

How Restricting Migrants' Job Options Affects Both Migrants and Residents

Wilbur Townsend, Harvard University
Corey Allan, Ministry of Business, Innovation and Employment

These results are not official statistics. They have been created for research purposes from the Integrated Data Infrastructure (IDI) which is carefully managed by Stats NZ. For more information about the IDI please visit <https://www.stats.govt.nz/integrated-data/>. The results are based in part on tax data supplied by Inland Revenue to Stats NZ under the Tax Administration Act 1994 for statistical purposes. Any discussion of data limitations or weaknesses is in the context of using the IDI for statistical purposes, and is not related to the data's ability to support Inland Revenue's core operational requirements. Access to the data used in this study was provided by Stats NZ under conditions designed to give effect to the security and confidentiality provisions of the Data and Statistics Act 2022. The results presented in this study are the work of the author, not Stats NZ or individual data suppliers.

Can work visas facilitate immigration while protecting residents' wages?

We focus on one aspect of work visa policy: **Requirements that migrants work in certain jobs.**

- ▶ 'High-skill' occupations.
- ▶ Rural areas or certain cities.
- ▶ The agriculture or medical sectors.
- ▶ Particular firms.

We ask how these requirements affect **migrant workers, resident workers, and their employers.**

Specifically: We study New Zealand's Essential Skills work visa. These migrants were **limited to positions for which no New Zealanders were believed to be available.**

Incidence will depend on how firms set wages:

Bargaining?

- ▶ Firms bargain with individual workers.
- ▶ Weaker outside options decrease migrants' bargaining power and so decrease their wages.
- ▶ **Residents only affected insofar as their marginal product is affected.**

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- ▶ Firms commit to position-specific wages.
- ▶ A firm's wage choice will depend on the distribution of outside options.
- ▶ **Restricting migrants' job options can decrease wages for residents.**

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A fundamental topic in labor economics with implications for the **causes of wage inequality**, the **effects of employment regulations** and the **allocative efficiency** of the labor market (Flinn & Mullins '21; Silbert & Townsend '24).

Today's talk

Reduced-form facts:

- ▶ Migrants earn less than other workers — but only because of the jobs they hold.
- ▶ Loosening an *individual* migrant's job restrictions does not affect their wages.
- ▶ But loosening the restrictions on *all* migrants within an occupation does increase wages.

These facts are consistent with a model in which wages are set through monopsonistic wage-posting.

We estimate such a model:

- ▶ Restricting migrants' job options decreases their wages by an average of 8%.
- ▶ Most residents are unaffected — but 2.1% have their wage decreased by $\geq 2\%$.
- ▶ The restrictions decrease aggregate welfare — mostly because of non-wage amenities.

Existing literature: work visas

- ▶ Individual workers' visa status matters little (Wang '21),
- ▶ While the overall migrant visa system matters a lot (Naidu, Nyarko & Wang '16; Ahrens, et. al '23; Borjas & Edo '23).

We reconcile these results by **combining analyses of individual- and market-level shocks.**

We also extend this literature by asking how **visa systems affect residents' wages** (Amior & Manning '20; Amior & Stuhler '23; Borjas & Edo '23).

Existing literature: theories of wage-setting

- ▶ How workers' outside job options affect their wages (Caldwell & Harmon '19; Caldwell & Danieli '22).
- ▶ Other predictions (Carvalho, da Fonseca & Santarossa '23; Bloesch, Larsen & Taska '22; Hall & Krueger '12; Roussille & Scuderi 24; Rubens & Delabastita '24).

We focus on a novel prediction of wage-posting models:

- ▶ *Among workers at the same firm*, an individual's *own* outside option has no effect on their wage,
- ▶ But the wage chosen by that firm will depend on the *distribution* of workers' outside options.

Existing literature: structural estimation of labor market power

A large literature estimates labor market power:

- ▶ By estimating labor supply elasticities (Manning '03; Lamadon, Mogstad & Setzler '21; Azar, Berry & Marinescu '22; Berger, Herkenhoff & Mongey '22; Kroft, et al. '23; Roussille & Scuderi '23, Sharma '22, ...),
- ▶ By estimating production functions (Lamadon, Mogstad & Setzler '21; Yeh, Macaluso & Hershbein '22; Kroft, et al. '23; Delabastita & Rubens '22).

We estimate both a labor supply system, and firm-specific production functions **in which different occupations are distinct inputs**:

- ▶ Calculate firm-by-occupation marginal products by inverting a markdown equation (Berry, Levinsohn & Pakes '95; Chan et al. 2024).
- ▶ We estimate **firms' ability to substitute across occupations**.

The Essential Skills Work Visa

- ▶ Existed between 2008 and 2022.
- ▶ Accounted for 80% of employer-assisted visas, and about 25% of all work visas.
- ▶ Covers a broad range of skilled and semi-skilled occupations. Most common: chef, carpenter, dairy cattle farmer.

The Essential Skills visa was limited to positions for which there were ‘no New Zealanders available’.

- ▶ For most occupations: must show a ‘genuine attempt’ to recruit New Zealanders
- ▶ ... but **occupations on an “Essential Skills in Demand list” are exempt.**

Data and sample

We use linked **census, tax and immigration** data:

- ▶ **Census data:** Employment status, hours, occupation, education and demographics, in 2013 and 2018.
- ▶ **Tax data:** Worker-month-firm panel of earnings & location.
- ▶ **Immigration data:** Spell data on visas held, border movements, and occupations. Entries into resident visa lotteries.

Between **August 2008 to March 2022**: 278m observations across **3.8m workers**.

Among **205k Essential Skills workers**:

- ▶ 32% live in Auckland (vs. 27% overall). 54% urban (vs. 51% overall).
- ▶ Average age 35. 71% are men.
- ▶ 37% have at least a Bachelor's degree (vs. 32% overall).

Today's talk

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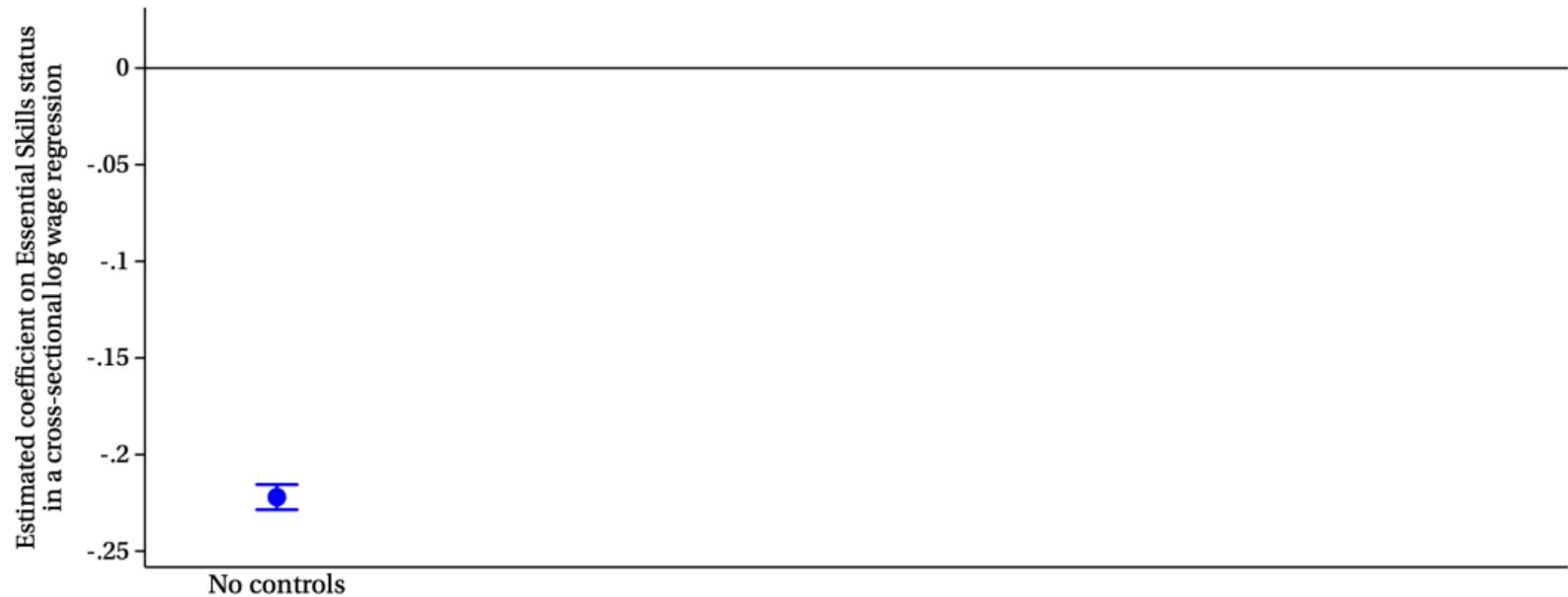
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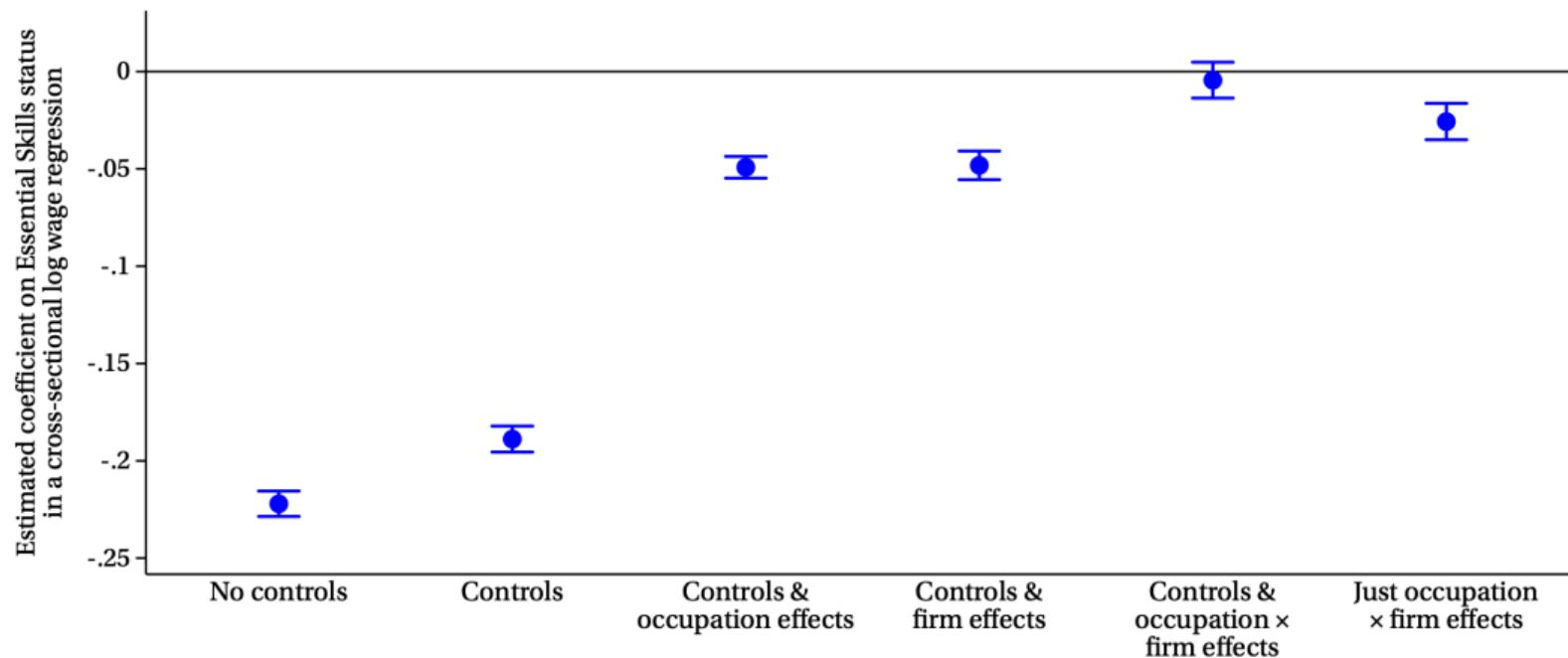
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Migrants earn less than other workers...



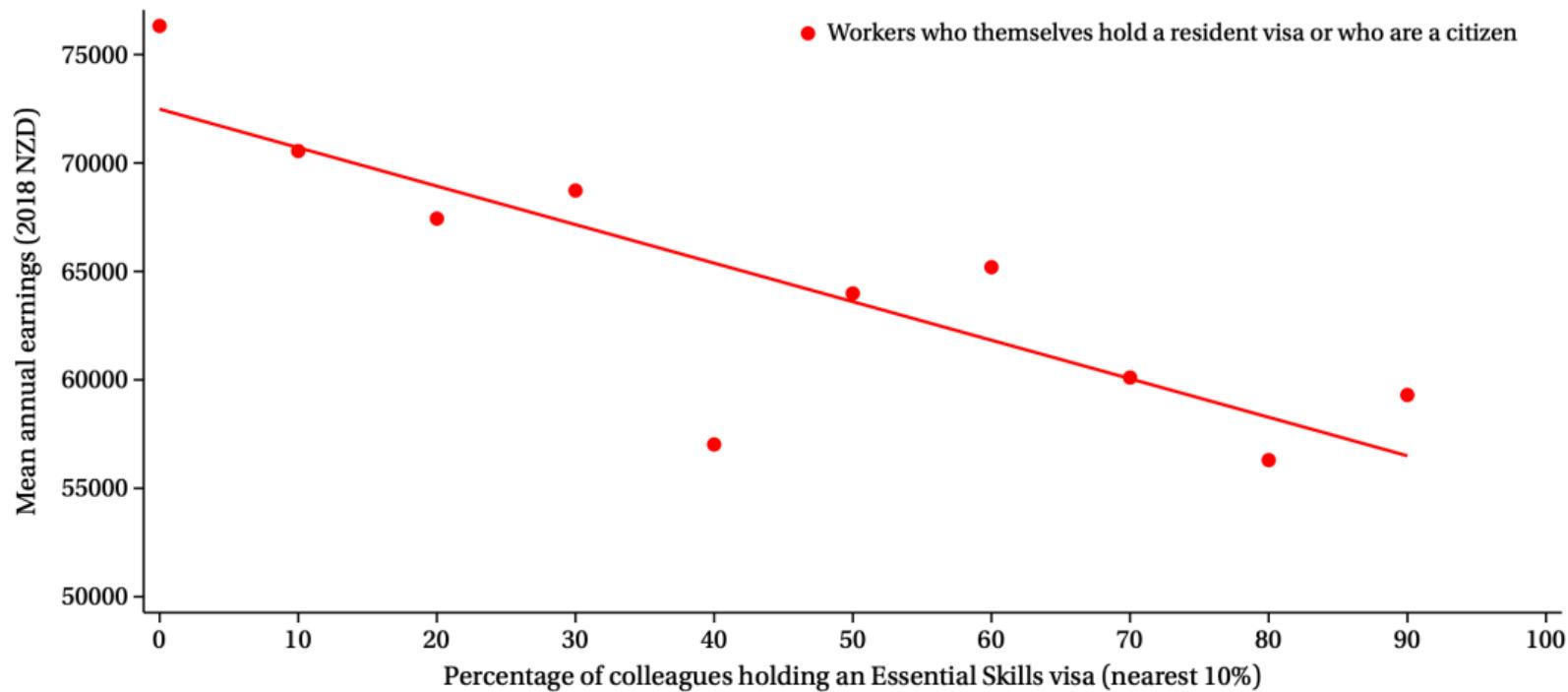
Sample: full-time employees with a unique firm in the 12 months before the 2018 Census. [Other samples](#).

Migrants earn less than other workers... but only because of their jobs



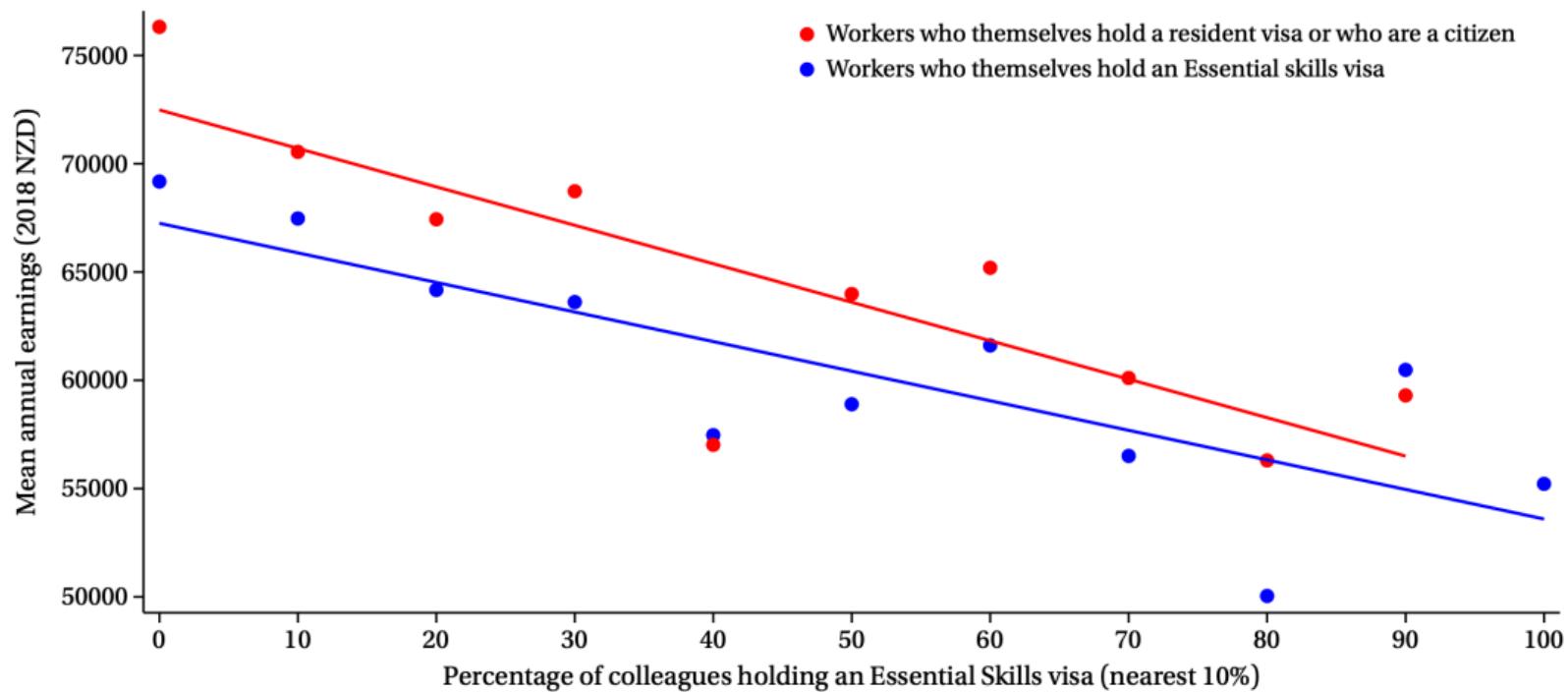
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Sample: full-time employees with a unique firm in the 12 months before the 2013 or 2018 censuses.

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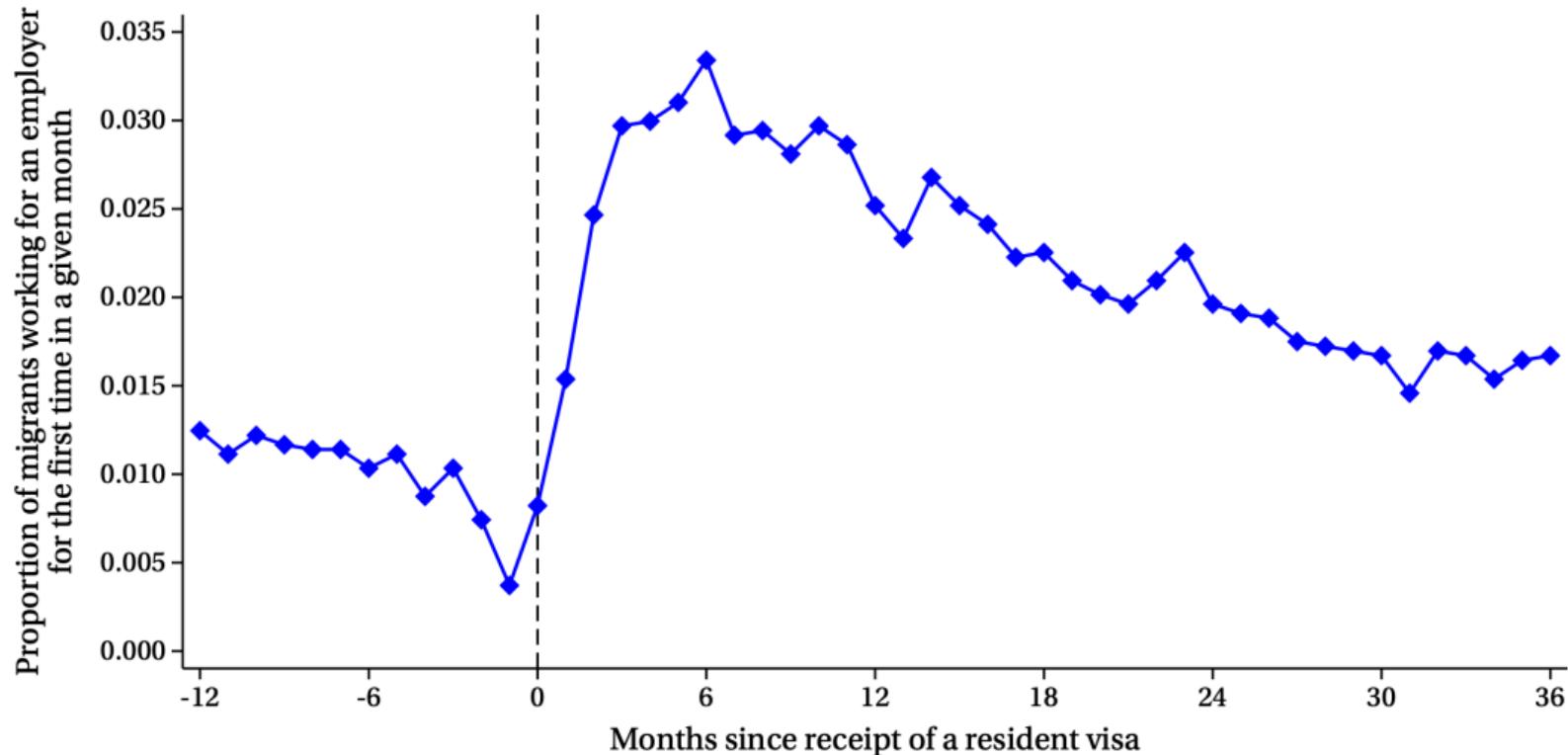
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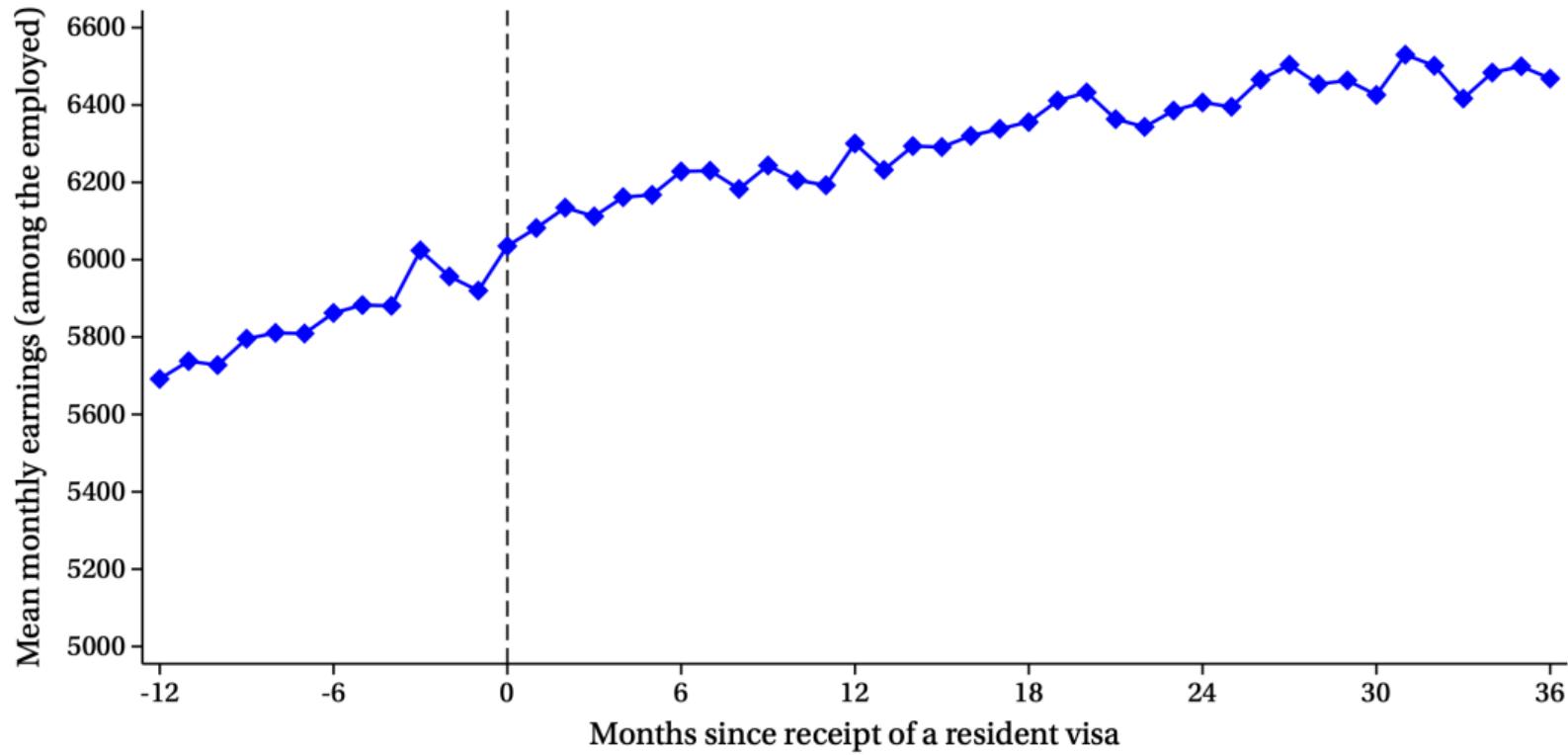
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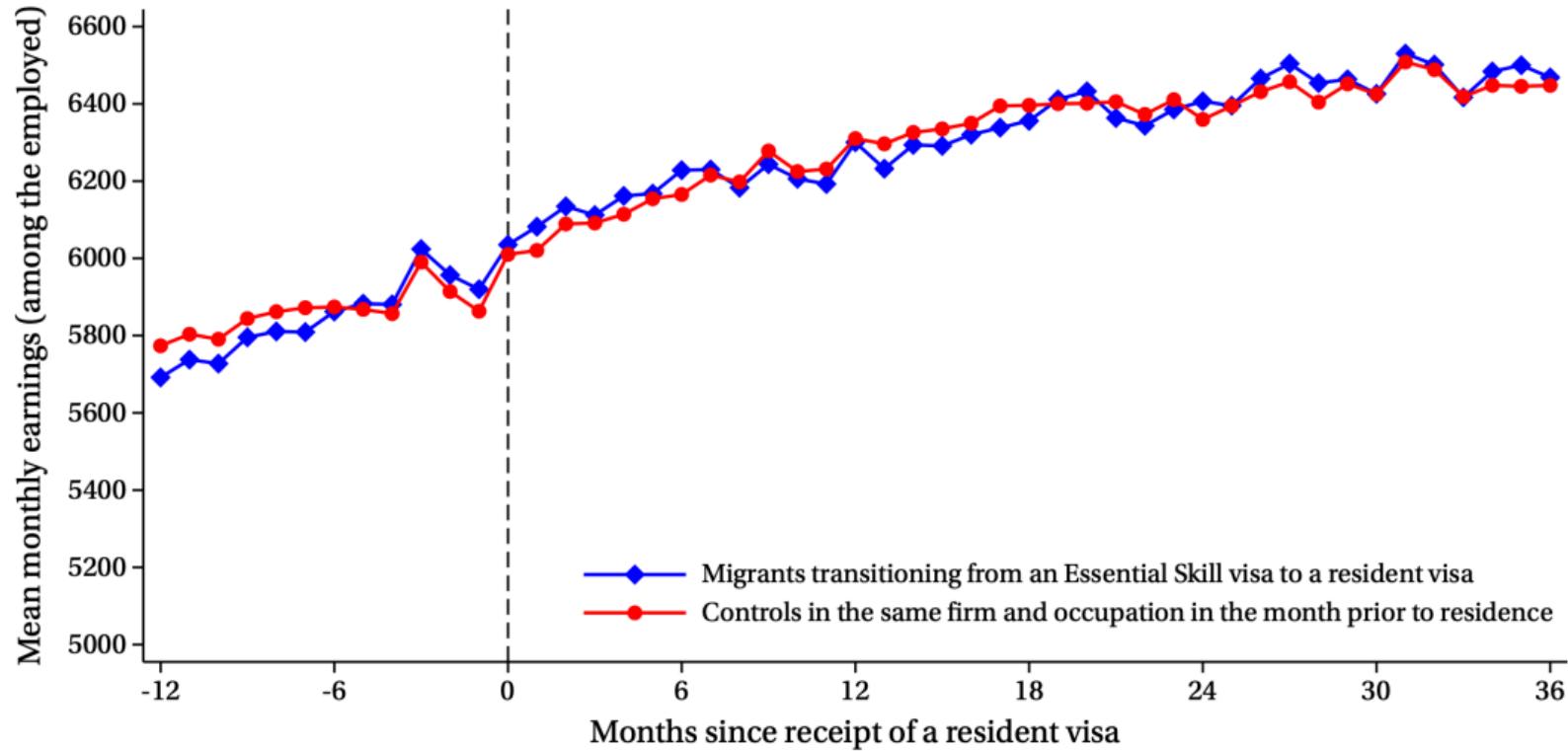
The Essential Skills visa restricts migrants' job options...



