

Lecture 10

Univariate Volatility Modeling

Reading: An introduction to analysis of financial data with R
(2013) by Ruey Tsay

MATH 8090 Time Series Analysis

October 19 & October 21, 2021

Background

ARCH Model

GARCH Model

IGARCH and
EGARCH Models

Stochastic Volatility
Model

Whitney Huang
Clemson University

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- 4 IGARCH and EGARCH Models
- 5 Stochastic Volatility Model

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Apple Inc

\$146.55 ↑ 13.24% +17.14 YTD

After Hours: **\$146.75** (↑ 0.14%) +0.20

Closed: Oct 18, 8:12:21 PM UTC-4 · USD · NASDAQ · Disclaimer



Source: Google Finance

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Log Returns of Apple Stock

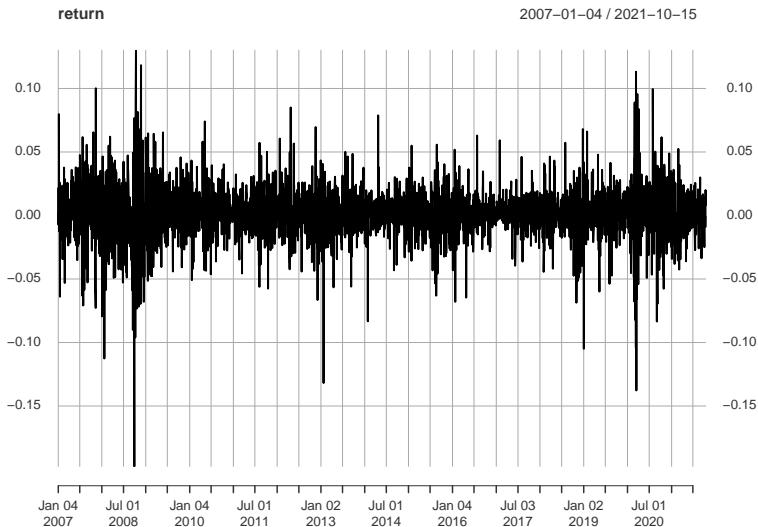
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Asset volatility is the degree of **variation** of a trading price series over time, usually measured by the **(conditional) standard deviation of (log) returns**

Why is volatility important?

- **Option pricing**, e.g., Black-Scholes formula
- **Risk management**, e.g., value at risk (VaR)
- **Asset allocation**, e.g., minimum-variance portfolio
- **Interval forecasts**

A key characteristic: Not directly observable

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We will take a **econometric approach** by modeling the conditional standard deviation (σ_t) of daily or monthly returns

Basic structure

$$r_t = \mu_t + a_t, \quad \mu_t = \phi_0 + \sum_{i=1}^p \phi_i r_{t-i} - \sum_{j=1}^q \theta_j a_{t-j}$$

Volatility models are concerned with time-evolution of

$$\sigma_t^2 = \text{Var}(r_t | F_{t-1}) = \text{Var}(a_t | F_{t-1}),$$

the conditional variance of a return

- Autoregressive conditional heteroscedastic (ARCH) model [Engle, 1982]
- Generalized Autoregressive Conditional Heteroskedasticity (GARCH) Model [Bollerslev, 1986]
- Integrated Generalized Autoregressive Conditional heteroskedasticity (IGARCH) model
- Exponential general autoregressive conditional heteroskedastic (EGARCH) model [Nelson, 1991]
- Asymmetric parametric ARCH models [Ding, Granger, and Engle, 1994]
- Stochastic volatility (SV) models [Melino and Turnbull, 1990; Harvey, Ruiz, and Shephard, 1994; Jacquier, Polson. and Rossi, 1994]

An ARCH(m) model:

$$a_t = \sigma_t \epsilon_t, \quad \sigma_t^2 = \alpha_0 + \alpha_1 a_{t-1}^2 + \cdots + \alpha_m a_{t-m}^2,$$

where $\{\epsilon_t\}$ is a sequence of i.i.d. r.v. with

- $\mathbb{E}(\epsilon_t) = 0$
- $\text{Var}(\epsilon_t) = 1$
- $\alpha_i \geq 0$ for $1 \leq i \leq m$
- Distribution: standard normal, standardize Student-t, generalized error distribution, or their skewed counterparts

