$ only improves the approximation ratio of our algorithms.

	adversarial order	random order
monotone	$\geq 1/2 - \varepsilon$ ($O(k/\varepsilon)$ [KMZ ⁺ 19]) $\leq 1/2 + \varepsilon$ (W+S [NTM ⁺ 18, FNSZ20]) $\leq 2/(2 + \sqrt{2})$ (S [HKMY20])	$\geq 1/2 + c$, $c < 10^{-13}$ ($O(k \log k)$ [NTM ⁺ 18]) $\geq 1 - 1/e - \varepsilon$ ($O(k \exp(\text{poly}(1/\varepsilon)))$ [ASS19]) $\leq 1 - 1/e + \varepsilon$ (S) $\geq 1 - 1/e - \varepsilon$ ($O(k/\varepsilon)$, this paper)
non-mono.	$\geq 1/2 - \varepsilon$ (unbounded comp, $O(k/\varepsilon^2)$ [AEF ⁺ 20]) ≥ 0.2779 ($O(k)$ [AEF ⁺ 20]) ≥ 0.1715 ($O(k)$ [FKK18]) $< 1/2 + \varepsilon$ (W [AEF ⁺ 20])	$\geq 1/e$ ($O(k/\varepsilon)$, this paper)

Table 1: A survey of results on monotone submodular streaming. Our contributions are highlighted in magenta. W = weak oracle, S = strong oracle, complexities refer to memory use. Lower-bounds propagate downwards in the table. Note that our polynomial-time algorithms require only weak oracle access, while our hardness results hold even for strong oracle.

monotone, and whether f has explicit structure (e.g. such as by presenting a set system for f in the coverage case), or accessible only via oracle queries. Table 1 describes all the relevant results.

Algorithmic results. Submodular maximization in the streaming setting was first considered by Badanidiyuru et al. [BMKK14] who gave a $(1/2 - \varepsilon)$ -approximation in $O(k \log k/\varepsilon)$ memory for monotone submodular functions under a cardinality constraint, using a thresholding idea with parallel enumeration of possible thresholds. This work led to a number of subsequent developments with the current best being a $(1/2 - \varepsilon)$ -approximation in $O(k/\varepsilon)$ memory [KMZ⁺19]. It turns out that the factor of $1/2$ is the best possible in the adversarial setting (with a weak oracle), but an improvement is possible in the random order model (the input is ordered by a uniformly random permutation). This was first shown by Norouzi-Fard et al. [NTM⁺18], who proved that the $1/2$ -hardness barrier for cardinality constraints can be broken, exhibiting a $(1/2 + c)$ -approximation in $O(k/\varepsilon \log k)$ memory where $c < 10^{-13}$. In a breakthrough work, Agrawal, Shadravan, and Stein [ASS19] gave a $(1 - 1/e - \varepsilon)$ -approximation using $k 2^{\text{poly}(1/\varepsilon)}$ memory and running time. We note that this is arbitrarily close to the optimal factor of $1 - 1/e$, but the algorithm is not practical, due to its dependence on ε (even for $\varepsilon = 0.1$, the resulting constants are astronomical).

Lower bounds. A few lower bounds are known for monotone functions in the *adversarial order model*: with a weak oracle, any $(1/2 + \varepsilon)$ -approximation would require $\Omega(n/k)$ memory [NTM⁺18]. Under a strong oracle, a lower bound of $\Omega(n/k^3)$ memory for any $(1/2 + \varepsilon)$ -approximation algorithm was shown in a recent paper by Feldman et al. [FNSZ20]. Another recent lower bound was proved by McGregor and Vu [MV19]: a $(1 - 1/e + \varepsilon)$ -approximation for coverage functions requires $\Omega(n/k^2)$ memory (this lower bound holds for explicitly given inputs, via communication complexity; we note that this is incomparable to the computational $(1 - 1/e + \varepsilon)$ -hardness of maximum coverage [Fei98]). For non-monotone functions, Alaluf et al. [AEF⁺20] proved an $\Omega(n)$ memory lower bound for the adversarial order model with unbounded computation.

In the random-order model, Agrawal et al. [ASS19] show that beating $7/8$ (for monotone submodular functions) requires $\Omega(n)$ memory. In contrast, we show that same construction as McGregor and Vu [MV19] also applies to randomly ordered streams: $(1 - 1/e + \varepsilon)$ for coverage functions requires $\Omega(n/k^2)$ memory even in the random-order model.

Submodular maximization in related models. A closely related model is the secretary with shortlists model [ASS19], where an algorithm is allowed to store a shortlist of more than k items (where k is the cardinality constraint). Unlike the streaming model however, once an element goes into the shortlist, it cannot be removed. Then, after seeing the entire stream, the algorithm chooses a subset of size k from the shortlist and returns that to the user. We note that the algorithms developed in this paper apply almost without modification to the shortlists model.

1.4 Overview of our techniques

Main algorithmic techniques. The primary impetus for our algorithmic work was an effort to avoid the extensive enumeration involved in the algorithm of Agrawal et al. [ASS19] which leads to memory requirements *exponential* in $\text{poly}(1/\epsilon)$.

To make things concrete, let us consider the input divided into disjoint *windows* of consecutive elements. The windows containing actual optimal elements play a special role — let’s call them *active windows* — these are the windows where we make quantifiable progress. When the stream is randomly ordered, we would ideally like to have each new element sampled independently and uniformly from the input. This leads to the intuition that the optimal elements are evenly spread out through all the windows. This cannot be literally true, since conditioned on the history of the stream, some elements have already appeared and cannot appear again. However, a key idea of Agrawal et al. [ASS19] allows us to circumvent this by reinserting the elements that we have already seen and that played a role in the selection process. What needs to be proved is that elements that were not selected can still appear in the future, conditioned on the history of the algorithm; that turns out to be true, provided that our algorithm operates in a certain greedy-like manner.

To ensure progress was made regardless of the positioning of the optimal elements, previous work made use of exponentially large enumerations to essentially guess which windows the optimal elements arrive in. Where we depart from previous work is the way we build our solution. The idea is to use an evolving family of solutions which are updated in parallel, so that we obtain a quantifiable gain regardless of where the optimal elements arrived. Specifically, we grow k solutions in parallel, where solution L_i has cardinality i . In each window, we attempt to extend a collection of solutions L_i (for varying i) by a new element e ; if e is beneficial on average to every L_i in the collection, we replace each L_{i+1} with the new solution $L_i \cup \{e\}$. Regardless of which windows happen to be active, we will show that the average gain over our evolving collection of solutions is analogous to the greedy algorithm. This is the basis of the analysis that leads to a factor of $1 - 1/e$.

In addition to our candidate solutions L_i , we maintain a pool of elements H that our algorithm has ever included in some candidate solution. We then use H to reintroduce elements artificially back into the input; this makes it possible to assume that every input element still appears in a future window with the same probability, which is key to the probabilistic analysis leading to $1 - 1/e$.

Non-monotone functions. Our algorithm for non-monotone submodular functions is similar, with the caveat that here we also have to be careful about not including any element in the solution with large probability. This is an important aspect of the randomized greedy algorithm for (offline) non-monotone submodular maximization [BFNS14] which randomly includes in each step one of the top k elements in terms of marginal values. We achieve a similar property by choosing the top element from the current window and a random subset of the pool of prior elements H .

Hardness results. Our hardness instances have the following general structure: there is a special subset of k *good* elements, and the remaining $n - k$ elements are *bad*. The good elements are indistinguishable from each other, and ditto for the bad elements. In the monotone case, any k bad elements are a $(1 - 1/e)$ -factor worse than the optimal solution (k good elements). Suppose furthermore that for parameter $r > 0$, as long as we never query the function on a subset with $\geq r$ good elements, the good elements are indistinguishable from bad elements. The only way to collect $\geq r$ good elements in the memory buffer is by chance – until we’ve collected the required number of good elements, they are indistinguishable from bad elements, so the subset in the memory buffer is random. The classic work of Nemhauser and Wolsey [NW78] constructs a pathological monotone submodular function with $r = \Omega(k)$, which we use to prove that without memory $\Omega(n)$ the algorithm cannot beat $1 - 1/e$. McGregor and Vu [MV19] construct a simple example of a coverage function with $r = 2$, which we use for our $\Omega(n/k^2)$ bound. For an exponential-size ground set, we extend their construction to $r = 3$ which translates to the improved lower bound of $\Omega(n/k^{3/2})$.

2 A $(1 - 1/e - \epsilon)$ -approximation in $O(k/\epsilon)$ memory

In this section, we develop a simple algorithm for optimizing a monotone submodular function with respect to a cardinality constraint. For the sake of exposition, we focus on the intuition behind the results and relegate full proofs to the appendix.

Our algorithm begins by randomly partitioning the stream E into αk contiguous *windows* of expected size $n/(\alpha k)$, where $\alpha > 1$ is a parameter controlling the memory dependence and approximation ratio. This is done by generating a random partition according to Algorithm 1. As the algorithm progresses, it maintains k partial solutions, the ℓ -th of which contains exactly ℓ elements. Within each window we process all the elements independently, and choose one candidate element e to extend the partial solutions by. We then add e to a collection of partial solutions L_ℓ at the end of the window. The range of partial solution sizes ℓ that we use roughly tracks the number of optimal elements we are expected to have seen so far in the stream.

Intuitively, our algorithm is guaranteed to make progress on windows that contain an element from the optimal solution O . Let us loosely call such windows *active* (a precise definition will be given later). Of course, the algorithm never knows which windows are active. However, the key idea of our analysis is that we are able to track the progress that our algorithm makes on active windows. Since the input stream is uniformly random, intuitively we expect to see i/α optimal elements after processing i windows. With high probability, the true number of optimal elements seen will be in the range $\mathcal{R} := [i/\alpha - \tilde{O}(\sqrt{k}), i/\alpha + \tilde{O}(\sqrt{k})]$. By focusing on the average improvement over levels in \mathcal{R} , we can show that each level ℓ in this range gains $\frac{1}{k}(f(O) - f(L_{\ell-1}))$ in expectation, whenever a (random) optimal element arrives.

For the analysis to work, ideally we would like each arriving optimal element to be selected uniformly among all optimal elements. This is not true conditioned on the history of decisions made by the algorithm. However, we can remedy this by re-inserting elements that we have selected before and subsampling the elements in the current window, with certain probabilities. A key lemma (Lemma 2.2) shows why this works, since the elements we have never included might still appear, given the history of the algorithm. Our basic algorithm is described in Algorithm 2, with the window partitioning procedure described in Algorithm 1.

Algorithm 1 Partitioning stream E into m windows.

