

Lecture 10

Univariate Volatility Modeling

Reading: An introduction to analysis of financial data with R
(2013) by Ruey Tsay

MATH 8090 Time Series Analysis
Week 10

Background

ARCH Model

GARCH Model

IGARCH and
EGARCH Models

Stochastic Volatility
Model

Whitney Huang
Clemson University

- 1 Background
- 2 ARCH Model
- 3 GARCH Model
- 4 IGARCH and EGARCH Models
- 5 Stochastic Volatility Model

Background

ARCH Model

GARCH Model

IGARCH and
EGARCH Models

Stochastic Volatility
Model

Apple Inc

\$171.95

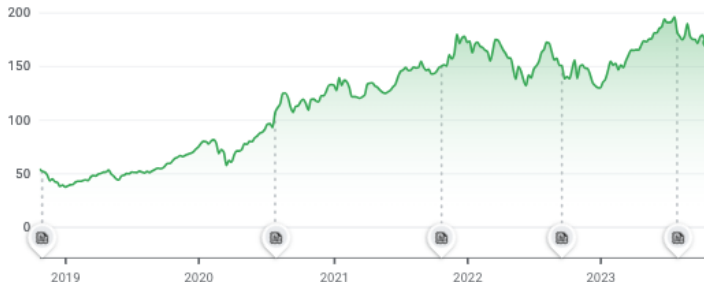
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+117.87 5Y

Oct 25, 10:26:13 AM UTC-4 · USD · NASDAQ · Disclaimer

1D 5D 1M 6M YTD 1Y 5Y MAX

× Key events



Source: Google Finance

Background

ARCH Model

GARCH Model

IGARCH and
EGARCH Models

Stochastic Volatility
Model

Log Returns of Apple Stock

$$r_t = \log\left(\frac{y_t}{y_{t-1}}\right), \quad \text{where } y_t \text{ is the price at time } t$$

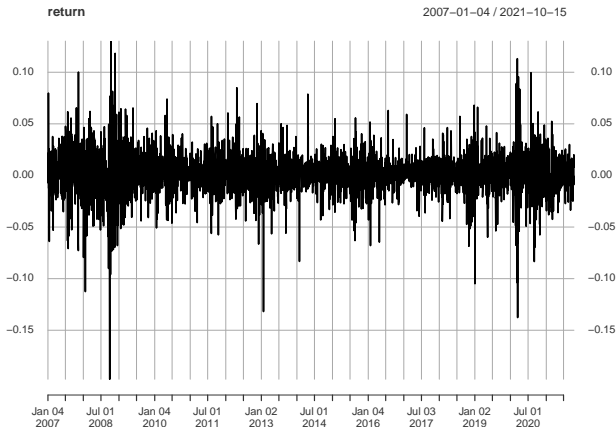
Background

ARCH Model

GARCH Model

IGARCH and
EGARCH Models

Stochastic Volatility
Model



Periods of high uncertainty or large price movements tend to cluster together \Rightarrow Volatility Clustering

Volatility measures the degree of **variation** in asset prices over time, typically quantified as the **(conditional) standard deviation of log returns**.

Why is volatility important?

- **Option pricing** - e.g., Black-Scholes model
- **Risk management** - e.g., Value-at-Risk (VaR)
- **Asset allocation** - e.g., minimum-variance portfolios
- **Forecasting** - e.g., interval prediction of returns

Key challenge: Volatility is **not directly observable**

Background

ARCH Model

GARCH Model

IGARCH and
EGARCH Models

Stochastic Volatility
Model

We will take an **econometric approach** by modeling the conditional standard deviation (σ_t) of daily or monthly returns

Basic structure

$$r_t = \mu_t + a_t, \quad \mu_t = \mathbb{E}(r_t | F_{t-1}) = \phi_0 + \sum_{i=1}^p \phi_i r_{t-i} + \sum_{j=1}^q \theta_j a_{t-j}$$

