

Lecture 14

Time Series Analysis

DSA 8020 Statistical Methods II
April 12-16, 2021

Time Series Data

Features of Times
Series

Means &
Autocovariances

Autoregressive Moving
Average (ARMA)
Models

A Case Study

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Agenda

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- 3 Means & Autocovariances
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- 5 A Case Study

Time Series Data

Features of Times Series

Means & Autocovariances

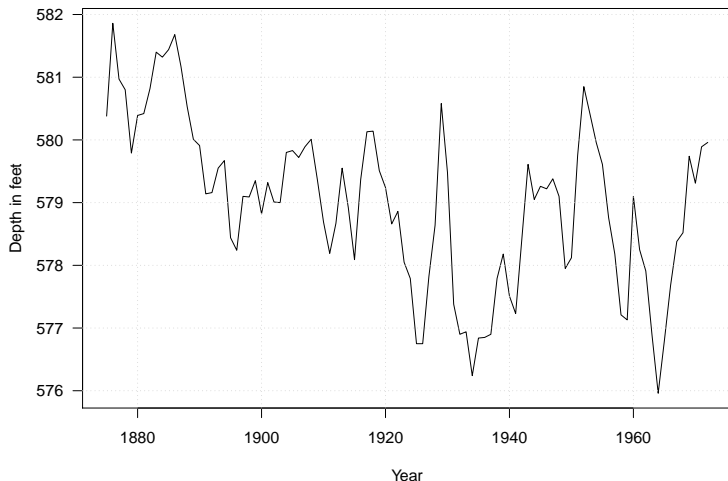
Autoregressive Moving Average (ARMA) Models

A Case Study

Level of Lake Huron 1875–1972

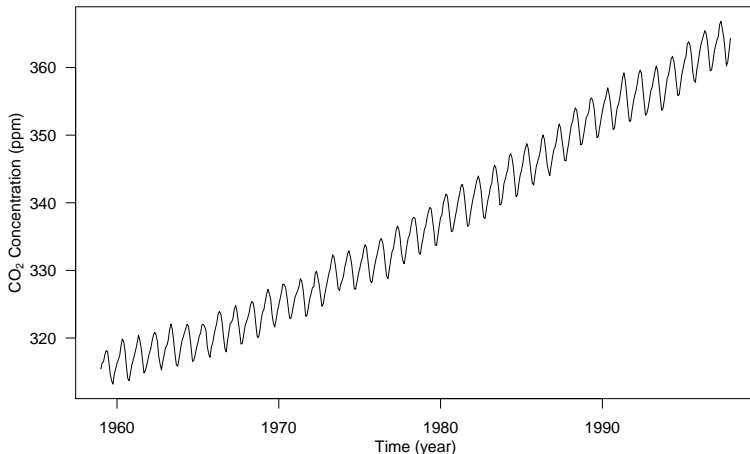
Annual measurements of the level of Lake Huron in feet.

[Source: [Brockwell & Davis, 1991](#)]

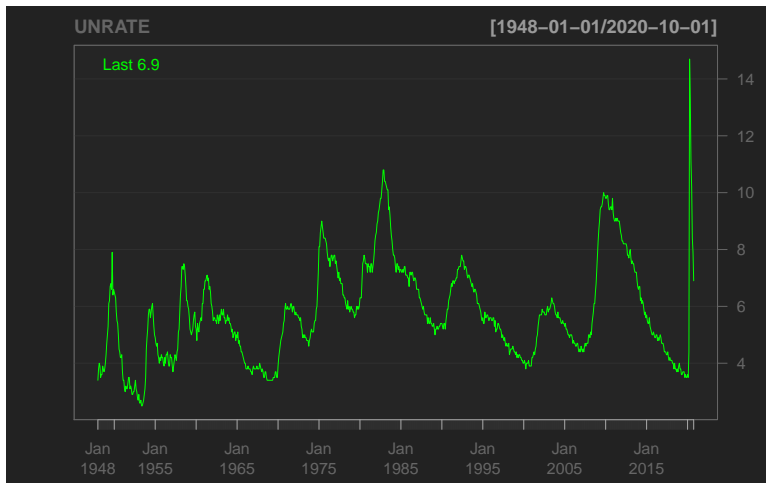


Mauna Loa Atmospheric CO₂ Concentration

Monthly atmospheric concentrations of CO₂ at the Mauna Loa Observatory [Source: [Keeling & Whorf, Scripps Institution of Oceanography \(SIO\)](#)]



US Unemployment Rate 1948 Jan. – 2020 Oct.



