

# Extreme Value Analysis Part I: Univariate Extreme Value Analysis

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# Agenda

Motivation

Extreme Value Theorem & Block Maxima Method

Peaks–Over–Threshold (POT) Method

Non-stationary Extreme Value Analysis

# Outline

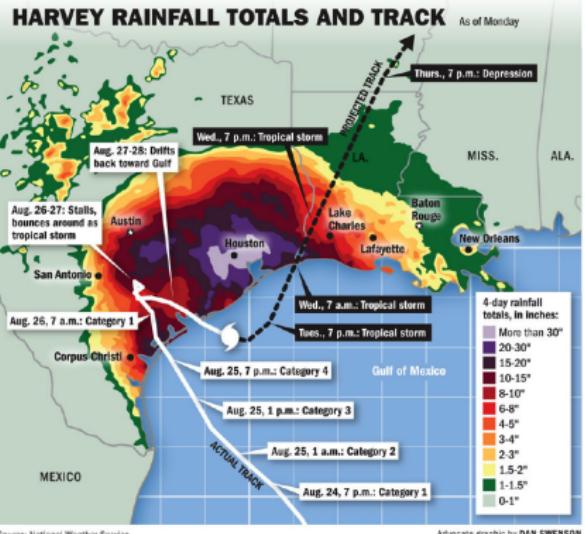
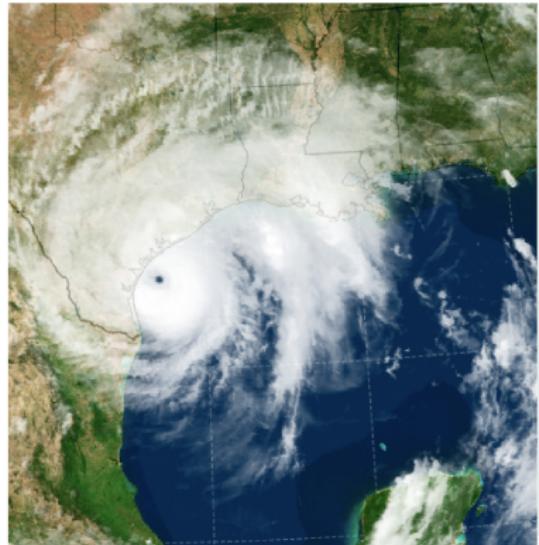
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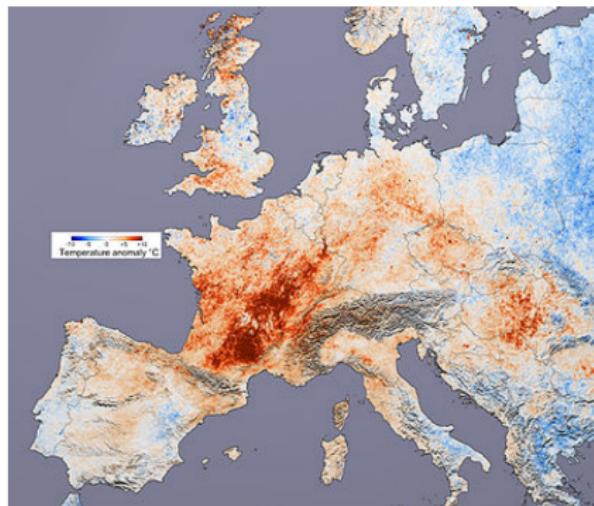
# Extreme Rainfall During Hurricane Harvey



Source: NASA (Left); National Weather Service (Right)

- ▶ “*A storm forces Houston, the limitless city, to consider its limits*” – The New York Times (8.31.17)

# Environmental Extremes: Heatwaves, Storm Surges, etc.



- ▶ **Heat wave:** The 2003 European heat wave led to the hottest summer on record in Europe since 1540 that resulted in at least **30,000 deaths**
- ▶ **Storm Surge:** Hurricane Katrina produced the highest storm surge ever recorded (**27.8 feet**) on the U.S. coast

# Why Study Extremes?

Although infrequent, extremes usually have large impact.

**Goal:** to quantify the tail behavior  $\Rightarrow$  often requires extrapolation.

## Applications:

- ▶ Hydrology: flooding
- ▶ Climate: temperature, precipitation, wind, ...
- ▶ Engineering: structural design, reliability
- ▶ Finance, Insurance/reinsurance

## Some Scientific Questions

- ▶ How to estimate the magnitude of extreme events (e.g. 100-year rainfall)?
- ▶ How extremes vary in space?
- ▶ How extremes may change in future climate conditions?

# Brief History

- ▶ 1920's: Foundations of asymptotic argument developed by Fisher and Tippett
- ▶ 1940's: Asymptotic theory unified and extended by Gnedenko and von Mises
- ▶ 1950's: Use of asymptotic distributions for statistical modelling by Gumbel and Jenkinson
- ▶ 1970's: Classic limit laws generalized by Pickands
- ▶ 1980's: Leadbetter (and others) extend theory to stationary processes
- ▶ 1990's: Multivariate and other techniques explored as a means to improve inference
- ▶ 2000's: Interest in spatial and spatio-temporal applications, and in finance