Volatility models are concerned with time-evolution of

$$\text{Var}(r_t | F_{t-1}) = \text{Var}(a_t | F_{t-1}) = \sigma_t^2,$$

the conditional variance of a return

Background

ARCH Model

GARCH Model

IGARCH and
EGARCH ModelsStochastic Volatility
Model

- Autoregressive conditional heteroscedastic (ARCH) model [Engle, 1982]
- Generalized Autoregressive Conditional Heteroskedasticity (GARCH) Model [Bollerslev, 1986]
- Integrated Generalized Autoregressive Conditional heteroskedasticity (IGARCH) model
- Exponential general autoregressive conditional heteroskedastic (EGARCH) model [Nelson, 1991]
- Asymmetric parametric ARCH models [Ding, Granger, and Engle, 1994]
- Stochastic volatility (SV) models [Melino and Turnbull, 1990; Harvey, Ruiz, and Shephard, 1994; Jacquier, Polson. and Rossi, 1994]

ARCH(m) model:

$$a_t = \sigma_t \epsilon_t, \quad \sigma_t^2 = \alpha_0 + \alpha_1 a_{t-1}^2 + \dots + \alpha_m a_{t-m}^2, \quad \alpha_i \geq 0 \text{ for } 1 \leq i \leq m$$

where $\{\epsilon_t\}$ is an i.i.d. sequence with:

- $\mathbb{E}[\epsilon_t] = 0$
- $\text{Var}[\epsilon_t] = 1$
- Common choices: standard normal, standardized Student-t, or generalized error distribution (and their skewed forms)

ARMA Analogy: ARCH(m) is analogous to an AR(m) model, but for the **squared innovations**

ARCH(1) model

$$a_t = \sigma_t \epsilon_t, \quad \sigma_t^2 = \alpha_0 + \alpha_1 a_{t-1}^2, \alpha_0 > 0, \quad \alpha_1 \geq 0$$

Main Properties:

- $\mathbb{E}(a_t) = 0$, $\text{Var}(a_t) = \frac{\alpha_0}{1 - \alpha_1}$, valid if $0 < \alpha_1 < 1$
- Under normality:

$$\mathbb{E}[a_t^4] = \frac{3\alpha_0^2(1 + \alpha_1)}{(1 - \alpha_1)(1 - 3\alpha_1^2)}$$

Finite only if $0 < \alpha_1^2 < \frac{1}{3}$

Implications:

- Finite variance: $\alpha_1 < 1$. Finite kurtosis: $\alpha_1^2 < \frac{1}{3}$
- Large $\alpha_1 \Rightarrow$ heavy-tailed a_t

Even with Gaussian errors, ARCH(1) produces heavy-tailed returns

Background

ARCH Model

GARCH Model

IGARCH and
EGARCH Models

Stochastic Volatility
Model

1 Model the mean and test for ARCH effects

Background

ARCH Model

GARCH Model

IGARCH and
EGARCH Models

Stochastic Volatility
Model

1 Model the mean and test for ARCH effects

- Fit a mean model: $a_t = y_t - \hat{\mu}_t$

Background

ARCH Model

GARCH Model

IGARCH and
EGARCH Models

Stochastic Volatility
Model

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- Test for conditional heteroscedasticity in $\{a_t\}$:

H_0 : no ARCH effects vs. H_1 : ARCH effects present

Background

ARCH Model

GARCH Model

IGARCH and
EGARCH Models

Stochastic Volatility
Model

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4 Check model adequacy

- Check Q -statistics for standardized residuals and their squares
- Assess skewness and kurtosis of residuals
- Verify no remaining ARCH effects

Advantages:

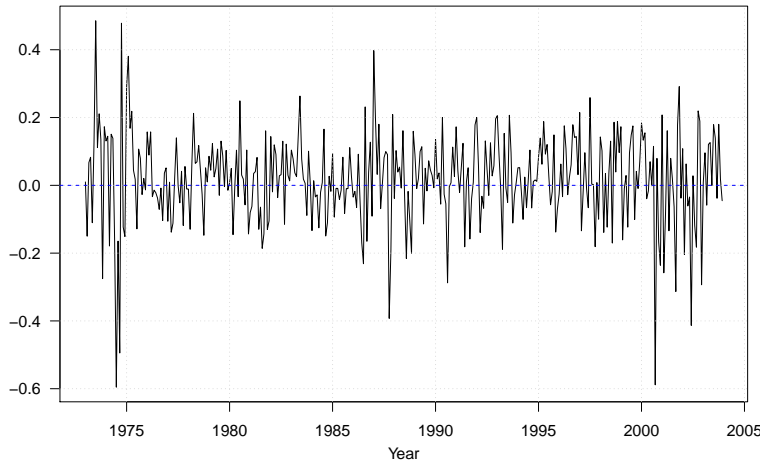
- **Simple and intuitive:** easy to understand and estimate
- **Captures volatility clustering:** periods of high and low variability
- **Produces heavy tails:** helps explain large market movements

