

Lecture 12

ARMA Case Study & Autoregressive Integrated Moving Average (ARIMA) Models

Reading: Cryer and Chen (2008): Chapter 5.1-5.3

MATH 4070: Regression and Time-Series Analysis

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Agenda

ARMA Case Study &
Autoregressive
Integrated Moving
Average (ARIMA)
Models



ARMA Case Study

ARMIA

1 ARMA Case Study

2 ARMIA

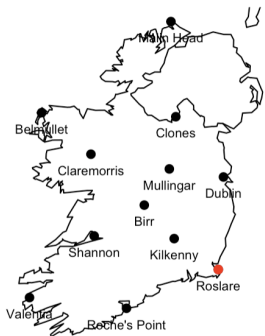
A Modeling Case Study of Ireland Wind Data

(Courtesy of Peter Craigmile's time series lecture notes)

Data Description [Haslett & Raftery, 1989 ¹]

Twelve wind stations collected daily readings over 18 years (from 1961 to 1978). Wind speeds were measured in knots (1 knot = $0.5148 \frac{m}{s}$)

We will focus on the wind data from 1965-1969 at the Rosslare station



Modeling procedure:

- Exploratory analysis
- Model and remove the trend and seasonal components
- ARMA model identification, fitting, and selection
- Perform forecast

¹ Haslett, J., & Raftery, A. E. (1989). Space-time modelling with long-memory dependence: Assessing Ireland's wind power resource. Journal of the Royal Statistical Society: Series C, 38(1), 1-21.

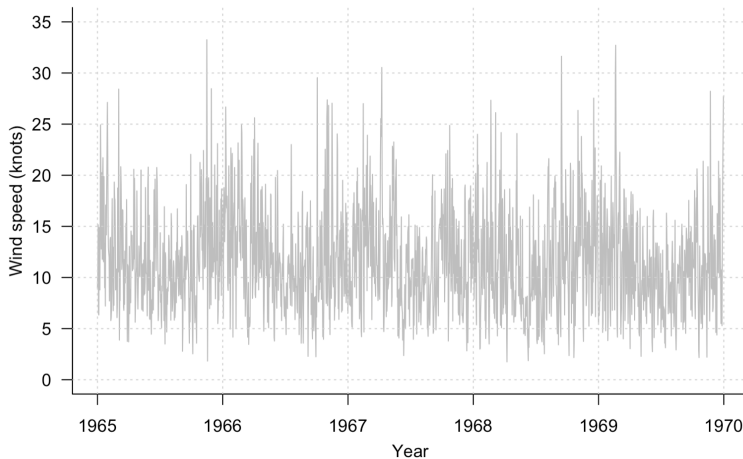
Wind Speed Time Series at Rosslare Station

ARMA Case Study &
Autoregressive
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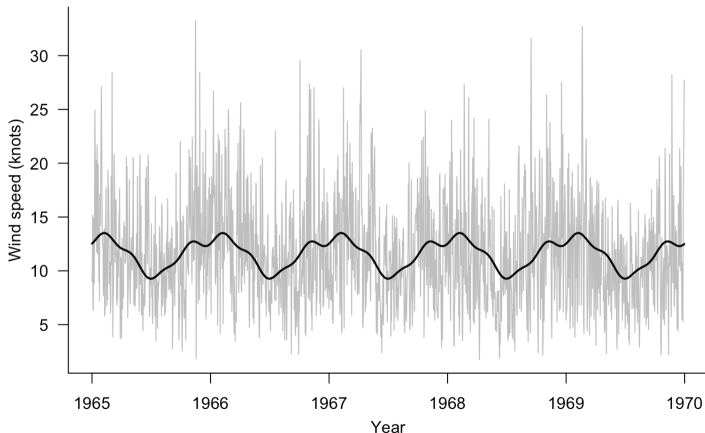
ARMA Case Study

ARIMA



- No clear trend
- Seasonal Pattern

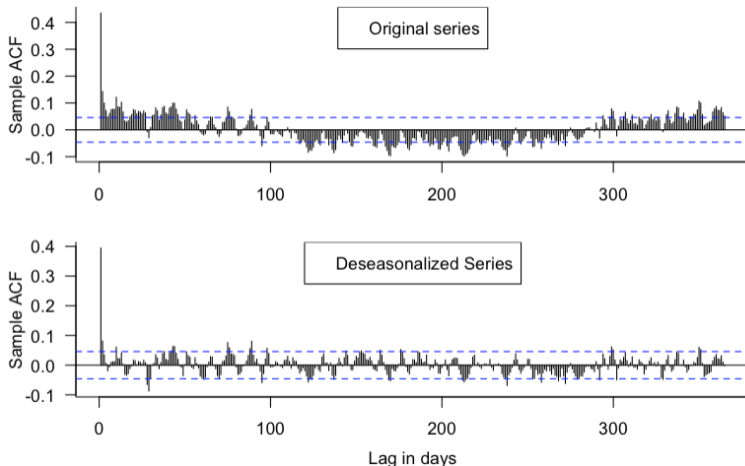
Estimating the Season Pattern



Here we use **harmonic regression** with 4 harmonics per year to model the seasonal components

