Lecture 15

Interpolation of Spatial Data II

DSA 8020 Statistical Methods II

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Outline

1. **Review: Spatial Interpolation**
2. **Parameter estimation**
3. **A Case Study of Paraná State Precipitation Data**

Conditional Distribution of Multivariate Normal

If

$$\begin{pmatrix} \boldsymbol{Y}_1 \\ \boldsymbol{Y}_2 \end{pmatrix} \sim \mathrm{N}\left(\begin{pmatrix} \boldsymbol{\mu_1} \\ \boldsymbol{\mu_2} \end{pmatrix}, \begin{pmatrix} \Sigma_{11} & \Sigma_{12} \\ \Sigma_{21} & \Sigma_{22} \end{pmatrix} \right)$$

Then

$$[\boldsymbol{Y}_1 | \boldsymbol{Y}_2 = \boldsymbol{y}_2] \sim \mathrm{N}\left(\boldsymbol{\mu_{1|2}}, \Sigma_{1|2}\right)$$

where

$$\boldsymbol{\mu_{1|2}} = \boldsymbol{\mu_1} + \Sigma_{12}\Sigma_{22}^{-1}\left(\boldsymbol{y}_2 - \boldsymbol{\mu_2}\right)$$
$$\Sigma_{1|2} = \Sigma_{11} - \Sigma_{12}\Sigma_{22}^{-1}\Sigma_{21}$$

GP-Based Spatial Interpolation: Kriging

If $\{Y(\boldsymbol{s})\}_{\boldsymbol{s}\in\mathcal{S}}$ follows a GP, then

$$\begin{pmatrix} Y_0 \\ \boldsymbol{Y} \end{pmatrix} \sim \mathrm{N}\left(\begin{pmatrix} m_0 \\ \boldsymbol{m} \end{pmatrix}, \begin{pmatrix} \sigma_0^2 & k^{\mathrm{T}} \\ k & \Sigma \end{pmatrix}\right)$$

We have

$$[Y_0|\boldsymbol{Y} = \boldsymbol{y}] \sim \mathrm{N}\left(m_{Y_0|\boldsymbol{Y}=\boldsymbol{y}}, \sigma^2_{Y_0|\boldsymbol{Y}=\boldsymbol{y}}\right)$$

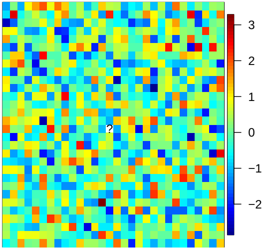
where

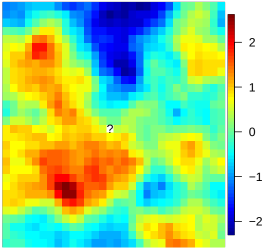
$$\begin{aligned} m_{Y_0|\boldsymbol{Y}=\boldsymbol{y}} &= m_0 + k^{\mathrm{T}}\Sigma^{-1}(\boldsymbol{y}-\boldsymbol{m}) \\ \sigma^2_{Y_0|\boldsymbol{Y}=\boldsymbol{y}} &= \sigma_0^2 - k^{\mathrm{T}}\Sigma^{-1}k \end{aligned}$$

Next, we are going to revisit our toy examples

Toy Examples Revisited

For simplicity, we assume $m(s) = 0$ for $s \in \mathcal{S}$, the spatial covariance only depends on distance





$$m_{Y_0|\boldsymbol{Y}=\boldsymbol{y}} = 0 + k^{\mathrm{T}}\Sigma^{-1}\left(\boldsymbol{y} - \boldsymbol{0}\right), \quad \sigma^2_{Y_0|\boldsymbol{Y}=\boldsymbol{y}} = \sigma_0^2 - k^{\mathrm{T}}\Sigma^{-1}k$$

Spatial uncorrelated field:

- $m_{Y_0|\boldsymbol{Y}} = 0$
- $\sigma^2_{Y_0|\boldsymbol{Y}=\boldsymbol{y}} = \sigma_0^2$

Spatial correlated field:

- $m_{Y_0|\boldsymbol{Y}} = k^{\mathrm{T}}\Sigma^{-1}\boldsymbol{y}$
- $\sigma^2_{Y_0|\boldsymbol{Y}=\boldsymbol{y}} = \sigma_0^2 - k^{\mathrm{T}}\Sigma^{-1}k$