- 1: Draw $|E|$ integers uniformly from $1, \dots, m$
 - 2: $n_i \leftarrow \#$ of integers equal to i
 - 3: $t_i \leftarrow \sum_{1 \leq j < i} n_j$
 - 4: **for** $i = 1, \dots, m$ **do**
 - 5: $w_i \leftarrow$ elements t_i to $t_i + n_i$ in E
 - 6: **return** $\{w_1, w_2, \dots, w_m\}$
-

Algorithm 2 MONOTONESTREAM(f, E, k, α)

- 1: Partition E into windows w_i for $i = 1, \dots, \alpha k$ with Algorithm 1.
 - 2: $L_\ell^0 \leftarrow \emptyset$ for $\ell = 0, 1, \dots, k$
 - 3: $H \leftarrow \emptyset$
 - 4: **for** $i = 1, \dots, \alpha k$ **do**
 - 5: $C_i \leftarrow w_i$
 - 6: $L_\ell^i \leftarrow L_\ell^{i-1}$ for $\ell = 0, 1, \dots, k$
 - 7: Sample each $e \in H$ with probability $\frac{1}{\alpha k}$ and add to a set R_i
 - 8: $C_i \leftarrow C_i \cup R_i$
 - 9: $z_l, z_h \leftarrow \max\{0, \lfloor i/\alpha \rfloor - 20\alpha\sqrt{k \log k}\}, \min\{k, \lceil i/\alpha \rceil + 20\alpha\sqrt{k \log k}\}$
 - 10: $e^* \leftarrow \operatorname{argmax}_{e \in C_i} \sum_{\ell=z_l}^{z_h} f(e|L_\ell^{i-1})$
 - 11: **if** $\sum_{\ell=z_l}^{z_h} f(L_\ell^{i-1} \cup \{e^*\}) > \sum_{\ell=z_l}^{z_h} f(L_{\ell+1}^{i-1})$ **then**
 - 12: $H = H \cup \{e^*\}$
 - 13: $L_{\ell+1}^i \leftarrow L_\ell^{i-1} \cup \{e^*\}$ for all $\ell \in [z_l, z_h]$
 - 14: **for** $\ell = 1, 2, \dots, k$ **do**
 - 15: **if** $f(L_\ell^i) \geq f(L_{\ell+1}^i)$ **then**
 - 16: $L_{\ell+1}^i \leftarrow L_\ell^i + \arg \max_{e \in L_{\ell+1}^i} f(e|L_\ell)$
 - 17: **return** $\arg \max_{1 \leq \ell \leq k} f(L_\ell^{\alpha k})$
-

In its most basic implementation, Algorithm 2 requires $O(\alpha k + k^2)$ memory (to store H and L_ℓ^i 's for $\ell = 0, \dots, k$). However, there are several optimizations we can make. Algorithm 2 can be

implemented in a way that the L_ℓ^i 's are not directly stored at all. To avoid storing the L_ℓ^i 's, we can augment H to contain not just e^* , but also the index of the window it was added in. The index of the window tells us the range of levels that e^* was inserted into, so all of the L_ℓ^i 's can be reconstructed from H as H contains a history of all the insertions. Thus the memory use of Algorithm 2 is the size of H at the end of the stream. Since there are αk windows and each window introduces at most 1 element to H , we have the following observation:

Observation 2.1. *Algorithm 2 uses at most $O(\alpha k)$ space and $O(n\alpha\sqrt{k\log k})$ time.*

When E is streamed in random order, our partitioning procedure (Algorithm 1) has a much simpler interpretation. (A similar lemma can be found in the appendix of Agrawal et al. [ASS19].)

Lemma 2.1. *Suppose E is streamed according to a permutation chosen at random and we partition E by Algorithm 1 into m windows. This is equivalent to assigning each $e \in E$ to one of m different buckets uniformly and independently at random.*

The algorithm's performance and behavior depends on the ordering of E . Let us define the *history* of the algorithm up to the i -th window, denoted \mathcal{H}_i , to be the sequence of all solutions produced up to that point. (Note that this history is only used in the analysis.) More precisely, we define the history as follows.

Definition 2.1. *Let H_i denote the state of the set H maintained by the algorithm, before processing window i . We define \mathcal{H}_i to be the set of all triples (e, ℓ, j) such that element $e \in H_i$ was added to solution L_ℓ in window j . In other words, \mathcal{H}_i contains all of the changes that the algorithm made to its state while processing the first i windows. For convenience, sometimes we treat \mathcal{H}_i as a set of elements and say that $e \in \mathcal{H}_i$ if $e \in H_i$.*

The history \mathcal{H}_i describes the entire memory state of the algorithm up to the end of window $i - 1$. In the following, we analyze the performance of the algorithm in the i -th window conditioned the history \mathcal{H}_i . Note that different random permutations of the input may produce this history, and we average over all of them in the analysis.

The next key lemma captures the intuition that elements not selected by the algorithm so far could still appear in the future, and bounds the probability with which this happens.

Lemma 2.2. *Fix a history \mathcal{H}_{i-1} . For any element $e \in E \setminus \mathcal{H}_{i-1}$, and any $i \leq i' \leq m = \alpha k$, we have $\Pr(e \in w_{i'} \mid \mathcal{H}_{i-1}) \geq 1/(\alpha k)$.*

Next, we define a set of *active* windows. In each active window, the algorithm is expected to make significant improvements to its candidate solution. The active windows will only be used in the analysis of the algorithm, and need not be computed in any way.

Definition 2.2. *Let O be the optimal solution. For window i , let p_e^i be the probability that $e \in w_i$ given \mathcal{H}_{i-1} . Define its active set A_i to be the union of R_i and the set obtained by sampling each $e \in w_i$ with probability $1/(\alpha k p_e^i)$. We call w_i an **active window** if $|O \cap A_i| \geq 1$ and we call $O \cap A_i$ the **active optimal elements** of window i .*

Note that the construction of active sets in Definition 2.2 is valid as Lemma 2.2 guarantees $1/(\alpha k p_e^i) \leq 1$. More importantly, the active window A_i subsamples the optimal elements so that each element appears in A_i with probability exactly $1/(\alpha k)$ regardless of the history \mathcal{H}_{i-1} . This allows us to tightly bound the number of active windows in the input, as we show in the next lemma.

Lemma 2.3. *Suppose we have streamed up to the $\alpha\beta$ -th window of the input for some $\beta > 0$. Then expected number of active windows seen so far satisfies*

$$\bar{Z}_{\alpha\beta} := \text{expected number of active windows} = \beta - \Theta(\beta/\alpha).$$

Furthermore, the actual number of windows concentrates around $\bar{Z}_{\alpha\beta}$ to within $\pm O\left(\sqrt{\beta \log \frac{1}{\delta}}\right)$ with probability $1 - \delta$.

Next we analyze the expected gain in the solution after processing each active window. Let \mathcal{A}_{i+1} the event that window w_{i+1} is active.

Lemma 2.4. *Let $\bar{\mathcal{L}}_i = \{z_l^i, z_l^i + 1, \dots, z_h^i\}$ where z_l^i and z_h^i are the values of z_l and z_h defined in Algorithm 2 on window i . Conditioned on a history \mathcal{H}_i and window $i + 1$ being active,*

$$\sum_{\ell \in \bar{\mathcal{L}}_{i+1}} \mathbf{E}[f(L_{\ell+1}^{i+1}) - f(L_\ell^i) \mid \mathcal{H}_i, \mathcal{A}_{i+1}] \geq \frac{1}{k} \sum_{\ell \in \bar{\mathcal{L}}_{i+1}} (f(O) - \mathbf{E}[f(L_\ell^i) \mid \mathcal{H}_i]). \quad (1)$$

Under an ideal partition, each element of O appears in a different window, with one optimal element appearing roughly once every α windows. Thus after k active windows, we expect to obtain a $1 - (1 - 1/k)^k \approx (1 - 1/e)$ -approximation (as in the standard greedy analysis).

Theorem 2.5. *The expected value of the best solution found by the algorithm is at least*

$$\left(1 - \frac{1}{e} - O\left(\frac{1}{\alpha} + \alpha\sqrt{\frac{\log k}{k}}\right)\right) OPT.$$

Setting $\alpha = \Theta(1/\varepsilon)$, we have a $(1 - 1/e - \varepsilon - o(1))$ -approximation using $O(k/\varepsilon)$ memory.

3 A $(1/e - \varepsilon)$ -approximation for non-monotone submodular maximization

In this section, we show that the basic algorithm described in Algorithm 2 can be altered to give a $1/e$ -approximation to the cardinality constrained non-monotone case (Algorithm 3).

Algorithm 3 NONMONOTONESTREAM(f, E, k, α)

```

1: Partition  $E$  into windows  $w_i$  for  $i = 1, \dots, \alpha k$  with Algorithm 1.
2:  $L_\ell^0 \leftarrow \emptyset$  for  $i = 0, \dots, k$ 
3:  $H \leftarrow \emptyset$ 
4:  $x_e^i \leftarrow \text{Unif}(0, 1)$  for  $e \in E, i \in [\alpha k]$ 
5: for  $i = 1, \dots, \alpha k$  do
6:    $C_i \leftarrow \emptyset$ 
7:   for  $e \in \text{window } w_i$  do
8:     for  $j = 1, \dots, i$  do
9:       Reconstruct  $L_\ell^{j-1}$  for all  $\ell$  from  $\mathcal{H}_{i-1}$ 
10:       $A_e^j = \{r \mid 1 \leq r < j, \sum_{\ell=z_r^i}^{z_h^i} f(e|L_\ell^{r-1}) < f_r \text{ or } x_e^r > q_e^r\}$  (see line 13 for  $f_r$ )
11:       $q_e^j = \frac{\alpha k - j + |A_e^j| + 1}{\alpha k}$ 
12:      if  $x_e^i \leq q_e^i$  then  $C_i \leftarrow C_i \cup \{e\}$ 
13:       $f_i \leftarrow \sum_{\ell=z_i^i}^{z_h^i} f(e^*|L_\ell^{i-1})$ 
14:   Perform lines 7 to 16 of Algorithm 2.
15: return  $\arg \max_\ell f(L_\ell^{\alpha k})$ 

```

Algorithm 3 uses the same kind of multi-level scheme as Algorithm 2. However, Algorithm 3 further sub-samples the elements of the input so that the probability of including any element is *exactly* $1/(\alpha k)$ lines 7–13 (coloured in orange). The sub-sampling allows us to bound the maximum probability that an element of the input is included in the solution. In particular, the sub-sampling is done by having the algorithm compute (on the fly) the conditional probability that an element e could have been selected had it appeared in the past. This gives us the ability to compute an appropriate sub-sampling probability to ensure that e does not appear in H with too high a probability. In terms of the proof, the sub-sampling allows us to perform a similar analysis to the RANDOMGREEDY algorithm of Buchbinder et al. [BFNS14].³

Since many of the main ideas are the same, we relegate the details of the analysis to the appendix.

Implementation of Algorithm 3 For clarity of exposition, we compute x_e^i up front in line 4. However, we can compute them on the fly in practice since each element only uses its value of x_e^i once (lines 10 and 12). This avoids an $O(n\alpha k)$ memory cost associated with storing each x_e^i . Finally, we assume that there are no ties when computing the best candidate element in each window. Ties can be handled by any arbitrary but consistent tie-breaking procedure. Any additional information used to break the ties (for example an ordering on the elements e) must be stored alongside f_i for the computation of A_e^j (line 10).

Theorem 3.1. *Algorithm 3 obtains a $(1/e - \varepsilon)$ -approximation for maximizing a non-monotone function f with respect to a cardinality constraint in $O(k/\varepsilon)$ memory.*

³A difference here is that instead of analysing a random element of the top- k marginals, we analyse the optimal set directly.

We remark that Algorithm 3 also achieves a guarantee of $1 - 1/e - \varepsilon$ for the monotone case, as Lemma 2.4 and Theorem 2.5 both still apply to Algorithm 3 when f is monotone. The main difference between the two is the sub-sampling (lines 7–13), which increases the running time of the algorithm.

4 $1 - 1/e$ hardness for monotone submodular maximization

The proofs of the following propositions and lemmas may be found in the appendix.

