

机器学习理论导引

作业三

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1 [20pts] Concentration Inequalities

- (1) [5pts] 请利用高斯函数的性质推导 $\Pr[\bar{X} - \mu \geq \epsilon]$ 上界.
- (2) [5pts] 请给出 Bernstein 不等式以 $1 - \delta$ 概率表达形式.
- (3) [10pts] 如果随机变量 $X \in [\alpha, \beta]$, 其中 $\beta > \alpha$ 是两个常数。试证明, 随机变量 X 是一个亚高斯随机变量 (sub-Gaussian Random Variable), 且满足其亚高斯系数 $b = (\beta - \alpha)^2/4$, 即对于任意 $\lambda > 0$, 有下式成立

$$\ln \mathbb{E}[\exp\{\lambda(X - \mu)\}] \leq \lambda^2(\beta - \alpha)^2/8, \quad (1.1)$$

其中 $\mu = \mathbb{E}[X]$.

Proof.

(1) $X \sim N(\mu, \sigma^2) \implies \sum_{i=1}^m X_i \sim N(m\mu, m\sigma^2) \implies \bar{X} = \frac{1}{m} \sum_{i=1}^m X_i \sim N(\mu, \frac{\sigma^2}{m})$, 所以 $\Pr(\bar{X} - \mu \geq \epsilon) = \Pr(\bar{X} \geq \mu + \epsilon) = \int_{\mu+\epsilon}^{+\infty} \frac{1}{\sqrt{2\pi\frac{\sigma^2}{m}}} \exp(-\frac{(x-\mu)^2}{2\frac{\sigma^2}{m}}) dx$, 令 $\sigma_0^2 = \frac{\sigma^2}{m}$, 于是有:

$$\begin{aligned} \Pr(\bar{X} \geq \mu + \epsilon) &= \int_{\mu+\epsilon}^{+\infty} \frac{1}{\sqrt{2\pi\sigma_0^2}} \exp(-\frac{(x-\mu)^2}{2\sigma_0^2}) dx \\ &= \int_{\epsilon}^{+\infty} \frac{1}{\sqrt{2\pi\sigma_0^2}} \exp(-\frac{x^2}{2\sigma_0^2}) dx \\ &= \int_0^{+\infty} \frac{1}{\sqrt{2\pi\sigma_0^2}} \exp(-\frac{(x+\epsilon)^2}{2\sigma_0^2}) dx \\ &= \int_0^{+\infty} \frac{1}{\sqrt{2\pi\sigma_0^2}} \exp(-\frac{x^2}{2\sigma_0^2}) \exp(-\frac{\epsilon^2}{2\sigma_0^2}) \exp(-\frac{2x\epsilon}{2\sigma_0^2}) dx \\ &= \exp(-\frac{\epsilon^2}{2\sigma_0^2}) \int_0^{+\infty} \frac{1}{\sqrt{2\pi\sigma_0^2}} \exp(-\frac{x^2}{2\sigma_0^2}) \exp(-\frac{2x\epsilon}{2\sigma_0^2}) dx \\ &\leq \exp(-\frac{\epsilon^2}{2\sigma_0^2}) \int_0^{+\infty} \frac{1}{\sqrt{2\pi\sigma_0^2}} \exp(-\frac{x^2}{2\sigma_0^2}) dx \\ &= \exp(-\frac{\epsilon^2}{2\sigma_0^2}) = \exp(-\frac{m\epsilon^2}{2\sigma^2}) \end{aligned}$$

(2) $\Pr(\bar{X}_n \geq \mu + \epsilon) \leq \exp(-\frac{n\epsilon^2}{2V+2b\epsilon}) = \delta \implies \epsilon = \frac{1}{n}(\sqrt{b^2 \ln^2 \delta - 2nV \ln \delta} - b \ln \delta)$, 所以 $\Pr(\bar{X}_n \geq \mu + \frac{\sqrt{b^2 \ln^2 \delta - 2nV \ln \delta} - b \ln \delta}{n}) \leq \delta$ 。

