

# Lecture 4

## Completely Randomized Designs: Model, Estimation, Inference

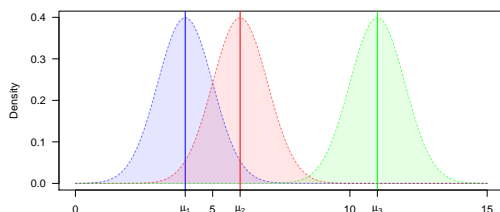
STAT 8050 Design and Analysis of Experiments  
January 21, 2020

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Clemson University

Notes

### Means Model

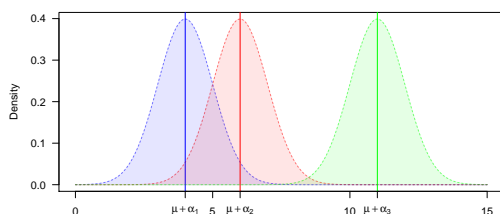
$$Y_{ij} = \mu_i + \epsilon_{ij}, \quad i = 1, \dots, g, \quad j = 1, \dots, n_i, \quad \epsilon_{ij} \sim N(0, \sigma^2)$$



Notes

### Effects Model

$$Y_{ij} = \mu + \alpha_i + \epsilon_{ij}, \quad i = 1, \dots, g, \quad j = 1, \dots, n_i, \quad \epsilon_{ij} \sim N(0, \sigma^2)$$

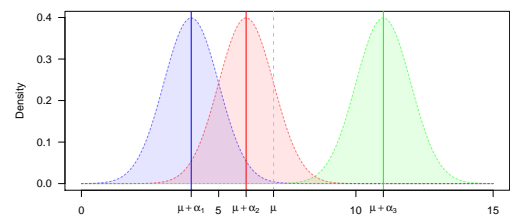


**Overparameterized.** Need to add a constraint so that the parameters are estimable.

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Effects Model Cont'd

Suppose we let  $\sum_{i=1}^g n_i \alpha_i = 0$



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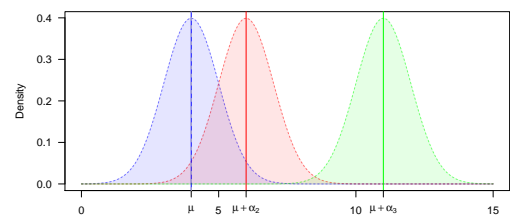
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Effects Model Cont'd

Suppose we let  $\alpha_1 = 0$



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Data Layout & the Dot Notation

$y_{ij}$  is the "observed" response for the  $j^{\text{th}}$  experimental unit to treatment  $i$ .

Treatment	Observations				Totals	Averages
1	$y_{11}$	$y_{12}$	$\cdots$	$y_{1n_1}$	$y_{1\cdot}$	$\bar{y}_{1\cdot}$
2	$y_{21}$	$y_{22}$	$\cdots$	$y_{2n_2}$	$y_{2\cdot}$	$\bar{y}_{2\cdot}$
$\vdots$	$\vdots$	$\vdots$	$\cdots$	$\vdots$	$\vdots$	$\vdots$
$g$	$y_{g1}$	$y_{g2}$	$\cdots$	$y_{gn_g}$	$y_{g\cdot}$	$\bar{y}_{g\cdot}$
				$y_{\cdot\cdot}$		$\bar{y}_{\cdot\cdot}$

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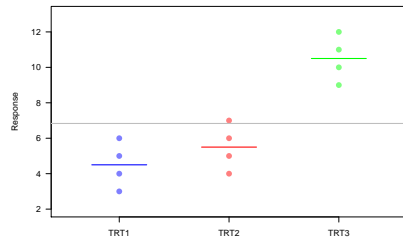
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ANOVA

