

Lecture 12

Discrimination and Classification

Readings: Zelterman, 2015, Chapters 10

DSA 8070 Multivariate Analysis

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Agenda

Overview

Binary Linear
Classification

1 Overview

2 Binary Linear Classification

Main objectives behind discrimination and classification are:

- **Discrimination**: separate distinct “populations” of observations
- **Classification**: classify new objectives into pre-defined populations

Examples

- Given measurements on the concentrations of five elements in bullet lead, find combinations of those concentrations that best describe bullets made by Cascade, Federal, Winchester and Remington ⇒ **discrimination**
- Using information on prisoners eligible for parole (good behavior, history of drug use, job skills, etc) can we successfully allocate a prisoners eligible for parole into two groups: those who will commit another crime or those who will not commit another crime? ⇒ **classification**

- **Data:**

$$\{\mathbf{X}_i, Y_i\}_{i=1}^n,$$

where Y_i is the class information for the i_{th} observation
 $\Rightarrow Y$ is a qualitative variable

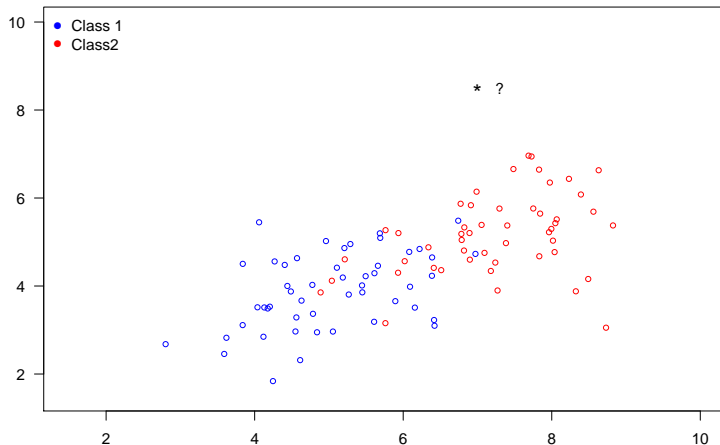
- **Classification** aims to classify a new observation (or several new observations) into one of those classes

Quantity of interest: $P(Y = k_{th} \text{ category} | \mathbf{X} = \mathbf{x})$

- In this lecture we will focus on **binary linear classification**

Toy Example

Wish to classify a new observation $x_i = (x_{1i}, x_{2i})$, denoted by $(*)$, into one of the two groups (**class 1** or **class 2**)



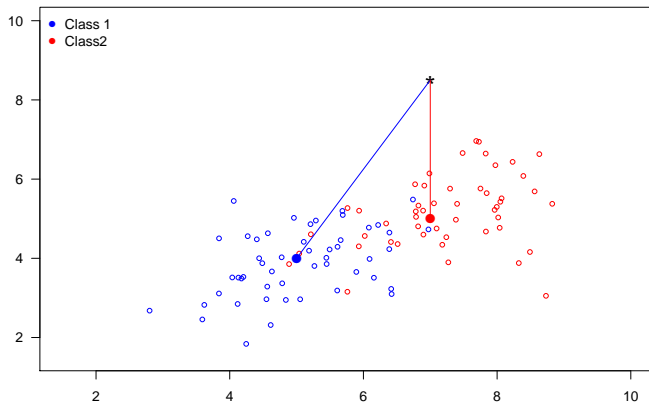
Toy Example Cont'd

We can compute the distances from this new observation $\mathbf{x} = (x_1, x_2)$ to the groups, for example,

$$d_1 = \sqrt{(x_1 - \mu_{11})^2 + (x_2 - \mu_{12})^2},$$

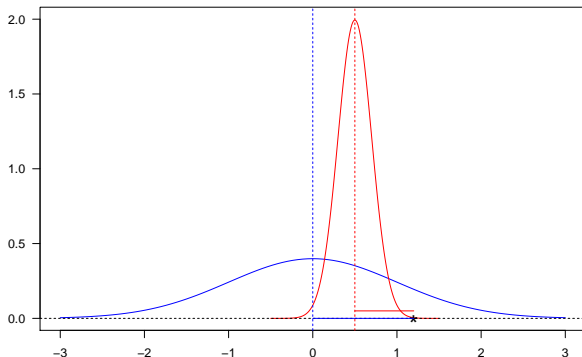
$$d_2 = \sqrt{(x_1 - \mu_{21})^2 + (x_2 - \mu_{22})^2}.$$

We can assign \mathbf{x} to the group with the smallest distance



Variance Corrected Distance

In this one-dimensional example, $d_1 = |x - \mu_1| > |x - \mu_2|$. Does that mean x is “closer” to group 2 (red) than group 1 (blue)?