...but has no effect on an individual migrant's earnings

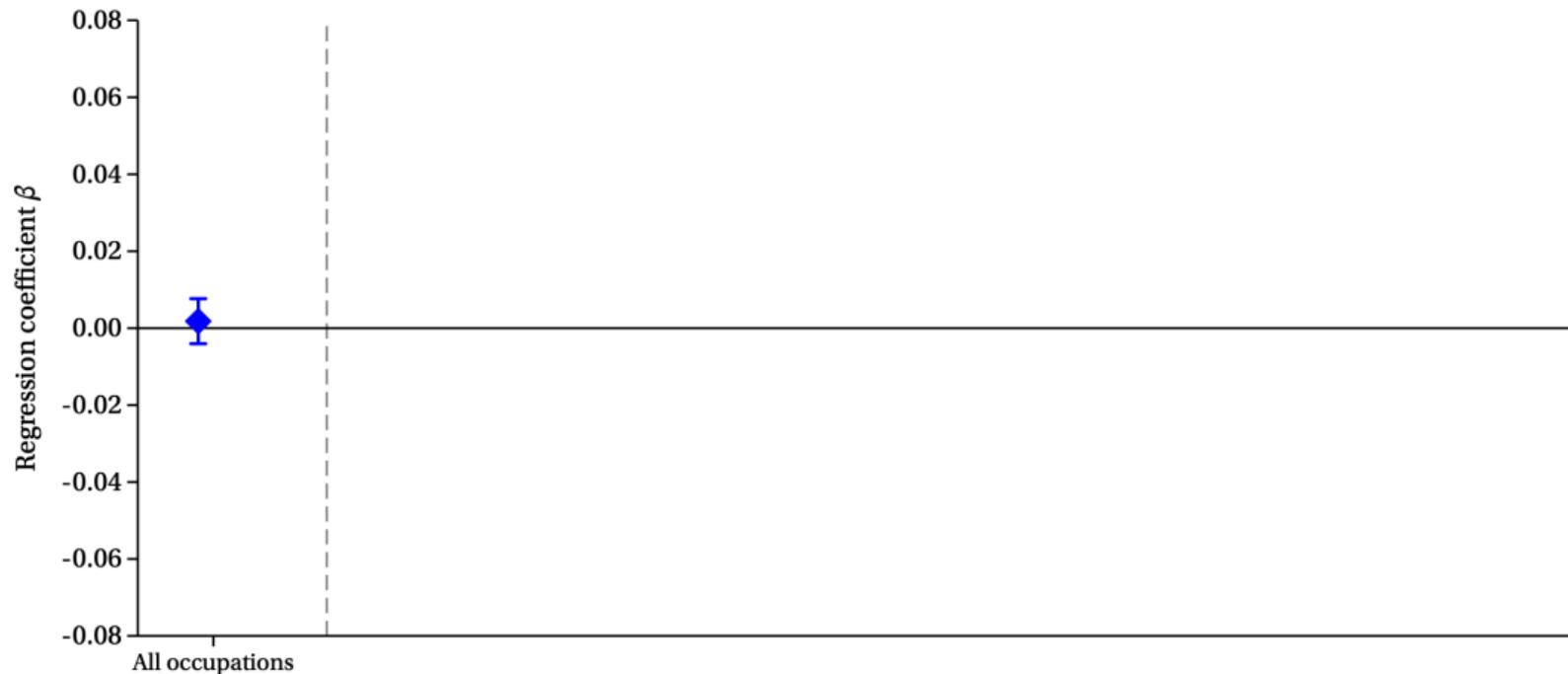


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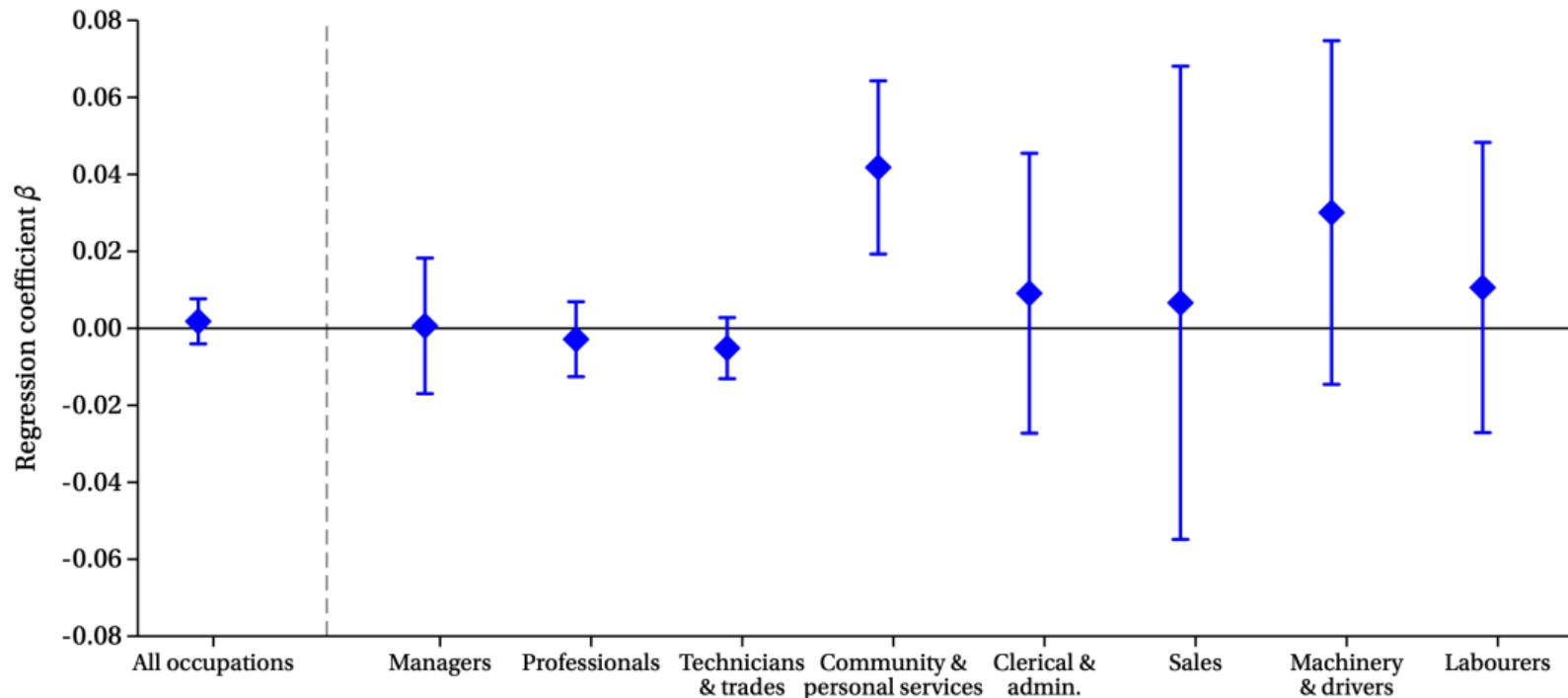
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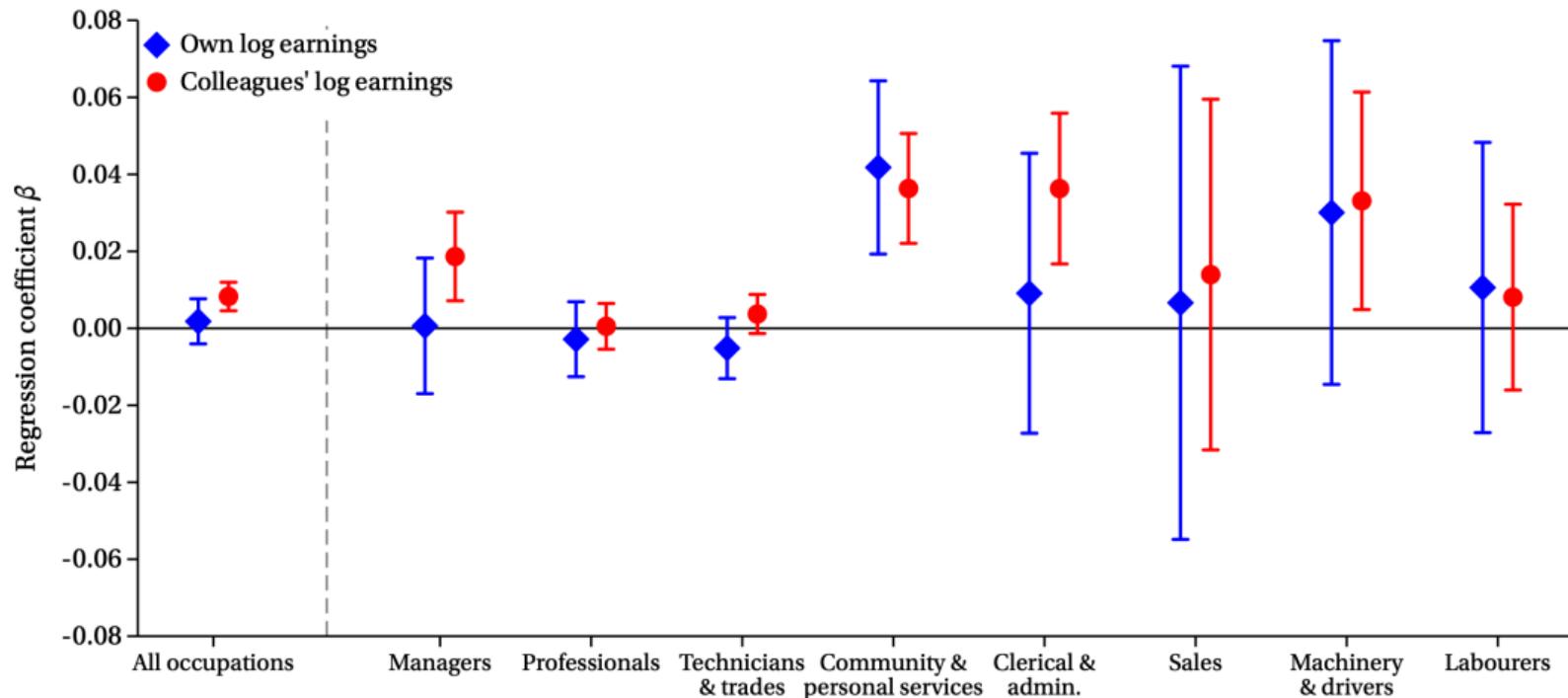
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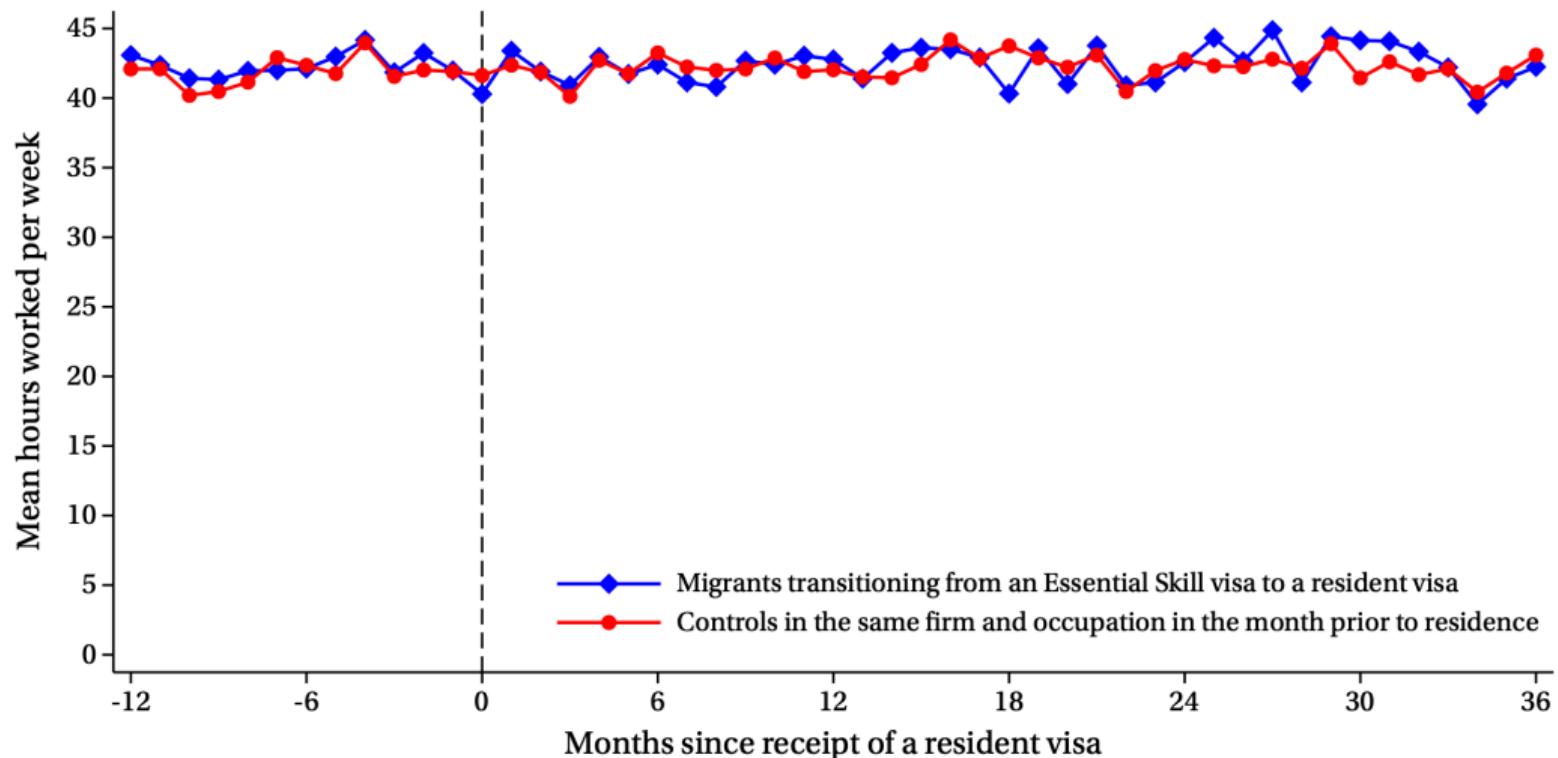


An individual migrant's job restriction does not affect their wage

Concern 1: New residents are able to negotiate better working conditions.

- Solution: *Study self-reported hours of work* in labor force surveys.

Obtaining a resident visa has no effect on hours of work



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Concern 2: Workers apply for residence in anticipation of **worker-specific** shocks.

- Solution: *Study entrants into visa lotteries.*

The Pacific lottery visas

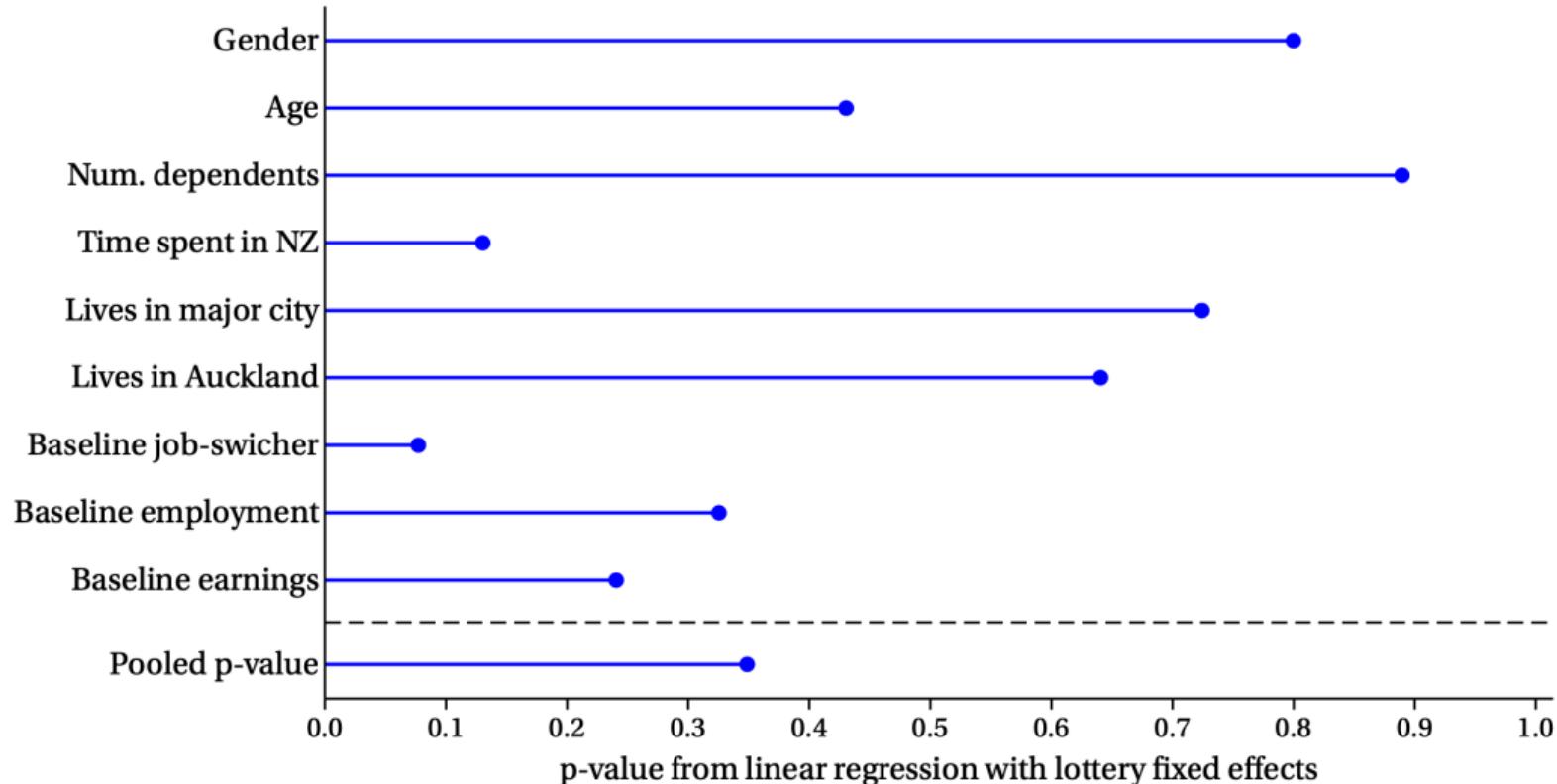
The **Pacific Access Category Resident Visa** and the **Samoan Quota Resident Visa** are New Zealand resident visas allocated by a **random lottery**.

- ▶ Citizens of Samoa, Kiribati, Tuvalu, Tonga and Fiji are eligible to enter.
- ▶ Typical entrants reside in their home country, but roughly 8% are in New Zealand when the lottery is drawn.

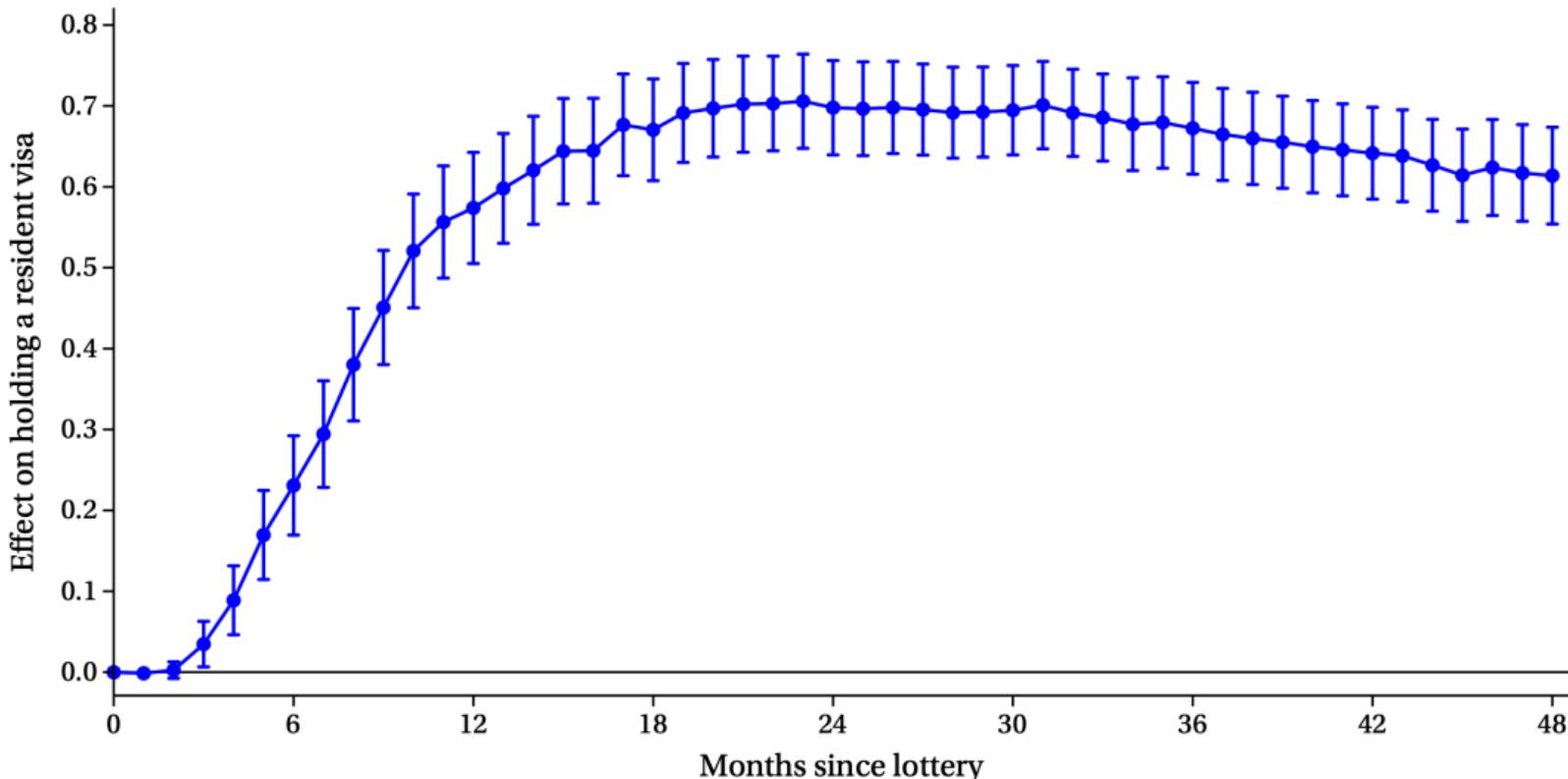
We study those 6501 lottery entries from 3891 individuals who

- ▶ **Entered the lottery for a Pacific lottery visa,**
- ▶ **While working in New Zealand on an Essential Skills visa.**

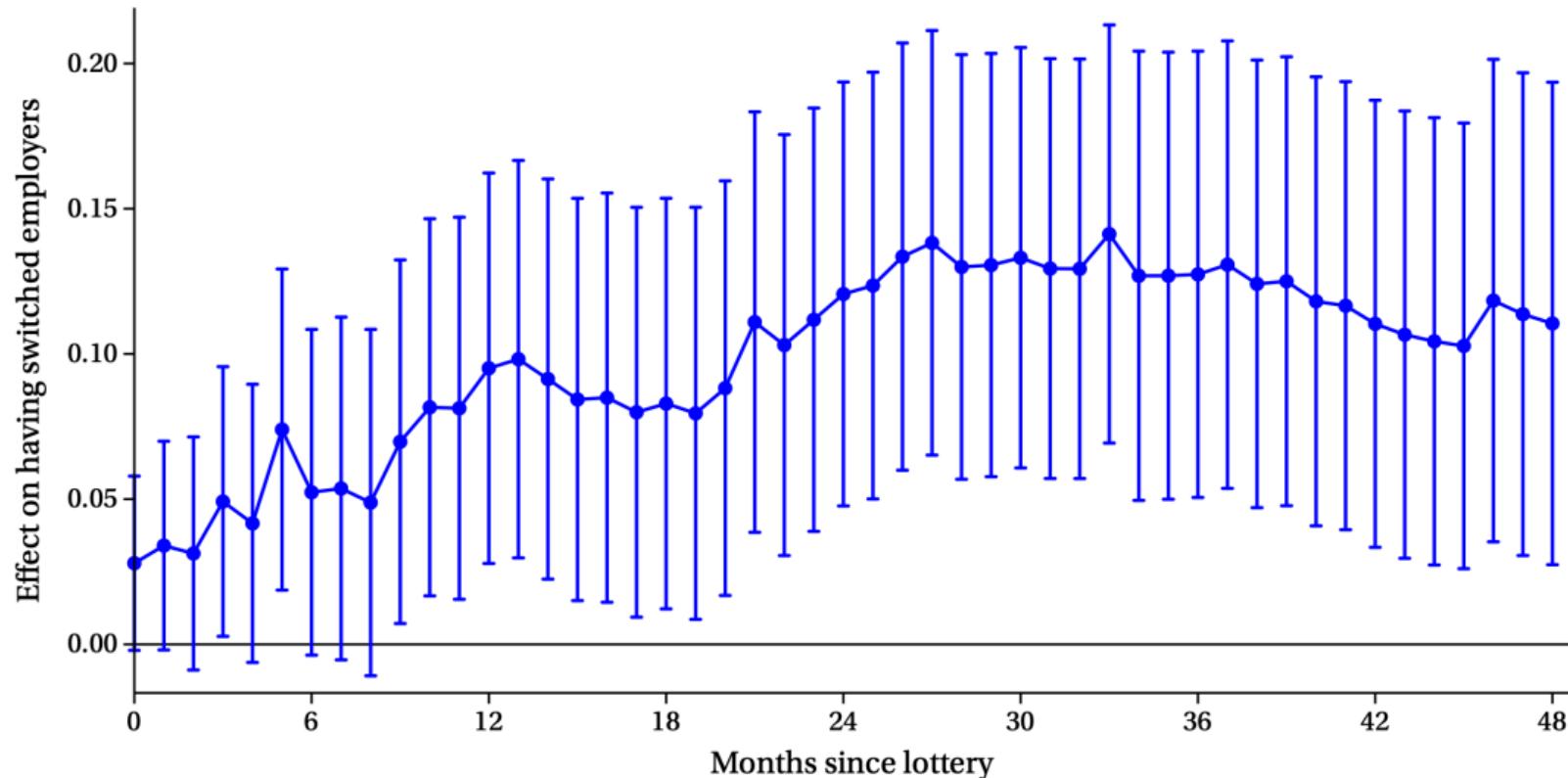
Lottery winners are similar to lottery losers



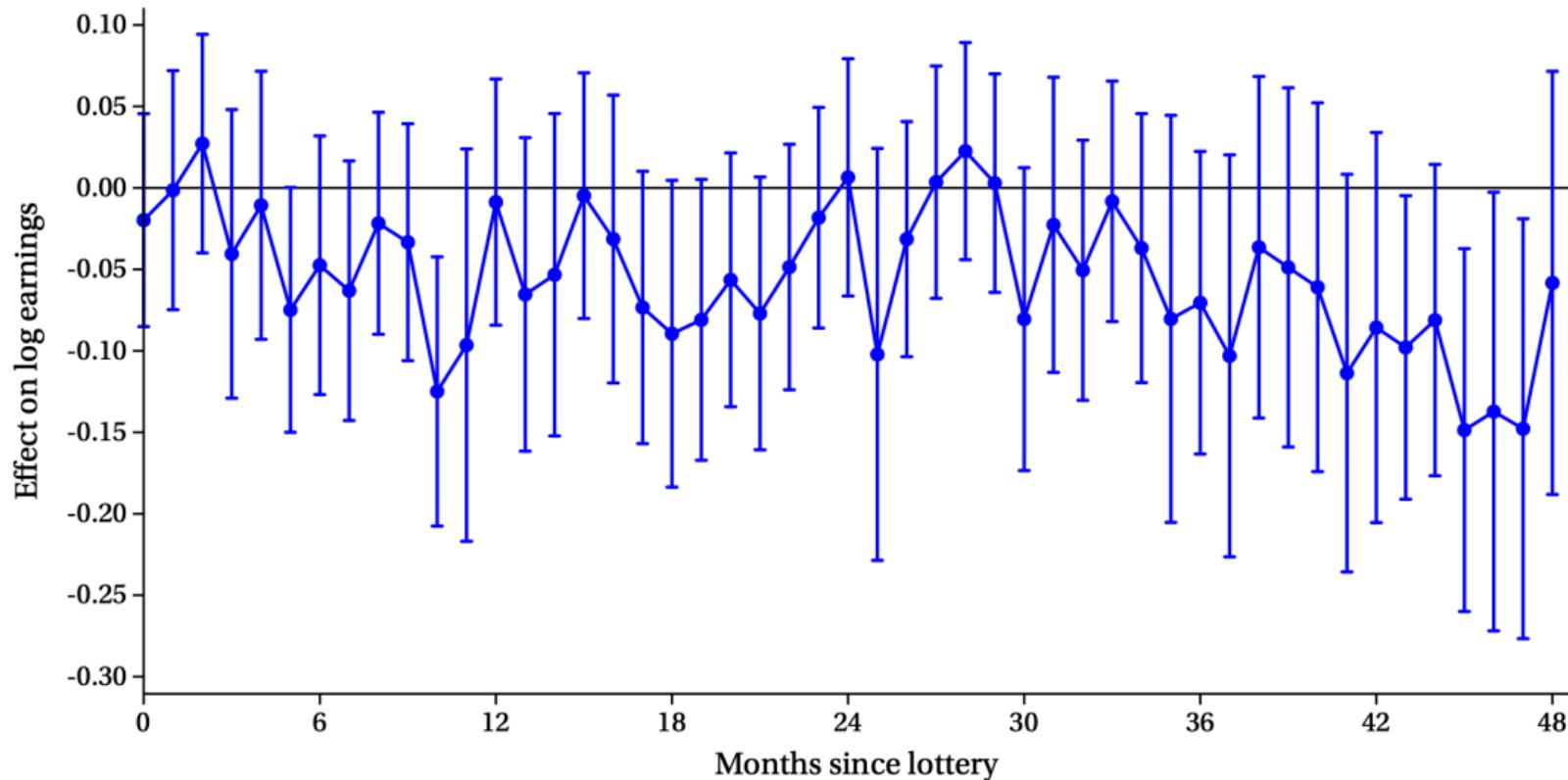
Lottery winners are much more likely to obtain resident visas



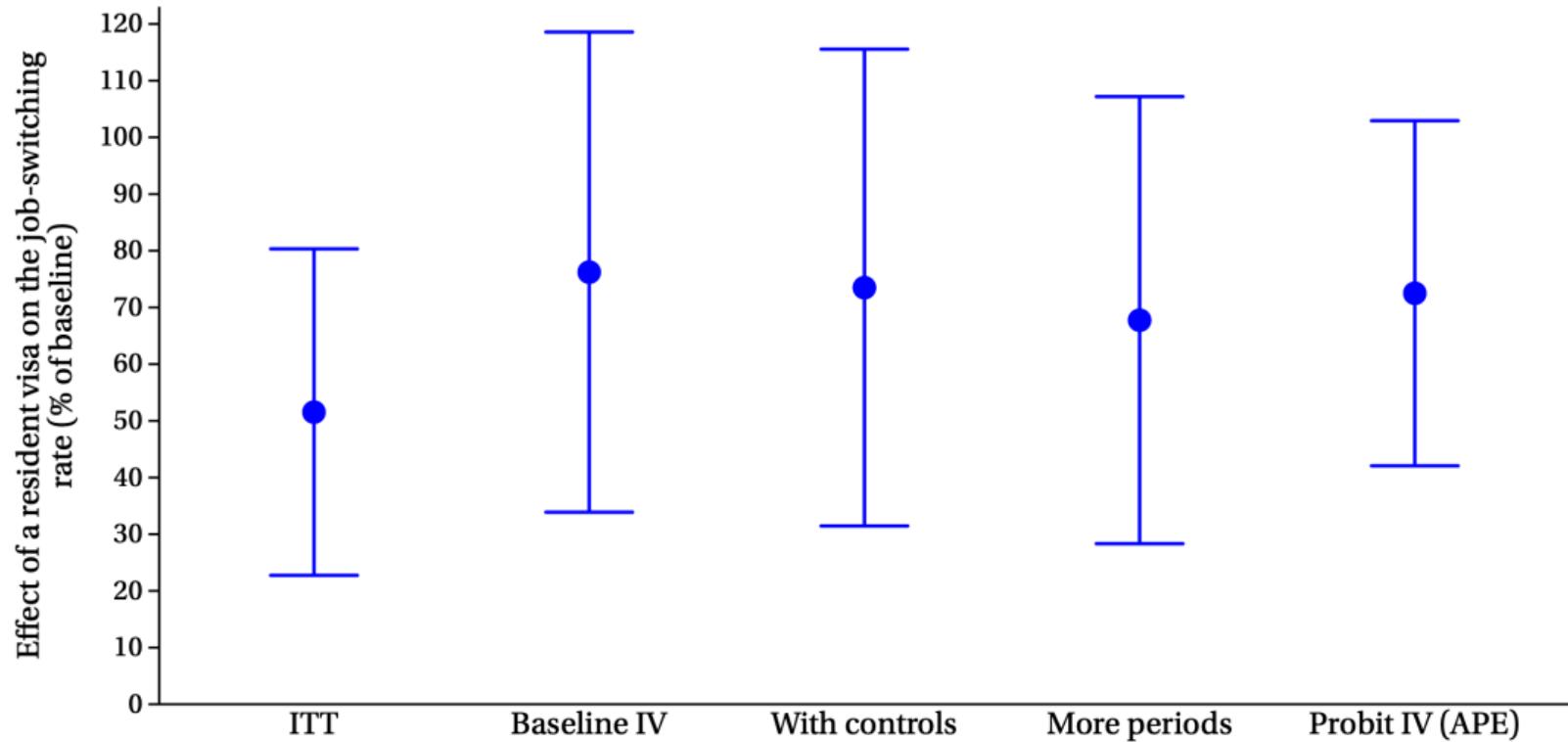
Lottery winners are more likely to switch jobs...



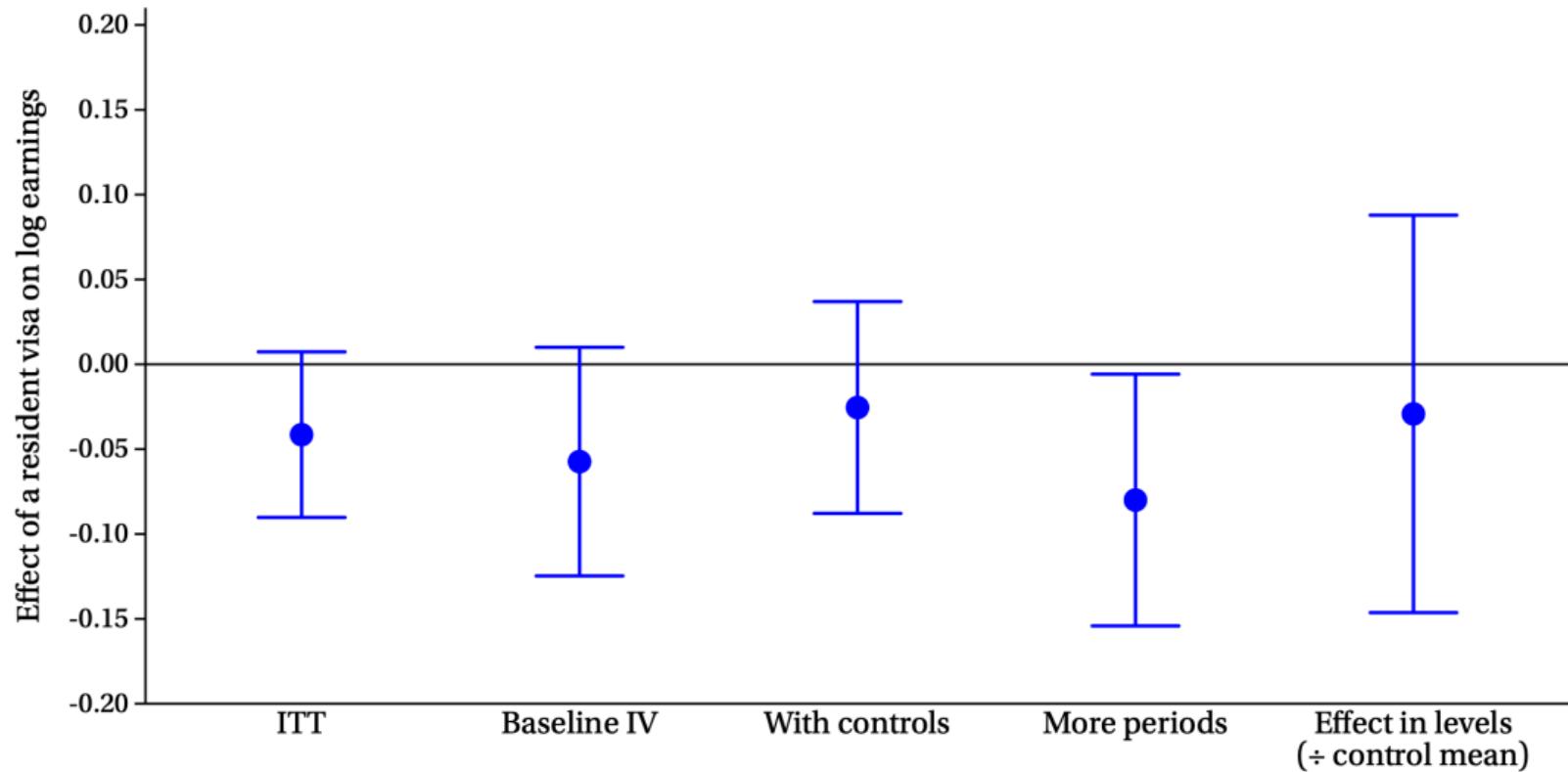
... but earn no more than lottery losers



IV estimates of the effect of a resident visa



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- ▶ Some new residents may move to **high-amenity but low-wage jobs**

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- ▶ Solution: Identify the *average treatment effect among job-stayers.*

Do better outside options matter if a worker stays at their original job?

Let $Y_{i,t}(W, M)$ denote i 's period- t log earnings, which depends on whether they won the lottery W_i and whether they moved $M_{i,t} = M_{i,t}(W_i)$:

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We are interested in the **average effect of winning a visa lottery, among those who stay at their initial job:**

$$\text{ATE}_{\text{stayers}} \equiv \mathbb{E}\left[Y_{i,t}(1, 0) - Y_{i,t}(0, 0) \mid M_{i,t}(1) = M_{i,t}(0) = 0\right].$$

Problem: Among job-stayers, the lottery is non-random.

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\therefore We can identify $\mathbb{E}\left[Y_{i,t}(1, 0) \mid M_{i,t}(1) = M_{i,t}(0) = 0\right] = \mathbb{E}\left[Y_{i,t} \mid M_{i,t} = 0; W_i = 1\right]$.

