

Lecture 12

Time Series Analysis I

DSA 8020 Statistical Methods II

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Agenda

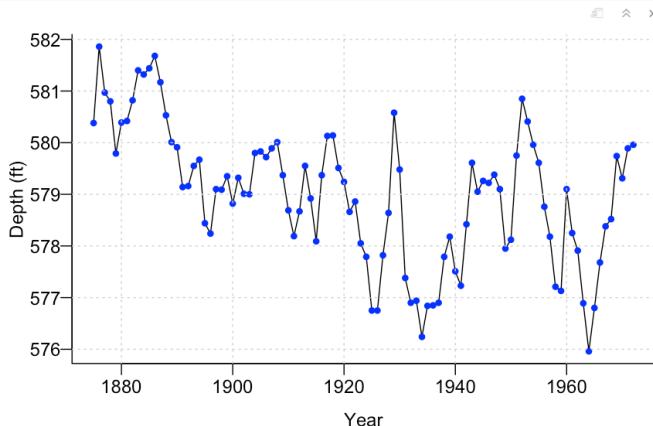
- 1 Time Series Data
- 2 Time Series Models
- 3 Objectives of Time Series Analysis
- 4 A Case Study

Level of Lake Huron 1875–1972

Annual measurements of the level of Lake Huron in feet.

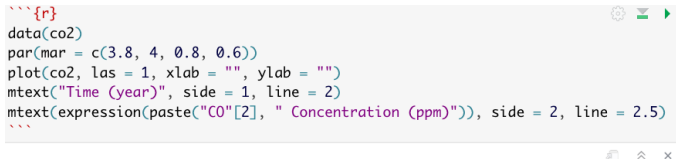
[Source: [Brockwell & Davis, 1991](#)]

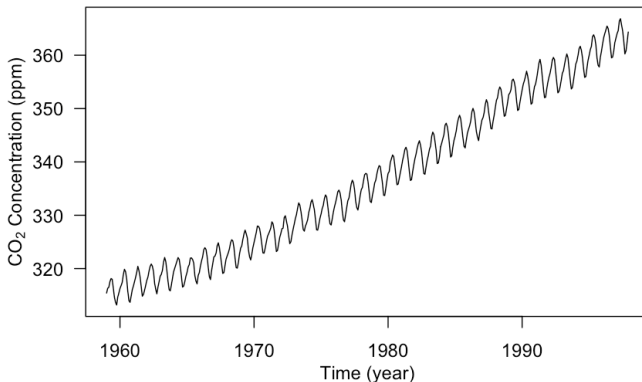
```
## {r}
par(mar = c(3.2, 3.2, 0.5, 0.5), mgp = c(2, 0.5, 0), bty = "L")
data(LakeHuron)
plot(LakeHuron, ylab = "Depth (ft)", xlab = "Year", las = 1)
points(LakeHuron, cex = 0.8, col = "blue", pch = 16)
grid()
```