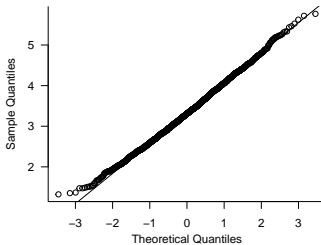
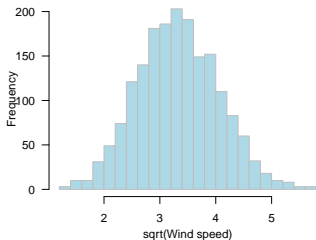
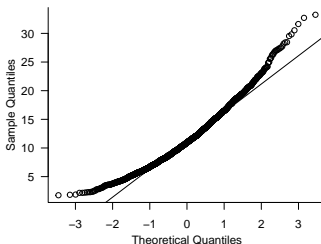
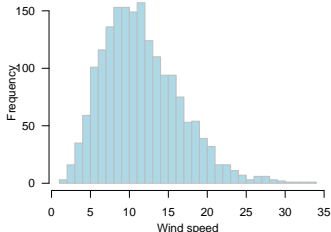
$$s_t = \beta_0 + \sum_{j=1}^4 (\beta_{1j} \cos(2\pi jt) + \beta_{2j} \sin(2\pi jt))$$

ACF Plots: Original and Deseasonalized Series



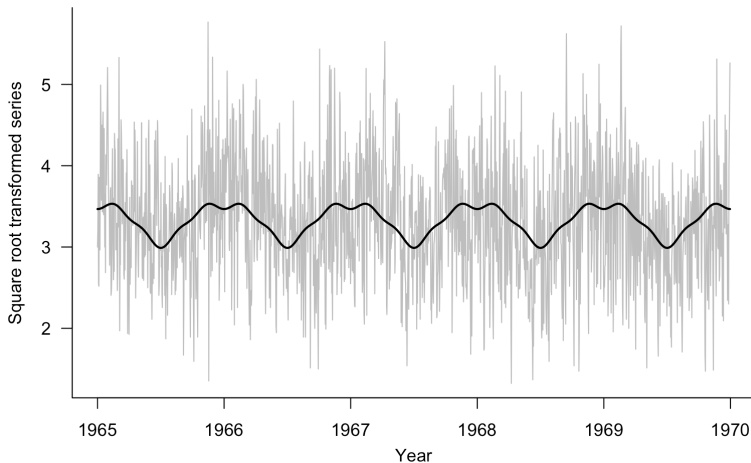
Seasonal modeling (via harmonic regression) effectively removes the oscillatory pattern in the ACF of the original series

Transform Data to Approximate Gaussian Distribution



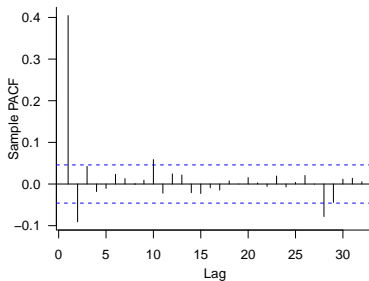
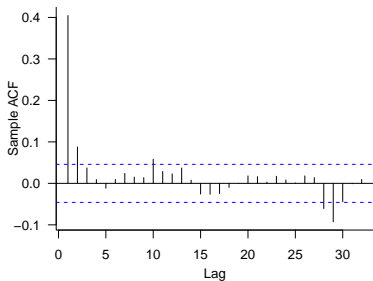
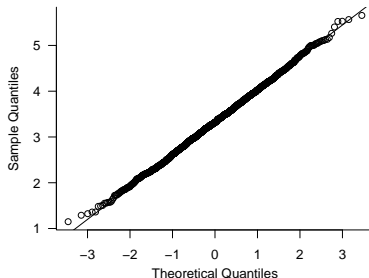
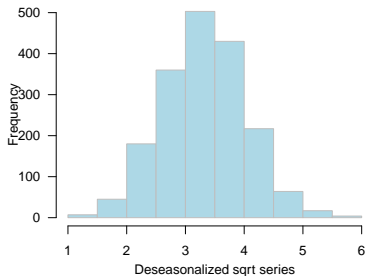
Square root transformation works! Now take the square root of the original data and deseasonalize again!