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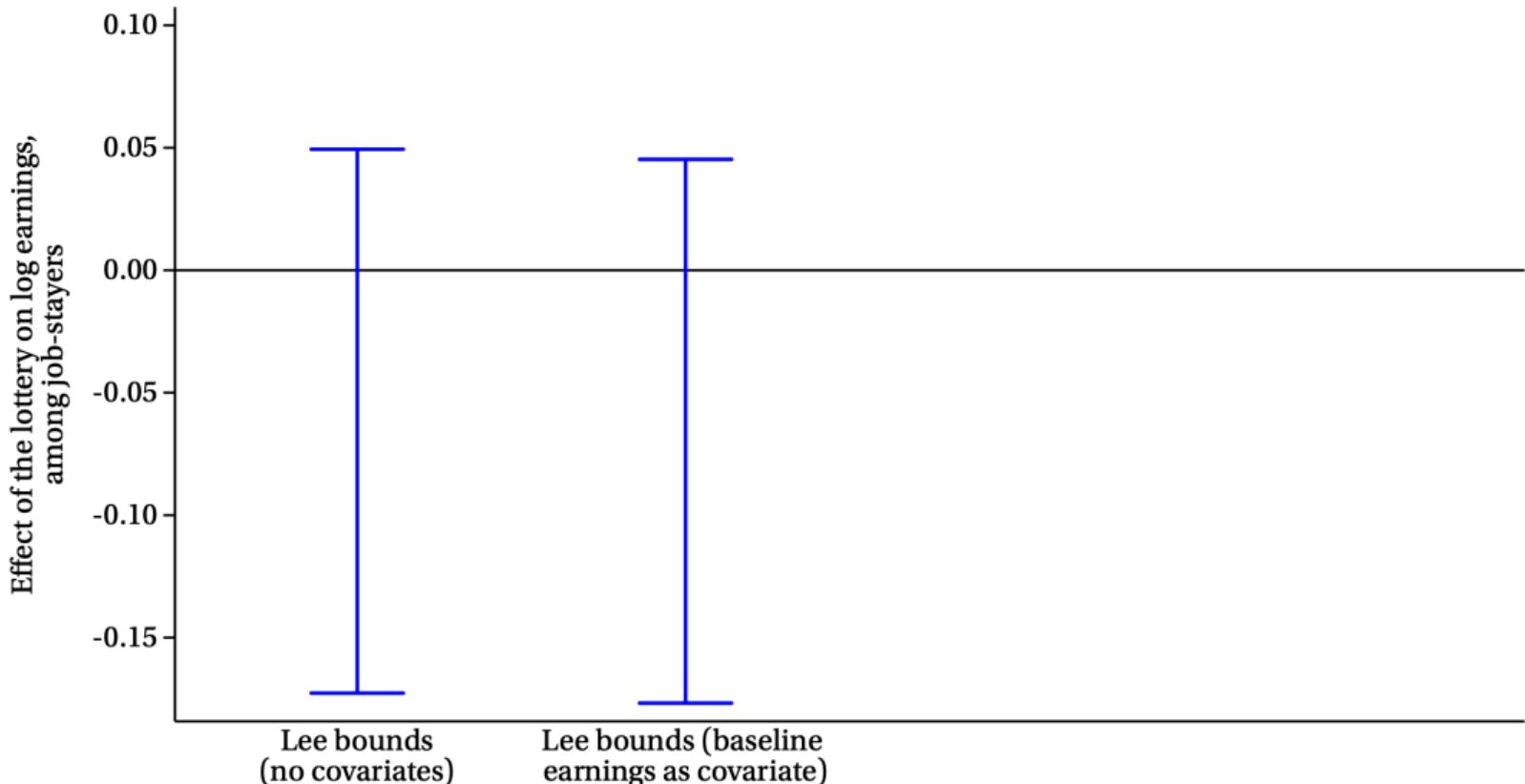
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Approach 1: Set-identify by dropping the highest or lowest $Y_{i,t}$ among lottery losers (Lee '09).

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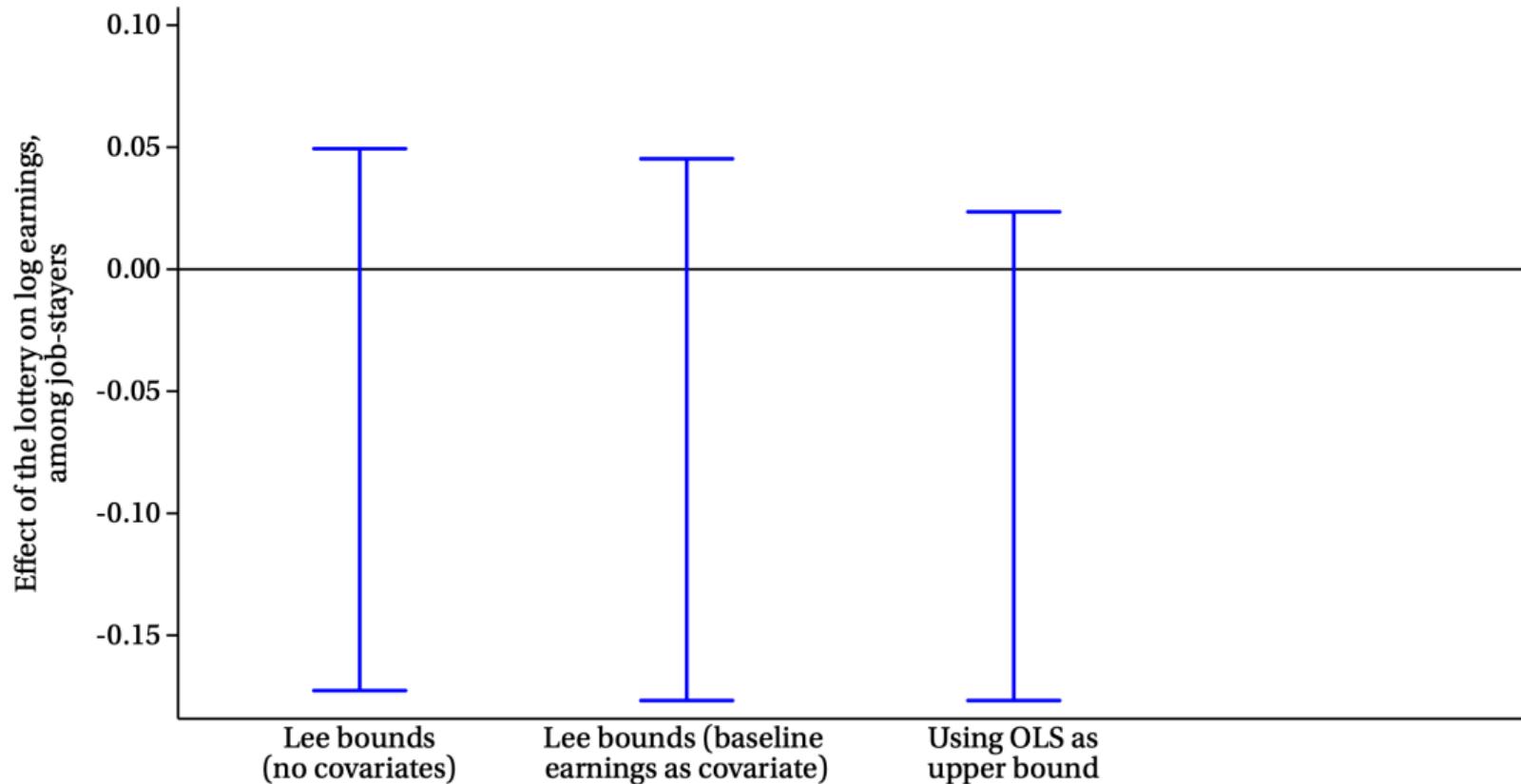
Approach 2: Sharpen the upper bound under the assumption that

$$\mathbb{E} \left[Y_{i,t}(0, 0) \middle| M_{i,t}(0) = 0 \right] \leq \mathbb{E} \left[Y_{i,t}(0, 0) \middle| M_{i,t}(1) = 0 \right],$$

which implies that OLS is upwards-biased.

$$\left(\text{We observe } \mathbb{E} \left[Y_{i,-1} \middle| M_{i,t}(0) = 0 \right] < \mathbb{E} \left[Y_{i,-1} \middle| M_{i,t}(1) = 0 \right]. \right)$$

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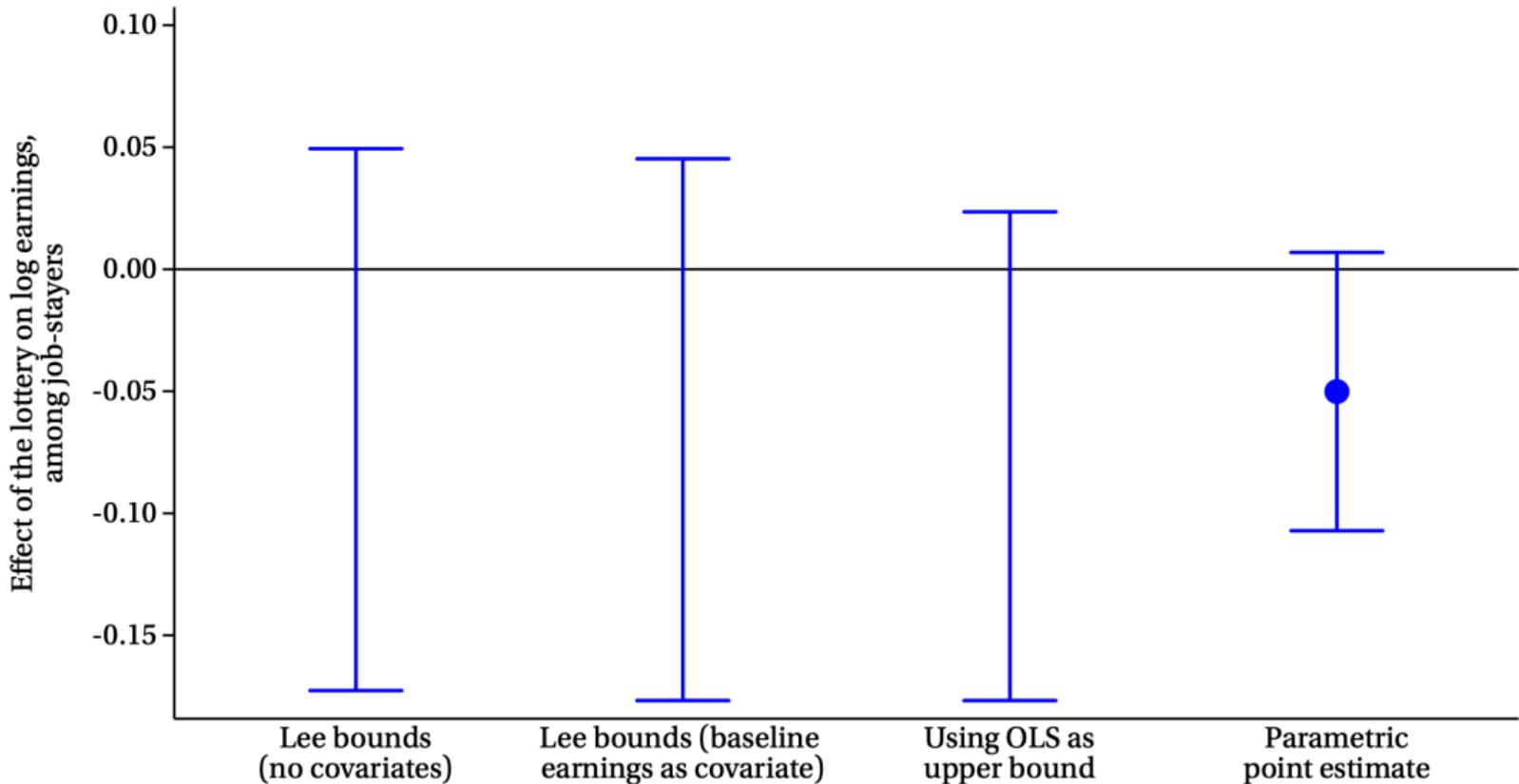
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Approach 3: Point-identify by modeling the decision to move.

Key assumption: job-switching is Markovian in counterfactual earnings

$$M_{i,t}(1), M_{i,t}(0) \perp \left(Y_{i,s}(0, 0) \right)_{s < t} \mid Y_{i,t}(0, 0).$$

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In an occupation, strengthening *all* migrants' job options increases wages

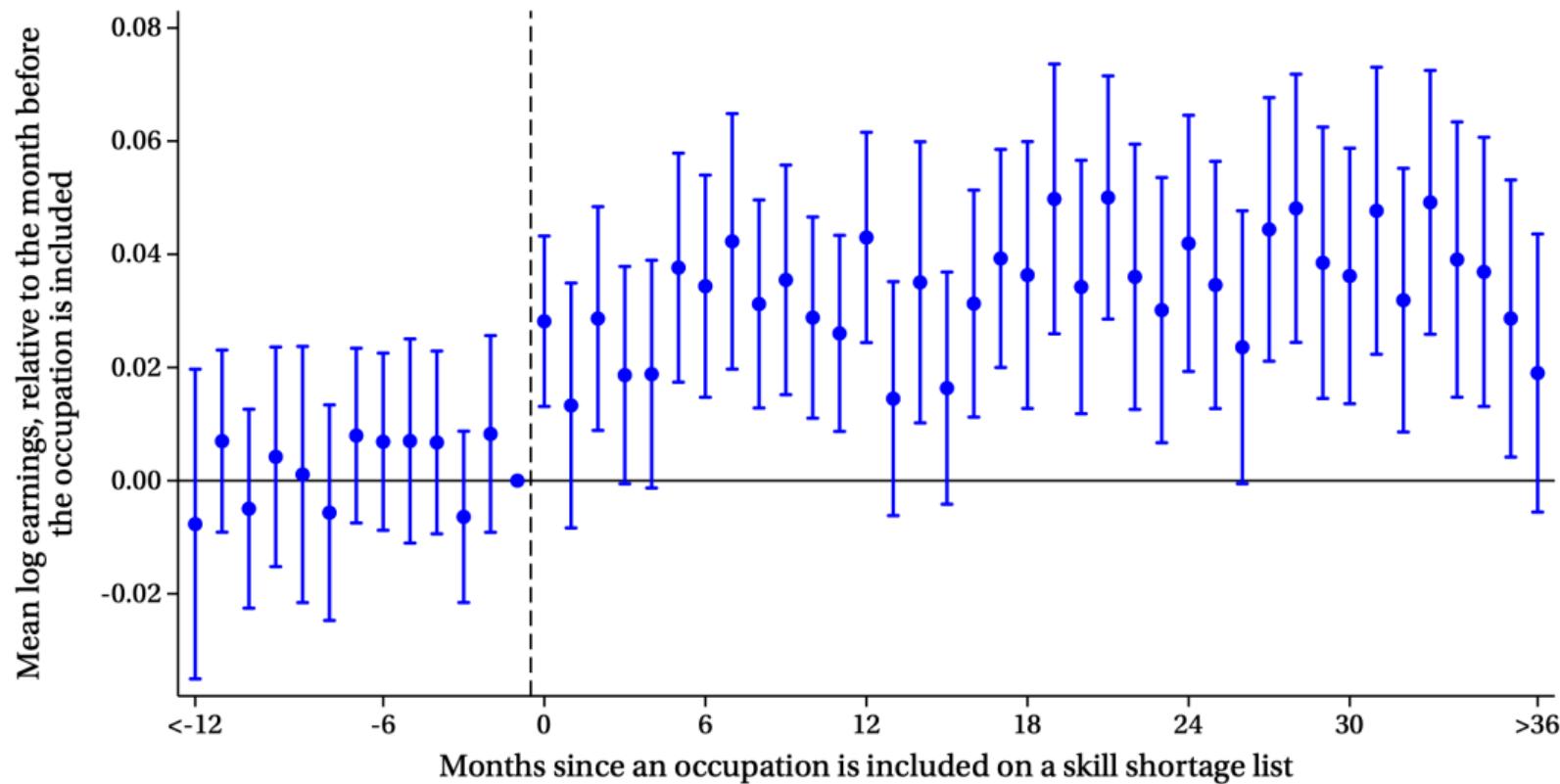
For **occupations on an 'Essential Skills in Demand' list:**

- ▶ Prospective employers need not show a 'genuine attempt' to recruit New Zealanders.
- ▶ So, **Essential Skills migrants' job options are unrestricted.**

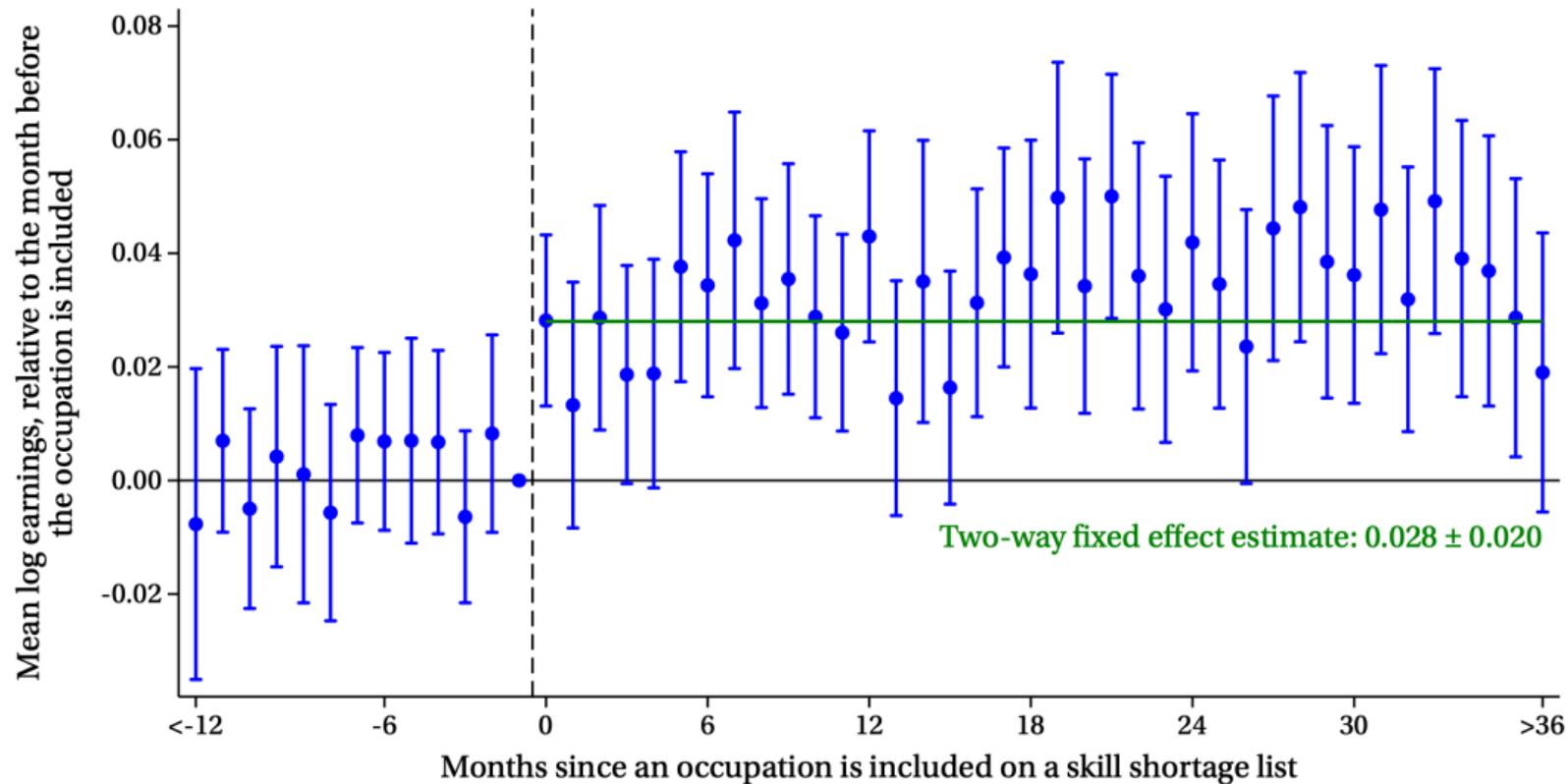
Occupations were added and removed from these lists regularly, often for reasons orthogonal to demand (e.g. occupation size).

We ask how wages change when an occupation is listed.

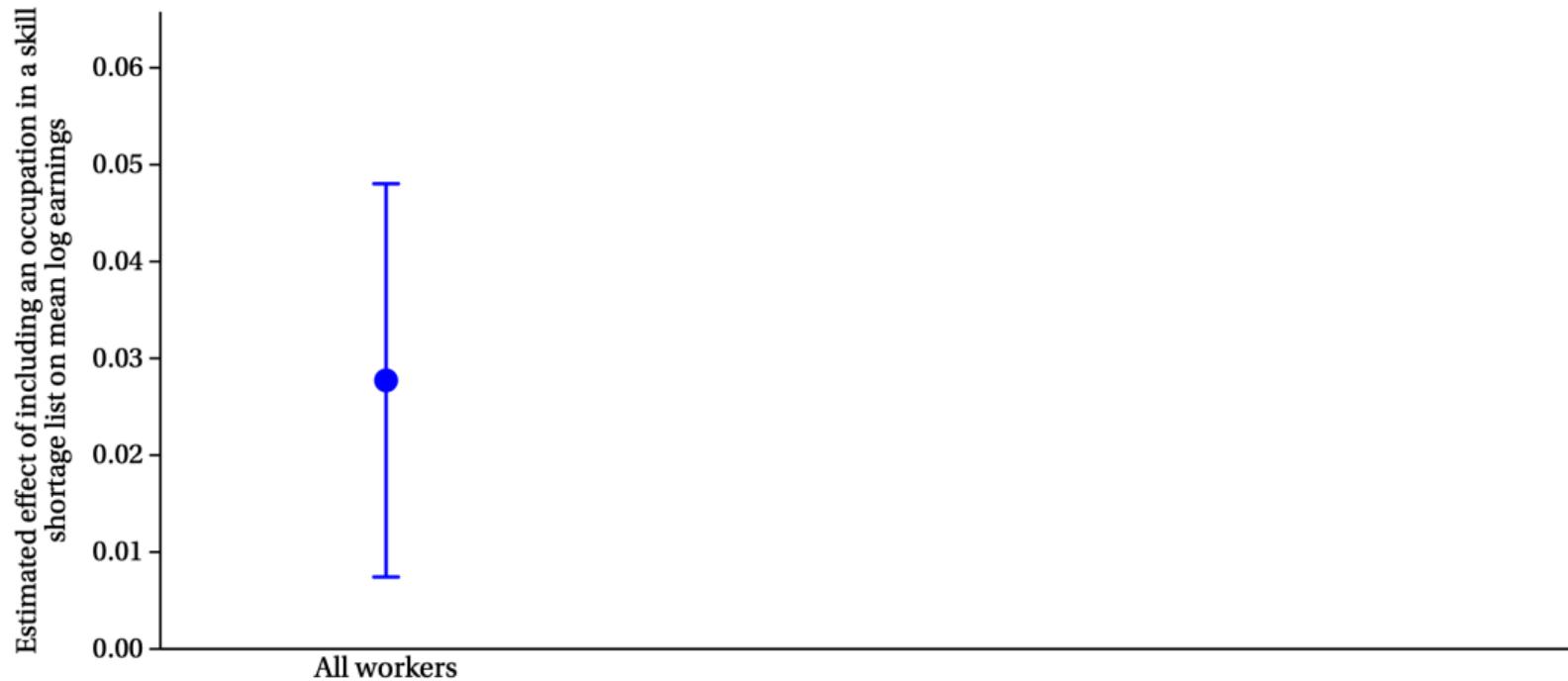
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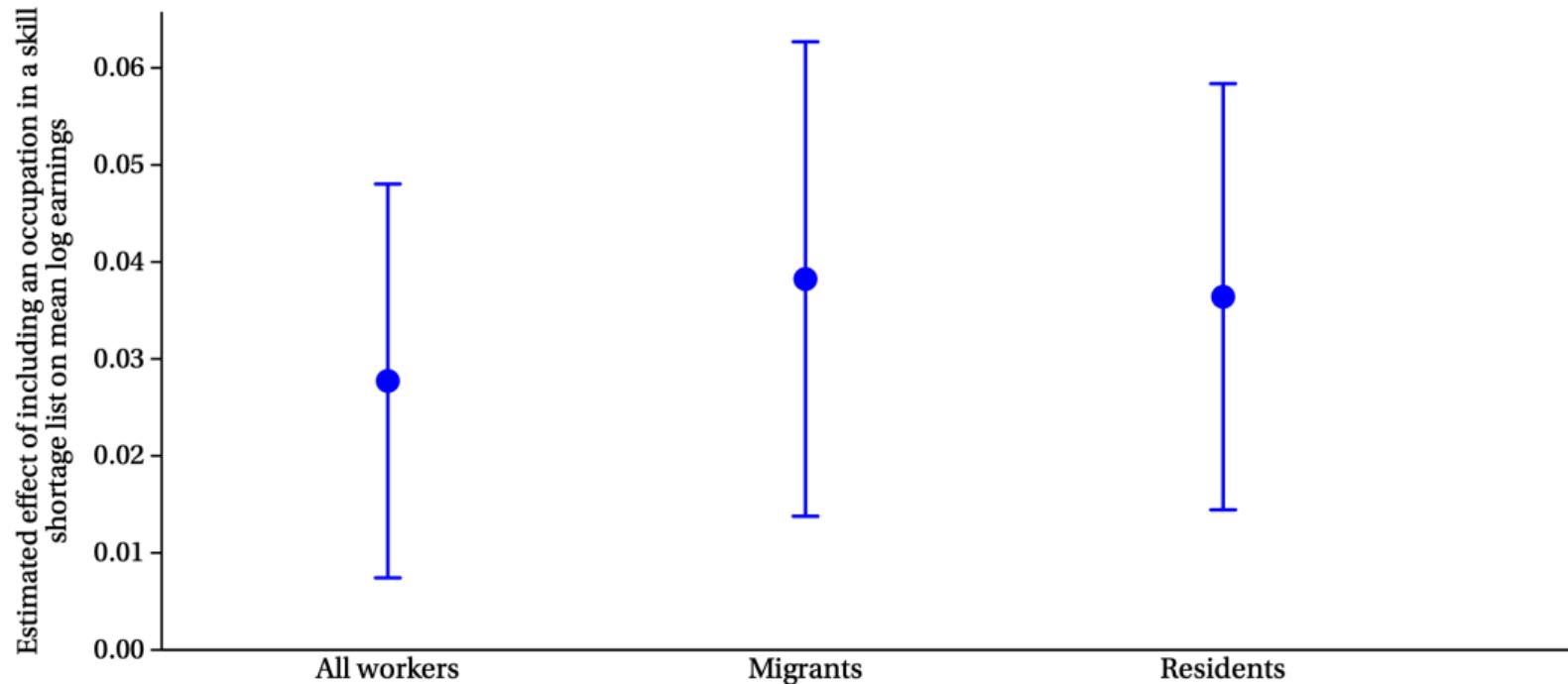
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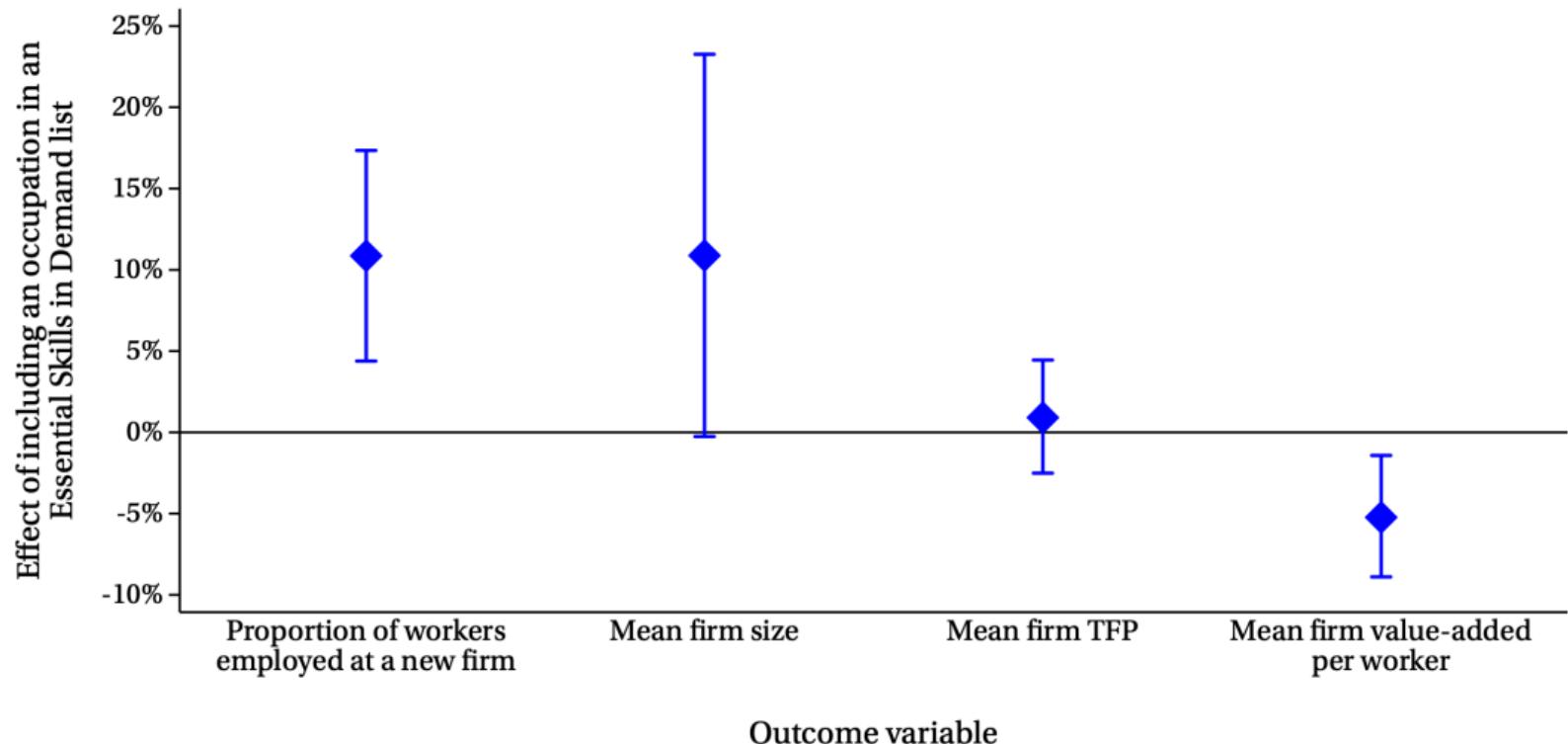
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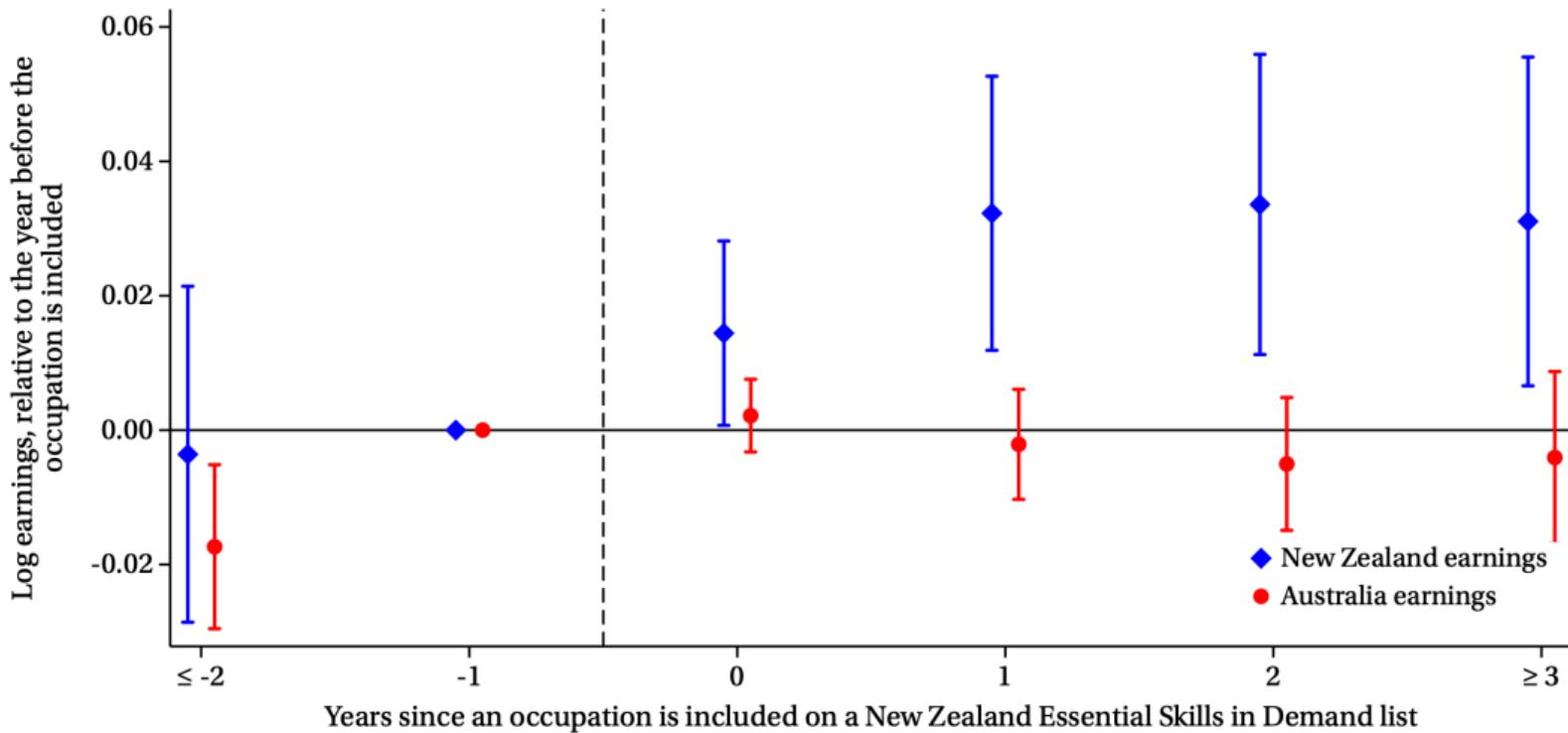
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Why does strengthening migrants' job options increase wages?



Using Australian data as a placebo outcome



Point estimate.

