Lecture 13

State-Space Models I

Readings: SS17 Chapter 6.1-6.2; BD Chapter 9.1-9.3

MATH 8090 Time Series Analysis
Week 13

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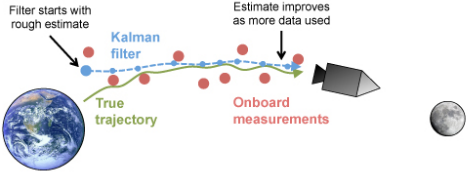
Agenda

Historical Background

- The original model emerged in the context of space tracking [Kalman, 1960, Kalman and Bucy, 1961]

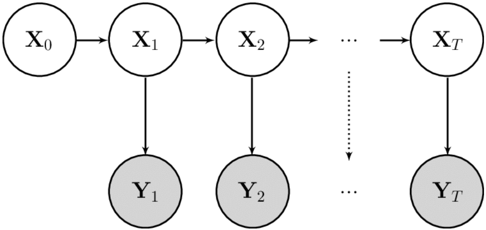
- The "state equation" defines the motion equations for the position of a spacecraft with location $\boldsymbol{x}_t$



- The data $\boldsymbol{y}_t$ reflect information that can be observed from a tracking device, such as velocity and azimuth

> The main goal was to retrieve the underling state $\{x_t\}$ based on observed data $\{y_t\}$

State-Space Model



$$\text{State: } \boldsymbol{X}_t = M_t \boldsymbol{X}_{t-1} + \boldsymbol{V}_t, \quad \boldsymbol{V}_t \overset{i.i.d.}{\sim} \mathrm{WN}(\mathbf{0}, Q_t), \quad t = 1, 2, \cdots$$

$$\text{Observation: } \boldsymbol{Y}_t = H_t \boldsymbol{X}_t + \boldsymbol{W}_t, \quad \boldsymbol{W}_t \overset{i.i.d.}{\sim} \mathrm{WN}(\mathbf{0}, R_t), \quad t = 1, 2, \cdots$$

- $\boldsymbol{X_t} \in \mathbb{R}^p$ and $\boldsymbol{Y_t} \in \mathbb{R}^q$ are the state vector and the observation vector at time t
- $\boldsymbol{M}_t$ is the $p \times p$ transition matrix, and $\boldsymbol{H}_t$ is the $q \times p$ observation matrix
- $\boldsymbol{V}_t$ and $\boldsymbol{W}_t$ are the state and observation noises

Additional Assumptions of State-Space Models

State equation:

$$\boldsymbol{X}_t = M_t \boldsymbol{X}_{t-1} + \boldsymbol{V}_t, \quad t = 1, 2 \cdots$$

Observation equation:

$$\boldsymbol{Y}_t = H_t \boldsymbol{X}_t + \boldsymbol{W}_t, \quad t = 1, 2, \cdots$$

- $\mathrm{E}(\boldsymbol{W}_s \boldsymbol{V}_t^T) = 0$ for all s and t, that is, every observation noise is uncorrelated with every state-transition noise

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- $\mathrm{E}(\boldsymbol{W}_s \boldsymbol{V}_t^T) = 0$ for all s and t, that is, every observation noise is uncorrelated with every state-transition noise
- Assuming $\mathrm{E}(\boldsymbol{X}_0) = \boldsymbol{\mu}_0$, $\mathrm{E}(\boldsymbol{X}_0 \boldsymbol{W}_t^T) = 0$ and $\mathrm{E}(\boldsymbol{X}_0 \boldsymbol{V}_t^T) = 0$ for all t, that is, initial state vector are uncorrelated with both observation and state transition noises

Applications of State-Space Models

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- The Kalman recursions for state-space models provide elegant solution for forecasting, filtering, and smoothing

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  - When $s < t \Rightarrow$ forecasting
  - When $s = t \Rightarrow$ filtering
  - When $s > t \Rightarrow$smoothing
- State-space models and Kalman recursions can be readily adapted to handle time series with missing values

AR(1) Process as a State-Space Model: I

- State-transition equation

$$\boldsymbol{X}_t = M_t \boldsymbol{X}_{t-1} + \boldsymbol{V}_t$$

is reminiscent of a causal AR(1) model:

$$Y_t = \phi Y_{t-1} + Z_t,$$

with $\{Z_t\} \sim \mathrm{WN}(0, \sigma^2)$ and $|\phi| < 1$

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  and by using a degenerate form of the observation equation: $\boldsymbol{Y}_t = H_t \boldsymbol{X}_t + \boldsymbol{W}_t$ in which $H_t = 1$ and $\boldsymbol{W}_t = 0$ so that $\boldsymbol{Y}_t = X_t$