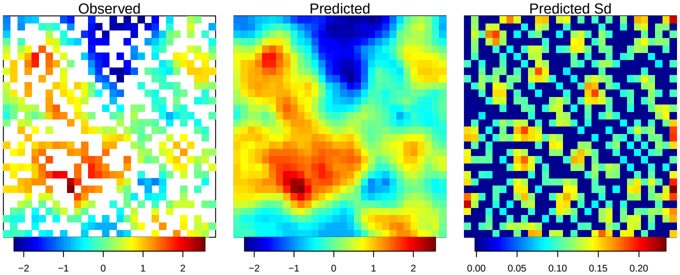
Interpolating Multiple Points in Space

In practice, we would like to predict the values at many locations. The Gaussian conditional distribution formula can still be used:

$$[\boldsymbol{Y}_0 | \boldsymbol{Y} = \boldsymbol{y}] \sim \mathrm{N}\left(\boldsymbol{m}_{\boldsymbol{Y}_0|\boldsymbol{Y}=\boldsymbol{y}}, \Sigma_{\boldsymbol{Y}_0|\boldsymbol{Y}=\boldsymbol{y}}\right)$$

where

$$\boldsymbol{m}_{\boldsymbol{Y}_0|\boldsymbol{Y}=\boldsymbol{y}} = \boldsymbol{m}_0 + \boldsymbol{k}^{\mathrm{T}} \Sigma^{-1} (\boldsymbol{y} - \boldsymbol{m})$$
$$\Sigma_{\boldsymbol{Y}_0|\boldsymbol{Y}=\boldsymbol{y}} = \Sigma_0 - \boldsymbol{k}^{\mathrm{T}} \Sigma^{-1} \boldsymbol{k}$$



GP-Based Spatial Interpolation: Kriging

If $\{Y(\boldsymbol{s})\}_{\boldsymbol{s}\in\mathcal{S}}$ follows a GP, then

$$\begin{pmatrix} \boldsymbol{Y}_0 \\ \boldsymbol{Y} \end{pmatrix} \sim \mathrm{N}\left(\begin{pmatrix} \boldsymbol{m}_0 \\ \boldsymbol{m} \end{pmatrix}, \begin{pmatrix} \Sigma_0 & \boldsymbol{k}^{\mathrm{T}} \\ \boldsymbol{k} & \Sigma \end{pmatrix} \right)$$

We have

$$[\boldsymbol{Y}_0 | \boldsymbol{Y} = \boldsymbol{y}] \sim \mathrm{N}\left(\boldsymbol{m}_{\boldsymbol{Y}_0|\boldsymbol{Y}=\boldsymbol{y}}, \Sigma_{\boldsymbol{Y}_0|\boldsymbol{Y}=\boldsymbol{y}}\right)$$

where

$$\begin{aligned} \boldsymbol{m}_{\boldsymbol{Y}_0|\boldsymbol{Y}=\boldsymbol{y}} &= \boldsymbol{m}_0 + \boldsymbol{k}^{\mathrm{T}}\Sigma^{-1}(\boldsymbol{y} - \boldsymbol{m}) \\ \Sigma_{\boldsymbol{Y}_0|\boldsymbol{Y}=\boldsymbol{y}} &= \Sigma_0 - \boldsymbol{k}^{\mathrm{T}}\Sigma^{-1}\boldsymbol{k} \end{aligned}$$

Question: what if we don't know $m(s;\boldsymbol{\beta}), c(h;\boldsymbol{\theta})$?

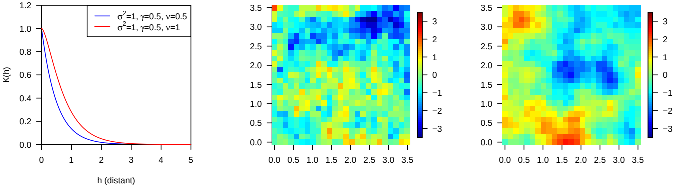
$\Rightarrow$ We need to estimate the mean and covariance from the data $\boldsymbol{y}$.

Outline

Recap: Gaussian Process

We assume that the observed data $\{y(\boldsymbol{s}_i)\}_{i=1}^{n}$ is one partial realization of a (continuously indexed) spatial stochastic process $\{Y(\boldsymbol{s})\}_{\boldsymbol{s}\in\mathcal{S}}$.