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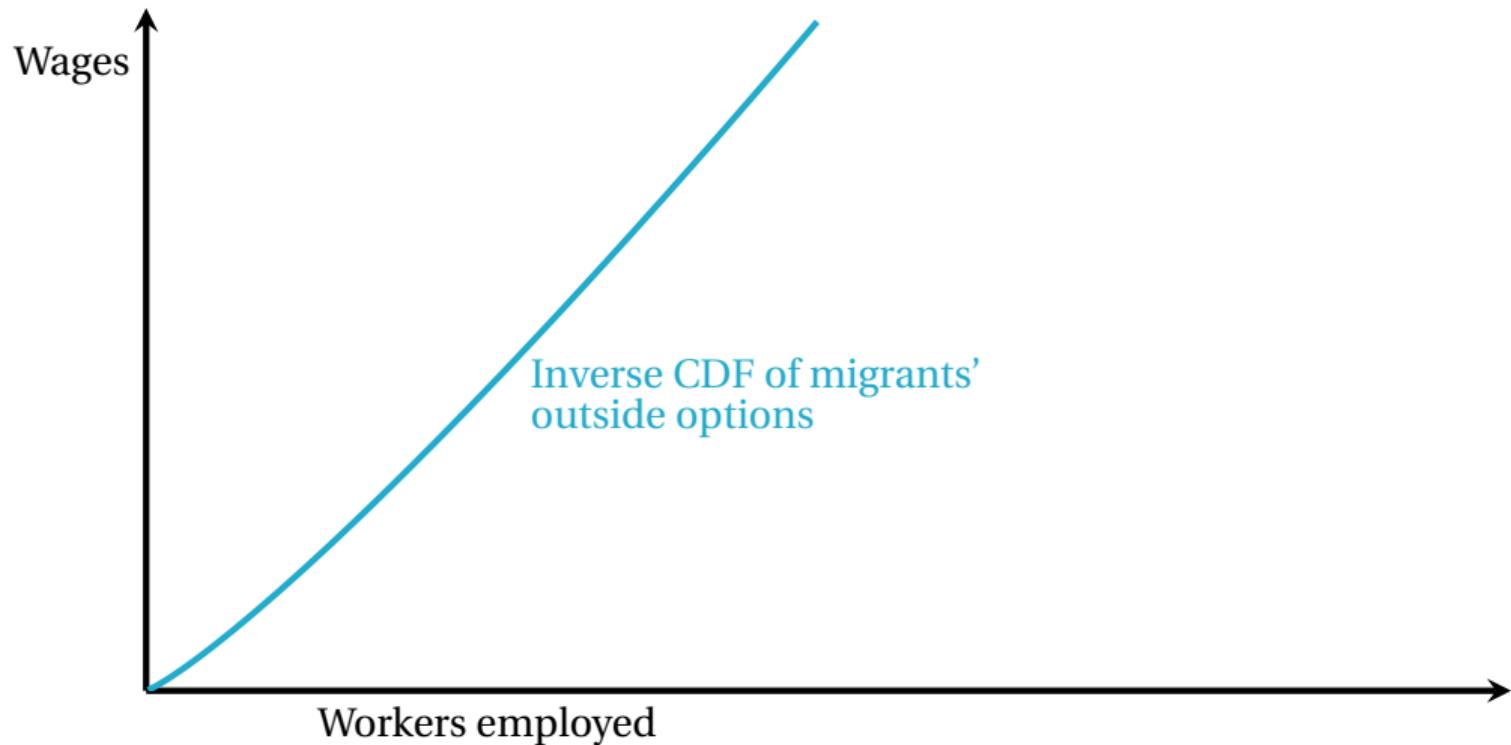
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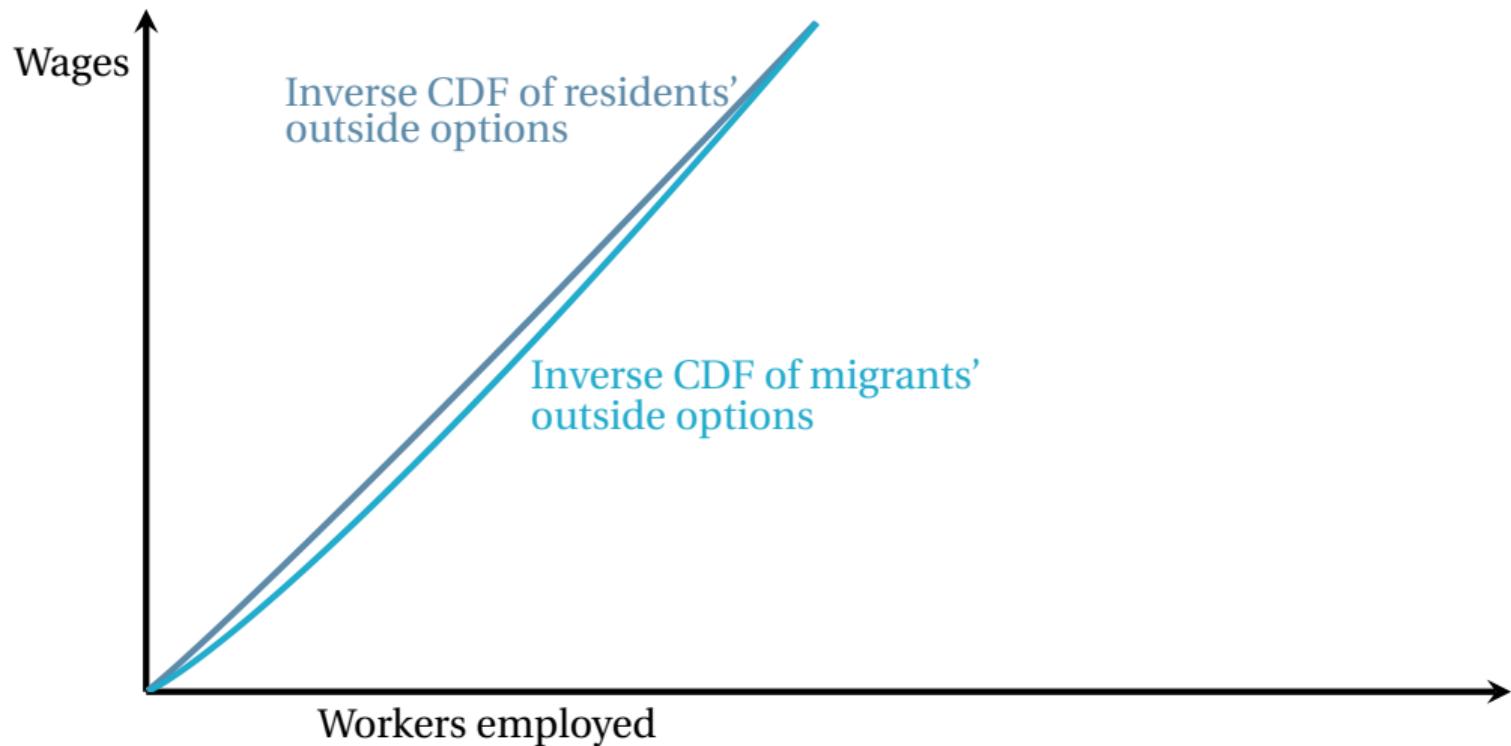
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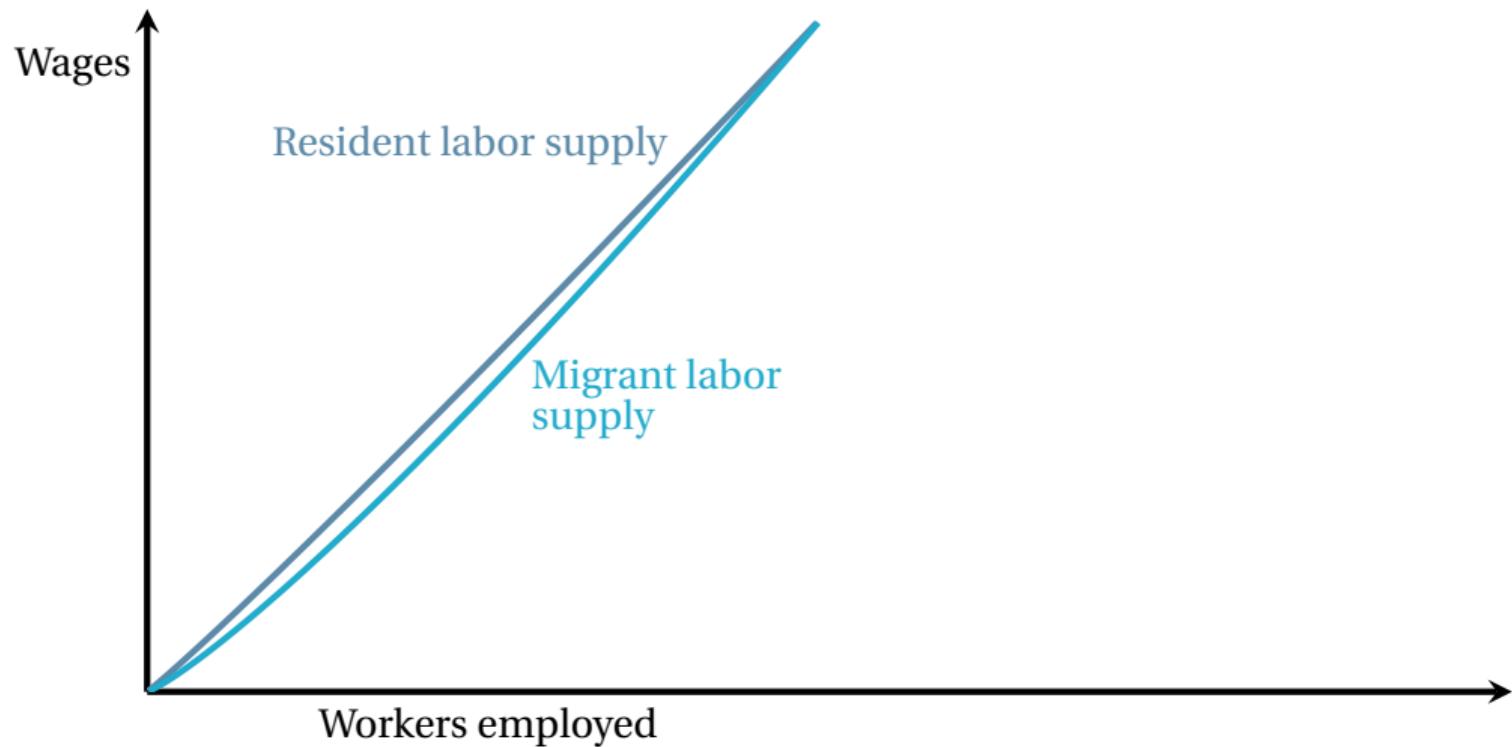
Posted wages when migrants' jobs are *unrestricted*



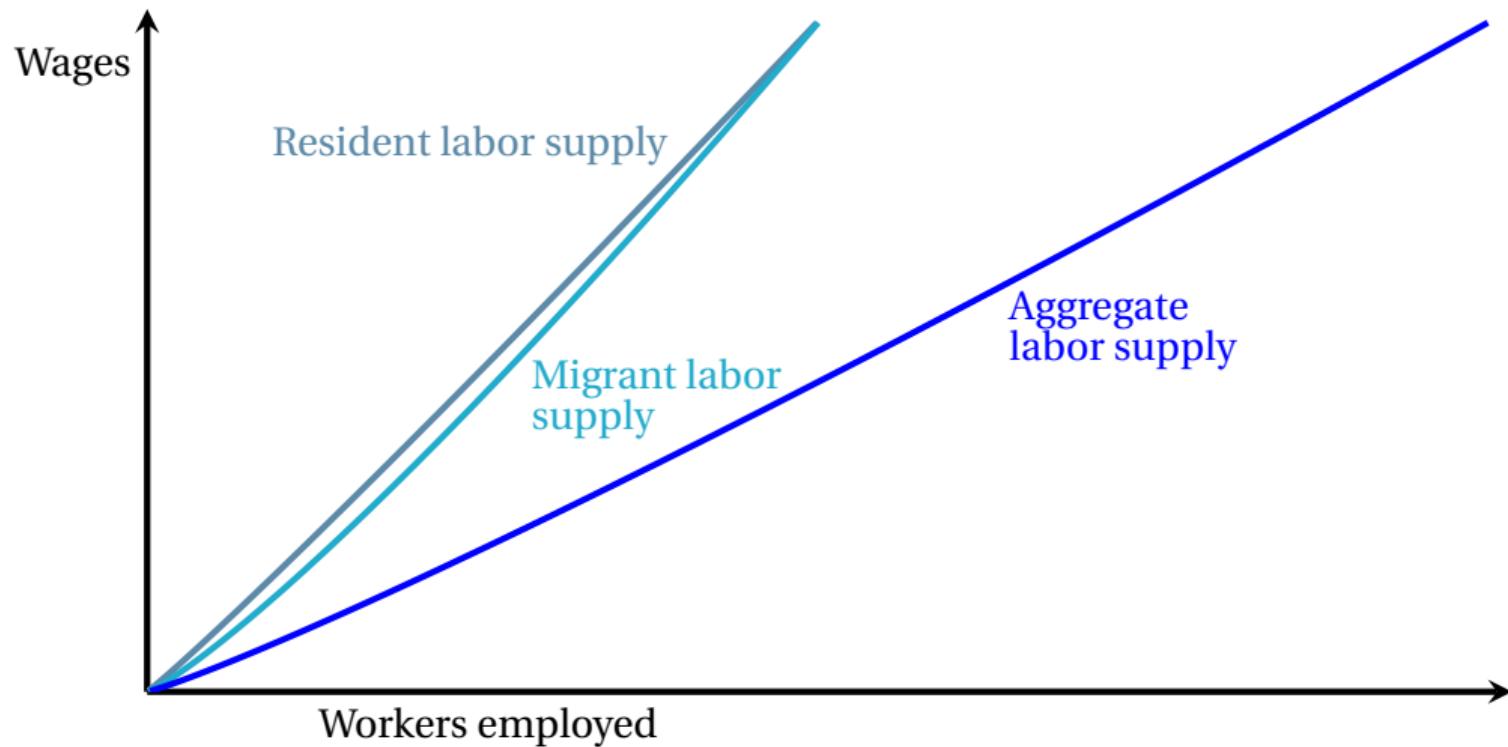
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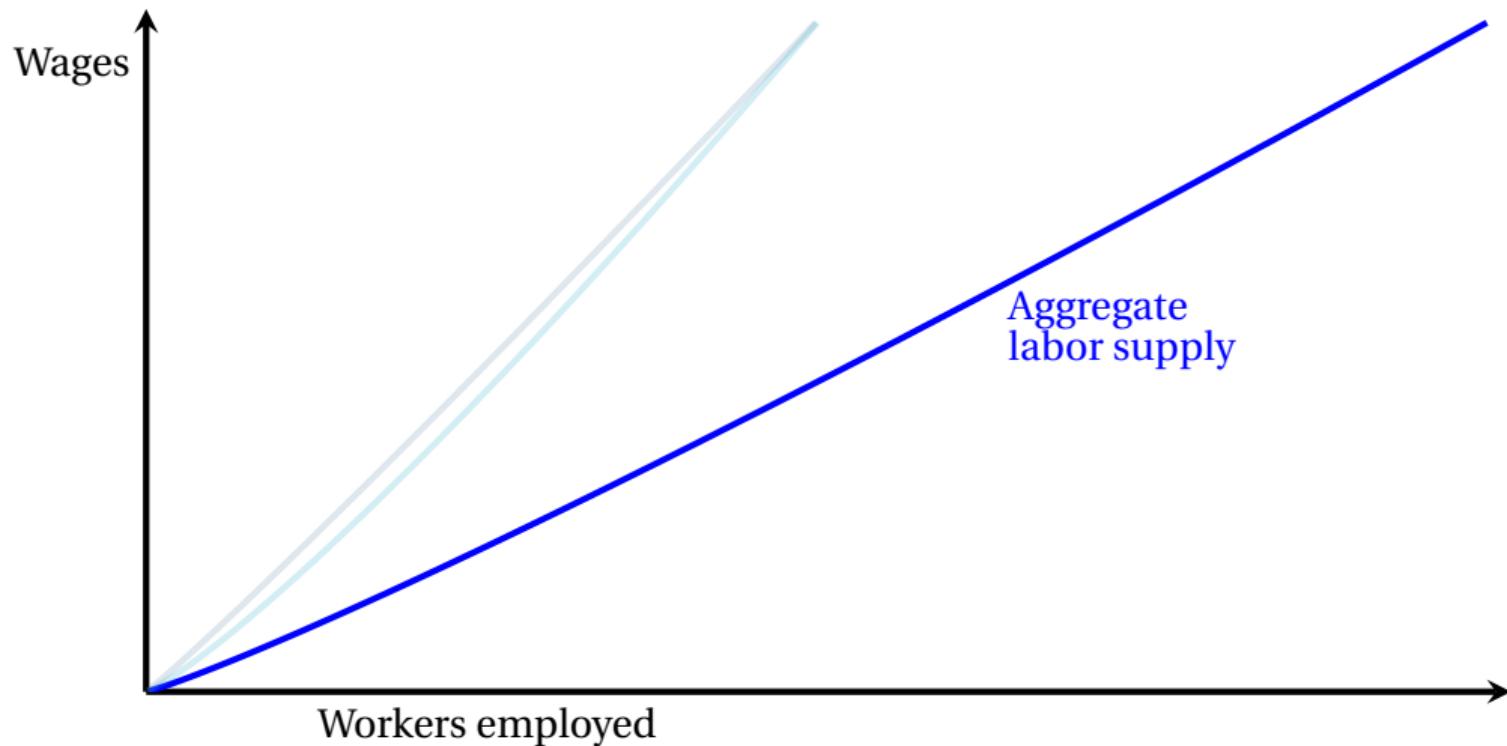
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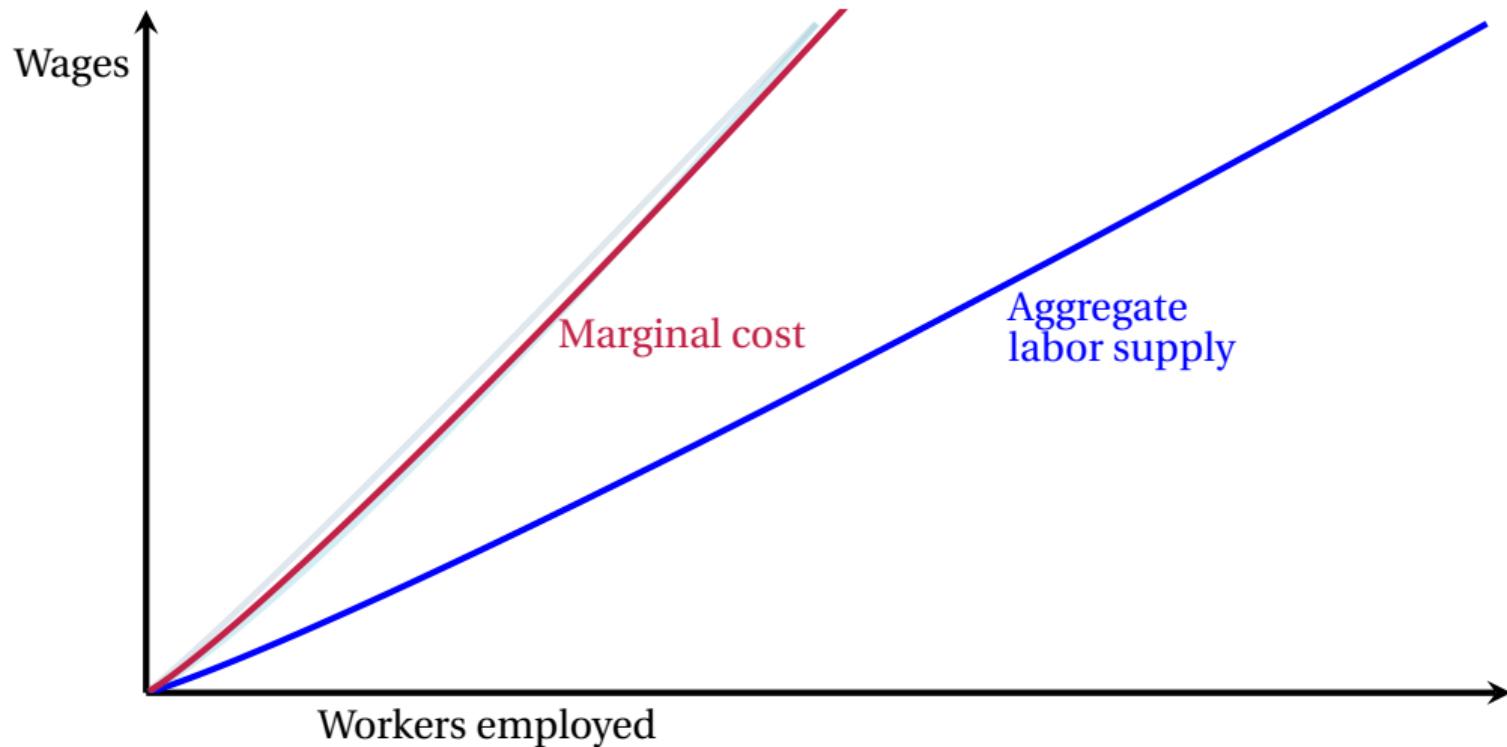
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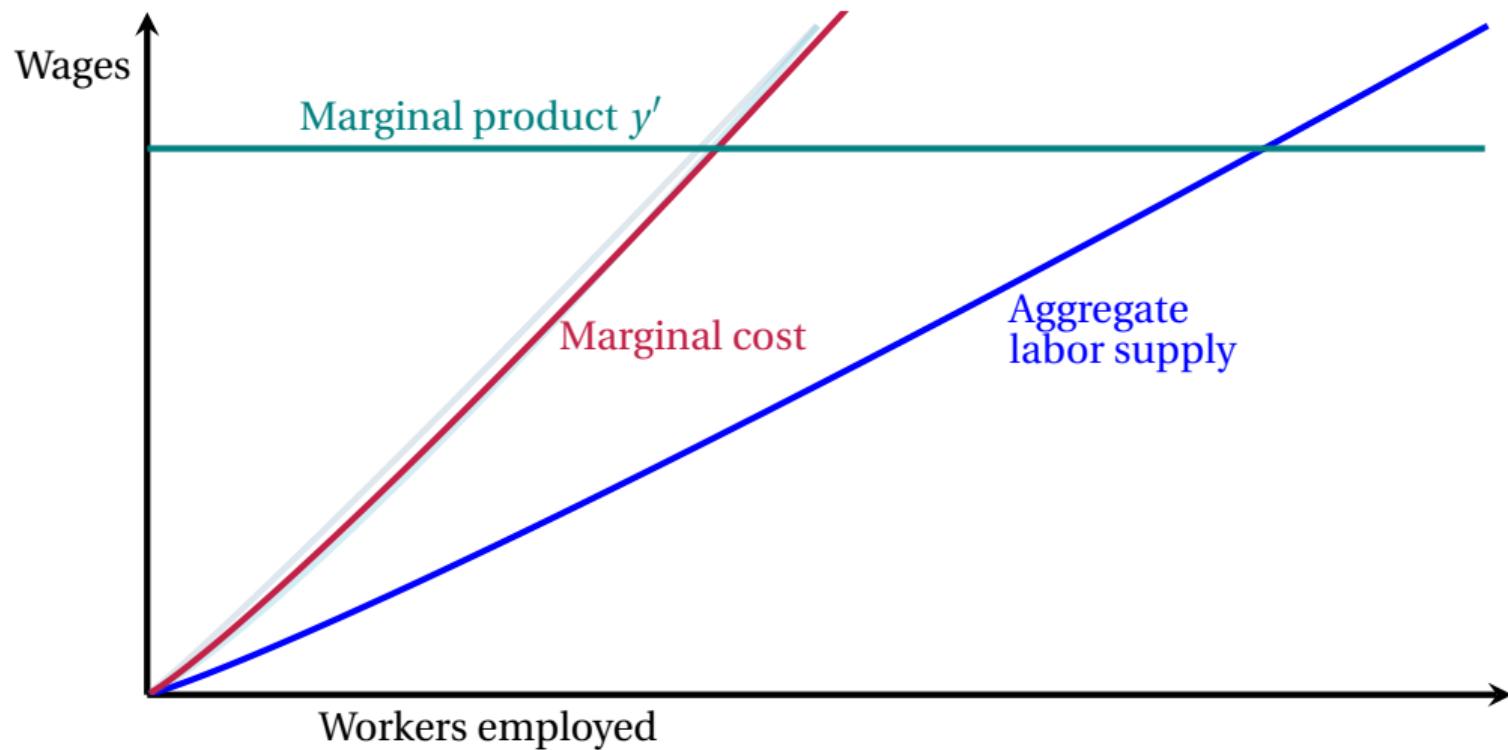
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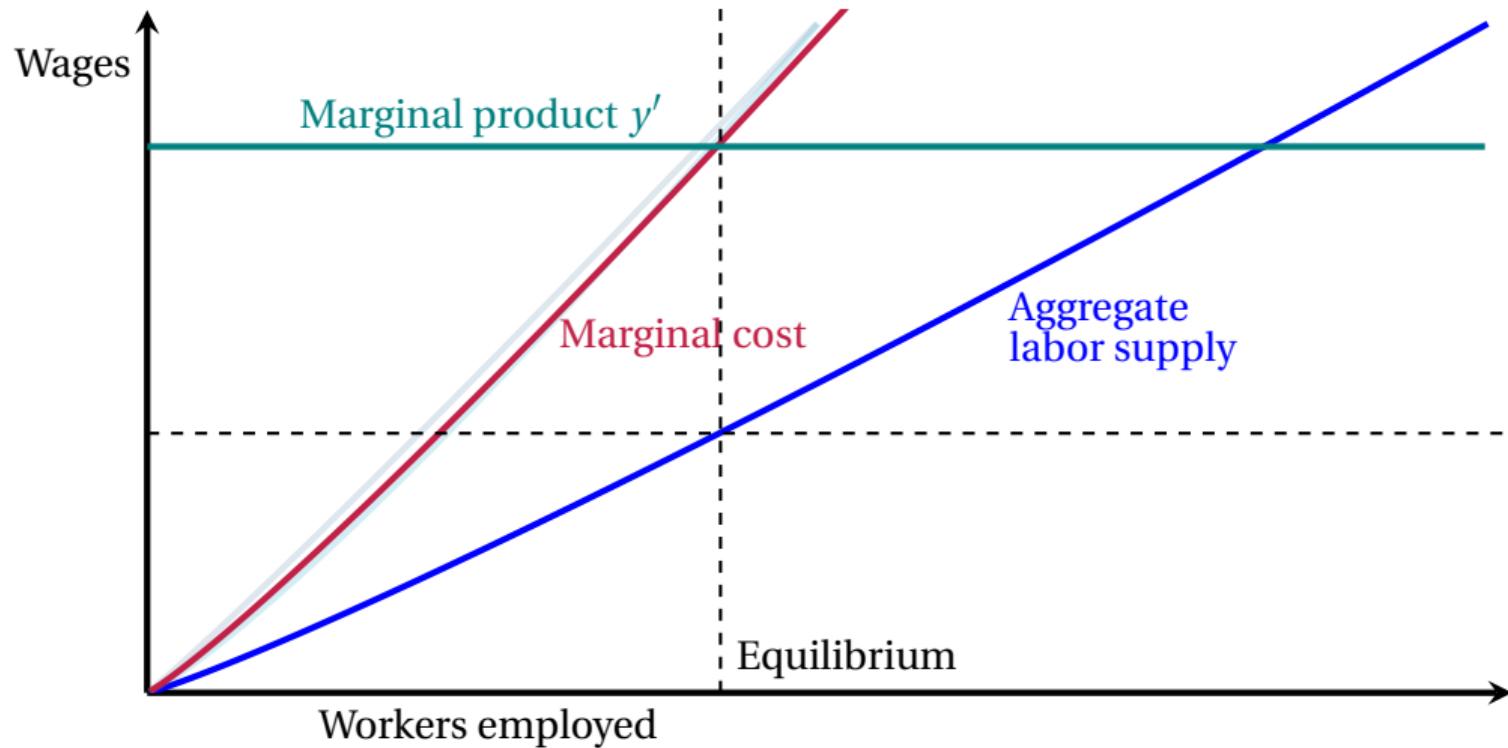
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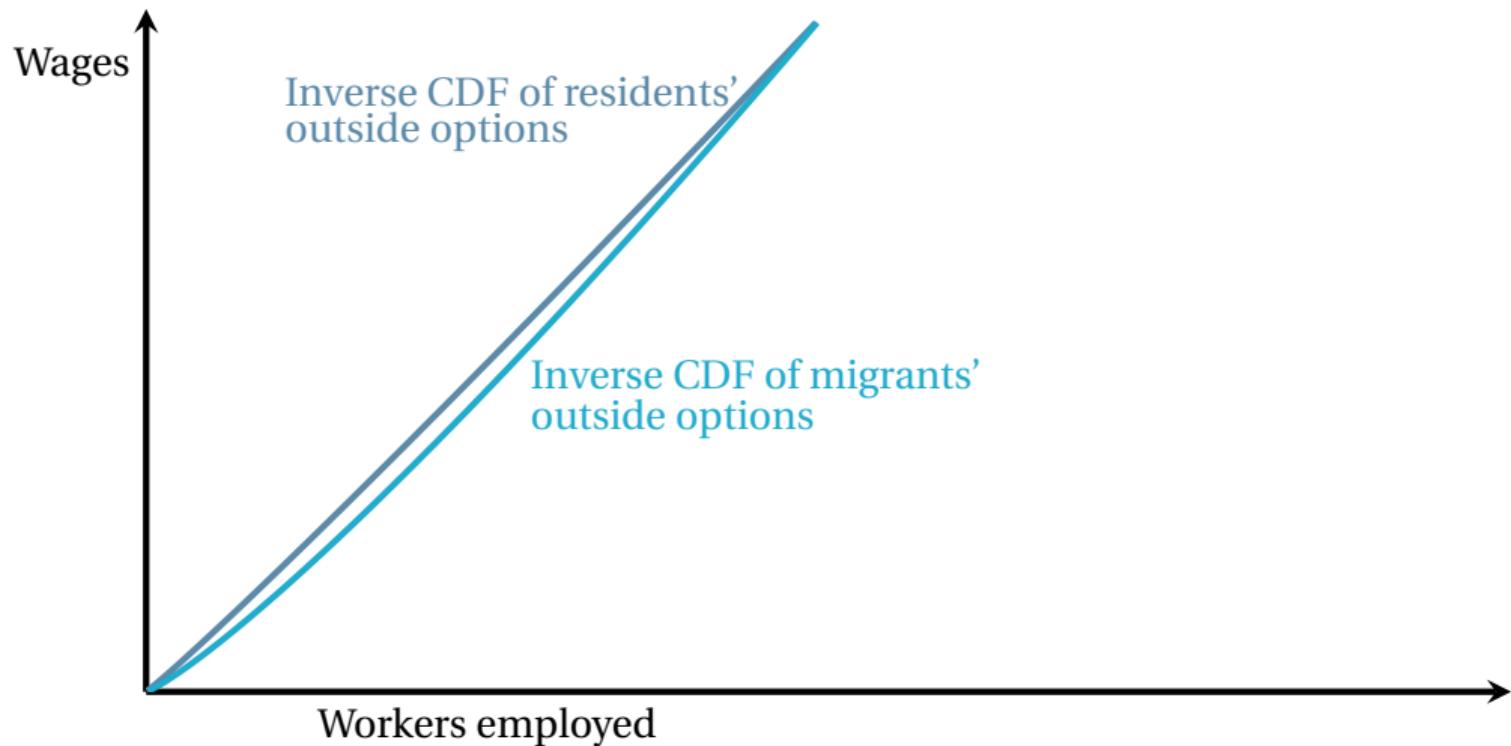
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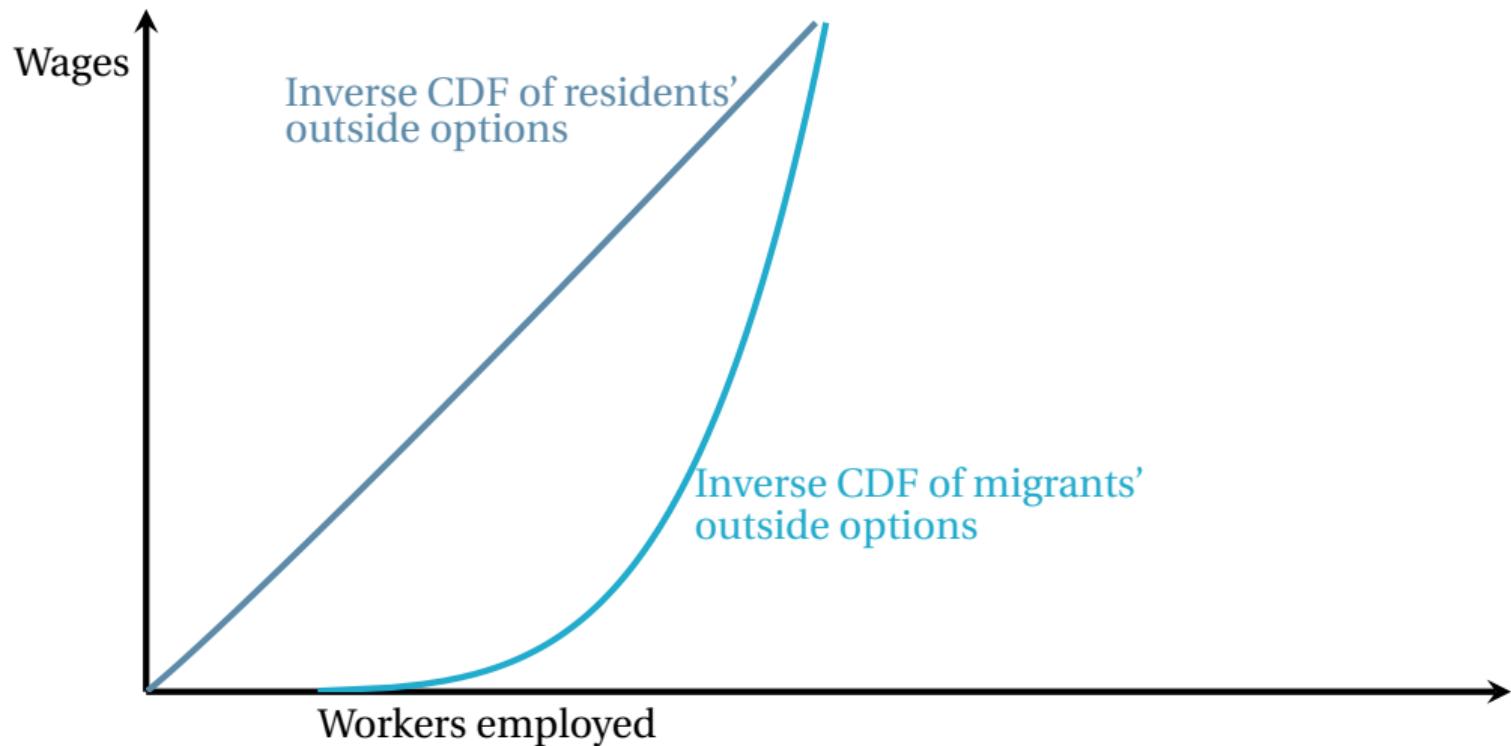
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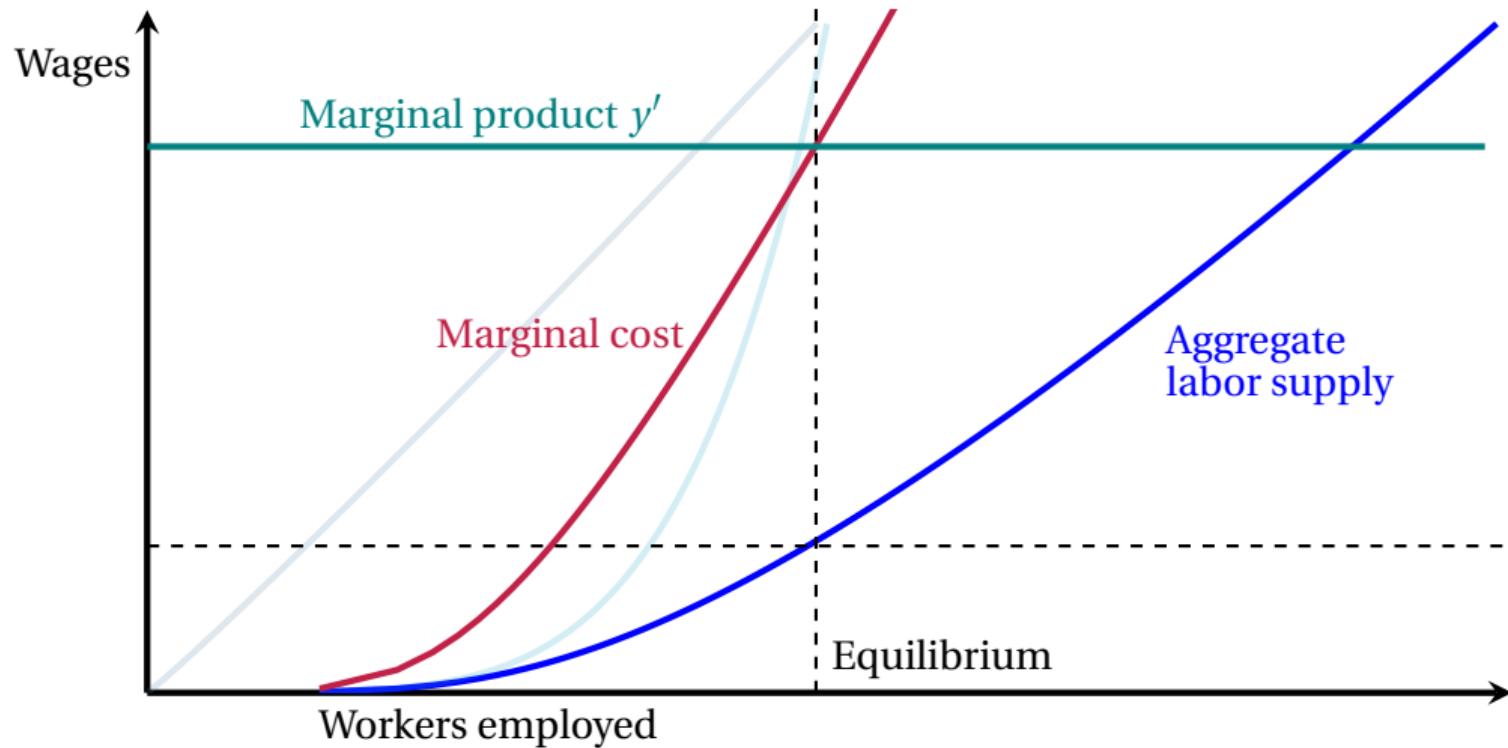
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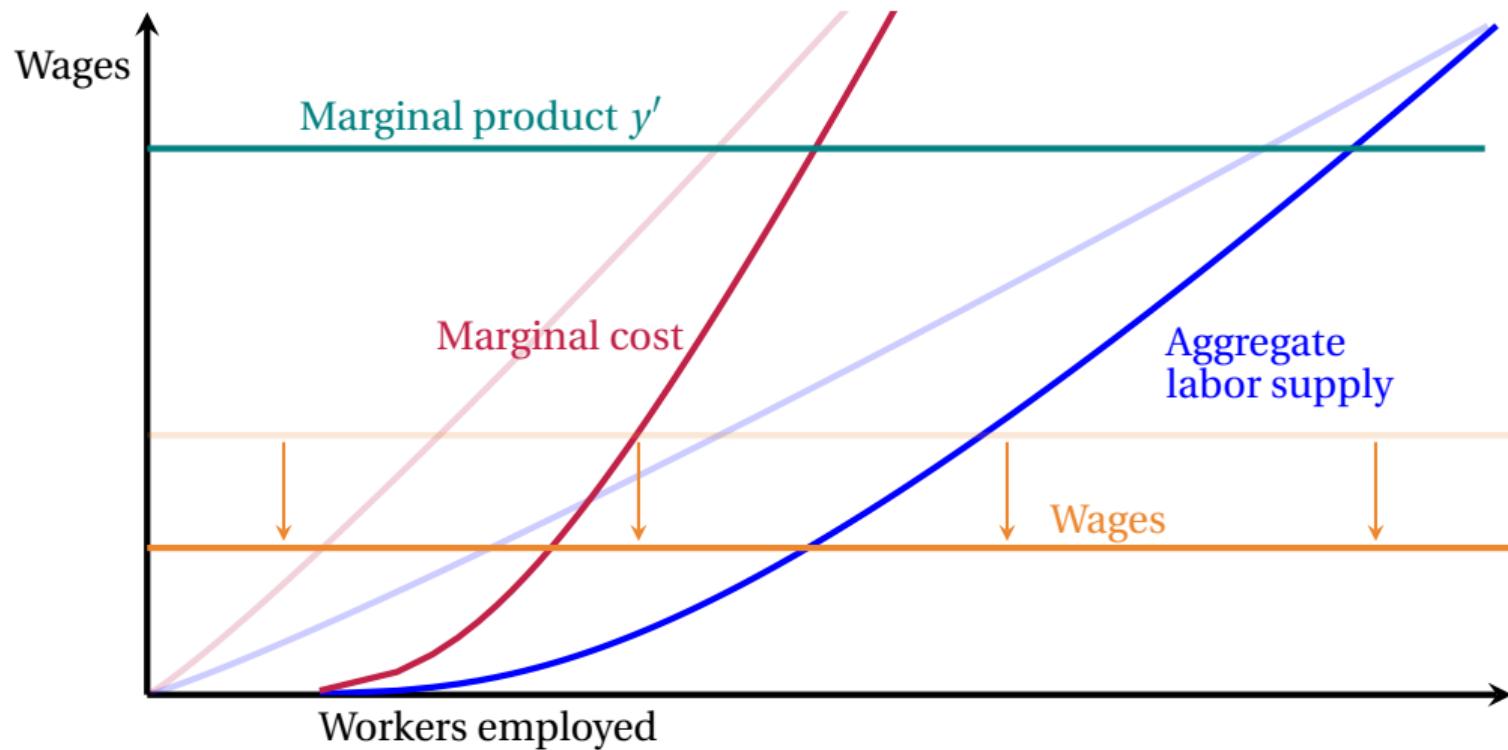
Posted wages when migrants' jobs are restricted



