

Lecture 11

Random and Mixed Effects Models, Computer Experiments

Reading: Oehlert Chapter 11.1-11.7; Dean-Voss-Draguljić
Chapters 17.1-17.3; 17.7& 20

DSA 8020 Statistical Methods II

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Agenda

Random and Mixed
Effects Models,
Computer
Experiments



Random and Mixed
Effects Models

Computer Experiments

1 Random and Mixed Effects Models

2 Computer Experiments

Everything we have done so far has dealt with **fixed effects**

CRD: $y_{ij} = \mu + \alpha_i + \epsilon_{ij}$

RCBD: $y_{ij} = \mu + \alpha_i + \beta_j + \epsilon_{ij}$

Factorial: $y_{ij} = \mu + \alpha_i + \beta_j + (\alpha\beta)_{ij} + \epsilon_{ij}$

- The treatment effects are unknown but constants \Rightarrow if we ran the experiment over again, would expect the same treatment effects
- We can increase the power of all of our tests by increasing the sample size n
- We perform inference on the treatment effects via t-tests and F-tests

Random effects models look very similar to fixed effects models. For example, we could have

$$y_{ij} = \mu + \alpha_i + \epsilon_{ij}.$$

The difference is in the **assumptions** we make for the treatment effects

Fixed Effects

Treatment effects α_i s are unknown **constants** that add to zero (or some other constraint)

Random Effects

- α_i s $\sim N(0, \sigma_\alpha^2)$
- α_i s are independent of ϵ_{ij}

How and Why Are Things So Different?

Fixed effects:

- The treatments are the treatments and they are unchanging
- If we rerun the experiment, we are still studying the same treatments

Random effects:

- The treatments are a random sample from a population of potential treatments
- If we rerun the experiment, we are looking at an entirely new sample of treatments
- Inference is on the population of potential treatments

Fixed effects:

- $\text{Var}(y_{ij}) = \sigma^2$
- All y_{ij} s are independent of each other
- Interest is about α_i s

Random effects:

- $\text{Var}(y_{ij}) = \sigma_\alpha^2 + \sigma^2$
- $$\text{Cor}(y_{ij}, y_{kl}) = \begin{cases} 0 & \text{if } i \neq k \\ \frac{\sigma_\alpha^2}{\sigma_\alpha^2 + \sigma^2} & \text{if } i = k; j \neq l \\ 1 & \text{if } i = k; j = l \end{cases}$$
- Interest is (mostly) about σ_α^2

An Example of Fixed Effects vs Random Effects

- 1 Compare reading ability of 10 2nd grade classes in NY:

Select $g = 10$ specific classes of interest. Randomly choose n students from each classroom. Want to compare α_i s (class-specific effects) \Rightarrow **Fixed effects**

- 2 Compare variability among all 2nd grade classes in NY:

Randomly choose $g = 10$ classes from large number of classes. Randomly choose n students from each classroom. Want to assess σ_α^2 (class to class variability) \Rightarrow **Random effects**

$$y_{ij} = \mu + \alpha_i + \epsilon_{ij},$$

where

- μ is the overall mean
- α_i : i th treatment effect and $\alpha_i \sim N(0, \sigma_\alpha^2)$
- $\{\alpha_i\}$ and $\{\epsilon_{ij}\}$ independent
- The hypotheses are:

$$H_0 : \sigma_\alpha^2 = 0$$

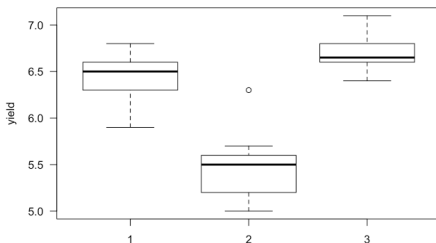
$$H_a : \sigma_\alpha^2 > 0$$

One can use either “old school” method (ANOVA) or “**new school**” method (REML) to make inference about σ_α^2

Random Effects Example

Suppose that an agronomist is studying a large number of varieties of soybeans for yield. The agronomist randomly selects three varieties, and then randomly assigns each of those varieties to 10 of 30 available plots.

Soybean	Yield
V1	6.6, 6.4, 5.9, 6.6, 6.2, 6.7, 6.3, 6.5, 6.5, 6.8
V2	5.6, 5.2, 5.3, 5.1, 5.7, 5.6, 5.6, 6.3, 5.0, 5.4
V3	6.9, 7.1, 6.4, 6.7, 6.5, 6.6, 6.6, 6.6, 6.8, 6.8



```
> fixef <- lm(yield ~ var)
```

```
> anova(fixef)
```

Analysis of Variance Table

Response: yield

	Df	Sum Sq	Mean Sq	F value	Pr(>F)
var	2	8.306	4.1530	49.593	9.114e-10 ***
Residuals	27	2.261	0.0837		