Consider an ARCH(1) model

$$a_t = \sigma_t \epsilon_t, \quad \sigma_t^2 = \alpha_0 + \alpha_1 a_{t-1}^2,$$

where $\alpha_0 > 0$ and $\alpha_1 \geq 0$. We have the following properties:

- $\mathbb{E}(a_t) = 0$
- $\text{Var}(a_t) = \frac{\alpha_0}{1-\alpha_1}$ if $0 < \alpha_1 < 1$
- Under normality,

$$m_4 = \frac{3\alpha_0^2(1 + \alpha_1)}{(1 - \alpha_1(1 - 3\alpha_1^2))}$$

provided $0 < \alpha_1^2 < \frac{1}{3}$. \Rightarrow this implies heavy tails

Advantages:

- Simplicity
- Generate volatility clustering
- Heavy tails

Weaknesses:

- Symmetric between positive and negative returns
- Restrictive
- Not sufficiently adaptive in prediction

- 1 Modeling the mean effect μ_t and testing for ARCH effects for a_t

H_0 : no ARCH effects versus H_1 : ARCH effects

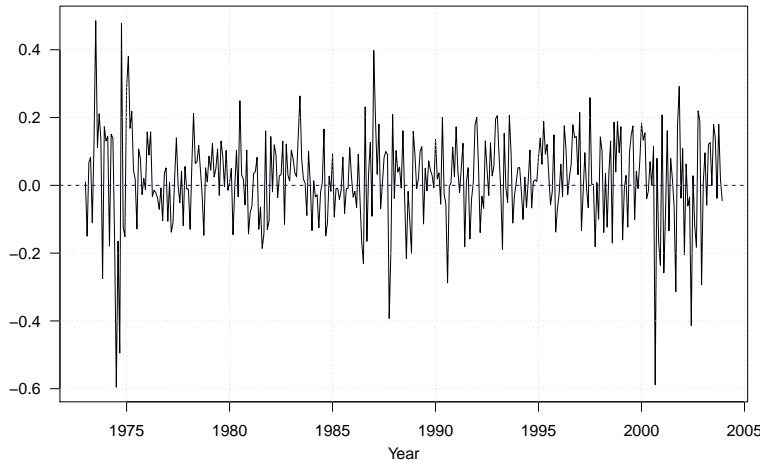
Use Q -statistics of squared residuals [Engle, 1982; McLeod and Li, 1983]

- 2 Order determination: use PACF of the squared residuals
- 3 Estimation: conditional MLE
- 4 Model checking: Q -statistics of standardized residuals and squared standardized residuals. Skewness and Kurtosis of standardized residuals

We use R package `fGarch` in this course

Example: Monthly Log Returns of Intel Stock

Here we use the monthly log returns of Intel stock to illustrate ARCH modeling



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Testing ARCH Effect

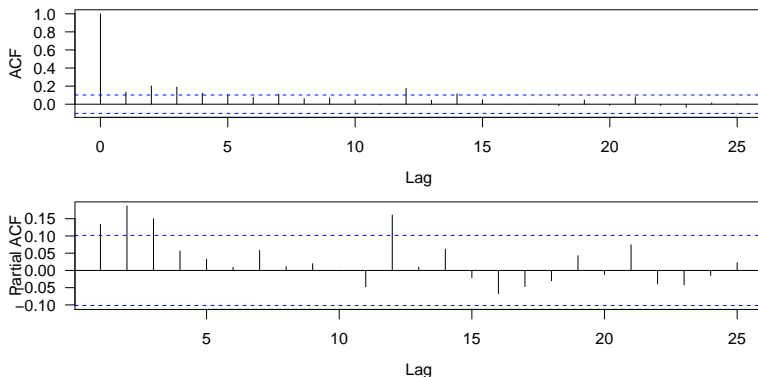
Here we test and examine the temporal pattern of the **squared residuals**

```
> Box.test(y^2, lag = 12, type = 'Ljung')
```

