

Lecture 5

Inferences about a Mean Vector

Readings: Zelterman, 2015, Chapters 5, 6, 7

DSA 8070 Multivariate Analysis

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Notes

Agenda

- 1 Confidence Intervals/Region for Population Means
- 2 Hypothesis Testing for Mean Vector
- 3 Multivariate Paired Hotelling's T-Square



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Overview

In this week we consider **estimation** and **inference** on population mean vector

We will explore the following questions:

- What is the sampling distribution of \bar{X}_n ?
- How to construct confidence intervals/region for population means
- How to conduct hypothesis testing for population means



Notes

Review: Sampling Distribution of Univariate Sample Mean \bar{X}_n

Suppose X_1, X_2, \dots, X_n is a random sample from a univariate population distribution with mean $\mathbb{E}(X) = \mu$ and variance $\text{Var}(X) = \sigma^2$. The sample mean \bar{X}_n is a function of random sample and therefore has a distribution

- $\bar{X}_n \sim N(\mu, \frac{\sigma^2}{n})$ when the sample size n is "sufficiently" large \Rightarrow This is the central limit theorem (CLT)
- The result above is exact if the population follows a normal distribution, i.e., $X \sim N(\mu, \sigma^2)$
- The standard error $\sqrt{\text{Var}(\bar{X}_n)} = \frac{\sigma}{\sqrt{n}}$ provides a measure estimation precision. In practice, we use $\frac{s}{\sqrt{n}}$ instead where s is the sample standard deviation

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Sampling Distribution of Multivariate Sample Mean Vector \bar{X}_n

Suppose X_1, X_2, \dots, X_n is a random sample from a multivariate population distribution with mean vector $\mathbb{E}(X) = \mu$ and covariance matrix Σ .

- $\bar{X}_n \sim N(\mu, \frac{1}{n}\Sigma)$ when the sample size n is "sufficiently" large \Rightarrow This is the multivariate version of CLT
- The result above is exact if the population follows a normal distribution, i.e., $X \sim N(\mu, \Sigma)$
- Again, the estimation precision improves with a larger sample size. Like the univariate case we would need to replace Σ by its estimate S , the sample covariacne matrix

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Review: Interval Estimation of Univariate Population Mean μ

The general format of a confidence interval (CI) estimate of a population mean is

Sample mean \pm multiplier \times standard error of mean.

For variable X , a CI estimate of its population mean μ is

$$\bar{X}_n \pm t_{n-1}(\frac{\alpha}{2}) \frac{s}{\sqrt{n}},$$

Here the multiplier value is a function of the confidence level, α , the sample size n

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Constructing Confidence Intervals for Mean Vector

We will still use the general recipe

Sample mean ± multiplier × standard error of mean.

The multiplier value also depends the strategy used for dealing with the multiple inference issue

- One at a Time CIs: a CI for μ_j is computed as

$$\bar{x}_j \pm t_{n-1}(\alpha/2) \frac{s_j}{\sqrt{n}}, \quad j = 1, \dots, p$$

- Bonferroni Method: a CI for μ_j is computed as

$$\bar{x}_j \pm t_{n-1}(\alpha/2p) \frac{s_j}{\sqrt{n}}, \quad j = 1, \dots, p$$

- Simultaneous CIs: a CI for μ_j is computed as

$$\bar{x}_j \pm \sqrt{\frac{(n-1)p}{n-p} F_{p,n-p}(\alpha)} \frac{s_j}{\sqrt{n}}, \quad j = 1, \dots, p$$

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Example: Mineral Content Measurements [source: Penn Stat Univ. STAT 505]

This example uses the dataset that includes mineral content measurements at two different arm bone locations for $n = 64$ women. We'll determine confidence intervals for the two different population means. Sample means and standard deviations for the two variables are:

Variable	Sample size	Mean	Std Dev
domradius (X_1)	$n = 64$	$\bar{x}_1 = 0.8438$	$s_1 = 0.1140$
domhumerus (X_2)	$n = 64$	$\bar{x}_2 = 1.7927$	$s_2 = 0.2835$

Let's apply the three methods we learned to construct 95% CIs

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Mineral Content Measurements Example Cont'd