Mauna Loa Atmospheric CO₂ Concentration

Monthly atmospheric concentrations of CO₂ at the Mauna Loa Observatory [Source: [Keeling & Whorf, Scripps Institution of Oceanography \(SIO\)](#)]

```
```\r\ndata(co2)\npar(mar = c(3.8, 4, 0.8, 0.6))\nplot(co2, las = 1, xlab = "", ylab = "")\nmtxt("Time (year)", side = 1, line = 2)\nmtxt(expression(paste("CO"[2], " Concentration (ppm)")), side = 2, line = 2.5)\n```\n
```



# US Unemployment Rate 1948 Jan. – 2021 July

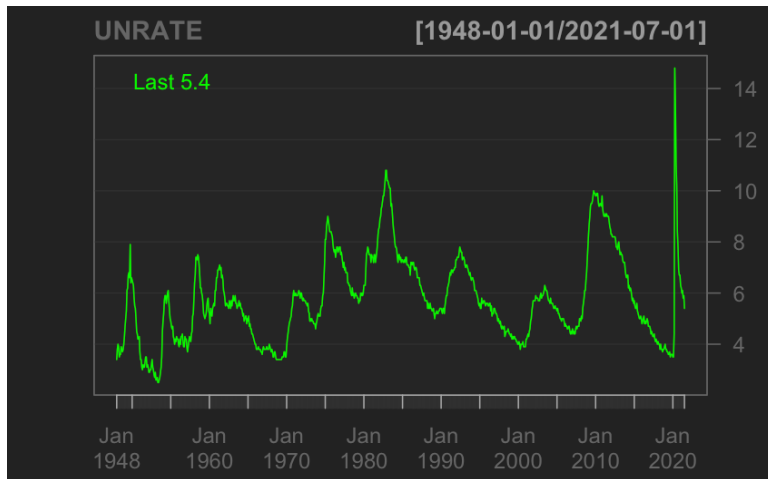
[Source: St. Louis Federal Reserve Bank's FRED system]

Time Series Data

Time Series Models

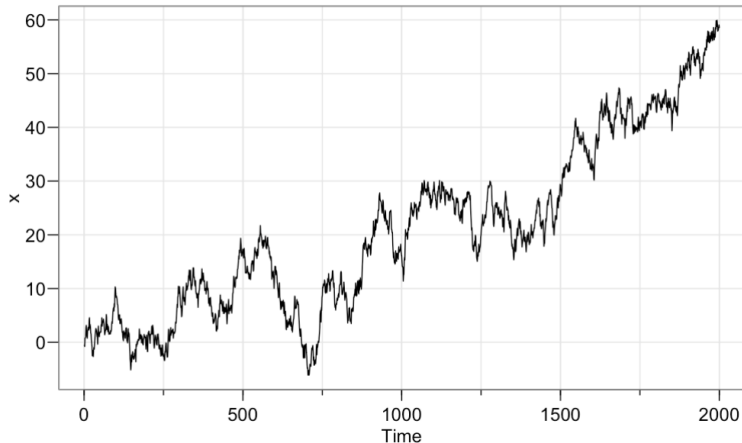
Objectives of Time  
Series Analysis

A Case Study




# A Simulated Time Series