Box-Ljung test

data: y^2

X-squared = 68.67, df = 12, p-value = 5.676e-10



Here we fit an ARCH(3) for the volatility:

$$r_t = \mu + a_t, \quad a_t = \sigma_t \epsilon_t, \quad \sigma_t^2 = \alpha_0 + \sum_{i=1}^3 \alpha_i a_{t-i}^2,$$

assuming $\epsilon_t \stackrel{i.i.d.}{\sim} N(0, 1)$.

Error Analysis:

	Estimate	Std. Error	t value	Pr(> t)
mu	0.016572	0.006423	2.580	0.00988 **
omega	0.012043	0.001579	7.627	2.4e-14 ***
alpha1	0.208649	0.129177	1.615	0.10626
alpha2	0.071837	0.048551	1.480	0.13897
alpha3	0.049045	0.048847	1.004	0.31536

Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

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Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

Let's fit a simplified ARCH(1) model

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ARCH(1) Model Fitting

Error Analysis:

	Estimate	Std. Error	t value	Pr(> t)	
mu	0.016570	0.006161	2.689	0.00716	**
omega	0.012490	0.001549	8.061	6.66e-16	***
alpha1	0.363447	0.131598	2.762	0.00575	**

Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

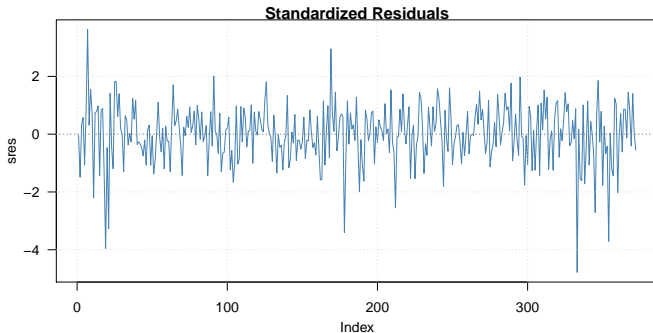
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Standardised Residuals Tests:

			Statistic	p-Value
Jarque-Bera Test	R	Chi ²	122.404	0
Shapiro-Wilk Test	R	W	0.9647625	8.273101e-08
Ljung-Box Test	R	Q(10)	13.72604	0.1858587
Ljung-Box Test	R	Q(15)	22.31714	0.09975386
Ljung-Box Test	R	Q(20)	23.88257	0.2475594
Ljung-Box Test	R ²	Q(10)	12.50025	0.25297
Ljung-Box Test	R ²	Q(15)	30.11276	0.01152131
Ljung-Box Test	R ²	Q(20)	31.46404	0.04935483
LM Arch Test	R	TR ²	22.036	0.0371183

ARCH(1) Model with Student-t Innovations

Error Analysis:

	Estimate	Std. Error	t value	Pr(> t)	
mu	0.021571	0.006054	3.563	0.000366	***
omega	0.013424	0.001968	6.820	9.09e-12	***
alpha1	0.259867	0.119901	2.167	0.030209	*
shape	5.985979	1.660030	3.606	0.000311	***

Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

Log Likelihood:

242.9678 normalized: 0.6531391

Description:

Mon Oct 18 15:27:27 2021 by user:

Standardised Residuals Tests:

			Statistic	p-Value
Jarque-Bera Test	R	Chi^2	130.8931	0
Shapiro-Wilk Test	R	W	0.9637533	5.744995e-08
Ljung-Box Test	R	Q(10)	14.31288	0.1591926
Ljung-Box Test	R	Q(15)	23.34043	0.07717449
Ljung-Box Test	R	Q(20)	24.87286	0.2063387
Ljung-Box Test	R^2	Q(10)	15.35917	0.1195054
Ljung-Box Test	R^2	Q(15)	33.96318	0.003446127
Ljung-Box Test	R^2	Q(20)	35.46828	0.01774746
LM Arch Test	R	TR^2	24.11517	0.01961957