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## Usual vs Extremes

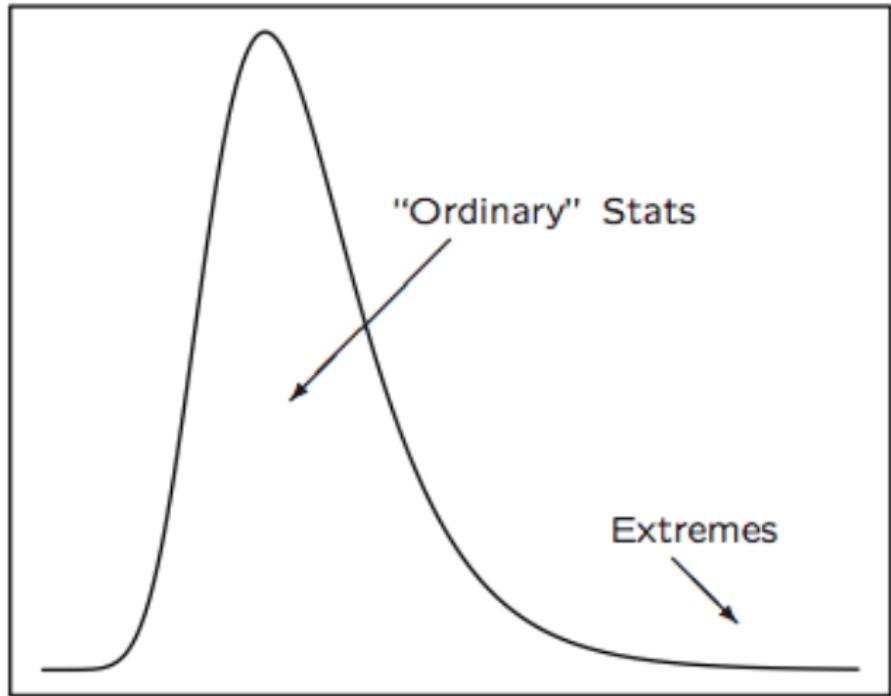


Figure courtesy of Dan Cooley

# Probability Framework

Let  $X_1, \dots, X_n \stackrel{i.i.d.}{\sim} F$  and define  $M_n = \max\{X_1, \dots, X_n\}$   
Then the distribution function of  $M_n$  is

$$\begin{aligned}\mathbb{P}(M_n \leq x) &= \mathbb{P}(X_1 \leq x, \dots, X_n \leq x) \\ &= \mathbb{P}(X_1 \leq x) \times \dots \times \mathbb{P}(X_n \leq x) = F^n(x)\end{aligned}$$

## Remark

$$F^n(x) \xrightarrow{n \rightarrow \infty} \begin{cases} 0 & \text{if } F(x) < 1 \\ 1 & \text{if } F(x) = 1 \end{cases}$$

⇒ the limiting distribution is degenerate.

# Asymptotic: Classical Limit Laws

Recall the **Central Limit Theorem**:

$$\frac{S_n - n\mu}{\sqrt{n}\sigma} \xrightarrow{d} N(0, 1),$$

where  $S_n = \sum_{i=1}^n X_i$

⇒ rescaling is the key to obtain a non-degenerate distribution

**Question:** Can we get the limiting distribution of

$$\frac{M_n - b_n}{a_n}$$

for suitable sequence  $\{a_n\} > 0$  and  $\{b_n\}$ ?

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## CLT in Action

1. Generate 100 ( $n$ ) random numbers from an Exponential distribution (population distribution)
2. Compute the **sample mean** of these 100 random numbers
3. Repeat this process 120 times

## Extremal Types Theorem (Fisher–Tippett 1928, Gnedenko 1943)

Define  $M_n = \max\{X_1, \dots, X_n\}$  where  $X_1, \dots, X_n \stackrel{\text{i.i.d.}}{\sim} F$ . If  $\exists a_n > 0$  and  $b_n \in \mathbb{R}$  such that, as  $n \rightarrow \infty$ , if

$$\mathbb{P}\left(\frac{M_n - b_n}{a_n} \leq x\right) \xrightarrow{d} G(x)$$

then  $G$  must be the same type of the following form:

$$G(x; \mu, \sigma, \xi) = \exp\left\{-\left[1 + \xi\left(\frac{x - \mu}{\sigma}\right)\right]_+^{-\frac{1}{\xi}}\right\}$$

where  $x_+ = \max(x, 0)$  and  $G(x)$  is the distribution function of the **generalized extreme value distribution (GEV( $\mu, \sigma, \xi$ )**)

- ▶  $\mu$  and  $\sigma$  are location and scale parameters
- ▶  $\xi$  is a shape parameter determining the rate of tail decay, with

•  $\xi < 0$ : Weibull distribution

•  $\xi = 0$ : Gumbel distribution

•  $\xi > 0$ : Fréchet distribution

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## Example: Exponential Maxima

Let  $X \sim \text{Exp}(\lambda = 1)$ . Set  $a_n = 1$ ,  $b_n = \log(n)$ . We want to show  $\frac{M_n - b_n}{a_n}$  converges to a GEV distribution, where  $M_n = \max_{i=1}^n X_i$ .

$$\begin{aligned}\mathbb{P}\left(\frac{M_n - b_n}{a_n} \leq x\right) &= \mathbb{P}(M_n \leq a_n x + b_n) \\ &= \mathbb{P}(M_n \leq x + \log(n)) \\ &= (1 - \exp(-x - \log(n)))^n \\ &= (1 - \frac{1}{n} \exp(-x))^n \\ &\xrightarrow{n \rightarrow \infty} \exp(-\exp(x))\end{aligned}$$

It is the cdf of the **standard Gumbel** distribution

## Extremal Types Theorem in Action

1. Generate 100 ( $n$ ) random numbers from an Exponential distribution (population distribution)
2. Compute the **sample maximum** of these 100 random numbers
3. Repeat this process 120 times

# Max-Stability and GEV

## Definition

A distribution  $G$  is said to be **max-stable** if

$$G^k(a_k x + b_k) = G(x), \quad k \in \mathbb{N}$$

for some constants  $a_k > 0$  and  $b_k$

- ▶ Taking powers of a distribution function results only in a change of location and scale
- ▶ A distribution is **max-stable**  $\iff$  it is a **GEV** distribution

