

# Lecture 5


## Time Series Regression

Reading: Forecasting, Time Series, and Regression (4th edition) by Bowerman, O'Connell, and Koehler: Chapter 6

MATH 4070: Regression and Time-Series Analysis

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Time Series Regression

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Time Series Data  
Trend Estimation  
Estimating Seasonality

5.1

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### Agenda

- 1 Time Series Data
- 2 Trend Estimation
- 3 Estimating Seasonality

Time Series Regression

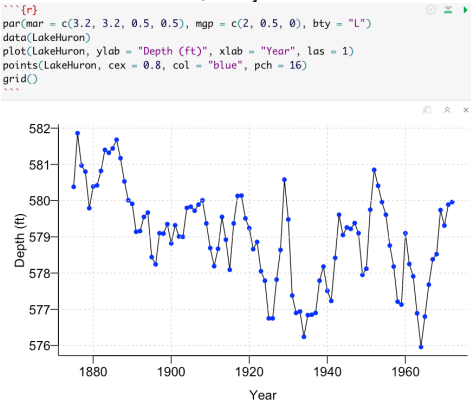
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### Level of Lake Huron 1875–1972

Annual measurements of the level of Lake Huron in feet.  
[Source: Brockwell & Davis, 1991]



Time Series Regression

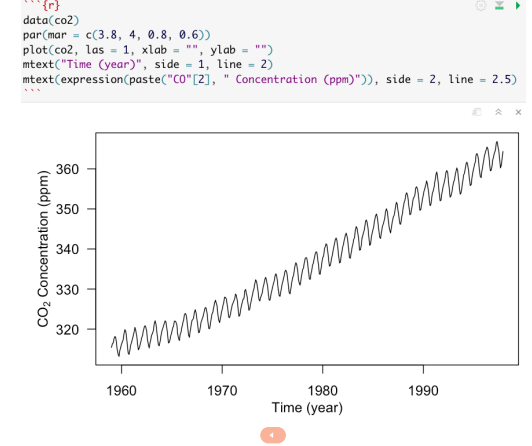
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Mauna Loa Monthly Atmospheric CO<sub>2</sub> Concentration

[Source: Keeling & Whorf, Scripps Institution of Oceanography]



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Time Series Data

- A **time series** is a collection of observations  $\{y_t, t \in T\}$  taken sequentially in time ( $t$ ) with the index set  $T$ 
  - $T = \{0, 1, 2, \dots, T\} \subset \mathbb{Z} \Rightarrow$  **discrete-time time series**
  - $T = [0, T] \subset \mathbb{R} \Rightarrow$  **continuous-time time series**
- A discrete-time time series might be intrinsically discrete or might arise from a underlying continuous-time time series via
  - sampling (e.g., instantaneous wind speed)
  - aggregation (e.g., daily accumulated precipitation amount)
  - extrema (e.g., daily maximum temperature)
- We will focus on dealing with **discrete-time real-valued** ( $Y_t \in \mathbb{R}$ ) **time series** in this course

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Exploratory Time Series Analysis

- Start with a **time series plot**, i.e., to plot  $y_t$  versus  $t$
- Look at the following:
  - Are there abrupt changes?
  - Are there "outliers"?
  - Is there a need to transform the data?
- Examine the **trend**, **seasonal components**, and the "noise" term

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Features of Times Series

- Trends ( $\mu_t$ )
  - $\mu_t$  represents continuous changes, usually in the mean, over longer time scales. "The essential idea of trend is that it shall be smooth." - [Kendall, 1973]
  - The form of the trend is typically unknown and needs to be estimated. Removing the trend yields a **detrended** series
- Seasonal or Periodic Components ( $s_t$ )
  - $s_t$  repeats consistently over time, i.e.,  $s_t = s_{t+kd}$
  - The form and period  $d$  of the seasonal component must be estimated to **deseasonalize** the series.
- The "Noise" Process ( $\eta_t$ )
  - $\eta_t$  represents the component that is neither trend nor seasonality
  - Focus on finding plausible statistical models for this process

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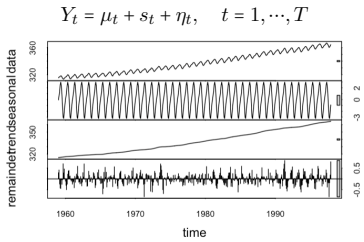
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Combining Trend, Seasonality, and Noise Together

