

Lecture 20

Review

STAT 8010 Statistical Methods I
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Inference for One Population Mean
Inference for Two Population Means
Inferences for Matched Pairs
ANOVA
Multiple Comparisons & Linear Contrasts
RCBD
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Notes

Agenda

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- 2 Inference for Two Population Means
- 3 Inferences for Matched Pairs
- 4 ANOVA
- 5 Multiple Comparisons & Linear Contrasts
- 6 RCBD

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Inference for One Population Mean

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Inferences for One Population Mean μ

Goal: To infer $\mu = \mathbb{E}(X)$ from a random sample $\{X_1, X_2, \dots, X_n\}$

- Point estimation:

$$\bar{X}_n = \frac{\sum_{i=1}^n X_i}{n}$$

- Interval Estimation: $100 \times (1 - \alpha)\%$ Confidence Interval (CI)

- $\sigma = \sqrt{\text{Var}(X)}$ is known:

$$\left(\bar{X}_n - z_{\alpha/2} \frac{\sigma}{\sqrt{n}}, \bar{X}_n + z_{\alpha/2} \frac{\sigma}{\sqrt{n}} \right)$$

- σ is unknown:

$$\left(\bar{X}_n - t_{\alpha/2, df=n-1} \frac{s}{\sqrt{n}}, \bar{X}_n + t_{\alpha/2, df=n-1} \frac{s}{\sqrt{n}} \right)$$

Assumptions

- Data is a random sample from the population
- i) sample size n is sufficiently large (e.g., $n > 30$) or the population X follows a normal distribution



Notes

Margin of Error & Sample Size Calculation

- Margin of error:

$$z_{\alpha/2} \frac{\sigma}{\sqrt{n}} \quad \text{if } \sigma \text{ known}$$

$$t_{\alpha/2, df=n-1} \frac{s}{\sqrt{n}} \quad \text{if } \sigma \text{ unknown}$$

\Rightarrow CI for $\mu = \bar{X}_n \pm \text{margin of error}$

- Sample size determination:

$$n = \left(\frac{z_{\alpha/2} \times \sigma}{\text{margin of error}} \right)^2,$$

if σ is given



Notes

Hypothesis Testing for μ

- 1 State the null and alternative hypotheses:

$$H_0 : \mu = \mu_0 \text{ vs. } H_a : \mu > \text{ or } \neq \text{ or } < \mu_0$$

- 2 Compute the test statistic:

$$z_{\text{obs}} = \frac{\bar{X}_n - \mu_0}{\sigma / \sqrt{n}}, \quad \sigma \text{ known}; \quad t_{\text{obs}} = \frac{\bar{X}_n - \mu_0}{s / \sqrt{n}}, \quad \sigma \text{ unknown}$$

- 3 Make the decision of the test:

Rejection Region/ P-Value Methods

- 4 Draw the conclusion of the test:

We (do/do not) have enough statistical evidence to conclude that (H_a in words) at α significant level.



Notes

Type I, II Error & Power

True State	Decision	
	Reject H_0	Fail to reject H_0
H_0 is true	Type I error	Correct
H_0 is false	Correct	Type II error

- Type I error: $\mathbb{P}(\text{Reject } H_0 | H_0 \text{ is true}) = \alpha$
 - Type II error: $\mathbb{P}(\text{Fail to reject } H_0 | H_0 \text{ is false}) = \beta$
 - The power (PWR): $\mathbb{P}(\text{Reject } H_0 | H_0 \text{ is false}) = 1 - \beta$.
- $\Rightarrow \text{PWR}(\mu_a) = 1 - \beta(\mu_a) = 1 - \mathbb{P}(z^* \leq z_\alpha - \frac{|\mu_0 - \mu_a|}{\sigma/\sqrt{n}})$

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Duality of Hypothesis Test with Confidence Interval

There is an interesting relationship between CIs and hypothesis tests. If $H_0 = \mu_0$ is rejected with significance level α then the corresponding confidence interval does not contain μ_0 with the confidence level $(1 - \alpha)$, and vice versa

Hypothesis testing at α level	$(1 - \alpha) \times 100\%$ CI
$H_0 : \mu = \mu_0$ vs. $H_a : \mu \neq \mu_0$	$\bar{X} \pm t_{\alpha/2, n-1} s / \sqrt{n}$
$H_0 : \mu = \mu_0$ vs. $H_a : \mu > \mu_0$	$(\bar{X} - t_{\alpha, n-1} s / \sqrt{n}, \infty)$
$H_0 : \mu = \mu_0$ vs. $H_a : \mu < \mu_0$	$(-\infty, \bar{X} + t_{\alpha, n-1} s / \sqrt{n})$

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Inferences for Two Population Means

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Statistical Inference for $\mu_1 - \mu_2$

- **Assumptions:** Populations X_1 and X_2 both follows a normal distribution (or their sample sizes are large enough); Data are random sample from their population
- Point estimation: $\bar{X}_1 - \bar{X}_2$
- Interval estimation:

$$\bar{X}_1 - \bar{X}_2 \pm \text{margin of error,}$$

where margin of error =

$$t_{\alpha/2, df^*} \sqrt{\frac{s_1^2}{n_1} + \frac{s_2^2}{n_2}}, \quad df^* = \frac{\left(\frac{s_1^2}{n_1} + \frac{s_2^2}{n_2}\right)^2}{\frac{(s_1^2/n_1)^2}{n_1-1} + \frac{(s_2^2/n_2)^2}{n_2-1}}$$