# Quantiles and Return Levels

## ► Quantiles of GEV

$$G(m_p) = \exp \left\{ - \left[ 1 + \xi \left( \frac{m_p - \mu}{\sigma} \right) \right]_+^{-\frac{1}{\xi}} \right\} = 1 - p$$
$$\Rightarrow m_p = \mu - \frac{\sigma}{\xi} \left[ 1 - \{-\log(1-p)^{-\xi}\} \right] \quad 0 < p < 1$$

- In the extreme value terminology,  $m_p$  is the **return level** associated with the **return period**  $\frac{1}{p}$

# Statistical Practice

Assume  $n$  is large enough so that

$$\begin{aligned}\mathbb{P}\left(\frac{M_n - b_n}{a_n} \leq x\right) &\approx \exp\left(-[1 + \xi x]^{-1/\xi}\right) \\ \Rightarrow \mathbb{P}(M_n \leq y) &\approx \exp\left(-\left[1 + \xi \left(\frac{y - b_n}{a_n}\right)^{-1/\xi}\right]\right) \\ &= \exp\left(-\left[1 + \xi \left(\frac{y - \mu}{\sigma}\right)\right]^{-1/\xi}\right)\end{aligned}$$

Then, we have a three-parameter estimation problem.  $\mu, \sigma, \xi$  can be estimated via **maximum likelihood**

## Maximum Likelihood Estimation

Let  $M_1, \dots, M_k \stackrel{\text{iid}}{\sim} \text{GEV}$ , then log-likelihood for  $(\mu, \sigma, \xi)$  when  $\xi \neq 0$  is

$$\begin{aligned}\ell(\mu, \sigma, \xi) = & -m \log \sigma - (1 + 1/\xi) \sum_{i=1}^k \log \left[ 1 + \xi \left( \frac{m_i - \mu}{\sigma} \right) \right] \\ & - \sum_{i=1}^k \left[ 1 + \xi \left( \frac{m_i - \mu}{\sigma} \right) \right]^{-1/\xi},\end{aligned}$$

provided that  $1 + \xi(\frac{m_i - \mu}{\sigma}) > 0$ , for  $i = 1, \dots, k$ .

When  $\xi = 0 \rightarrow$  use the Gumbel limit of the GEV

Maximum likelihood estimate (MLE) is obtained by (numerically) maximization of log-likelihood shown above

## Non-Regular Models [Smith, 1985]

- ▶ When  $\xi > -0.5$ , maximum likelihood estimators are regular, in the sense of having the usual asymptotic properties
- ▶ When  $-1 < \xi < -0.5$ , maximum likelihood estimators are generally obtainable, but do not have the standard asymptotic properties
- ▶ When  $\xi < -1$ , maximum likelihood estimators are unlikely to be obtainable

In most environmental problems  $\hat{\xi} \in [-0.5, 0.5]$  so maximum likelihood appears to work fine