Limitations:

- **Symmetric response:** treats positive and negative shocks equally
- **Parameter constraints:** e.g., ARCH(1) requires $\alpha_1^2 \in [0, 1/3]$
- **Limited flexibility:** volatility forecasts can change too slowly

Example: Monthly Log Returns of Intel Stock

Here we use the monthly log returns of Intel stock to illustrate ARCH modeling



Background

ARCH Model

GARCH Model

IGARCH and
EGARCH Models

Stochastic Volatility
Model

Testing ARCH Effect

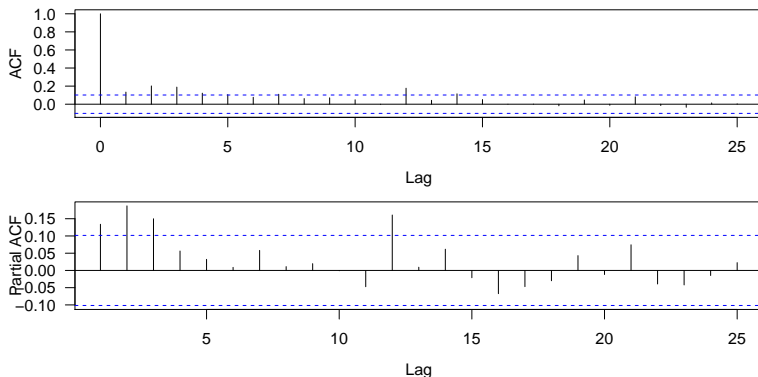
Here we test and examine the temporal pattern of the **squared residuals**

```
> Box.test(y^2, lag = 12, type = 'Ljung')
```

Box-Ljung test

data: y^2

X-squared = 68.67, df = 12, p-value = 5.676e-10



Here we fit an ARCH(3) for the volatility:

$$r_t = \mu + a_t, \quad a_t = \sigma_t \epsilon_t, \quad \sigma_t^2 = \alpha_0 + \sum_{i=1}^3 \alpha_i a_{t-i}^2,$$

assuming $\epsilon_t \stackrel{i.i.d.}{\sim} N(0, 1)$.

Error Analysis:

	Estimate	Std. Error	t value	Pr(> t)
mu	0.016572	0.006423	2.580	0.00988 **
omega	0.012043	0.001579	7.627	2.4e-14 ***
alpha1	0.208649	0.129177	1.615	0.10626
alpha2	0.071837	0.048551	1.480	0.13897
alpha3	0.049045	0.048847	1.004	0.31536

Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

[Background](#)[ARCH Model](#)[GARCH Model](#)[IGARCH and
EGARCH Models](#)[Stochastic Volatility
Model](#)

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Let's fit a simplified ARCH(1) model

[Background](#)[ARCH Model](#)[GARCH Model](#)[IGARCH and
EGARCH Models](#)[Stochastic Volatility
Model](#)

ARCH(1) Model Fitting

Error Analysis:

	Estimate	Std. Error	t value	Pr(> t)	
mu	0.016570	0.006161	2.689	0.00716	**
omega	0.012490	0.001549	8.061	6.66e-16	***
alpha1	0.363447	0.131598	2.762	0.00575	**

Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

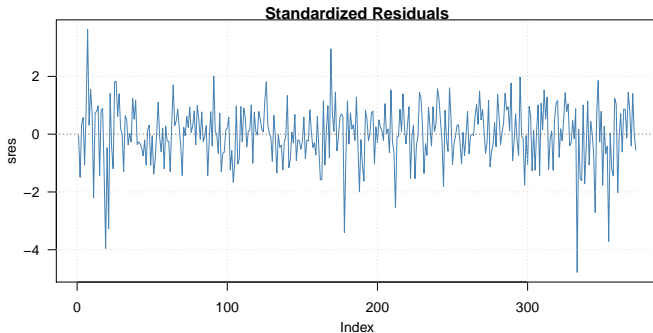
Background

ARCH Model

GARCH Model

IGARCH and
EGARCH Models

Stochastic Volatility
Model



Standardised Residuals Tests:

			Statistic	p-Value
Jarque-Bera Test	R	Chi ²	122.404	0
Shapiro-Wilk Test	R	W	0.9647625	8.273101e-08
Ljung-Box Test	R	Q(10)	13.72604	0.1858587
Ljung-Box Test	R	Q(15)	22.31714	0.09975386
Ljung-Box Test	R	Q(20)	23.88257	0.2475594
Ljung-Box Test	R ²	Q(10)	12.50025	0.25297
Ljung-Box Test	R ²	Q(15)	30.11276	0.01152131
Ljung-Box Test	R ²	Q(20)	31.46404	0.04935483
LM Arch Test	R	TR ²	22.036	0.0371183

- Jarque-Berg & Shapiro-Wilk Tests: Normality
- Lagrange multiplier (LM) Test: ARCH Effects

[Background](#)[ARCH Model](#)[GARCH Model](#)[IGARCH and
EGARCH Models](#)[Stochastic Volatility
Model](#)

ARCH(1) Model with Student-t Innovations

Error Analysis:

	Estimate	Std. Error	t value	Pr(> t)	
mu	0.021571	0.006054	3.563	0.000366	***
omega	0.013424	0.001968	6.820	9.09e-12	***
alpha1	0.259867	0.119901	2.167	0.030209	*
shape	5.985979	1.660030	3.606	0.000311	***

Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

Log Likelihood:

242.9678 normalized: 0.6531391

Description:

Mon Oct 18 15:27:27 2021 by user:

Standardised Residuals Tests:

			Statistic	p-Value
Jarque-Bera Test	R	Chi^2	130.8931	0
Shapiro-Wilk Test	R	W	0.9637533	5.744995e-08
Ljung-Box Test	R	Q(10)	14.31288	0.1591926
Ljung-Box Test	R	Q(15)	23.34043	0.07717449
Ljung-Box Test	R	Q(20)	24.87286	0.2063387
Ljung-Box Test	R^2	Q(10)	15.35917	0.1195054
Ljung-Box Test	R^2	Q(15)	33.96318	0.003446127
Ljung-Box Test	R^2	Q(20)	35.46828	0.01774746
LM Arch Test	R	TR^2	24.11517	0.01961957

Background

ARCH Model

GARCH Model

IGARCH and
EGARCH Models

Stochastic Volatility
Model

GARCH: Generalized Autoregressive Conditional Heteroskedasticity

For a log-return series r_t , let $a_t = r_t - \mu_t$ denote the innovation at time t .

$$a_t = \sigma_t \varepsilon_t, \quad \sigma_t^2 = \alpha_0 + \sum_{i=1}^m \alpha_i a_{t-i}^2 + \sum_{j=1}^s \beta_j \sigma_{t-j}^2$$

where $\varepsilon_t \sim \text{i.i.d.}(0, 1)$, $\alpha_0 > 0$, $\alpha_i, \beta_j \geq 0$, and $\sum_{i=1}^{\max(m,s)} (\alpha_i + \beta_i) < 1$ ensures stationarity

Interpretation: Past shocks (a_{t-i}^2) and past volatility (σ_{t-j}^2) drive current volatility

ARCH reacts only to past shocks, while **GARCH** also accounts for past volatility - giving smoother and more realistic dynamics in financial returns

Background

ARCH Model

GARCH Model

IGARCH and
EGARCH Models

Stochastic Volatility
Model

Model:

$$\sigma_t^2 = \alpha_0 + \alpha_1 a_{t-1}^2 + \beta_1 \sigma_{t-1}^2$$

Key Properties:

- **Stationarity:** holds if $0 \leq \alpha_1, \beta_1 \leq 1$ and $\alpha_1 + \beta_1 < 1$
- **Volatility clustering:** large shocks tend to be followed by large shocks
- **Heavy tails:**

$$\frac{\mathbb{E}(a_t^4)}{[\mathbb{E}(a_t^2)]^2} = \frac{3[1 - (\alpha_1 + \beta_1)^2]}{1 - (\alpha_1 + \beta_1)^2 - 2\alpha_1^2} > 3$$

- **1-step forecast:**

$$\sigma_t^2(1) = \alpha_0 + \alpha_1 a_t^2 + \beta_1 \sigma_t^2$$

For multi-step ahead forecasts, use $a_t^2 = \sigma_t^2 \epsilon_t^2$ and rewrite the model as

$$\sigma_{t+1}^2 = \alpha_0 + (\alpha_1 + \beta_1)\sigma_t^2 + \alpha_1\sigma_t^2(\epsilon_t^2 - 1)$$

We have 2-step ahead volatility forecast

$$\sigma_t^2(2) = \alpha_0 + (\alpha_1 + \beta_1)\sigma_t^2(1)$$

In general, we have

$$\begin{aligned}\sigma_t^2(\ell) &= \alpha_0 + (\alpha_1 + \beta_1)\sigma_t^2(\ell - 1), \quad \ell > 1 \\ &= \frac{\alpha_0[1 - (\alpha_1 + \beta_1)^{\ell-1}]}{1 - \alpha_1 - \beta_1} + (\alpha_1 + \beta_1)^{\ell-1}\sigma_t^2(1)\end{aligned}$$

Therefore

$$\sigma_t^2(\ell) \rightarrow \frac{\alpha_0}{1 - \alpha_1 - \beta_1}, \quad \text{as } \ell \rightarrow \infty$$

[Background](#)[ARCH Model](#)[GARCH Model](#)[IGARCH and
EGARCH Models](#)[Stochastic Volatility
Model](#)

Intel Example Revisited

Error Analysis:

	Estimate	Std. Error	t value	Pr(> t)
mu	0.0163276	0.0062624	2.607	0.00913 **
omega	0.0010918	0.0005291	2.063	0.03907 *
alpha1	0.0802716	0.0281162	2.855	0.00430 **
beta1	0.8553014	0.0461374	18.538	< 2e-16 ***

Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

Log Likelihood:

239.5189 normalized: 0.6438681

Description:

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Standardised Residuals Tests:

			Statistic	p-Value
Jarque-Bera Test	R	Chi^2	156.5138	0
Shapiro-Wilk Test	R	W	0.9676933	2.471139e-07
Ljung-Box Test	R	Q(10)	9.805485	0.4577215
Ljung-Box Test	R	Q(15)	16.54435	0.346824
Ljung-Box Test	R	Q(20)	17.8005	0.6005484
Ljung-Box Test	R^2	Q(10)	0.5130171	0.9999925
Ljung-Box Test	R^2	Q(15)	10.24557	0.8040151
Ljung-Box Test	R^2	Q(20)	11.77988	0.9234441
LM Arch Test	R	TR^2	9.334459	0.6741288

Background

ARCH Model

GARCH Model

IGARCH and
EGARCH Models

Stochastic Volatility
Model

Volatility Series and Standardized Residuals

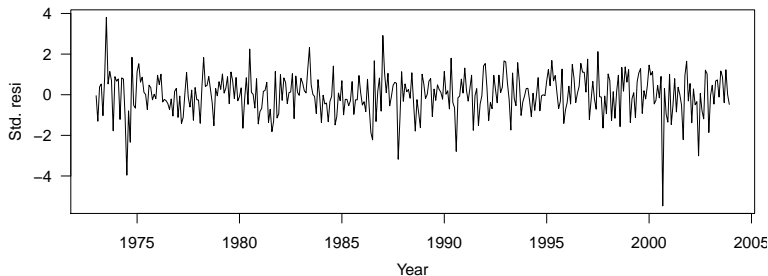
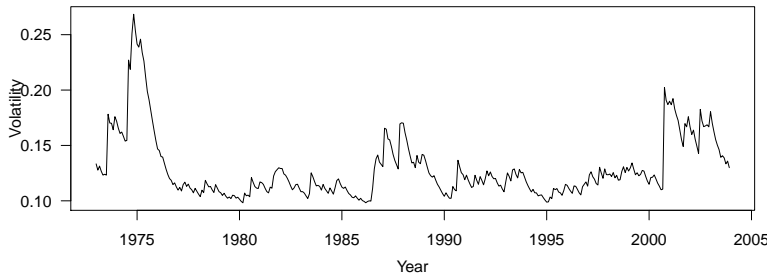
Background