Posted wages when migrants' jobs are restricted



Restricting migrants' jobs reduces posted wages



Structural model: motivation

Research question: **how does restricting migrants job options affect wages, profits, and welfare?**

Limitation of reduced-form analyses:

- ▶ Can't identify the distribution of effects across workers.
- ▶ Firm-level data is noisy and often missing: can't estimate profit effects.
- ▶ Don't observe non-wage amenities: can't identify welfare effects.

Structural model: overview

The model is designed to capture two mechanisms.

Restricting migrants' job options will **decrease the elasticity of their labor supply, inducing firms to set less generous wages.**

- ▶ Empirical question: how restricted are migrants' job options in practice?
- ▶ Estimate workers' willingness to move across firms and space.

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Restricting migrants' job options will **segregate them into certain firms, decreasing their marginal product.**

- ▶ Empirical question: How concave is production? How easily can firms substitute across occupations?
- ▶ Estimate production functions which account for workers' occupations.

Structural model: primitives

Perfect information, partial equilibrium. Two periods, but decision-making is static.
In each period, the model comprises a continuum of workers I_t and finite firms F_t .

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Each worker $i \in \mathbf{I}_t$ has an **exogenous occupation** $o_{i,t}$ and supplies an **exogenous endowment of labor** $l_{i,t}$ to

$$f_{i,t} = \arg \max_{f \in \mathbf{F}_{i,t}} u_{i,t}(w_{f,o_{i,t},t}, f),$$

where $\mathbf{F}_{i,t}$ is the worker's choice set:

- ▶ If i is a *resident*: $\mathbf{F}_{i,t} = \mathbf{F}_{o_{i,t}}$;
- ▶ If i is a *migrant*: $\mathbf{F}_{i,t} = \mathbf{F}_{o_{i,t}}^{migrant} \subseteq \mathbf{F}_{o_{i,t}}$.

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A type- x worker has **nested-logit preferences** over firms f in locations c_f with wages w :

$$u_{i,t}(w, f) = \tau \log(l_{i,t} w) + \bar{\xi}_{c_f, o_{i,x}, t} + \frac{1}{\lambda} \zeta_{i, c_f, t} + \xi_{f, o_{i,x}, t} + \epsilon_{i,f,t}; \quad \zeta_{i, c_f, t}, \epsilon_{i,f,t} \perp l_{i,t}.$$

Structural model: primitives

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Firms have **CES production functions**:

$$y_{f,t} \left((L_{f,o,t})_{o \in \mathbf{O}_{f,t}} \right) = \left(\sum_{o \in \mathbf{O}_{f,t}} e^{\phi_{f,o,t}} L_{f,o,t}^\rho \right)^{\frac{v}{\rho}}, \quad L_{f,o,t} \equiv \int_{i:f_i=f, o_i=o} l_{i,t} di.$$

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Firms post wages $w_{f,o,t}$ in **Bertrand competition**:

$$\vec{w}_{f,t} \in \arg \max_{\vec{w} \in \mathbf{R}^{|\mathbf{O}_{f,t}|}} \{ y_f (\vec{L}_{f,t} (\vec{w}; \vec{w}_{-f,t})) - \vec{w}' \vec{L}_{f,t} (\vec{w}; \vec{w}_{-f,t}) \}.$$

Structural model: calculating structural residuals

Given a candidate parameter vector τ, λ, ρ, ν :

1. Calculate the **location and firm amenity values** $\bar{\xi}_{c,o,x,t}, \xi_{f,o,x,t}$ given observed employment, wages and the parameters τ, λ (Berry '94).

$$\log N_{f,o,x,t}^{\text{resident}} = \text{Constant}_{x,c_f,t} + \tau \log(w_{f,o,t}) + \xi_{f,o,x,t}$$

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2. Calculate $\eta_{f,o,x,t}^\nu$: **the labor supply elasticity of type- x occupation- o workers with visa-status ν** , given observed market shares and the parameters τ, λ .

$$\eta_{f,o,x,t}^\nu = \tau \left(1 - (1 - \lambda) \text{regional market share}_{f,o,x,\nu,t} - \lambda \text{national market share}_{f,o,x,\nu,t} \right)$$

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3. Calculate $\eta_{f,o,t}$: **the labor supply elasticity of occupation- o workers** to firm f , by aggregating over visa statuses and worker types.

$$\eta_{f,o,t} = \frac{1}{N_{f,o,t}} \sum_{x \in X} \left(N_{f,o,x,t}^{\text{resident}} \eta_{f,o,x,t}^{\text{resident}} + N_{f,o,x,t}^{\text{migrant}} \eta_{f,o,x,t}^{\text{migrant}} \right)$$

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3. Calculate $\eta_{f,o,t}$: **the labor supply elasticity of occupation- o workers** to firm f , by aggregating over visa statuses and worker types.
4. Calculate the **marginal product of labor for each occupation at each firm** by inverting the markdown equation, given wages and the labor supply elasticities $\eta_{f,o,t}$.

$$w_{f,o,t} = \frac{\partial y_{f,t}}{\partial L_{f,o,t}} \left(1 + \frac{1}{\eta_{f,o,t}}\right)^{-1}$$

Structural model: calculating structural residuals

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5. Use these marginal products, observed employment, and the production parameters ρ, ν **to calculate productivity residuals** $\phi_{f,o,t}$.

$$\frac{\partial y_{f,t}}{\partial L_{f,o,t}} = \nu \left(\sum_{o' \in \mathbf{O}_{f,t}} e^{\phi_{f,o',t}} L_{f,o',t}^\rho \right)^{\frac{\nu-\rho}{\rho}} e^{\phi_{f,o,t}} L_{f,o,t}^{\rho-1}$$

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Explicit form for structural residuals.

Structural model: data

Sample: Paid employees with a unique firm during the three months prior to the 2013 and 2018 censuses.

Wages: The log wage in an occupation-by-firm is the mean log earnings less mean log hours.

Worker types: Born in New Zealand vs. born overseas.

Locations: Functional urban areas.

Firms: Enterprise \times location.

Structural model: moment conditions

Key assumption: shocks to the amenity-value of employment for NZ-born workers $\Delta\xi_{f,o,NZ,t}$ are orthogonal to both

- ▶ productivity shocks $\Delta\phi_{f,o,t}$, and
- ▶ baseline productivity $\phi_{f,o,t-1}$.

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- ▶ Unobserved worker productivity generates measurement error in wages.
- ▶ Measurement error in wages introduces a spurious correlation.

Solution: calculate wages, productivity and amenities in each of two random split-samples. Require e.g. productivity shocks in one sample be orthogonal to amenity shocks in the other sample.

Explicit form for moment conditions.

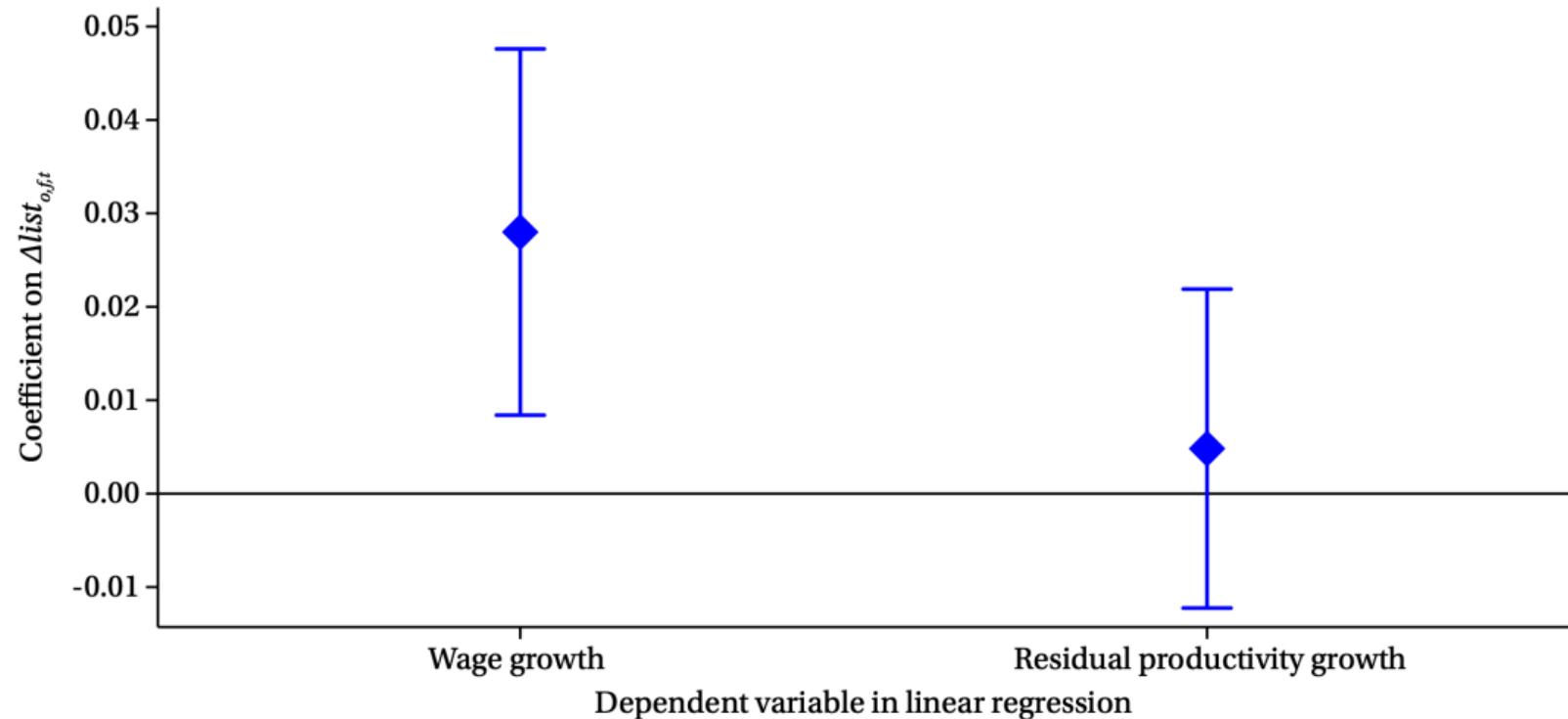
Structural model: estimates

	Baseline spec.
τ (atomistic labor supply)	3.446 [0.544]
λ (workers substitute across space)	0.483 [0.242]
ρ (firms substitute across occs.)	0.970 [0.026]
v (returns to scale)	0.541 [0.097]

Structural model: estimates

	Baseline spec.	Richer demog.	Randomize workers	Coefficients by sector			Ignore hours
τ (atomistic labor supply)	3.446 [0.544]	3.547 [0.614]	3.564 [0.614]			3.356 [0.537]	3.314 [0.544]
λ (workers substitute across space)	0.483 [0.242]	0.513 [0.305]	0.682 [0.294]			0.463 [0.225]	0.456 [0.216]
				Primary	Manuf.	Services	
ρ (firms substitute across occs.)	0.970 [0.026]	0.963 [0.029]	0.919 [0.032]	1.023 [0.080]	0.930 [0.046]	0.975 [0.029]	0.969 [0.026]
v (returns to scale)	0.541 [0.097]	0.536 [0.089]	0.473 [0.093]	0.822 [0.132]	0.541 [0.178]	0.547 [0.101]	0.550 [0.092]

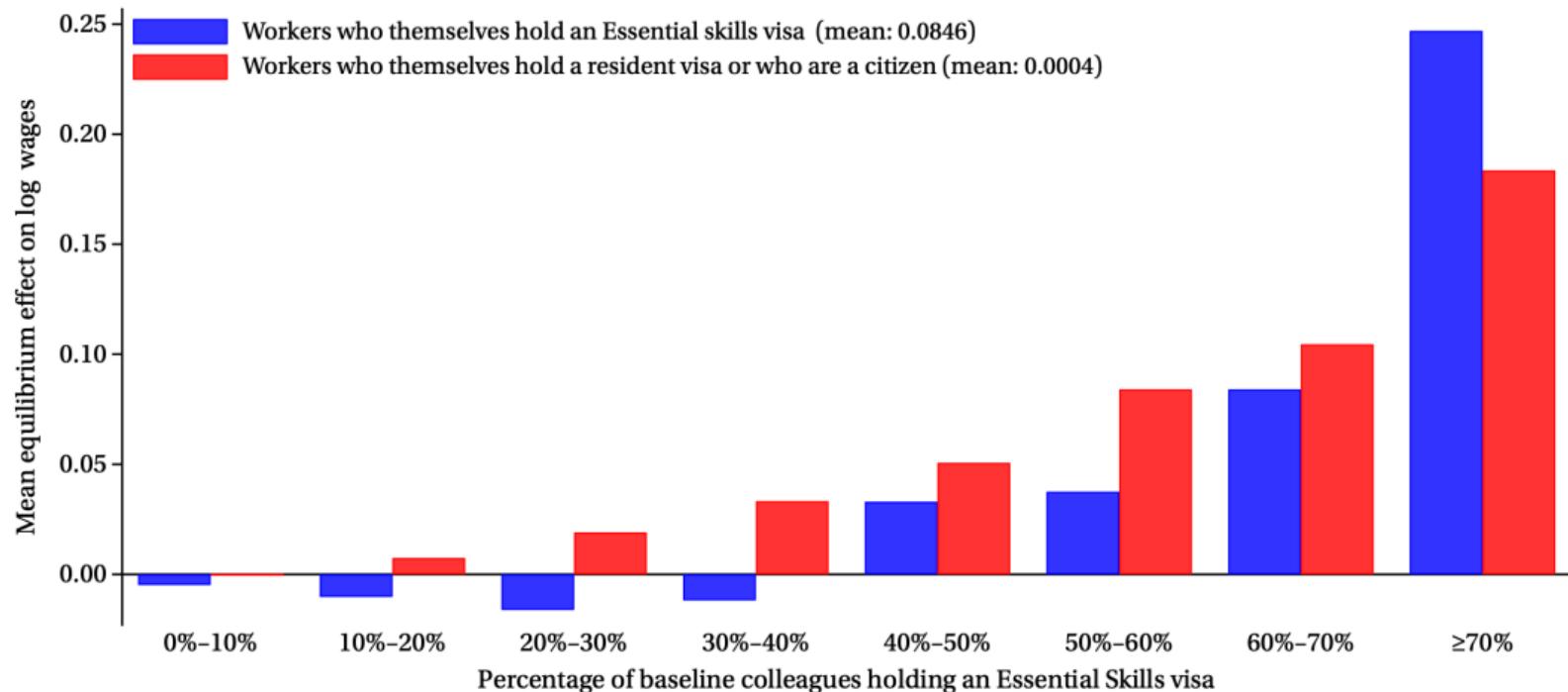
Structural model: over-identification



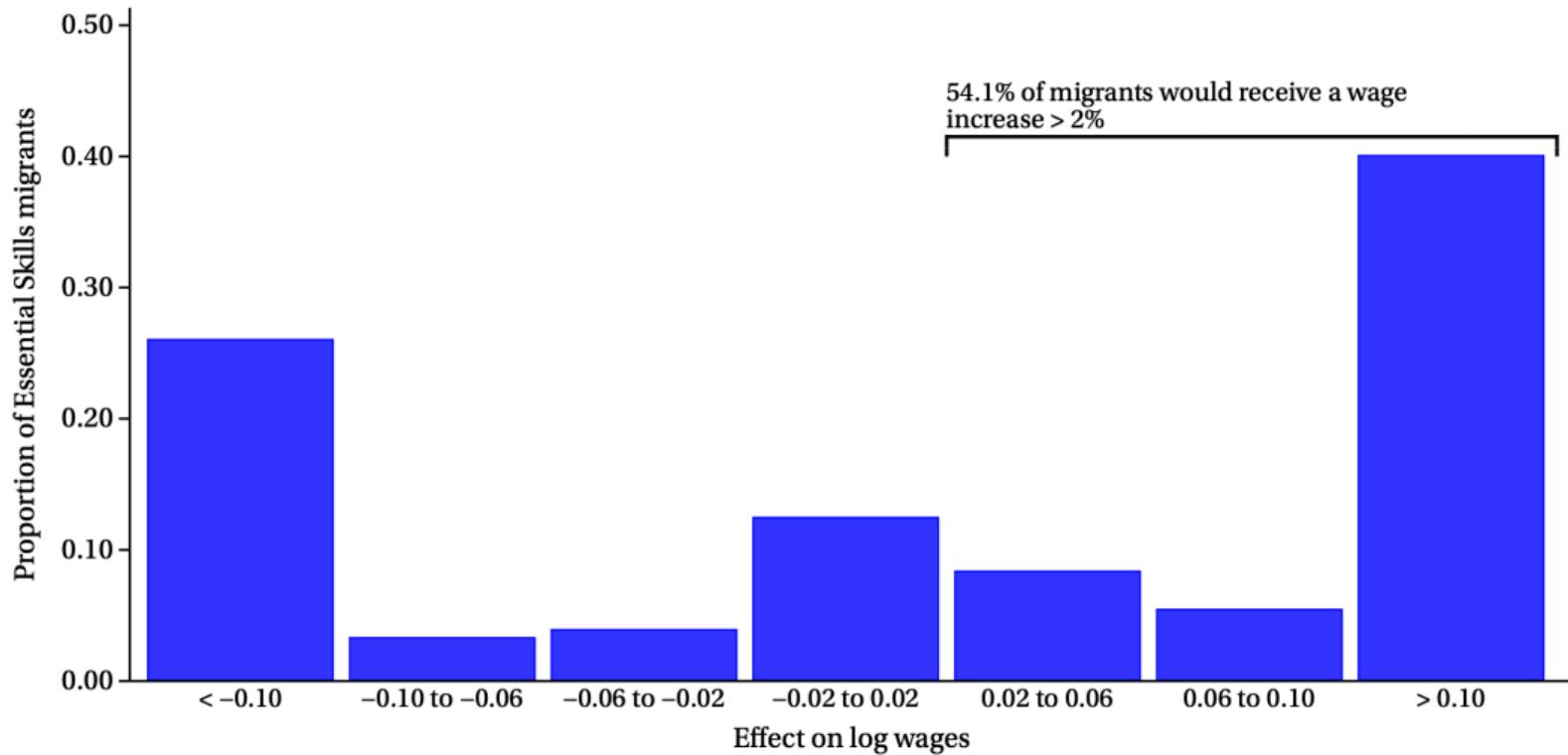
The equilibrium effect of allowing migrants to work at any firm

1. Use only 2018 cross-section, assume constant worker productivity.
2. Calculate productivity as in estimation.
3. Calculate amenities – imputing when necessary.
4. Treating labor as continuous, solve for the actual equilibrium and a counterfactual equilibrium in which migrants' job options are unrestricted.
5. Allocate discrete workers to actual and counterfactual firms using a
NOVEL ALGORITHM FOR NESTED EXTREME VALUE RANDOM VARIABLES.

The equilibrium effect of allowing migrants to work at any firm

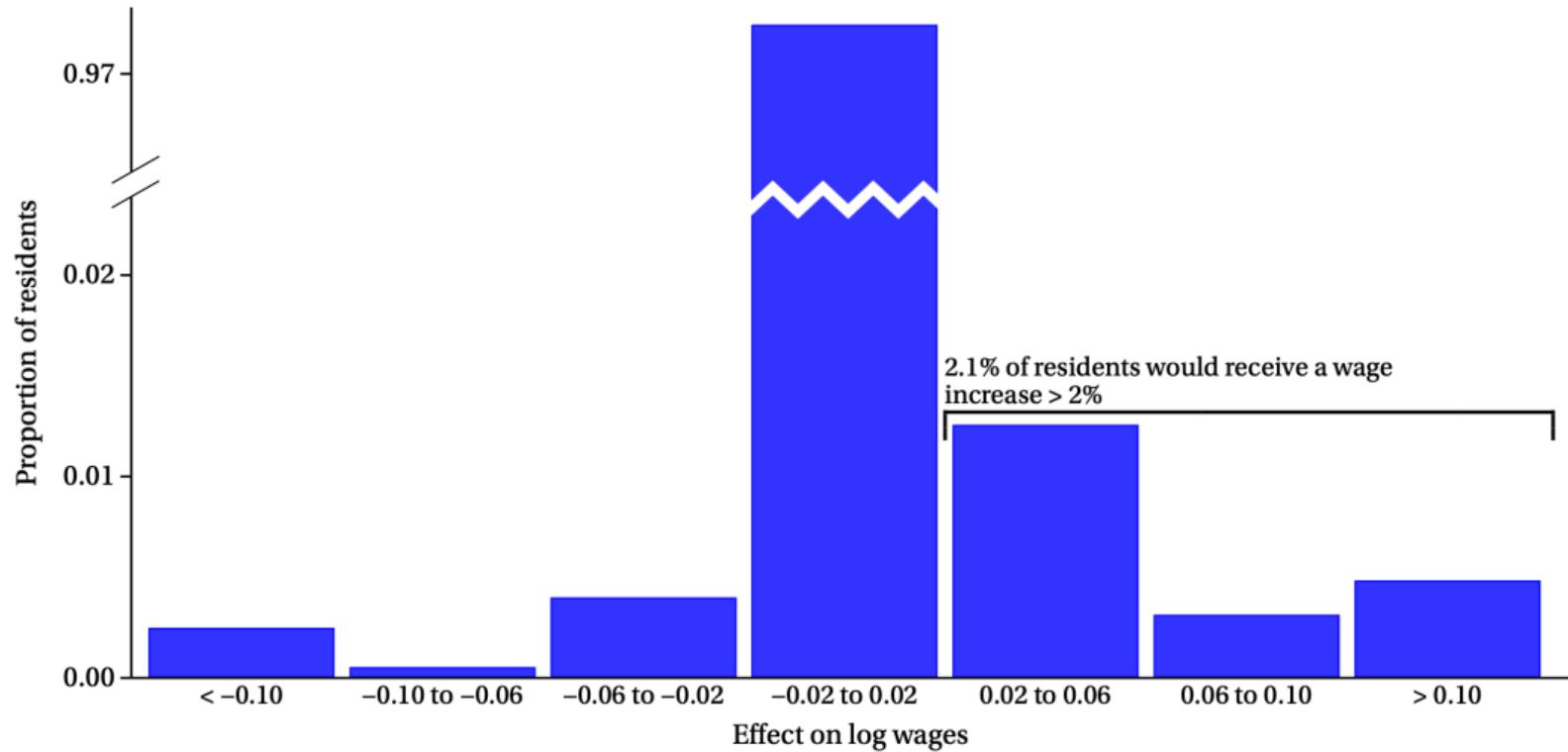


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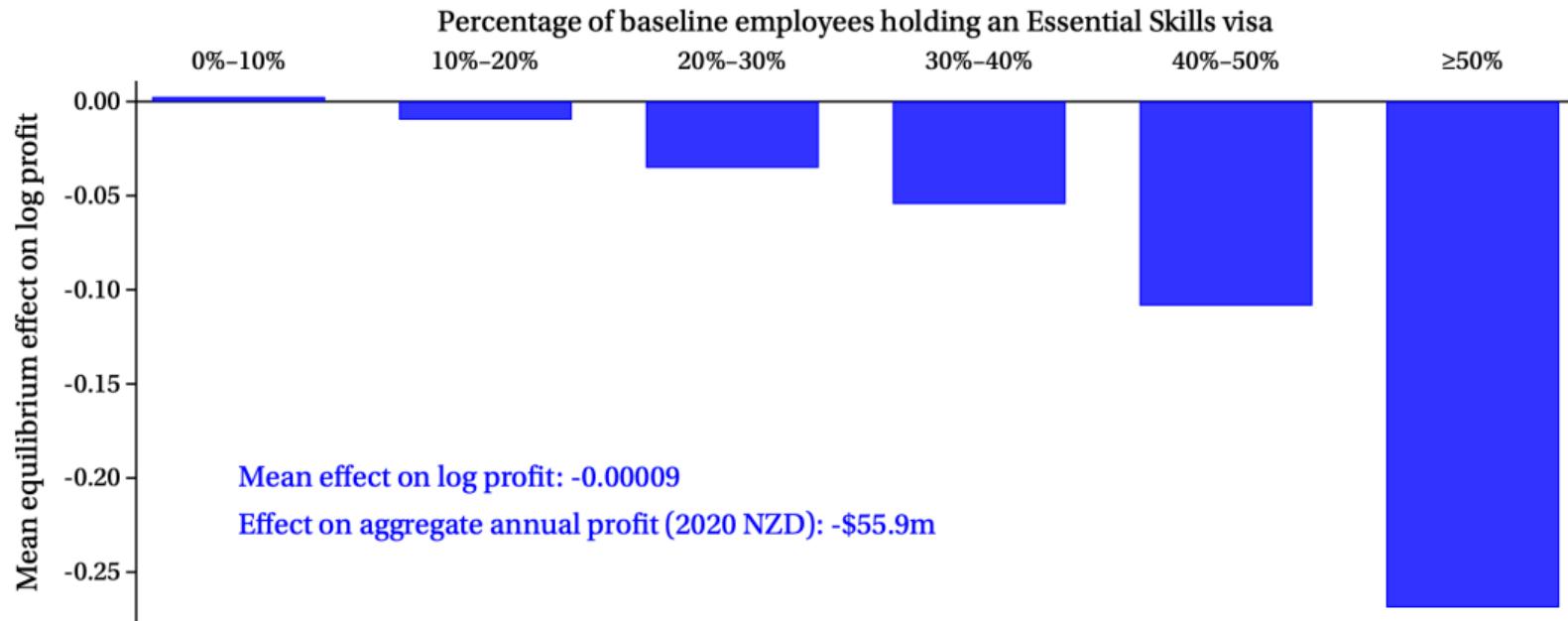
Unpack.