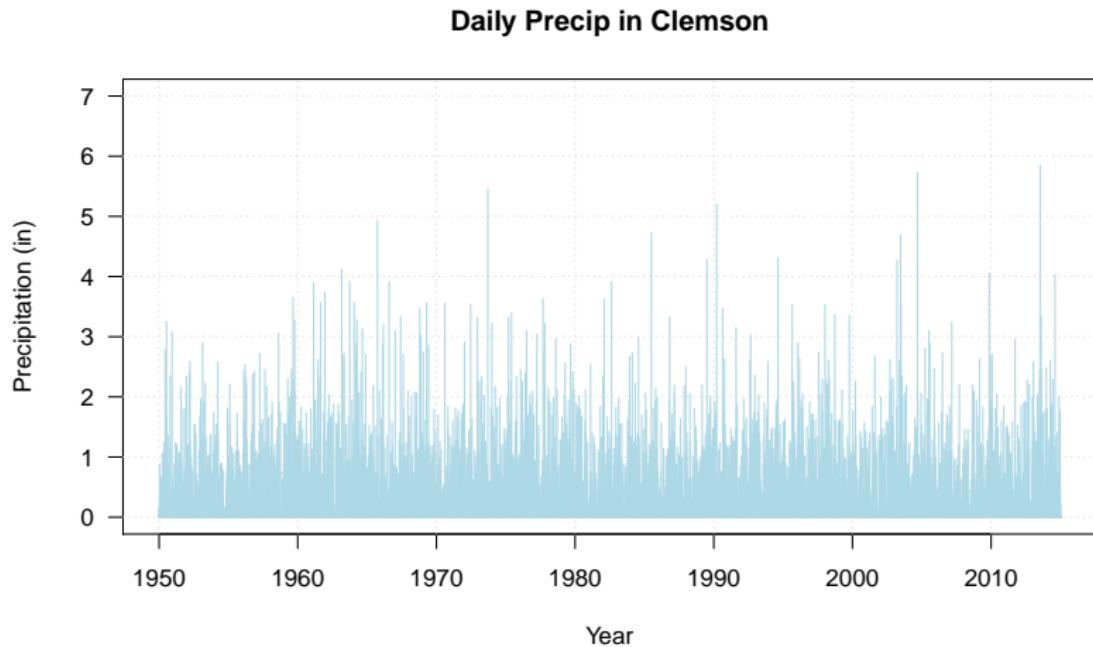
# Inference for GEV

- ▶ Delta method
- ▶ Profile likelihood method

# Minima

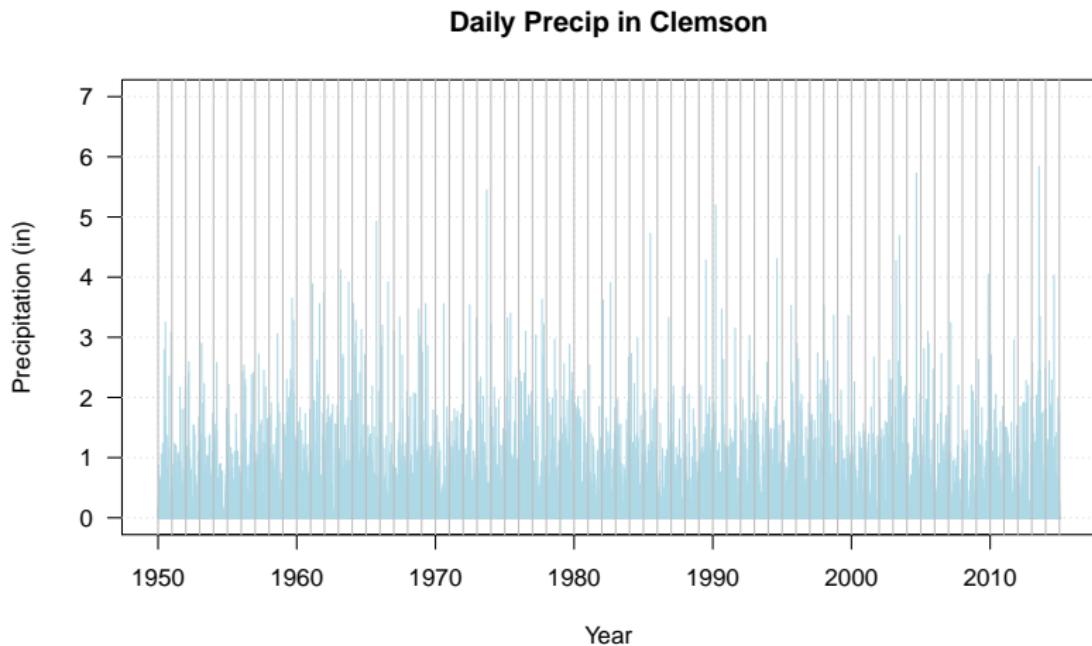
$$\min(X_1, \dots, X_n) = -\max(-X_1, \dots, -X_n)$$

# Clemson Daily Precipitation [Data Source: USHCN]



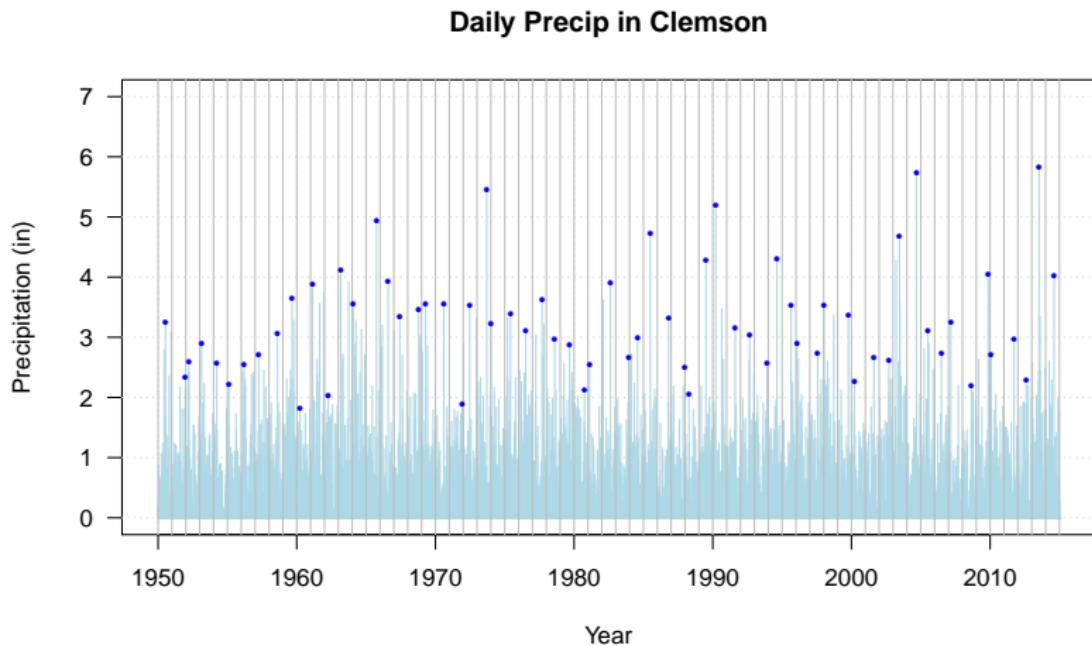
# Block Maxima Method (Gumbel 1958)