Proposition 4.1. *Fix subsets G, B of elements (denoting “good” and “bad”) such that $|G| = k$ and $|B| = n - k$; let $r \in [0, k]$ be some parameter. Let m denote the size of the memory buffer, and let p denote the probability that a random subset of size m contains at least $r - 1$ good elements. Let $f : G \cup B \rightarrow \mathbb{R}$ be a function that satisfies the following symmetries:*

- *f is symmetric over good (resp. bad) elements, namely there exists \hat{f} such that $f(S) = \hat{f}(|S \cap G|, |S \cap B|)$.*
- *For any set S with $\leq r - 1$ good elements, f does not distinguish between good and bad elements, namely for $g \leq r - 1$, $\hat{f}(g, b) = \hat{f}(0, b + g)$.*

Then any algorithm has expected value at most

$$ALG \leq (1 - pk)\hat{f}(0, k) + pk \cdot OPT. \quad (2)$$

We now consider a few different f ’s that satisfy the desiderata of Proposition 4.1.

Lemma 4.1 (monotone submodular function [NW78]). *There exists a monotone submodular f that satisfies the desiderata of Proposition 4.1 for $r = 2\varepsilon k$, and such that:*

- $\hat{f}(0, k) = 1 - 1/e + O(\varepsilon)$.
- $OPT = f(G) = 1$.

Lemma 4.2 (polynomial-universe coverage function [MV19]). *There exists a (monotone submodular) coverage function f over a polynomial universe U that satisfies the desiderata of Proposition 4.1 for $r = 2$, and such that:*

- $\hat{f}(0, k) = (1 - 1/e + o(1))|U|$.
- $OPT = f(G) = |U|$.

Lemma 4.3 (exponential-universe coverage function (new construction)). *There exists a (monotone submodular) coverage function f over an exponential universe U that satisfies the desiderata of Proposition 4.1 for $r = 3$, and such that:*

- $\hat{f}(0, k) = (1 - 1/e + o(1))|U|$.
- $OPT = f(G) = |U|$.

Our main hardness result follows from the lemmas above:

Theorem 4.4. *Any $(1 - 1/e + \varepsilon)$ -approximation algorithm in the random order strong oracle model must use the following memory:*

- $\Omega(n)$ for a general monotone submodular function.
- $\Omega(n/k^2)$ for a coverage function over a polynomial universe.
- $\Omega(n/k^{3/2})$ for a coverage function over an exponential universe.

5 Experimental results

In the following section, we give experimental results for our monotone streaming algorithm. Due to space limitations, the experiments for the non-monotone algorithm can be found in the appendix.

Our main goal is to show that our algorithm performs well in a practical setting and is simple to implement. In fact, we show that our algorithm is on par with *offline* algorithms in performance, and returns competitive solutions across a variety of datasets. All experiments were performed on a 2.7 GHz dual-core Intel i7 CPU with 16 GB of RAM.

We compare the approximation ratios obtained by our algorithm with three benchmarks:

- The *offline* LAZYGREEDY algorithm [Min78], which is both theoretically optimal and obtains the same solution as greedy (in faster time). Note that we don't expect to outperform it with a streaming algorithm; but as we hoped, our algorithm comes close.
- The SIEVESTREAMING algorithm of Badanidiyuru et al. [BMKK14], which is the first algorithm to appear for adversarial streaming submodular optimization.
- The SALSA algorithm of Norouzi-Fard et al. [NTM⁺18], which is the first “beyond 1/2” approximation algorithm for random-order streams. This algorithm runs several variants of SIEVESTREAMING in parallel with thresholds that change as the algorithm progresses through the stream.⁴ for *adversarial* order streaming. As we would expect, our algorithm performs better on random arrival streams.

Note that in terms of memory use, our algorithm is strictly more efficient. The analysis in previous sections show that the memory is $O(k/\varepsilon)$ (with a small constant), versus $O(k \log k/\varepsilon)$ for both SIEVESTREAMING and SALSA. Thus in the experiments below, we focus on the approximation ratio obtained by our algorithm.

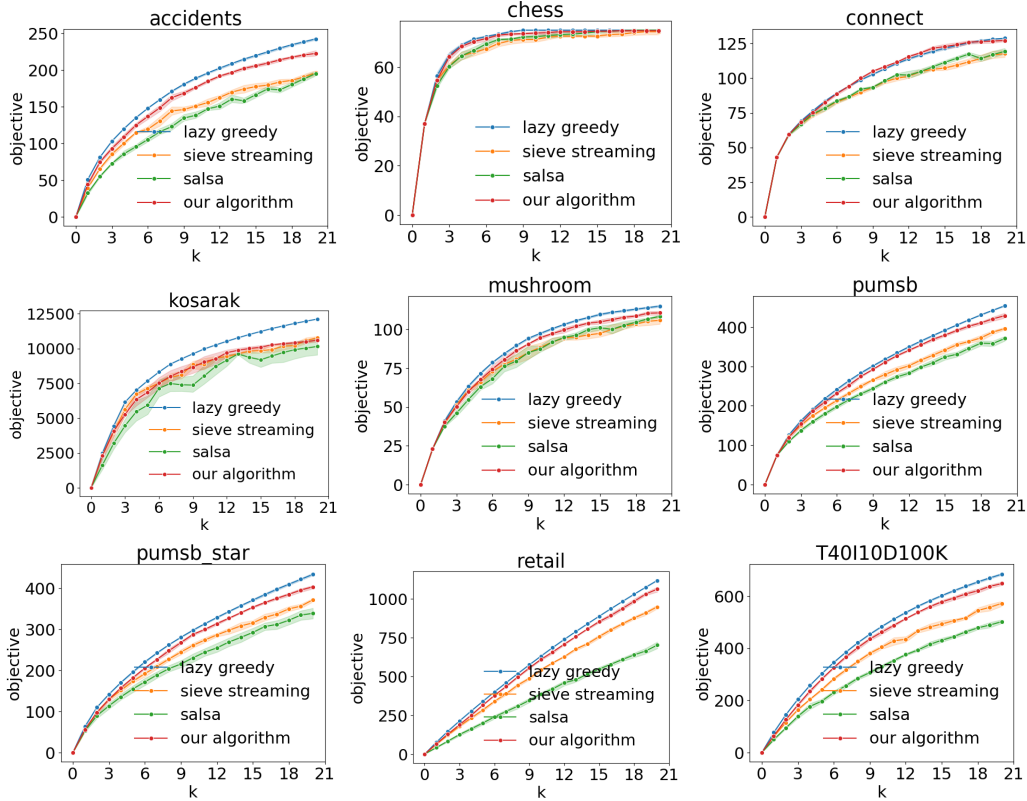


Figure 1: Performance of **lazy greedy**, **sieve streaming**, and **our algorithm** on each data set (averaged across 10 runs, shaded regions represent variance across different random orderings).

⁴SIEVESTREAMING is also known as threshold greedy in the literature [BV14]. Note that the later SIEVESTREAMING++ algorithm of [KMZ⁺19] is more efficient, but for approximation ratio SIEVESTREAMING is a stronger benchmark.

Datasets Our datasets are drawn from set coverage instances from the 2003 and 2004 workshops on Frequent Itemset Mining Implementations [oDM04] and the Steiner triple instances of Beasley [Bea87]. For each data set we run the three algorithms for cardinality constraints varying from $k = 1$ to 20. The results of the algorithms are averaged across 10 random stream orderings. Table 2 describes the data sources. Figure 1 shows the performance of the three algorithms on each data set. All code can be found at <https://github.com/where-is-paul/submodular-streaming> and all datasets can be found at <https://tinyurl.com/neurips-21>.

dataset	source	# of sets	universe size
accidents	(anonymized) traffic accident data	340183	468
chess	UCI ML Repository	3196	75
connect	UCI ML Repository	67557	129
kosarak	(anonymized) click-stream data	990002	41270
mushroom	UCI ML Repository	8124	119
pumsb	census data for population and housing	49046	7116
pumsb_star	census data for population and housing	49046	7116
retail	(anonymized) retail market basket data	88162	16469
T40I10D100K	generator from IBM Quest research	100000	999

Table 2: Description of data sources for monotone experiments.

6 Conclusion and Future Work

In this work, we have presented memory-optimal algorithms for the problem maximizing submodular functions with respect to cardinality constraints in the random order streaming model. Our algorithms achieve an optimal approximation factor of $1 - 1/e$ for the monotone submodular case, and an approximation factor of $1/e$ for the non-monotone case. In addition to theoretical guarantees, we show that the algorithm outperforms existing state-of-the-art on a variety of datasets.

We close with a few open questions that would make for interesting future work. Although our algorithm is memory-optimal, it is not runtime-optimal. In particular, the SIEVESTREAMING [BMKK14] and SALSA [NTM⁺18] algorithms both run in time $O(n \log k/\epsilon)$, whereas our algorithm runs in time $O(n\sqrt{k} \log k/\epsilon)$. The non-monotone variant of our algorithm runs even slower, as it needs to perform sub-sampling operations that take at least $O(\alpha^3 k^{5/2})$ per stream element in its current form. Improving this runtime would greatly improve the practicality of our algorithm for extremely large cardinality constraints. Finally, there has been recent interest in examining streaming algorithms for streams under “adversarial injection” [GKRS20]. In such streams, the optimal elements of the stream are randomly ordered, while adversarial elements can be injected between the optimal elements with no constraints. Despite the seemingly large power of the adversary, the $1/2$ approximation barrier can still be broken in this model. It would be interesting to see if the work in this paper can be extended to such a setting.

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Appendix

A Missing proofs from Section 2

We begin with a simple lemma showing that the values of the levels are *monotone*:

Lemma A.1. *For all i and ℓ , $f(L_\ell^i) \leq f(L_{\ell+1}^{i+1})$ and $f(L_\ell^i) \leq f(L_{\ell+1}^i)$.*

Proof. First, we note that the second part of the lemma holds by lines 15–16. Let z_l^i and z_h^i be the value of z_l and z_h in Algorithm 2 on line 9 on window i . Consider a window $i+1$. There are two cases, depending on whether an element e^* was added to the solutions or not. Suppose no element e^* was added to the solution. Then all the levels remain the same. Line 15 guarantees that $f(L_\ell^i) \leq f(L_{\ell+1}^{i+1})$. Since no elements were added, $f(L_{\ell+1}^i) = f(L_{\ell+1}^{i+1})$ so $f(L_\ell^i) \leq f(L_{\ell+1}^{i+1})$ for every level ℓ . Now suppose an element e^* was added in window $i+1$. For levels $\ell > z_h^{i+1}$ and $\ell < z_l^{i+1}$, $L_{\ell+1}^{i+1} = L_{\ell+1}^i$, so $f(L_{\ell+1}^{i+1}) = f(L_{\ell+1}^i) \geq f(L_\ell^i)$. For levels $\ell \in [z_l^{i+1}, z_h^{i+1}]$, $L_{\ell+1}^{i+1} = L_\ell^i \cup \{e^*\}$ at the end of line 13, and its value only improves through lines 15–16. Thus $f(L_{\ell+1}^{i+1}) \geq f(L_\ell^i \cup \{e^*\}) \geq f(L_\ell^i)$. \square

The rest of the proofs below correspond directly to unproven lemmas in Section 2.