- One at a Time CIs: $\bar{x}_j \pm t_{n-1}(\alpha/2) \frac{s_j}{\sqrt{n}}, \quad j = 1, \dots, p$.
Therefore 95% CIs for μ_1 and μ_2 are:

$$\mu_1 : 0.8438 \pm \underbrace{1.998}_{t_{63}(0.025)} \times \frac{0.1140}{\sqrt{64}} = [0.815, 0.872]$$

$$\mu_2 : 1.7927 \pm 1.998 \times \frac{0.2835}{\sqrt{64}} = [1.722, 1.864]$$

- Bonferroni Method:

$$\bar{x}_j \pm t_{n-1}(\alpha/2p) \frac{s_j}{\sqrt{n}}, \quad j = 1, \dots, p.$$

$$\mu_1 : 0.8438 \pm \underbrace{2.296}_{t_{63}(0.025)} \times \frac{0.1140}{\sqrt{64}} = [0.811, 0.877]$$

$$\mu_2 : 1.7927 \pm 2.296 \times \frac{0.2835}{\sqrt{64}} = [1.711, 1.874]$$

- Simultaneous CIs:

$$\bar{x}_j \pm \sqrt{\frac{(n-1)p}{n-p} F_{p,n-p}(\alpha)} \frac{s_j}{\sqrt{n}}, \quad j = 1, \dots, p$$

$$\mu_1 : 0.8438 \pm 2.528 \times \frac{0.1140}{\sqrt{64}} = [0.808, 0.880]$$

$$\mu_2 : 1.7927 \pm 2.528 \times \frac{0.2835}{\sqrt{64}} = [1.703, 1.882]$$

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Confidence Intervals/Region for Population Means

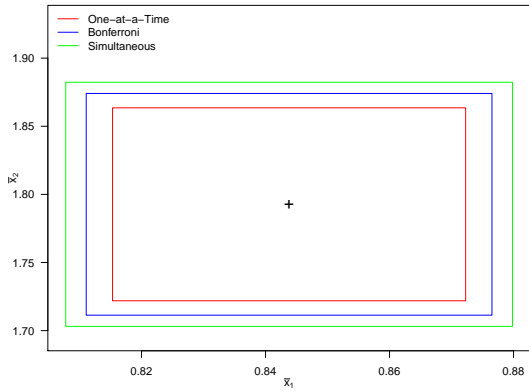
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95 % CIs Based on Three Methods

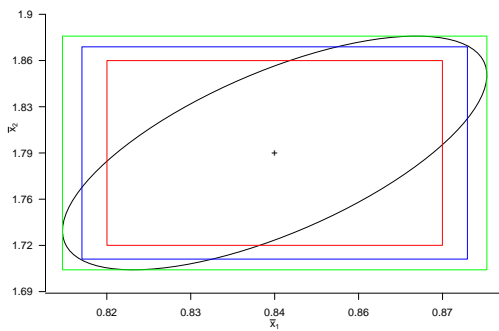


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Confidence Ellipsoid

A confidence ellipsoid for μ is the set of μ satisfying

$$n(\bar{X}_n - \mu)^T S^{-1}(\bar{X}_n - \mu) \leq \frac{(n-1)p}{n-p} F_{p, n-p}(\alpha)$$



Notes

Hypothesis Testing for Mean

- Recall: for univariate data, t statistic

$$t = \frac{\bar{X}_n - \mu_0}{s/\sqrt{n}} \Rightarrow t^2 = \frac{(\bar{X}_n - \mu_0)^2}{s^2/n} = n(\bar{X}_n - \mu_0)(s^2)^{-1}(\bar{X}_n - \mu_0)$$

Under $H_0 : \mu = \mu_0$

$$t \sim t_{n-1}, \quad t^2 \sim F_{1, n-1}$$

- Extending to multivariate by analogy:

$$T^2 = n(\bar{X}_n - \mu_0)^T S^{-1}(\bar{X}_n - \mu_0)$$

Under $H_0 : \mu = \mu_0$

$$\frac{(n-p)}{(n-1)p} T^2 \sim F_{p, n-p}$$

Note: T^2 here is the so-called **Hotelling's T-Square**

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Hypothesis Testing for Mean Vector μ

- 1 State the null
$$H_0 : \mu = \mu_0$$

and the alternative
$$H_a : \mu \neq \mu_0$$
- 2 Compute the test statistic
$$F = \frac{n-p}{(n-1)p} n (\bar{X}_n - \mu_0)^T S^{-1} (\bar{X}_n - \mu_0)$$
- 3 Compute the P-value. Under $H_0 : F \sim F_{p,n-p}$
- 4 Draw a conclusion: We do (or do not) have enough statistical evidence to conclude $\mu \neq \mu_0$ at α significant level

