

4/25/18

Note Title

1/20/2000

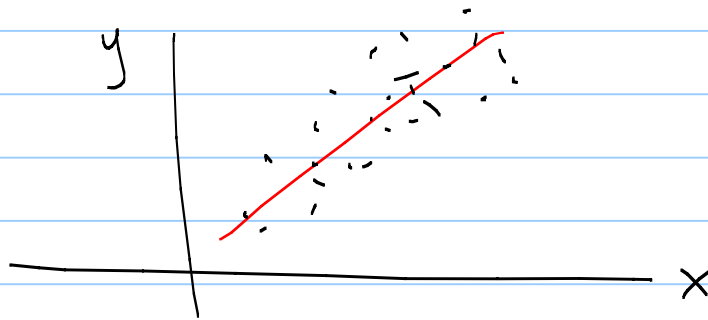
Stochastic Approaches to Linear Regression

Regression data

$$\{(\bar{x}_1, y_1), (\bar{x}_2, y_2), \dots, (\bar{x}_p, y_p)\}, \quad x \in \mathbb{R}^N$$

We model directly the distribution $P_r(y/\bar{x})$
(discriminative methods)

Goal: predict the posterior distribution

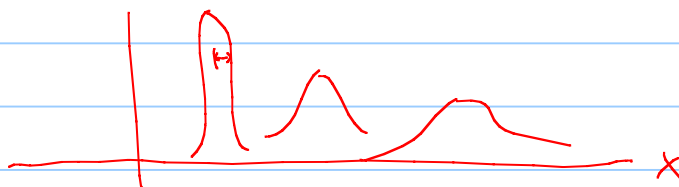


\therefore

$$P_r(y_p / \bar{x}_p, \bar{\theta}) = \text{Norm}_{y_p} [b + \bar{x}_p^T \bar{w}, \sigma^2], \quad \bar{\theta} = \{b, \bar{w}, \sigma^2\}$$

Univariate Normal or Gaussian

$$P_r(x) = \frac{1}{\sqrt{2\pi}\sigma^2} \exp\left[-\frac{(x-\mu)^2}{2\sigma^2}\right] = \text{Norm}_x[\mu, \sigma^2]$$



$$\tilde{x}_p = [1 \ \bar{x}_p]^T, \quad \tilde{w} = [b \ \bar{w}]^T$$

$$P_r(y_p / \tilde{x}_p, \bar{\theta}) = \text{Norm}_{y_p} [\tilde{x}_p^T \tilde{w}, \sigma^2]$$

Learning

$$\text{Find } \bar{\theta} = \{\bar{w}, \sigma^2\}$$

Multivariate Normal

$$(x \in \mathbb{R}^N) \quad \Pr(\bar{x}) = \frac{1}{(2\pi)^{N/2} |\Sigma|^{1/2}} \exp\left[-\frac{1}{2} (\bar{x} - \bar{\mu})^T \Sigma^{-1} (\bar{x} - \bar{\mu})\right]$$
$$= \text{Norm}_{\bar{x}}[\bar{\mu}, \Sigma]$$

- covariance matrix Σ , $N \times N$, symmetric, positive definite
- examples (2D)

$$\Sigma_{\text{spher}} = \begin{bmatrix} \sigma^2 & 0 \\ 0 & \sigma^2 \end{bmatrix} = \sigma^2 I, \quad \Sigma_{\text{diag}} = \begin{bmatrix} \sigma_1^2 & 0 \\ 0 & \sigma_2^2 \end{bmatrix}, \quad \Sigma_{\text{full}} = \begin{bmatrix} \sigma_{11}^2 & \sigma_{12}^2 \\ \sigma_{21}^2 & \sigma_{22}^2 \end{bmatrix}$$
$$\sigma_{12} = \sigma_{21}$$