Lemma 2.1. *Suppose E is streamed according to a permutation chosen at random and we partition E by Algorithm 1 into m windows. This is equivalent to assigning each $e \in E$ to one of m different buckets uniformly and independently at random.*

Proof. The way we define the window sizes is equivalent to placing each element independently into a random bucket, and letting n_i be the number of elements in bucket i . Hence the distribution of window sizes is correct. Conditioned on the window sizes, the assignment of elements into windows is determined by a random permutation; any partition is equally likely. Therefore the distribution of elements into windows is equivalent to placing each element into a random window independently. \square

Definition A.1. *Let $P : E \rightarrow [\alpha k]$ be the mapping of E to their window indices; i.e. if e is in the j -th window, then $P(e) = j$. A partition P is \mathcal{H}_i -**compatible** if the algorithm produces history \mathcal{H}_i when streaming the first i windows partitioned by P .*

Lemma 2.2. *Fix a history \mathcal{H}_{i-1} . For any element $e \in E \setminus \mathcal{H}_{i-1}$, and any $i \leq i' \leq m = \alpha k$, we have $\Pr(e \in w'_i \mid \mathcal{H}_{i-1}) \geq 1/(\alpha k)$.*

Proof. Let e be any element of $E \setminus \mathcal{H}_{i-1}$ and choose any $j, j' \in [m]$. For each \mathcal{H}_{i-1} -compatible partition P with $P(e) = j'$, we begin by showing that we can create another \mathcal{H}_{i-1} -compatible partition \tilde{P} by setting $\tilde{P}(e) = j$ and all other values of \tilde{P} equal to P . In other words, any \mathcal{H}_{i-1} -compatible partition where e is in window j can be mapped to another where e is in window j' .

Observe that because $e \notin \mathcal{H}_{i-1}$, e must not have been chosen in windows j or j' by the algorithm so far. There are two possible reasons for this: first, windows j or j' could be greater than (further in the future) or equal to window i , in which case window j, j' trivially does not affect \mathcal{H}_{i-1} . Second, consider the case when windows j or j' is less than i . If this is the case, then element e already arrived in the stream but was not selected by the algorithm in any solution. Hence e was either never the maximum element found in line 10, or if it was, its marginal value was not sufficient to replace the current solution. In either case, removing or adding e to windows j and j' will not change the history \mathcal{H}_{i-1} : if a different element e' was chosen for the update in window j' , this will still be the case; and if no update occurred, this will also still be the case. Finally, observe that the pool of elements H for re-insertion will also not be affected, since element e was not part of it.

Thus, we may change $P(e)$ from j' to j and maintain a \mathcal{H}_{i-1} -compatible partition. Since \tilde{P} is equal to P everywhere except on e , this maps each such partition P to a unique partition \tilde{P} (and vice versa), establishing a bijection between \mathcal{H}_{i-1} -compatible partitions with $e \in w_j$ and $e \in w_{j'}$. Consequently, the number of partitions with $P(e) = j$ compatible with \mathcal{H}_{i-1} is equal to the number of partitions with $P(e) = j'$.

Let J_e be the set of indices j where there exists \mathcal{H}_{i-1} -compatible partitions with $e \in w_j$. The argument above applies to any windows $j \in J_e$. In particular, J_e contains all windows greater than or equal to i , since these windows clearly do not affect \mathcal{H}_{i-1} . For any element $e \notin \mathcal{H}_{i-1}$, and $j, j' \in J_e$, we have

$$\begin{aligned} \Pr(e \in w_j \mid \mathcal{H}_{i-1}) &= \frac{\#\text{partitions } P \text{ with } P(e) = j \text{ and } P \text{ is } \mathcal{H}_{i-1}\text{-compatible}}{\#\text{partitions } P \text{ where } P \text{ is } \mathcal{H}_{i-1}\text{-compatible}} \\ &= \frac{\#\text{partitions } P \text{ with } P(e) = j' \text{ and } P \text{ is } \mathcal{H}_{i-1}\text{-compatible}}{\#\text{partitions } P \text{ where } P \text{ is } \mathcal{H}_{i-1}\text{-compatible}} \\ &= \Pr(e \in w_{j'} \mid \mathcal{H}_{i-1}). \end{aligned}$$

The first and last lines follow from Lemma 2.1, since any partition happens with uniform probability.

Any element $e \notin \mathcal{H}_{i-1}$ must appear in some window, and it is equally likely to be in any of the windows where it could be present without affecting the history \mathcal{H}_{i-1} . So we have

$$1 = \sum_{s=1}^{\alpha k} \Pr(e \in w_s \mid \mathcal{H}_{i-1}) = |J_e| \cdot \Pr(e \in w_i \mid \mathcal{H}_{i-1})$$

The lemma follows from noting that $|J_e| \leq \alpha k$. \square

Proof. By the definition of an active set, $\Pr(\mathbb{1}_{o \in A_i} \mid \mathcal{H}_{i-1}) = 1/(\alpha k)$ for any \mathcal{H}_{i-1} and $o \in O \setminus \mathcal{H}_{i-1}$. For $o \in \mathcal{H}_{i-1} \cap O$, these elements are reintroduced by the algorithm with probability $1/(\alpha k)$. Hence, $\Pr(\mathbb{1}_{o \in A_i}) = 1/(\alpha k)$ for any $o \in O$, without conditioning on \mathcal{H}_{i-1} . Since the input permutation is uniformly random, $\mathbb{1}_{o_a \in A_i}$ and $\mathbb{1}_{o_b \in A_i}$ are independent for $a \neq b$ in any window i . Letting $\mathbb{1}_i = \cup_{o \in O} \mathbb{1}_{o_a \in A_i}$, we have

$$\Pr(\mathbb{1}_i) = 1 - \left(1 - \frac{1}{\alpha k}\right)^k = 1/\alpha - 1/(2\alpha^2) + O(1/\alpha^3)$$

for large enough α and k . Thus,

$$\mathbf{E} \sum_{i=1}^{\alpha\beta} \mathbb{1}_i = \alpha\beta \cdot \mathbf{E} \mathbb{1}_1 = (1 - 1/(2\alpha) + O(1/\alpha^2))\beta.$$

Next, note that $\mathbb{1}_i$ and $\mathbb{1}_j$ are negatively dependent: conditioning on a window being active decreases the number of optimal elements available to the other windows (and conditioning on a window not being active increases the number available to other windows). Thus we have:

$$\begin{aligned} \mathbf{Var} \left(\sum \mathbb{1}_i \right) &\leq \alpha\beta (\mathbf{E} \mathbb{1}_1 - \mathbf{E} \mathbb{1}_1^2) \\ &= (1 - 3/(2\alpha) + O(1/\alpha^2))\beta. \end{aligned}$$

Now can apply the lower-tail bound Hoeffding's inequality ([Doe20], Theorem 1.10.12) to get

$$\begin{aligned} \Pr \left(\sum_{i=1}^{\alpha\beta} \mathbb{1}_i \leq -c\sqrt{\beta \log \frac{1}{\delta}} + \mathbf{E} \sum_{i=1}^{\alpha\beta} \mathbb{1}_i \right) &\leq \exp \left(-\frac{c^2 \log \frac{1}{\delta}}{3} \right) \\ &\leq \delta/2 \end{aligned}$$

for a large enough constant c .

Similarly, we may apply the upper-tail bound of Hoeffding's inequality [Doe20] (Corollary 1.10.13), to obtain:

$$\Pr \left(\sum_{i=1}^{\alpha\beta} \mathbb{1}_i \leq c\sqrt{\beta \log \frac{1}{\delta}} + \mathbf{E} \sum_{i=1}^{\alpha\beta} \mathbb{1}_i \right) \leq \delta/2.$$

for a large enough constant c .

Since $\mathbf{E} \sum_{i=1}^{\alpha\beta} \mathbb{1}_i = \beta - \Theta(\beta/\alpha)$, the number of active windows in the first $\alpha\beta$ windows is at least $\beta - O\left(\sqrt{\beta \log \frac{1}{\delta}}\right) - \Theta(\beta/\alpha)$ and at most $\beta + O\left(\sqrt{\beta \log \frac{1}{\delta}}\right) - \Theta(\beta/\alpha)$ with probability at least $1 - \delta$. \square

Lemma 2.3. Suppose we have streamed up to the $\alpha\beta$ -th window of the input for some $\beta > 0$. Then expected number of active windows seen so far satisfies

$$\bar{Z}_{\alpha\beta} := \text{expected number of active windows} = \beta - \Theta(\beta/\alpha).$$

Furthermore, the actual number of windows concentrates around $\bar{Z}_{\alpha\beta}$ to within $\pm O\left(\sqrt{\beta \log \frac{1}{\delta}}\right)$ with probability $1 - \delta$.

Lemma 2.4. Let $\bar{\mathcal{L}}_i = \{z_l^i, z_l^i + 1, \dots, z_h^i\}$ where z_l^i and z_h^i are the values of z_l and z_h defined in Algorithm 2 on window i . Conditioned on a history \mathcal{H}_i and window $i + 1$ being active,

$$\sum_{\ell \in \bar{\mathcal{L}}_{i+1}} \mathbf{E}[f(L_{\ell+1}^{i+1}) - f(L_\ell^i) \mid \mathcal{H}_i, \mathcal{A}_{i+1}] \geq \frac{1}{k} \sum_{\ell \in \bar{\mathcal{L}}_{i+1}} (f(O) - \mathbf{E}[f(L_\ell^i) \mid \mathcal{H}_i]). \quad (3)$$

Proof. As in the previous lemma, we first note that by the construction of the active set A_{i+1} , any $e \in E$ appears in A_{i+1} with probability exactly $1/(\alpha k)$, so $p := \Pr(o_j \in A_{i+1} \mid \mathcal{H}_i) = 1/(\alpha k)$ for any $o_j \in O$. Also, the appearances of different elements are mutually independent. In particular, we have $\Pr(\mathcal{A}_{i+1} \mid \mathcal{H}_i) = \Pr(\exists o \in O \cap A_{i+1} \mid \mathcal{H}_i) = 1 - (1 - p)^k$.

Order the o_j 's so that $\sum_{\ell \in \bar{\mathcal{L}}_{i+1}} f(o_j \mid L_\ell^i) \geq \sum_{\ell \in \bar{\mathcal{L}}_{i+1}} f(o_{j'} \mid L_\ell^i)$ for $j \leq j'$. Given that $o \in A_{i+1}$ for some random $o \in O$, we have

$$\begin{aligned} \sum_{\ell \in \bar{\mathcal{L}}_{i+1}} \mathbf{E}[f(L_{\ell+1}^{i+1}) - f(L_\ell^i) \mid \mathcal{H}_i, \mathcal{A}_{i+1}] &\geq \max_{e \in A_{i+1}} \sum_{\ell \in \bar{\mathcal{L}}_{i+1}} \mathbf{E}[f(e \mid L_\ell^i) \mid \mathcal{H}_i, \mathcal{A}_{i+1}] \\ &\geq \sum_{j=1}^k \frac{p(1-p)^{j-1}}{1 - (1-p)^k} \sum_{\ell \in \bar{\mathcal{L}}_{i+1}} \mathbf{E}[f(o_j \mid L_\ell^i) \mid \mathcal{H}_i, \mathcal{A}_{i+1}] \\ &\geq \sum_{\ell \in \bar{\mathcal{L}}_{i+1}} \mathbf{E}\left[\frac{1}{k} \sum_{j=1}^k f(o_j \mid L_\ell^i) \mid \mathcal{H}_i, \mathcal{A}_{i+1}\right] \\ &\geq \sum_{\ell \in \bar{\mathcal{L}}_{i+1}} \mathbf{E}\left[\frac{1}{k} (f(O \cup L_\ell^i) - f(L_\ell^i)) \mid \mathcal{H}_i, \mathcal{A}_{i+1}\right] \quad (\text{Submodularity}) \\ &\geq \sum_{\ell \in \bar{\mathcal{L}}_{i+1}} \mathbf{E}\left[\frac{1}{k} (f(O) - f(L_\ell^i)) \mid \mathcal{H}_i, \mathcal{A}_{i+1}\right] \quad (\text{Monotonicity}) \\ &\geq \sum_{\ell \in \bar{\mathcal{L}}_{i+1}} \mathbf{E}\left[\frac{1}{k} (f(O) - f(L_\ell^i)) \mid \mathcal{H}_i\right]. \end{aligned} \quad (4)$$