```
##{r}
set.seed(123)
w <- rnorm(2000); x <- cumsum(w); tsplot(x, las = 1)
##
```



- A **time series** is a set of observations  $\{y_t, t \in T\}$  made sequentially in time ( $t$ ) with the index set  $T$ 
  - $T = \{0, 1, 2, \dots, T\} \subset \mathbb{Z} \Rightarrow$  **discrete-time time series**
  - $T = [0, T] \subset \mathbb{R} \Rightarrow$  **continuous-time time series**
- A discrete-time time series might be intrinsically discrete or might arise from a underlying continuous-time time series via
  - sampling (e.g., instantaneous wind speed)
  - aggregation (e.g., daily accumulated precipitation amount)
  - extrema (e.g., daily maximum temperature)
- We will focus on dealing with **discrete-time real-valued** ( $Y_t \in \mathbb{R}$ ) **time series**

- Start with a **time series plot**, i.e., to plot  $y_t$  versus  $t$  
- Look at the following:
  - Are there abrupt changes?
  - Are there “outliers”?
  - Is there a need to transform the data?
- Examine the **trend**, **seasonal components**, and the “noise” term



## ● Trends

- One can think of trend,  $\mu_t$ , as continuous changes, usually in the mean, over longer time scales  $\Rightarrow$  *“the essential idea of trend is that it shall be smooth”* - [Kendall, 1973]
- Usually the form of the trend is unknown and needs to be estimated. When the trend is removed, we obtain a **detrended** series

## ● Seasonal or periodic components

- A seasonal component  $s_t$  constantly repeats itself in time, i.e.,  $s_t = s_{t+kd}$
- We need to estimate the form and/or the period  $d$  of the seasonal component to **deseasonalize** the series

## ● The “noise” process

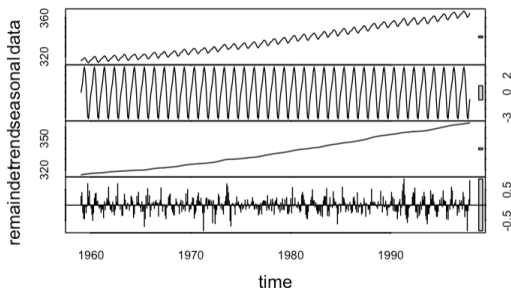
- The noise process,  $\eta_t$ , is the component that is neither trend nor seasonality
- We will focus on finding plausible (typically stationary) statistical models for this process

# Combining Trend, Seasonality, and Noise Together

There are two commonly used approaches

- Additive model:

$$y_t = \mu_t + s_t + \eta_t, \quad t = 1, \dots, T$$



