

Lecture 5


Time Series Regression

Reading: Forecasting, Time Series, and Regression (4th edition) by Bowerman, O'Connell, and Koehler: Chapter 6

MATH 4070: Regression and Time-Series Analysis

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Time Series Regression

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Time Series Data
Trend Estimation
Estimating Seasonality

5.1

Notes


Notes

Notes

Agenda

- 1 Time Series Data
- 2 Trend Estimation
- 3 Estimating Seasonality

Time Series Regression

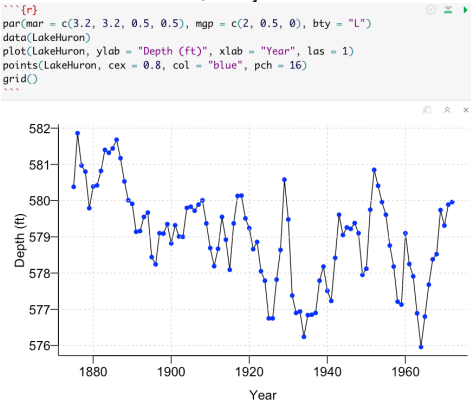
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Estimating Seasonality


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Level of Lake Huron 1875–1972

Annual measurements of the level of Lake Huron in feet.
[Source: Brockwell & Davis, 1991]



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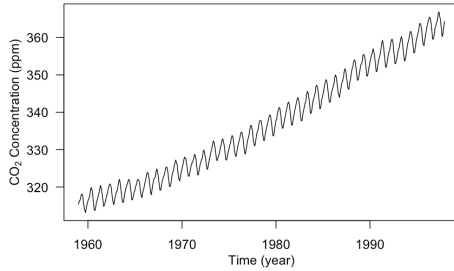
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Mauna Loa Atmospheric CO₂ Concentration

Monthly atmospheric concentrations of CO₂ at the Mauna Loa Observatory [Source: [Keeling & Whorf](#), [Scripps Institution of Oceanography](#)]

```
## [r]
data(co2)
par(mar = c(3.8, 4, 0.8, 0.6))
plot(co2, las = 1, xlab = "", ylab = "")
mtext("Time (year)", side = 1, line = 2)
mtext(expression(paste("CO"[2], " Concentration (ppm)")), side = 2, line = 2.5)
```



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Time Series Data

- A **time series** is a collection of observations $\{y_t, t \in T\}$ taken sequentially in time (t) with the index set T
 - $T = \{0, 1, 2, \dots, T\} \subset \mathbb{Z} \Rightarrow$ **discrete-time time series**
 - $T = [0, T] \subset \mathbb{R} \Rightarrow$ **continuous-time time series**
- A discrete-time time series might be intrinsically discrete or might arise from a underlying continuous-time time series via
 - sampling (e.g., instantaneous wind speed)
 - aggregation (e.g., daily accumulated precipitation amount)
 - extrema (e.g., daily maximum temperature)
- We will focus on dealing with **discrete-time real-valued** ($Y_t \in \mathbb{R}$) **time series** in this course

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Exploratory Time Series Analysis

- Start with a **time series plot**, i.e., to plot y_t versus t
- Look at the following:
 - Are there abrupt changes?
 - Are there "outliers"?
 - Is there a need to transform the data?
- Examine the **trend**, **seasonal components**, and the "noise" term

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Features of Times Series

- Trends (μ_t)
 - μ_t represents continuous changes, usually in the mean, over longer time scales. *"The essential idea of trend is that it shall be smooth."* - [Kendall, 1973]
 - The form of the trend is typically unknown and needs to be estimated. Removing the trend yields a **detrended** series
- Seasonal or Periodic Components (s_t)
 - s_t repeats consistently over time, i.e., $s_t = s_{t+kd}$
 - The form and period d of the seasonal component must be estimated to **deseasonalize** the series.
- The "Noise" Process (η_t)
 - η_t represents the component that is neither trend nor seasonality
 - Focus on finding plausible statistical models for this process