The first line follows from the fact that $A_{i+1} \subseteq C_{i+1}$, since C_{i+1} contains R_{i+1} and the entirety of w_{i+1} . The numerator of $\frac{p(1-p)^{j-1}}{1 - (1-p)^k}$ follows from the fact that if o_j is the maximum, then no elements in O valued higher than it can appear in A_{i+1} . The denominator is the probability that window w_{i+1} is active. (All events are conditioned on \mathcal{H}_i .) The third line is subtle and follows from Chebyshev's sum inequality. Let $\alpha_j = \frac{p(1-p)^{j-1}}{1 - (1-p)^k}$ and $\beta_j = \sum_{\ell \in \bar{\mathcal{L}}_{i+1}} \mathbf{E}[f(o_j \mid L_\ell^i) \mid \mathcal{H}_i, \mathcal{A}_{i+1}]$. Since α_j and β_j are both decreasing sequences, Chebyshev's sum inequality gives:

$$\sum_{j=1}^k \alpha_j \beta_j \geq \frac{1}{k} \sum_{j=1}^k \alpha_j \sum_{j=1}^k \beta_j = \frac{1}{k} \sum_{j=1}^k \sum_{\ell \in \bar{\mathcal{L}}_{i+1}} \mathbf{E}[f(o_j \mid L_\ell^i) \mid \mathcal{H}_i, \mathcal{A}_{i+1}]. \quad \square$$

Finally, note that \mathcal{H}_i is independent of \mathcal{A}_{i+1} , as $\Pr(\mathcal{A}_{i+1} \mid \mathcal{H}_i)$ is constructed to be the same regardless of \mathcal{H}_i . Hence the conditioning on \mathcal{A}_{i+1} can be removed due to the independence and the fact that L_ℓ^i only depends on \mathcal{H}_i and not \mathcal{A}_{i+1} .

Theorem 2.5. The expected value of the best solution found by the algorithm is at least

$$\left(1 - \frac{1}{e} - O\left(\frac{1}{\alpha} + \alpha \sqrt{\frac{\log k}{k}}\right)\right) OPT.$$

Setting $\alpha = \Theta(1/\varepsilon)$, we have a $(1 - 1/e - \varepsilon - o(1))$ -approximation using $O(k/\varepsilon)$ memory.

Proof. Let $Z_i = \sum_{j=1}^i \mathbb{1}_{\mathcal{A}_j}$ be a random variable which indicates the number of active windows up to window i . Recall from Lemma 2.3 that $\bar{Z}_i := \mathbf{E}[Z_i] = i/\alpha - \Theta(i/\alpha^2)$, and furthermore Z_i concentrates around its expectation: For $\sigma := 20\sqrt{k \log k}$, we have that w.h.p. $(1 - \text{poly}(1/k))$,

$$\forall i \quad Z_i \in [\bar{Z}_i \pm \sigma]. \quad (5)$$

By the definition of O , $f(L_\ell^i) \leq f(O)$ for any ℓ and i since $|L_\ell^i| \leq k$. Since (5) fails only with probability $\text{poly}(1/k)$, we can condition on it while only having negligible effect of $\pm \text{poly}(1/k)f(O)$ on $\mathbf{E}[f(L_\ell^i)]$ for any i, ℓ . We will henceforth simply assume that it always holds; the (in)equalities for the rest of the proof of this lemma hold up to this $\pm \text{poly}(1/k)f(O)$ term.

Our goal is to use Lemma 2.4 to argue about the progress of $\mathbf{E}[f(L_\ell^i)]$ for $\ell = Z_i$. However, we don't know how to prove an analogue of Equation (1) that also conditions on $\ell = Z_i$. We circumvent this issue by tracking the average of $\mathbf{E}[f(L_\ell^i)]$ over ℓ in an interval \mathcal{L}_i around Z_i .

Let $\bar{\mathcal{L}}_i := [\bar{Z}_i \pm \alpha\sigma]$ and $\mathcal{L}_i := [Z_i \pm (\alpha\sigma + \sigma)]$. To keep both nonnegative, we denote $i' := i - \alpha(\alpha\sigma + \sigma) - \Theta(1)$ and consider the contribution from i s.t. $i' \geq 0$. By Equation (5), the interval \mathcal{L}_i has a large overlap with an interval $\bar{\mathcal{L}}_i$ around \bar{Z}_i . Since $\bar{\mathcal{L}}_i$ does not depend on Z_i (or the history in general), we can safely apply Lemma 2.4 for all $\ell \in \bar{\mathcal{L}}_i$.

We will argue that on average over $\ell \in \mathcal{L}_i$ and the randomness of the stream, $\mathbf{E}[f(L_\ell^i)]$ increases like

$$\frac{1}{|\mathcal{L}_i|} \sum_{\ell \in \mathcal{L}_i} \mathbf{E}[f(L_\ell^i)] \geq \left(1 - \left(1 - \frac{1}{k}(1 - e^{-1/\alpha})\right)^{i'} - O\left(\frac{i}{\alpha^2 k}\right)\right) f(O). \quad (6)$$

In particular, for $\tau := \alpha(k - \alpha\sigma - 2\sigma)$, we have that

$$\frac{1}{|\mathcal{L}_\tau|} \sum_{\ell \in \mathcal{L}_\tau} \mathbf{E}[f(L_\ell^\tau)] \geq \left(1 - 1/e - O\left(\frac{1}{\alpha} + \frac{\alpha\sigma}{k}\right)\right) f(O).$$

Assuming Equation (5), the maximum level in \mathcal{L}_τ is at most $Z_\tau + \alpha\sigma + \sigma \leq k$ w.h.p. The values of the levels are monotonically increasing due to line algorithm 2 and therefore level k satisfies

$$\mathbf{E}[f(L_k^\tau)] \geq \left(1 - 1/e - O\left(\frac{1}{\alpha} + \frac{\alpha\sigma}{k}\right)\right) f(O).$$

We now prove Equation (6). For i s.t. $i' \geq 0$, we have:

$$\begin{aligned} \frac{1}{|\mathcal{L}_{i+1}|} \sum_{\ell \in \mathcal{L}_{i+1}} \mathbf{E}[f(L_\ell^{i+1}) \mid \mathcal{H}_i, \mathcal{A}_{i+1}] &= \frac{1}{|\mathcal{L}_i|} \sum_{\ell \in \mathcal{L}_i} \mathbf{E}[f(L_{\ell+1}^{i+1}) \mid \mathcal{H}_i, \mathcal{A}_{i+1}] && \text{(Def. of } \mathcal{L}_i) \\ &= \frac{1}{|\mathcal{L}_i|} \sum_{\ell \in \mathcal{L}_i \setminus \bar{\mathcal{L}}_{i+1}} \mathbf{E}[f(L_{\ell+1}^{i+1}) \mid \mathcal{H}_i, \mathcal{A}_{i+1}] + \frac{1}{|\mathcal{L}_i|} \sum_{\ell' \in \bar{\mathcal{L}}_{i+1}} \mathbf{E}[f(L_{\ell'+1}^{i+1}) \mid \mathcal{H}_i, \mathcal{A}_{i+1}] && \text{(Equation (5))} \\ &\geq \frac{1}{|\mathcal{L}_i|} \sum_{\ell \in \mathcal{L}_i \setminus \bar{\mathcal{L}}_{i+1}} \mathbf{E}[f(L_\ell^i) \mid \mathcal{H}_i, \mathcal{A}_{i+1}] + \frac{1}{|\mathcal{L}_i|} \sum_{\ell' \in \bar{\mathcal{L}}_{i+1}} \mathbf{E}[f(L_{\ell'+1}^{i+1}) \mid \mathcal{H}_i, \mathcal{A}_{i+1}] && \text{(Lemma A.1)} \\ &\geq \frac{1}{|\mathcal{L}_i|} \sum_{\ell \in \mathcal{L}_i} \mathbf{E}[f(L_\ell^i) \mid \mathcal{H}_i, \mathcal{A}_{i+1}] + \frac{1}{|\mathcal{L}_i|} \sum_{\ell' \in \bar{\mathcal{L}}_{i+1}} \frac{1}{k} \mathbf{E}[f(O) - f(L_{\ell'}^i) \mid \mathcal{H}_i] && \text{(Lemma 2.4)} \\ &= \frac{1}{|\mathcal{L}_i|} \sum_{\ell \in \mathcal{L}_i} \mathbf{E}[f(L_\ell^i) \mid \mathcal{H}_i] + \frac{1}{|\mathcal{L}_i|} \sum_{\ell' \in \bar{\mathcal{L}}_{i+1}} \frac{1}{k} \mathbf{E}[f(O) - f(L_{\ell'}^i) \mid \mathcal{H}_i] && (\mathcal{A}_{i+1} \text{ indep. } L_\ell^i \forall \ell) \\ &\geq \frac{1}{|\mathcal{L}_i|} \sum_{\ell \in \mathcal{L}_i} \left(\frac{1}{k} f(O) + \frac{k-1}{k} \mathbf{E}[f(L_\ell^i) \mid \mathcal{H}_i] \right) - O\left(\frac{\sigma}{k|\mathcal{L}_i|}\right) f(O) && (|\mathcal{L}_i| - |\bar{\mathcal{L}}_{i+1}| = 2\sigma) \\ &= \frac{1}{k} f(O) + \frac{k-1}{k} \left(\frac{1}{|\mathcal{L}_i|} \sum_{\ell \in \mathcal{L}_i} \mathbf{E}[f(L_\ell^i) \mid \mathcal{H}_i] \right) - O\left(\frac{\sigma}{k|\mathcal{L}_i|}\right) f(O). \end{aligned}$$

The most subtle line of the argument is the one invoking the independence of L_ℓ^i and \mathcal{A}_{i+1} . This independence is true when conditioned on \mathcal{H}_i : \mathcal{A}_{i+1} and \mathcal{H}_i are independent⁵ and L_ℓ^i is determined by \mathcal{H}_i .

⁵Recall that active sets are defined so that $\Pr(\mathcal{A}_{i+1} \mid \mathcal{H}_i) = 1 - (1 - \frac{1}{\alpha k})^k$ for any \mathcal{H}_i .