AR(1) Process as a State-Space Model: II

Need to define the initial state X_0 in order to complete the model:

- A natural choice is

$$X_0 = \sum_{j=1}^{\infty} \phi^j Z_{1-j}, \quad \text{for which } \mathrm{Var}(X_0) = \frac{\sigma^2}{1-\phi^2}$$

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- Could also set $X_0 = Z_0 \frac{\sigma}{\sqrt{1-\phi^2}}$ to get a AR(1) process, but using $X_0 = Z_0$ would lead to a valid state-space model that is **not** a true AR(1) model

AR(1) Process as a State-Space Model: III

AR(1) process with $0 < \phi < 1$ is known as “red noise”, red noise is related to a 1st order stochastic differential equation, rendering it a model for various geophysical processes:

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- Can modify this setup by changing observational noise from $\boldsymbol{W}_t = 0$ to $\boldsymbol{W}_t = W_t \sim \mathrm{WN}(0, \sigma_W^2)$, where W_t is uncorrelated with Z_t's
- The observation and state-transition equations become

$$Y_t = X_t + W_t \text{ and } X_t = \phi X_{t-1} + Z_t$$

ARMA(1,1) Process as a State-Space Model: I

Recall ARMA(1,1) process $Y_t - \phi Y_{t-1} = Z_t + \theta Z_{t-1}$

- Expressing ARMA(1,1) as $\phi(B)Y_t = \theta(B)Z_t$, note that one can create Y_t by taking causal AR(1) process $X_t = \phi^{-1}(B)Z_t$ and subjecting it to a $\theta(B)$ filter to obtain output $Y_t = \theta(B)X_t = \theta(B)\phi^{-1}(B)Z_t$

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- Can express filtering of AR(1) process by

$$Y_t = \begin{bmatrix} 1 & \theta \end{bmatrix} \begin{bmatrix} X_t \\ X_{t-1} \end{bmatrix},$$

which matches up with observation equation

$$\boldsymbol{Y}_t = H_t \boldsymbol{X}_t + \boldsymbol{W}_t$$

if $\boldsymbol{Y}_t = Y_t$, $H_t = \begin{bmatrix} 1 & \theta \end{bmatrix}$, $\boldsymbol{X}_t = \begin{bmatrix} X_t \\ X_{t-1} \end{bmatrix}$ and $\boldsymbol{W}_t = 0$

ARMA(1,1) Process as a State-Space Model: II

- Given $\boldsymbol{X}_t = \begin{bmatrix} X_t & X_{t-1} \end{bmatrix}^T$, can express $X_t = \phi X_{t-1} + Z_t$ in the 1st row of matrix equation

$$\begin{bmatrix} X_t \\ X_{t-1} \end{bmatrix} = \begin{bmatrix} \phi & 0 \\ 1 & 0 \end{bmatrix} \begin{bmatrix} X_{t-1} \\ X_{t-2} \end{bmatrix} + \begin{bmatrix} Z_t \\ 0 \end{bmatrix},$$

which matches up with state-transition equation

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if $M_t = \begin{bmatrix} \phi & 0 \\ 1 & 0 \end{bmatrix}$ and $\boldsymbol{V}_t = \begin{bmatrix} Z_t \\ 0 \end{bmatrix}$ with

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- to complete the model, let

$$\boldsymbol{X}_0 = \begin{bmatrix} X_0 \\ X_{-1} \end{bmatrix} = \begin{bmatrix} \sum_{j=1}^{\infty} \phi^j Z_{1-j} \\ \sum_{j=1}^{\infty} \phi^j Z_{-j} \end{bmatrix},$$

  noting that $\boldsymbol{X}_0$ and $\boldsymbol{V}_t$ for $t \geq 1$ are uncorrelated, as required

ARMA(1,1) Process as a State-Space Model: III

Since

$$\mathrm{E}(\boldsymbol{X}_0\boldsymbol{X}_0^T) = \begin{bmatrix}\gamma(0) & \gamma(1)\\ \gamma(1) & \gamma(0)\end{bmatrix} = \frac{\sigma^2}{1-\phi^2}\begin{bmatrix}1 & \phi\\ \phi & 1\end{bmatrix},$$

can alternatively stipulate

$$\boldsymbol{X}_0 = \begin{bmatrix}1 & \frac{\phi}{\sqrt{1-\phi^2}}\\ 0 & \frac{\phi}{\sqrt{1-\phi^2}}\end{bmatrix}\begin{bmatrix}Z_0\\ Z_{-1}\end{bmatrix},$$