1. Determine the block size and extract the block maxima



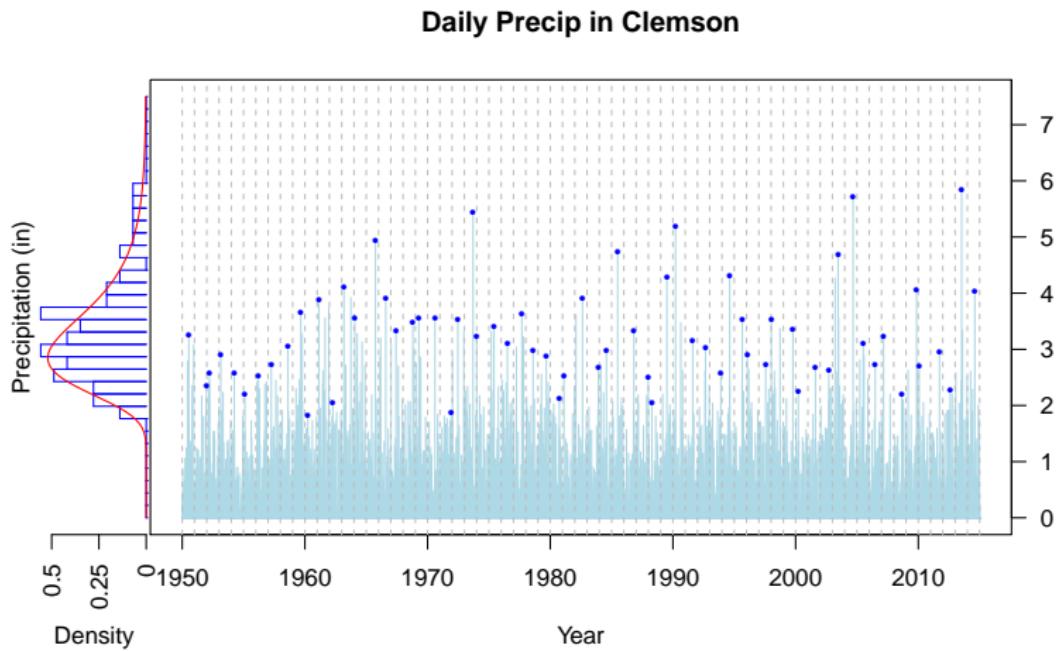
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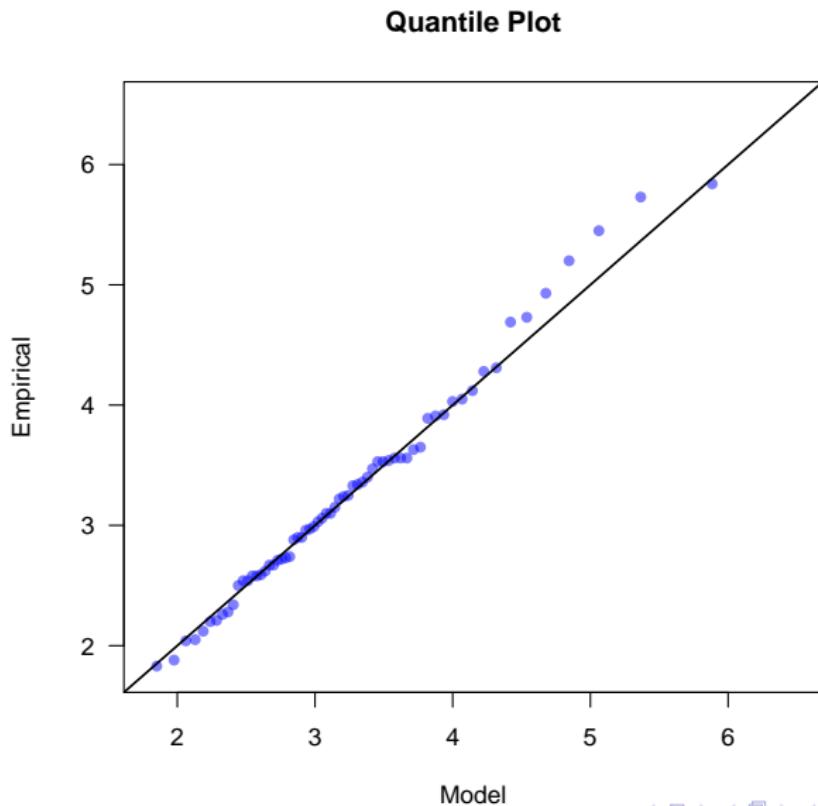
## Block Maxima Method (Gumbel 1958)

2. Fit the GEV to the maximal and assess the fit



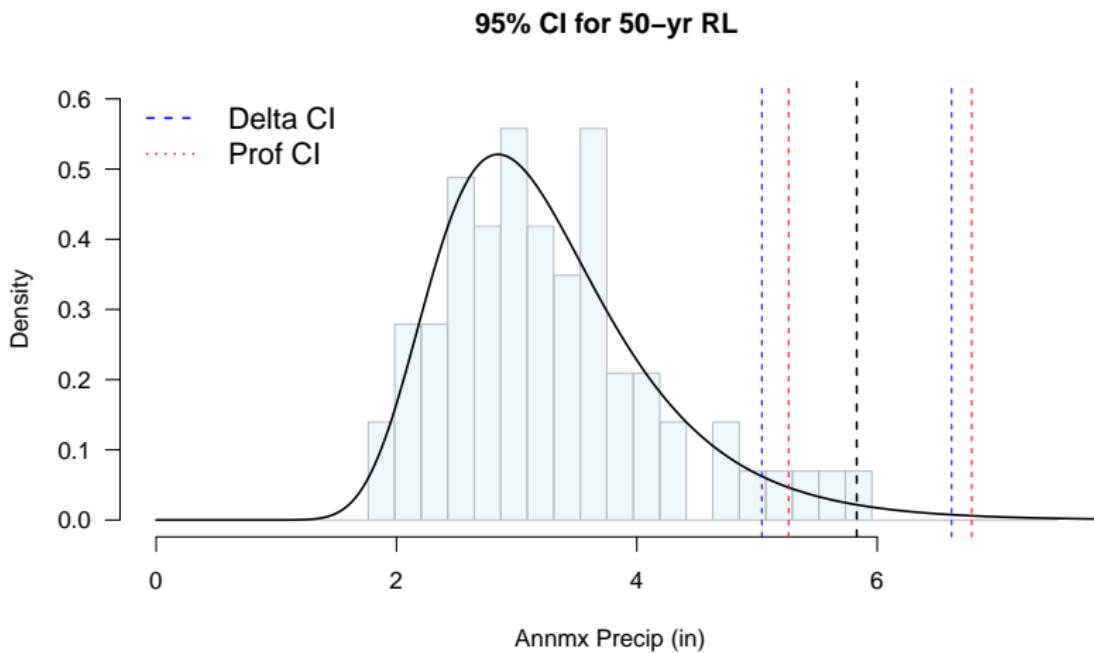
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## Block Maxima Method (Gumbel 1958)