The equilibrium effect of allowing migrants to work at any firm



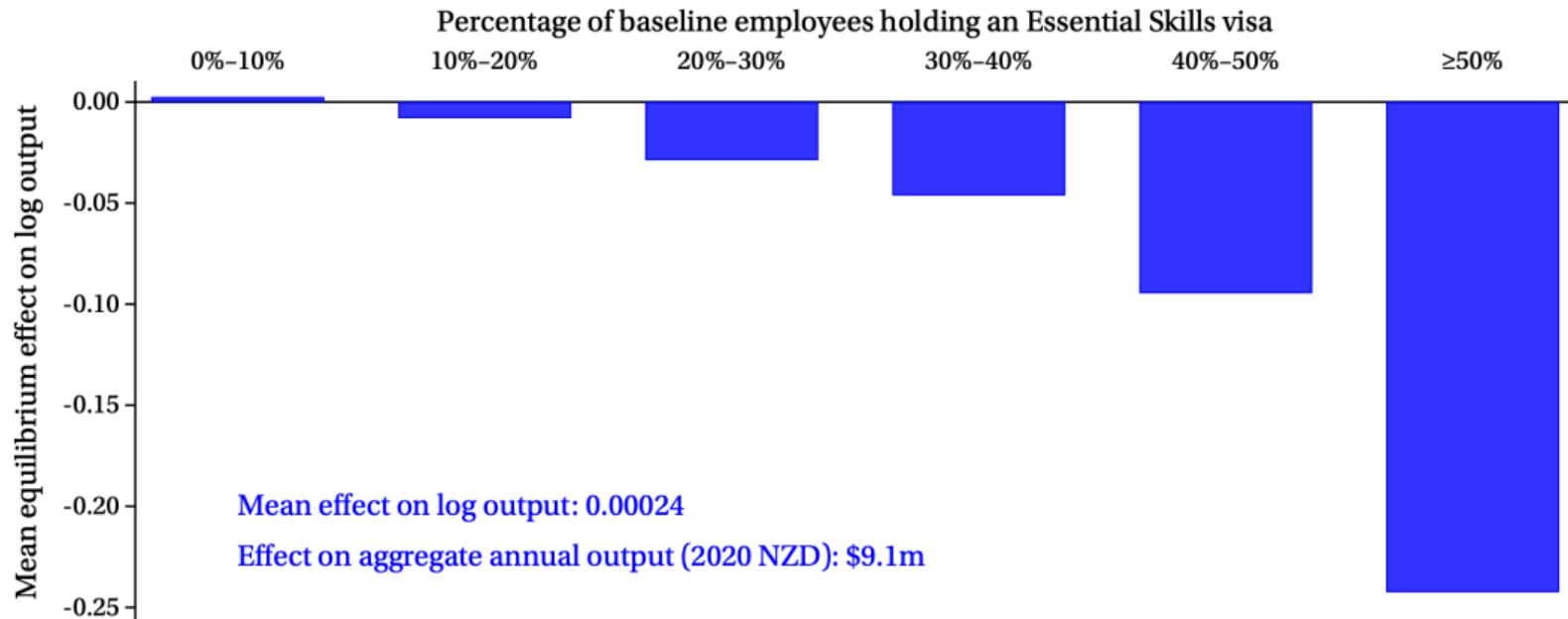
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The equilibrium effect of allowing migrants to work at any firm



How restricting migrants' job options affects welfare

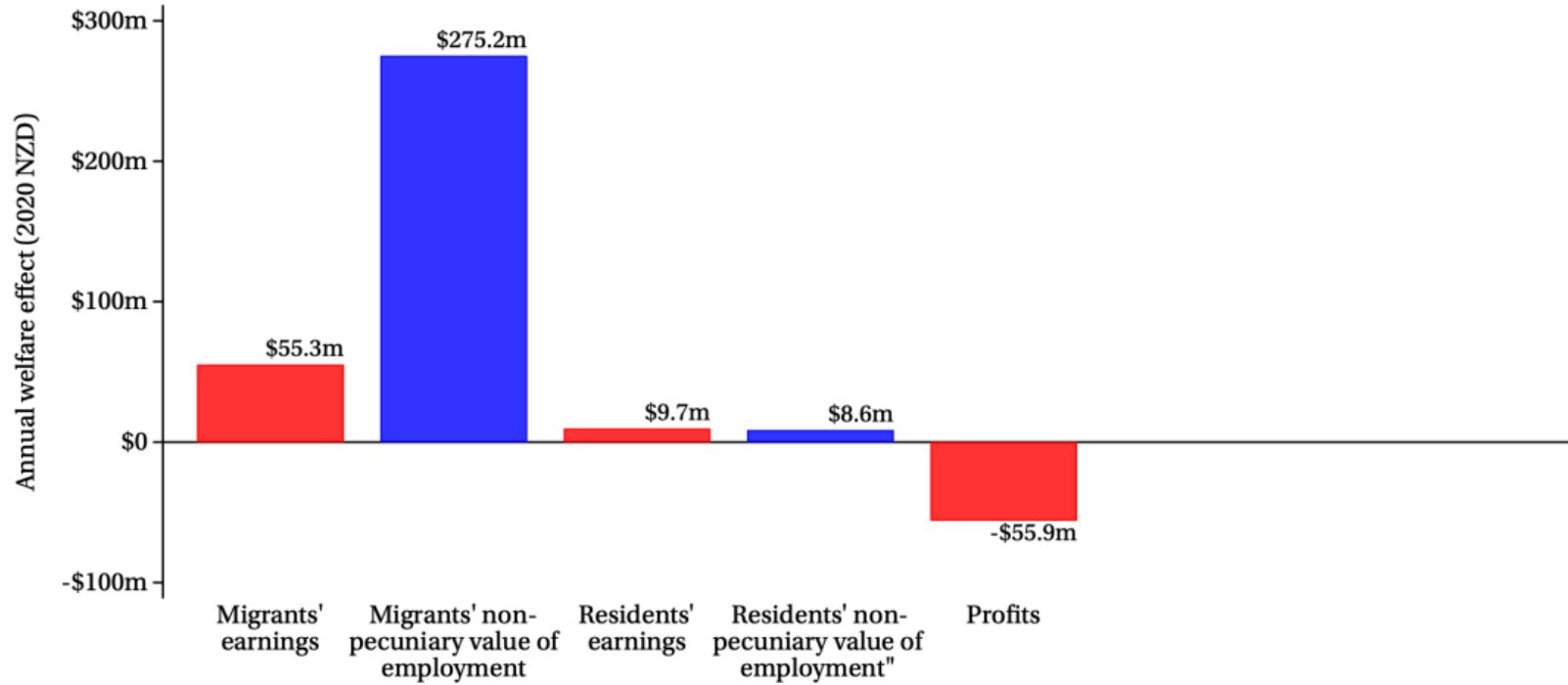
Let WTP_i denote i 's money-metric welfare gain from the counterfactual equilibrium w', f' , in which migrants can work at any firm:

$$u_i(w'_{o_i, f'_i} - WTP_i, f'_i) = u_i(w_{o_i, f_i}, f_i),$$

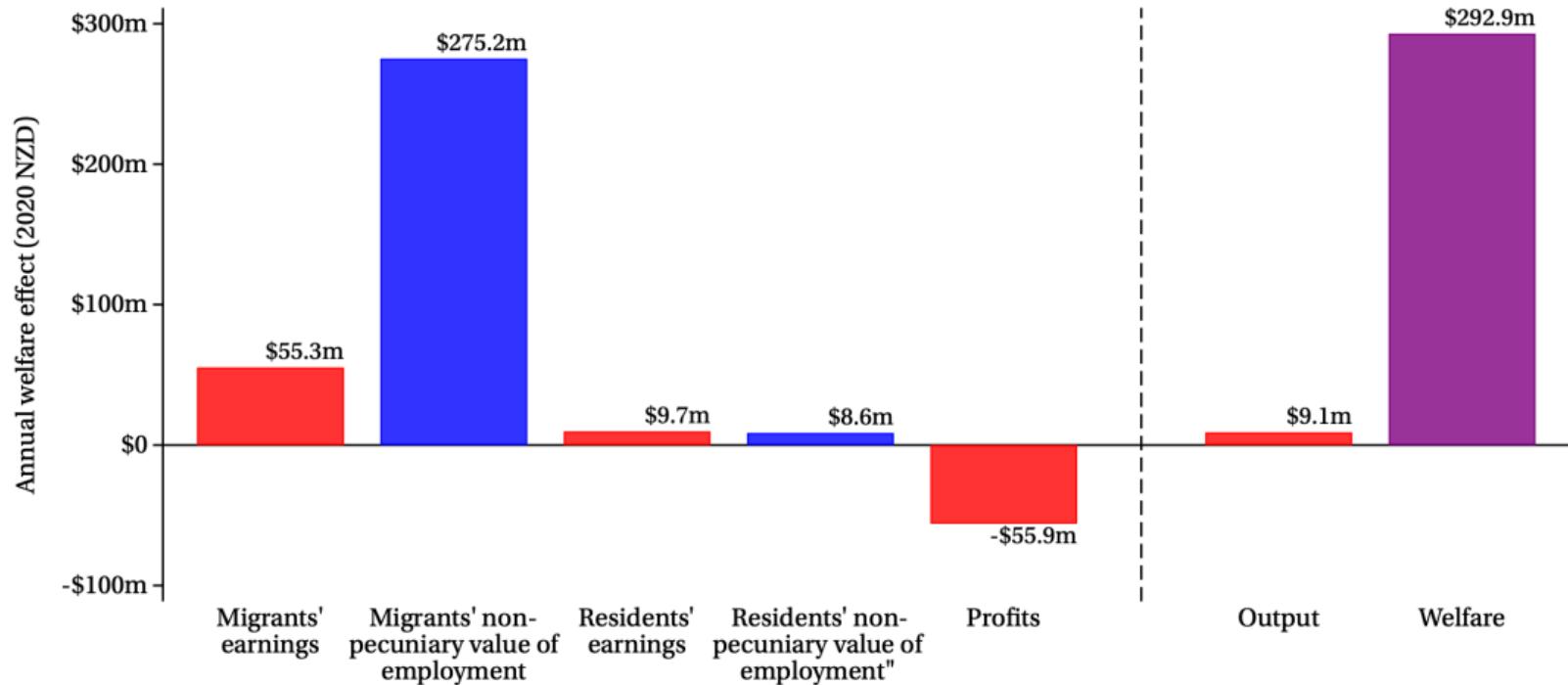
which we can decompose into

- ▶ an earnings effect $w'_{o_i, f'_i} - w_{o_i, f_i}$, and
- ▶ a non-penuniary effect $WTP_i - (w'_{o_i, f'_i} - w_{o_i, f_i})$.

How restricting migrants' job options affects welfare



How restricting migrants' job options affects welfare



Some additional structural results which we definitely won't have time to discuss

- ▶ Does the Essential Skills visa protect residents from the wage effects of increased migration?
- ▶ Who are the migrants whose wages would increase or decrease?
- ▶ Who are the residents whose wages would increase or decrease?
- ▶ Which are the firms whose profits would increase or decrease?

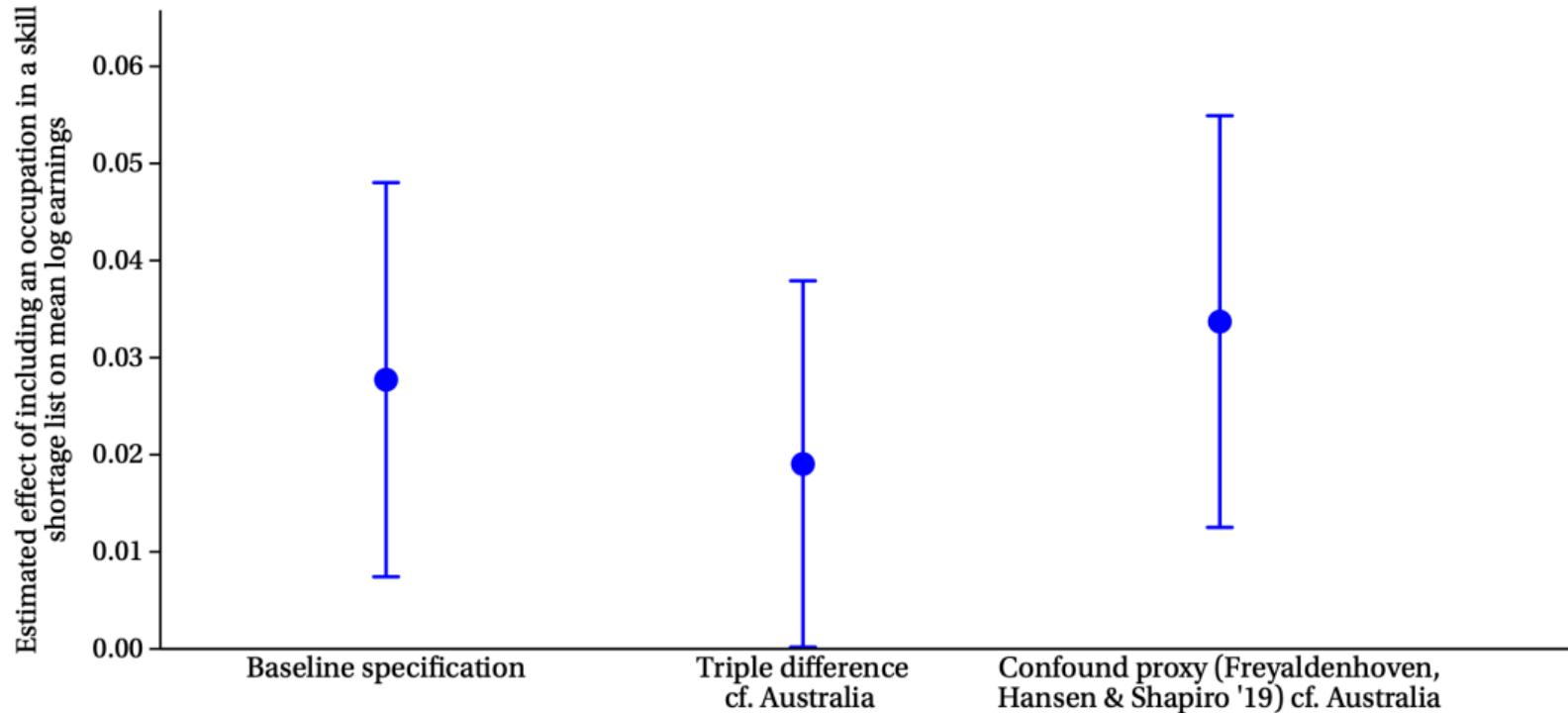
Restricting migrants' job options affects everyone's wages

Firms don't *specifically* discriminate against migrants,

But firms do account for migrants' weaker job options when setting everyone's wages.

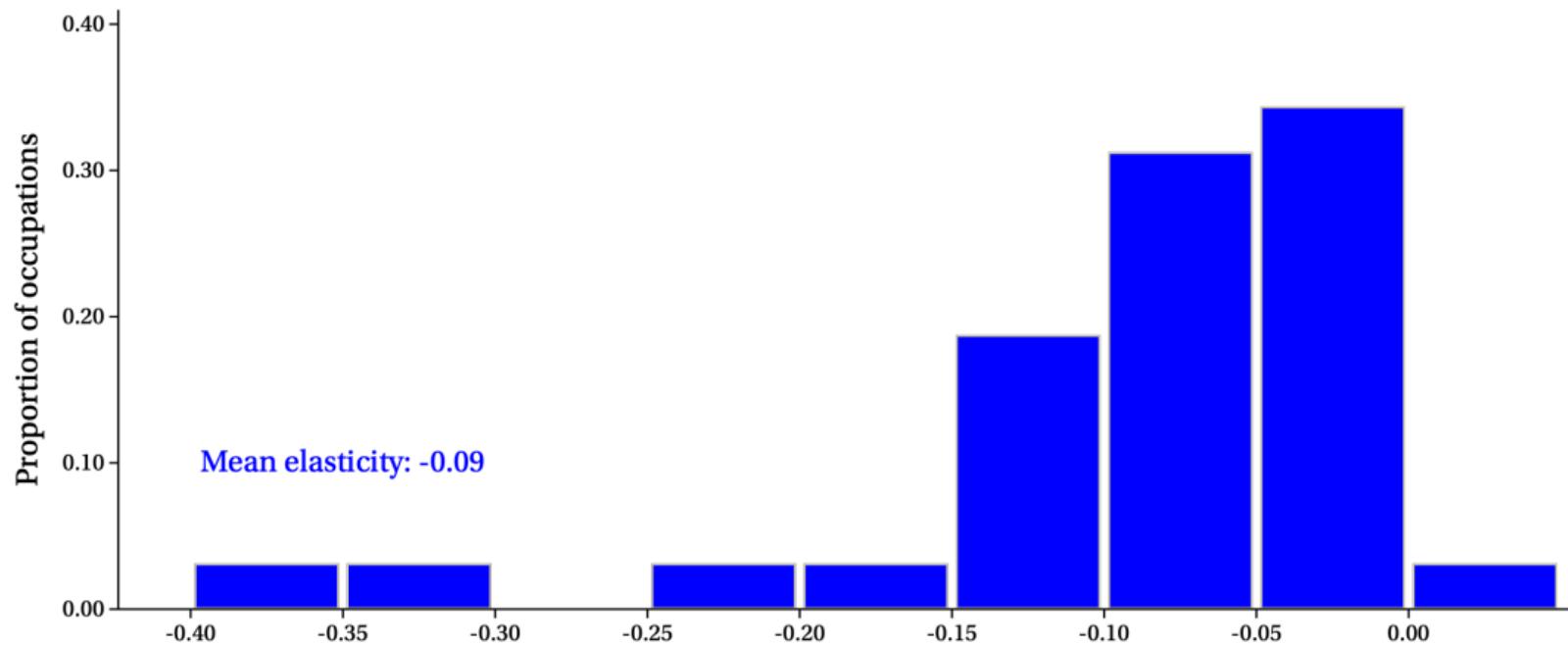
Restricting migrant's job options reduces both their wages and the wages of their colleagues.

Using Australian data as a placebo outcome



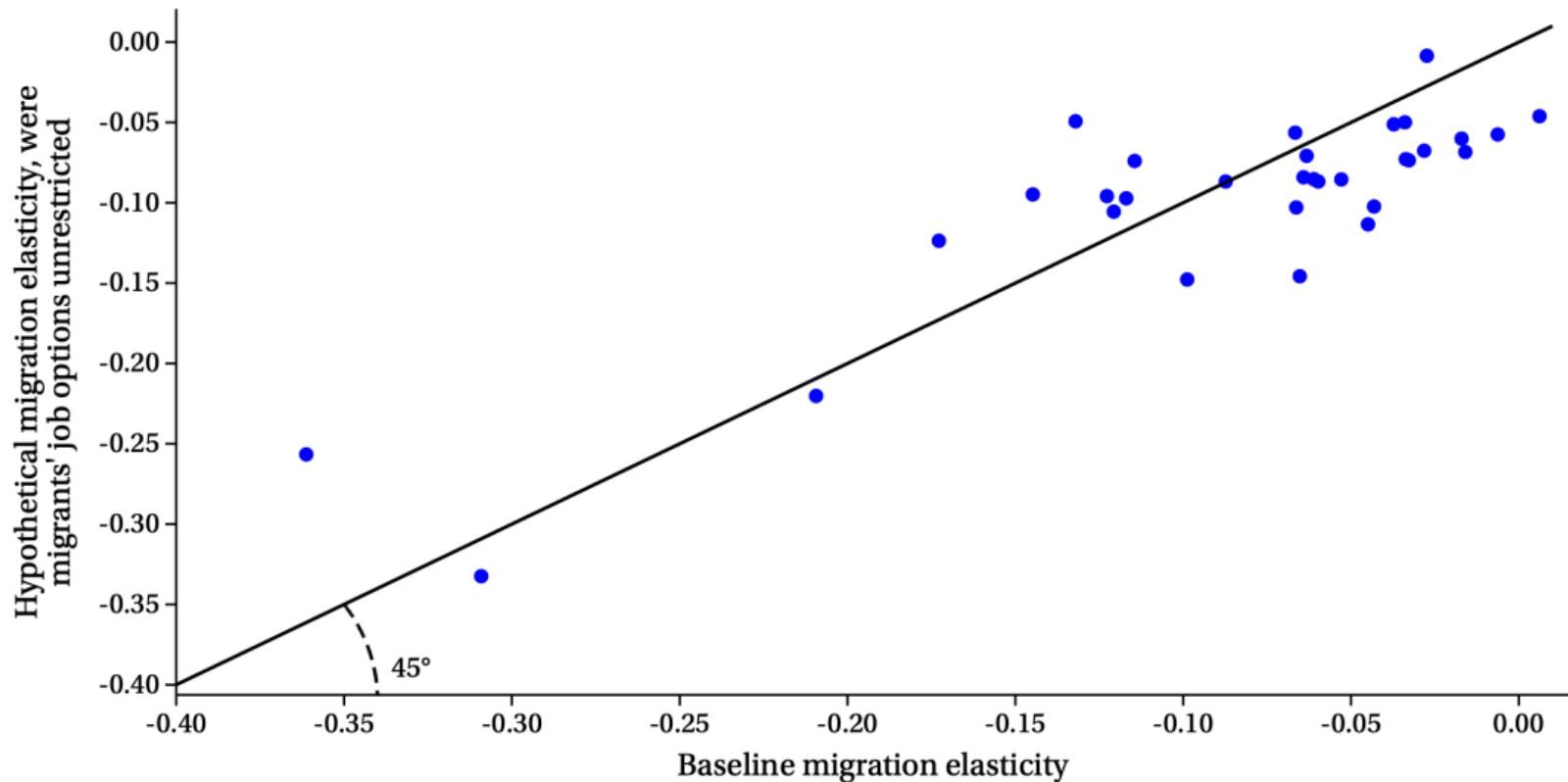
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Does restricting migrants' job options protect residents from immigration?

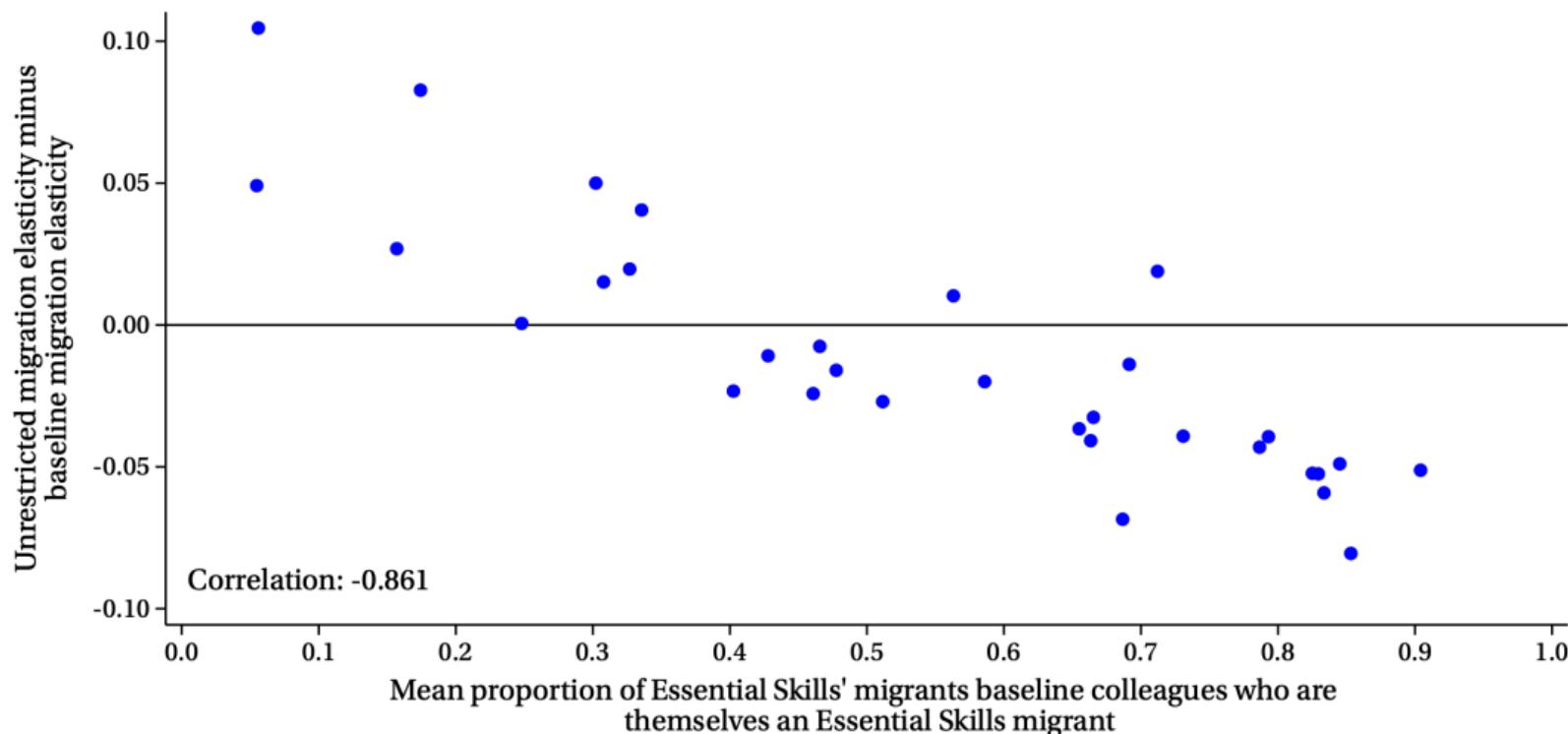


$\frac{\text{change in residents' mean log wage, given a 5\% increase in the number of migrants}}{\text{change in the log number of workers, given a 5\% increase in the number of migrants}}$.

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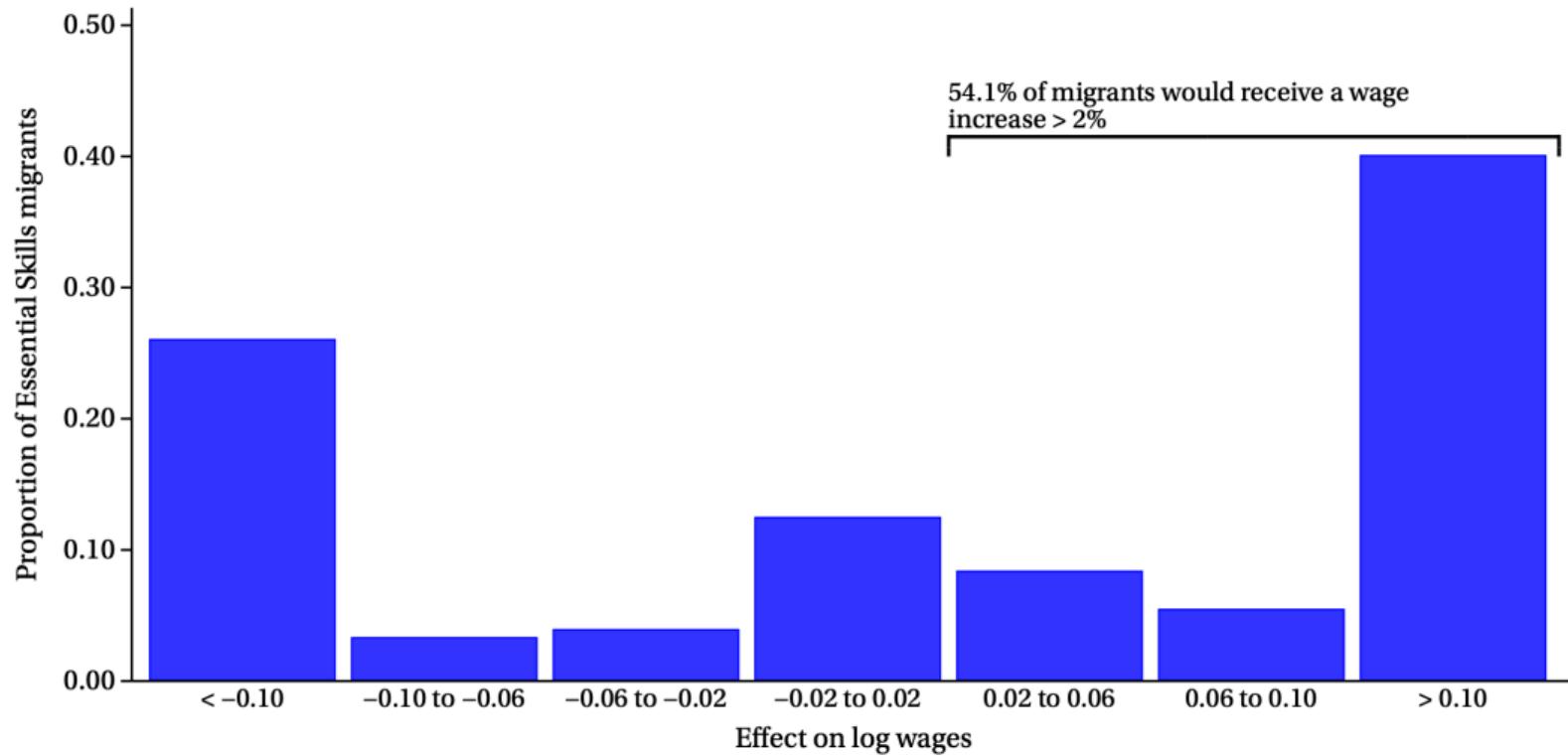


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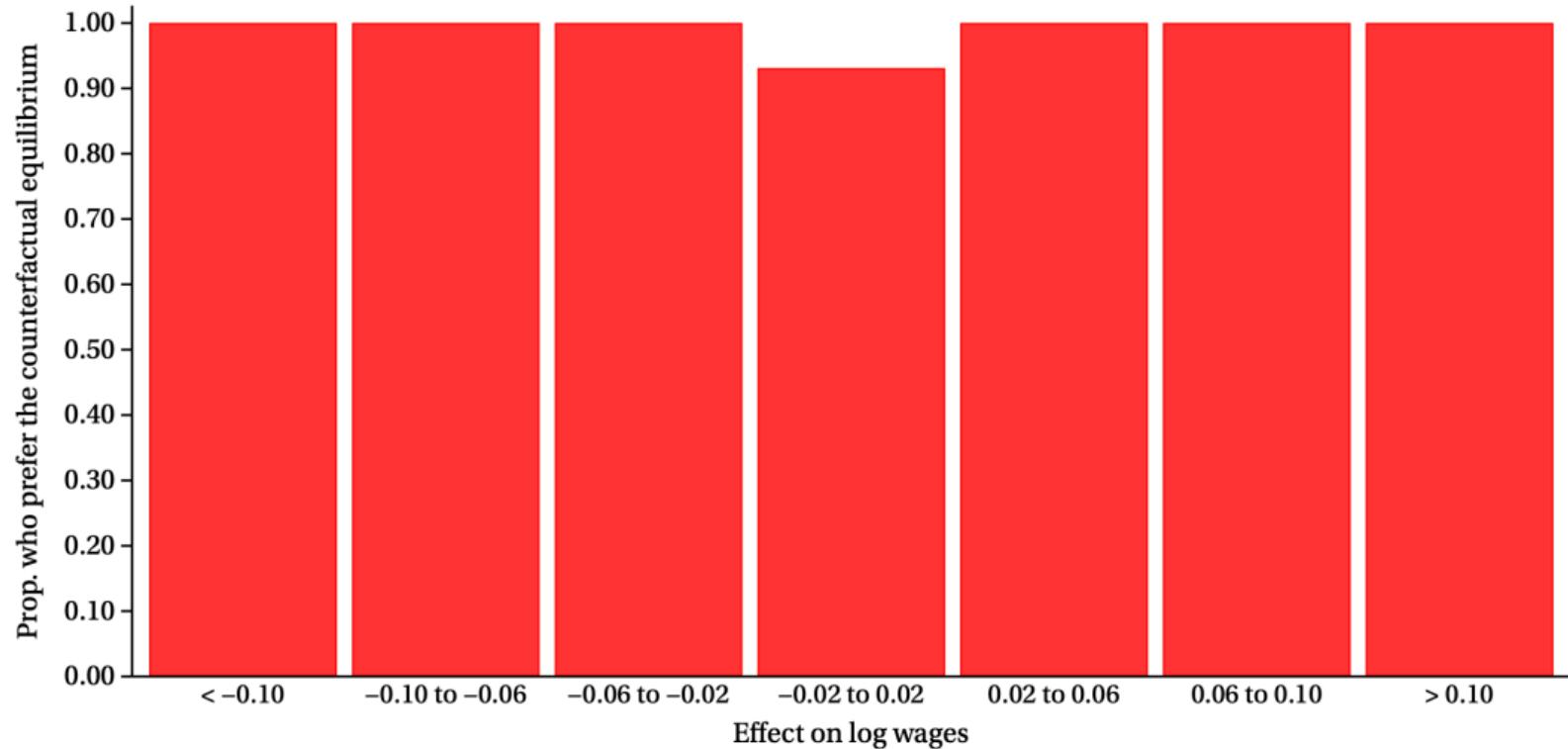
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Who are the migrants whose wages would increase or decrease?