yielding

$$\begin{aligned}\mathrm{E}(\boldsymbol{X}_0\boldsymbol{X}_0^T) &= \begin{bmatrix}1 & \frac{\phi}{\sqrt{1-\phi^2}}\\ 0 & \frac{\phi}{\sqrt{1-\phi^2}}\end{bmatrix}\begin{bmatrix}\sigma^2 & 0\\ 0 & \sigma^2\end{bmatrix}\begin{bmatrix}1 & 0\\ \frac{\phi}{\sqrt{1-\phi^2}} & \frac{1}{\sqrt{1-\phi^2}}\end{bmatrix}\\ &= \frac{\sigma^2}{1-\phi^2}\begin{bmatrix}1 & \phi\\ \phi & 1\end{bmatrix}\end{aligned}$$

as required

The Linear Gaussian State-Space Model

- State equation:

$$\boldsymbol{X}_t = M_t \boldsymbol{X}_{t-1} + \boldsymbol{V}_t,$$

where $\boldsymbol{V}_t \overset{iid}{\sim} \mathrm{N}(\mathbf{0}, Q_t)$ with $\boldsymbol{X}_0 \sim \mathrm{N}(\boldsymbol{\mu}_0, \Sigma_0)$

- Observation equation:

$$\boldsymbol{Y}_t = H_t \boldsymbol{X}_t + \boldsymbol{W}_t,$$

where $\boldsymbol{W}_t \overset{iid}{\sim} \mathrm{N}(\mathbf{0}, R_t)$

- Additional assumptions: $\boldsymbol{X}_0$, $\{\boldsymbol{V}_t\}$, and $\{\boldsymbol{W}_t\}$ are uncorrelated

Forecasting, Filtering, and Smoothing

Goal: To estimate the underlying unobserved signal $\boldsymbol{X}_t$, given the data $\boldsymbol{y}_{1:s} = \{\boldsymbol{y}_1, \boldsymbol{y}_2, \cdots, \boldsymbol{y}_s\}$:

- When $s < t$, the problem is called forecasting or prediction
- When $s = t$, the problem is called filtering
- When $s > t$, the problem is called smoothing

In addition to these estimates, we would also want to measure their precision. The solution to these problems is accomplished via the Kalman filter and Kalman smoother

The Kalman Filter: General Results

Assume the filtering distribution at time $t-1$ is

$$[\boldsymbol{X}_{t-1}|\boldsymbol{y}_{1:t-1}] \sim \mathrm{N}\left(\boldsymbol{\mu}_{t-1}^{a}, \Sigma_{t-1}^{a}\right)$$

- **Forecast Step:** Gives the forecast distribution at time t:

$$[\boldsymbol{X}_{t}|\boldsymbol{y}_{1:t-1}] \sim \mathrm{N}\left(\boldsymbol{\mu}_{t}^{f}, \Sigma_{t}^{f}\right),$$

where $\boldsymbol{\mu}_{t}^{f} = M_t \boldsymbol{\mu}_{t-1}^{a}$, and $\Sigma_{t}^{f} = M_t \Sigma_{t-1}^{a} M_t^T + Q_t$.

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  where $\boldsymbol{\mu}_t^f = M_t\boldsymbol{\mu}_{t-1}^a$, and $\Sigma_t^f = M_t\Sigma_{t-1}^a M_t^T + Q_t$.
- **Update Step:** updates the forecast distribution using new data $\boldsymbol{y}_t$

  $$[\boldsymbol{X}_t|\boldsymbol{y}_{1:t}] \sim \mathrm{N}\left(\boldsymbol{\mu}_t^a, \Sigma_t^a\right),$$

  where $\boldsymbol{\mu}_t^a = \boldsymbol{\mu}_t^f + K_t\left(\boldsymbol{y}_t - H_t\boldsymbol{\mu}_t^f\right)$, and $\Sigma_t^a = \left(I - K_tH_t^T\right)\Sigma_t^f$, and

  $$K_t = \Sigma_t^f H_t^T\left(H_t\Sigma_t^f H_t^T + R_t\right)^{-1}$$

  is the Kalman gain matrix

Local Level Model: Part I

Let's begin with a particularly simple example of a state space model: the local level model. We will develop the basic state space techniques for this model.