(3) 令 $Y = X - \mu \in [\alpha - \mu, \beta - \mu]$, 其中 $\mu = E(X)$, 记 $a = \alpha - \mu$, $b = \beta - \mu$, 则 $Y \in [a, b]$ 。令 $Z = \frac{Y-a}{b-a}$, 则 $Y = (b-a)Z + a = Zb + (1-Z)a$ 。因为 $Z = \frac{Y-a}{b-a} \in [0, 1]$, 由 Jensen 不等式有 $e^{\lambda Y} = e^{Z\lambda b + (1-Z)\lambda a} \leq Ze^{\lambda b} + (1-Z)e^{\lambda a}$ 。对不等式两边求取一下期望:

$$\begin{aligned} \mathbb{E}(e^{\lambda Y}) &\leq e^{\lambda b} \mathbb{E}(Z) + e^{\lambda a} (1 - \mathbb{E}(Z)) \\ &= e^{\lambda b} \frac{b-a}{b-a} + e^{\lambda a} \frac{a}{b-a} \\ &= e^{\lambda a} \left(\frac{b}{b-a} - \frac{a}{b-a} e^{\lambda(b-a)} \right) \\ &= e^{\frac{a}{b-a} \lambda(b-a)} \left(1 + \frac{a}{b-a} - \frac{a}{b-a} e^{\lambda(b-a)} \right) \end{aligned}$$

对两边求一下对数, 则有:

$$\ln \mathbb{E}(e^{\lambda Y}) \leq \frac{a}{b-a} \lambda(b-a) + \ln \left(1 + \frac{a}{b-a} - \frac{a}{b-a} e^{\lambda(b-a)} \right)$$

令 $x = \lambda(b-a)$, $t = \frac{a}{b-a}$, 则 $\frac{a}{b-a} \lambda(b-a) + \ln \left(1 + \frac{a}{b-a} - \frac{a}{b-a} e^{\lambda(b-a)} \right) = tx + \ln(1+t-te^x)$, 令 $h(x) = tx + \ln(1+t-te^x)$, 则有 $h'(x) = t + \frac{-te^x}{1+t-te^x}$, 进一步有 $h''(x) = \frac{-te^x(1+t-te^x) - (-te^x)^2}{(1+t-te^x)^2} = -\left(\frac{te^x}{1+t-te^x}\right) - \left(\frac{te^x}{1+t-te^x}\right)^2 = -\left(\frac{te^x}{1+t-te^x} + \frac{1}{2}\right)^2 + \frac{1}{4} \leq \frac{1}{4}$, 由泰勒展开定理知存在 $\xi \in [0, x]$ 使得:

$$\begin{aligned} h(x) &= h(0) + xh'(0) + \frac{x^2}{2}h''(\xi) \\ &= 0 + x \cdot 0 + \frac{x^2}{2}h''(\xi) \\ &\leq \frac{x^2}{2} \cdot \frac{1}{4} = \frac{\lambda^2(b-a)^2}{8} \end{aligned}$$

代回则有:

$$\ln \mathbb{E}(e^{\lambda Y}) = \ln \mathbb{E}(e^{\lambda(X-\mu)}) \leq h(x) \leq \frac{\lambda^2(b-a)^2}{8} = \frac{\lambda^2(\beta-\alpha)^2}{8}$$

□

2 [30pts] McDiarmid Inequality

请利用 Chernoff 引理证明如下的 McDiarmid 不等式:

