0.0.1 Inference

Testing models Consider the linear model

$$\mathbb{Y} = \mathbb{X}\boldsymbol{\beta} + \boldsymbol{\varepsilon} \quad \cdots \quad \text{full model}$$

where
$$\gamma(X) = \gamma \le p$$
, $E(\varepsilon) = 0$, $Var(\varepsilon) = \sigma^2 I$

Consider the linear model

$$y = \mathbb{W}\gamma + \varepsilon$$
 ··· redued model

where
$$C(\mathbb{W}) \subset C(\mathbb{X})$$
, $E(\varepsilon) = 0$, $Var(\varepsilon) = \sigma^2 I$

note:

- 1. the estimation space of R.M. is smaller than in the F.M.
- 2. The goal is to test whether or not the reduced model is also correct.
 - If R.M. is correct, there is no reason not to use it.
 - Smaller models are easier to interpret.
- 3. Let $P_{\mathbb{X}}$ and $P_{\mathbb{W}}$ denote the ppm onto $C(\mathbb{X})$ and $C(\mathbb{W})$. Because $C(\mathbb{W}) \subset C(\mathbb{X})$, we know that $P_{\mathbb{X}} P_{\mathbb{W}}$ is the ppm onto $C(P_{\mathbb{X}} P_{\mathbb{W}}) = \underline{C(\mathbb{W})^{\perp} C(\mathbb{X})}$

Lemma 1 Assume $P_{\mathbb{X}}$ and $P_{\mathbb{W}}$ are projection matries and $P_2 - P_1$ is P.D., then

- 1. $P_1P_2 = P_2P_1 = P_2$
- 2. $P_1 P_2$ is a p.m.

note:

- 1. P_1 is a p.m onto C(X)
- 2. P_2 is a p.m onto $C(\mathbb{W}) \subset C(\mathbb{X})$
- 3. $P_1 P_2$ is a p.m onto the orthogonal complement of W with in X

figure

Under F.M., our estimate for $E(y) = x\beta$ is $P_{\mathbb{X}}\mathbb{Y}$ Under R.M., our estimate is $E(y) = \mathbb{W}\gamma$ is $P_{\mathbb{W}}\mathbb{Y}$

1. If R.M. is correct, then the $P_{\mathbb{X}}$ and $P_{\mathbb{W}}$ are estimates the same thing. (i.e. $P_{\mathbb{X}}\mathbb{Y} - P_{\mathbb{W}}\mathbb{Y} = (P_{\mathbb{X}} - P_{\mathbb{W}})\mathbb{Y}$ should be small.)

2. the size of $(P_{\mathbb{X}} - P_{\mathbb{W}}) \, \mathbb{Y}$ is

$$\frac{\left[\left(P_{\mathbb{X}}-P_{\mathbb{W}}\right)\mathbb{Y}\right]^{\dagger}\left[\left(P_{\mathbb{X}}-P_{\mathbb{W}}\right)\mathbb{Y}\right]}{\gamma\left(P_{\mathbb{X}}-P_{\mathbb{W}}\right)}=\frac{\mathbb{Y}^{\dagger}\left(P_{\mathbb{X}}-P_{\mathbb{W}}\right)\mathbb{Y}}{\gamma\left(P_{\mathbb{X}}-P_{\mathbb{W}}\right)}$$

If R.M. is correct, then

$$E\left(\frac{\mathbb{Y}^{\dagger}\left(P_{\mathbb{X}}-P_{\mathbb{W}}\right)\mathbb{Y}}{\gamma\left(P_{\mathbb{X}}-P_{\mathbb{W}}\right)}\right) = \frac{1}{\gamma^{*}}E\left(\mathbb{Y}^{\dagger}\left(P_{\mathbb{X}}-P_{\mathbb{W}}\right)\mathbb{Y}\right)$$

$$= \frac{1}{\gamma^{*}}\left\{tr\left[\left(P_{\mathbb{X}}-P_{\mathbb{W}}\right)\sigma^{2}I\right] + \left(W\gamma\right)^{\dagger}\left(P_{\mathbb{X}}-P_{\mathbb{W}}\right)\left(W\gamma\right)\right\}$$

$$= \frac{1}{\gamma^{*}}\left\{\sigma^{2}tr\left(P_{\mathbb{X}}-P_{\mathbb{W}}\right) + \gamma^{\dagger}W^{\dagger}\left(P_{\mathbb{X}}-P_{\mathbb{W}}\right)W\gamma\right\}$$

$$= \frac{1}{\gamma^{*}}\left\{\sigma^{2}\gamma\left(P_{\mathbb{X}}-P_{\mathbb{W}}\right)\right\}$$

$$= \sigma^{2} \quad \left(\text{U.E. of } \sigma^{2}\right)$$

If R.M. is not correct, then

$$\begin{split} E\left(\frac{\mathbb{Y}^{\dagger}\left(P_{\mathbb{X}}-P_{\mathbb{W}}\right)\mathbb{Y}}{\gamma^{*}}\right) &= \frac{1}{\gamma^{*}}\left\{tr\left[\left(P_{\mathbb{X}}-P_{\mathbb{W}}\right)\sigma^{2}I\right]\right. \\ &+ \left(\mathbb{X}\beta\right)^{\dagger}\left(P_{\mathbb{X}}-P_{\mathbb{W}}\right)\left(\mathbb{X}\beta\right)\right\} \\ &= \sigma^{2}+\left(\mathbb{X}\beta\right)^{\dagger}\left(P_{\mathbb{X}}-P_{\mathbb{W}}\right)\mathbb{X}\beta \\ \\ MSE &= \frac{\mathbb{Y}^{\dagger}\left(I-P_{\mathbb{X}}\right)\mathbb{Y}}{\gamma\left(I-P_{\mathbb{X}}\right)} &= \frac{1}{n-\gamma}\mathbb{Y}^{\dagger}\left(I-P_{\mathbb{X}}\right)\mathbb{Y} \\ E\left(MSE\right) &= \sigma^{2} \end{split}$$

Test Statistic

To test R.M. vs F.M., we use

$$F = \frac{\mathbb{Y}^{\dagger} (I - P_{\mathbb{X}}) \, \mathbb{Y} / \gamma^{*}}{\mathbb{Y}^{\dagger} (I - P_{\mathbb{X}}) \, \mathbb{Y} / (n - \gamma)}$$

$$= \begin{cases} \simeq 1 & \text{, R.M. is correct} \\ > 1 & \text{, R.M. is not correct} \end{cases}$$

$$= \frac{\text{MSR}}{\text{MSE}}$$