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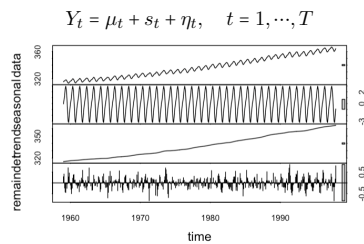
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Notes

Combining Trend, Seasonality, and Noise Together

There are two commonly used approaches

- Additive model:



Estimating Trend for Nonseasonal Model

Assuming $s_t = 0$ (i.e., there is no “seasonal” variation), we have


$$Y_t = \mu_t + \eta_t,$$

with $\mathbb{E}(\eta_t) = 0$

Methods for **estimating trends**

- Least squares regression
- Smoothing

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Notes

Trend Estimation: Linear Regression

- The additive nonseasonal time series model for $\{Y_t\}$ is


$$Y_t = \mu_t + \eta_t,$$

where the trend is assumed to be a linear combination of known covariate series $\{x_{it}\}_{i=1}^p$

$$\mu_t = \beta_0 + \sum_{i=1}^p \beta_i x_{it}.$$

- Here we want to **estimate** $\beta = (\beta_0, \beta_1, \dots, \beta_p)^T$ from the data $\{y_t, \{x_{it}\}_{i=1}^p\}_{t=1}^T$
- You're likely quite familiar with this formulation already \Rightarrow **Regression Analysis**

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Some Examples of Covariate Series $\{x_{it}\}$

- **Simple linear regression model:**

$$\mu_t = \beta_0 + \beta_1 x_t,$$

for example, the temperature trend at time t could be a constant (β_0) plus a multiple (β_1) of the carbon dioxide level at time t (x_t)


- **Polynomial regression model:**

$$\mu_t = \beta_0 + \sum_{i=1}^p \beta_i t^i$$

- **Change point model:**

$$\mu_t = \begin{cases} \beta_0 & \text{if } t \leq t^*; \\ \beta_0 + \beta_1 & \text{if } t \geq t^*. \end{cases}$$

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Notes

Parameter Estimation: Ordinary Least Squares

- Like in the linear regression setting, we can estimate the parameters via **ordinary least squares (OLS)**

- Specifically, we minimize the following objective function:

$$\ell_{ols} = \sum_{t=1}^T (y_t - \beta_0 - \sum_{k=1}^p x_{kt} \beta_k)^2.$$

- The estimates $\beta = (\beta_0, \beta_1, \dots, \beta_p)^T$ minimizing the above objective function are called the **OLS estimates of β** \Rightarrow they are easiest to express in **matrix form**

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The Model and Parameter Estimates in Matrix Form

- Matrix representation:

$$Y = X\beta + \eta,$$

where $Y = \begin{bmatrix} Y_1 \\ \vdots \\ Y_T \end{bmatrix}$, $X = \begin{bmatrix} 1 & x_{11} & x_{12} & \cdots & x_{1p} \\ 1 & x_{21} & x_{22} & \cdots & x_{2p} \\ \vdots & \vdots & \vdots & \cdots & \vdots \\ 1 & x_{T1} & x_{T2} & \cdots & x_{Tp} \end{bmatrix}$, and

$$\eta = \begin{bmatrix} \eta_1 \\ \vdots \\ \eta_T \end{bmatrix}$$

- Assuming $X^T X$ is **invertible**, the OLS estimate of β can be shown to be

$$\hat{\beta} = (X^T X)^{-1} X^T Y,$$

and the `lm` function in R calculates OLS estimates

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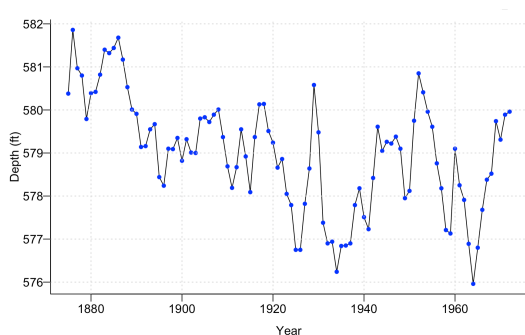
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Notes