定理 1 (McDiarmid Inequality). X_1, X_2, \dots, X_m 为定义在空间 \mathcal{X} 上的 m 个独立的随机变量. 假设 $f: \mathcal{X}^m \mapsto \mathbb{R}$ 是一个关于 X_1, \dots, X_m 的实值函数, 并满足对任意的 $x_1, x_2, \dots, x_m, x'_i \in \mathcal{X}$, 都有

$$|f(x_1, \dots, x_i, \dots, x_m) - f(x_1, \dots, x'_i, \dots, x_m)| \leq c_i$$

成立, 其中 $c_i > 0$. 那么对任意 $\epsilon > 0$, 有

$$\Pr[f(X_1, \dots, X_m) - \mathbb{E}[f(X_1, \dots, X_m)] \geq \epsilon] \leq e^{-2\epsilon^2 / \sum_{i=1}^m c_i^2}.$$

请写出详细的证明过程。

Proof. slides 中的证明跳了一些步骤我没能补齐, 这里的证明方法参考自一篇 *stack exchange* 上的回答 [1]。

记 $f(X_{1:m}) = f(X_1, \dots, X_m)$. 令 $V_i = \mathbb{E}[f(X_{1:m}) | X_1, \dots, X_i] - \mathbb{E}[f(X_{1:m}) | X_1, \dots, X_{i-1}]$, $1 < i \leq m$, 对于边界情况有 $V_1 = \mathbb{E}[f(X_{1:m}) | X_1] - \mathbb{E}[f(X_{1:m})]$, 易见 $f(X_{1:m}) - \mathbb{E}[f(X_{1:m})] = \sum_{i=1}^m V_i$. 由马尔科夫不等式有:

$$\begin{aligned} \Pr[f(X_{1:m}) - \mathbb{E}[f(X_{1:m})] \geq \epsilon] &= \Pr[\lambda(f(X_{1:m}) - \mathbb{E}[f(X_{1:m})]) \geq \lambda\epsilon] \\ &= \Pr[\exp\{\lambda(f(X_{1:m}) - \mathbb{E}[f(X_{1:m})])\} \geq \exp(\lambda\epsilon)] \\ &= \Pr\left[\exp\left\{\lambda \sum_{i=1}^m V_i\right\} \geq \exp(\lambda\epsilon)\right] \\ &\leq e^{-\lambda\epsilon} \mathbb{E}\left[\exp\left\{\lambda \sum_{i=1}^m V_i\right\}\right] \end{aligned}$$

先给 $\mathbb{E}[e^{\lambda V_i}]$ 一个上界, 令

$$\begin{aligned} H_i &= \sup V_i = \sup_u (\mathbb{E}[f(X_{1:m} | X_1, \dots, X_{i-1}, u)] - \mathbb{E}[f(X_{1:m} | X_1, \dots, X_{i-1})]) \\ L_i &= \inf V_i = \inf_v (\mathbb{E}[f(X_{1:m} | X_1, \dots, X_{i-1}, v)] - \mathbb{E}[f(X_{1:m} | X_1, \dots, X_{i-1})]) \end{aligned}$$

对两者逐差有:

$$\begin{aligned} H_i - L_i &= \sup_u \mathbb{E}[f(X_{1:m} | X_1, \dots, X_{i-1}, u)] - \inf_v \mathbb{E}[f(X_{1:m} | X_1, \dots, X_{i-1}, v)] \\ &= \int_{-\infty}^{+\infty} \dots \int_{-\infty}^{+\infty} \sup_u f(x_1, \dots, x_{i-1}, u, x_{i+1}, \dots, x_m) f_{X_{i+1}}(x_{i+1}) \dots f_{X_m}(x_m) dx_{i+1} \dots dx_m \\ &\quad - \int_{-\infty}^{+\infty} \dots \int_{-\infty}^{+\infty} \inf_v f(x_1, \dots, x_{i-1}, v, x_{i+1}, \dots, x_m) f_{X_{i+1}}(x_{i+1}) \dots f_{X_m}(x_m) dx_{i+1} \dots dx_m \\ &= \int_{-\infty}^{+\infty} \dots \int_{-\infty}^{+\infty} \sup_{u,v} (f(x_1, \dots, x_{i-1}, u, x_{i+1}, \dots, x_m) \\ &\quad - f(x_1, \dots, x_{i-1}, v, x_{i+1}, \dots, x_m)) f_{X_{i+1}}(x_{i+1}) \dots f_{X_m}(x_m) dx_{i+1} \dots dx_m \end{aligned}$$