Decomposition of  $y_{ij}$ :  $y_{ij} = \bar{y}_{..} + (\bar{y}_{i.} - \bar{y}_{..}) + (y_{ij} - \bar{y}_{i.})$

$$\Rightarrow \sum_{i=1}^g \sum_{j=1}^{n_i} (y_{ij} - \bar{y}_{..})^2 = \underbrace{\sum_{i=1}^g n_i (\bar{y}_{i.} - \bar{y}_{..})^2}_{SS_{TRT}} + \underbrace{\sum_{i=1}^g \sum_{j=1}^{n_i} (y_{ij} - \bar{y}_{i.})^2}_{SS_E}$$

$SS_T$                        $SS_{TRT}$                        $SS_E$



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ANOVA Table

Source	df	SS	MS	EMS
Treatment	$g - 1$	$SS_{TRT}$	$MS_{TRT} = \frac{SS_{TRT}}{g - 1}$	$\sigma^2 + \frac{\sum_{i=1}^g n_i \alpha_i^2}{g - 1}$
Error	$N - g$	$SS_E$	$MS_E = \frac{SS_E}{N - g}$	$\sigma^2$
Total	$N - 1$	$SS_T$		

$$SS_T = \sum_{i=1}^g \sum_{j=1}^{n_i} (y_{ij} - \bar{y}_{..})^2 = \sum_{i=1}^g \sum_{j=1}^{n_i} y_{ij}^2 - \frac{y_{..}^2}{N}$$
$$SS_{TRT} = \sum_{i=1}^g n_i (\bar{y}_{i.} - \bar{y}_{..})^2 = \sum_{i=1}^g \frac{y_{i.}^2}{n_i} - \frac{y_{..}^2}{N}$$
$$SS_E = \sum_{i=1}^g \sum_{j=1}^{n_i} (y_{ij} - \bar{y}_{i.})^2 = \sum_{i=1}^g \sum_{j=1}^{n_i} y_{ij}^2 - \sum_{i=1}^g \frac{y_{i.}^2}{n_i} = SS_T - SS_{TRT}$$

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F-Test

Testing for treatment effects

$$H_0 : \alpha_i = 0 \quad \text{for all } i$$
$$H_a : \alpha_i \neq 0 \quad \text{for some } i$$

**Test statistics:**  $F = \frac{MS_{TRT}}{MS_E}$ . Under  $H_0$ , the test statistic follows an F-distribution with  $g - 1$  and  $N - g$  degrees of freedom  
Reject  $H_0$  if

$$F_{obs} > F_{g-1, N-g; \alpha}$$

for an  $\alpha$ -level test,  $F_{g-1, N-g; \alpha}$  is the  $100 \times (1 - \alpha)\%$  percentile of a **central F-distribution** with  $g - 1$  and  $N - g$  degrees of freedom.

The **P-value** of the F-test is the probability of obtaining  $F$  at least as extreme as  $F_{obs}$ , that is,  $P(F > F_{obs}) \Rightarrow$  reject  $H_0$  if P-value  $< \alpha$ .

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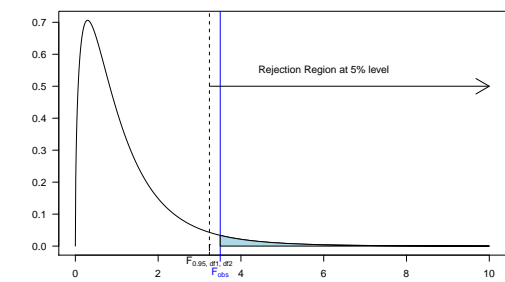
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F Distribution and the F-Test



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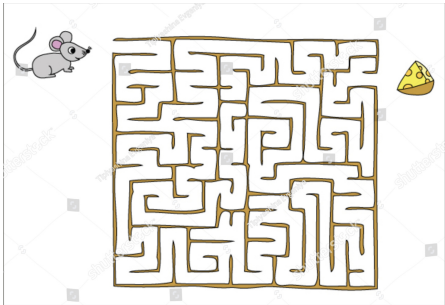
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Example



Source: <https://www.shutterstock.com/image-vector/find-your-way-cheese-mouse-maze-232569073>

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