Lake Huron Example Revisited



Let's **assume** there is a **linear trend in time** \Rightarrow we need to estimate the **intercept** β_0 and **slope** β_1

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The R Output


```
Call:
lm(formula = LakeHuron ~ yr)

Residuals:
    Min       1Q   Median       3Q      Max
-2.50997 -0.72726  0.00083  0.74402  2.53565

Coefficients:
            Estimate Std. Error t value Pr(>|t|)
(Intercept)  625.554918   7.764293   80.568 < 2e-16 ***
yr          -0.024201   0.004036  -5.996 3.55e-08 ***
---
Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

Residual standard error: 1.13 on 96 degrees of freedom
Multiple R-squared:  0.2725,    Adjusted R-squared:  0.2649
F-statistic: 35.95 on 1 and 96 DF,  p-value: 3.545e-08
```

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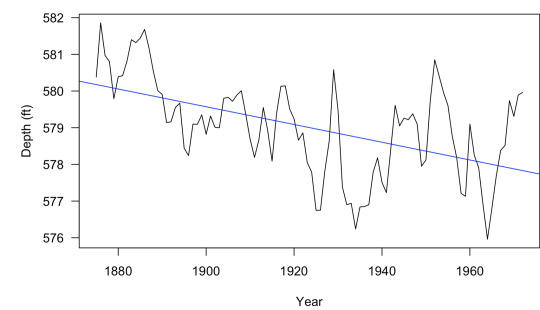
Trend Estimation

Estimating Seasonality

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Notes

Plot the (Estimated) Trend $\hat{\mu}_t = \hat{\beta}_0 + \hat{\beta}_1 t$



$\hat{\beta}_1 = -0.0242$ (ft/yr) \Rightarrow there seems to be a decreasing trend

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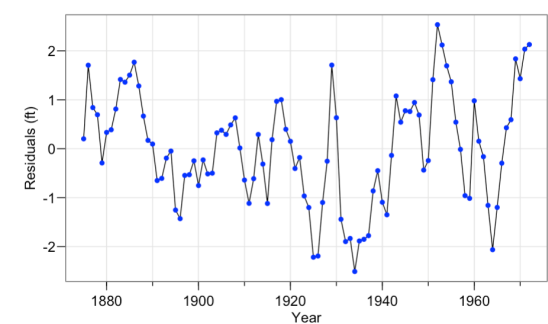
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Notes

Plot the Residuals $\{\hat{\eta}_t = y_t - \hat{\beta}_0 - \hat{\beta}_1 t\}$



$\{\hat{\eta}_t\}$ seems to exhibit some temporal dependence structure, should we worry about the results we have (recall OLS makes an i.i.d. assumption)?

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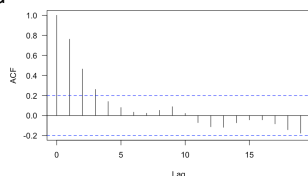
5.18

Notes

Statistical Properties of the OLS Estimates with Correlated Errors

- Assume the components of X are not random, the OLS estimates $\hat{\beta}$ are **unbiased** for β
Proof:

- Since $\{\eta_t\}$ is typically not an i.i.d. process (see the acf plot below), statistical inferences regarding β will be invalid



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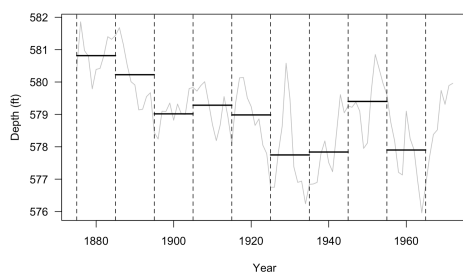
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Notes

Smoothing or Local Averaging

In some cases, we may relax the trend assumption using a 'non-parametric' approach.