- **Observation equation**:

$$Y_t = X_t + W_t, \quad \{W_t\} \overset{iid}{\sim} \mathrm{N}(0, \sigma_W^2)$$

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- Assume $\mathrm{E}(X_0) = \mu_0$ and $\mathrm{Var}(X_0) = \sigma_0^2$ and X_0 is uncorrected with W_t's and V_t's

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  i.e., if state variable is at a certain 'level' at time t, we can expect no change in its level at time $t + 1$

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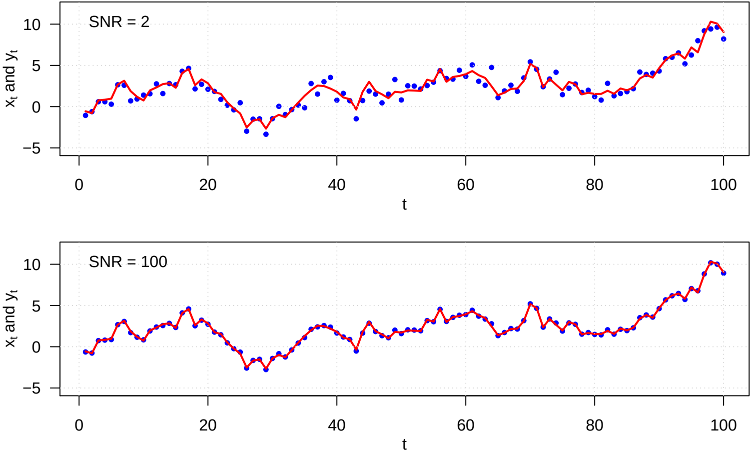
- When $\sigma_W^2 > 0$, trend is corrupted by noise, so ability to pick out trend depends upon "signal to noise" ratio (SNR) $\frac{\sigma_V^2}{\sigma_W^2}$

Local Level Model: Examples of Different SNR



Four Problems in State-Space Models

Given observations $\{Y_i\}_{i=1}^{t}$ of a local level process,

1. Filtering: what is best predictor of state X_t?
2. Forecasting: what is best predictor of state X_{t+1}?
3. Smoothing: what is best predictor of state X_s for $s < t$?
4. Estimation: what are best estimates of model parameters $\sigma_W^2, \sigma_V^2, \mu_0, \sigma_0^2$?

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To facilitate discussion, let’s assume that X_0, V_t’s, and W_t are normals, implying that Y_t and the remaining X_t’s share this property.

Regression Lemma I

- Suppose random vectors $\boldsymbol{X}$ and $\boldsymbol{Y}$ are jointly normal with mean vector $\boldsymbol{\mu}$ and covariance matrix Σ, to be denoted by

$$\begin{bmatrix} \boldsymbol{X} \\ \boldsymbol{Y} \end{bmatrix} \sim \mathrm{N}(\boldsymbol{\mu}, \Sigma)$$

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- Can partition both $\boldsymbol{\mu}$ and Σ:

$$\begin{bmatrix} \boldsymbol{X} \\ \boldsymbol{Y} \end{bmatrix} \sim \mathrm{N}\left(\begin{bmatrix} \boldsymbol{\mu}_{\boldsymbol{X}} \\ \boldsymbol{\mu}_{\boldsymbol{Y}} \end{bmatrix}, \begin{bmatrix} \Sigma_{\boldsymbol{XX}} & \Sigma_{\boldsymbol{XY}} \\ \Sigma_{\boldsymbol{YX}} & \Sigma_{\boldsymbol{YY}} \end{bmatrix} \right),$$

where $\boldsymbol{\mu}_{\boldsymbol{X}}$ ($\boldsymbol{\mu}_{\boldsymbol{Y}}$) and $\Sigma_{\boldsymbol{XX}}$ ($\Sigma_{\boldsymbol{YY}}$) are mean and covariance matrix for $\boldsymbol{X}$ ($\boldsymbol{Y}$); $\Sigma_{\boldsymbol{XY}}$ is the cross-covariance matrix between $\boldsymbol{X}$ and $\boldsymbol{Y}$

Regression Lemma II

- Conditional distribution of $\boldsymbol{X}$ given $\boldsymbol{Y} = \boldsymbol{y}$ is multivariate normal with mean vector

$$\boldsymbol{\mu}_{\boldsymbol{X}|\boldsymbol{y}} = \boldsymbol{\mu}_{\boldsymbol{X}} + \Sigma_{\boldsymbol{X}\boldsymbol{Y}} \Sigma_{\boldsymbol{Y}\boldsymbol{Y}}^{-1} (\boldsymbol{y} - \boldsymbol{\mu}_{\boldsymbol{Y}})$$

and covariance matrix

$$\Sigma_{\boldsymbol{X}|\boldsymbol{y}} = \Sigma_{\boldsymbol{X}\boldsymbol{X}} - \Sigma_{\boldsymbol{X}\boldsymbol{Y}} \Sigma_{\boldsymbol{Y}\boldsymbol{Y}}^{-1} \Sigma_{\boldsymbol{X}\boldsymbol{Y}}^{T}$$