- Multiplicative model:

$$y_t = \mu_t s_t \eta_t, \quad t = 1, \dots, T$$

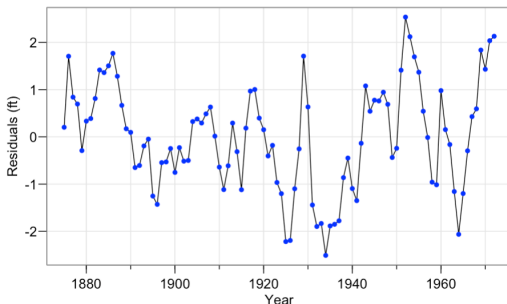
If all  $\{y_t\}$  are positive then we obtain the additive model by taking logarithms:

$$\log y_t = \log \mu_t + \log s_t + \log \eta_t, \quad t = 1, \dots, T$$

# Time Series Models

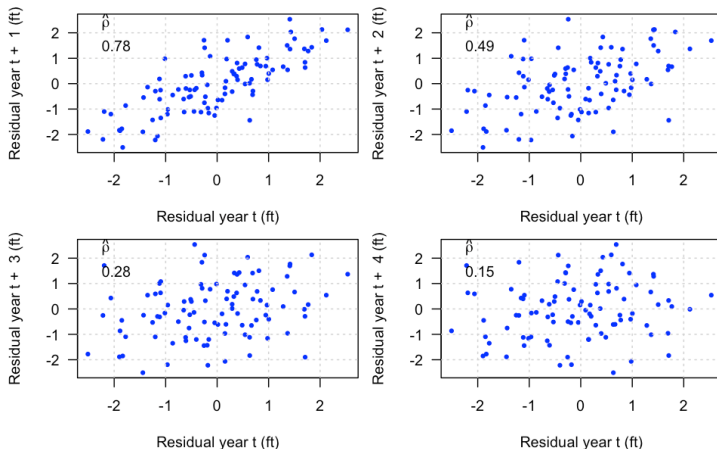
## Lake Huron Time Series

- **Time series analysis** is the area of statistics which deals with the analysis of **dependency** between different observations (typically  $\{\eta_t\}$ )
- Some key features of the Lake Huron time series:
  - decreasing trend
  - some “random” fluctuations around the decreasing trend
- We extract the “noise” component by assuming a linear trend



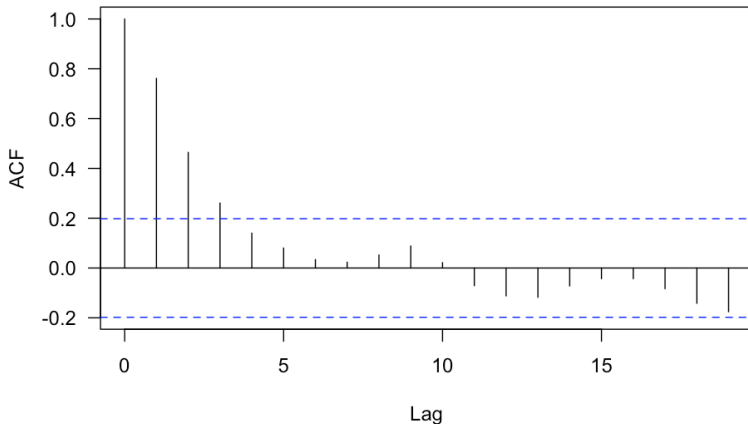
# Exploring the Temporal Dependence Structure of $\{\eta_t\}$

$\{\eta_t\}$  exhibit some temporal dependence structure, that is, the nearby (in time) values tend to be more alike than those far apart values. To see this, let's make a few time lag plots



## Further Exploration of the Temporal Dependence Structure

Let's plot the correlation as a function of the time lag



We will use this information to suggest an appropriate model

- A **time series model** is a probabilistic model for  $\{Y_t : t \in T\}$  that describes ways that the series data  $\{y_t\}$  could have been generated
- Will try to keep our models for  $\{Y_t\}$  simple by assuming **stationarity**  $\Rightarrow$  characteristic of the distribution of  $\{Y_t\}$  does not depend on the time points, only on the “time lag”
- While most time series are not stationary, one either remove or model the non-stationary parts (e.g., de-trend or de-seasonalization) so that we are only left with a stationary component  $\{\eta_t\}$ .

- The **mean function** of  $\{\eta_t\}$  is

$$\mu_t = E[\eta_t], \quad t \in T$$

- The **autocovariance function** of  $\{\eta_t\}$  is

$$\gamma(t, t') = \text{Cov}(\eta_t, \eta_{t'}) = E[(\eta_t - \mu_t)(\eta_{t'} - \mu_{t'})], \quad t, t' \in T,$$