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Generalized Autoregressive Conditional Heteroskedasticity (GARCH) Model

For a log return series r_t , let $a_t = r_t - \mu_t$ be the innovation at time t . Then a_t follows a GARCH(m, s) model if

$$a_t = \sigma_t \epsilon_t, \quad \sigma_t^2 = \alpha_0 + \sum_{i=1}^m \alpha_i a_{t-i}^2 + \sum_{j=1}^s \beta_j \sigma_{t-j}^2,$$

where $\{\epsilon_t\}$ is defined as before, $\alpha_0 > 0$, $\beta_j \geq 0$, and $\sum_{i=1}^{\max(m,s)} (\alpha_i + \beta_i) < 1$

Re-parameterization:

Let $\eta_t = a_t^2 - \sigma_t^2$. The GARCH model becomes

$$a_t^2 = \alpha_0 + \sum_{i=1}^{\max(m,s)} (\alpha_i + \beta_i) a_{t-i}^2 + \eta_t - \sum_{j=1}^s \beta_j \eta_{t-j}$$

This is an ARMA form for the squared series a_t^2

Model:

$$\sigma_t^2 = \alpha_0 + \alpha_1 a_{t-1}^2 + \beta_1 \sigma_{t-1}^2,$$

Properties

- Weak stationarity if $0 \leq \alpha_1, \beta_1 \leq 1, (\alpha_1 + \beta_1) < 1$
- Volatility clusters
- Heavy tails if $1 - 2\alpha_1^2 - (\alpha_1 + \beta_1)^2 > 0$, as

$$\frac{\mathbb{E}(a_t^4)}{[\mathbb{E}(a_t^2)]^2} = \frac{3[1 - (\alpha_1 + \beta_1)^2]}{1 - (\alpha_1 + \beta_1)^2 - 2\alpha_1^2} > 3$$

- 1-step ahead forecast

$$\sigma_h^2(1) = \alpha_0 + \alpha_1 a_h^2 + \beta_1 \sigma_h^2$$

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For multi-step ahead forecasts, use $a_t^2 = \sigma_t^2 \epsilon_t^2$ and rewrite the model as

$$\sigma_{t+1}^2 = \alpha_0 + (\alpha_1 + \beta_1)\sigma_t^2 + \alpha_1\sigma_t^2(\epsilon_t^2 - 1)$$

We have 2-step ahead volatility forecast

$$\sigma_h^2(2) = \alpha_0 + (\alpha_1 + \beta_1)\sigma_h^2(1)$$

In general, we have

$$\begin{aligned}\sigma_h^2(\ell) &= \alpha_0 + (\alpha_1 + \beta_1)\sigma_h^2(\ell - 1), \quad \ell > 1 \\ &= \frac{\alpha_0[1 - (\alpha_1 + \beta_1)^{\ell-1}]}{1 - \alpha_1 - \beta_1} + (\alpha_1 + \beta_1)^{\ell-1}\sigma_h^2(1)\end{aligned}$$

Therefore

$$\sigma_h^2(\ell) \rightarrow \frac{\alpha_0}{1 - \alpha_1 - \beta_1}, \quad \text{as } \ell \rightarrow \infty$$

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Intel Example Revisited

Error Analysis:

	Estimate	Std. Error	t value	Pr(> t)
mu	0.0163276	0.0062624	2.607	0.00913 **
omega	0.0010918	0.0005291	2.063	0.03907 *
alpha1	0.0802716	0.0281162	2.855	0.00430 **
beta1	0.8553014	0.0461374	18.538	< 2e-16 ***

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Log Likelihood:

239.5189 normalized: 0.6438681

Description:

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Standardised Residuals Tests:

			Statistic	p-Value
Jarque-Bera Test	R	Chi^2	156.5138	0
Shapiro-Wilk Test	R	W	0.9676933	2.471139e-07
Ljung-Box Test	R	Q(10)	9.805485	0.4577215
Ljung-Box Test	R	Q(15)	16.54435	0.346824
Ljung-Box Test	R	Q(20)	17.8005	0.6005484
Ljung-Box Test	R^2	Q(10)	0.5130171	0.9999925
Ljung-Box Test	R^2	Q(15)	10.24557	0.8040151
Ljung-Box Test	R^2	Q(20)	11.77988	0.9234441
LM Arch Test	R	TR^2	9.334459	0.6741288

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Volatility Series and Standardized Residuals

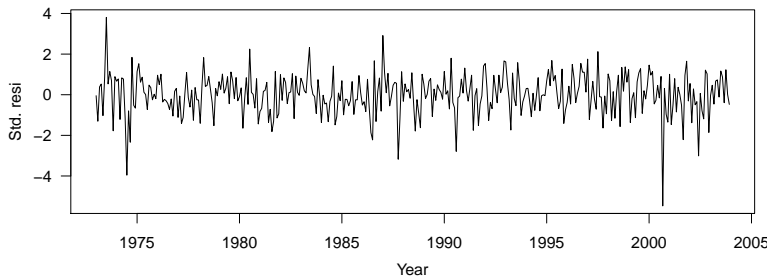
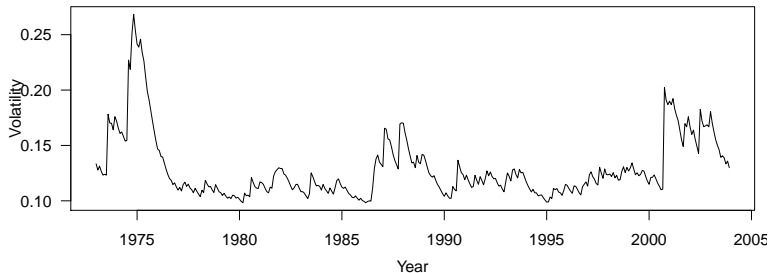
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GARCH Model Checking: ACF and PACF

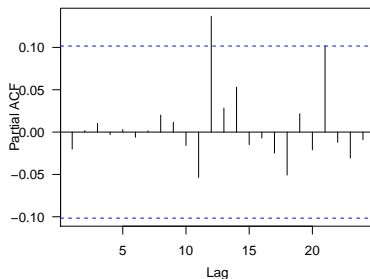
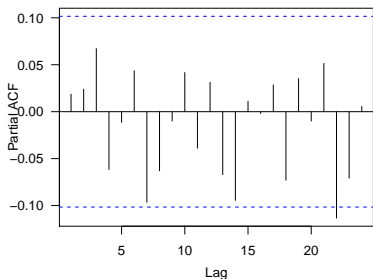
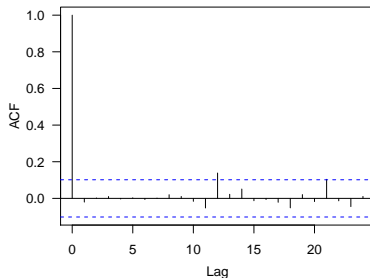
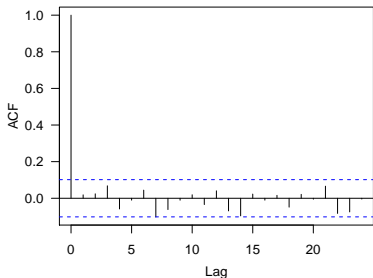
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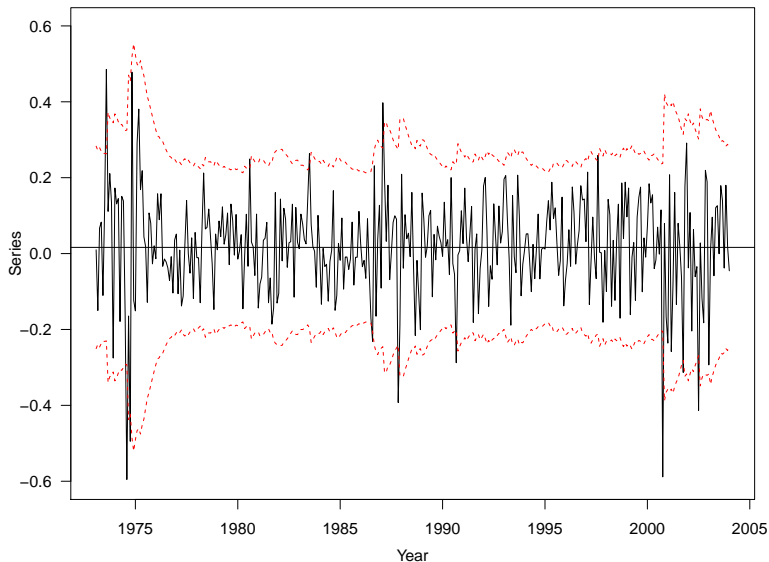
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95% Pointwise Prediction Intervals