- A **time series** is a set of observations made sequentially in time
- **Time series analysis** is the area of statistics which deals with the analysis of **dependency** between different observations in time series data
- A **time series model** is a probabilistic model that describes ways that the series data $\{y_t\}$ could have been generated
- More specifically, a time series model is usually a probability model for $\{Y_t : t \in T\}$, **a collection of random variables indexed in time**

Some Objectives of Time Series Analysis

- Find a **statistical model** that adequately explains the **dependence** observed in a time series
- To conduct **statistical inferences**, e.g., Is there evidence of a decreasing trend in the Lake Huron depths?
- To **forecast** future values of the time series based on those we have already observed

● Trends

- One can think of trend, μ_t as continuous changes, usually in the mean, over longer time scales
- Usually the form of the trend is unknown and needs to be estimated. When the trend is removed, we obtain a **detrended** series

● Seasonal or periodic components

- A seasonal component s_t constantly repeats itself in time, i.e., $s_t = s_{t+kd}$
- We need to estimate the form and/or the period d of the seasonal component to **deseasonalize** the series

● The “noise” process

- The noise process, η_t , is the component that is neither trend nor seasonality
- We will focus on finding plausible (typically stationary) statistical models for this process

There are two commonly used approaches

- Additive model:

$$y_t = \mu_t + s_t + \eta_t$$

- Multiplicative model:

$$y_t = \mu_t s_t \eta_t$$

If all $\{y_t\}$ are positive then we obtain the additive model by taking logarithms:

$$\log y_t = \log \mu_t + \log s_t + \log \eta_t$$

- The **mean function** of $\{Y_t\}$ is

$$\mu_t = E[Y_t], \quad t \in T$$

- The **autocovariance function** of $\{Y_t\}$ is

$$\gamma(t, t') = \text{Cov}(Y_t, Y_{t'}) = E[(Y_t - \mu_t)(Y_{t'} - \mu_{t'})], \quad t, t' \in T$$

When $t = t'$ we obtain $\gamma(t, t') = \text{Cov}(Y_t, Y_t) = \text{Var}(Y_t) = \sigma_t^2$,
the variance function of Y_t

The **autocorrelation function (ACF)** of $\{Y_t\}$ is

$$\rho(t, t') = \text{Corr}(Y_t, Y_{t'}) = \frac{\gamma(t, t')}{\sqrt{\gamma(t, t)\gamma(t', t')}}}$$

It measures the strength of **linear association** between Y_t and $Y_{t'}$

Properties:

- 1 $-1 \leq \rho(t, t') \leq 1, \quad t, t' \in T$
- 2 $\rho(t, t') = \rho(t', t), \quad \forall t, t' \in T; \rho(t, t) = 1, \quad \forall t \in T$
- 3 $\rho(t, t')$ is a **non-negative definite** function

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We will still try to keep our models for $\{\eta_t\}$ as simple as possible by assuming **stationarity**, meaning that some characteristic of $\{\eta_t\}$ does not depend on the time points, only on the “time lag” between time points:

- $E[\eta_t] = 0, \quad \forall t \in T$
- $\text{Cov}(\eta_t, \eta_{t'}) = \gamma(t' - t) = \text{Cov}(\eta_{t+s}, \eta_{t'+s})$

\Rightarrow autocorrelation function (ACF):

$$\rho(h) = \frac{\gamma(h)}{\gamma(0)}$$

Let $\{Z_t\}$ be independent and identical random variables that follow $N(0, \sigma^2)$

- Moving Average Processes (MA(q)):

$$\eta_t = Z_t + \theta_1 Z_{t-1} + \theta_2 Z_{t-2} \cdots + \theta_q Z_{t-q}$$

Let $\{Z_t\}$ be independent and identical random variables that follow $N(0, \sigma^2)$

- Moving Average Processes (MA(q)):

$$\eta_t = Z_t + \theta_1 Z_{t-1} + \theta_2 Z_{t-2} \cdots + \theta_q Z_{t-q}$$

- Autoregressive Processes (AR(p)):

$$\eta_t = \phi_1 \eta_{t-1} + \phi_2 \eta_{t-2} + \cdots + \phi_p \eta_{t-p} + Z_t$$

Let $\{Z_t\}$ be independent and identical random variables that follow $N(0, \sigma^2)$