ARCH Model

GARCH Model

IGARCH and
EGARCH Models

Stochastic Volatility
Model



GARCH Model Checking: ACF and PACF

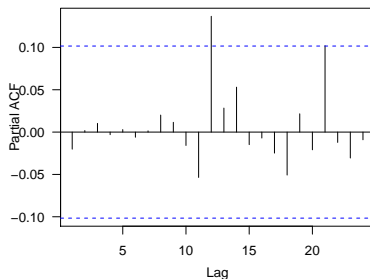
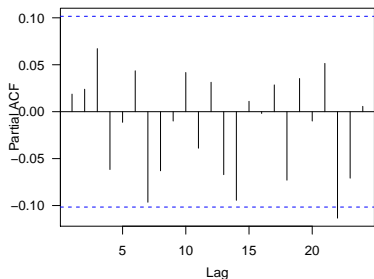
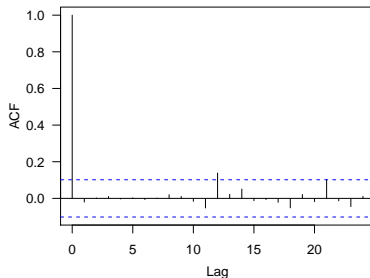
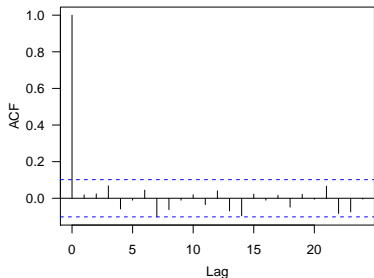
Background

ARCH Model

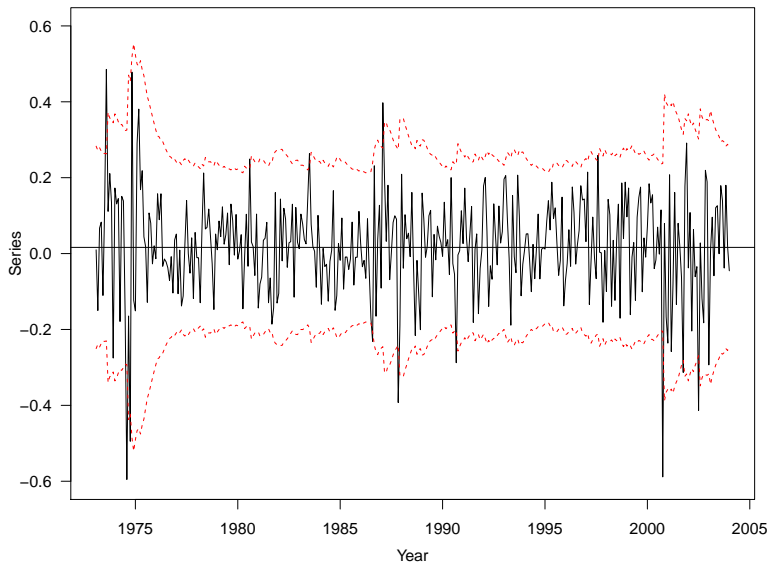
GARCH Model

IGARCH and
EGARCH Models

Stochastic Volatility
Model



95% Pointwise Prediction Intervals



Background

ARCH Model

GARCH Model

IGARCH and
EGARCH Models

Stochastic Volatility
Model

If the AR polynomial of the GARCH representation has **unit root** then we have an IGARCH model

An IGARCH(1, 1) model:

$$a_t = \sigma_t \epsilon_t, \quad \sigma_t^2 = \alpha_0 + \beta_1 \sigma_{t-1}^2 + (1 - \beta_1) a_{t-1}^2$$

ℓ -step ahead forecasts

$$\sigma_t^2(\ell) = \sigma_t(1)^2 + (\ell - 1)\alpha_0, \quad \ell \geq 1$$

\Rightarrow the effect of $\sigma_t^2(1)$ on future volatilities is persistent, and the volatility forecasts form a straight line with slope α_0

Background

ARCH Model

GARCH Model

IGARCH and
EGARCH Models

Stochastic Volatility
Model

Exponential GARCH (EGARCH) Model [Nelson, 1991]

Motivation: EGARCH captures asymmetric volatility effects-negative shocks often increase volatility more than positive shocks of the same size.