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If the AR polynomial of the GARCH representation has **unit root** then we have an IGARCH model

An IGARCH(1, 1) model:

$$a_t = \sigma_t \epsilon_t, \quad \sigma_t^2 = \alpha_0 + \beta_1 \sigma_{t-1}^2 + (1 - \beta_1) a_{t-1}^2$$

ℓ -step ahead forecasts

$$\sigma_h^2(\ell) = \sigma_h^2(1) + (\ell - 1)\alpha_0, \quad \ell \geq 1$$

\Rightarrow the effect of $\sigma_h^2(1)$ on future volatilities is persistent, and the volatility forecasts form a straight line with slope α_0

The EGARCH model is able to capture **asymmetric effects** between positive and negative asset returns by considering the weight innovation

$$g(\epsilon_t) = \theta\epsilon_t + \gamma [|\epsilon_t| - \mathbb{E}(|\epsilon_t|)],$$

with $\mathbb{E}[g(\epsilon_t)] = 0$

We can see the asymmetry of $g(\epsilon_t)$ by rewriting it as

$$g(\epsilon_t) = \begin{cases} (\theta + \gamma)\epsilon_t - \gamma\mathbb{E}(|\epsilon_t|) & \text{if } \epsilon_t \geq 0, \\ (\theta - \gamma)\epsilon_t - \gamma\mathbb{E}(|\epsilon_t|) & \text{if } \epsilon_t < 0 \end{cases}$$

An EGARCH(m, s) model can be written as

$$a_t = \sigma_t \epsilon_t, \quad \log(\sigma_t^2) = \alpha_0 + \frac{1 + \beta_1 B + \dots + \beta_{s-1} B^{s-1}}{1 - \alpha_1 B - \dots - \alpha_m B^m} g(\epsilon_{t-1})$$

Model:

$$a_t = \sigma_t \epsilon_t, \quad (1 - \alpha B) \log(\sigma_t^2) = (1 - \alpha) \alpha_0 + g(\epsilon_{t-1}),$$

where the ϵ_t are i.i.d. standard normal. In this case,
 $\mathbb{E}(|\epsilon_t|) = \sqrt{\frac{2}{\pi}}$ and the model for $\log(\sigma_t^2)$ becomes

$$(1 - \alpha B) \log(\sigma_t^2) = \begin{cases} (1 - \alpha) \alpha_0 - \sqrt{\frac{2}{\pi}} \gamma + (\gamma + \theta) \epsilon_{t-1} & \text{if } \epsilon_{t-1} \geq 0, \\ (1 - \alpha) \alpha_0 - \sqrt{\frac{2}{\pi}} \gamma + (\gamma - \theta) (-\epsilon_{t-1}) & \text{if } \epsilon_{t-1} < 0 \end{cases}$$