$$\begin{aligned}
& -f(x_1, \dots, x_{i-1}, v, x_{i+1}, \dots, x_m)) f_{X_{i+1}}(x_{i+1}) \dots f_{X_m}(x_m) dx_{i+1} \dots dx_m \\
& \leq \int_{-\infty}^{+\infty} \dots \int_{-\infty}^{+\infty} c_i f_{X_{i+1}}(x_{i+1}) \dots f_{X_m}(x_m) dx_{i+1} \dots dx_m \\
& = c_i
\end{aligned}$$

又由 *chernoff* 引理, 有 $\ln \mathbb{E}[e^{\lambda V_i}] \leq \frac{\lambda^2(H_i - L_i)^2}{8} \leq \frac{\lambda^2 c_i^2}{8}$, 于是:

$$\begin{aligned}
\mathbb{E} \left[\exp \left\{ \lambda \sum_{i=1}^m V_i \right\} \right] &= \mathbb{E} \left[\mathbb{E} \left[\exp \left\{ \lambda \sum_{i=1}^m V_i \right\} \mid X_1, \dots, X_{m-1} \right] \right] \\
&= \mathbb{E} \left[\exp \left\{ \lambda \sum_{i=1}^{m-1} V_i \right\} \mathbb{E} [e^{\lambda V_m} \mid X_1, \dots, X_{m-1}] \right] \\
&\leq \mathbb{E} \left[\exp \left\{ \lambda \sum_{i=1}^{m-1} V_i \right\} \exp \left[\frac{\lambda^2 c_m^2}{8} \right] \right] \\
&\leq \dots \leq \exp \left(\frac{\lambda^2}{8} \sum_{i=1}^m c_i^2 \right)
\end{aligned}$$

令 $\lambda = \frac{4\epsilon}{\sum_{i=1}^m c_i^2}$, 则有:

$$\begin{aligned}
\Pr[f(X_{1:m}) - \mathbb{E}[f(X_{1:m})] \geq \epsilon] &\leq e^{-\lambda \epsilon} \mathbb{E} \left[\exp \left\{ \lambda \sum_{i=1}^m V_i \right\} \right] \\
&\leq e^{-\lambda \epsilon} \exp \left(\frac{\lambda^2}{8} \sum_{i=1}^m c_i^2 \right) \\
&= e^{-2\epsilon^2 / \sum_{i=1}^m c_i^2}
\end{aligned}$$

□

3 [20pts] Basic Inequalities

假设 $\xi_1, \xi_2, \dots, \xi_n$ 是 n 个伯努利随机变量, 且满足 $\xi_i \sim B(1/i)$, 设 $T_n = \sum_{i=1}^n \xi_i$, 试证明

$$\begin{aligned}\mathbb{E}[T_n] &= \sum_{i=1}^n \frac{1}{i} \geq \ln n. \\ \mathbb{V}[T_n] &= \sum_{i=2}^n \frac{1}{i} \left(1 - \frac{1}{i}\right) \leq \ln n + 1.\end{aligned}$$

Proof.

(1) $\mathbb{E}(T_n) = \mathbb{E}(\sum_{i=1}^n \xi_i) = \sum_{i=1}^n \mathbb{E}(\xi_i) = \sum_{i=1}^n \frac{1}{i}$, 而 $\ln n = \int_1^n \frac{1}{x} dx \leq \sum_{i=1}^n \frac{1}{i}$, 所以有 $\mathbb{E}(T_n) = \sum_{i=1}^n \frac{1}{i} \geq \ln n$.