Innovation function:

$$g(\varepsilon_t) = \theta\varepsilon_t + \gamma(|\varepsilon_t| - \mathbb{E}|\varepsilon_t|), \quad \text{where } \mathbb{E}[g(\varepsilon_t)] = 0$$

Interpretation:

- If $\gamma > 0$, negative returns ($\varepsilon_t < 0$) cause higher volatility than positive ones.
- The term $|\varepsilon_t| - \mathbb{E}|\varepsilon_t|$ ensures mean zero.

Asymmetric form:

$$g(\varepsilon_t) = \begin{cases} (\theta + \gamma)\varepsilon_t - \gamma\mathbb{E}|\varepsilon_t|, & \varepsilon_t \geq 0, \\ (\theta - \gamma)\varepsilon_t - \gamma\mathbb{E}|\varepsilon_t|, & \varepsilon_t < 0 \end{cases}$$

EGARCH(m, s) model:

$$a_t = \sigma_t \varepsilon_t, \quad \log(\sigma_t^2) = \alpha_0 + \frac{1 + \beta_1 B + \dots + \beta_{s-1} B^{s-1}}{1 - \alpha_1 B - \dots - \alpha_m B^m} g(\varepsilon_{t-1})$$

Model:

$$a_t = \sigma_t \epsilon_t, \quad (1 - \alpha B) \log(\sigma_t^2) = (1 - \alpha) \alpha_0 + g(\epsilon_{t-1}),$$

where the ϵ_t are i.i.d. standard normal. In this case,
 $\mathbb{E}(|\epsilon_t|) = \sqrt{\frac{2}{\pi}}$ and the model for $\log(\sigma_t^2)$ becomes

$$(1 - \alpha B) \log(\sigma_t^2) = \begin{cases} (1 - \alpha) \alpha_0 - \sqrt{\frac{2}{\pi}} \gamma + (\gamma + \theta) \epsilon_{t-1} & \text{if } \epsilon_{t-1} \geq 0, \\ (1 - \alpha) \alpha_0 - \sqrt{\frac{2}{\pi}} \gamma + (\gamma - \theta)(-\epsilon_{t-1}) & \text{if } \epsilon_{t-1} < 0 \end{cases}$$

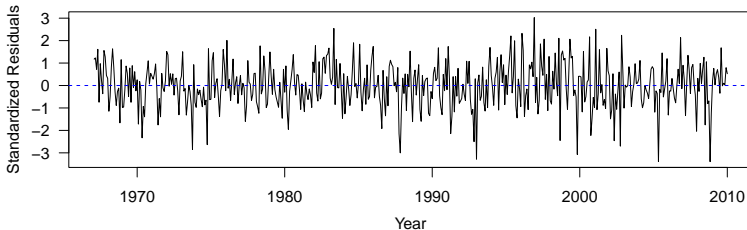
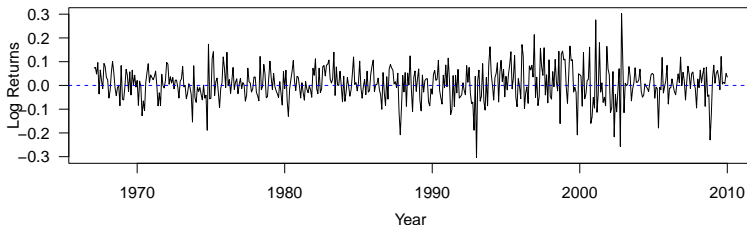
Finally, we have

$$\sigma_t^2 = \sigma_{t-1}^{2\alpha} \exp \left((1 - \alpha) \alpha_0 - \sqrt{\frac{2}{\pi}} \gamma \right) \begin{cases} \exp \left[(\gamma + \theta) \frac{a_{t-1}}{\sigma_{t-1}} \right] & \text{if } a_{t-1} \geq 0, \\ \exp \left[(\gamma - \theta) \frac{|a_{t-1}|}{\sigma_{t-1}} \right] & \text{if } a_{t-1} < 0. \end{cases}$$