There are two commonly used approaches

- Additive model:




- Multiplicative model:

$$Y_t = \mu_t s_t \eta_t, \quad t = 1, \dots, T$$

If all  $\{y_t\}$  are positive then we obtain the additive model by taking logarithms:

$$\log Y_t = \log \mu_t + \log s_t + \log \eta_t, \quad t = 1, \dots, T$$

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The (Additive) Decomposition Model


- The additive model for a time series  $\{Y_t\}$  is

$$Y_t = \mu_t + s_t + \eta_t,$$

where

- $\mu_t$  is the **trend** component
- $s_t$  is the **seasonal** component
- $\eta_t$  is the **random (noise)** component with  $\mathbb{E}(\eta_t) = 0$
- Standard procedure:
  - Estimate/remove the trend and seasonal components
  - Analyze the remainder, the residuals  
 $\hat{\eta}_t = y_t - \hat{\mu}_t - \hat{s}_t$
- We will focus on (1) for this week

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Estimating Trend for Nonseasonal Model

Assuming  $s_t = 0$  (i.e., there is no “seasonal” variation), we have


$$Y_t = \mu_t + \eta_t,$$

with  $\mathbb{E}(\eta_t) = 0$

Methods for **estimating trends**

- Least squares regression
- Smoothing

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Trend Estimation: Linear Regression

- The additive nonseasonal time series model for  $\{Y_t\}$  is


$$Y_t = \mu_t + \eta_t,$$

where the trend is assumed to be a linear combination of known covariate series  $\{x_{it}\}_{i=1}^p$

$$\mu_t = \beta_0 + \sum_{i=1}^p \beta_i x_{it}.$$

- Here we want to **estimate**  $\beta = (\beta_0, \beta_1, \dots, \beta_p)^T$  from the data  $\{y_t, \{x_{it}\}_{i=1}^p\}_{t=1}^T$
- You're likely quite familiar with this formulation already  $\Rightarrow$  [Regression Analysis](#)

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Some Examples of Covariate Series  $\{x_{it}\}$

- **Simple linear regression model:**

$$\mu_t = \beta_0 + \beta_1 x_t,$$

for example, the temperature trend at time  $t$  could be a constant ( $\beta_0$ ) plus a multiple ( $\beta_1$ ) of the carbon dioxide level at time  $t$  ( $x_t$ )


- **Polynomial regression model:**

$$\mu_t = \beta_0 + \sum_{i=1}^p \beta_i t^i$$

- **Change point model:**

$$\mu_t = \begin{cases} \beta_0 & \text{if } t \leq t^*; \\ \beta_0 + \beta_1 & \text{if } t \geq t^*. \end{cases}$$

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## Parameter Estimation: Ordinary Least Squares

- Like in the linear regression setting, we can estimate the parameters via **ordinary least squares (OLS)**

- Specifically, we minimize the following objective function:

$$\ell_{ols} = \sum_{t=1}^T (y_t - \beta_0 - \sum_{k=1}^p x_{kt} \beta_k)^2.$$

- The estimates  $\beta = (\beta_0, \beta_1, \dots, \beta_p)^T$  minimizing the above objective function are called the **OLS estimates of  $\beta$**   $\Rightarrow$  they are easiest to express in **matrix form**

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## The Model and Parameter Estimates in Matrix Form

- Matrix representation:

$$Y = X\beta + \eta,$$

where  $Y = \begin{bmatrix} Y_1 \\ \vdots \\ Y_T \end{bmatrix}$ ,  $X = \begin{bmatrix} 1 & x_{11} & x_{12} & \cdots & x_{1p} \\ 1 & x_{21} & x_{22} & \cdots & x_{2p} \\ \vdots & \vdots & \vdots & \cdots & \vdots \\ 1 & x_{T1} & x_{T2} & \cdots & x_{Tp} \end{bmatrix}$ , and

$$\eta = \begin{bmatrix} \eta_1 \\ \vdots \\ \eta_T \end{bmatrix}$$

- Assuming  $X^T X$  is **invertible**, the OLS estimate of  $\beta$  can be shown to be