(2) 令 $\mu = \mathbb{E}(T_n)$,

$$\begin{aligned}\mathbb{V}(T_n) &= \mathbb{E}[(T_n - \mathbb{E}(T_n))^2] \\ &= \mathbb{E}[T_n^2 - 2\mu T_n + \mu^2] \\ &= \mathbb{E}(T_n^2) - 2\mu \mathbb{E}(T_n) + \mathbb{E}(\mu^2) \\ &= \mathbb{E}\left[\sum_{i=1}^n \sum_{j=1}^n \xi_i \xi_j\right] - \mu^2 \\ &= \sum_{i=1}^n \sum_{j=1}^n [\mathbb{I}_{i=j} \frac{1}{i} + \mathbb{I}_{i \neq j} \frac{1}{ij}] - \mu^2 \\ &= \sum_{i=1}^n \sum_{j=1}^n [\frac{1}{ij}] + \sum_{i=1}^n [\frac{1}{i} - \frac{1}{i^2}] - \mu^2 \\ &= \sum_{i=1}^n \frac{1}{i} \sum_{j=1}^n [\frac{1}{j}] + \sum_{i=1}^n [\frac{1}{i} - \frac{1}{i^2}] - \mu^2 \\ &= \sum_{i=1}^n \frac{1}{i} \mu + \sum_{i=1}^n [\frac{1}{i} - \frac{1}{i^2}] - \mu^2 \\ &= \mu^2 + \sum_{i=1}^n [\frac{1}{i} - \frac{1}{i^2}] - \mu^2 \\ &= \sum_{i=1}^n [\frac{1}{i} - \frac{1}{i^2}] = \sum_{i=1}^n \frac{1}{i} [1 - \frac{1}{i}] \\ &\leq \sum_{i=2}^n \frac{1}{i} < \int_1^n \frac{1}{x} dx = \ln n < \ln n + 1\end{aligned}$$

□

4 [30pts] Rademacher Complexity based Margin Bounds

记 $S = \{(x_1, y_1), (x_2, y_2), \dots, (x_m, y_m)\}$ 是一个大小为 m 的训练集, 其中每个样本 (x_i, y_i) 是独立同分布地从分布 \mathcal{D} 中采样得到. 试证明, 对任意 $\delta > 0$ 和 $\theta > 0$, 任意分类器 $H \in \mathcal{C}(\mathcal{H})$ 至少以 $1 - \delta$ 的概率满足:

$$\Pr_{(x,y) \sim \mathcal{D}}[yH(x) < 0] \leq \sum_{i=1}^m \mathbb{I}[y_i H(x_i) \leq \theta] + \frac{2}{\theta} \hat{\mathfrak{R}}_S(\mathcal{H}) + 3\sqrt{\frac{\ln(1/\delta)}{2m}}. \quad (4.1)$$

其中, 函数 $\mathbb{I}[\cdot]$, 仅当 \cdot 为真时取值为 1, 否则为 0; $\mathcal{C}(\mathcal{H})$ 为函数空间 \mathcal{H} 的凸包, 具体形式为

$$\mathcal{C}(\mathcal{H}) = \left\{ H = \sum_i \alpha_i h_i : h_i \in \mathcal{H}, \alpha_i \geq 0 \text{ and } \sum_i \alpha_i = 1 \right\}.$$

Proof.

令 $\mathcal{H}_C = \mathcal{C}(\mathcal{H})$, $\tilde{\mathcal{H}}_C = \{(x, y) \mapsto yh(x), h \in \mathcal{H}_C\}$, 记 $\Phi_\rho(x) = \mathbb{I}_{x \leq 0} + (1 - \frac{x}{\rho})\mathbb{I}_{0 < x \leq \rho}$, 同时记 $\tilde{\mathcal{H}}_{C,\rho} = \Phi_\rho \circ \tilde{\mathcal{H}}_C$, 易见 $\tilde{\mathcal{H}}_{C,\rho}$ 中函数均是 $\frac{1}{\rho}$ -lipshitz 的, 且 $\Phi_\rho(x) \in [0, 1]$, 由 slides 泛化性一章结论可知:

$$\mathbb{E}[\Phi_\rho(yh(x))] \leq \frac{1}{m} \sum_{i=1}^m \Phi_\rho(yh(x_i)) + 2\hat{\mathfrak{R}}(\tilde{\mathcal{H}}_{C,\rho}) + 3\sqrt{\frac{\ln 1/\delta}{2m}}$$

对于要证的不等式左侧, $\Pr_{(x,y) \sim \mathcal{D}}[yH(x) < 0] = \mathbb{E}[\mathbb{I}_{yH(x) < 0}] \leq \mathbb{E}[\Phi_\rho(yH(x))]$.