3. Perform inference for return levels, probabilities, etc. Two methods to quantify estimation uncertainty: 1) Delta-method (relies on asymptotic normality); 2) Profile likelihood



# Outline

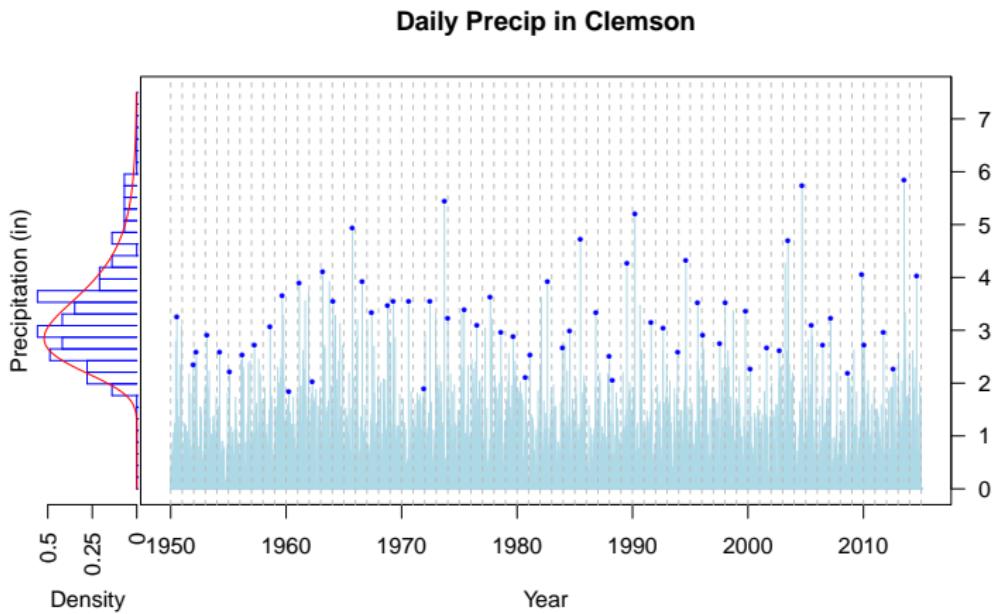
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Peaks–Over–Threshold (POT) Method

Non-stationary Extreme Value Analysis

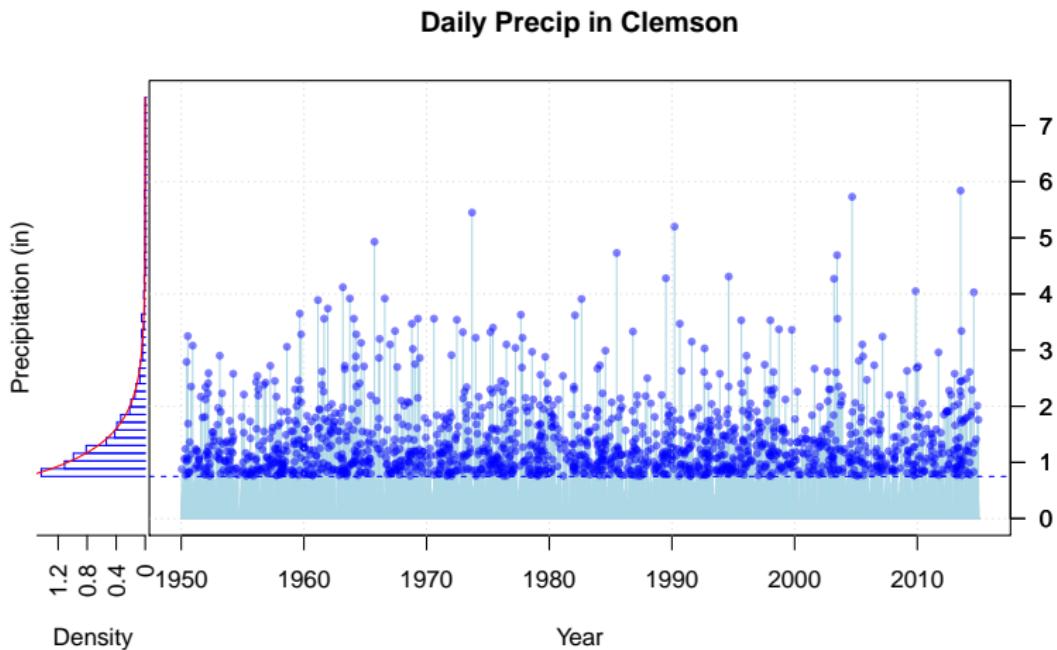
# Recall the Block Maxima Method



**Question:** Can we use data more efficiently?

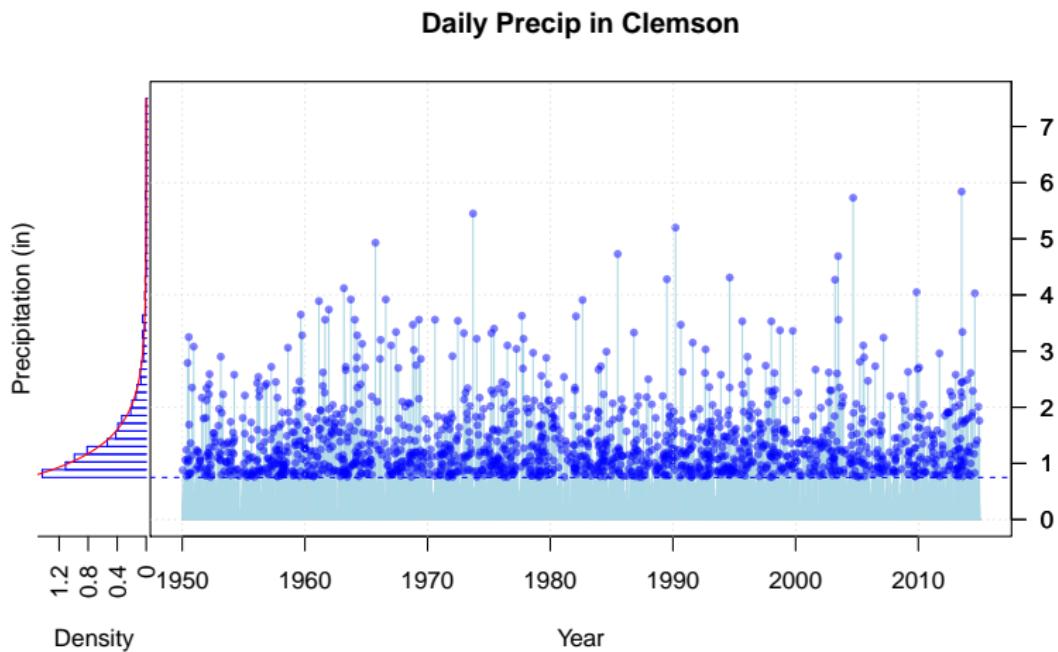
## Peaks-over-threshold (POT) method [Davison & Smith 1990]