when  $t = t'$  we obtain  $\gamma(t, t') = \text{Cov}(\eta_t, \eta_t) = \text{Var}(\eta_t) = \sigma_t^2$ ,  
the variance function of  $Y_t$



The autocorrelation function (ACF) of  $\{\eta_t\}$  is

$$\rho(t, t') = \text{Corr}(\eta_t, \eta_{t'}) = \frac{\gamma(t, t')}{\sqrt{\gamma(t, t)\gamma(t', t')}}.$$

It measures the strength of linear association between  $Y_t$  and  $Y_{t'}$

## Properties:

- 1  $-1 \leq \rho(t, t') \leq 1, \quad t, t' \in T$
- 2  $\rho(t, t') = \rho(t', t), \quad \forall t, t' \in T; \rho(t, t) = 1, \quad \forall t \in T$
- 3  $\rho(t, t')$  is a non-negative definite function

Partial autocorrelation function (PACF) is a conditional correlation, i.e., the correlation at two time points given the information at all other time points

We will try to keep our **models for  $\{\eta_t\}$**  as simple as possible by assuming **stationarity**, meaning that characteristic of  $\{\eta_t\}$  does not depend on the time points, only on the “time lag”:

- $E[\eta_t] = 0, \quad \forall t \in T$
- $\text{Cov}(\eta_t, \eta_{t'}) = \gamma(t' - t) = \text{Cov}(\eta_{t+s}, \eta_{t'+s})$

$\Rightarrow$  autocorrelation function (ACF):

$$\rho(h) = \frac{\gamma(h)}{\gamma(0)}$$

Let  $\{Z_t\}$  be independent and identical random variables that follow  $N(0, \sigma^2)$

- Moving Average Processes (MA(q)):

$$\eta_t = Z_t + \theta_1 Z_{t-1} + \theta_2 Z_{t-2} \cdots + \theta_q Z_{t-q}$$

Let  $\{Z_t\}$  be independent and identical random variables that follow  $N(0, \sigma^2)$

- Moving Average Processes (MA(q)):

$$\eta_t = Z_t + \theta_1 Z_{t-1} + \theta_2 Z_{t-2} \cdots + \theta_q Z_{t-q}$$

- Autoregressive Processes (AR(p)):

$$\eta_t = \phi_1 \eta_{t-1} + \phi_2 \eta_{t-2} + \cdots + \phi_p \eta_{t-p} + Z_t$$

Let  $\{Z_t\}$  be independent and identical random variables that follow  $N(0, \sigma^2)$

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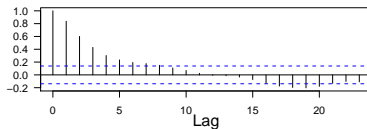
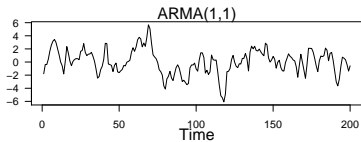
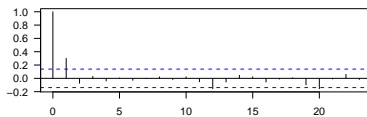
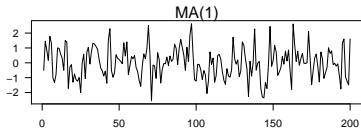