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Example: Women's Dietary Intake [source: Penn Stat Univ. STAT 505]

The recommended intake and a sample mean for all women between 25 and 50 years old are given below:

Variable	Recommended Intake (μ_0)	Sample Mean (\bar{x}_n)
Calcium	1000 mg	624.0 mg
Iron	15 mg	11.1 mg
Protein	60 g	65.8 g
Vitamin A	800 μ g	839.6 μ g
Vitamin C	75 mg	78.9 mg

Here we would like to test, at $\alpha = 0.01$ level, if the $\mu = \mu_0$

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Women's Dietary Intake Example Analysis

- 1 State the null
$$H_0 : \mu = \mu_0$$

and the alternative
$$H_a : \mu \neq \mu_0$$
- 2 Compute the test statistic
$$F = \frac{n-p}{(n-1)p} n (\bar{x}_n - \mu_0)^T S^{-1} (\bar{x}_n - \mu_0) = 349.80$$
- 3 Compute the P-value. Under $H_0 : F \sim F_{p,n-p} \Rightarrow$ p-value
$$= \Pr(F_{p,n-p} > 349.80) = 3 \times 10^{-191} < \alpha = 0.01$$
- 4 Draw a conclusion: We do have enough statistical evidence to conclude $\mu \neq \mu_0$ at α significant level

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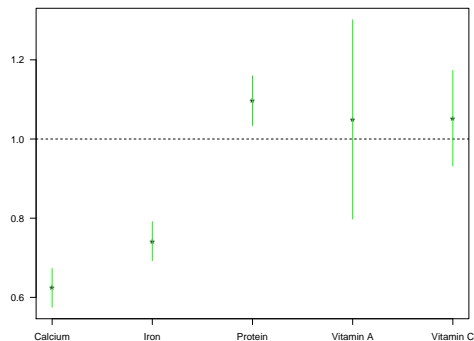
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Profile Plots

- 1 Standardize each of the observations by dividing their hypothesized means
- 2 Plot either simultaneous or Bonferroni CIs for the population mean of these standardized variables



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Spouse Survey Data Example

A sample ($n = 30$) of husband and wife pairs are asked to respond to each of the following questions:

- 1 What is the level of passionate love you feel for your partner?
- 2 What is the level of passionate love your partner feels for you?
- 3 What is the level of companionate love you feel for your partner?
- 4 What is the level of companionate love your partner feels for you?

Responses were recorded on a typical five-point scale: 1) None at all 2) Very little 3) Some 4) A great deal 5) Tremendous amount.

We will try to address the following question: Do the husbands respond to the questions in the same way as their wives?



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Multivariate Paired Hotelling's T-Square

Let X_F and X_M be the responses to these 4 questions for females and males, respectively. Here the quantities of interest are $\mathbb{E}(D) = \mu_D$, the average differences across all husband and wife pairs.

- 1 State the null $H_0 : \mu_D = 0$ and the alternative hypotheses $H_a : \mu_D \neq 0$
- 2 Compute the test statistic

$$F = \frac{n-p}{(n-1)p} n \bar{D}_n^T S_D^{-1} \bar{D}_n$$

- 3 **Compute the P-value.** Under $H_0 : F \sim F_{p, n-p}$
- 4 **Draw a conclusion:** We do (or do not) have enough statistical evidence to conclude $\mu_D \neq 0$ at α significant level



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Spouse Survey Data Example Analysis

- 1 State the null

$H_0 : \mu_D = \mathbf{0}$

and the alternative

$H_a : \mu_D \neq \mathbf{0}$

- 2 Compute the test statistic

$$F = \frac{n-p}{(n-1)p} n \bar{D}_n^T S_D^{-1} \bar{D}_n = 2.942$$

- 3 Compute the P-value. Under $H_0 : F \sim F_{p,n-p} \Rightarrow$
p-value = $\mathbb{P}_\mathbb{R}(F_{p,n-p} >) = 0.0394 < \alpha = 0.05$

- 4 Draw a conclusion: We do have enough statistical evidence to conclude $\mu_D \neq \mathbf{0}$ at 0.05 significant level

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