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- Best (under MSE) predictor of $\boldsymbol{X}$ given $\boldsymbol{Y}$ is

$$\mathrm{E}(\boldsymbol{X}|\boldsymbol{Y}) = \boldsymbol{\mu}_{\boldsymbol{X}|\boldsymbol{Y}} = \boldsymbol{\mu}_{\boldsymbol{X}} + \Sigma_{\boldsymbol{XY}}\Sigma_{\boldsymbol{YY}}^{-1}(\boldsymbol{Y} - \boldsymbol{\mu}_{\boldsymbol{Y}})$$

Regression Lemma III

- Recall that, if random vector $\boldsymbol{U}$ has covariance matrix $\Sigma_{\boldsymbol{U}}$, then covariance matrix for $A\boldsymbol{U}$ is $A\Sigma_{\boldsymbol{U}}A^T$

  $\Rightarrow$ covariance matrix of $c + A(\boldsymbol{U} - \mu_{\boldsymbol{U}})$ is also $A\Sigma_{\boldsymbol{U}}A^T$

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- Covariance matrix for

$$\mathrm{E}(\boldsymbol{X}|\boldsymbol{Y}) = \boldsymbol{\mu}_{\boldsymbol{X}|\boldsymbol{Y}} = \boldsymbol{\mu}_{\boldsymbol{X}} + \Sigma_{\boldsymbol{XY}}\Sigma_{\boldsymbol{YY}}^{-1}(\boldsymbol{Y} - \boldsymbol{\mu}_{\boldsymbol{Y}})$$

  is thus

$$\Sigma_{\boldsymbol{XY}}\Sigma_{\boldsymbol{YY}}^{-1}\Sigma_{\boldsymbol{YY}}\Sigma_{\boldsymbol{YY}}^{-1}\Sigma_{\boldsymbol{XY}}^T = \Sigma_{\boldsymbol{XY}}\Sigma_{\boldsymbol{YY}}^{-1}\Sigma_{\boldsymbol{XY}}^T$$

Note: it is not the same as $\Sigma_{\boldsymbol{X}|\boldsymbol{y}} = \Sigma_{\boldsymbol{XX}} - \Sigma_{\boldsymbol{XY}}\Sigma_{\boldsymbol{YY}}^{-1}\Sigma_{\boldsymbol{XY}}^T$

Regression Lemma IV

Consider prediction error $\boldsymbol{U}$ associated with best linear predictor of $\boldsymbol{X}$:

$$\boldsymbol{U} = \boldsymbol{X} - \mathrm{E}(\boldsymbol{X}|\boldsymbol{Y})$$

- Since $\mathrm{E}\left[\mathrm{E}\left(\boldsymbol{X}|\boldsymbol{Y}\right)\right] = \boldsymbol{\mu}_{\boldsymbol{X}} \Rightarrow \mathrm{E}(\boldsymbol{U}) = \mathbf{0}$

Regression Lemma IV

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- Since $\mathrm{E}\left[\mathrm{E}\left(\boldsymbol{X}|\boldsymbol{Y}\right)\right] = \boldsymbol{\mu}_{\boldsymbol{X}} \Rightarrow \mathrm{E}(\boldsymbol{U}) = \mathbf{0}$
- Covariance matrix for $\boldsymbol{U}$ is given by

$$\begin{aligned}\mathrm{E}(\boldsymbol{U}\boldsymbol{U}^T) =& \mathrm{E}\left([\boldsymbol{X} - \mathrm{E}(\boldsymbol{X}|\boldsymbol{Y})]\,[\boldsymbol{X} - \mathrm{E}(\boldsymbol{X}|\boldsymbol{Y})]^T\right)\\ =& \mathrm{E}(\boldsymbol{X}\boldsymbol{X}^T) + \mathrm{E}[\mathrm{E}(\boldsymbol{X}|\boldsymbol{Y})\mathrm{E}(\boldsymbol{X}|\boldsymbol{Y})^T]\\ &- \mathrm{E}[\boldsymbol{X}\mathrm{E}(\boldsymbol{X}|\boldsymbol{Y})^T] - \mathrm{E}[\mathrm{E}(\boldsymbol{X}|\boldsymbol{Y})\boldsymbol{X}^T]\\ =& \Sigma_{\boldsymbol{X}\boldsymbol{X}} - \Sigma_{\boldsymbol{X}\boldsymbol{Y}}\Sigma_{\boldsymbol{Y}\boldsymbol{Y}}^{-1}\Sigma_{\boldsymbol{X}\boldsymbol{Y}}^T,\end{aligned}$$

which is equal to $\Sigma_{\boldsymbol{X}|\boldsymbol{y}}$, the conditional covariance matrix