We should take the “spread” of each group into account.

$$\tilde{d}_1 = |x - \mu_1|/\sigma_1 < \tilde{d}_2 = |x - \mu_2|/\sigma_2$$

General Covariance Adjusted Distance: Mahalanobis Distance

The [Mahalanobis distance](#) [Mahalanobis, 1936] is a measure of the distance between a point x and a multivariate distribution of X :

$$D_M(x) = \sqrt{(x - \mu)^T \Sigma^{-1} (x - \mu)},$$

where μ is the mean vector and Σ is the variance-covariance matrix of X

One can use the Mahalanobis distance, by computing the Mahalanobis distance between an observations x_i and the “center” of the k_{th} population μ_k , to carry out classification

Binary Classification with Multivariate Normal Populations

Assume $X_1 \sim \text{MVN}(\mu_1, \Sigma)$, $X_2 \sim \text{MVN}(\mu_2, \Sigma)$, that is,
 $\Sigma_1 = \Sigma_2 = \Sigma$

- Maximum Likelihood of group membership:

Group 1 if $\ell(x, \mu_1, \Sigma) > \ell(x, \mu_2, \Sigma)$

- Linear Discriminant Function:

Group 1 if $(\mu_1 - \mu_2)^T \Sigma^{-1} x - \frac{1}{2}(\mu_1 - \mu_2)^T \Sigma^{-1}(\mu_1 + \mu_2) > 0$

- Minimize Mahalanobis distance:

Group 1 if $(x - \mu_1)^T \Sigma^{-1}(x - \mu_1) < (x - \mu_2)^T \Sigma^{-1}(x - \mu_2)$

All the methods above are equivalent

In addition to the observed characteristics of units $\{x_i\}_{i=1}^n$, other features of classification rules are:

- Prior probability:

If one population is more prevalent than the other, chances are higher that a new unit came from the larger population. Stronger evidence would be needed to allocate the unit to the population with the smaller prior probability.

- Costs of misclassification:

It may be more costly to misclassify a seriously ill subject as healthy than to misclassify a healthy subject as being ill.

Classification Regions

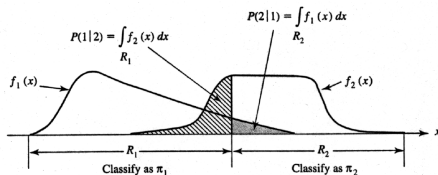
Ignore for now the prior probabilities of each population and the potentially different misclassification costs.

- The probability of misclassifying an object into π_2 when it belongs in π_1 is

$$P(2|1) = \mathbb{P}(X \in \mathcal{R}_2 | \pi_1)$$

- The probability of misclassifying an object into π_1 when it belongs in π_2 is

$$P(1|2) = \mathbb{P}(X \in \mathcal{R}_1 | \pi_2)$$



Source: Figure 11.3 from Applied Multivariate Statistical Analysis, 6th Ed (Johnson & Wichern). Visualization is for $p = 1$ variable.

Let p_1 and p_2 denote the prior probabilities of π_1, π_2 , and $c(1|2), c(2|1)$ be the costs of misclassification:

- Then probabilities of the four possible outcomes are:

$$\mathbb{P}(\text{correctly classified as } \pi_1) = \mathbb{P}(\mathbf{X} \in \mathcal{R}_1 | \pi_1) \mathbb{P}(\pi_1) = P(1|1)p_1$$

$$\mathbb{P}(\text{incorrectly classified as } \pi_1) = \mathbb{P}(\mathbf{X} \in \mathcal{R}_1 | \pi_2) \mathbb{P}(\pi_2) = P(1|2)p_2$$

$$\mathbb{P}(\text{correctly classified as } \pi_2) = \mathbb{P}(\mathbf{X} \in \mathcal{R}_2 | \pi_2) \mathbb{P}(\pi_2) = P(2|2)p_2$$

$$\mathbb{P}(\text{incorrectly classified as } \pi_2) = \mathbb{P}(\mathbf{X} \in \mathcal{R}_2 | \pi_1) \mathbb{P}(\pi_1) = P(2|1)p_1$$