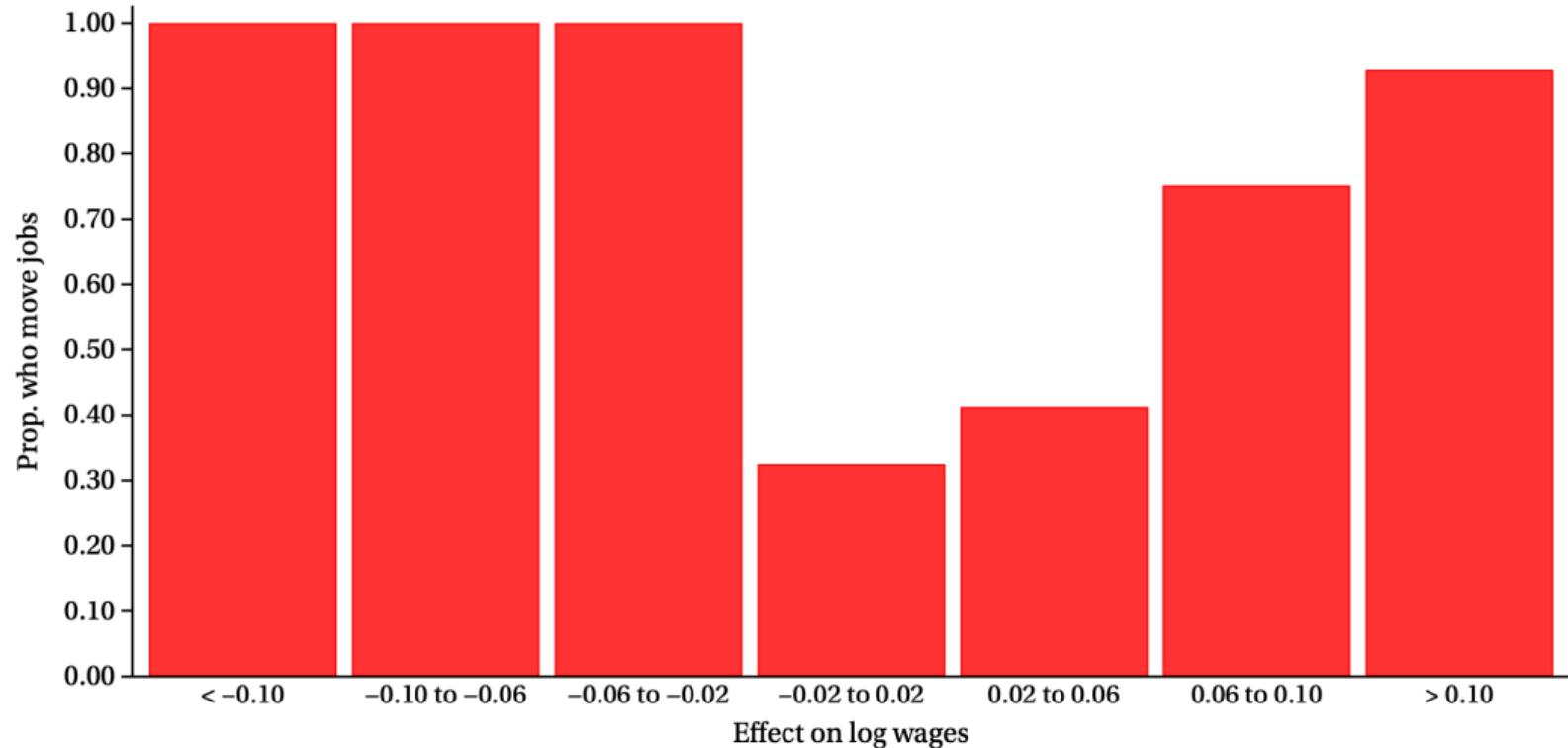
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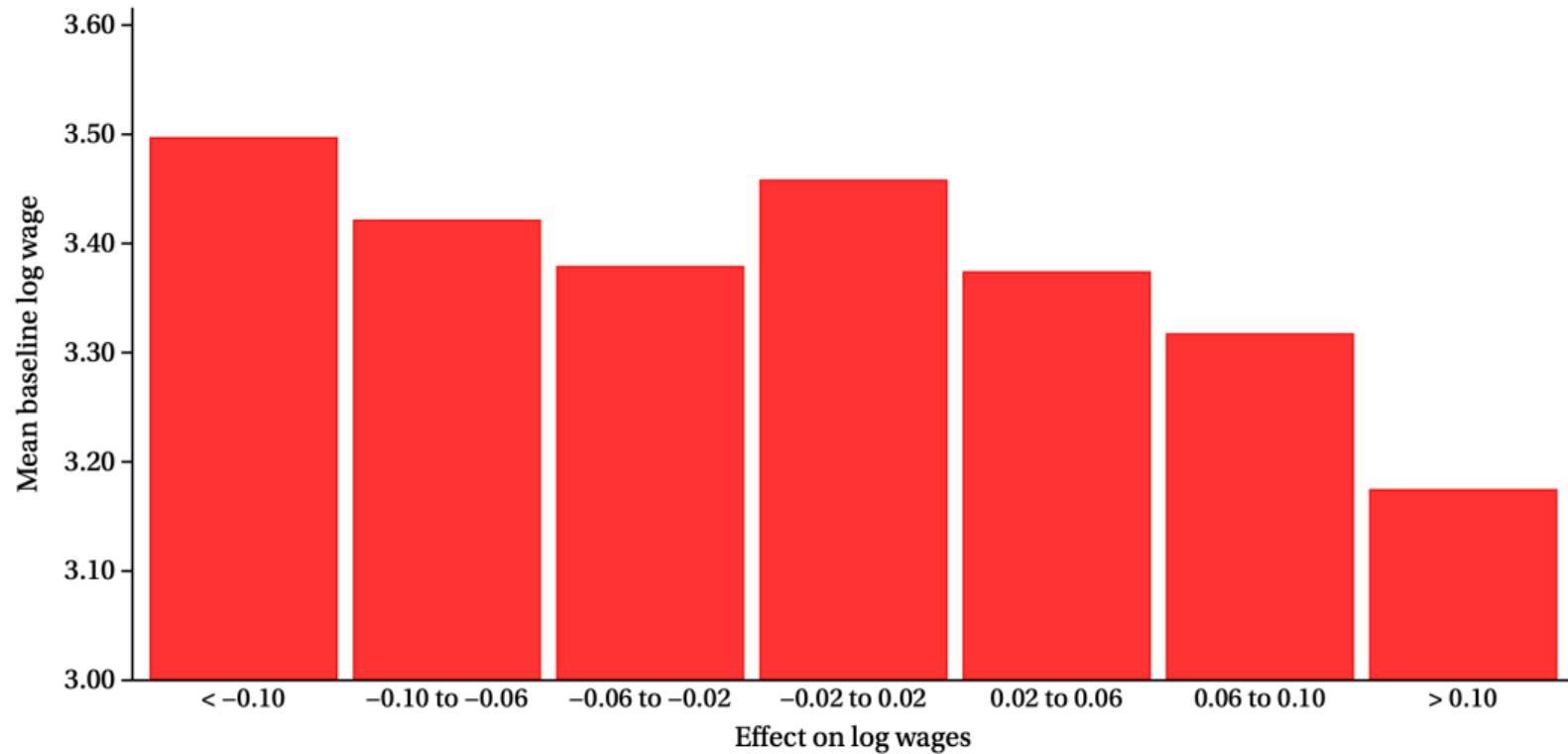
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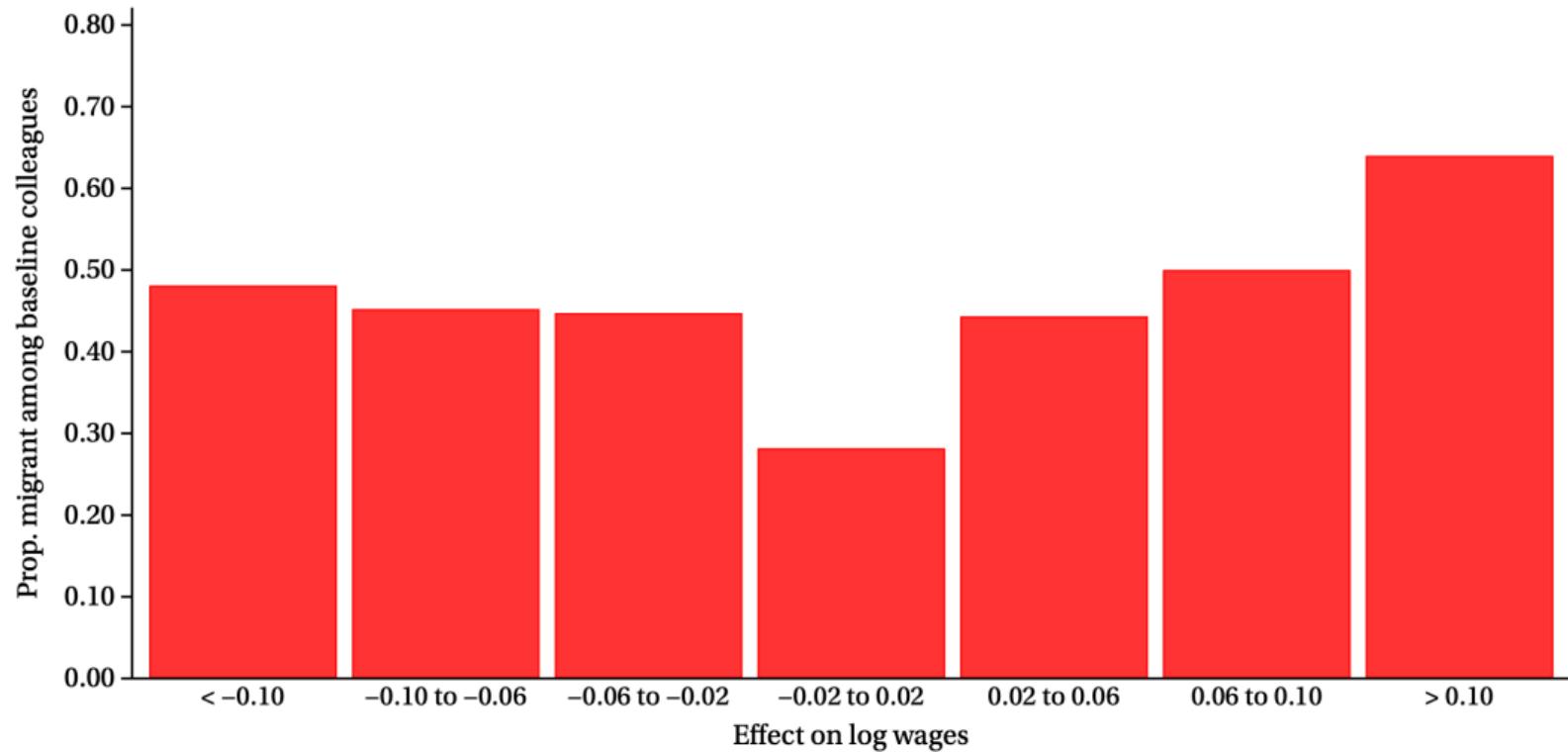
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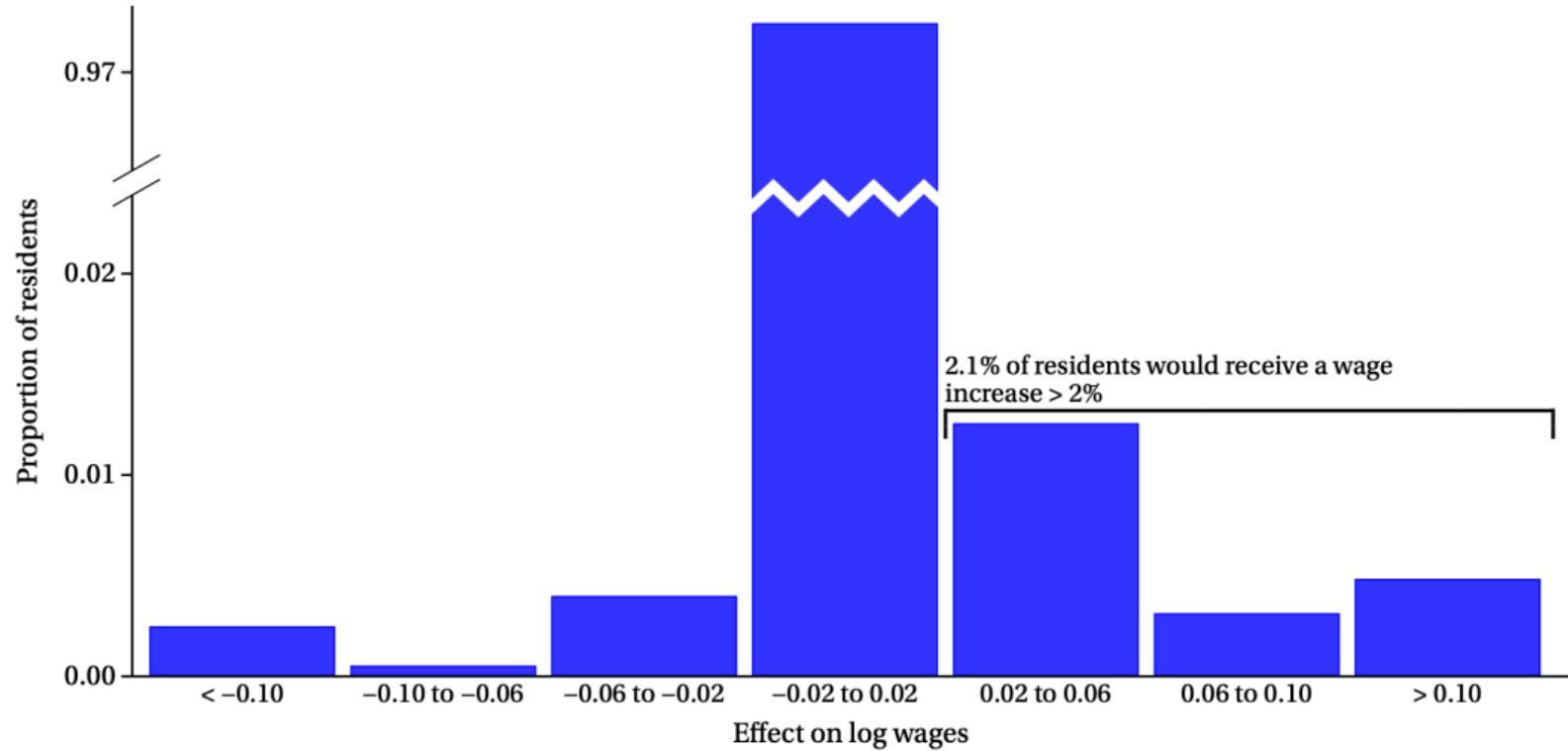
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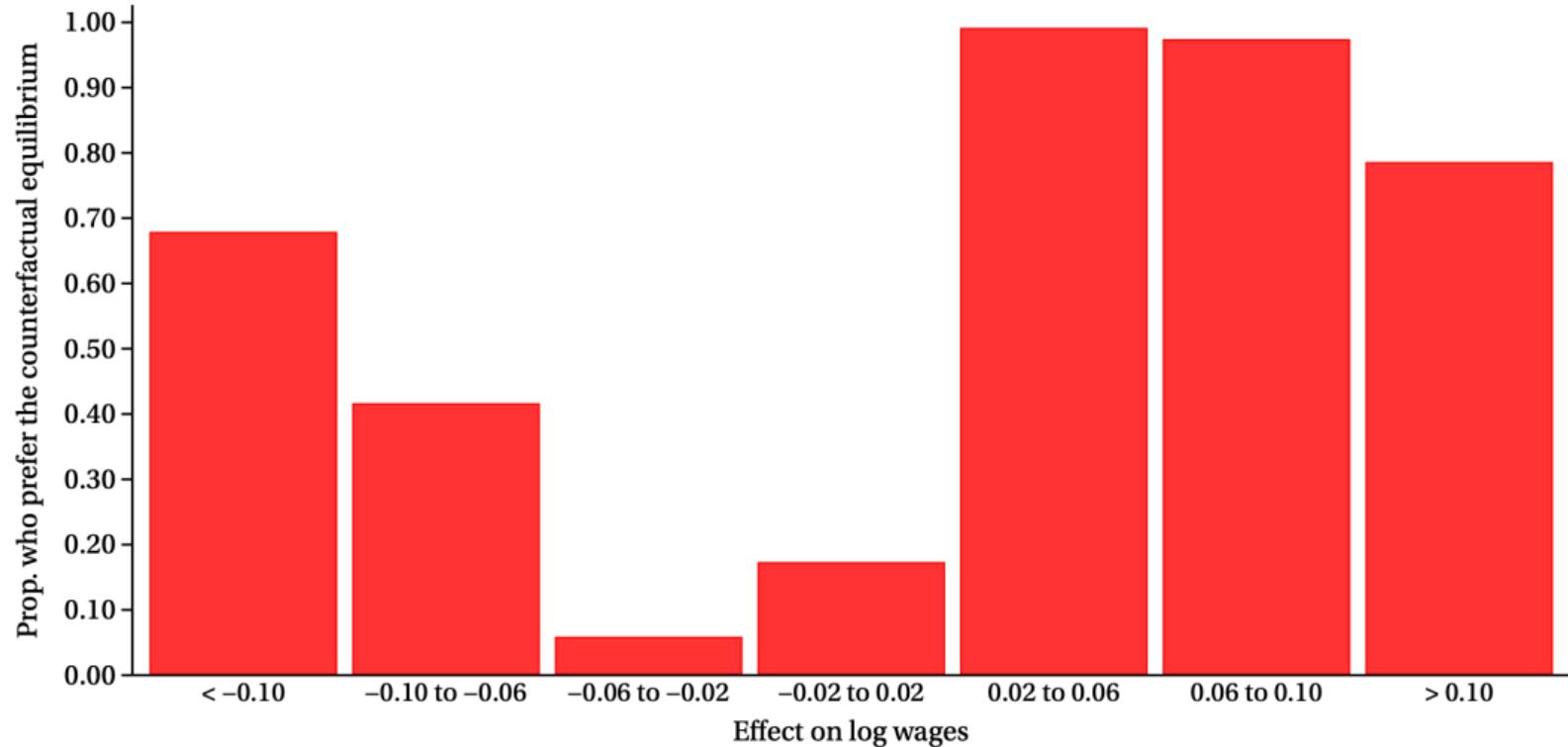
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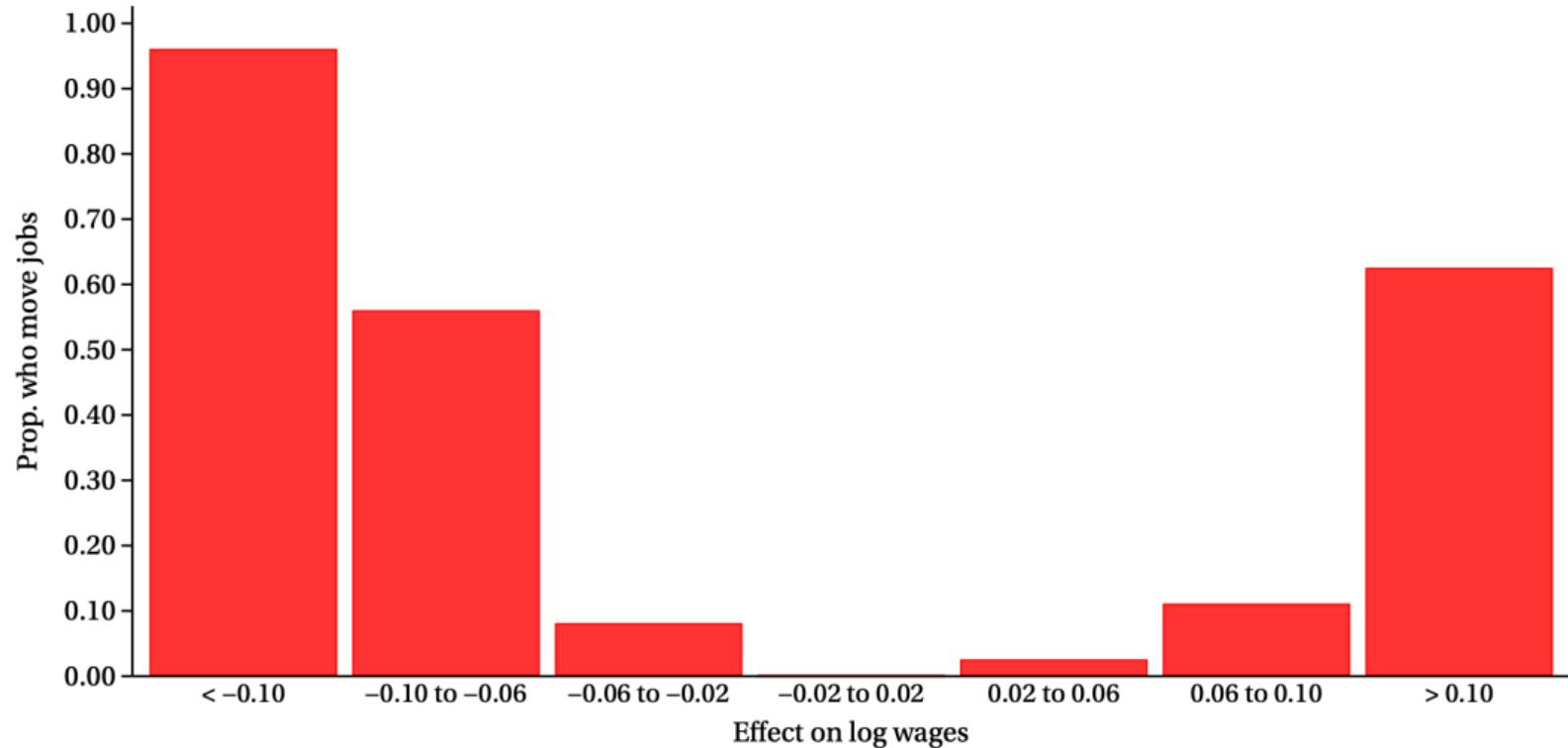
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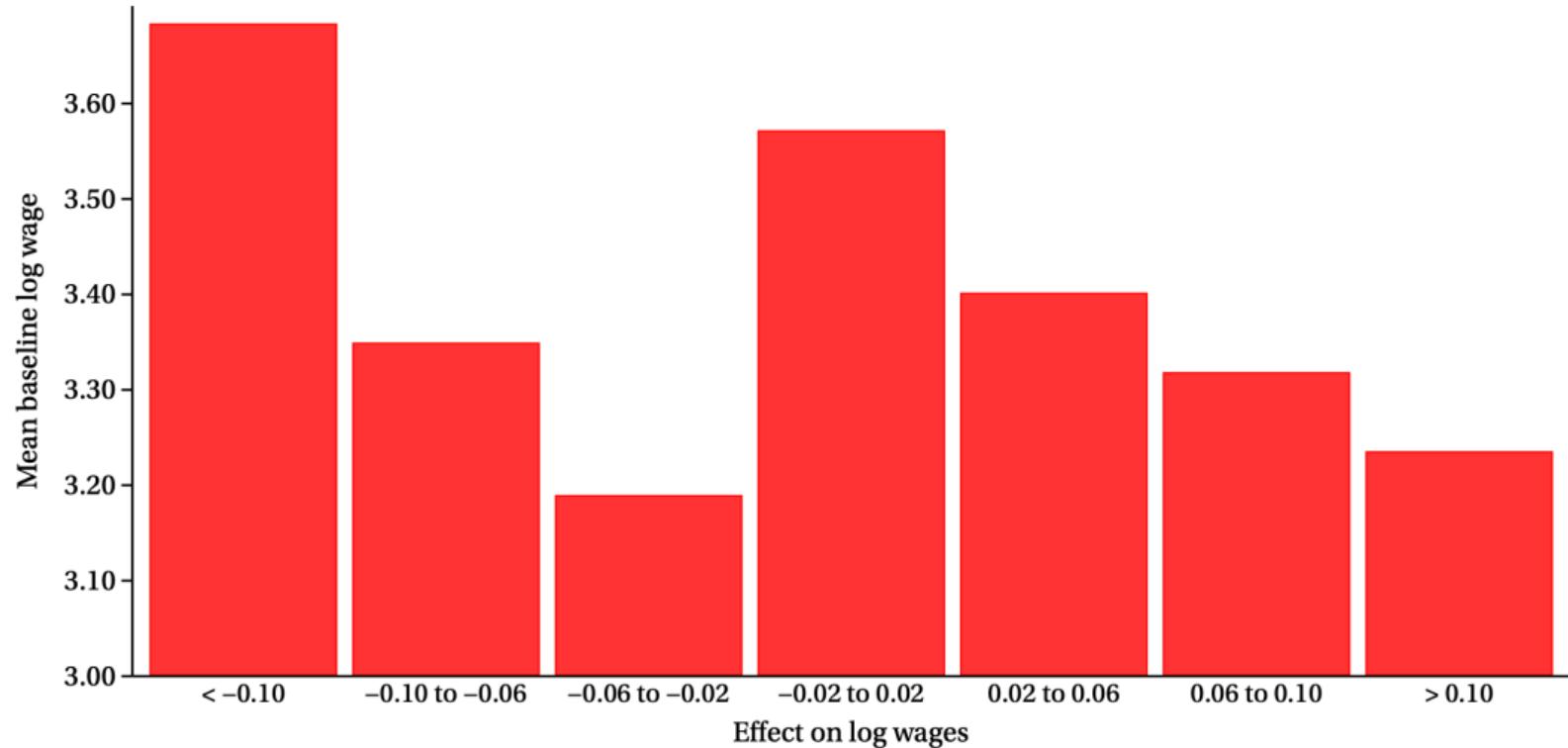
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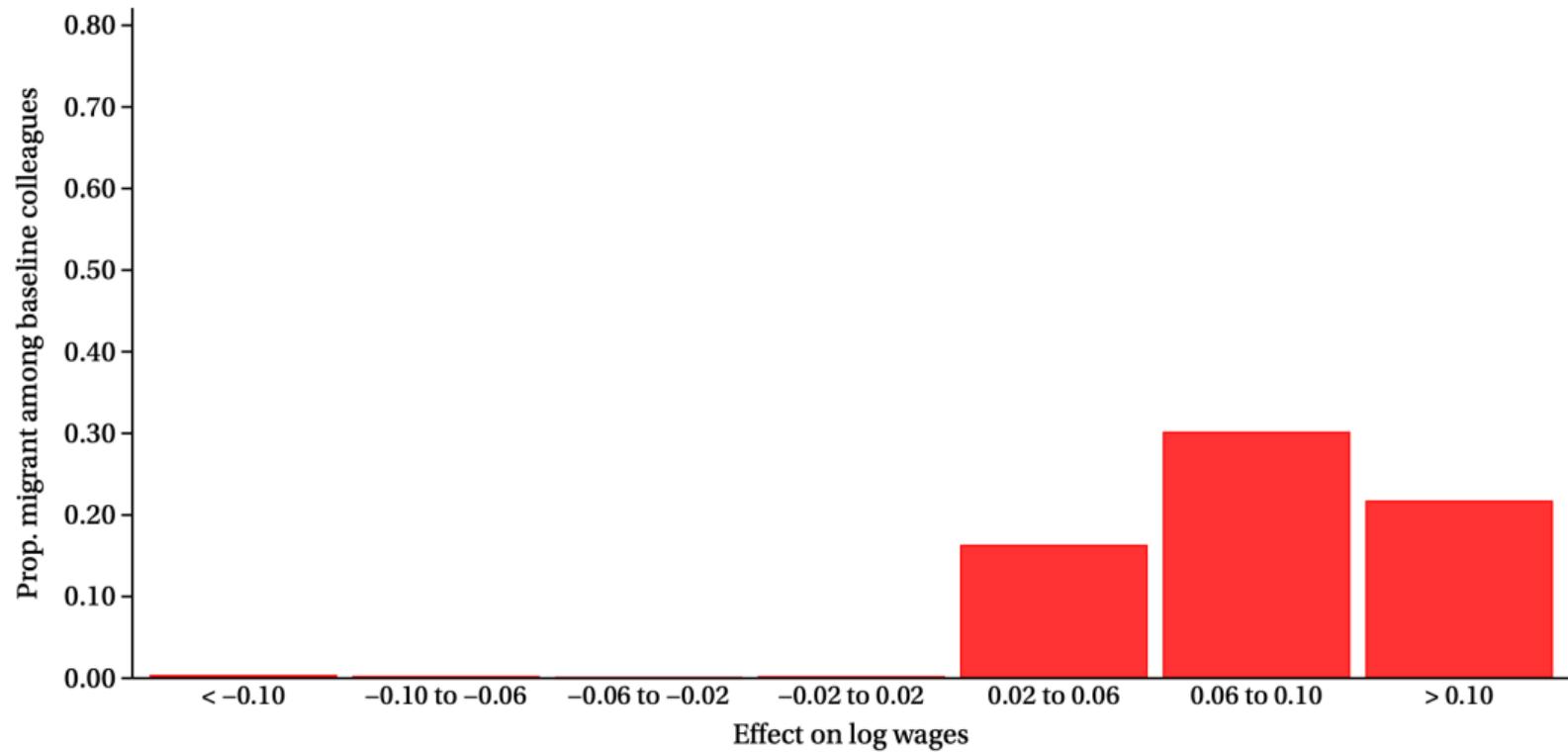
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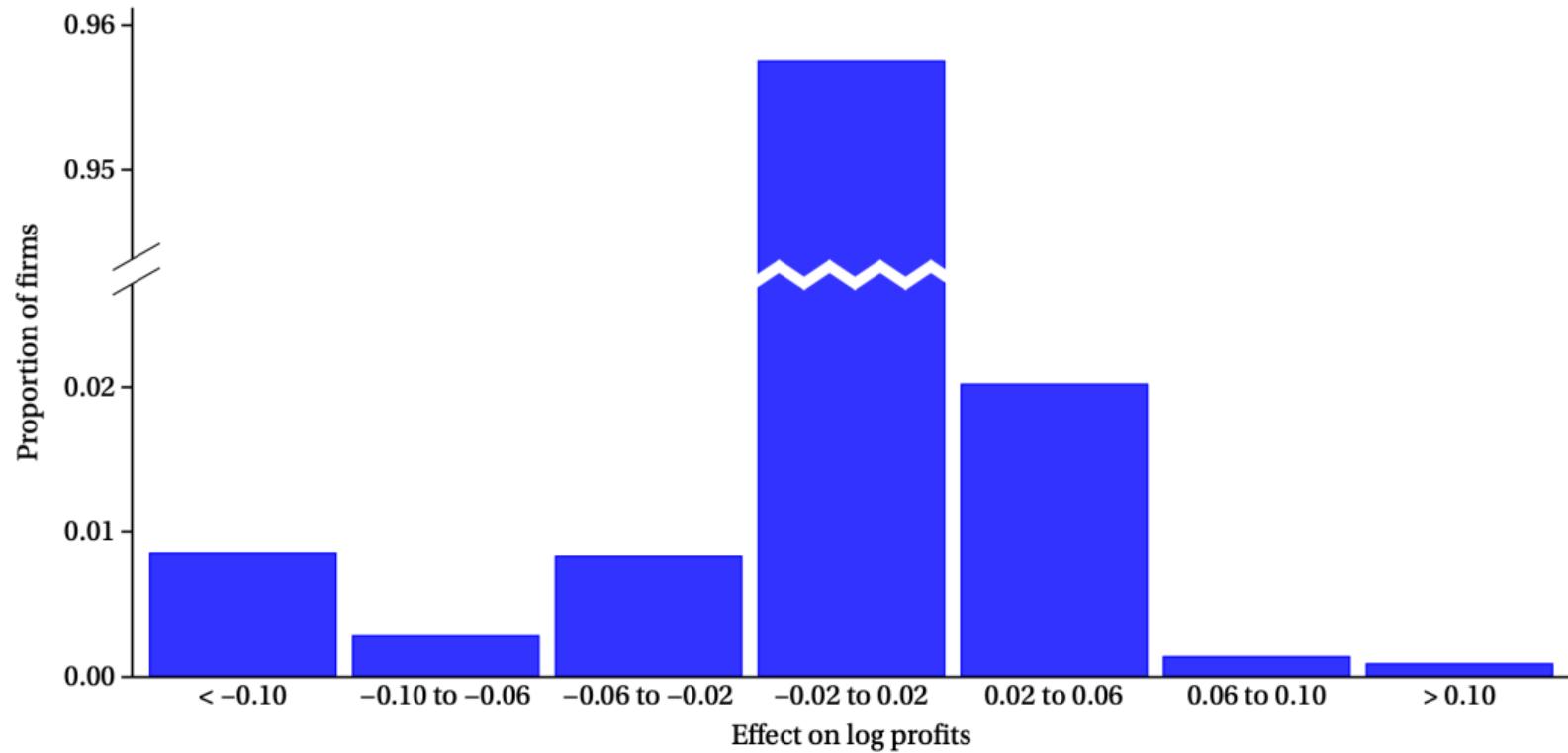
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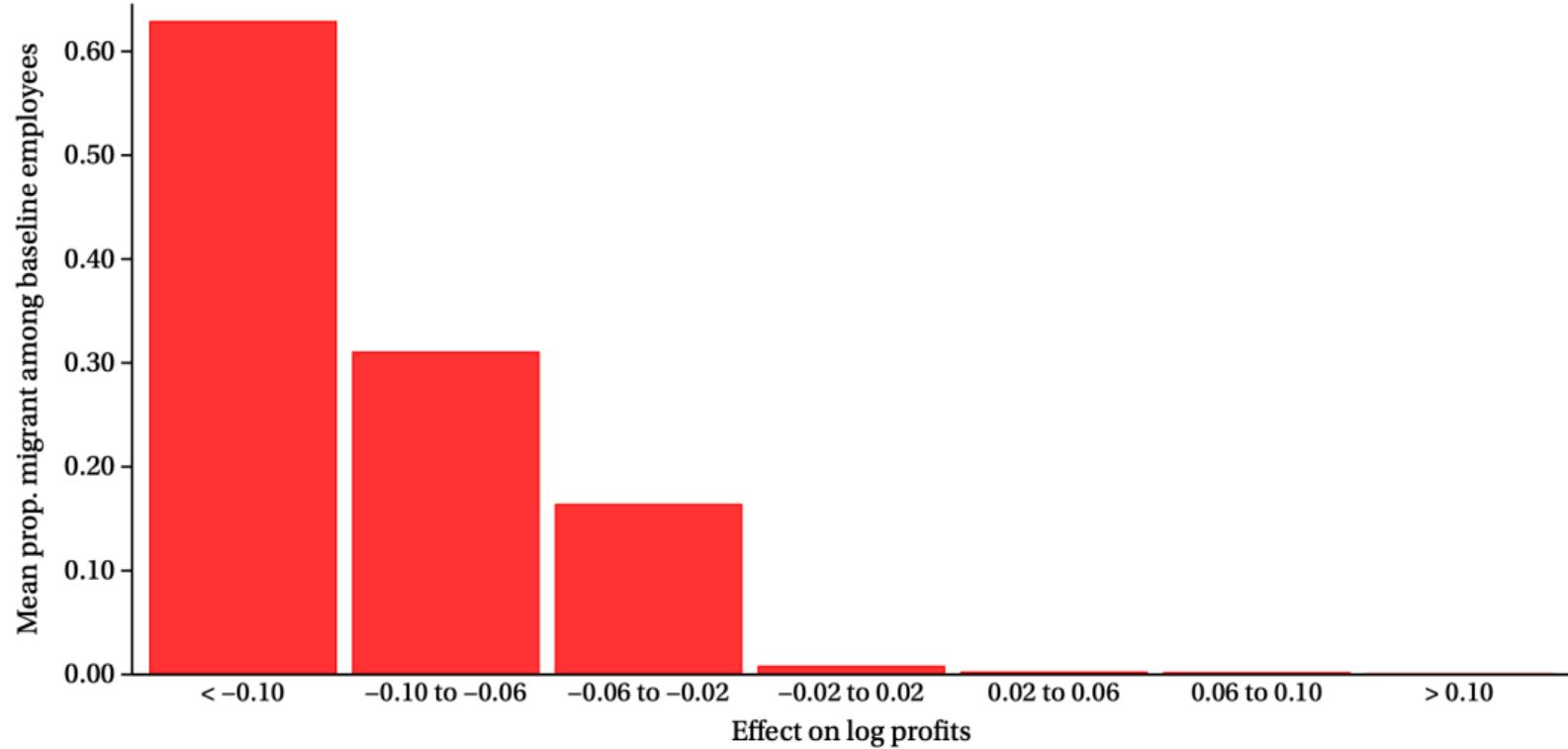
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Which are the firms whose profits would increase or decrease?