Signif. codes:

0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

```
> coefficients(fixef)
```

(Intercept)	var2	var3
6.45	-0.97	0.25

Random Effects Analysis

```
> library(lme4)
> randef <- lmer(yield ~ 1 + (1|var), REML = TRUE)
> summary(mod1)
```

Linear mixed model fit by maximum likelihood . t-tests
use Satterthwaite's method [lmerModLmerTest]
Formula: yield ~ 1 + (1 | var)

AIC	BIC	logLik	deviance	df.resid
27.2	31.4	-10.6	21.2	27

Scaled residuals:

Min	1Q	Median	3Q	Max
-1.8755	-0.6033	0.1245	0.5068	2.7574

Random effects:

Groups	Name	Variance	Std.Dev.
var	(Intercept)	0.26849	0.5182
Residual		0.08374	0.2894

Number of obs: 30, groups: var, 3

Fixed effects:

	Estimate	Std. Error	df	t value	Pr(> t)
(Intercept)	6.2100	0.3038	3.0000	20.44	0.000256

Concrete Cylinder Example Revisited

Suppose you are manufacturing concrete cylinders for bridge supports. There are three ways of drying concrete (say A, B, and C), and you want to find the one that gives you the best compressive strength. The concrete is mixed in batches that are large enough to produce exactly three cylinders, and your production engineer believes that there is substantial variation in the quality of the concrete from batch to batch.

Treatment	Batch					Trt Sum
	1	2	3	4	5	
A	52	47	44	51	42	236
B	60	55	49	52	43	259
C	56	48	45	44	38	231
Batch Mean	168	150	138	147	123	726

If we were treat the batch effects as random effects, then we have a **Mixed Effects Model**

Concrete Cylinder Example: Mixed Effects Analysis

```
> randef <- lmer(x ~ trt + (1|blk), REML = TRUE, data = dat)
> summary(randef)
```

Linear mixed model fit by REML. t-tests use

Satterthwaite's method [lmerModLmerTest]

Formula: x ~ trt + (1 | blk)

Data: dat

REML criterion at convergence: 71.1

Scaled residuals:

Min	1Q	Median	3Q	Max
-1.1417	-0.6147	-0.1494	0.5772	1.3390

Random effects:

Groups	Name	Variance	Std.Dev.
blk	(Intercept)	28.35	5.324
Residual		5.85	2.419

Number of obs: 15, groups: blk, 5

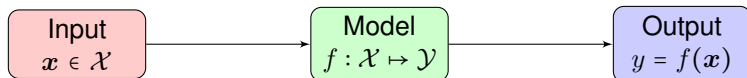
Fixed effects:

	Estimate	Std. Error	df	t value	Pr(> t)
(Intercept)	47.200	2.615	5.054	18.047	8.76e-06
trtB	4.600	1.530	8.000	3.007	0.0169
trtC	-1.000	1.530	8.000	-0.654	0.5316

What is a Computer Experiment

In some situations it is economically, ethically, or simply not possible to run a **physical experiment**. Instead, the following scenario might be feasible:

- the physical process can be described by a mathematical model (e.g., a system of differential equations)
- computer code (simulator) can be written to compute the response from the mathematical model

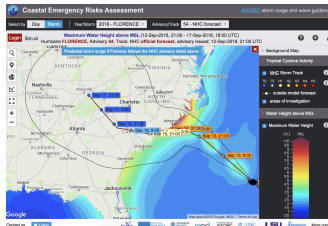
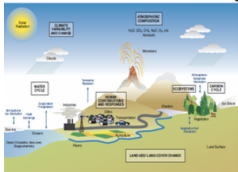


In this case, a researcher can conduct a **computer experiment** by running the computer code, which serves as a proxy for the physical process, to compute a “response” at any combination of values of the inputs

Random and Mixed Effects Models, Computer Experiments

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Computer Experiments



Computer Experiments vs. Physical Experiments

- *“Experimental results are believed by everyone, except for the person who ran the experiment”*
- *“Computational results are believed by no one, except the person who wrote the code”*

Replication, randomization and blocking are irrelevant for a computer experiment because many **computer codes are deterministic** and **all the inputs to the code are known and can be controlled**

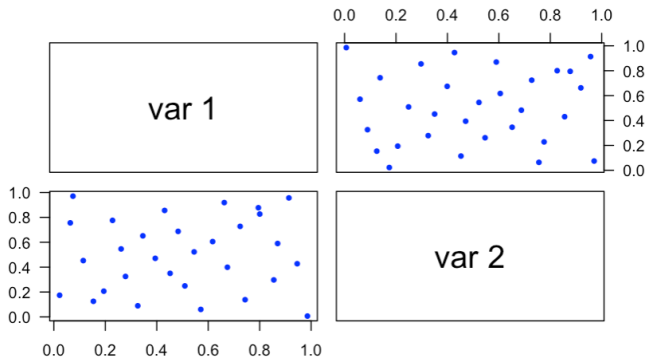
Here we are concerning about **design** and **analysis** of computer experiments:

- **Design**: Which configurations of $\{\mathbf{x}_i\}_{i=1}^n$ to run the computer model
- **Analysis**: How to estimate the input-output relationship $y = f(\mathbf{x})$ using data $\{\mathbf{x}_i, y_i\}_{i=1}^n$ from a computer experiment

Design of Computer Experiments

Question: where to make the runs, i.e., the selection of inputs $\{x_i\}_{i=1}^n$ for a given computational budget n .