Regression Corollary

Now specialize to the case where $\boldsymbol{X}$ has just one element, say, X

- Corollary: conditional distribution of X given $\boldsymbol{Y} = \boldsymbol{y}$ is normal with mean

$$\mu_X + \Sigma_{X\boldsymbol{Y}}^T \Sigma_{\boldsymbol{Y}\boldsymbol{Y}}^{-1} (\boldsymbol{y} - \boldsymbol{\mu}_{\boldsymbol{Y}})$$

and conditional variance

$$\Sigma_{X|\boldsymbol{y}} = \sigma_X^2 - \Sigma_{X\boldsymbol{Y}}^T \Sigma_{\boldsymbol{Y}\boldsymbol{Y}}^{-1} \Sigma_{X\boldsymbol{Y}},$$

where $\sigma_X^2 = \mathrm{Var}(X)$ and $\Sigma_{X\boldsymbol{Y}}$ is a column vector containing covariance between X and $\boldsymbol{Y}$

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- Since conditional variance is same as MSE for X, will refer to $\Sigma_{X|\boldsymbol{y}}$ as MSE

Aside – Revisiting Time Series Prediction: I

Suppose $\{X_t\}$ is zero mean stationary process with ACF $\gamma(h)$

- Set X to X_{n+1} and put $X_1, \cdots, X_n$ into $\boldsymbol{Y}$
- Corollary says best linear predictor $\hat{X}_{n+1}$ of X_{n+1} given $X_1, \cdots, X_n$ is

$$\hat{X}_{n+1} = \Sigma_{X\boldsymbol{Y}}^T \Sigma_{\boldsymbol{Y}\boldsymbol{Y}}^{-1} \boldsymbol{Y} = \gamma_n^T \Gamma_n^{-1} \boldsymbol{Y} \stackrel{\text{def}}{=} \boldsymbol{c}_n^T \boldsymbol{Y},$$

where

1. $\gamma_n = [\gamma(1), \gamma(2), \cdots, \gamma(n)]^T = \Sigma_{X\boldsymbol{Y}}$
2. (i, j)th entry of matrix $\Gamma_n = \Sigma_{\boldsymbol{Y}\boldsymbol{Y}}$ is $\gamma(i - j)$
3. $\boldsymbol{c}_n^T \stackrel{\text{def}}{=} \gamma_n^T \Gamma_n^{-1}$ and hence $\boldsymbol{c}_n = \Gamma_n^{-1} \gamma_n$

Aside – Revisiting Time Series Prediction: II

Recall that MSE for $\hat{X}_{n+1}$ is

$$\begin{aligned} v_n &= \mathrm{Var}(X_{n+1}) - \boldsymbol{c}_n^T \boldsymbol{\gamma}_n \\ &= \sigma_X^2 - \boldsymbol{\gamma}_n^T \Gamma_n^{-1} \boldsymbol{\gamma}_n \\ &= \sigma_X^2 - \Sigma_{X\boldsymbol{Y}}^T \Sigma_{\boldsymbol{Y}\boldsymbol{Y}}^{-1} \Sigma_{X\boldsymbol{Y}} \\ &= \Sigma_{X|\boldsymbol{y}} \end{aligned}$$

This is a special case of regression corollary

Filtering for Local Level Model: I

Let's begin with a particularly simple example of a state space model: the local level model

- Local level model:

$$Y_t = X_t + W_t, \quad \{W_t\} \sim \mathrm{N}(0, \sigma_W^2)$$
$$X_t = X_{t-1} + V_t, \quad \{V_t\} \sim \mathrm{N}(0, \sigma_V^2)$$

  and X_0 is a R.V. that

  - is uncorrelated with W_t's and V_t's
  - has $\mathrm{E}(X_0) = \mu_0$ and $\mathrm{Var}(X_0) = \sigma_0^2$

- Filtering problem is to predict unknown state X_t based on data up to time t, i.e., $\boldsymbol{Y}_{1:t} = (y_1, \cdots, y_t)^T$