When s_1 and s_2 “similar enough”, we replace

$$\sqrt{\frac{s_1^2}{n_1} + \frac{s_2^2}{n_2}} \text{ by } s_p \sqrt{\frac{1}{n_1} + \frac{1}{n_2}}, \text{ where}$$
$$s_p = \sqrt{\frac{(n_1-1)s_1^2 + (n_2-1)s_2^2}{n_1+n_2-2}}, df = n_1 + n_2 - 2$$

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Hypothesis Testing for $\mu_1 - \mu_2$

- State the null and alternative hypotheses:
 - Upper-tailed test: $H_0 : \mu_1 - \mu_2 = 0$ vs. $H_a : \mu_1 - \mu_2 > 0$
 - Lower-tailed test: $H_0 : \mu_1 - \mu_2 = 0$ vs. $H_a : \mu_1 - \mu_2 < 0$
 - Two-tailed test: $H_0 : \mu_1 - \mu_2 = 0$ vs. $H_a : \mu_1 - \mu_2 \neq 0$
- Compute the test statistic:

$$t_{obs} = \frac{\bar{X}_1 - \bar{X}_2}{s_p \sqrt{\frac{1}{n_1} + \frac{1}{n_2}}}, \quad \sigma_1 = \sigma_2$$
$$\frac{\bar{X}_1 - \bar{X}_2}{\sqrt{\frac{s_1^2}{n_1} + \frac{s_2^2}{n_2}}}, \quad \sigma_1 \neq \sigma_2$$

- Make the decision of the test:
Rejection Region/ P-Value Methods
- Draw the conclusion of the test

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Inferences for Matched Pairs

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Notes

Paired T-Tests

- **When to use:** before/after study, pairing subjects, study on twins, etc
- $H_0 : \mu_{diff} = 0$ vs. $H_a : \mu_{diff} > 0$ or $\mu_{diff} < 0$ or $\mu_{diff} \neq 0$, where μ_{diff} is the population mean of the paired difference
- Test statistic: $t_{obs} = \frac{\bar{x}_{diff} - 0}{\frac{s_{diff}}{\sqrt{n}}}$

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ANOVA

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ANOVA and Overall F Test

Overall F-Test

- $H_0 : \mu_1 = \mu_2 = \dots = \mu_J$
 H_a : At least one mean is different

- ANOVA Table:

Source	df	SS	MS	F statistic
Treatment	$J - 1$	SSTr	$MSTr = \frac{SSTr}{J - 1}$	$F = \frac{MSTr}{MSE}$
Error	$N - J$	SSE	$MSE = \frac{SSE}{N - J}$	
Total	$N - 1$	SSTo		

- Test Statistic: $F_{obs} = \frac{MSTr}{MSE}$. Under H_0 ,
 $F^* \sim F_{df_1 = J - 1, df_2 = N - J}$

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Multiple Comparisons & Linear Contrasts

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Family-Wise Error Rate (FWER) and Multitple Comparisons

- **Family-Wise Error Rate (FWER)** $\bar{\alpha}$: the probability of making 1 or more type I errors in a set of hypothesis tests
- **Bonferroni Correction**: Adjust the significant level for each of the m tests to be $\frac{\alpha}{m}$ to control the **FWER**
- **Fisher's LSD** and **Tukey's HSD**

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Linear Contrasts

- **Definition**: Let c_1, c_2, \dots, c_J are constants where $\sum_{j=1}^J c_j = 0$, then $L = \sum_{j=1}^J c_j \mu_j$ is called a **linear contrast** of the population means.
- Point Estimation:

$$\hat{L} = \sum_{j=1}^J c_j \bar{X}_j$$

- Interval Estimation:

$$(\hat{L} - t_{\alpha/2, df=N-J} \hat{se}_{\hat{L}}, \hat{L} + t_{\alpha/2, df=N-J} \hat{se}_{\hat{L}}),$$

where $\hat{se}_{\hat{L}} = \sqrt{MSE \left(\frac{c_1^2}{n_1} + \dots + \frac{c_J^2}{n_J} \right)}$

- Hypothesis Testing for linear contrasts

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- ANOVA table

Source	df	SS	MS	F statistic
Treatment	$J - 1$	SS_{trt}	$MS_{trt} = \frac{SS_{trt}}{J-1}$	$F_{trt} = \frac{MS_{trt}}{MS_{err}}$
Block	$B - 1$	SS_{blk}	$MS_{blk} = \frac{SS_{blk}}{B-1}$	
Error	$(B - 1)(J - 1)$	SS_{err}	$MS_{err} = \frac{SS_{err}}{(B-1)(J-1)}$	
Total	$N - 1$	SS_{tot}		

- Why we may want to do blocking (See the example in Lecture 17)?
- Use interaction plot to assess the additivity assumption (i.e., treatment effects are consistent across blocks)

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