对于不等式右侧的第一项, $\frac{1}{m} \sum_{i=1}^m \Phi_\rho(yh(x_i)) \leq \frac{1}{m} \sum_{i=1}^m \mathbb{I}_{yh(x_i) \leq \rho} \leq \sum_{i=1}^m \mathbb{I}_{yh(x_i) \leq \rho}$

对于不等式右侧的第二项,

$$\begin{aligned} \hat{\mathfrak{R}}(\tilde{\mathcal{H}}_{C,\rho}) &= \hat{\mathfrak{R}}(\Phi_\rho \circ \tilde{\mathcal{H}}_C) \leq \frac{1}{\rho} \hat{\mathfrak{R}}(\tilde{\mathcal{H}}_C) \\ &= \frac{1}{\rho} \mathbb{E}_\sigma \left[\sup_{h \in \tilde{\mathcal{H}}_C} \frac{1}{m} \sum_{i=1}^m \sigma_i h(x_i) y_i \right] \\ &= \frac{1}{\rho} \mathbb{E}_\sigma \left[\sup_{h \in \mathcal{H}_C} \frac{1}{m} \sum_{i=1}^m \sigma_i h(x_i) \right] \\ &= \frac{1}{\rho} \mathbb{E}_\sigma \left[\sup_{h_1 \in \mathcal{H}, \dots, h_d \in \mathcal{H}} \frac{1}{m} \sum_{i=1}^m \sigma_i \sum_{j=1}^d \alpha_j h_j(x_i) \right] \\ &= \frac{1}{\rho} \mathbb{E}_\sigma \left[\sum_{j=1}^d \alpha_j \sup_{h_j \in \mathcal{H}} \frac{1}{m} \sum_{i=1}^m \sigma_i h_j(x_i) \right] \\ &= \frac{1}{\rho} \sum_{j=1}^d \alpha_j \mathbb{E}_\sigma \left[\sup_{h_j \in \mathcal{H}} \frac{1}{m} \sum_{i=1}^m \sigma_i h_j(x_i) \right] \\ &= \frac{1}{\rho} \sum_{j=1}^d \alpha_j \hat{\mathfrak{R}}(\mathcal{H}) \\ &= \frac{1}{\rho} \hat{\mathfrak{R}}(\mathcal{H}) \end{aligned}$$

将上述三个结合，即有

$$\begin{aligned}
 \Pr_{(x,y) \sim \mathcal{D}}[yH(x) < 0] &\leq \mathbb{E}[\Phi_\rho(yH(x))] \\
 &\leq \frac{1}{m} \sum_{i=1}^m \Phi_\rho(yh(x)) + 2\hat{\mathfrak{R}}(\tilde{\mathcal{H}}_{\mathcal{C},\rho}) + 3\sqrt{\frac{\ln 1/\delta}{2m}} \\
 &\leq \sum_{i=1}^m \mathbb{I}_{yh(x) \leq \rho} + 2\frac{1}{\rho}\hat{\mathfrak{R}}(\mathcal{H}) + 3\sqrt{\frac{\ln 1/\delta}{2m}}
 \end{aligned}$$

□

参考文献

- [1] Anand (<https://stats.stackexchange.com/users/2513/anand>). Understanding proof of mc-diarmid's inequality. Cross Validated. URL:<https://stats.stackexchange.com/q/21513> (version: 2017-04-13).