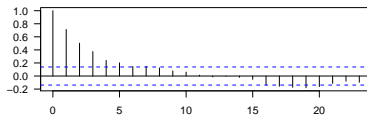
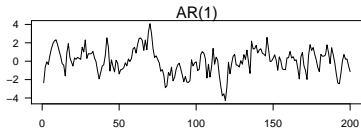
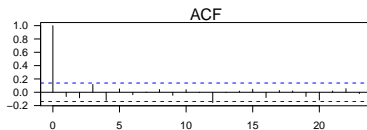
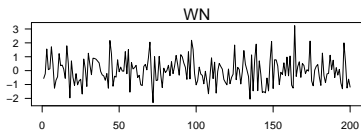
- Autoregressive Processes (AR(p)):

$$\eta_t = \phi_1 \eta_{t-1} + \phi_2 \eta_{t-2} + \cdots + \phi_p \eta_{t-p} + Z_t$$

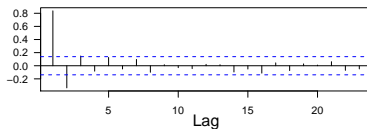
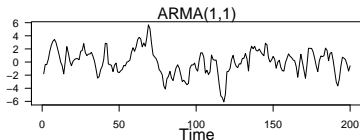
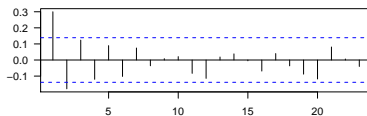
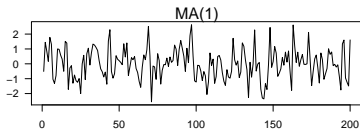
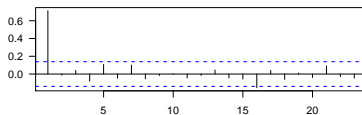
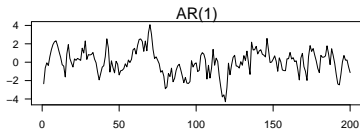
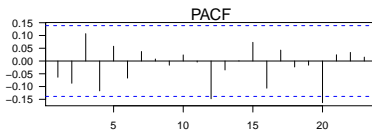
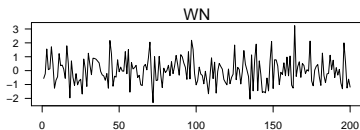
- Autoregressive Moving Average Processes ARMA(p,q):

$$\eta_t = \phi_1 \eta_{t-1} + \phi_2 \eta_{t-2} + \cdots + \phi_p \eta_{t-p} + Z_t + \theta_1 Z_{t-1} + \theta_2 Z_{t-2} + \cdots + \theta_q Z_{t-q}$$

# ACF Plots



# PACF Plots



## Identification of ARMA Models using ACF/PACF Plots

Use the ACF and PACF together to identify possible models. The following table gives some rough guidelines. Unfortunately, it's not a well-defined process and some guesswork is usually needed

	ACF	PACF
$AR(p)$	Tails off	Cuts off after lag $p$
$MA(q)$	Cuts off after lag $q$	Tails off
$ARMA(p, q)$	Tails off	Tails off



We wish to test:

$H_0 : \{e_1, e_2, \dots, e_T\}$  is an i.i.d. noise sequence;

$H_1 : H_0$  is false,

where  $\{e_t\}$  are the residuals after fitting a model to  $\{\eta_t\}$

**Test statistic:**

$$Q_{LB} = T(T-2) \sum_{h=1}^k \frac{\hat{\rho}^2(h)}{T-h} \sim \chi_k^2.$$

Ljung-Box test can be carried out in R using the function

`Box.test`

# Objectives of Time Series Analysis

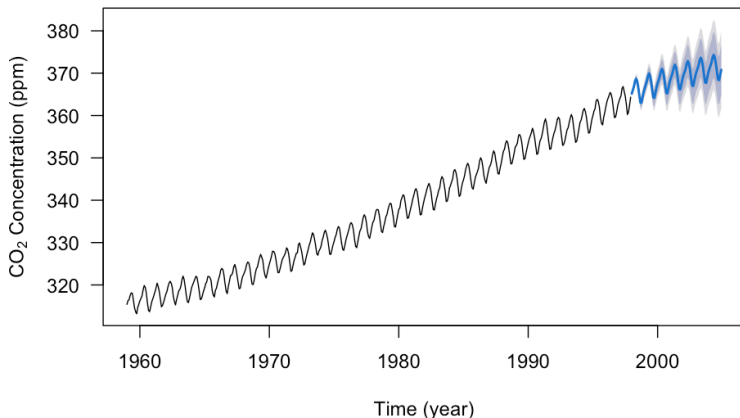
# Some Objectives of Time Series Analysis

- **Modeling:** Find a **statistical model** that adequately explains the observed time series
- For example, identify a model which can account for the fact that the depths of Lake Huron are correlated with different years and with a decreasing long-term trend
- The fitted model can be used for further **statistical inference**, for instance, to answer the question like: **Is there evidence of decreasing trend in the Lake Huron depths?**