Example: $x_i = (x_{i1}, x_{i2})^T$ and $n = 30$



This is an example of [space-filling designs](#), which allow for a spread of points encompassing the entire design space

Goal: fit a statistical model to the computer model inputs-output $\{y_i, \mathbf{x}_i\}_{i=1}^n$ to “emulate” the simulator and to quantify the prediction uncertainty for $y(\mathbf{x}_{\text{new}})$ via a **Gaussian Process Model** GP $(m(\cdot), K(\cdot, \cdot))$, where

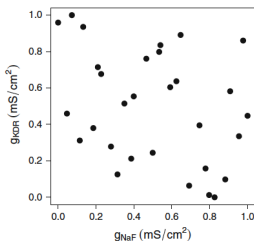
- $m(\mathbf{x}) = E[y(\mathbf{x})]$ is the **mean function**, usually takes a simple form, e.g., $m(\mathbf{x}) = \mu$
- $K(\mathbf{x}, \mathbf{x}') = \text{Cov}(y(\mathbf{x}), y(\mathbf{x}'))$ is the **covariance function**, usually parametrized by “distance”. e.g.,
$$K(\mathbf{x}, \mathbf{x}') = C(\mathbf{x}, \mathbf{x}'; \boldsymbol{\theta}) = \sigma^2 \prod_{j=1}^p C_j(d(\mathbf{x}_j, \mathbf{x}'_j); \theta_j).$$

Parameters $(\mu, \sigma^2, \{\theta_j\}_{j=1}^p)$ can be estimated by fitting a GP model to $\{y_i, \mathbf{x}_{i=1}^n\}$ via **maximum likelihood method**. The prediction (e.g., predicting $y(\mathbf{x}_{\text{new}})$) and prediction uncertainty can be carried out using the (Gaussian) conditional distribution formula

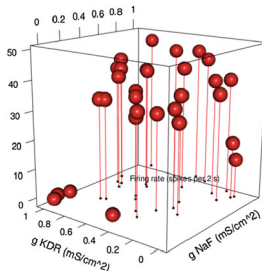
Neuron Experiment [pp.776-778, Dean-Voss-Draguljić, 2017]

The firing rate of a neuron at +380 pA current injection of a young monkey is modeled as a deterministic function of two input variables:

- x_1 g_{NaF} : maximal conductance of the transient sodium
- x_2 g_{KDR} : maximal conductance of the delayed-rectifier potassium



(a) 30×2 LHD



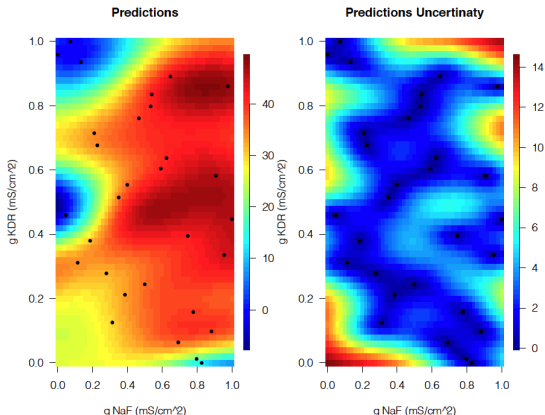
(b) Firing rates

Source: Fig. 20.6, Dean-Voss-Draguljić, 2017

The goal here is to reconstruct the 2D firing rate surface within the input space

Neuron Experiment Result

- A GP with squared exponential covariance function (i.e., $C(\mathbf{x}, \mathbf{x}') = \sigma^2 e^{-[\theta_1(x_1 - x'_1)^2 + \theta_2(x_2 - x'_2)^2]}$) is fitted to $\{y_i, \mathbf{x}_i\}_{i=1}^n$ with the estimated parameters $\hat{\mu} = 27.61$ $\hat{\sigma}^2 = 251.86$, $\hat{\theta}_{\text{NaF}} = 5.03$, $\hat{\theta}_{\text{KDR}} = 50.22$.
- With these estimated parameters one can calculate the predictions (**Left**) and their prediction uncertainties (**Right**)



This slides cover:

- Random and Mixed Effects Models
- Computer Experiments: Concepts, Design and Analysis