

# Lecture 7

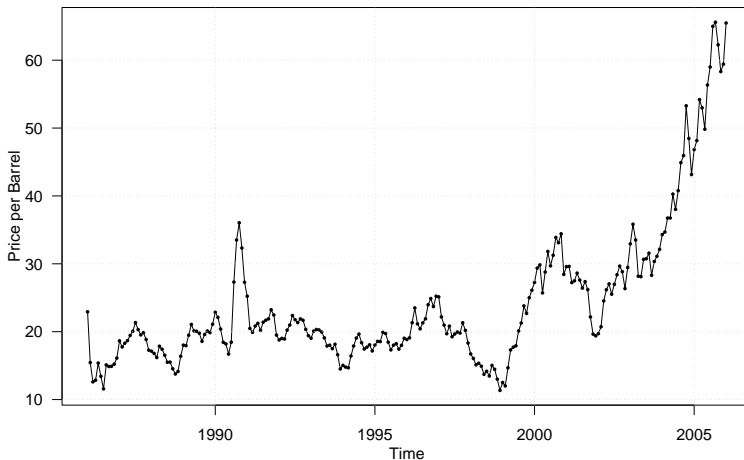
## Nonstationary Time Series Models

Readings: Cryer & Chan Ch 5; Brockwell & Davis Ch 6.1-6.4;  
Shumway & Stoffer Ch 3.6-3.7

*MATH 8090 Time Series Analysis*  
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# Monthly Price of Oil: January 1986–January 2006



Recall the random walk process

$$X_t = Z_1 + Z_2 + \cdots + Z_t = \sum_{j=1}^t Z_j,$$

where  $\{Z_t\} \sim \text{WN}(0, \sigma^2)$

$\{X_t\}$  is a **nonstationary process**

- We can obtain a **stationary** process by **differencing**

$$\nabla X_t = X_t - X_{t-1} = (1 - B)X_t = Z_t$$

- $\{X_t\}$  is an example of an **autoregressive integrated moving average** (ARIMA) process— ARIMA(0, 1, 0) process

An ARIMA model is an ARMA process after differencing

- Let  $d$  be a non-negative integer. Then  $X_t$  is an ARIMA( $p, d, q$ ) process if

$$Y_t = \nabla^d X_t = (1 - B)^d X_t$$

is a **causal** ARMA process

- Let  $\phi(B)$  be the AR polynomial and  $\theta(B)$  be the MA polynomial. Then for  $\{Z_t\} \sim \text{WN}(0, \sigma^2)$

$$\phi(B)Y_t = \theta(B)Z_t,$$

and since  $Y_t = (1 - B)^d X_t$

$$\phi(B)(1 - B)^d X_t = \theta(B)Z_t$$

**Example: ARIMA(1, 1, 0)**

Let  $\phi(z) = 1 - \phi_1 z$ ,  $\theta(z) = 1$  and  $d = 1$ . For a **causal stationary solution** (after differencing) we need to assume  $|\phi_1| < 1$ . Then  $\{X_t\}$  is an ARIMA (1, 1, 0) process,

$$(1 - \phi_1 B)(1 - B)X_t = Z_t,$$

where  $\{Z_t\} \sim \text{WN}(0, \sigma^2)$

Now let  $Y_t = (1 - B)X_t = X_t - X_{t-1}$ , after some rearrangements we have

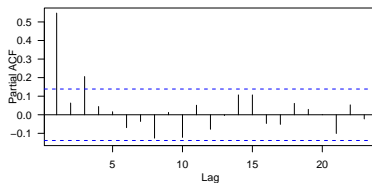
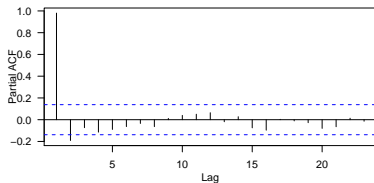
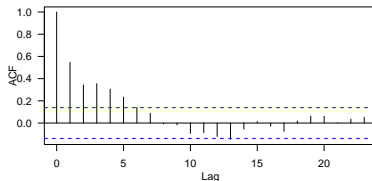
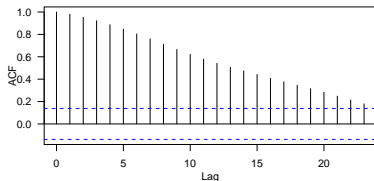
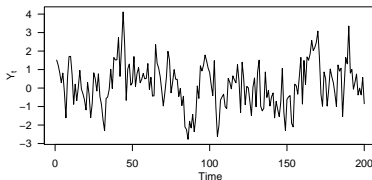
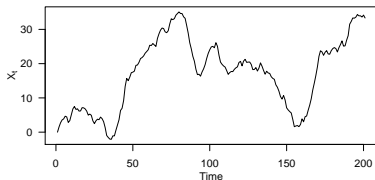
$$\begin{aligned} X_t &= X_{t-1} + Y_t \\ &= (X_{t-2} + Y_{t-1}) + Y_t \\ &\vdots \\ &= X_0 + \sum_{j=1}^t Y_j \end{aligned}$$