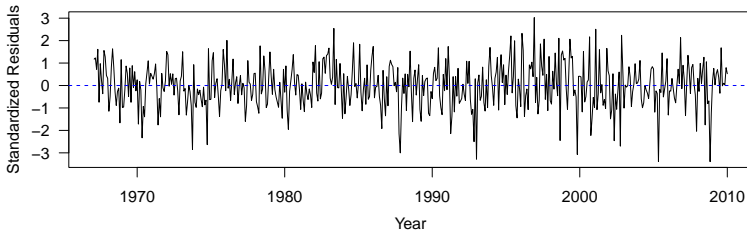
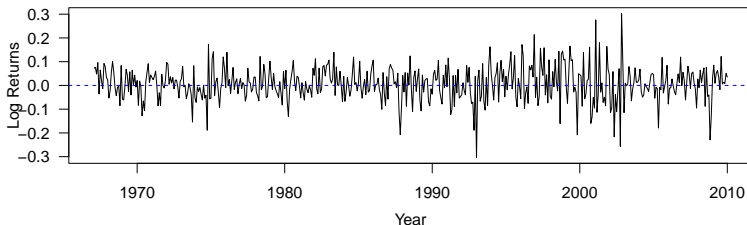
Finally, we have

$$\sigma_t^2 = \sigma_{t-1}^{2\alpha} \exp \left((1 - \alpha) \alpha_0 - \sqrt{\frac{2}{\pi}} \gamma \right) \begin{cases} \exp \left[(\gamma + \theta) \frac{a_{t-1}}{\sigma_{t-1}} \right] & \text{if } a_{t-1} \geq 0, \\ \exp \left[(\gamma - \theta) \frac{|a_{t-1}|}{\sigma_{t-1}} \right] & \text{if } a_{t-1} < 0. \end{cases}$$

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IBM Stock Example

We consider the monthly log returns of IBM stock from January 1967 to December 2009



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$$r_t = 0.067 + a_t, \quad a_t = \sigma_t \epsilon_t, \quad \epsilon_t \sim N(0, 1)$$

$$\log(\sigma_t^2) = -0.598 + 0.218(|\epsilon_{t-1}| - 0.424\epsilon_{t-1}) + 0.920 \log(\sigma_{t-1}^2)$$

Therefore, we have

$$\sigma_t^2 = \sigma_{t-1}^{2 \times 0.920} \exp(-0.598) \times \begin{cases} \exp(0.125) & \text{if } \epsilon_{t-1} \geq 0, \\ \exp(-0.310) & \text{if } \epsilon_{t-1} < 0. \end{cases}$$

For example, for a standardized shock with magnitude 2 (i.e., two standard deviations), we have

$$\frac{\sigma_t^2(\epsilon_{t-1} = -2)}{\sigma_t^2(\epsilon_{t-1} = 2)} = \frac{\exp(-0.31 \times (-2))}{\exp(0.125 \times 2)} = e^{0.37} = 1.448$$

Therefore, the impact of a negative shock of size two standard deviations is about 44.8% higher than that of a positive shock of the same size

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Stochastic Volatility (SV) Model

A (simple) SV model is

$$a_t = \sigma_t \epsilon_t, \quad (1 - \alpha_1 B - \dots - \alpha_m B^m) \log(\sigma_t^2) = \alpha_0 + \nu_t,$$

where ϵ_t 's are i.i.d. $N(0, 1)$, ν_t 's are i.i.d. $N(0, \sigma_\nu^2)$, $\{\epsilon_t\}$ and $\{\nu_t\}$ are independent

Long-memory SV Model:

$$a_t = \sigma_t \epsilon_t, \quad \sigma_t = \sigma \exp(u_t/2), \quad (1 - B)^d u_t = \eta_t,$$

where $\sigma > 0$, ϵ_t 's are i.i.d. $N(0, 1)$, η_t 's are i.i.d. $N(0, \sigma_\eta^2)$ and independent of ϵ_t , and $0 < d < 0.5$.

In LMSV, we have

$$\begin{aligned} \log(a_t^2) &= \log(\sigma^2) + u_t + \log(\epsilon_t^2) \\ &= [\log(\sigma^2) + \mathbb{E}(\log(\epsilon_t^2))] + u_t + [\log(\epsilon_t^2) - \mathbb{E}(\log(\epsilon_t^2))] \\ &= \mu + u_t + e_t \end{aligned}$$

Thus, the $\log(a_t^2)$ series is a **Gaussian long-memory** signal plus a non-Gaussian white noise