In order to avoid cluttering the notation below, every expectation below is conditioned on \mathcal{H}_i . To address the conditioning on \mathcal{A}_{i+1} , consider

$$\begin{aligned}
\frac{1}{|\mathcal{L}_{i+1}|} \sum_{\ell \in \mathcal{L}_{i+1}} \mathbb{E}[f(L_\ell^{i+1})] &= \frac{1 - \Pr[\mathcal{A}_{i+1}]}{|\mathcal{L}_{i+1}|} \sum_{\ell \in \mathcal{L}_{i+1}} \mathbb{E}[f(L_\ell^{i+1}) \mid \neg \mathcal{A}_{i+1}] + \frac{\Pr[\mathcal{A}_{i+1}]}{|\mathcal{L}_{i+1}|} \sum_{\ell \in \mathcal{L}_{i+1}} \mathbb{E}[f(L_\ell^{i+1}) \mid \mathcal{A}_{i+1}] \\
&\geq \left(1 - \frac{\Pr[\mathcal{A}_{i+1}]}{k}\right) \frac{1}{|\mathcal{L}_i|} \sum_{\ell \in \mathcal{L}_i} \mathbb{E}[f(L_\ell^i)] + \frac{1}{k} \Pr[\mathcal{A}_{i+1}] f(O) - O\left(\frac{\sigma}{k|\mathcal{L}_i|}\right) \Pr[\mathcal{A}_{i+1}] f(O) \\
&\geq \left(1 - \frac{1}{\alpha k}\right) \frac{1}{|\mathcal{L}_i|} \sum_{\ell \in \mathcal{L}_i} \mathbb{E}[f(L_\ell^i)] + \frac{1}{\alpha k} f(O) - O\left(\frac{\sigma}{\alpha k |\mathcal{L}_i|}\right) f(O).
\end{aligned}$$

where the last line comes from the inequality $1 - e^{-1/\alpha} \leq \Pr[\mathcal{A}_{i+1}] = 1 - (1 - \frac{1}{\alpha k})^k \leq \frac{1}{\alpha}$ and $\sum_{\ell \in \mathcal{L}_{i+1}} \mathbb{E}[f(L_\ell^{i+1}) \mid \neg \mathcal{A}_{i+1}] \geq \sum_{\ell \in \mathcal{L}_{i+1}} \mathbb{E}[f(L_\ell^i) \mid \neg \mathcal{A}_{i+1}]$ due to line 11 and the fact that $\tilde{\mathcal{L}}_{i+1} \subseteq \mathcal{L}_i$ with high probability. Finally, given that $|\mathcal{L}_i| = 2(\alpha + 1)\sigma$, the last error term is $O(\frac{1}{\alpha^2 k}) f(O)$. Now we may remove the implicit conditioning on \mathcal{H}_i (as we've conditioned everything on it so far). Equation (6) then follows by induction on i . \square

Remark A.1. For any setting of ε , the approximation factor is at best $1 - \frac{1}{e} - \left(\frac{\log k}{k}\right)^{1/4}$, so we might as well choose $\varepsilon \geq \left(\frac{\log k}{k}\right)^{1/4}$. Agrawal et al. [ASS19] have a similar issue; the formal guarantee shown by Agrawal et al. is an approximation factor of $1 - 1/e - \varepsilon - \log(1/\varepsilon)/(\varepsilon^4 k)$. This implies that in Agrawal et al., the approximation is never better than $1 - 1/e - \tilde{O}(1/k^{1/5})$.

B Missing proofs from Section 4

Proposition 4.1. Fix subsets G, B of elements (denoting “good” and “bad”) such that $|G| = k$ and $|B| = n - k$; let $r \in [0, k]$ be some parameter. Let m denote the size of the memory buffer, and let p denote the probability that a random subset of size m contains at least $r - 1$ good elements. Let $f : G \cup B \rightarrow \mathbb{R}$ be a function that satisfies the following symmetries:

- f is symmetric over good (resp. bad) elements, namely there exists \hat{f} such that

$$f(S) = \hat{f}(|S \cap G|, |S \cap B|).$$

- For any set S with $\leq r - 1$ good elements, f does not distinguish between good and bad elements, namely for $g \leq r - 1$,

$$\hat{f}(g, b) = \hat{f}(0, b + g).$$

Then any algorithm has expected value at most

$$ALG \leq (1 - pk) \hat{f}(0, k) + pk \cdot OPT. \quad (7)$$

Lemma 4.3 (exponential-universe coverage function (**new construction**)). *There exists a (monotone submodular) coverage function f over an exponential universe U that satisfies the desiderata of Proposition 4.1 for $r = 3$, and such that:*

- $\hat{f}(0, k) = (1 - 1/e + o(1))|U|$.
- $OPT = f(G) = |U|$.

Proof. The universe of elements to be covered U is the n -dimensional k -side-length hypercube $[k]^n$. Let B_i denote the set $\{\mathbf{x} \in [k]^n \mid x_i = 1\}$, namely, the set of all the vectors of which the i -th coordinate is 1. Let G_i denote the set $\{\mathbf{x} \in [k]^n \mid x_1 = i, x_2 \neq k\} \cup \{\mathbf{x} \in [k]^n \mid x_1 = k, x_2 = k\}$. Our set system consists of k good sets G_1, \dots, G_k and $n - k$ bad sets B_{k+1}, \dots, B_n . We make the following three observations:

- For any $b \in [n - k]$ and $g \in \{0, 1, 2\}$, any distinct $i_1, \dots, i_b \in \{k + 1, \dots, n\}$ and $j_1, j_2 \in [k]$, we have that

$$|(\cup_{t=1}^b B_{i_t}) \cup (\cup_{t=1}^g G_{j_t})| = \left(1 - \left(1 - \frac{1}{k}\right)^{b+g}\right) k^n.$$

- Moreover, it holds that

$$|\cup_{j=1}^k G_j| = \left(1 - \frac{1}{k} + \frac{1}{k^2}\right) k^n.$$

- Finally, the output of the coverage function is fully determined by the number of good sets and the number of bad sets in the input. Hence, there is a succinct encoding of all the possible values of this coverage function, which uses $O(\log n)$ bits. \square

Theorem 4.4. Any $(1 - 1/e + \varepsilon)$ -approximation algorithm in the random order strong oracle model must use the following memory:

- $\Omega(n)$ for a general monotone submodular function.
- $\Omega(n/k^2)$ for a coverage function over a polynomial universe.
- $\Omega(n/k^{3/2})$ for a coverage function over an exponential universe.

Proof. Each case follows by combining Proposition 4.1 with Lemmata 4.1-4.3 respectively. For each case, we need to compute a bound on m such that the probability p of observing $r - 1$ good elements in a random sample of m is $p \leq \varepsilon/k$.

Case $r = 2\varepsilon k$: For $m = \varepsilon n$, the expected number of good elements is $km/n = \varepsilon k$. By Chernoff bound, probability of deviating by $\approx \varepsilon k$ is exponentially small.

Case $r = 2$: For $m = \varepsilon n/k^2$, the expected number of good elements is $km/n = \varepsilon/k$. By Markov's inequality, the probability of having at least $r - 1 = 1$ good element in memory is $p < \varepsilon/k$.

Case $r = 3$: For $m = \sqrt{\varepsilon}n/k^{3/2}$, each good element appears in memory with probability $\sqrt{\varepsilon}k^{-3/2}$. The probability that any fixed pair of good elements appear in memory is $\leq \varepsilon k^{-3}$. Taking a union bound over $\binom{k}{2}$ pairs, we have that $p < \varepsilon/k$. \square

C A $(1/e - \varepsilon)$ -approximation for non-monotone submodular maximization

In this section, we show that the basic algorithm described in Algorithm 2 can be altered to give a $1/e$ -approximation to the cardinality constrained non-monotone case (Algorithm 4).

Algorithm 4 uses the same kind of multi-level scheme as Algorithm 2 but differs in two ways.

First, Algorithm 4 further sub-samples the elements of the input so that the probability of including any element is *exactly* $1/(\alpha k)$ lines 8–13 (coloured in orange). The sub-sampling allows us to bound the maximum probability that an element of the input is included in the solution. In particular, the sub-sampling is done by having the algorithm compute (on the fly) the conditional probability that an element e could have been selected had it appeared in the past. This gives us the ability to compute an appropriate sub-sampling probability to ensure that e does not appear in H with too high a probability. In terms of the proof, the sub-sampling allows us to perform a similar analysis to the RANDOMGREEDY algorithm of Buchbinder et al. [BFNS14].⁶

Second, the addition of elements to levels in $[z_l, z_h]$ may cause a decrease in the function value, meaning that we no longer maintain the nesting property in Lemma A.1 (for similar reasons, line 22 also differs from the monotone case by simply copying level ℓ into $\ell + 1$). Fortunately, we only require the nesting property to hold on levels outside of $[z_l, z_h]$. We show that this remains true for Algorithm 4.

⁶A difference here is that instead of analysing a random element of the top- k marginals, we analyse the optimal set directly.

Algorithm 4 NONMONOTONESTREAM(f, E, k, α)

```
1: Partition  $E$  into windows  $w_i$  for  $i = 1, \dots, \alpha k$  with Algorithm 1.
2:  $L_\ell^0 \leftarrow \emptyset$  for  $i = 0, \dots, k$ 
3:  $H \leftarrow \emptyset$ 
4:  $x_e^i \leftarrow \text{Unif}(0, 1)$  for  $e \in E, i \in [\alpha k]$ 
5: for  $i = 1, \dots, \alpha k$  do
6:    $C_i \leftarrow \emptyset$ 
7:    $z_l^i, z_h^i \leftarrow \max\{0, \lfloor i/\alpha \rfloor - 20\alpha\sqrt{k \log k}\}, \min\{k, \lceil i/\alpha \rceil + 20\alpha\sqrt{k \log k}\}$ 
8:   for  $e \in \text{window } w_i$  do
9:     for  $j = 1, \dots, i$  do
10:      Reconstruct  $L_\ell^{j-1}$  for all  $\ell$  from  $\mathcal{H}_{i-1}$ 
11:       $A_e^j = \{r \mid 1 \leq r < j, \sum_{\ell=z_l^r}^{z_h^r} f(e|L_\ell^{r-1}) < f_r \text{ or } x_e^r > q_e^r\}$  (see line 16 for  $f_r$ )
12:       $q_e^j = \frac{\alpha k - j + |A_e^j| + 1}{\alpha k}$ 
13:      if  $x_e^i \leq q_e^i$  then  $C_i \leftarrow C_i \cup \{e\}$ 
14:      Sample each  $e \in H$  with probability  $\frac{1}{\alpha k}$  and add to  $C_i$ 
15:       $e^* \leftarrow \text{argmax}_{e \in C_i} \sum_{\ell=z_l^i}^{z_h^i} f(e|L_\ell^{i-1})$ 
16:       $f_i \leftarrow \sum_{\ell=z_l^i}^{z_h^i} f(e^*|L_\ell^{i-1})$ 
17:      if  $\sum_{\ell=z_l^i}^{z_h^i} f(L_\ell^{i-1} \cup \{e^*\}) > \sum_{\ell=z_l^i}^{z_h^i} f(L_\ell^{i-1})$  then
18:         $H = H \cup \{e^*\}$ 
19:         $L_{\ell+1}^i \leftarrow L_\ell^{i-1} \cup \{e^*\}$  for all  $\ell \in [z_l, z_h]$ 
20:        for  $\ell = 1, 2, \dots, k$  do
21:          if  $f(L_\ell^i) \geq f(L_{\ell+1}^i)$  then
22:             $L_{\ell+1}^i \leftarrow L_\ell^i$ 
23: return  $\arg \max_\ell f(L_\ell^{\alpha k})$ 
```

Lemma C.1. Let z_l^i and z_h^i be the value of z_l and z_h in Algorithm 4 on window i . For all i and $\ell \notin [z_l^{i+1}, z_h^{i+1}]$, $f(L_\ell^i) \leq f(L_{\ell+1}^{i+1})$.