Estimating Transformed Series Seasonality



Next, we need to check if the deseasonalized series Gaussian like

Marginal and ACF/PACF of the Deseasonalized Series



Based on ACF/PACF, which ARMA model would you choose?

Potential Model 1: AR(1)

```
> ar1.model <- arima(sqrt.rosslare.ds, order = c(1, 0, 0))  
> ar1.model
```

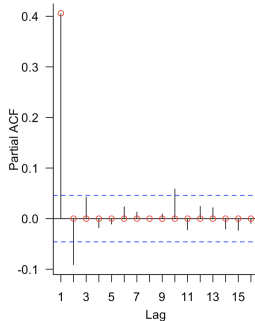
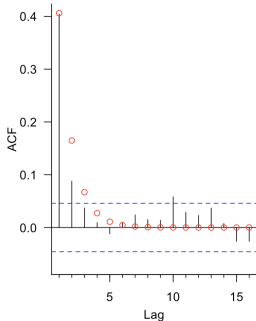
Call:

```
arima(x = sqrt.rosslare.ds, order = c(1, 0, 0))
```

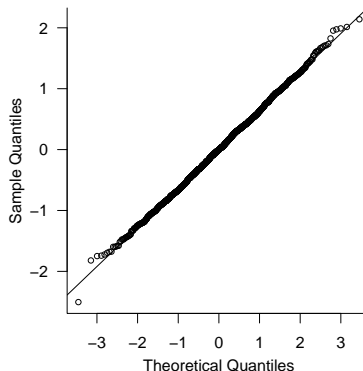
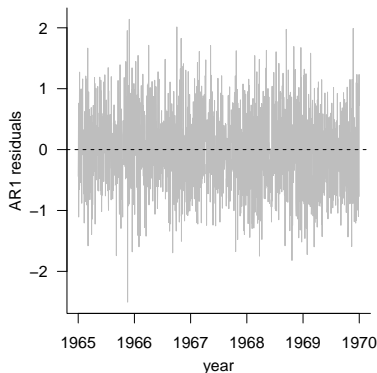
Coefficients:

	ar1	intercept
	0.4060	3.3257
s.e.	0.0214	0.0254

σ^2 estimated as 0.4148: log likelihood = -1787.72, aic = 3581.43



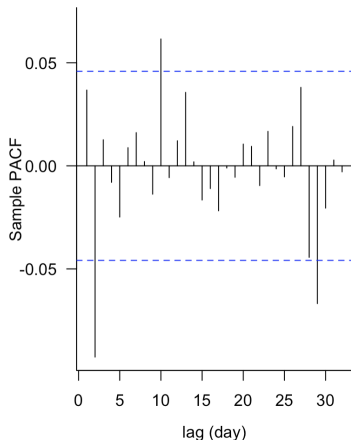
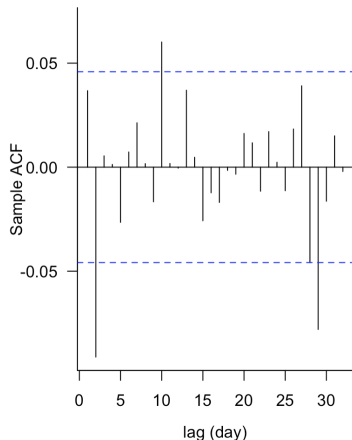
Residual Plots for the AR(1) Model



Normality assumption seems reasonable.

Next check the [ACF/PACF](#) and perform a [Box test](#) to assess if the AR(1) fit adequately account for temporal dependence structure

Diagnostic for the AR(1) Model



```
> Box.test(ar1.resids, lag = 32, fitdf = 1, type = "Ljung-Box")
```

Box-Ljung test

```
data: ar1.resids
```

```
X-squared = 53.142, df = 31, p-value = 0.00794
```

Potential Model 2: AR(2)

```
> (ar2.model <- arima(sqrt.rosslare.ds, order = c(2, 0, 0)))
```

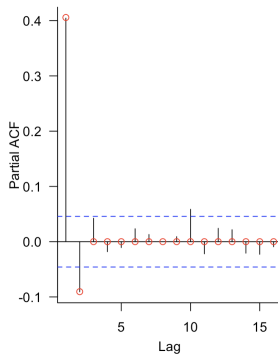
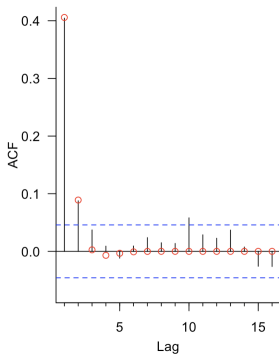
Call:

```
arima(x = sqrt.rosslare.ds, order = c(2, 0, 0))
```