Filtering for Local Level Model: II

Best linear predictor of X_t given $\boldsymbol{Y}_{1:t}$ is

$$\mu_t^a \stackrel{\text{def}}{=} \mathrm{E}(X_t|\boldsymbol{Y}_{1:t}) = \mu_t + \Sigma_{t,t}^T \Sigma_{Y,t}^{-1} (\boldsymbol{Y}_{1:t} - \boldsymbol{\mu}_{1:t}),$$

where

- $\mu_t = \mathrm{E}(X_t)$, $\boldsymbol{\mu}_{1:t}$ is a vector containing, for $j = 1, \cdots, t$,

$$\mu_j \stackrel{\text{def}}{=} \mathrm{E}(X_j) = \mathrm{E}(X_j + W_j) = \mathrm{E}(Y_j)$$

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- Vector $\Sigma_{t,t}$ contains covarinces between X_t and $\boldsymbol{Y}_{1:t}$
- (i, j)th element of matrix $\Sigma_{Y,t}$ is covariance between Y_i and Y_j

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- With $\sigma_t^2 \stackrel{\text{def}}{=} \mathrm{Var}(X_t)$, MSE for predictor is

$$\mathrm{E}[(X_t - \mu_t^a)^2] = \sigma_t^2 - \Sigma_{t,t}^T \Sigma_{Y,t}^{-1} \Sigma_{t,t} \stackrel{\text{def}}{=} \Sigma_t^a$$

Forecasting for Local Level Model: I

Forecasting: estimate X_{t+1} given $\boldsymbol{Y}_{1:t}$

- Best linear predictor of X_{t+1} given $\boldsymbol{Y}_{1:t}$ is

$$\mu^f_{t+1} \stackrel{\text{def}}{=} \mathrm{E}(X_{t+1}|\boldsymbol{Y}_{1:t}) = \mu_{t+1} + \Sigma^T_{t+1,t}\Sigma^{-1}_{Y,t}(\boldsymbol{Y}_{1:t} - \boldsymbol{\mu}_{1:t}),$$

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- MSE for predictor is

$$\mathrm{E}[(X_{t+1} - \mu_{t+1}^f)^2] = \sigma_{t+1}^2 - \Sigma_{t+1,t}^T \Sigma_{Y,t}^{-1} \Sigma_{t+1,t} \stackrel{\text{def}}{=} \Sigma_{t+1}^f$$

Forecasting for Local Level Model: II

- Let's also consider best linear predictor of Y_{t+1} given $\boldsymbol{Y}_{1:t}$:

$$Y_{t+1}^{t} \stackrel{\text{def}}{=} \mathrm{E}(Y_{t+1}|\boldsymbol{Y}_{1:t}) = \mu_{Y,t+1} + \tilde{\Sigma}_{t+1,t}^{T}\Sigma_{Y,t}^{-1}(\boldsymbol{Y}_{1:t} - \boldsymbol{\mu}_{Y,1:t}),$$

where the vector $\tilde{\Sigma}_{t+1,t}$ has covaraínces between Y_{t+1} and $\boldsymbol{Y}_{1:t}$

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  where the vector $\tilde{\Sigma}_{t+1,t}$ has covarainces between Y_{t+1} and $\boldsymbol{Y}_{1:t}$

- However, note that, for $j = 1, \cdots, t$

$$\mathrm{Cov}(Y_{t+1}, Y_j) = \mathrm{Cov}(X_{t+1} + W_{t+1}, Y_j) = \mathrm{Cov}(X_{t+1}, Y_j)$$

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- Thus $\tilde{\Sigma}_{t+1,t} = \Sigma_{t+1,t}$, yielding

$$Y_{t+1}^{t} = \mu_{Y,t+1} + \Sigma_{t+1,t}^{T}\Sigma_{Y,t}^{-1}(\boldsymbol{y}_{1:t} - \boldsymbol{\mu}_{Y,1:t}) = \mu_{t+1}^{f}$$

  $\Rightarrow$ difference between Y_{t+1} and X_{t+1} is W_{t+1}, therefore they have the same estimator, but their MSEs differ:

$$\mathrm{E}\left[(Y_{t+1} - Y_{t+1}^{f})^2\right] = \Sigma_{t+1}^{f} + \sigma_W^2$$

Filtering for Local Level Model: III

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- To compute Σ_t^a, i.e., MSE for μ_t^a, need $\sigma_t^2 = \mathrm{Var}(X_t)$ in addition to 2 and 3 above

- Since $X_t = X_{t-1} + V_t$ and $Y_t = X_t + W_t$, telescoping yields $X_j = X_0 + \sum_{l=1}^{j} V_l$ and $Y_j = X_0 + \sum_{l=1}^{j} V_l + W_j$, $j = 1, \cdots, t$