- Moving Average Processes (MA(q)):

$$\eta_t = Z_t + \theta_1 Z_{t-1} + \theta_2 Z_{t-2} \cdots + \theta_q Z_{t-q}$$

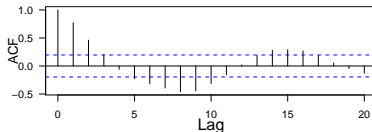
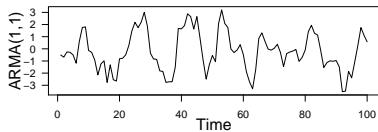
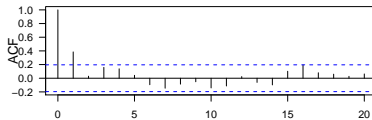
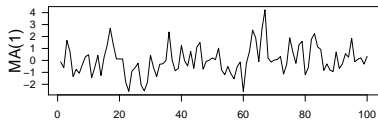
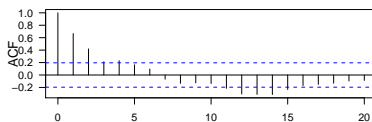
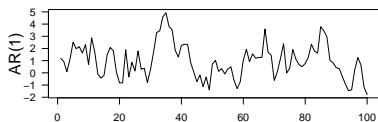
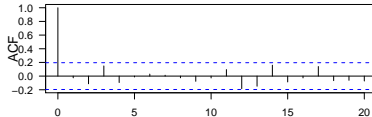
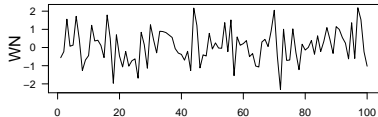
- Autoregressive Processes (AR(p)):

$$\eta_t = \phi_1 \eta_{t-1} + \phi_2 \eta_{t-2} + \cdots + \phi_p \eta_{t-p} + Z_t$$

- Autoregressive Moving Average Processes ARMA(p,q):

$$\eta_t = \phi_1 \eta_{t-1} + \phi_2 \eta_{t-2} + \cdots + \phi_p \eta_{t-p} + Z_t + \theta_1 Z_{t-1} + \theta_2 Z_{t-2} + \cdots + \theta_q Z_{t-q}$$

Autocorrelation Plot



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Lake Huron Case Study



Source: <https://www.worldatlas.com/articles/what-states-border-lake-huron.html>

- Detrending
- Model fitting and selection
- Forecasting

Time Series Data

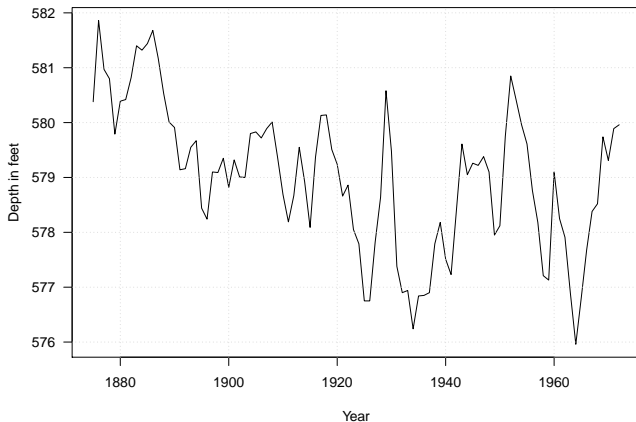
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Annual Measurements of the Level of Lake Huron

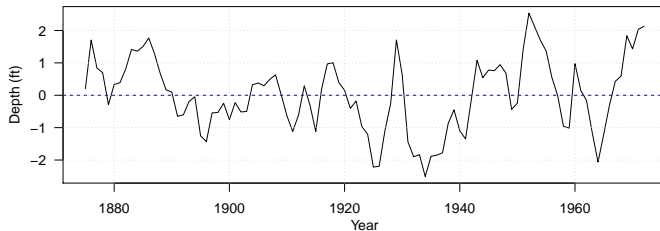
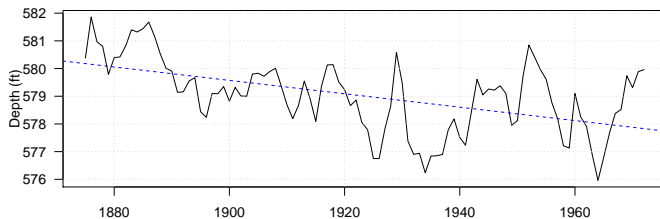


There seems to be a decreasing trend \Rightarrow need to estimate the trend to get the detrended series