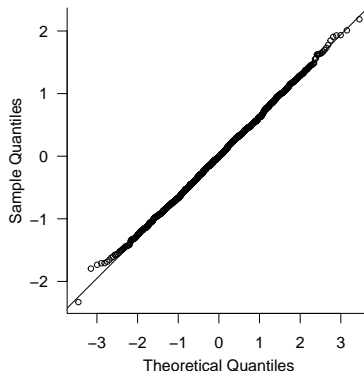
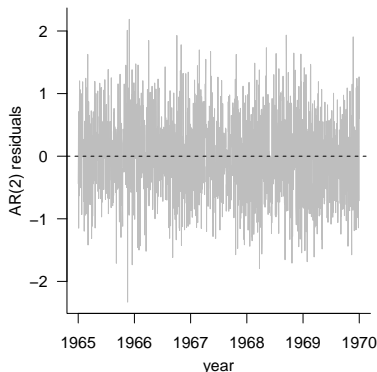
Coefficients:

	ar1	ar2	intercept
	0.4425	-0.0905	3.3254
s.e.	0.0233	0.0233	0.0232

σ^2 estimated as 0.4114: log likelihood = -1780.23, aic = 3568.46



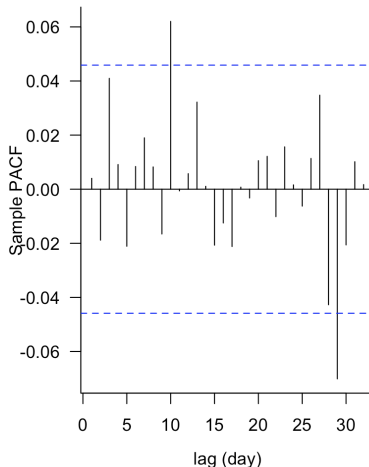
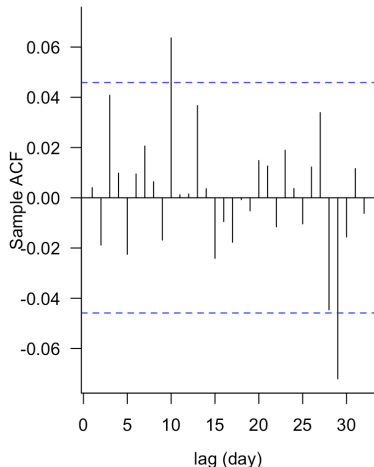
Residual Plots for the AR(2) Model



Normality assumption seems reasonable.

Next check the [ACF/PACF](#) and perform a [Box test](#) to assess if the AR(2) fit adequately account for temporal dependence structure

Diagnostic for the AR(2) Model



```
> Box.test(ar2.resids, lag = 32, fitdf = 2, type = "Ljung-Box")
```

Box-Ljung test

```
data: ar2.resids
```

```
X-squared = 36.548, df = 30, p-value = 0.1907
```


Potential Model 3: ARMA(1, 1)

```
> (arma11.model <- arima(sqrt.rosslare.ds, order = c(1, 0, 1)))
```

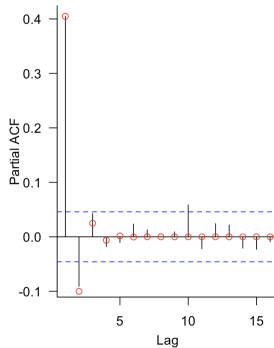
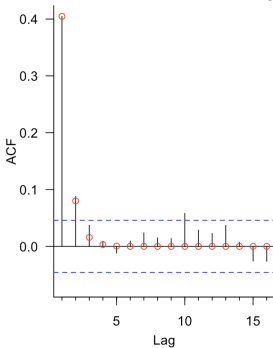
Call:

```
arima(x = sqrt.rosslare.ds, order = c(1, 0, 1))
```