$$\hat{\beta} = (X^T X)^{-1} X^T Y,$$

and the `lm` function in R calculates OLS estimates

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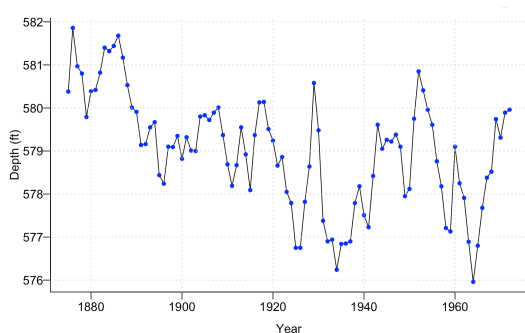
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## Lake Huron Example Revisited



Let's **assume** there is a **linear trend in time**  $\Rightarrow$  we need to estimate the **intercept**  $\beta_0$  and **slope**  $\beta_1$

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The R Output

```
Call:
lm(formula = LakeHuron ~ yr)

Residuals:
    Min       1Q   Median       3Q      Max
-2.50997 -0.72726  0.00083  0.74402  2.53565

Coefficients:
            Estimate Std. Error t value Pr(>|t|)
(Intercept)  625.554918   7.764293   80.568 < 2e-16 ***
yr          -0.024201   0.004036  -5.996 3.55e-08 ***
---
Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

Residual standard error: 1.13 on 96 degrees of freedom
Multiple R-squared:  0.2725,    Adjusted R-squared:  0.2649
F-statistic: 35.95 on 1 and 96 DF,  p-value: 3.545e-08
```

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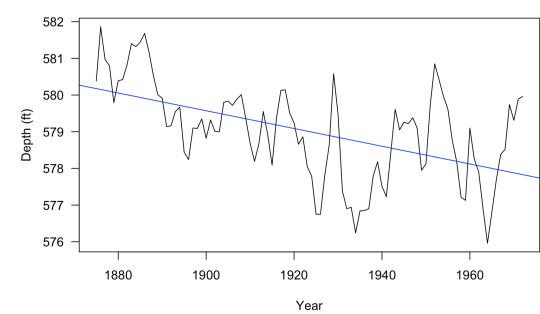
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Plot the (Estimated) Trend  $\hat{\mu}_t = \hat{\beta}_0 + \hat{\beta}_1 t$



$\hat{\beta}_1 = -0.0242$  (ft/yr)  $\Rightarrow$  there seems to be a decreasing trend

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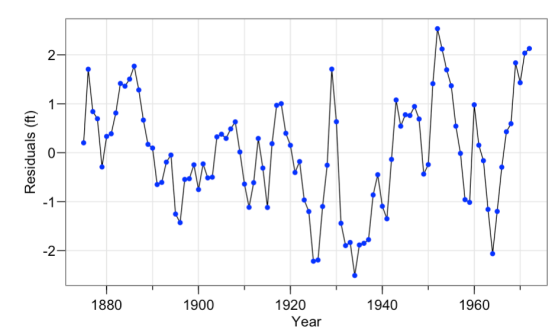
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Plot the Residuals  $\{\hat{\eta}_t = y_t - \hat{\beta}_0 - \hat{\beta}_1 t\}$



$\{\hat{\eta}_t\}$  seems to exhibit some temporal dependence structure, should we worry about the results we have (recall OLS makes an i.i.d. assumption)?

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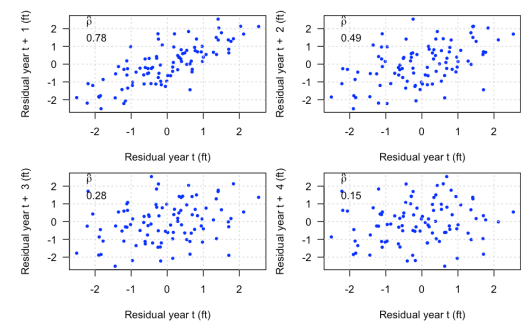
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Exploring the Dependence Structure of “Noise”  $\{\eta_t\}$

$\{\eta_t\}$  exhibit a temporal dependence structure, meaning that the nearby (in time) values tend to be more alike than those that are far part. To observe this, let's create a few time lag plots