Proof. Consider a window $i+1$. Regardless of whether an element e^* was added to the solution or not, levels $\ell > z_h^{i+1}$ and $\ell < z_l^{i+1}$ are not changed. Thus $L_{\ell+1}^{i+1} = L_{\ell+1}^i$, so $f(L_{\ell+1}^{i+1}) = f(L_{\ell+1}^i) \geq f(L_\ell^i)$ by line 21. Thus $f(L_{\ell+1}^{i+1}) \geq f(L_\ell^i)$ for $\ell \notin [z_l^{i+1}, z_h^{i+1}]$. \square

Implementation of Algorithm 4 For clarity of exposition, we compute x_e^i up front in line 4. However, we can compute them on the fly in practice since each element only uses its value of x_e^i once (lines 11 and 13). This avoids an $O(n\alpha k)$ memory cost associated with storing each x_e^i . Finally, we assume that there are no ties when computing the best candidate element in each window (line 15). Ties can be handled by any arbitrary but consistent tie-breaking procedure. Any additional information used to break the ties (for example an ordering on the elements e) must be stored alongside f_i for the computation of A_e^j (line 11).

Next, we show that the probability an element e is in a candidate set C_i is exactly $1/(\alpha k)$ for any $e \in E$. The proof is conceptually very similar to Lemma C.2, in which we showed that $p_e^i := \Pr(e \in w_i \mid \mathcal{H}_{i-1}) \geq 1/(\alpha k)$. However, we make the additional observation that the proof of Lemma C.2 also offers a way for the algorithm to compute $\Pr(e \in w_i \mid \mathcal{H}_{i-1})$ exactly. By computing this probability and sub-sampling e with probability $1/(\alpha k p_e^i)$, we ensure that e is included in C_i (line 13) with probability exactly $1/(\alpha k)$.

Lemma C.2. Fix a history \mathcal{H}_{i-1} . For any element $e \in E$, we have $\Pr(e \in C_i \mid \mathcal{H}_{i-1}) = 1/(\alpha k)$.

Proof. When $e \in \mathcal{H}_{i-1}$, e is added to C_i with probability $1/(\alpha k)$ exactly. Thus we assume $e \in E \setminus \mathcal{H}_{i-1}$.

Let $\mathcal{T} = A_e^i \cup \{i, i+1, \dots, \alpha k\}$ (where A_e^i is defined on line 11). Fix any \mathcal{H}_{i-1} -compatible partition P and $j \in \mathcal{T}$. As in Lemma C.2, we begin by showing that we can create another \mathcal{H}_{i-1} -compatible

partition \tilde{P} by setting $\tilde{P}(e) = j$ and all other values of \tilde{P} equal to P . Since \tilde{P} is equal to P everywhere except on e , this maps each such partition P to a unique partition \tilde{P} . Consequently, the map from P to \tilde{P} is a bijection, and so the number of \mathcal{H}_{i-1} -compatible partitions with $P(e) = j$ is equal for any $j \in \mathcal{T}$.

Observe that because $e \notin \mathcal{H}_{i-1}$, e can be removed from window $P(e)$ without changing \mathcal{H}_{i-1} . We now separate the argument into two cases, for $j \geq i$ and $j \in A_e^i$. If $j \geq i$, the mapping of P to \tilde{P} does not change \mathcal{H}_{i-1} , since window j is not included in the computations determining \mathcal{H}_{i-1} . Thus \tilde{P} is trivially \mathcal{H}_{i-1} -compatible. If $j \in A_e^i$, this means that either $\sum_{\ell=z_l^j}^{z_h^j} f(e|L_\ell^{j-1})$ was too small, or it was probabilistically ignored because x_e^j was too big. Either way, this means that $e \notin C_j$ and hence $e \notin L_\ell^j$ for any ℓ . Consequently, adding e to window j does not change \mathcal{H}_{i-1} . As a result, \tilde{P} is \mathcal{H}_{i-1} -compatible for any $j \in \mathcal{T}$.

For any windows $j, j' \in \mathcal{T}$, we then have

$$\begin{aligned} \Pr(e \in w_j | \mathcal{H}_{i-1}, e \notin \mathcal{H}_{i-1}) &= \frac{\#\text{partitions } P \text{ with } P(e) = j \text{ and } P \text{ is } \mathcal{H}_{i-1}\text{-compatible}}{\#\text{partitions } P \text{ is } \mathcal{H}_{i-1}\text{-compatible}} \\ &= \frac{\#\text{partitions } P \text{ with } P(e) = j' \text{ and } P \text{ is } \mathcal{H}_{i-1}\text{-compatible}}{\#\text{partitions } P \text{ is } \mathcal{H}_{i-1}\text{-compatible}} \\ &= \Pr(e \in w_{j'} | \mathcal{H}_{i-1}, e \notin \mathcal{H}_{i-1}). \end{aligned}$$

We now have the ingredients to compute $\Pr(e \in C_j | \mathcal{H}_{i-1}) = 1/(\alpha k)$. Any element $e \notin \mathcal{H}_{i-1}$ must appear in some window, so we have

$$\begin{aligned} 1 &= \sum_{s=1}^{\alpha k} \Pr(e \in w_s | \mathcal{H}_{i-1}, e \notin \mathcal{H}_{i-1}) \\ &= \sum_{s \in \mathcal{T}} \Pr(e \in w_s | \mathcal{H}_{i-1}, e \notin \mathcal{H}_{i-1}) \\ &= |\mathcal{T}| \Pr(e \in w_i | \mathcal{H}_{i-1}, e \notin \mathcal{H}_{i-1}) \end{aligned}$$

Since $\Pr(e \in C_i | \mathcal{H}_{i-1}) = \Pr(e \in w_i | \mathcal{H}_{i-1}, e \notin \mathcal{H}_{i-1}) q_e^i$, we have $\Pr(e \in C_i | \mathcal{H}_{i-1}) = \frac{1}{|\mathcal{T}|} \cdot |\mathcal{T}|/(\alpha k) = 1/(\alpha k)$. \square

We are now ready to show the approximation guarantees of Algorithm 4. To do this, we borrow the following lemma from Buchbinder et al. [BFNS14]:

Lemma C.3 (Lemma 2.2 [BFNS14]). *Let $f : 2^E \rightarrow \mathbb{R}_+$ be a submodular function. Further, let R be a random subset of $T \subseteq E$ in which every element occurs with probability at most p (not necessarily independently). Then, $\mathbf{E}f(R) \geq (1-p)f(\emptyset)$.*

First, we note that the analysis of Lemma 2.4 up to the application of submodularity in Equation (4) still applies, leading to the following observation:

Observation C.1. *Let $\tilde{\mathcal{L}}_i = \{z_l, z_l + 1, \dots, z_h\}$ where z_l and z_h are defined in Algorithm 2. Conditioned on a history \mathcal{H}_i and window $i+1$ being active (the event \mathcal{A}_{i+1}),*

$$\sum_{\ell \in \tilde{\mathcal{L}}_{i+1}} \mathbf{E}[f(L_{\ell+1}^{i+1}) - f(L_\ell^i) | \mathcal{H}_i, \mathcal{A}_{i+1}] \geq \frac{1}{k} \sum_{\ell \in \tilde{\mathcal{L}}_{i+1}} \mathbf{E}[f(O \cup L_\ell^i) - f(L_\ell^i) | \mathcal{H}_i].$$

Now we relate the value of $f(O \cup L_\ell^i)$ to $f(O)$. As in Buchbinder et al. [BFNS14], this will involve showing that no element of the ground set is included into any of the levels L_ℓ^i with too high of a probability.

Lemma C.4. *For every i and every ℓ , the $\mathbf{E}[f(O \cup L_\ell^i)] \geq (1 - 1/(\alpha k))^i f(O)$ for all $\ell \leq k$.*

Proof. To be inserted into a partial solution L_ℓ^i , an element e must appear as part of the candidate set C_i in Algorithm 4 or have appeared in \mathcal{H}_{i-1} . As shown in Lemma C.2, when conditioned on e not

appearing in \mathcal{H}_{i-1} , e appears in C_i with probability exactly $1/(\alpha k)$. Thus $\Pr(e \in \mathcal{H}_i \mid e \notin \mathcal{H}_{i-1}) \leq 1/(\alpha k)$. By induction, we have

$$\begin{aligned} \Pr(e \notin \mathcal{H}_i) &= \Pr(e \notin \mathcal{H}_i \mid e \notin \mathcal{H}_{i-1}) \Pr(e \notin \mathcal{H}_{i-1}) \\ &\geq \left(1 - \frac{1}{\alpha k}\right) \Pr(e \notin \mathcal{H}_{i-1}) \\ &\geq \left(1 - \frac{1}{\alpha k}\right)^i \end{aligned}$$

where we assume by induction that $\Pr(e \notin \mathcal{H}_{i-1}) \geq \left(1 - \frac{1}{\alpha k}\right)^{i-1}$. After i windows, any particular element of the input is in \mathcal{H}_i (and L_ℓ^i for any ℓ) with probability at most $1 - \left(1 - 1/(\alpha k)\right)^i$.

Define the submodular function $g(S) = f(S \cup O)$. By Lemma C.3 and the reasoning above,

$$\mathbf{E}[g(L_\ell^i)] \geq (1 - 1/(\alpha k))^i f(O). \quad \square$$

Theorem 3.1. *Algorithm 4 obtains a $(1/e - \varepsilon)$ -approximation for maximizing a non-monotone function f with respect to a cardinality constraint.*

Proof. Let $Z_i = \sum_{j=1}^i \mathbb{1}_{\mathcal{A}_i}$ be a random variable measuring the number of active windows up to window i . We follow an analysis similar to Theorem 2.5. For $\sigma := 20\sqrt{k \log k}$, Lemma 2.3 shows that w.h.p. $(1 - \text{poly}(1/k))$,

$$\forall i \quad Z_i \in [\bar{Z}_i \pm \sigma]. \quad (8)$$

By the definition of O , $f(L_\ell^i) \leq f(O)$ for any ℓ and i since $|L_\ell^i| \leq k$. Again, since (8) fails only with probability $\text{poly}(1/k)$, we can condition on it while only having negligible effect of $\pm \text{poly}(1/k)f(O)$ on $\mathbf{E}[f(L_\ell^i)]$ (for any i, ℓ). For the rest of the proof, the inequalities will hold up to a $\pm \text{poly}(1/k)f(O)$ term (which has no effect on the final asymptotic error guarantee).

Let $\bar{\mathcal{L}}_i := [\bar{Z}_i \pm \alpha\sigma]$ and $\mathcal{L}_i := [Z_i \pm (\alpha\sigma + \sigma)]$. To keep both non-negative, we denote $i' := i - \alpha(\alpha\sigma + \sigma) - \Theta(1)$ and consider the contribution from i s.t. $i' \geq 0$.

In the non-monotone case, our goal will be to argue that:

$$\frac{1}{|\mathcal{L}_i|} \sum_{\ell \in \mathcal{L}_i} \mathbf{E}[f(L_\ell^i) \mid \mathcal{A}_{i+1}] \geq \frac{i+1}{\alpha k} \left(1 - \frac{1}{\alpha k}\right)^i f(O) - O\left(\frac{i+1}{\alpha^2 k}\right) f(O). \quad (9)$$

In particular, for $\tau := \alpha(k - \alpha\sigma - 2\sigma)$, we have that

$$\frac{1}{|\mathcal{L}_i|} \sum_{\ell \in \mathcal{L}_\tau} \mathbf{E}[f(L_\ell^\tau)] \geq \left(1/e - O\left(\frac{1}{\alpha} + \frac{\alpha\sigma}{k}\right)\right) f(O).$$

Since the solutions L_ℓ^i increase in value as ℓ increases, we also have $L_k^\tau \geq (1/e - O(\frac{1}{\alpha} + \frac{\alpha\sigma}{k})) f(O)$.

