

Lecture 24

Computer Experiments & Principal Component Analysis

STAT 8020 Statistical Methods II
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Notes

Agenda

① Computer Experiments

② Multivariate Analysis

③ Principal component analysis (PCA)

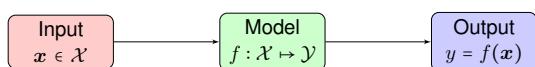


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What is a Computer Experiment

In some situations it is economically, ethically, or simply not possible to run a **physical experiment**. Instead, the following scenario might be feasible:

- the physical process can be described by a mathematical model (e.g., a system of differential equations)
- computer code (simulator) can be written to compute the response from the mathematical model

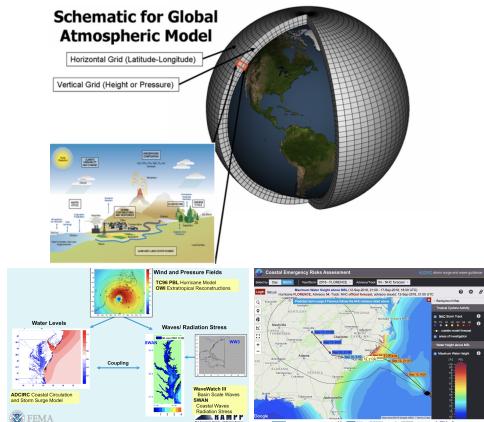


In this case, a researcher can conduct a **computer experiment** by running the computer code, which serves as a proxy for the physical process, to compute a "response" at any combination of values of the inputs



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Examples of Computer Models



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Computer Experiments vs. Physical Experiments

- "Experimental results are believed by everyone, except for the person who ran the experiment"
- "Computational results are believed by no one, except the person who wrote the code"

Replication, randomization and blocking are irreverent for a computer experiment because many computer codes are deterministic and all the inputs to the code are known and can be controlled



Notes

Design & Analysis of Computer Experiments

- **Design:**
where to make the runs, i.e., the selection of inputs $\{x_i\}_{i=1}^n$ where $x_i = (x_{1,i}, x_{2,i}, \dots, x_{d,i})$
- **Analysis:**
fit a statistical model using the model inputs-output $\{y_i, x_i\}_{i=1}^n$ to "emulate" the simulator and to quantify the prediction uncertainty for $y(x_{\text{new}})$, usually via a Gaussian Process Model GP $(m(\cdot), K(\cdot, \cdot))$, where

- $m(x) = E[y(x)]$ is the mean function
- $K(x, x') = \text{Cov}(y(x), y(x'))$ is the covariance function



Notes

An Overview of Multivariate Analysis

- In many studies, observations are collected on **several variables** on each experimental/observational unit

- Multivariate analysis** is a collection of statistical methods for analyzing these multivariate data sets

Common Objectives

- Dimensionality reduction
- Classification
- Grouping (Clustering)



Notes

Multivariate Data

We display a multivariate data that contains n units on p variables using a matrix

$$\mathbf{X} = \begin{pmatrix} X_{1,1} & X_{2,1} & \cdots & X_{p,1} \\ X_{1,2} & X_{2,2} & \cdots & X_{p,2} \\ \vdots & \cdots & \ddots & \vdots \\ X_{1,n} & X_{2,n} & \cdots & X_{p,n} \end{pmatrix}$$

Summary Statistics

- Mean Vector:** $\bar{\mathbf{X}} = (\bar{X}_1, \bar{X}_2, \dots, \bar{X}_p)^T$
- Covariance Matrix:** $\Sigma = \{\sigma_{ij}\}_{i,j=1}^p$, where $\sigma_{ii} = \text{Var}(X_i)$, $i = 1, \dots, p$ and $\sigma_{ij} = \text{Cov}(X_i, X_j)$, $i \neq j$

Next, we are going to introduce **Principal Component Analysis (PCA)**, a useful tool for conducting dimension reduction



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Example: Monthly Sea Surface Temperatures



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Sea Surface Temperatures and Anomalies

- The “data” are gridded at a 2° by 2° resolution from $124^\circ E - 70^\circ W$ and $30^\circ S - 30^\circ N$. The dimension of this SST data set is
2303 (number of grid points in space) \times
552 (monthly time series from 1970 Jan. to 2015 Dec.)
- Sea-surface temperature anomalies are the temperature differences from the climatology (i.e. long-term monthly mean temperatures)
- We will demonstrate the use of Empirical Orthogonal Function (EOF) analysis to uncover the low-dimensional structure of this spatio-temporal data set



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The Empirical Orthogonal Function (EOF) Decomposition

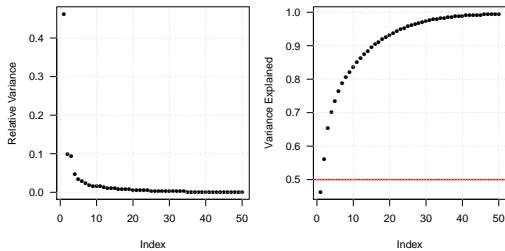
Empirical orthogonal functions (EOFs) are the geophysicist's terminology for the eigenvectors in the eigen-decomposition of an empirical covariance matrix. In its discrete formulation, EOF analysis is simply **Principal Component Analysis (PCA)**. EOFs are usually used

- To find principal spatial structures
- To reduce the dimension (spatially or temporally) in large spatio-temporal datasets



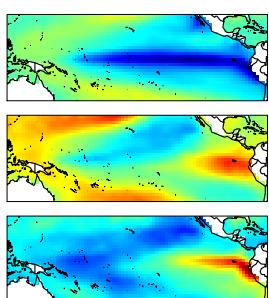
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Screen Plot for EOFs