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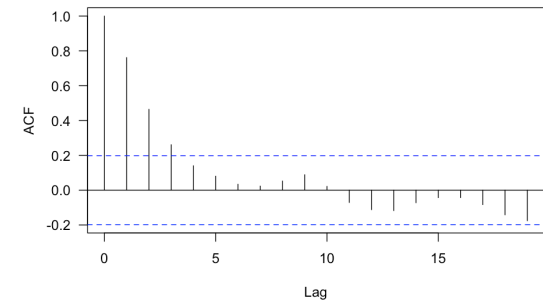
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Further Exploration of the Temporal Dependence Structure

Let's plot the correlation as a function of the time lag



We will learn how to use this information to suggest an appropriate model

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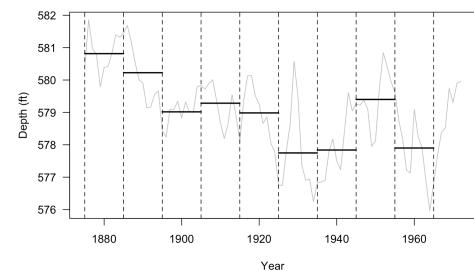
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Smoothing or Local Averaging

In some cases, we may relax the trend assumption using a 'non-parametric' approach.

We divide the time series into small blocks (each with 10 years of data) and average each block.



Doing this gives a very rough estimate of the trend. **Can we do better?**

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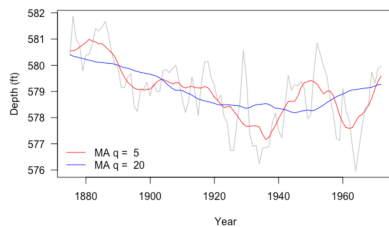
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## Moving Average Smoother

A **moving average smoother** estimates the trend at time  $t$  by averaging the current observation and the  $q$  nearest observations from either side. That is

$$\hat{\mu}_t = \frac{1}{2q+1} \sum_{j=-q}^q y_{t-j}$$



$q$  is the “smoothing” parameter, which controls the smoothness of the estimated trend  $\hat{\mu}_t$

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## Exponential Smoothing

- Let  $\alpha \in [0, 1]$  be some fixed constant, defined

$$\hat{\mu}_t = \begin{cases} Y_1 & \text{if } t = 1; \\ \alpha Y_t + (1 - \alpha)\hat{\mu}_{t-1} & t = 2, \dots, T. \end{cases}$$

- For  $t = 2, \dots, T$ , we can rewrite  $\hat{\mu}_t$  as

$$\sum_{j=0}^{t-2} \alpha(1 - \alpha)^j Y_{t-j} + (1 - \alpha)^{t-1} Y_1.$$

$\Rightarrow$  it is a one-sided moving average filter with **exponentially decreasing weights**. One can alter  $\alpha$  to control the amounts of smoothing (see next slide for an example)

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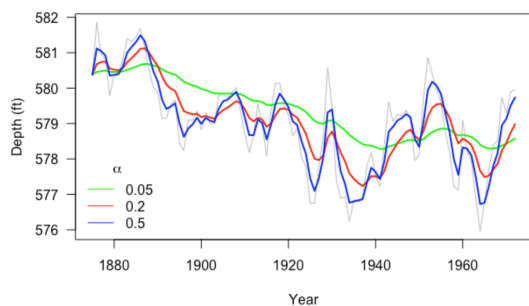
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## $\alpha$ is the Smoothing Parameter for Exponential Smoothing



The smaller the  $\alpha$ , the smoother the resulting trend

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Seasonal Component Estimation

Let's consider a situation where a time series consists of only a seasonal component (assuming the trend has been estimated/removed). In this scenario,

Y\_t = s\_t + η\_t,

with {s\_t} having period d (i.e., s\_{t+jd} = s\_t for all integers j and t), ∑\_{t=1}^d s\_t = 0 and E(η\_t) = 0

Two methods to estimate {s\_t}

- Harmonic regression
- Seasonal mean model

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Harmonic Regression

- A harmonic regression model has the form

s\_t = ∑\_{j=1}^k A\_j cos(2πf\_j + φ\_j).