1. Select a “sufficiently large” threshold  $u$ , extract the exceedances



## Peaks-over-threshold (POT) method [Davison & Smith 1990]

2. Fit an appropriate model to exceedances



## GPD for Exceedances

If  $M_n = \max_{i=1,\dots,n} X_i$  (for a large  $n$ ) can be approximated by a GEV( $\mu, \sigma, \xi$ ), then for sufficiently large  $u$ ,

$$\begin{aligned}\mathbb{P}(X_i > x + u | X_i > u) &= \frac{n\mathbb{P}(X_i > x + u)}{n\mathbb{P}(X_i > u)} \\ &\rightarrow \left( \frac{1 + \xi \frac{x+u-b_n}{a_n}}{1 + \xi \frac{u-b_n}{a_n}} \right)^{-\frac{1}{\xi}} \\ &= \left( 1 + \frac{\xi x}{a_n + \xi(u - b_n)} \right)^{-\frac{1}{\xi}}\end{aligned}$$

⇒ Survival function of **generalized Pareto distribution**

## Pickands–Balkema–de Haan Theorem (1974, 1975)

If  $M_n = \max_{1 \leq i \leq n} \{X_i\} \approx \text{GEV}(\mu, \sigma, \xi)$ , then, for a “large”  $u$  (i.e.,  $u \rightarrow x_F = \sup\{x : F(x) < 1\}$ ),

$$\mathbb{P}(X > u) \approx \frac{1}{n} \left[ 1 + \xi \left( \frac{u - \mu}{\sigma} \right) \right]^{\frac{-1}{\xi}}$$

$F_u = \mathbb{P}(X - u < y | X > u)$  is well approximated by the **generalized Pareto distribution (GPD)**. That is:

$$F_u(y) \xrightarrow{d} H_{\tilde{\sigma}, \xi}(y) \quad u \rightarrow x_F$$

where

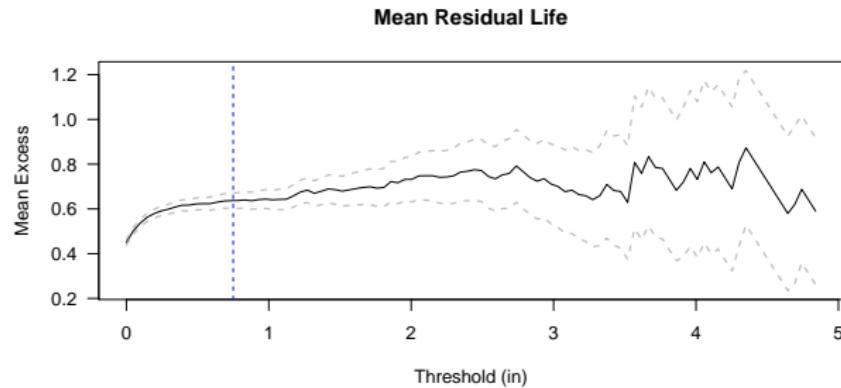
$$H_{\tilde{\sigma}, \xi}(y) = \begin{cases} 1 - (1 + \xi y / \tilde{\sigma})^{-1/\xi} & \xi \neq 0; \\ 1 - \exp(-y / \tilde{\sigma}) & \xi = 0. \end{cases}$$

and  $\tilde{\sigma} = \sigma + \xi(u - \mu)$

# How to Choose the Threshold?

## Bias-variance tradeoff:

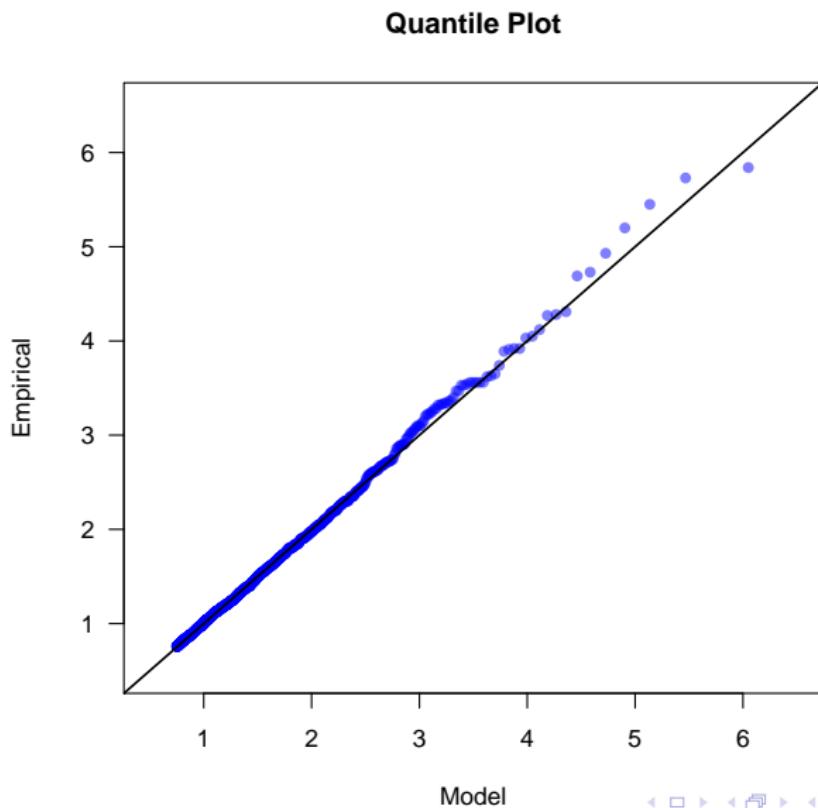
- ▶ Threshold too low  $\Rightarrow$  bias because of the model asymptotics being invalid
- ▶ Threshold too high  $\Rightarrow$  variance is large due to few data points