Plots of the Trend and Residuals

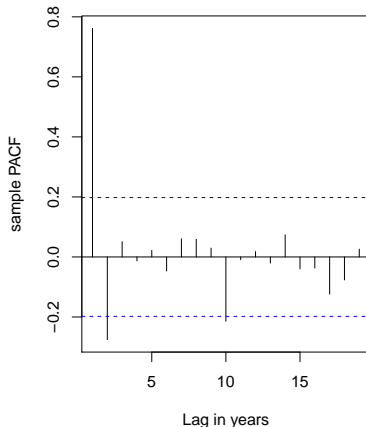
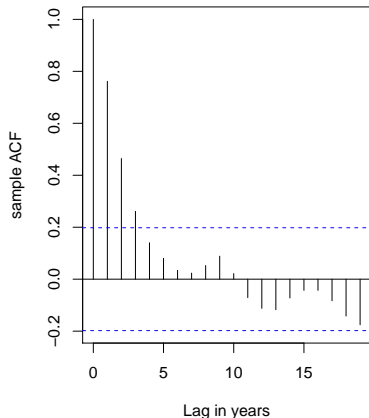
$$y_t = \underbrace{\mu_t}_{\text{trend}} + \underbrace{\eta_t}_{\text{residual}}$$

where we **assume** $\mu_t = \alpha + \beta t$, i.e., a **linear trend in time**

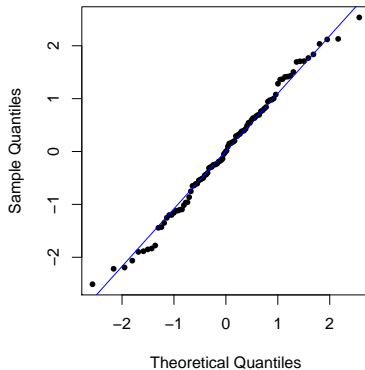
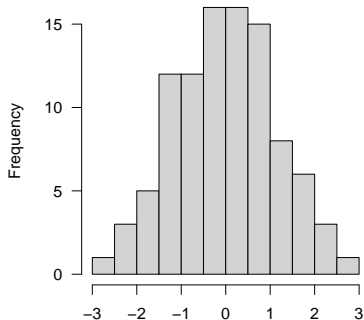


ACF and PACF Plots

- Tapering pattern in ACF \Rightarrow need to include AR terms
- Significant PACF values at the first 2 lags \Rightarrow a AR(2) may be appropriate



Assessing Normality Assumption for η_t



Fitting AR(2)

```
> (ar2.model <- arima(deTrend, order = c(2, 0, 0)))
```

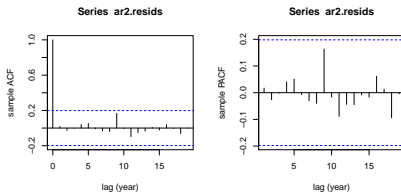
Call:

```
arima(x = deTrend, order = c(2, 0, 0))
```

Coefficients:

	ar1	ar2	intercept
	1.0047	-0.2919	0.0196
s.e.	0.0977	0.1004	0.2351

σ^2 estimated as 0.4571: log likelihood = -101.25, aic = 210.5



```
> Box.test(ar2.resids, type = "Ljung-Box")
```

Box-Ljung test

data: ar2.resids

X-squared = 0.029966, df = 1, p-value = 0.8626

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We can conduct model selection by using, for example, AIC

```
> ar1.model <- arima(deTrend, order = c(1, 0, 0))  
> ar2.model <- arima(deTrend, order = c(2, 0, 0))  
> arma21.model <- arima(deTrend, order = c(2, 0, 1))  
> AIC(ar1.model); AIC(ar2.model); AIC(arma21.model)  
[1] 216.5835  
[1] 210.5032  
[1] 212.1784
```

Fitting AR(2) + a Linear Trend

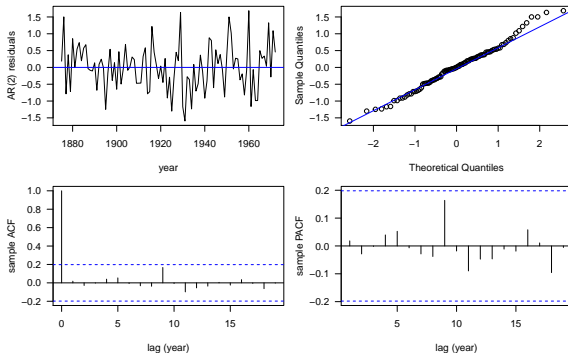
```
> library(forecast)
> (fit <- Arima(LakeHuron, order = c(2, 0, 0), include.drift = T))
Series: LakeHuron
ARIMA(2,0,0) with drift
```

Coefficients:

	ar1	ar2	intercept	drift
	1.0048	-0.2913	580.0915	-0.0216
s.e.	0.0976	0.1004	0.4636	0.0081

sigma² estimated as 0.476: log likelihood=-101.2

AIC=212.4 AICc=213.05 BIC=225.32



10-Year-Ahead Forecasts