We divide the time series into small blocks (each with 10 years of data) and average each block.



Doing this gives a very rough estimate of the trend. **Can we do better?**

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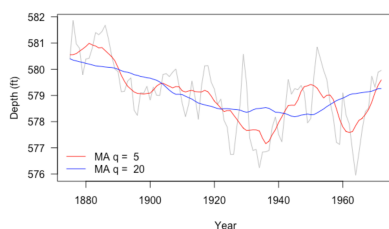
5.20

Notes

Moving Average Smoother

A **moving average smoother** estimates the trend at time t by averaging the current observation and the q nearest observations from either side. That is

$$\hat{\mu}_t = \frac{1}{2q+1} \sum_{j=-q}^q y_{t-j}$$



q is the "smoothing" parameter, which controls the smoothness of the estimated trend $\hat{\mu}_t$

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Notes

Exponential Smoothing

- Let $\alpha \in [0, 1]$ be some fixed constant, defined


$$\hat{\mu}_t = \begin{cases} Y_1 & \text{if } t = 1; \\ \alpha Y_t + (1 - \alpha)\hat{\mu}_{t-1} & t = 2, \dots, T. \end{cases}$$

- For $t = 2, \dots, T$, we can rewrite $\hat{\mu}_t$ as

$$\sum_{j=0}^{t-2} \alpha(1 - \alpha)^j Y_{t-j} + (1 - \alpha)^{t-1} Y_1.$$

⇒ it is a one-sided moving average filter with exponentially decreasing weights. One can alter α to control the amounts of smoothing (see next slide for an example)

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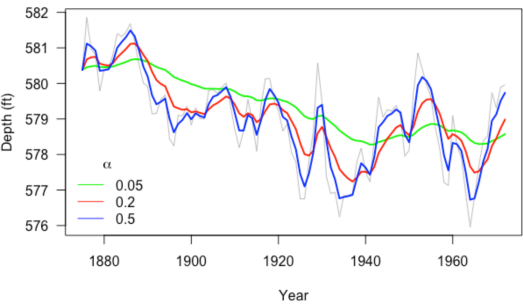
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Notes

α is the Smoothing Parameter for Exponential Smoothing



The smaller the α , the smoother the resulting trend

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Notes

Seasonal Component Estimation

Let's consider a situation where a time series consists of only a seasonal component (assuming the trend has been estimated/removed). In this scenario,


$$Y_t = s_t + \eta_t,$$

with $\{s_t\}$ having period d (i.e., $s_{t+jd} = s_t$ for all integers j and t), $\sum_{t=1}^d s_t = 0$ and $\mathbb{E}(\eta_t) = 0$

Two methods to estimate $\{s_t\}$

- Harmonic regression
- Seasonal mean model

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Notes

Harmonic Regression

- A harmonic regression model has the form

$$s_t = \sum_{j=1}^k A_j \cos(2\pi f_j t + \phi_j).$$

For each $j = 1, \dots, k$:

- $A_j > 0$ is the amplitude of the j -th cosine wave
 - f_j controls the frequency of the j -th cosine wave (how often waves repeats)
 - $\phi_j \in [-\pi, \pi]$ is the phase of the j -th wave (where it starts)
- The above can be expressed as

$$\sum_{j=1}^k (\beta_{1j} \cos(2\pi f_j) + \beta_{2j} \sin(2\pi f_j)),$$

where $\beta_{1j} = A_j \cos(\phi_j)$ and $\beta_{2j} = A_j \sin(\phi_j) \Rightarrow$ if $\{f_j\}_{j=1}^k$ are known, we can use regression techniques to estimate the parameters $\{\beta_{1j}, \beta_{2j}\}_{j=1}^k$

Time Series Regression

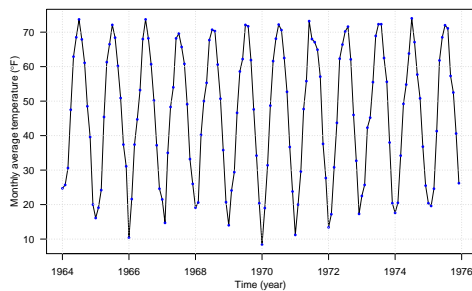
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Notes

Monthly Average Temperature in Dubuque, IA [Cryer & Chan, 2008]



Let's assume that there is no trend in this time series. In this context, our goal is to estimate s_t , the seasonal component.