For each j = 1, ..., k:

- A\_j > 0 is the amplitude of the j-th cosine wave
  - f\_j controls the frequency of the j-th cosine wave (how often waves repeats)
  - φ\_j ∈ [-π, π] is the phase of the j-th wave (where it starts)
- The above can be expressed as

∑\_{j=1}^k (β\_{1j} cos(2πf\_j) + β\_{2j} sin(2πf\_j)),

where β\_{1j} = A\_j cos(φ\_j) and β\_{2j} = A\_j sin(φ\_j) ⇒ if {f\_j}\_{j=1}^k are known, we can use regression techniques to estimate the parameters {β\_{1j}, β\_{2j}}\_{j=1}^k

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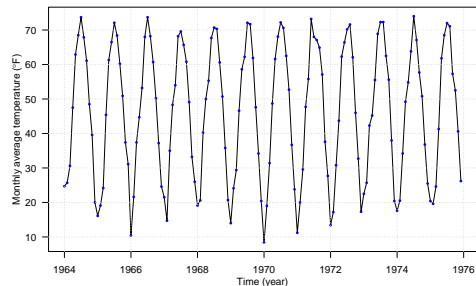
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Monthly Average Temperature in Dubuque, IA [Cryer & Chan, 2008]



Let's assume that there is no trend in this time series. In this context, our goal is to estimate s\_t, the seasonal component.

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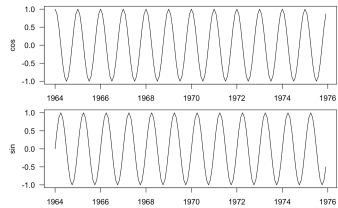
Use a Harmonic Regression to Model Annual Cycles

**Model:**  $s_t = \beta_0 + \beta_1 \cos(2\pi t) + \beta_2 \sin(2\pi t)$


⇒ annual cycles can be modeled by a linear combination of cos and sin with 1-year period.

In R, we can easily create these harmonics using the harmonic function in the TSA package

```
harmonics <- harmonic(tempdub, 1)
```



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R Code & Output


```
##{r}
harReg <- lm(tempdub ~ harmonics)
summary(harReg)
```

```
Call:
lm(formula = tempdub ~ harmonics)

Residuals:
    Min       1Q   Median       3Q      Max
-11.1580  -2.2756  -0.1457   2.3754  11.2671

Coefficients:
            Estimate Std. Error t value Pr(>|t|)
(Intercept)    46.2660     0.3088  149.816 < 2e-16 ***
harmonicscos(2*pi*t) -26.7079     0.4367  -61.154 < 2e-16 ***
harmonicssin(2*pi*t)  -2.1697     0.4367   -4.968 1.93e-06 ***
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Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
```

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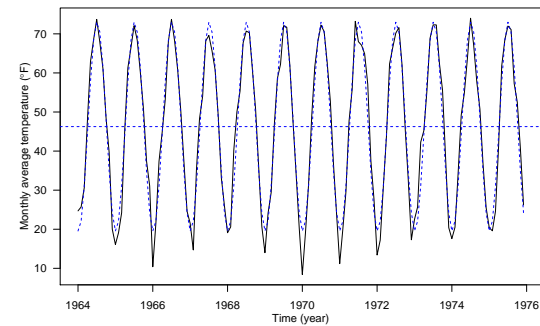
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The Harmonic Regression Model Fit



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Seasonal Means Model


• **Harmonics regression** assumes the seasonal pattern has a regular shape, i.e., the height of the peaks is the same as the depth of the troughs

• A less restrictive approach is to model  $\{s_t\}$  as

$$s_t = \begin{cases} \beta_1 & \text{for } t = 1, 1 + d, 1 + 2d, \dots; \\ \beta_2 & \text{for } t = 2, 2 + d, 2 + 2d, \dots; \\ \vdots & \vdots \\ \beta_d & \text{for } t = d, 2d, 3d, \dots \end{cases}$$

• This is the **seasonal means** model, the parameters  $(\beta_1, \beta_2, \dots, \beta_d)^T$  can be estimated under the linear model framework (think about ANOVA)