**Task:** To choose a  $u_0$  s.t. the Mean Residual Life curve behaves linearly  $\forall u > u_0$

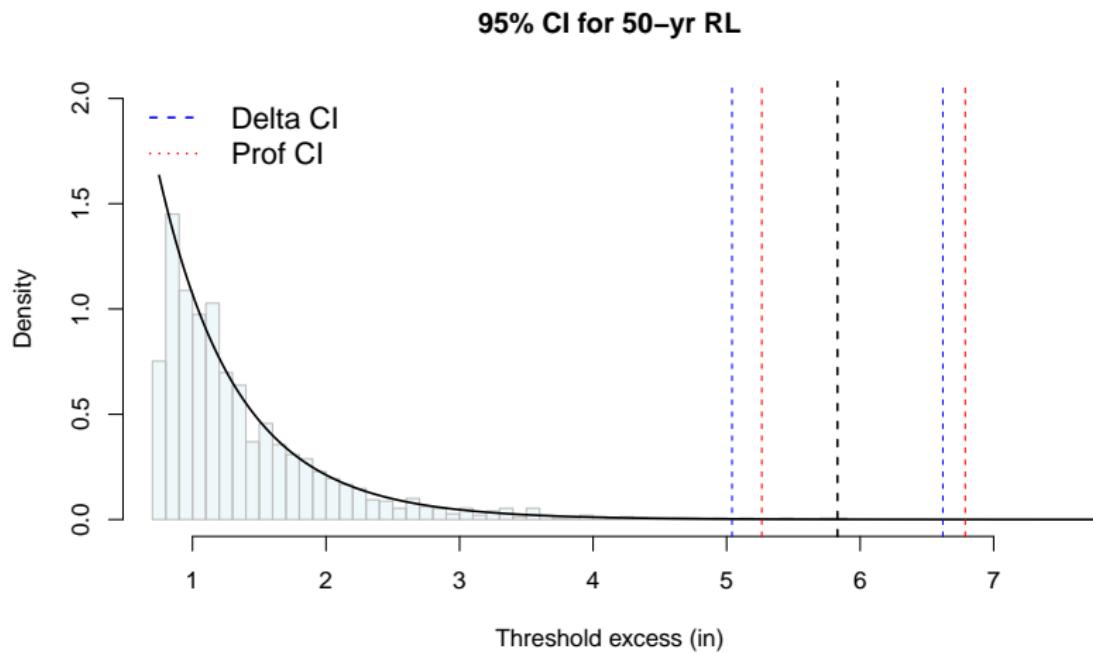
## Peaks-over-threshold (POT) method [Davison & Smith 1990]

2. Fit an appropriate model to exceedances and assess the fit



# Peaks-over-threshold (POT) method [Davison & Smith 1990]

3. Perform inference for return levels, probabilities, etc



## Temporal Dependence

**Question:** Is the GEV still the limiting distribution for block maxima of a stationary (but not independent) sequence  $\{X_i\}$ ?

**Answer:** Yes, as long as mixing conditions hold. ([Leadbetter et al., 1983](#))

What does this mean for inference?

**Block maximum approach:** GEV still correct for marginal. Since block maximum data likely have negligible dependence, proceed as usual

**Threshold exceedance approach:** GPD is correct for the marginal. If extremes occur in clusters, estimation affected as likelihood assumes independence of threshold exceedances

## Remarks on Univariate Extremes

- ▶ To estimate the tail, EVT uses only extreme observations
- ▶ Shape parameter  $\xi$  is extremely important but hard to estimate
- ▶ Threshold exceedance approaches allow the user to retain more data than block-maximum approaches, thereby reducing the uncertainty with parameter estimates
- ▶ Temporal dependence in the data is more of an issue in threshold exceedance models. One can either decluster, or alternatively, adjust inference

# Outline

Motivation

Extreme Value Theorem & Block Maxima Method

Peaks–Over–Threshold (POT) Method

Non-stationary Extreme Value Analysis

# Modeling Non-Stationary GEV

- ▶  $M_t \sim \text{GEV}(\mu(t), \sigma(t), \xi(t))$
- ▶ Typically assume “simple” structure for  $\mu(t)$  and  $\sigma(t)$ , and  $\xi(t)$  be a constant
- ▶  $\mu(t)$  and  $\sigma(t)$  could depend on some relevant factors

## Summary & Discussion

- ▶ Extreme value theory provides a framework to model extreme values
  - ▶ GEV for fitting block maxima
  - ▶ GPD for fitting threshold exceedances
  - ▶ Return level for communicating risk
- ▶ Practical Issues: seasonality, temporal dependence, non-stationarity, ...

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- ▶ Extreme value theory provides a framework to model extreme values
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