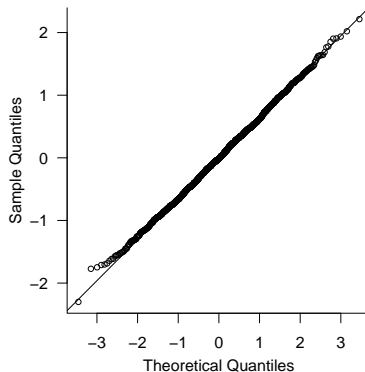
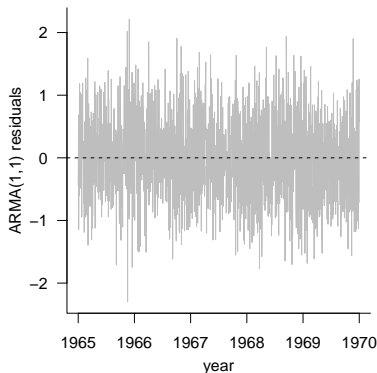
Coefficients:

	ar1	ma1	intercept
	0.1978	0.2502	3.3254
s.e.	0.0556	0.0553	0.0234

sigma² estimated as 0.4108: log likelihood = -1778.82, aic = 3565.64



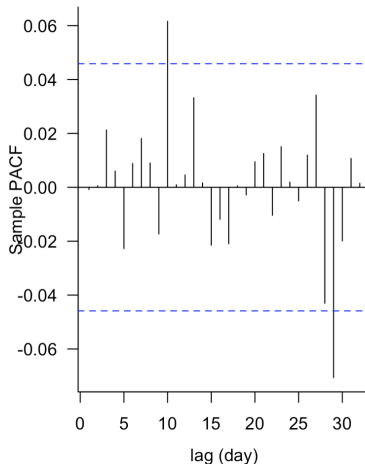
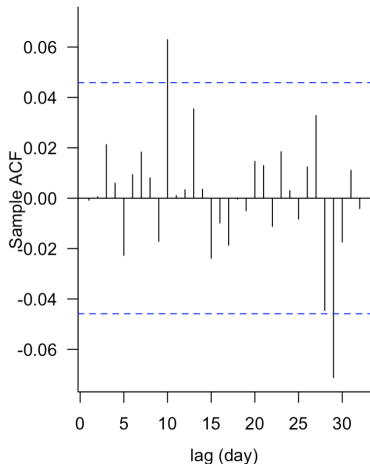
Residual Plots for the ARMA(1, 1) Model



Normality assumption seems reasonable.

Next check the [ACF/PACF](#) and perform a [Box test](#) to assess if the ARMA(1, 1) fit adequately account for temporal dependence structure

Diagnostic for the ARMA(1, 1) Model



```
> Box.test(arma11.resids, lag = 32, fitdf = 2, type = "Ljung-Box")
```

Box-Ljung test

```
data: arma11.resids  
X-squared = 32.757, df = 30, p-value = 0.3332
```

Potential Model 4: ARMA(2, 1)

```
> (arma21.model <- arima(sqrt.rosslare.ds, order = c(2, 0, 1)))
```

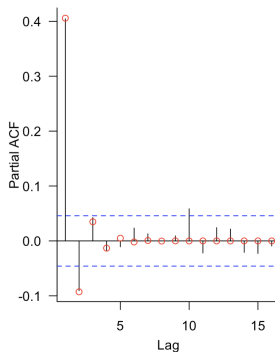
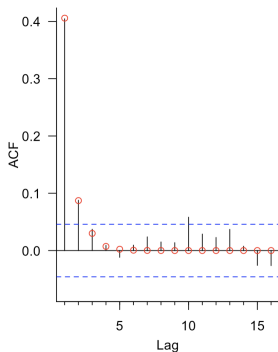
Call:

```
arima(x = sqrt.rosslare.ds, order = c(2, 0, 1))
```

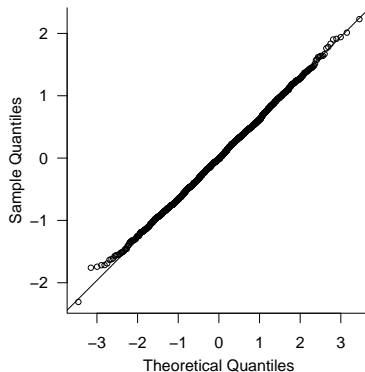
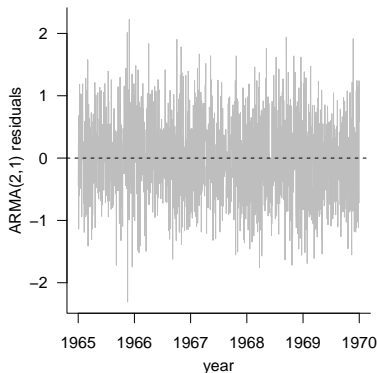
Coefficients:

	ar1	ar2	ma1	intercept
	0.0703	0.0587	0.3768	3.3253
s.e.	0.1691	0.0772	0.1663	0.0237