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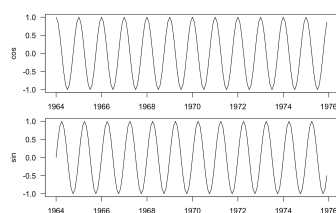
Use a Harmonic Regression to Model Annual Cycles

Model: $s_t = \beta_0 + \beta_1 \cos(2\pi t) + \beta_2 \sin(2\pi t)$

\Rightarrow annual cycles can be modeled by a linear combination of cos and sin with 1-year period.

In R, we can easily create these harmonics using the harmonic function in the TSA package

```
harmonics <- harmonic(tempdub, 1)
```



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Notes

R Code & Output

```
##{r}
harReg <- lm(tempdub ~ harmonics)
summary(harReg)

Call:
lm(formula = tempdub ~ harmonics)

Residuals:
    Min       1Q   Median       3Q      Max
-11.1580  -2.2756  -0.1457   2.3754  11.2671

Coefficients:
              Estimate Std. Error t value Pr(>|t|)
(Intercept)    46.2660     0.3088  149.816 < 2e-16 ***
harmonicscos(2*pi*t) -26.7079     0.4367  -61.154 < 2e-16 ***
harmonicssin(2*pi*t)  -2.1697     0.4367   -4.968 1.93e-06 ***
---
Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
```

Time Series Regression

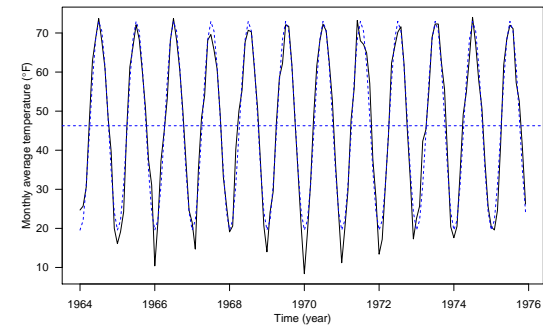
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The Harmonic Regression Model Fit



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Notes

Seasonal Means Model

- **Harmonics regression** assumes the seasonal pattern has a regular shape, i.e., the height of the peaks is the same as the depth of the troughs
- A less restrictive approach is to model $\{s_t\}$ as

$$s_t = \begin{cases} \beta_1 & \text{for } t = 1, 1 + d, 1 + 2d, \dots \\ \beta_2 & \text{for } t = 2, 2 + d, 2 + 2d, \dots \\ \vdots & \vdots \\ \beta_d & \text{for } t = d, 2d, 3d, \dots \end{cases}$$

- This is the **seasonal means** model, the parameters $(\beta_1, \beta_2, \dots, \beta_d)^T$ can be estimated under the linear model framework (think about ANOVA)

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Notes

R Output

```
Call:
lm(formula = tempdub ~ month - 1)

Residuals:
    Min       1Q   Median       3Q      Max
-8.2750 -2.2479  0.1125  1.8896  9.8250

Coefficients:
            Estimate Std. Error t value Pr(>|t|)
monthJanuary    16.608      0.987   16.83  <2e-16 ***
monthFebruary    20.650      0.987   20.92  <2e-16 ***
monthMarch       32.475      0.987   32.90  <2e-16 ***
monthApril       46.525      0.987   47.14  <2e-16 ***
monthMay         58.092      0.987   58.86  <2e-16 ***
monthJune        67.500      0.987   68.39  <2e-16 ***
monthJuly        71.717      0.987   72.66  <2e-16 ***
monthAugust      69.333      0.987   70.25  <2e-16 ***
monthSeptember   61.025      0.987   61.83  <2e-16 ***
monthOctober     50.975      0.987   51.65  <2e-16 ***
monthNovember    36.650      0.987   37.13  <2e-16 ***
monthDecember    23.642      0.987   23.95  <2e-16 ***
---
Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
```