Filtering for Local Level Model: IV

Using

$$X_j = X_0 + \sum_{l=1}^{j} V_l \text{ and } Y_j = X_0 + \sum_{l=1}^{j} + W_j, \quad j = 1, \cdots, t,$$

get $\mu_j = \mathrm{E}[X_j] = \mathrm{E}[X_0] = \mu_0$ and (assuming $j \le k \le t$)

$$\begin{aligned}
\mathrm{Cov}(X_t, Y_j) &= \mathrm{Cov}\left(X_0 + \sum_{l=1}^{t} V_l, X_0 + \sum_{l=1}^{j} V_l + W_j\right) \\
&= \sigma_0^2 + j\sigma_V^2 \\
\mathrm{Cov}(Y_j, Y_k) &= \mathrm{Cov}\left(X_0 + \sum_{l=1}^{j} V_l + W_j, X_0 + \sum_{l=1}^{k} V_l + W_k\right) \\
&= \sigma_0^2 + \min(j,k)\sigma_V^2 + \delta_{jk}\sigma_W^2,
\end{aligned}$$

where $\delta_{jk} = 1$ if $j = k$ and $\delta_{jk} = 0$ if $j \ne k$

Filtering for Local Level Model: V

Using

$$X_t = X_0 + \sum_{l=1}^{t} V_l,$$

get

$$\sigma_t^2 = \mathrm{Var}(X_t) = \sigma_0^2 + t\sigma_V^2$$

- Now we have all the pieces needed to form μ_t^a and its MSE Σ_t^a
- **Note**: similar argument leads to pieces needed to form forecast μ_{t+1}^f and its MSE Σ_{t+1}^f

Kalman Recursions for Filtering/Forecasting: I

While straightforward conceptually, forming

$$\mu_t^a = \mu_t + \Sigma_{t,t}^T \Sigma_{Y,t}^{-1} (\boldsymbol{Y}_{1:t} - \boldsymbol{\mu}_{1:t})$$

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via these equations requires inversion of matrix $\Sigma_{Y,t}$ whose dimension $t \times t$ becomes problematic as t gets large ☹

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> ⇒ The celebrated Kalman recursions give a recipe that avoids explicit matrix inversion

- **Idea**: at time $t-1$, we have 4 quantities of interest: fitted value μ_{t-1}^a, and forecast μ_t^f and their associated MSEs Σ_{t-1}^a and Σ_t^f
- **Note**: $\mu_{t-1}^a = \mu_t^f$ for local level model (but not others)

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  1. steps 1 and 2 are preparatory

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  - fitted values μ_t^a and forecast μ_{t+1}^f
  - associated MSEs Σ_t^a and Σ_{t+1}^f
- There are six steps in the Kalman recursion:
  1. steps 1 and 2 are preparatory
  2. steps 3 and 4 yield μ_t^a and Σ_t^a (filtering)

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- At time t, new observation Y_t becomes available
- Kalman recursion takes μ_t^f, Σ_t^f and Y_t and yields
  - fitted values μ_t^a and forecast μ_{t+1}^f
  - associated MSEs Σ_t^a and Σ_{t+1}^f
- There are six steps in the Kalman recursion:
  1. steps 1 and 2 are preparatory
  2. steps 3 and 4 yield μ_t^a and Σ_t^a (filtering)
  3. steps 5 and 6 yield μ_{t+1}^f and Σ_{t+1}^f (forecasting)

Kalman Recursions for Filtering/Forecasting: III

1. Compute innovation:

$$U_t = Y_t - Y_t^{t-1} = Y_t - \mu_t^f$$

2. Compute MSE for Y_t^{t-1}:

$$\Sigma_t^f + \sigma_W^2 \stackrel{\text{def}}{=} F_t$$

3. Compute new filtered value:

$$\mu_t^a = \mu_t^f + K_t U_t,$$

where $K_t \stackrel{\text{def}}{=} \Sigma_t^f / F_t$ is the so-called Kalman gain

4. Compute MSE for new filtered value:

$$\Sigma_t^a = \Sigma_t^f (1 - K_t)$$

Kalman Recursions for Filtering/Forecasting: IV

5. Compute new forecast:

$$\mu_{t+1}^f = \mu_t^f + K_t U_t = \mu_t^a$$

6. Compute MSE for new forecast:

$$\Sigma_{t+1}^f = \Sigma_t(1 - K_t) + \sigma_W^2 = \Sigma_t^a + \sigma_W^2$$

Recursions are carried out for $t = 0, \cdots, n$ with inputs $\mathrm{E}[X_0] = \mu_0$, $\mathrm{Var}(X_0) = \sigma_0^2$ and $Y_t's$