σ^2 estimated as 0.4107: log likelihood = -1778.56, aic = 3567.11



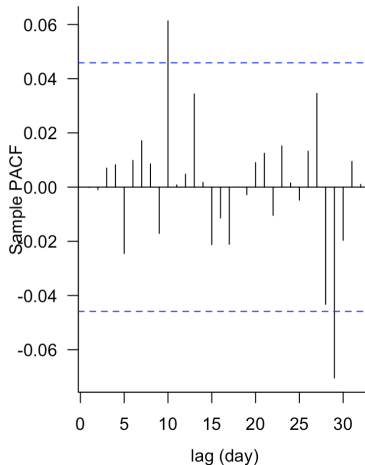
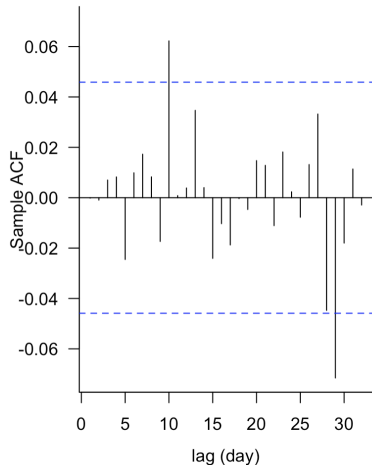
Residual Plots for the ARMA(2, 1) Model



Normality assumption seems reasonable.

Next check the [ACF/PACF](#) and perform a [Box test](#) to assess if the ARMA(2, 1) fit adequately account for temporal dependence structure

Diagnostic for the ARMA(2, 1) Model



```
> Box.test(arma21.resids, lag = 32, fitdf = 3, type = "Ljung-Box")
```

Box-Ljung test

data: arma21.resids

X-squared = 32.171, df = 29, p-value = 0.3124

Comparing Models via Information Criteria

Model	AIC	AICc
AR(1)	3583.817	3583.824
AR(2)	3570.650	3570.663
ARMA(1, 1)	3567.833	3567.847
ARMA(2, 1)	3569.319	3569.341

ARMA Case Study

ARMA

Which model would you pick?

Question: How do we predict wind speeds on the original scale, including the seasonality that was previously estimated?

- Suppose we want to predict the next 7 days of wind speed values. We base our forecasts on the chosen ARMA(1,1) model.
- We need to reverse the order of our modeling process: \Rightarrow forecast under the transformed scale \rightarrow add the estimated seasonal component \rightarrow back-transform to the original scale.

Forecasting Future Wind Speeds, continued

- The **forecasts** for the next 7 days of deseasonalized square root values are:

```
> round(sqrt.rosslare.forecast$pred, 3)
```

Time Series:

```
Start = c(1970, 1)
```

```
End = c(1970, 7)
```

```
Frequency = 365
```

```
[1] 3.997 3.458 3.352 3.331 3.326 3.326 3.325
```

- The **standard error** for the forecasts are:

```
> round(sqrt.rosslare.forecast$se, 3)
```

Time Series:

```
Start = c(1970, 1)
```

```
End = c(1970, 7)
```

```
Frequency = 365
```

```
[1] 0.641 0.702 0.705 0.705 0.705 0.705 0.705
```

Forecasting future wind speeds, continued

- Next, we add back in the seasonality to get:

```
> adj.forecast <- fitted(harm.model)[1:h] + sqrt.rosslare.forecast$pred  
> round(adj.forecast, 3)
```

Time Series:

Start = c(1970, 1)

End = c(1970, 7)

Frequency = 365

1	2	3	4	5	6	7
4.139	3.600	3.494	3.473	3.470	3.470	3.470

- Finally, we transform back to the original scale

```
> round(adj.forecast^2, 3)
```

Time Series:

Start = c(1970, 1)

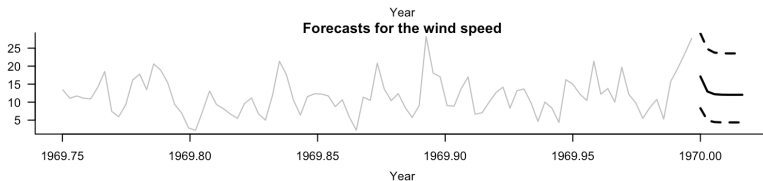
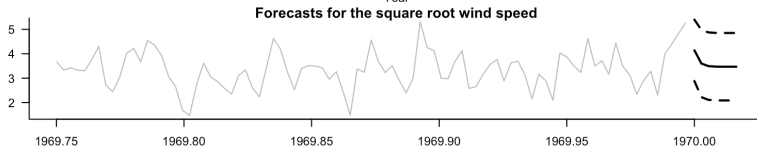
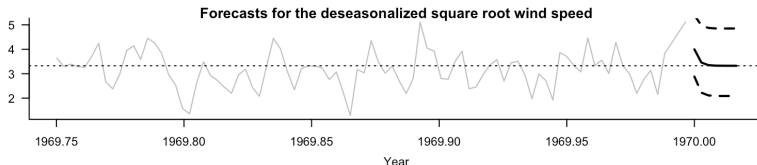
End = c(1970, 7)

