Completely Randomized Designs: Model, Estimation, Inference



Lecture 4

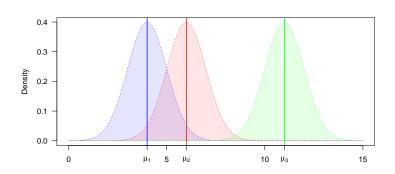
Completely Randomized Designs: Model, Estimation, Inference

STAT 8050 Design and Analysis of Experiments January 21, 2020

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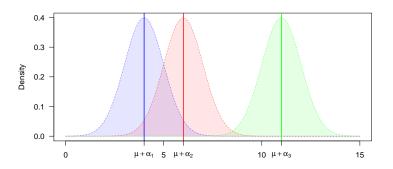


$$Y_{ij} = \mu_i + \epsilon_{ij}, \quad i = 1, \dots, g, \quad j = 1, \dots, n_i, \quad \epsilon_{ij} \sim N(0, \sigma^2)$$





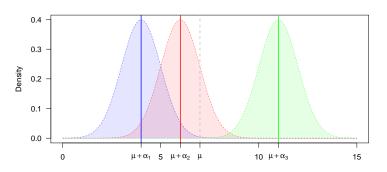
$$Y_{ij} = \mu + \alpha_i + \epsilon_{ij}, \quad i = 1, \dots, g, \quad j = 1, \dots, n_i, \quad \epsilon_{ij} \sim N(0, \sigma^2)$$



Overparameterized. Need to add a constraint so that the parameters are estimable.



Suppose we let $\sum_{i=1}^{g} n_i \alpha_i = 0$

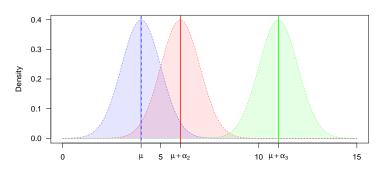


Effects Model Cont'd

Completely
Randomized
Designs: Model,
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Suppose we let $\alpha_1 = 0$



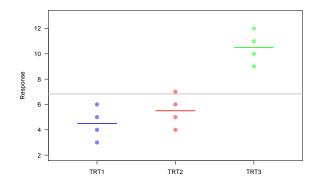
Data Layout & the Dot Notation

Randomized
Designs: Model,
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 y_{ij} is the "observed" response for the j^{th} experimental unit to treatment i.

| Treatment | C | Observ | vatio | ns | Totals | Averages |
|-----------|------------------------|------------------------|-------|------------|-------------------------|---------------|
| 1 | <i>y</i> ₁₁ | <i>y</i> ₁₂ | ••• | y_{1n_1} | <i>y</i> ₁ . | \bar{y}_1 . |
| 2 | <i>y</i> ₂₁ | y ₂₂ | ••• | y_{2n_2} | <i>y</i> ₂ . | $ar{y}_2$. |
| : | ÷ | ÷ | ••• | ÷ | ÷ | : |
| g | y_{g1} | y_{g2} | ••• | y_{gn_g} | y_g . | \bar{y}_g . |
| | | | | | у | <u> </u> |

$$\Rightarrow \underbrace{\sum_{i=1}^{g} \sum_{j=1}^{n_{i}} (y_{ij} - \bar{y}_{..})^{2}}_{SS_{T}} = \underbrace{\sum_{i=1}^{g} n_{i} (\bar{y}_{i\cdot} - \bar{y}_{..})^{2}}_{SS_{TRT}} + \underbrace{\sum_{i=1}^{g} \sum_{j=1}^{n_{i}} (y_{ij} - \bar{y}_{i\cdot})^{2}}_{SS_{E}}$$





ANOVA Table

| , | Source | df | SS | MS | EMS |
|---|-----------|--------------|-------------------|-----------------------------------|--|
| - | Treatment | g-1 | SS _{TRT} | $MS_{TRT} = \frac{SS_{TRT}}{g-1}$ | $\sigma^2 + \frac{\sum_{i=1}^g n_i \alpha_i^2}{g-1}$ |
| | | | | $MS_E = \frac{SS_E}{N-g}$ | |
| - | Total | <i>N</i> – 1 | SS_T | | |

$$SS_{T} = \sum_{i=1}^{g} \sum_{j=1}^{n_{i}} (y_{ij} - \bar{y}_{..})^{2} = \sum_{i=1}^{g} \sum_{j=1}^{n_{i}} y_{ij}^{2} - \frac{y_{..}^{2}}{N}$$

$$SS_{TRT} = \sum_{i=1}^{g} n_{i} (\bar{y}_{i} - \bar{y}_{..})^{2} = \sum_{i=1}^{g} \frac{y_{i}^{2}}{n_{i}} - \frac{y_{..}^{2}}{N}$$

$$SS_{E} = \sum_{i=1}^{g} \sum_{j=1}^{n_{i}} (y_{ij} - \bar{y}_{i}.)^{2} = \sum_{i=1}^{g} \sum_{j=1}^{n_{i}} y_{ij}^{2} - \sum_{i=1}^{g} \frac{y_{i}^{2}}{n_{i}} = SS_{T} - SS_{TRT}$$



F-Test

Testing for treatment effects

$$H_0: \alpha_i = 0$$
 for all i
 $H_a: \alpha_i \neq 0$ for some i

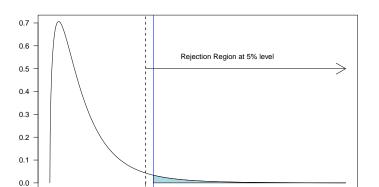
Test statistics: $F = \frac{\text{MS}_{TRT}}{\text{MS}_E}$. Under H_0 , the test statitic follows an F-distribution with g-1 and N-g degrees of freedom Reject H_0 if

$$F_{obs} > F_{g-1,N-g;\alpha}$$

for an α -level test, $F_{g-1,N-g;\alpha}$ is the $100 \times (1-\alpha)\%$ percentile of a central F-distribution with g-1 and N-g degrees of freedom.

The P-value of the F-test is the probability of obtaining F at least as extreme as F_{obs} , that is, $P(F > F_{obs}) \Rightarrow \text{reject } H_0$ if P-value $< \alpha$.

F Distribution and the F-Test



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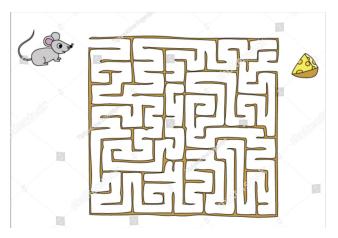
F_{0.95, df1, df2} F_{obs} 4

2

0



Example



Source: https://www.shutterstock.com/image-vector/find-your-way-cheese-mouse-maze-232569073