## Some Objectives of Time Series Analysis, Cont'd

**Forecasting** is perhaps the most common objective. One observe a time series of given length and wish to **predict** or **forecast** future values of the time series based on those already observed.

Forecasts from TBATS(1, {3,1}, -, {<12,5>})



## Some Objectives of Time Series Analysis, Cont'd

- **Adjustment:** an example would be [seasonal adjustment](#), where the seasonal component is estimated and then removed in order to better understand the underlying trend
- **Simulation:** use a time series model (which adequately describes a physical process) as a surrogate to *simulate repeatedly in order to approximate how the physical process behaves*
- **Control:** adjust various [input \(control\)](#) parameters so that the time series fits closer to a given standard (many examples from statistical quality control)

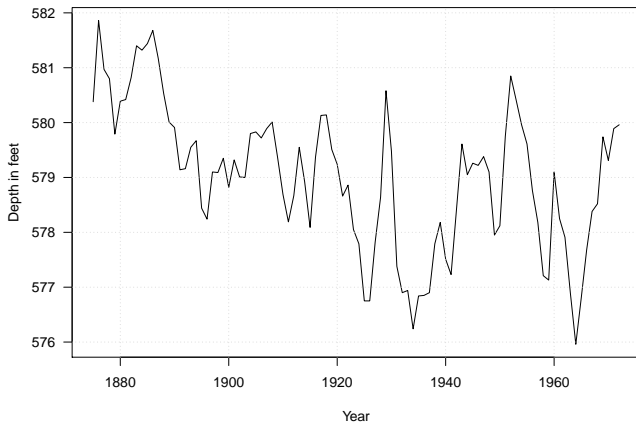
# Lake Huron Case Study



**Source:** <https://www.worldatlas.com/articles/what-states-border-lake-huron.html>

- Detrending
- Model fitting and selection
- Forecasting

# Annual Measurements of the Level of Lake Huron

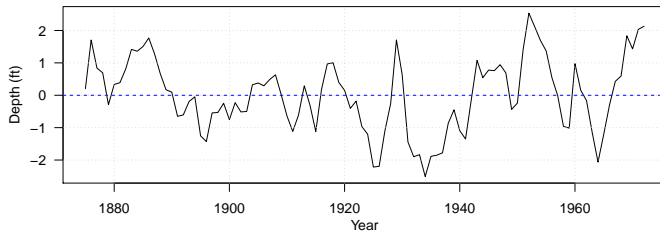
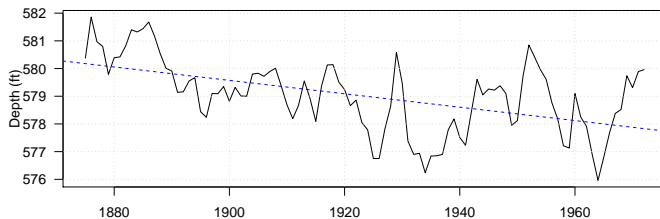


There seems to be a decreasing trend  $\Rightarrow$  need to estimate the trend to get the detrended series

# Plots of the Trend and Residuals

$$y_t = \underbrace{\mu_t}_{\text{trend}} + \underbrace{\eta_t}_{\text{residual}}$$

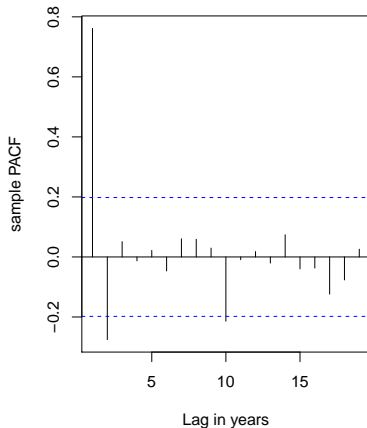
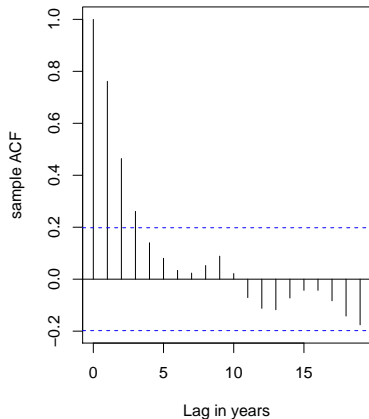
where we **assume**  $\mu_t = \alpha + \beta t$ , i.e., a **linear trend in time**



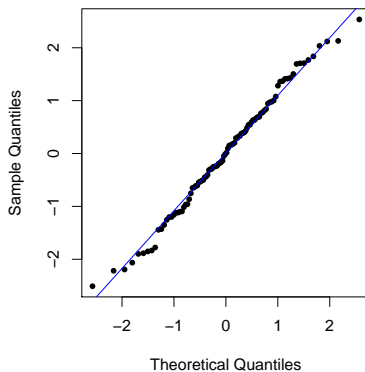
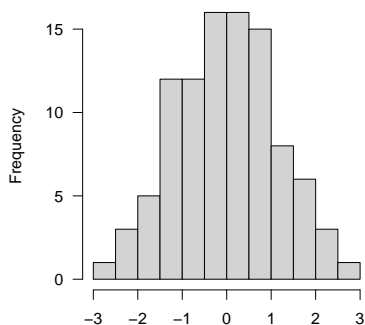


## ACF and PACF Plots

- Tapering pattern in ACF  $\Rightarrow$  need to include AR terms
- Significant PACF values at the first 2 lags  $\Rightarrow$  a AR(2) may be appropriate



# Assessing Normality Assumption for $\eta_t$



# Fitting AR(2)