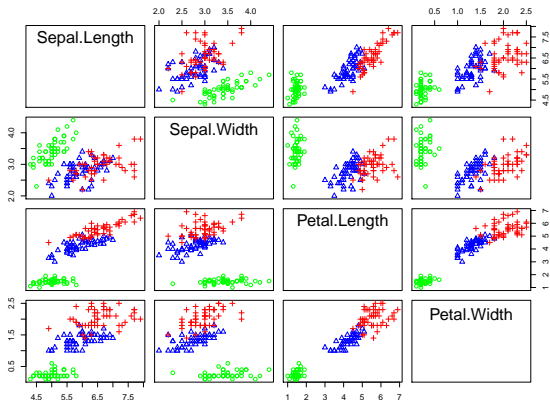
- Classification rules are often evaluated in terms of the **expected cost of misclassification (ECM)**:

$$\text{ECM} = c(2|1)p(2|1)p_1 + c(1|2)P(1|2)P(1|2)p_2,$$

and we seek rules that **minimize the ECM**

Example: Fisher's Iris Data

4 variables (sepal length and width and petal length and width),
3 species (**setosa**, **versicolor**, and **virginica**)



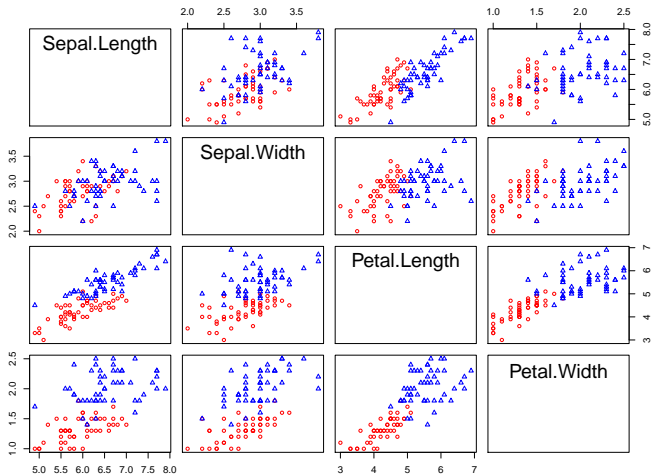
Task: Classify flowers into different species based on lengths and widths of sepal and petal

Fisher's Iris Data Cont'd

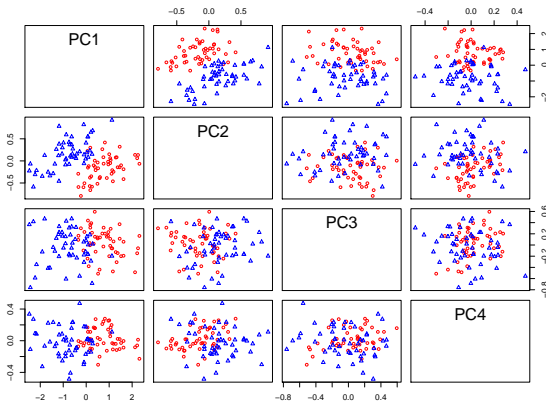
Let's focus on the latter two classes (**versicolor**, and **virginica**)