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Perform EOF Decomposition and Plot the First Three Modes



EOF1: The classic ENSO pattern

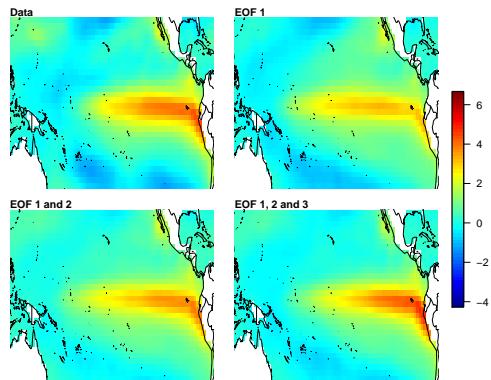
EOF2: A modulation of the center

EOF3: Messing with the coast of SA and the Northern Pacific.



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1998 Jan El Niño Event



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Principal Component Analysis

Given a random sample from a p -dimensional random vector $\mathbf{X}_i = \{X_{1,i}, X_{2,i}, \dots, X_{p,i}\}$, $i = 1, \dots, n$

- Dimension reduction technique
 - Large number of variables (p)
 - Number of variables (p) may be greater than number of observations (n)
- Create new, uncorrelated variables (principal components) for the follow up analysis
 - Principal Component Regression
 - Interpretation of principal components can be difficult in some situations



Notes

Finding Principal Components

Principal Components (PC) are uncorrelated linear combinations $\tilde{X}_1, \tilde{X}_2, \dots, \tilde{X}_p$ determined sequentially, as follows:

- The first PC is the linear combination $\tilde{X}_1 = \mathbf{c}_1^T \mathbf{X} = \sum_{i=1}^p c_{1i} X_i$ that maximize $\text{Var}(\tilde{X}_1)$ subject to $\mathbf{c}_1^T \mathbf{c}_1 = 1$

- The second PC is the linear combination $\tilde{X}_2 = \mathbf{c}_2^T \mathbf{X} = \sum_{i=1}^p c_{2i} X_i$ that maximize $\text{Var}(\tilde{X}_2)$ subject to $\mathbf{c}_2^T \mathbf{c}_2 = 1$ and $\mathbf{c}_2^T \mathbf{c}_1 = 0$

⋮

- The j_{th} PC is the linear combination $\tilde{X}_j = \mathbf{c}_j^T \mathbf{X} = \sum_{i=1}^p c_{ji} X_i$ that maximize $\text{Var}(\tilde{X}_j)$ subject to $\mathbf{c}_j^T \mathbf{c}_j = 1$ and $\mathbf{c}_j^T \mathbf{c}_k = 0 \forall k < j$



Notes

Principal Components

- Let Σ , the covariance matrix of \mathbf{X} , have eigenvalue-eigenvector pairs $(\lambda_i, \mathbf{e}_i)_{i=1}^p$ with $\lambda_1 \geq \lambda_2 \geq \dots \geq \lambda_p \geq 0$. Then, the k_{th} principal component is given by

$$\tilde{X}_k = \mathbf{e}_k^T \mathbf{X} = e_{k1} X_1 + e_{k2} X_2 + \dots + e_{kp} X_p$$

- Then,

$$\text{Var}(\tilde{X}_i) = \lambda_i, \quad i = 1, \dots, p$$

$$\text{Cov}(\tilde{X}_j, \tilde{X}_k) = 0, \quad \forall j \neq k$$



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PCA and Proportion of Variance Explained

- It can be shown that

$$\sum_{i=1}^p \text{Var}(\tilde{X}_i) = \lambda_1 + \lambda_2 + \dots + \lambda_p = \sum_{i=1}^p \text{Var}(X_i)$$

- The proportion of the total variance associated with the k_{th} principal component is given by

$$\frac{\lambda_k}{\lambda_1 + \lambda_2 + \dots + \lambda_p}$$

- If a large proportion of the total population variance (say 80% or 90%) is explained by the first k PCs, then we can restrict attention to the first k PCs without much loss of information

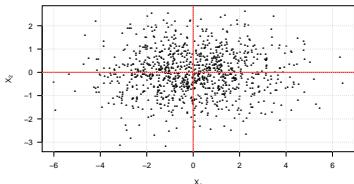


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Toy Example 1

Suppose we have $\mathbf{X} = (X_1, X_2)^T$ where $X_1 \sim N(0, 4)$, $X_2 \sim N(0, 1)$ are independent

- Total variation = $\text{Var}(X_1) + \text{Var}(X_2) = 5$
- X_1 axis explains 80% of total variation
- X_2 axis explains the remaining 20% of total variation

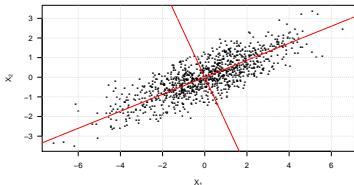


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Toy Example 2

Suppose we have $\mathbf{X} = (X_1, X_2)^T$ where $X_1 \sim N(0, 4)$, $X_2 \sim N(0, 1)$ and $\text{Cor}(X_1, X_2) = 0.8$

- Total variation
 $= \text{Var}(X_1) + \text{Var}(X_2) = \text{Var}(\tilde{X}_1) + \text{Var}(\tilde{X}_2) = 5$
- $\tilde{X}_1 = .9175X_1 + .3975X_2$ explains 93.9% of total variation
- $\tilde{X}_2 = .3975X_1 - .9175X_2$ explains the remaining 6.1% of total variation



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