Frequency = 365

1	2	3	4	5	6	7
17.132	12.962	12.208	12.064	12.039	12.039	12.044

- To get the prediction limits, we need to transform the lower and upper prediction limits on the sqrt scale

Visualizing the Forecasts

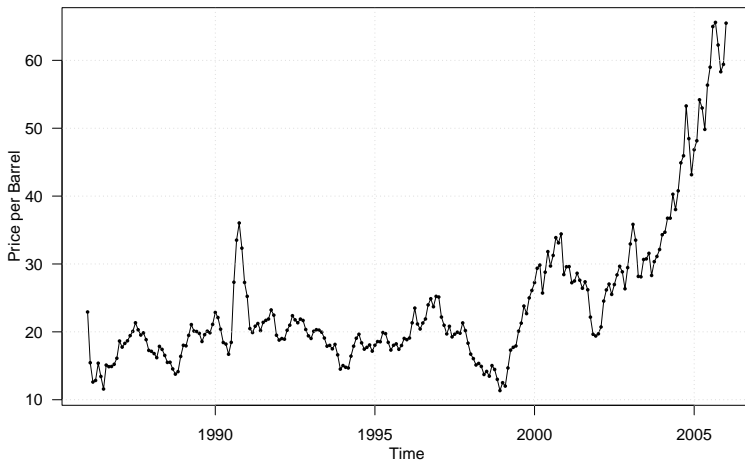


Further Questions

- What is the full model for our time series data?
- Is there a better description for the trend than just a constant term? What about alternative seasonal modeling?
- How well do we forecast? What about forecast uncertainty?

Autoregressive Integrated Moving Average (ARIMA) Models

Monthly Price of Oil: January 1986–January 2006



A stationary model does not seem to be reasonable. However, it is also not clear which (deterministic) trend model is appropriate 😞

Recall the random walk process

$$X_t = Z_1 + Z_2 + \cdots + Z_t = \sum_{j=1}^t Z_j,$$

where $\{Z_t\} \sim \text{WN}(0, \sigma^2)$

$\{X_t\}$ is a **nonstationary process**

- We can obtain a **stationary** process by **differencing**

$$\nabla X_t = X_t - X_{t-1} = (1 - B)X_t = Z_t$$

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- We can obtain a **stationary** process by **differencing**

$$\nabla X_t = X_t - X_{t-1} = (1 - B)X_t = Z_t$$

- $\{X_t\}$ is an example of an **autoregressive integrated moving average** (ARIMA) process— ARIMA(0, 1, 0) process

An ARIMA model is an ARMA process after differencing

- Let d be a non-negative integer. Then X_t is an ARIMA(p, d, q) process if

$$Y_t = \nabla^d X_t = (1 - B)^d X_t$$

is a **causal** ARMA process

An ARIMA model is an ARMA process after differencing

- Let d be a non-negative integer. Then X_t is an ARIMA(p, d, q) process if

$$Y_t = \nabla^d X_t = (1 - B)^d X_t$$

is a **causal** ARMA process

- Let $\phi(B)$ be the AR polynomial and $\theta(B)$ be the MA polynomial. Then for $\{Z_t\} \sim \text{WN}(0, \sigma^2)$

$$\phi(B)Y_t = \theta(B)Z_t,$$

and since $Y_t = (1 - B)^d X_t$, we have

$$\phi(B)(1 - B)^d X_t = \theta(B)Z_t$$

Example: ARIMA(1, 1, 0)

Let $\phi(z) = 1 - \phi_1 z$, $\theta(z) = 1$ and $d = 1$. For a **causal stationary solution** (after differencing) we need to assume $|\phi_1| < 1$. Then $\{X_t\}$ is an ARIMA (1, 1, 0) process,