- Gaussian Processes $\mathrm{GP}\left(m\left(\cdot\right), K\left(\cdot,\cdot\right)\right)$ are widely used in modeling spatial stochastic processes
- Spatial statisticians often focus on the covariance function. e.g. $K(h) = \sigma^2 \frac{\left(\sqrt{2\nu}h/\gamma\right)^{\nu} \mathcal{K}_\nu\left(\sqrt{2\nu}h/\gamma\right)}{\Gamma(\nu)2^{\nu-1}}$



Variogram, Semivariogram, and Covariance Function

Under the stationary and isotropic assumptions

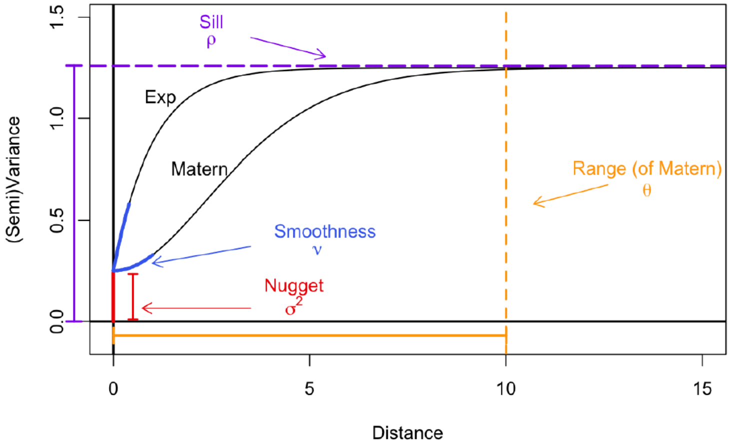
Variogram:

$$
\begin{aligned}
2\gamma(\boldsymbol{s}_i, \boldsymbol{s}_j) &= \mathrm{Var}\left(Y(\boldsymbol{s}_i) - Y(\boldsymbol{s}_j)\right) \\
&= \mathrm{E}\left\{\left((Y(\boldsymbol{s}_i) - \mu(\boldsymbol{s}_i)) - (Y(\boldsymbol{s}_j) - \mu(\boldsymbol{s}_j))\right)^2\right\} \\
&= \mathrm{E}\left\{\left(Y(\boldsymbol{s}_i) - Y(\boldsymbol{s}_j)\right)^2\right\} \\
&= 2\gamma(\|\boldsymbol{s}_i - \boldsymbol{s}_j\|) = 2\gamma(h)
\end{aligned}
$$

Semivariogram and covariance function:

$$
\gamma(h) = C(0) - C(h)
$$

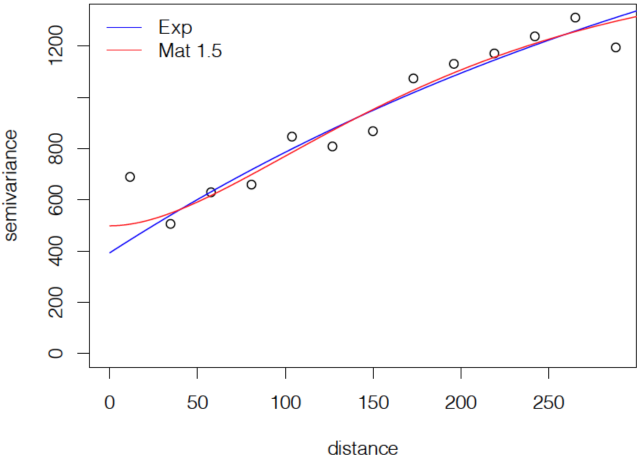
Semivariogram $\left\{\frac{1}{2}\mathrm{Var}\left(\varepsilon\left(\boldsymbol{s}_i\right)-\varepsilon\left(\boldsymbol{s}_j\right)\right)\right\}_{i,j}$



Source: `fields` vignette by Wiens and Krock, 2019

Estimation: Weighted Least Squares Method

$$\underset{\boldsymbol{\theta}}{\text{argmin}} \sum_{u \in \mathcal{U}} \frac{N(h_u)}{[\gamma(h_u;\boldsymbol{\theta})]^2} \left[\hat{\gamma}\left(h_u\right) - \gamma\left(h_u;\boldsymbol{\theta}\right)\right]^2$$



Maximum Likelihood Estimation (MLE)

Log-likelihood:

Given data $\boldsymbol{y} = (y(\boldsymbol{s}_1), \cdots, y(\boldsymbol{s}_n))^{\mathrm{T}}$

$$\ell_n(\boldsymbol{\beta}, \boldsymbol{\theta}; \boldsymbol{y}) \propto -\frac{1}{2}\log|\boldsymbol{\Sigma}_{\boldsymbol{\theta}}| - \frac{1}{2}(\boldsymbol{y} - \boldsymbol{X}^{\mathrm{T}}\boldsymbol{\beta})^{\mathrm{T}}[\boldsymbol{\Sigma}_{\boldsymbol{\theta}}]^{-1}_{n\times n}(\boldsymbol{y} - \boldsymbol{X}^{\mathrm{T}}\boldsymbol{\beta})$$

where $\boldsymbol{\Sigma}_{\boldsymbol{\theta}}(i,j) = \sigma^2\rho_{\rho,\nu}(\|\boldsymbol{s}_i - \boldsymbol{s}_j\|) + \tau^2 1_{\{\boldsymbol{s}_i=\boldsymbol{s}_j\}}, i,j = 1,\cdots,n$

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where $\boldsymbol{\Sigma}_{\boldsymbol{\theta}}(i,j) = \sigma^2\rho_{\rho,\nu}(\|\boldsymbol{s}_i - \boldsymbol{s}_j\|) + \tau^2 1_{\{\boldsymbol{s}_i = \boldsymbol{s}_j\}}, i, j = 1, \cdots, n$

for any fixed $\boldsymbol{\theta}_0 \in \Theta$ the unique value of $\boldsymbol{\beta}$ that maximizes ℓ_n is given by

$$\hat{\boldsymbol{\beta}} = \left(\boldsymbol{X}^{\mathrm{T}}\Sigma^{-1}_{\boldsymbol{\theta}_0}\boldsymbol{X}\right)^{-1}\boldsymbol{X}^{\mathrm{T}}\Sigma_{\boldsymbol{\theta}_0}\boldsymbol{y}$$

Then we obtain the profile log likelihood

$$\ell_n(\boldsymbol{\theta}; \boldsymbol{y}) \propto -\frac{1}{2}\log|\boldsymbol{\Sigma}_{\boldsymbol{\theta}}| - \frac{1}{2}\boldsymbol{y}^{\mathrm{T}}P(\boldsymbol{\theta})\boldsymbol{y}$$

where

$$P(\boldsymbol{\theta}) = \Sigma^{-1}_{\boldsymbol{\theta}} - \Sigma^{-1}_{\boldsymbol{\theta}}\boldsymbol{X}\left(\boldsymbol{X}^{\mathrm{T}}\Sigma^{-1}_{\boldsymbol{\theta}}\boldsymbol{X}\right)^{-1}\boldsymbol{X}^{\mathrm{T}}\Sigma_{\boldsymbol{\theta}}$$

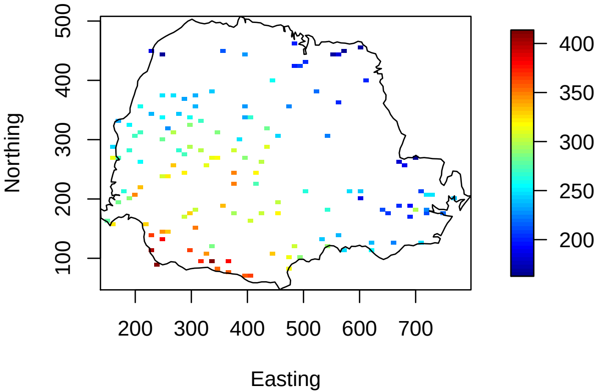
Solve the maximization problem above to get the MLE

Remarks on Likelihood-based estimation

- Maximizing $\ell_n(\boldsymbol{\theta}; \boldsymbol{y})$ is a constrained nonlinear optimization problem $\Rightarrow$ ML estimates must be obtained numerically
- Restricted (or residual) maximum likelihood (REML) can be used instead