Thus  $\{X_t\}$  is a “sort of random walk”—we **cumulatively sum** an AR(1) process,  $\{Y_t\}$

# Simulated ARIMA and Differenced ARMA Process

We simulate an ARIMA(1,1,0):

$$(1 - 0.5B)(1 - B)X_t = Z_t, \quad \{Z_t\} \sim N(0, 1)$$



## Adding a Polynomial Trend

For  $d \geq 1$ , let  $\{X_t\}$  be an  $\text{ARIMA}(p, d, q)$  process. Then  $\{X_t\}$  satisfies the equation

$$\phi(B)(1-B)^d X_t = \theta(B)Z_t$$

- Let  $\mu_t$  be a polynomial of degree  $(d-1)$ , i.e.,  $\mu_t = \sum_{j=0}^{d-1} a_j t^j$  for constants  $\{a_j\}$
- Now let  $V_t = \mu_t + X_t$ , then

$$\begin{aligned}\phi(B)(1-B)^d V_t &= \phi(B)(1-B)^d (\mu_t + X_t) \\ &= \phi(B)(1-B)^d \mu_t + \phi(B)(1-B)^d X_t \\ &= 0 + \phi(B)(1-B)^d X_t \\ &= \theta(B)Z_t\end{aligned}$$

- Takeaway:**  $\text{ARIMA}(p, d, q)$  are useful for modeling data with **polynomial trends**, due to the inherent differencing that can be used to remove trends

# Typical Steps for Modeling ARIMA Processes: Exploratory Data Analysis

- Plot the data, ACF, PACF and Q-Q plots
  - Check for unusual features of the data
  - Check for stationarity
  - Do we need to transform the data?
- Eliminate trend
  - Estimating the trend and removing it from the series
  - Or, differencing the series (i.e., select  $d$  in the ARIMA model)
- Plot the sample ACF/PACF for the stationary component
  - Identify candidate values of  $p$  and  $q$



# Typical Steps for Modeling ARIMA Processes: Model Estimation

- Estimate the ARMA process parameters for the candidate models
- Check the goodness of fit: Are the time series residuals,  $\{r_t\}$  a sample of *i.i.d.* noise?
- Model selection:
  - Using information criteria such as AIC and AICC
  - Test model parameters to compare between the “full” model and the “subset” model

We need more assumptions to forecast  $\text{ARIMA}(p, d, q)$  processes. Let us start with the case of  $d = 1$ , i.e.,

$$\phi(B)(1 - B)X_t = \theta(B)Z_t,$$

where  $\{Z_t\} \sim \text{WN}(0, \sigma^2)$

- **Note:**  $Y_t = (1 - B)X_t = X_t - X_{t-1}$  is an  $\text{ARMA}(p, q)$  process
- We want to find the **best linear predictor** (BLP) of  $X_{n+1}$  based on  $X_0, X_1, \dots, X_n$ 
  - We know that  $X_{n+1} = X_n + Y_{n+1} \Rightarrow$  only need to figure out the BLP of  $Y_{n+1}$  based on  $\{X_0, Y_1, \dots, Y_n\}$
  - We need to know  $\mathbb{E}(X_0^2)$  and  $\mathbb{E}(X_0 Y_j)$  for  $j = 1, \dots, n + 1$

**Problem:** What is  $\mathbb{E}(X_0 Y_j)$ ?

- We **assume** that  $X_0$  is **uncorrelated** with  $Y_1, Y_2, \dots$
- Then the BLP of  $X_{n+1}$  based on  $\{X_0, X_1, \dots, X_n\}$  is the same as the BLP of  $X_{n+1}$  based on  $\{Y_1, Y_2, \dots, Y_n\}$
- This extends to ARIMA( $p, d, q$ ) processes:

If we **assume** that  $\{X_{1-d}, \dots, X_0\}$  is **uncorrelated** with  $Y_1, Y_2, \dots$ , then the BLP of  $Y_{n+1}$  based on  $\{X_{1-d}, \dots, X_0, \dots, X_n\}$  is the same as the BLP based on  $\{Y_1, Y_2, \dots, Y_n\}$

## Percentage Changes and Logarithms

Suppose  $X_t$  tends to have relatively stable **percentage changes** from one time period to the next. Specifically, assume that

$$X_t = (1 + Y_t)X_{t-1},$$

where  $100Y_t$  is the percentage change from  $X_{t-1}$  to  $X_t$ . Then

$$\log(X_t) - \log(X_{t-1}) = \log\left(\frac{X_t}{X_{t-1}}\right) = \log(1 + Y_t).$$

If  $Y_t$  is restricted to, say,  $|Y_t| < 0.2$  (ie., the percentage changes are at most  $\pm 20\%$ ), then, to a good approximation,  $\log(1 + Y_t) \approx Y_t$ . Consequently

$$\Delta[\log(X_t)] \approx Y_t$$

will be relatively stable and perhaps well-modeled by a stationary process.

**In financial literature, the differences of the (natural) logarithms are usually called **returns****

# Time Series Plots of Monthly US Electricity Production