- When the covariance is spherical or diagonal, the individual variables are independent
- e.g. 2D diagonal

joint $\rightarrow \Pr(x_1, x_2) = \frac{1}{2\pi \sqrt{|\Sigma|}} \exp\left[-\frac{1}{2} (x_1, x_2) \Sigma^{-1} \begin{pmatrix} x_1 \\ x_2 \end{pmatrix}\right]$

$$= \frac{1}{2\pi \sqrt{|\Sigma|}} \exp\left[-\frac{1}{2} (x_1, x_2) \begin{pmatrix} \sigma_1^2 & 0 \\ 0 & \sigma_2^2 \end{pmatrix} \begin{pmatrix} x_1 \\ x_2 \end{pmatrix}\right]$$
$$= \frac{1}{\sqrt{2\pi\sigma_1^2}} \exp\left[-\frac{x_1^2}{2\sigma_1^2}\right] \cdot \frac{1}{\sqrt{2\pi\sigma_2^2}} \exp\left[-\frac{x_2^2}{2\sigma_2^2}\right]$$
$$= \Pr(x_1) \cdot \Pr(x_2)$$

$|\Sigma| = \sigma_1^2 \cdot \sigma_2^2$

Form matrices & vector

$$\bar{y} = [y_1, \dots, y_p]^T, \quad \tilde{X} = [\tilde{x}_1, \tilde{x}_2, \dots, \tilde{x}_p] = \begin{bmatrix} 1 & 1 & 1 \\ \bar{x}_1 & \bar{x}_2 & \dots & \bar{x}_p \end{bmatrix}$$

$$\tilde{W} = [b \ w_1 \ \dots \ w_N]^T_{(N+1) \times 1} \quad (N+1) \times P$$

$$Pr(\bar{y} | \tilde{X}, \bar{\theta}) = \text{Norm}_{\bar{y}} [\tilde{X}^T \tilde{W}, \sigma^2 I]$$

Learning

$$\hat{\bar{\theta}} = \arg \max_{\bar{\theta}} [Pr(\bar{y} | \tilde{X}, \bar{\theta})] = \arg \max_{\bar{\theta}} \left\{ \log [Pr(\bar{y} | \tilde{X}, \bar{\theta})] \right\}$$

$$\hat{\tilde{W}}, \hat{\sigma}^2 = \arg \max_{\tilde{W}, \sigma^2} \left\{ \log \left[\frac{1}{(2\pi)^{P/2} (\sigma^2)^{P/2}} \exp \left[-\frac{1}{2\sigma^2} (\bar{y} - \tilde{X}\tilde{W})^T (\bar{y} - \tilde{X}\tilde{W}) \right] \right] \right\}$$

$$= \arg \max_{\tilde{W}, \sigma^2} \left\{ \underbrace{-\frac{P}{2} \log(2\pi) - \frac{P}{2} \log(\sigma^2) - \frac{(\bar{y} - \tilde{X}\tilde{W})^T (\bar{y} - \tilde{X}\tilde{W})}{2\sigma^2}}_a \right\}$$

$$\nabla_{\tilde{W}} a = 0$$

$$\Rightarrow 2\tilde{X}\tilde{X}^T \tilde{W} - 2\tilde{X}\bar{y} = 0 \Rightarrow \tilde{X}\tilde{X}^T \tilde{W} = \tilde{X}\bar{y}$$

$$\Rightarrow \tilde{W} = (\tilde{X}\tilde{X}^T)^{-1} \tilde{X}\bar{y}$$

if $\tilde{X}\tilde{X}^T$ invertible

EXACTLY the same expression we derived using least-squares.

$$\frac{da}{d\sigma^2} = 0$$

$$\Rightarrow -\frac{P}{2} \frac{1}{\sigma^2} + \frac{(\bar{y} - \tilde{X}\tilde{W})^2}{2\sigma^4} = 0 \Rightarrow \sigma^2 = \frac{(\bar{y} - \tilde{X}\tilde{W})^2}{P}$$

3 issues

1. predictions over-confident
2. only linear function
- 3) all \bar{x} are involved