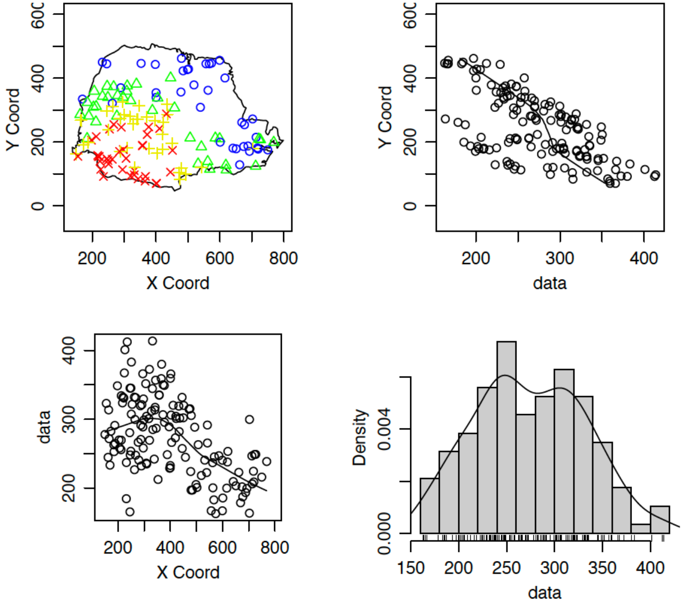
Paraná State Precipitation Data

We look at the average winter (May-June, dry season) rainfall at 143 locations throughout Paraná, Brazil

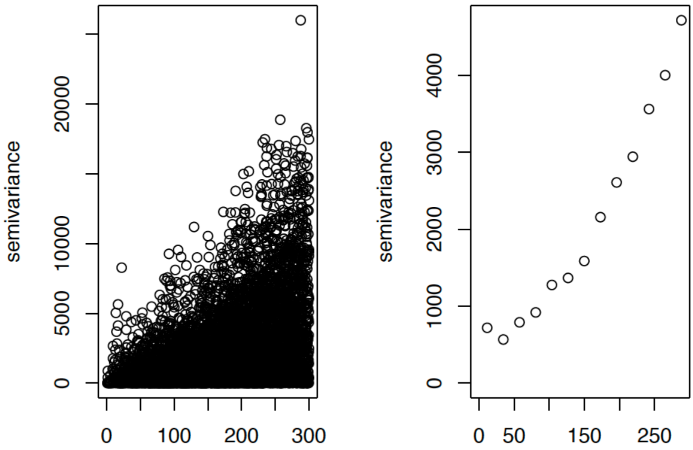


Goal: To interpolate the values in the spatial domain

Exploratory Data Analysis



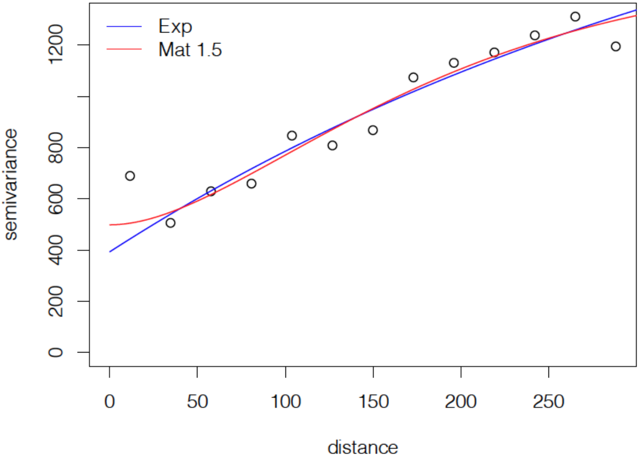
Variogram Analysis



An increasing variogram suggests spatial dependence structure

Estimating Spatial Covariance Via Varogram

```
parana.vtfit.exp <- variofit(parana.variot)
parana.vtfit.mat1.5  <- variofit(parana.variot, kappa = 1.5)
```



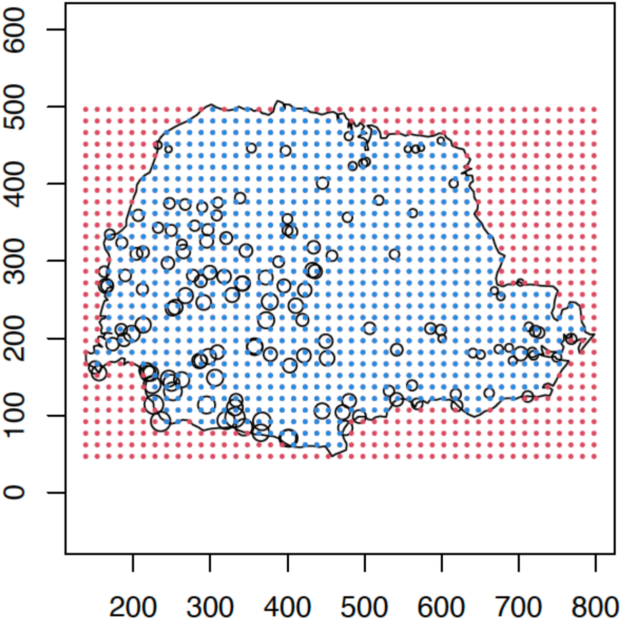
Maximum Likelihood Estimation of Paraná Rainfall

```
(parana.ml1 <- likfit(parana, trend = "1st", ini = c(1000, 50), nug = 100))
```