Overview

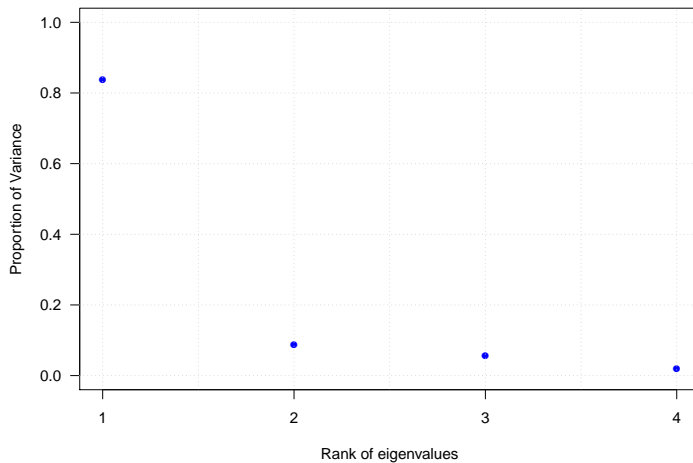
Binary Linear
Classification



To further simplify the matter, let's focus on the first two PCs of X



Screen Plot

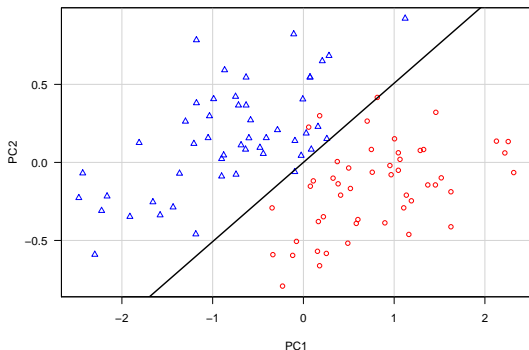


Linear Discriminant Analysis

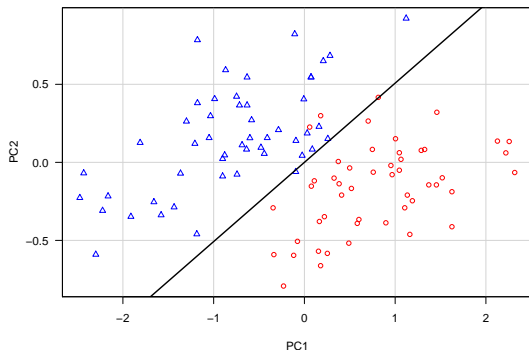
Main idea: Use Bayes rule to compute

$$P(Y = k | \mathbf{X} = \mathbf{x}) = \frac{P(Y = k)P(\mathbf{X} = \mathbf{x} | Y = k)}{P(\mathbf{X} = \mathbf{x})} = \frac{\pi_k f_k(\mathbf{x})}{\sum_{k=1}^K \pi_k f_k(\mathbf{x})}.$$

Assuming $f_k(\mathbf{x}) \sim \text{MVN}(\boldsymbol{\mu}_k, \Sigma)$, $k = 1, \dots, K$ and use $\hat{\pi}_k = \frac{n_k}{n} \Rightarrow$ it turns out the resulting classifier is **linear in \mathbf{X}**



Classification Performance Evaluation



```
fit.LDA
```

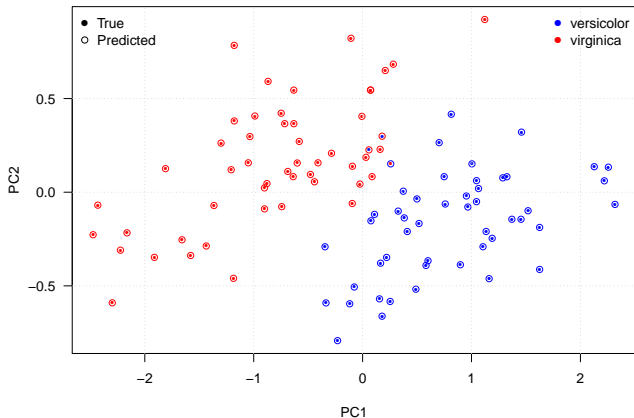
```
versicolor virginica
```

```
versicolor      47      3
```

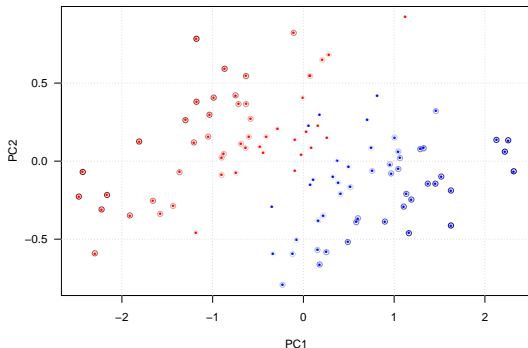
```
virginica       1      49
```

Logistic Regression Classifier

Main idea: Model the logit $\log\left(\frac{P(Y=1)}{1-P(Y=1)}\right)$ as a linear function in \mathbf{x}



Logistic Regression Classifier Cont'd



	logisticPred	
	versicolor	virginica
versicolor	48	2
virginica	1	49

For a binary classification problem, one can show that both linear discriminant analysis (LDA) and logistic regression are **linear classifiers**. The difference is in how the parameters are estimated:

- Logistic regression uses the conditional likelihood based on $P(Y|X = x)$
- LDA uses the full likelihood based on multivariate normal assumption on X
- Despite these differences, in practice the results are often very similar

Quadratic Discriminant Analysis

In linear discriminant analysis, we **assume** $\{f_k(x)\}_{k=1}^K$ are normal densities and $\Sigma_1 = \Sigma_2$, therefore we obtain a **linear classifier**. What if $\Sigma_1 \neq \Sigma_2 \Rightarrow$ we get **quadratic discriminant analysis**

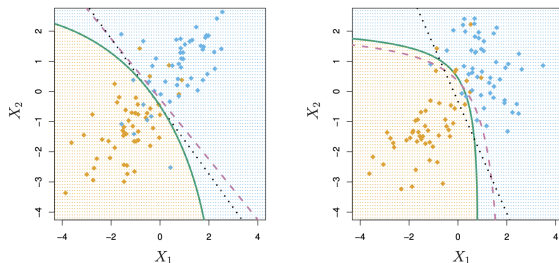


Figure: Figure courtesy of [An Introduction of Statistical Learning](#) by G. James et al. pp. 150