## STOCHASTIC PROCESSES LECTURE#6

## 6.1 Non-decreasing families of $\sigma$ -algebras in $\Omega$

With any stochastic process  $(X_t, t \in T)$  we can associate a family of  $\sigma$ -algebras in the sample space  $\Omega$ 

$$\mathcal{F}_t = \sigma(X_u, u \le t) = \sigma\{\{\omega \in \Omega : X_u(\omega) \in C\}, u \le t, C \in \mathcal{S}\}$$
  $t \in T$ 

The family (6–1) is non-decreasing:  $\mathcal{F}_{t_1} \subseteq \mathcal{F}_{t_2}$  for  $t_1 < t_2$ ; it is often referred to as a filtration.

## 6.2 THE WIENER PROCESS

A Wiener process is defined as a stochastic process  $W_t$ , where the time parameter t changes in some interval, satisfying the following conditions:

1) For  $t_0 < t_1 < \ldots < t_n$  the increments

$$W(t_1) - W(t_0)$$
  $W(t_2) - W(t_1)$  ...  $W(t_n) - W(t_{n-1})$ 

are independent random variables.

- 2) For two time moments s < t, the increment  $W_t W_s$  has the normal distribution with parameters (0, t s) (i.e., with the expectation equal to 0 and the variance equal to t s).
- 3) The trajectories  $W_{\bullet}(\omega)$  are continuous functions of t for every  $\omega \in \Omega$ .

Question: is the Wiener process determined uniquely by the requirements 1) and 2) in the sense of its finite-dimensional distributions? The answer is NO because if  $W_t$  is a Wiener process, so is  $W_t + C$  where C is an arbitrary constant. But if we impose on the Wiener process the condition that it is equal to some constant  $x_0$  at some time moment  $t_0$  then its finite-dimensional distributions are determined in a unique way by the requirements 1–2).

## THEOREM 6.1

Let  $W_t$  be the Wiener process defined for  $t \in [t_0, +\infty[$  and such that

$$W(t_0) = x_0$$

The *n*-dimensional distribution  $\mu_{t_1,...,t_n}$  for *n* time moments  $t_1 < ... < t_n$  strictly greater than  $t_0$  is a continuous distribution characterized by the following probability density function:

$$p_{t_1,\dots,t_n}(x_1,\dots,x_n) = \frac{1}{\sqrt{2\pi(t_1-t_0)}} e^{-\frac{(x_1-x_0)^2}{2(t_1-t_0)}} \times \frac{1}{\sqrt{2\pi(t_2-t_1)}} e^{-\frac{(x_2-x_1)^2}{2(t_2-t_1)}} \times \dots \times \frac{1}{\sqrt{2\pi(t_n-t_{n-1})}} e^{-\frac{(x_n-x_{n-1})^2}{2(t_n-t_{n-1})}}$$

Now the question that remains is: does such a stochastic process exist? Here the Kolmogorov consistency condition comes into play (see Lecture #2).

THE CONSISTENCY CONDITION: It is enough to verify that integrating with respect to  $x_i$ 

$$(6\text{--2}) \quad \int_{-\infty}^{+\infty} \left[ \frac{1}{\sqrt{2\pi(t_i-t_{i-1})}} e^{-\frac{(x_i-x_{i-1})^2}{2(t_i-t_{i-1})}} \times \frac{1}{\sqrt{2\pi(t_{i+1}-t_i)}} e^{-\frac{(x_{i+1}-x_i)^2}{2(t_{i+1}-t_i)}} \right] dx_i$$

results in

$$\frac{1}{\sqrt{2\pi(t_{i+1}-t_{i-1})}}e^{-\frac{(x_{i+1}-x_{i-1})^2}{2(t_{i+1}-t_{i-1})}}$$

The rest of the factors in the densities

$$p_{t_1,\ldots,t_{i-1},t_i,t_{i+1},\ldots,t_n}(x_1,\ldots,x_{i-1},x_i,x_{i+1},\ldots,x_n)$$

and

$$p_{t_1,\ldots,t_{i-1},t_{i+1},\ldots,t_n}(x_1,\ldots,x_{i-1},x_{i+1},\ldots,x_n)$$

are the same.

The integral in Eq.(6–2) is the formula for the distribution density of the sum of two independent random variables

$$p_{U+V}(v) = \int_{-\infty}^{+\infty} p_U(u) p_V(v-u) du$$
  $(u = x_i, v = x_{i+1})$ 

where U has the normal distribution with parameters  $(x_{i-1}, [t_i - t_{i-1}])$  and V has the normal distribution with parameters  $(0, [t_{i+1} - t_i])$ . U + V being the sum of two independent normal random variables, it also has the normal distribution with expectation  $x_{i-1} + 0$  and variance  $(t_i - t_{i-1}) + (t_{i+1} - t_i) = t_{i+1} - t_{i-1}$ .

Back to the Wiener process definition: Let  $\{\mathcal{F}_t\}_{t\geq t_0}$  be the non-decreasing family of  $\sigma$ -algebras (or filtration) generated by the Wiener process  $(W_t, t \geq t_0)$  starting from  $x_0$ . The condition 1) can be reformulated as

1') For s < t the increments  $W_t - W_s$  is independent of the  $\sigma$ -algebra  $\mathcal{F}_s = \sigma(W_u, u \le s)$ .