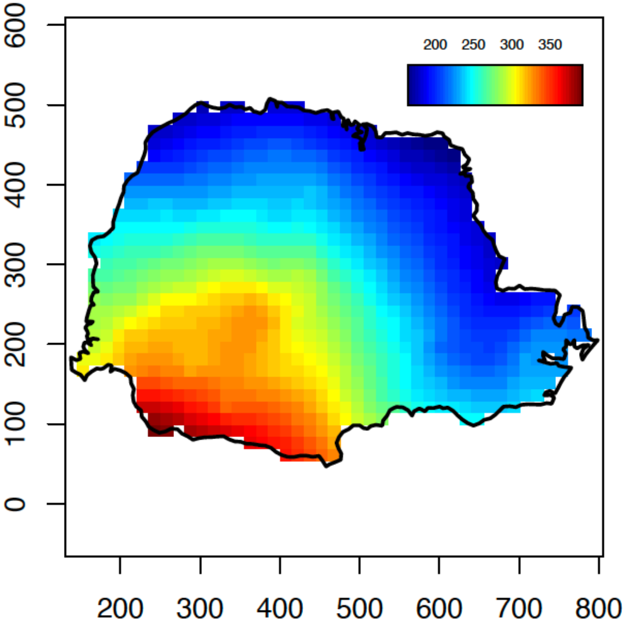
```
## ---------------------------------------------------------------
## likfit: likelihood maximisation using the function optim.
## likfit: Use control() to pass additional
##          arguments for the maximisation function.
##         For further details see documentation for optim.
## likfit: It is highly advisable to run this function several
##         times with different initial values for the parameters.
## likfit: WARNING: This step can be time demanding!
## ---------------------------------------------------------------
## likfit: end of numerical maximisation.

## likfit: estimated model parameters:
##      beta0      beta1      beta2      tausq    sigmasq        phi
## "416.4984" " -0.1375" " -0.3997" "385.5180" "785.6904" "184.3863"
## Practical Range with cor=0.05 for asymptotic range: 552.3719
##
## likfit: maximised log-likelihood = -663.9
```

Next, we will use these information to conduct spatial interpolation

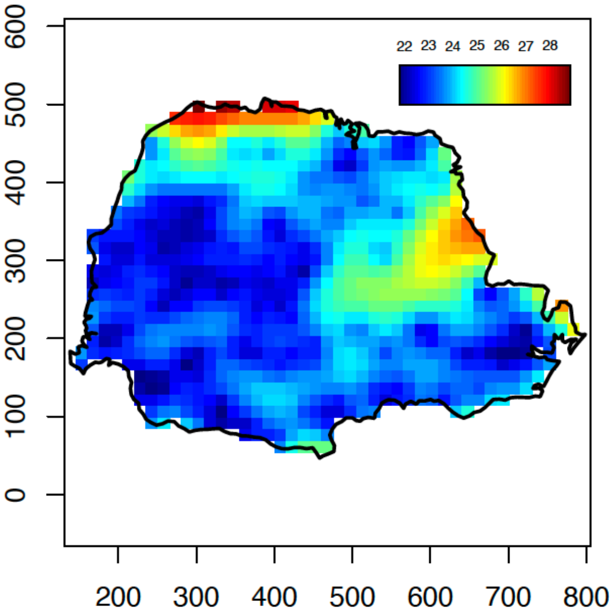
Setting Up the Spatial Grids for Prediction



Spatial Predicted Map



Prediction Uncertainty Map



Summary

This slides cover:

- Parameter Estimation for Gaussian Process Spatial Models
- Spatial predictions using Gaussian Process Spatial Models