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R Output

```
Call:
lm(formula = tempdub ~ month - 1)

Residuals:
    Min       1Q   Median       3Q      Max
-8.2750 -2.2479  0.1125  1.8896  9.8250

Coefficients:
            Estimate Std. Error t value Pr(>|t|)
monthJanuary    16.608      0.987   16.83 <2e-16 ***
monthFebruary    20.650      0.987   20.92 <2e-16 ***
monthMarch       32.475      0.987   32.90 <2e-16 ***
monthApril       46.525      0.987   47.14 <2e-16 ***
monthMay         58.092      0.987   58.86 <2e-16 ***
monthJune        67.500      0.987   68.39 <2e-16 ***
monthJuly        71.717      0.987   72.66 <2e-16 ***
monthAugust      69.333      0.987   70.25 <2e-16 ***
monthSeptember   61.025      0.987   61.83 <2e-16 ***
monthOctober     50.975      0.987   51.65 <2e-16 ***
monthNovember    36.650      0.987   37.13 <2e-16 ***
monthDecember    23.642      0.987   23.95 <2e-16 ***
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Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
```

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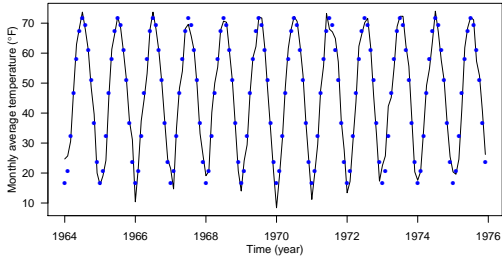
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The Seasonal Means Model Fit



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Trend Estimation  
Estimating Seasonality

5.33

Notes

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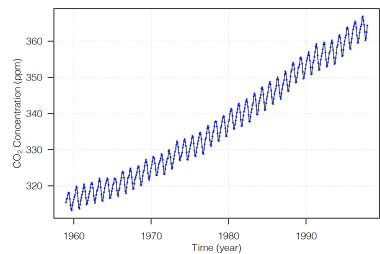
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Estimating the Trend and Seasonal variation Together



Let's perform a regression analysis to model both  $\mu_t$  (assuming a linear time trend) and  $s_t$  (using cos and sin)

```
##{r}
time <- as.numeric(time(co2))
harmonics <- harmonic(co2, 1)

lm_trendSeason <- lm(co2 ~ time + harmonics)
summary(lm_trendSeason)
```

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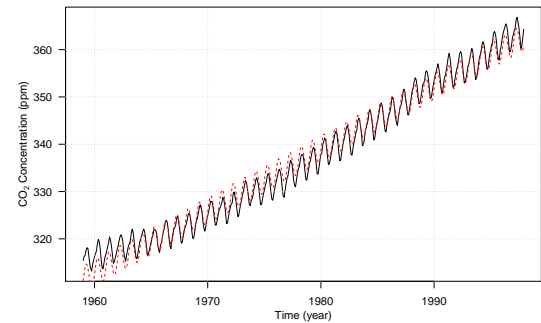
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The Regression Fit



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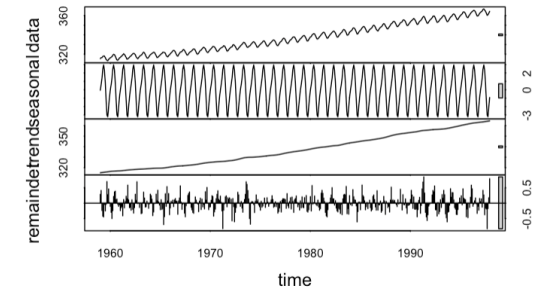
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Seasonal and Trend decomposition using Loess [Cleveland, et. al., 1990]

```
##{r}
# Seasonal and Trend decomposition using Loess (STL)
par(mar = c(4, 3.6, 0.8, 0.6))
stl <- stl(co2, s.window = "periodic")
plot(stl, las = 1)
```



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Summary

These slides cover:

- Main features of a time series: trend, seasonality, and “noise”
- Estimating trends using multiple linear regression and “nonparametric” smoothing
- Estimating seasonality using harmonic regression and the seasonal mean model

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R Functions to Know

- Visualizing time series data: `plot` (for `ts` objects), `ts.plot`, `tsplot` (`astsa` package)
- Fitting time series regression: `lm`, `harmonic` (`TSA` package) for creating harmonic predictors, `filter` for smoothing
- Seasonal and trend decomposition: `stl`

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