[Background](#)[ARCH Model](#)[GARCH Model](#)[IGARCH and
EGARCH Models](#)[Stochastic Volatility
Model](#)

IBM Stock Example

We consider the monthly log returns of IBM stock from January 1967 to December 2009



Background

ARCH Model

GARCH Model

IGARCH and
EGARCH Models

Stochastic Volatility
Model

$$r_t = 0.067 + a_t, \quad a_t = \sigma_t \epsilon_t, \quad \epsilon_t \sim N(0, 1)$$

$$\log(\sigma_t^2) = -0.598 + 0.218(|\epsilon_{t-1}| - 0.423\epsilon_{t-1}) + 0.920 \log(\sigma_{t-1}^2)$$

Therefore, we have

$$\sigma_t^2 = \sigma_{t-1}^{2 \times 0.920} \exp(-0.598) \times \begin{cases} \exp(0.125) & \text{if } \epsilon_{t-1} \geq 0, \\ \exp(-0.310) & \text{if } \epsilon_{t-1} < 0. \end{cases}$$

For example, for a standardized shock with magnitude 2 (i.e., two standard deviations), we have

$$\frac{\sigma_t^2(\epsilon_{t-1} = -2)}{\sigma_t^2(\epsilon_{t-1} = 2)} = \frac{\exp(-0.31 \times (-2))}{\exp(0.125 \times 2)} = e^{0.37} = 1.448$$

Therefore, the impact of a negative shock of size two standard deviations is about 44.8% higher than that of a positive shock of the same size

[Background](#)[ARCH Model](#)[GARCH Model](#)[IGARCH and
EGARCH Models](#)[Stochastic Volatility
Model](#)

A simple stochastic volatility (SV) model assumes time-varying volatility driven by a latent process:

$$a_t = \sigma_t \varepsilon_t, \quad (1 - \alpha_1 B - \dots - \alpha_m B^m) \log(\sigma_t^2) = \alpha_0 + \nu_t$$

where:

- $\varepsilon_t \sim \text{i.i.d. } N(0, 1)$ are shocks to returns
- $\nu_t \sim \text{i.i.d. } N(0, \sigma_\nu^2)$ are shocks to volatility
- ε_t and ν_t are independent

Key idea: Volatility follows its own stochastic process rather than being a deterministic function of past shocks (as in GARCH)

Background

ARCH Model

GARCH Model

IGARCH and
EGARCH Models

Stochastic Volatility
Model

Long-Memory Stochastic Volatility (LMSV) Model

Model:

$$a_t = \sigma_t \varepsilon_t, \quad \sigma_t = \sigma \exp(u_t/2), \quad (1 - B)^d u_t = \eta_t, \quad 0 < d < 0.5$$

where $\varepsilon_t \sim N(0, 1)$, $\eta_t \sim N(0, \sigma_\eta^2)$, and they are independent

By taking logs,

$$\begin{aligned} \log(a_t^2) &= \log(\sigma^2) + u_t + \log(\varepsilon_t^2) \\ &= [\log(\sigma^2) + \mathbb{E}(\log(\varepsilon_t^2))] + u_t + [\log(\varepsilon_t^2) - \mathbb{E}(\log(\varepsilon_t^2))] \\ &= \mu + u_t + e_t \end{aligned}$$

so $\log(a_t^2)$ is a **Gaussian long-memory signal** plus a non-Gaussian noise

Inference: Because σ_t (or u_t) is **unobserved**, the likelihood is intractable. Hence, estimation is typically done using **MCMC or particle filtering**

In the R session, the `stochvol` package uses **MCMC sampling** to approximate the posterior of volatility paths and parameters

Background

ARCH Model

GARCH Model

IGARCH and
EGARCH Models

Stochastic Volatility
Model