$$(1 - \phi_1 B)(1 - B)X_t = Z_t,$$

where $\{Z_t\} \sim \text{WN}(0, \sigma^2)$

Now let $Y_t = (1 - B)X_t = X_t - X_{t-1}$, after some rearrangements we have

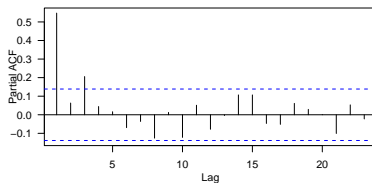
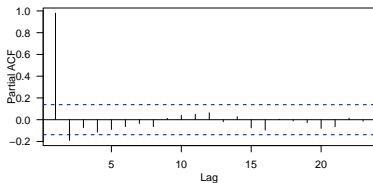
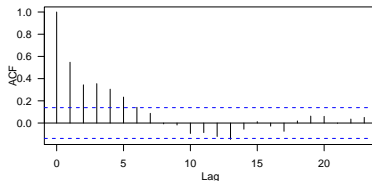
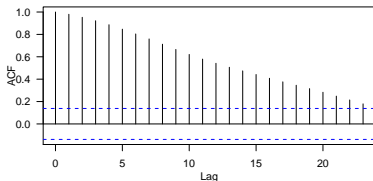
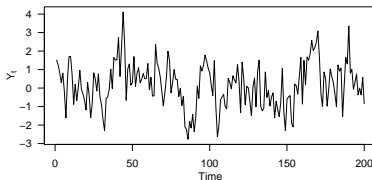
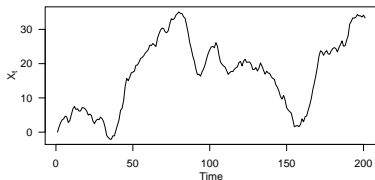
$$\begin{aligned} X_t &= X_{t-1} + Y_t \\ &= (X_{t-2} + Y_{t-1}) + Y_t \\ &\vdots \\ &= X_0 + \sum_{j=1}^t Y_j \end{aligned}$$

Thus $\{X_t\}$ is a “sort of random walk”—we **cumulatively sum** an AR(1) process, $\{Y_t\}$

Simulated ARIMA and Differenced ARMA Process

We simulate an ARIMA(1,1,0):

$$(1 - 0.5B)(1 - B)X_t = Z_t, \quad \{Z_t\} \sim N(0, 1)$$



Adding a Polynomial Trend

For $d \geq 1$, let $\{X_t\}$ be an $\text{ARIMA}(p, d, q)$ process. Then $\{X_t\}$ satisfies the equation

$$\phi(B)(1 - B)^d X_t = \theta(B)Z_t$$

- Let μ_t be a polynomial of degree $(d - 1)$, i.e., $\mu_t = \sum_{j=0}^{d-1} a_j t^j$ for constants $\{a_j\}$

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$$\phi(B)(1-B)^d X_t = \theta(B)Z_t$$

- Let μ_t be a polynomial of degree $(d-1)$, i.e., $\mu_t = \sum_{j=0}^{d-1} a_j t^j$ for constants $\{a_j\}$
- Now let $V_t = \mu_t + X_t$, then

$$\begin{aligned}\phi(B)(1-B)^d V_t &= \phi(B)(1-B)^d (\mu_t + X_t) \\ &= \phi(B)(1-B)^d \mu_t + \phi(B)(1-B)^d X_t \\ &= 0 + \phi(B)(1-B)^d X_t \\ &= \theta(B)Z_t\end{aligned}$$

Adding a Polynomial Trend

For $d \geq 1$, let $\{X_t\}$ be an $\text{ARIMA}(p, d, q)$ process. Then $\{X_t\}$ satisfies the equation

$$\phi(B)(1 - B)^d X_t = \theta(B)Z_t$$

- Let μ_t be a polynomial of degree $(d - 1)$, i.e., $\mu_t = \sum_{j=0}^{d-1} a_j t^j$ for constants $\{a_j\}$
- Now let $V_t = \mu_t + X_t$, then

$$\begin{aligned}\phi(B)(1 - B)^d V_t &= \phi(B)(1 - B)^d (\mu_t + X_t) \\ &= \phi(B)(1 - B)^d \mu_t + \phi(B)(1 - B)^d X_t \\ &= 0 + \phi(B)(1 - B)^d X_t \\ &= \theta(B)Z_t\end{aligned}$$

- Takeaway:** $\text{ARIMA}(p, d, q)$ are useful for modeling data with **polynomial trends**, due to the inherent differencing that can be used to remove trends

Steps for Modeling ARIMA Processes: Exploratory Analysis

- Plot the data, ACF, PACF and Q-Q plots
 - Check for unusual features of the data
 - Check for stationarity
 - Do we need to transform the data?
- Eliminate trend
 - Estimating the trend and removing it from the series
 - Or, differencing the series (i.e., select d in the ARIMA model)
- Plot the sample ACF/PACF for the stationary component
 - Identify candidate values of p and q

Steps for Modeling ARIMA Processes: Estimation and Model Checking

- Estimate the ARMA process parameters for the candidate models
- Check the goodness of fit: Are the time series residuals, $\{r_t\}$ a sample of *i.i.d.* noise?
- Model selection:
 - Using information criteria such as AIC and AICC
 - Test model parameters to compare between the “full” model and the “subset” model