```
> (ar2.model <- arima(deTrend, order = c(2, 0, 0)))
```

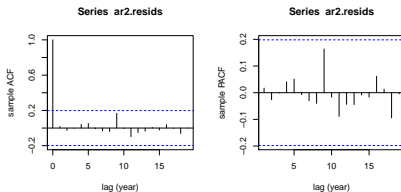
Call:

```
arima(x = deTrend, order = c(2, 0, 0))
```

Coefficients:

	ar1	ar2	intercept
	1.0047	-0.2919	0.0196
s.e.	0.0977	0.1004	0.2351

sigma^2 estimated as 0.4571: log likelihood = -101.25, aic = 210.5



```
> Box.test(ar2.resids, type = "Ljung-Box")
```

Box-Ljung test

data: ar2.resids

X-squared = 0.029966, df = 1, p-value = 0.8626

We can conduct model selection by using, for example, AIC

```
> ar1.model <- arima(deTrend, order = c(1, 0, 0))
> ar2.model <- arima(deTrend, order = c(2, 0, 0))
> arma21.model <- arima(deTrend, order = c(2, 0, 1))
> AIC(ar1.model); AIC(ar2.model); AIC(arma21.model)
[1] 216.5835
[1] 210.5032
[1] 212.1784
```

# Fitting AR(2) + a Linear Trend

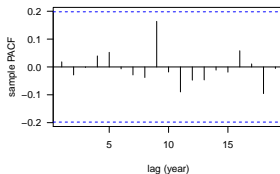
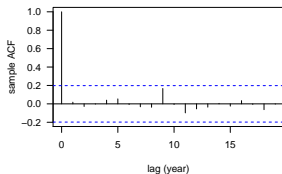
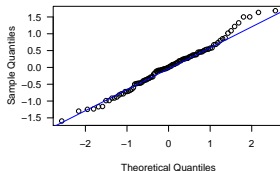
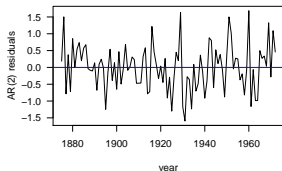
```
> library(forecast)
> (fit <- Arima(LakeHuron, order = c(2, 0, 0), include.drift = T))
Series: LakeHuron
ARIMA(2,0,0) with drift
```

Coefficients:

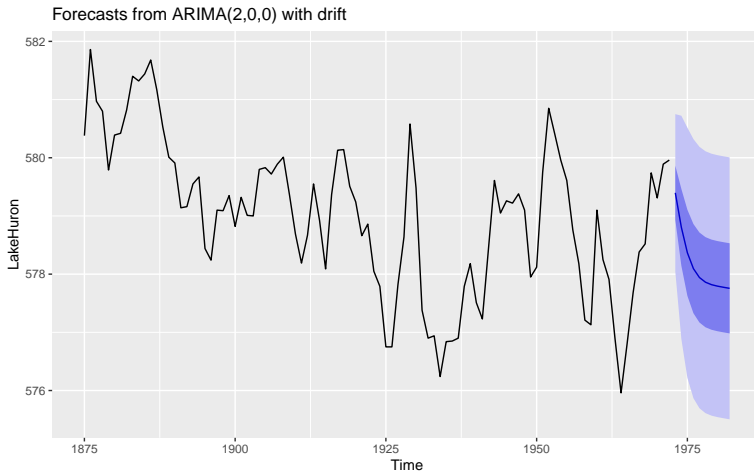
	ar1	ar2	intercept	drift
	1.0048	-0.2913	580.0915	-0.0216
s.e.	0.0976	0.1004	0.4636	0.0081

sigma<sup>2</sup> estimated as 0.476: log likelihood=-101.2

AIC=212.4 AICc=213.05 BIC=225.32



# 10-Year-Ahead Forecasts



This slides cover:

- Basic concepts of time series analysis
- A widely used class of models: **ARMA**
- ARMA model identification, estimation/prediction, inference