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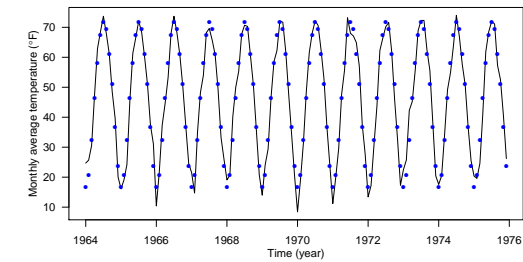
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The Seasonal Means Model Fit



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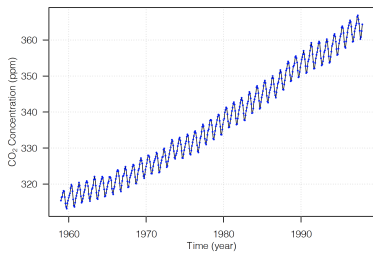
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Notes


Estimating the Trend and Seasonal variation Together



```
Let's perform a regression analysis to model both  $\mu_t$ 
(assuming a linear time trend) and  $s_t$  (using cos and sin)
```{r}
time <- as.numeric(time(co2))
harmonics <- harmonic(co2, 1)

lm_trendSeason <- lm(co2 ~ time + harmonics)
summary(lm_trendSeason)
```

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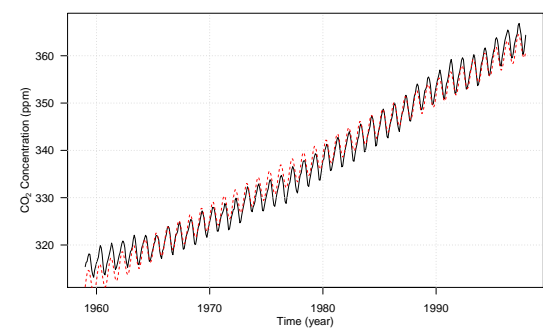
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The Regression Fit



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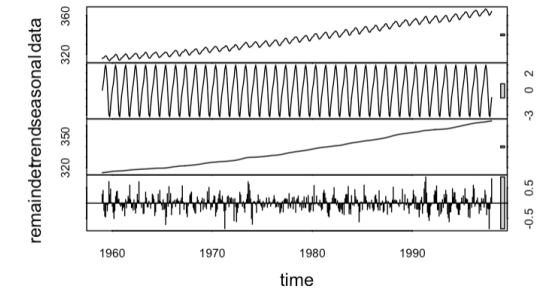
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Seasonal and Trend decomposition using Loess  
[Cleveland, et. al., 1990]

```
##{r}
Seasonal and Trend decomposition using Loess (STL)
par(mfrow = c(4, 3.6, 0.8, 0.6))
stl <- stl(co2, s.window = "periodic")
plot(stl, las = 1)
```



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Summary

These slides cover:

- Main features of a time series: trend, seasonality, and "noise"
- Estimating trends using multiple linear regression and "nonparametric" smoothing
- Estimating seasonality using harmonic regression and the seasonal mean model

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Notes

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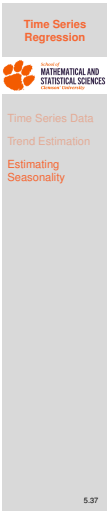
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R Functions to Know

- Visualizing time series data: `plot` (for `ts` objects), `ts.plot`, `tsplot` (`astsa` package)
- Fitting time series regression: `lm`, `harmonic` (`TSA` package) for creating harmonic predictors, `filter` for smoothing
- Seasonal and trend decomposition: `stl`



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