We now prove Equation (9). To avoid cluttering the notation, every expectation in the equation below is conditioned on \mathcal{H}_{i-1} . The reasoning below follows along the same lines as the monotone case. Roughly speaking, \mathcal{L}_i and $\bar{\mathcal{L}}_i$ largely overlap, so their averages are close up to $1/\text{poly}(k)$ factors. However, since \mathcal{L}_i is deterministically defined, this allows us to apply Observation C.1. For i

s.t. $i' \geq 0$, we have:

$$\begin{aligned}
\frac{1}{|\mathcal{L}_{i+1}|} \sum_{\ell \in \mathcal{L}_{i+1}} \mathbf{E}[f(L_\ell^{i+1}) \mid \mathcal{A}_{i+1}] &= \frac{1}{|\mathcal{L}_i|} \sum_{\ell \in \mathcal{L}_i} \mathbf{E}[f(L_\ell^{i+1}) \mid \mathcal{A}_{i+1}] && \text{(Def. of } \mathcal{L}_i) \\
&= \frac{1}{|\mathcal{L}_i|} \left(\sum_{\ell \in \mathcal{L}_i \setminus \bar{\mathcal{L}}_{i+1}} \mathbf{E}[f(L_\ell^{i+1}) \mid \mathcal{A}_{i+1}] + \sum_{\ell' \in \bar{\mathcal{L}}_{i+1}} \mathbf{E}[f(L_{\ell'}^{i+1}) \mid \mathcal{A}_{i+1}] \right) && \text{(Eq. (8))} \\
&= \frac{1}{|\mathcal{L}_i|} \left(\sum_{\ell \in \mathcal{L}_i \setminus \bar{\mathcal{L}}_{i+1}} \mathbf{E}[f(L_\ell^i) \mid \mathcal{A}_{i+1}] + \sum_{\ell' \in \bar{\mathcal{L}}_{i+1}} \mathbf{E}[f(L_{\ell'}^{i+1}) \mid \mathcal{A}_{i+1}] \right) && \text{(Lemma C.1)} \\
&\geq \frac{1}{|\mathcal{L}_i|} \left(\sum_{\ell \in \mathcal{L}_i} \mathbf{E}[f(L_\ell^i) \mid \mathcal{A}_{i+1}] + \sum_{\ell' \in \bar{\mathcal{L}}_{i+1}} \frac{1}{k} \mathbf{E}[(f(O \cup L_{\ell'}^i) - f(L_{\ell'}^i)) \mid \mathcal{A}_{i+1}] \right) && \text{(Observation C.1)} \\
&= \frac{1}{|\mathcal{L}_i|} \left(\sum_{\ell \in \mathcal{L}_i} \mathbf{E}[f(L_\ell^i)] + \sum_{\ell' \in \bar{\mathcal{L}}_{i+1}} \frac{1}{k} \mathbf{E}[(f(O \cup L_{\ell'}^i) - f(L_{\ell'}^i))] \right) && (\mathcal{A}_{i+1} \text{ indep. of } L_\ell^i \ \forall \ell) \\
&\geq \frac{1}{|\mathcal{L}_i|} \left(\sum_{\ell \in \mathcal{L}_i} \mathbf{E}[f(L_\ell^i)] + \sum_{\ell' \in \bar{\mathcal{L}}_{i+1}} \frac{1}{k} \mathbf{E}[(1 - (\alpha k)^{-1})^i f(O) - f(L_{\ell'}^i)] \right) && \text{(Lemma C.4)} \\
&\geq \frac{1}{|\mathcal{L}_i|} \sum_{\ell \in \mathcal{L}_i} \left(\frac{(1 - (\alpha k)^{-1})^i}{k} f(O) + \frac{k-1}{k} \mathbf{E}[f(L_\ell^i)] \right) - O\left(\frac{\sigma}{k|\mathcal{L}_i|}\right) f(O) && (|\mathcal{L}_i| - |\bar{\mathcal{L}}_{i+1}| = 2\sigma) \\
&= \frac{(1 - (\alpha k)^{-1})^i}{k} f(O) + \left(1 - \frac{1}{k}\right) \frac{1}{|\mathcal{L}_i|} \sum_{\ell \in \mathcal{L}_i} \mathbf{E}[f(L_\ell^i)] - O\left(\frac{\sigma}{k|\mathcal{L}_i|}\right) f(O).
\end{aligned}$$

The main difference between the non-monotone and monotone case is the application of Lemma C.4 (to bound the maximum amount an element may hurt the solution).

To address the conditioning on \mathcal{A}_{i+1} , consider

$$\begin{aligned}
\frac{1}{|\mathcal{L}_{i+1}|} \sum_{\ell \in \mathcal{L}_{i+1}} \mathbf{E}[f(L_\ell^{i+1})] &= \frac{1 - \Pr[\mathcal{A}_{i+1}]}{|\mathcal{L}_{i+1}|} \sum_{\ell \in \mathcal{L}_{i+1}} \mathbf{E}[f(L_\ell^{i+1}) \mid \neg \mathcal{A}_{i+1}] + \frac{\Pr[\mathcal{A}_{i+1}]}{|\mathcal{L}_{i+1}|} \sum_{\ell \in \mathcal{L}_{i+1}} \mathbf{E}[f(L_\ell^{i+1}) \mid \mathcal{A}_{i+1}] \\
&\geq \left(1 - \frac{\Pr[\mathcal{A}_{i+1}]}{k}\right) \frac{1}{|\mathcal{L}_i|} \sum_{\ell \in \mathcal{L}_i} \mathbf{E}[f(L_\ell^i)] + \frac{(1 - (\alpha k)^{-1})^i}{k} \Pr[\mathcal{A}_{i+1}] f(O) - O\left(\frac{\sigma}{k|\mathcal{L}_i|}\right) \Pr[\mathcal{A}_{i+1}] f(O) \\
&\geq \left(1 - \frac{1}{\alpha k}\right) \frac{1}{|\mathcal{L}_i|} \sum_{\ell \in \mathcal{L}_i} \mathbf{E}[f(L_\ell^i)] + \frac{(1 - (\alpha k)^{-1})^i}{\alpha k} f(O) - O\left(\frac{1}{\alpha^2 k}\right) f(O).
\end{aligned}$$

The second line requires $\mathbf{E}[f(L_\ell^{i+1}) \mid \neg \mathcal{A}_{i+1}] \geq \mathbf{E}[f(L_\ell^i) \mid \neg \mathcal{A}_{i+1}]$. Fortunately, this is true as Algorithm 4 only increases the average of values in \mathcal{L}_i : levels $\bar{\mathcal{L}}_i \subseteq \mathcal{L}_i$ are only updated if an increase is detected on line 17. The last line follows from plugging in $|\mathcal{L}_i| = 2(\alpha + 1)\sigma$ and $1 - e^{-1/\alpha} \leq \Pr[\mathcal{A}_{i+1}] = 1 - (1 - \frac{1}{\alpha k})^k \leq \frac{1}{\alpha}$. Now we may remove the (implicit) conditioning on \mathcal{H}_{i-1} , as everything has been conditioned on it so far.

Let $\bar{f}_i = \frac{1}{|\mathcal{L}_i|} \sum_{\ell \in \mathcal{L}_i} \mathbf{E}[f(L_\ell^i)]$. Now we show by induction that $\bar{f}_i \geq \frac{i}{\alpha k} (1 - \frac{1}{\alpha k})^{i-1} f(O) - O\left(\frac{i}{\alpha^2 k}\right) f(O)$. The base case is clearly true, as the first window has probability $1/(\alpha k)$ of catching any optimal element. By our analysis from above:

$$\begin{aligned}
\bar{f}_{i+1} &\geq \left(1 - \frac{1}{\alpha k}\right) \bar{f}_i + \frac{(1 - (\alpha k)^{-1})^i}{\alpha k} f(O) - O\left(\frac{1}{\alpha^2 k}\right) f(O) \\
&\geq \left(1 - \frac{1}{\alpha k}\right) \left(\frac{i}{\alpha k} \left(1 - \frac{1}{\alpha k}\right)^{i-1} - O\left(\frac{i}{\alpha^2 k}\right) \right) f(O) + \frac{(1 - \frac{1}{\alpha k})^i}{\alpha k} f(O) - O\left(\frac{1}{\alpha^2 k}\right) f(O) \\
&\geq \frac{i+1}{\alpha k} \left(1 - \frac{1}{\alpha k}\right)^i f(O) - O\left(\frac{i+1}{\alpha^2 k}\right) f(O).
\end{aligned}$$

In particular, for $\tau := \alpha(k - \alpha\sigma - 2\sigma)$, we have that

$$\frac{1}{|\mathcal{L}_i|} \sum_{\ell \in \mathcal{L}_\tau} \mathbb{E}[f(L_\ell^\tau)] \geq \left(1/e - O\left(\frac{1}{\alpha} + \frac{\alpha\sigma}{k}\right)\right) f(O).$$

Setting $\alpha = \Theta(1/\varepsilon)$ where $\varepsilon = \omega\left(\frac{\log k}{k}\right)^{1/4}$ gives the desired result. \square

We remark that Algorithm 4 also achieves a guarantee of $1 - 1/e - \varepsilon$ for the monotone case, as Lemma 2.4 and Theorem 2.5 both still apply to Algorithm 4 when f is monotone. The main difference between the two is the sub-sampling procedure (lines 8–13), which increases the running time of the algorithm.

D Experiments

All code can be found at <https://github.com/where-is-paul/submodular-streaming> and all datasets can be found at <https://tinyurl.com/neurips-21>.

Experiments for non-monotone submodular streaming

For non-monotone submodular functions, we compare against the offline random greedy algorithm of Buchbinder et al. [BFNS14].

Datasets Our datasets are drawn from diversity maximization tasks described in [LSK⁺21]. Here, given an $n \times n$ matrix L , the task is to find a subset S of k indices such that $\log \det(L_S)$ is maximized.⁷ Since our non-monotone algorithm is significantly more expensive to run than our monotone algorithm, we created substreams from these datasets by sampling a consecutive run of 1024 stream elements at random and then permuting them. As in the monotone case, for each data set we run the standard offline algorithm (random greedy) and compare against our streaming algorithm with k varying from 1 to 10. Table 3 describes the data sources. Figure 2 shows the performance of the three algorithms on each data set.

dataset	source
gowalla	gowalla geolocation data
yahoo!	yahoo front page visit data
bing	anonymized search data from the bing search engine

Table 3: Description of data sources for non-monotone experiments.

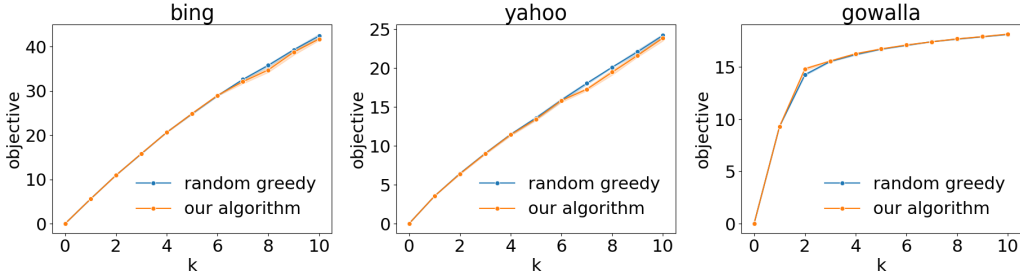


Figure 2: Performance of **standard greedy**, **random greedy**, and **our algorithm** on each data set (averaged across 10 runs, shaded regions represent variance across different random orderings).

⁷ L_S is the $k \times k$ submatrix formed by taking entries from L with rows and columns in S .