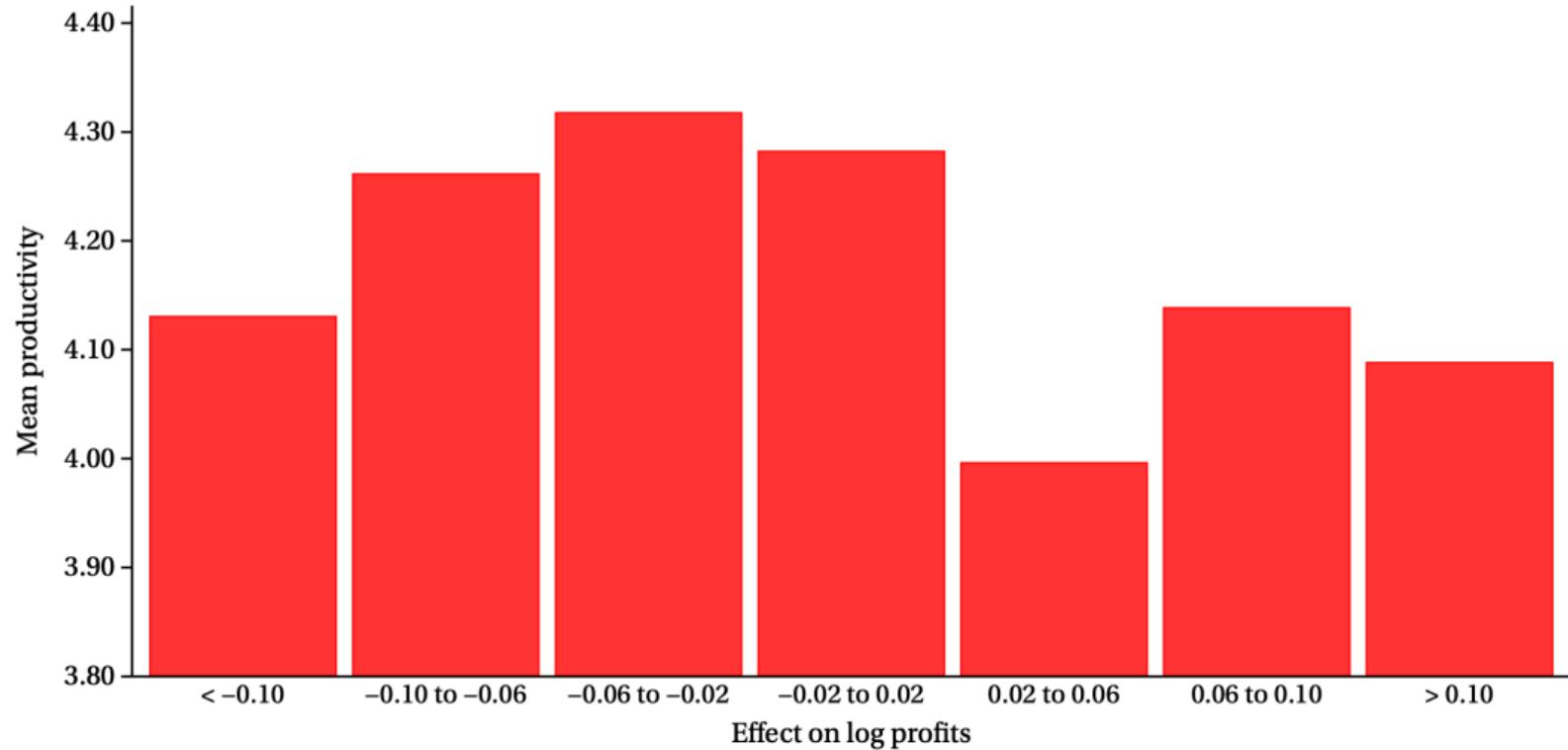
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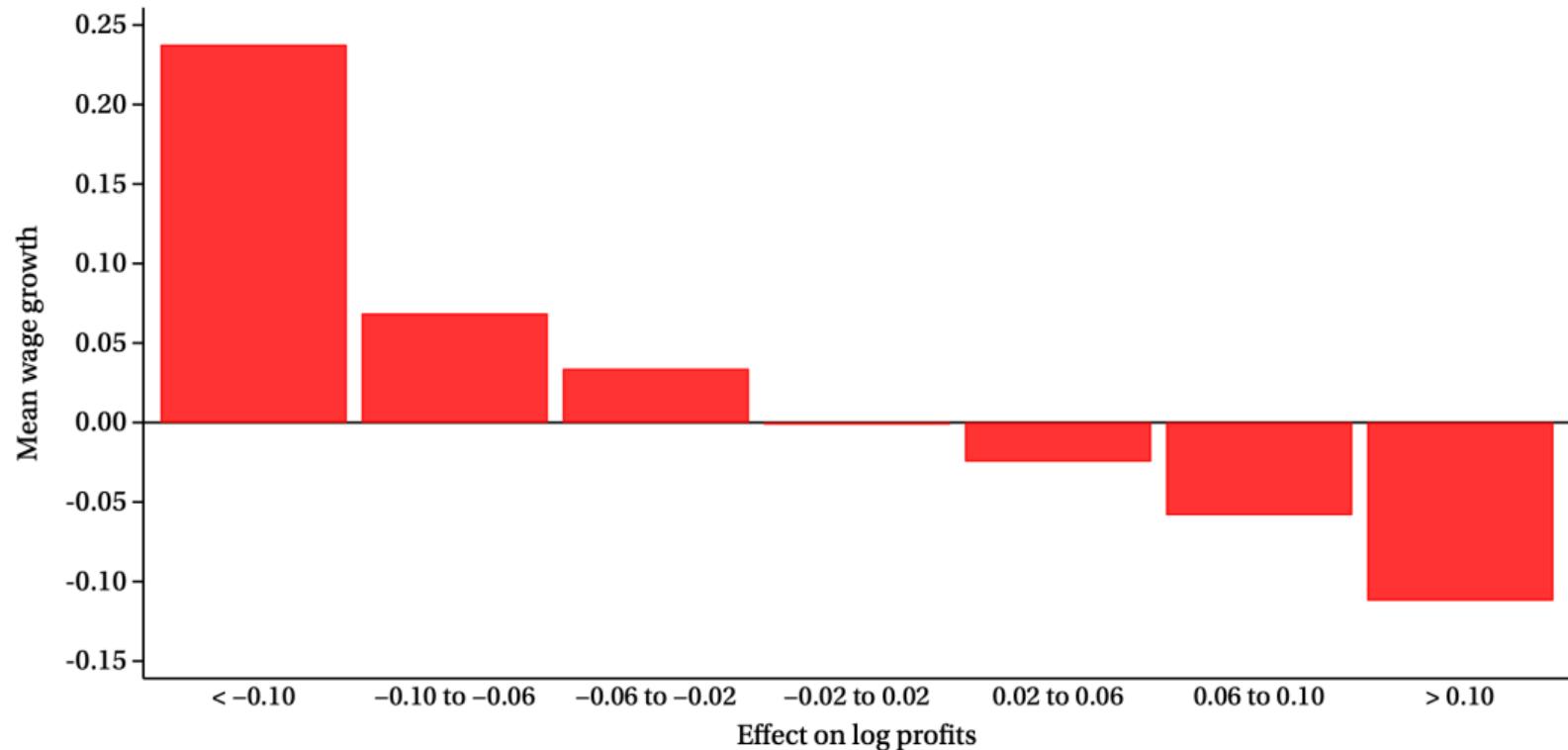
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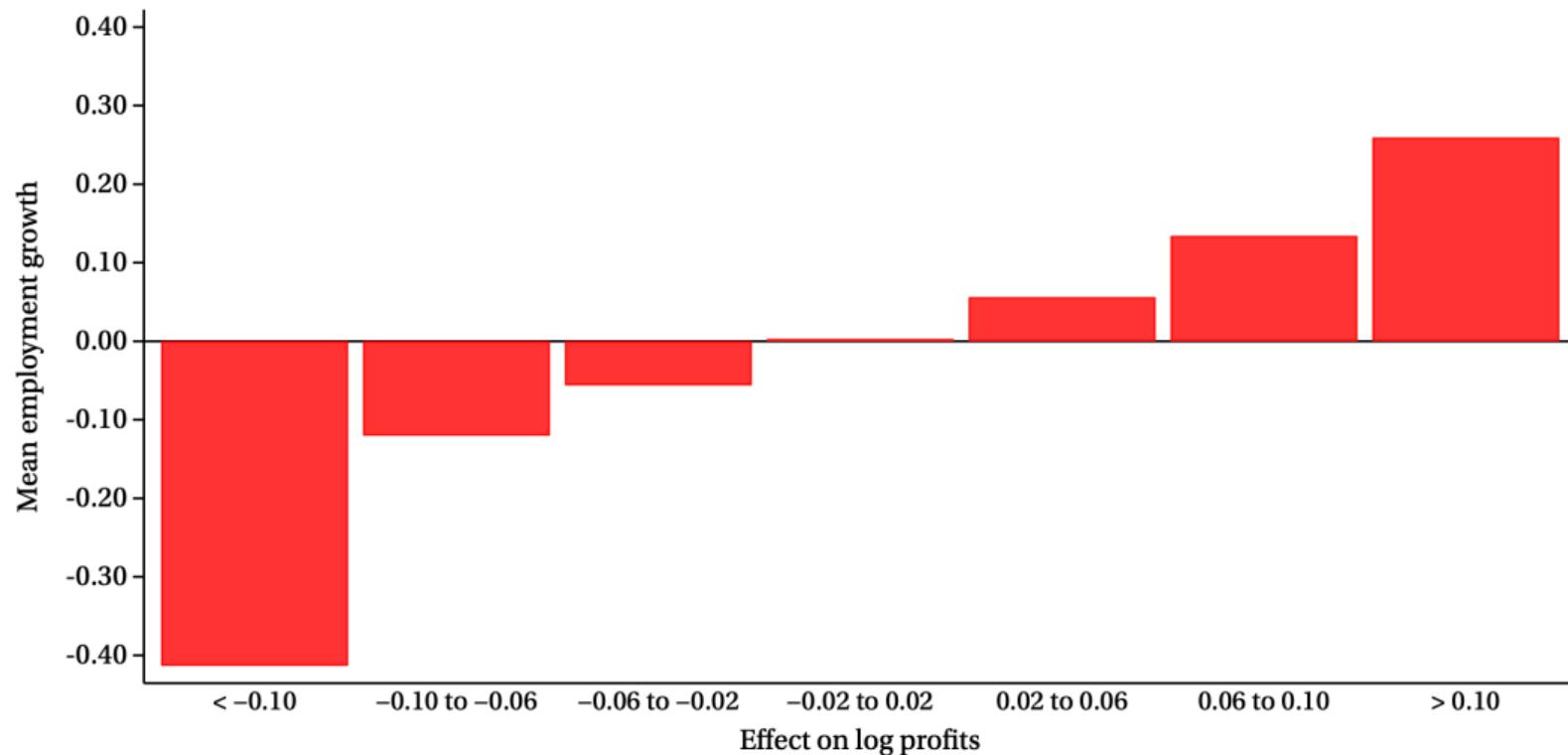
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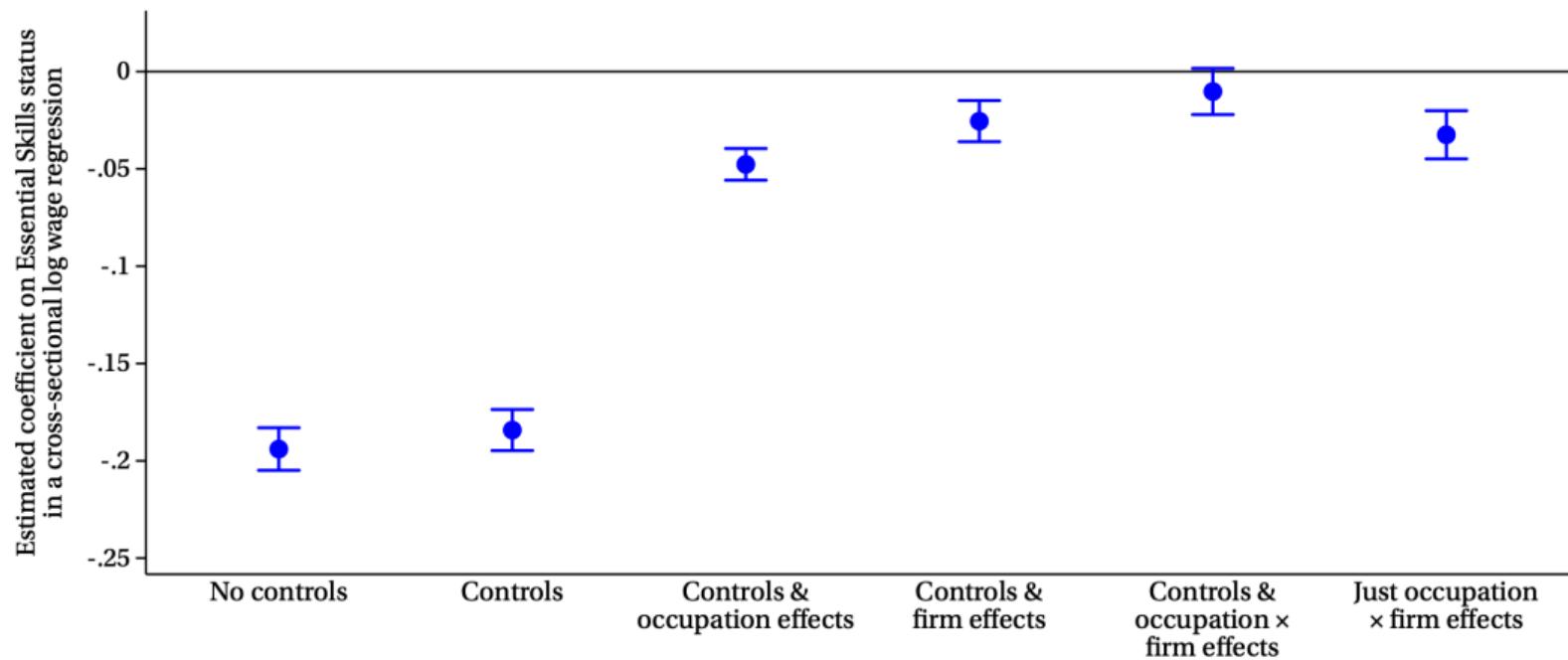
Back.

Which are the firms whose profits would increase or decrease?



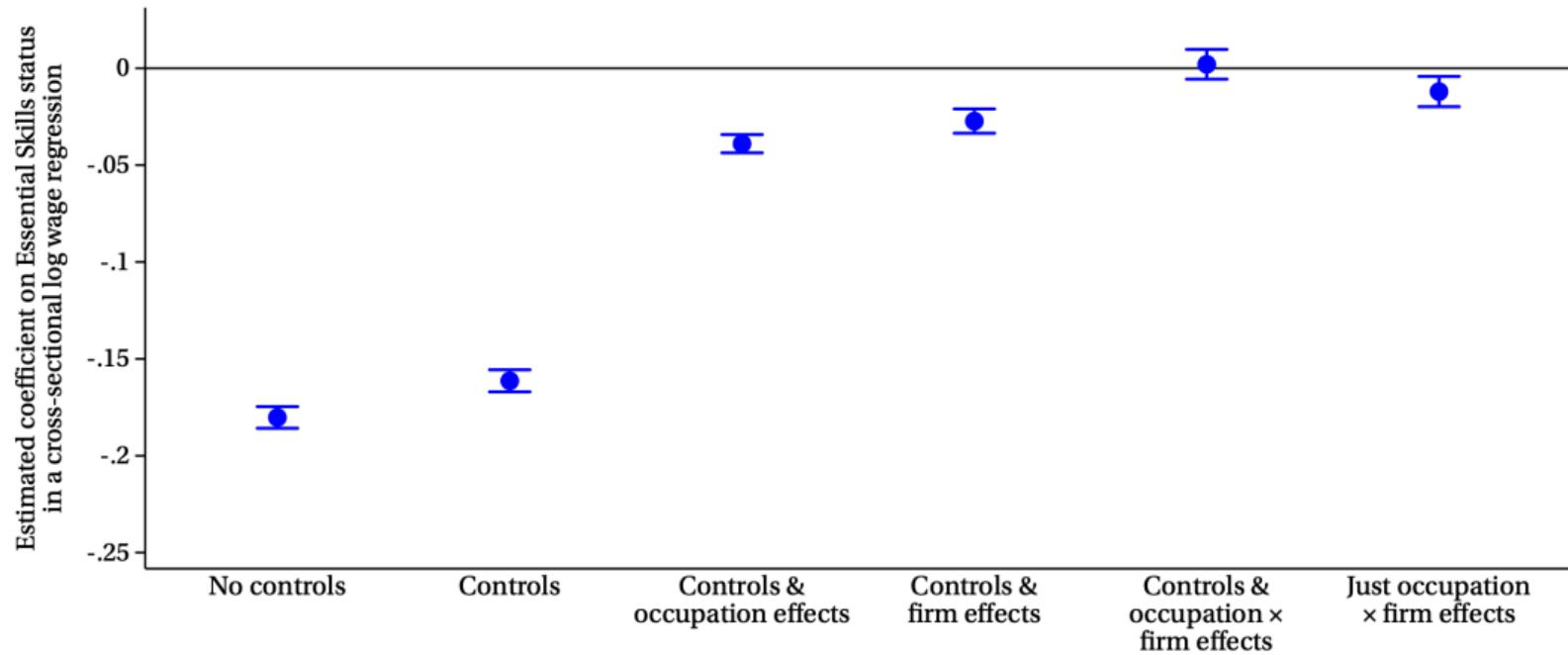
Back.

Migrants earn less than other workers... but only because of their jobs



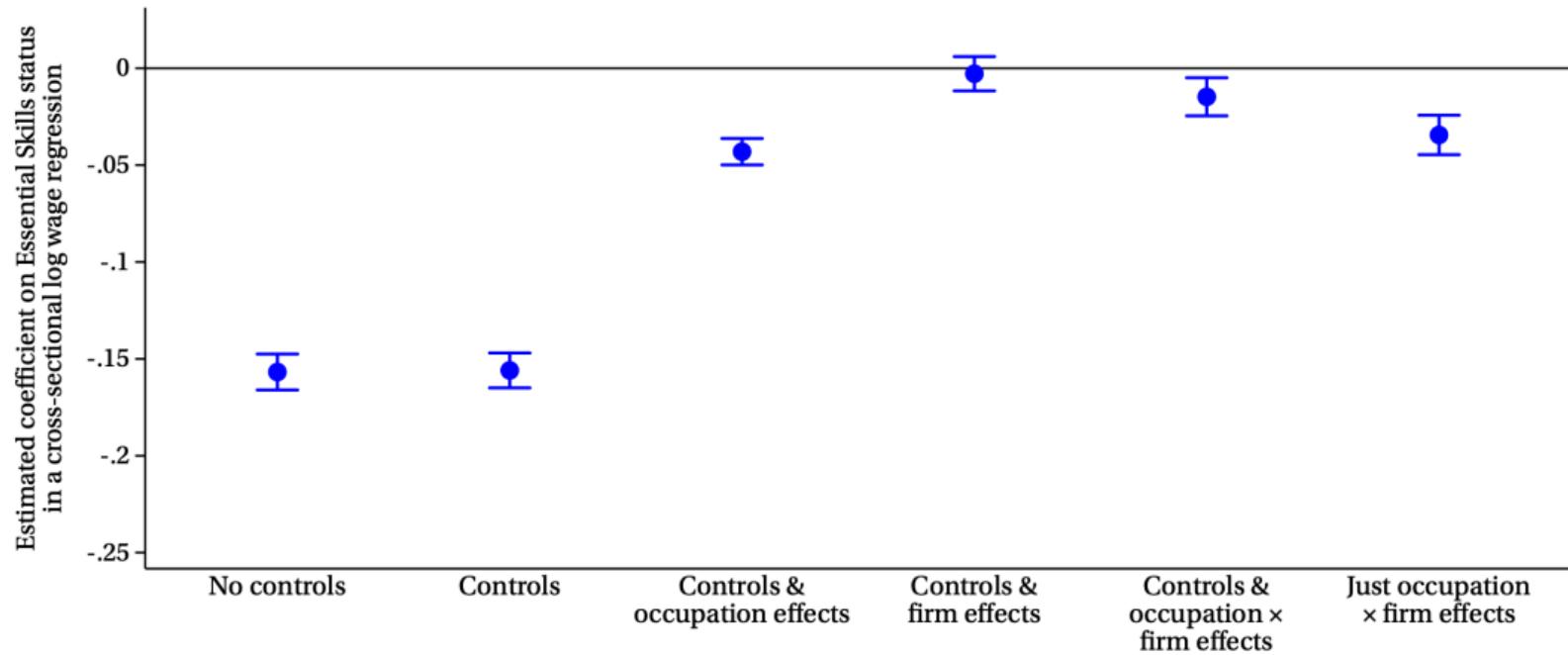
Sample: full-time employees with a unique firm in the 12 months before the 2013 Census. [Back](#).

Migrants earn less than other workers... but only because of their jobs



Sample: full-time employees with a unique firm in the 6 months before the 2018 Census. [Back](#).

Migrants earn less than other workers... but only because of their jobs



Sample: full-time employees with a unique firm in the 6 months before the 2013 Census. [Back](#).

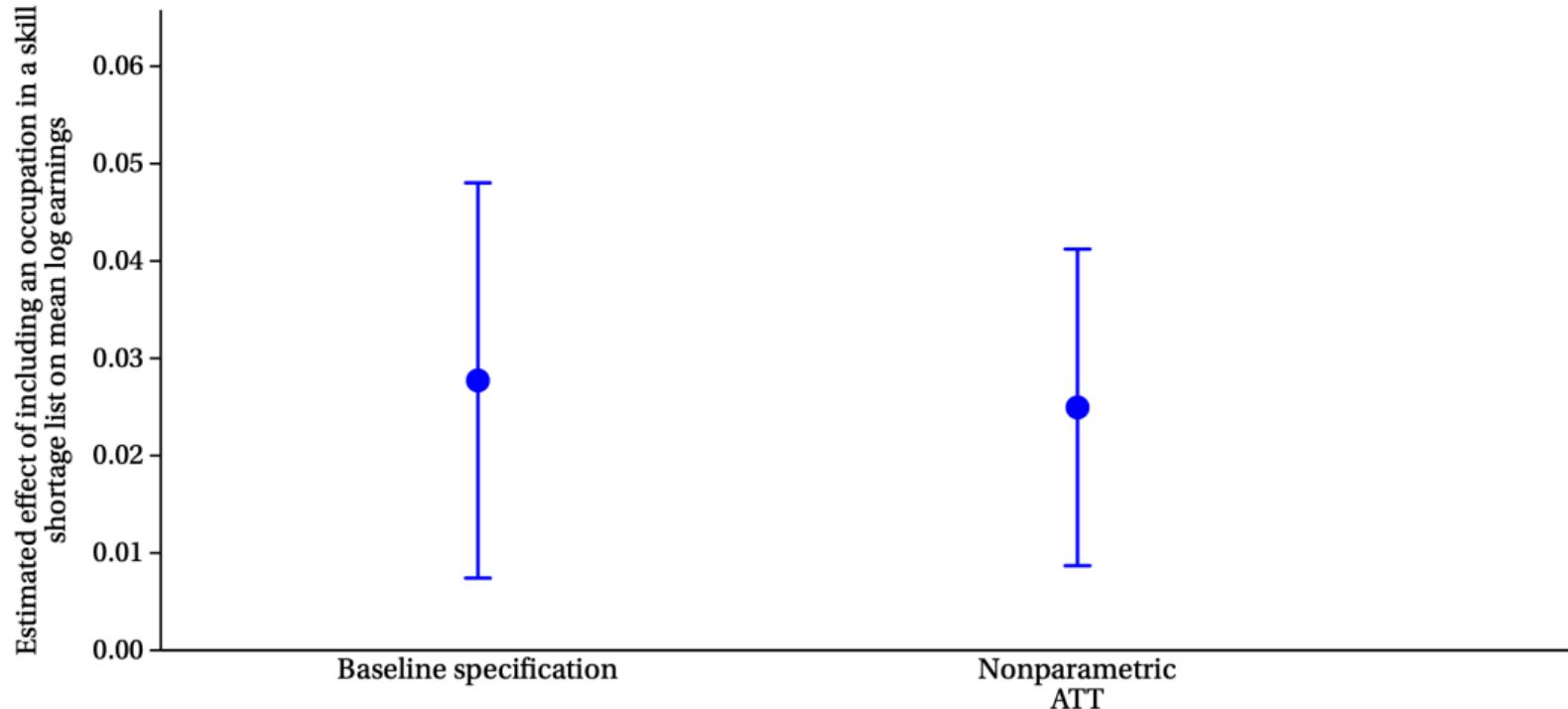
Accounting for treatment effect heterogeneity

A nonparametric ATT estimate *a la* de Chaisemartin & D'Haultfoeuille (2022).

- ▶ Assume that dynamic effects have stabilized after 2 years.
- ▶ Calculate a simple difference-in-difference estimate for each occupation added to or removed from a skill-shortage list, after 2 years of unchanged skill-shortage status, using each available control occupation and up to 36 post-treatment months.
- ▶ Average across control occupations, post-treatment months, and treatment occupations.

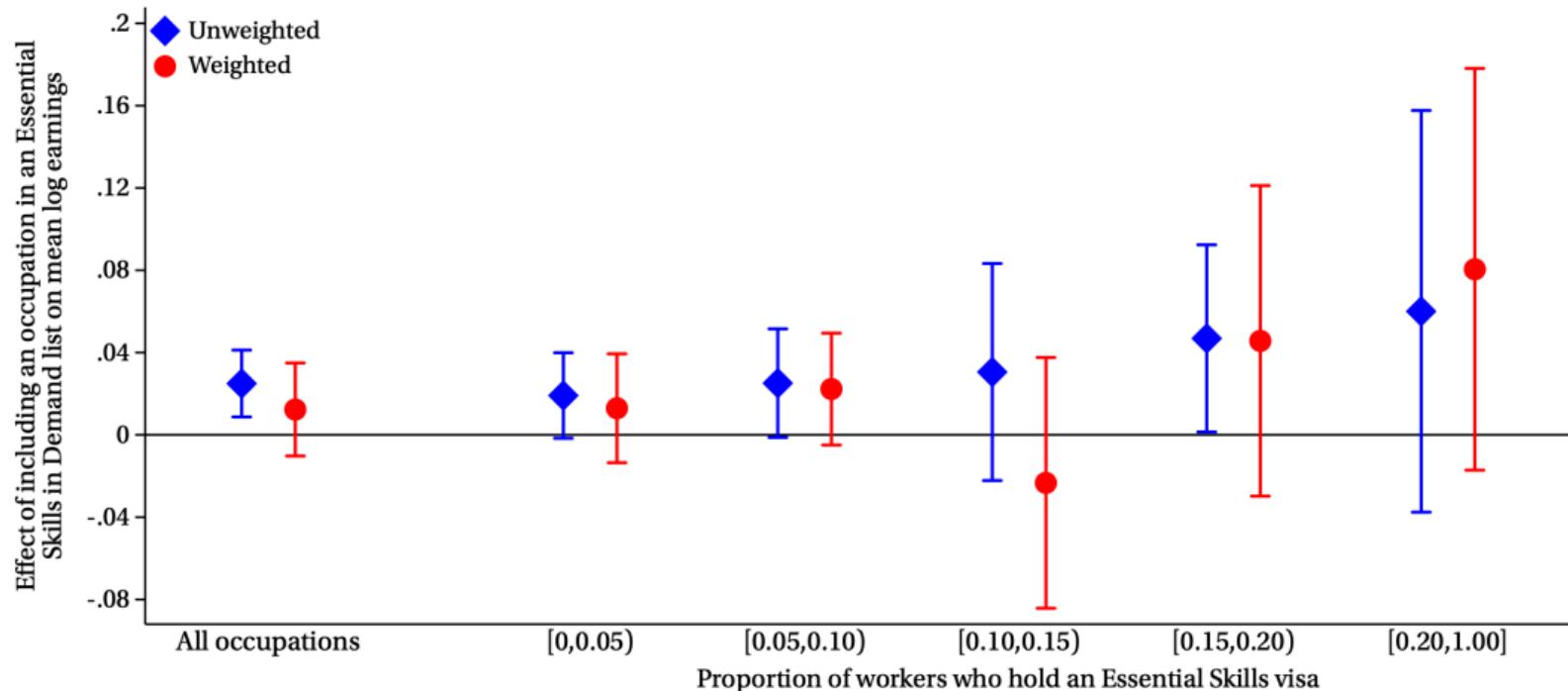
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Identifying effects for job-stayers

Consider the potential outcomes framework for worker i in period $t \in \{0, 1\}$:

$$Y_{it} = Y_{it}(W_i, M_i),$$

- ▶ Y_{it} denotes log earnings,
- ▶ W_i denotes being a lottery winner,
- ▶ M_i denotes moving jobs, which itself depends on the lottery: $M_i = M_i(W_i)$.

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Assume that the lottery, and the opportunity to switch jobs, occurs in period 1:

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$$\mathbb{E}[Y_{i1}(1, 0) - Y_{i1}(0, 0) | M_i(1) = 0].$$

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Assumption 1: The lottery is random: $W_i \perp M_i(\cdot), Y_{i0}(\cdot, \cdot), Y_{i1}(\cdot, \cdot)$.

Assumption 2: Lottery losers never switch jobs: $\forall i : M_i(0) = 0$.

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$$\therefore \underbrace{\mathbb{E}[Y_{i1}(1, 0) \mid M_i(1) = 0]}_{\text{Earnings for treated job-stayers}} = \underbrace{\mathbb{E}[Y_{i1} \mid M_i = 0, W_i = 1]}_{\text{This is observed}}.$$

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Remaining job: identify average counterfactual earnings for job-stayers:

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∴ We identify the DGP of $Y_{it}(0, 0)$ from lottery losers:

$$F_{Y_{i1}(0, 0)}(y_1 \mid Y_{i0}(0, 0) = y_0) = F_{Y_{i1}}(y_1 \mid Y_{i0} = y_0, W_i = 0)$$

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$$\therefore \mathbb{P}[M_i(1) = 1 \mid Y_{i0}(0,0)] = \mathbb{P}[M_i(1) = 1 \mid Y_{i0}(0,0), W_i = 1]$$

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Assumption 3: The job-switching decision is Markovian in counterfactual earnings:

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With earlier results, we have the Fredholm integral equation of the first kind:

$$\begin{aligned} \mathbb{P}[M_i(1) = 1 \mid Y_{i0} = y_0, W_i = 1] \\ = \int \mathbb{P}[M_i(1) = 1 \mid Y_{i1}(0, 0) = y_1] dF_{Y_{i1}(0, 0)}(y_1 \mid Y_{i0} = y_0, W_i = 0). \end{aligned}$$

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Under *regularity and completeness conditions*, the integral equation can be solved for

$$\mathbb{P}[M_i(1) = 1 \mid Y_{i1}(0, 0) = y_1].$$

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Remaining job: identify $\mathbb{E}[Y_{i1}(0,0) \mid W_i(1) = 0]$.

Having identified $\mathbb{P}[M_i(1) = 1 \mid Y_{i1}(0,0) = y_1]$, we can find $dF_{Y_{i1}(0,0)}(y_1 \mid M_i(1) = 1)$ using Bayes' rule:

$$dF_{Y_{i1}(0,0)}(y_1 \mid M_i(1) = 1) = \mathbb{P}[M_i(1) = 1 \mid Y_{i1}(0,0) = y_1] \frac{dF_{Y_{i1}(0,0)}(y_1)}{\mathbb{P}[M_i(1) = 1]}$$

where

- ▶ $dF_{Y_{i1}(0,0)}$ can be identified from lottery losers;
- ▶ $\mathbb{P}[M_i(1) = 1]$ can be identified from lottery winners.

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Proposition: The average treatment effect for job-stayers is identified from the joint distribution of earnings, moving indicators and lottery-winning indicators.

Back.

Estimating effects for job-stayers

We hope to estimate the **average treatment effect for job-stayers**:

$$\mathbb{E}[Y_{it}(1,0) - Y_{it}(0,0) \mid M_{it}(1) = 0].$$

Our setting has additional complications:

- ▶ Our data spans more than 2 periods.
- ▶ Winning the lottery is only random *conditional on the lottery entered*.
- ▶ Lottery winners can switch firms.

Estimating effects for job-stayers

We hope to estimate the **average treatment effect for job-stayers**:

$$\mathbb{E}[Y_{it}(1,0) - Y_{it}(0,0) \mid M_{it}(1) = 0].$$

We impose the monotonicity assumption: $\forall i, t : M_{it}(1) \geq M_{it}(0)$.

- We observe $\mathbb{E}[Y_{it}(1,0) \mid M_{it}(1) = 0] = \mathbb{E}[Y_{it} \mid W_{it} = 1, M_{it} = 0]$.

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We hope to estimate the **average treatment effect for job-stayers**:

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Let $b(i)$ denote i 's lottery. We estimate $\mathbb{E}[Y_{it}(0,0) \mid S_{it}(1) = 0]$ assuming:

- Job-moving is probit in $W_i, Y_{i,t}(0,0)$, with a lottery effect:

$$\mathbb{P}[M_{i,t} = 1 \mid W_i, (Y_{i,s}(0,0))_{s \leq t}, M_{i,t-1} = 0] = \Phi(\beta_1 W_i + \beta_2 Y_{i,t}(0,0) + \beta_3 W_i Y_{i,t}(0,0) + \delta_{b(i),t});$$

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- Counterfactual earnings $Y_{i,t}(0,0)$ are AR(1), with another lottery effect:

$$Y_{i,t}(0,0) = \alpha Y_{i,t-1}(0,0) + \gamma_{b(i),t} + \epsilon_{it}; \quad \epsilon_{it} \sim N(0, \sigma_\epsilon);$$

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- Lottery effects are joint-normal:

$$\delta_b, \gamma_b \sim N(\mu_{\delta\gamma}, \Sigma).$$

Estimating effects for job-stayers

We hope to estimate the **average treatment effect for job-stayers**:

$$\mathbb{E}[Y_{it}(1,0) - Y_{it}(0,0) \mid M_{it}(1) = 0].$$

We calculate our estimator using a Gibbs sampler algorithm with state

$$[\{m_{it}\}, \{Y_{i,t}(0,0)\}, \beta, \delta, \alpha, \gamma, \sigma_\delta, \sigma_\gamma, \sigma_\epsilon],$$

where m_{it} is the Probit latent variable:

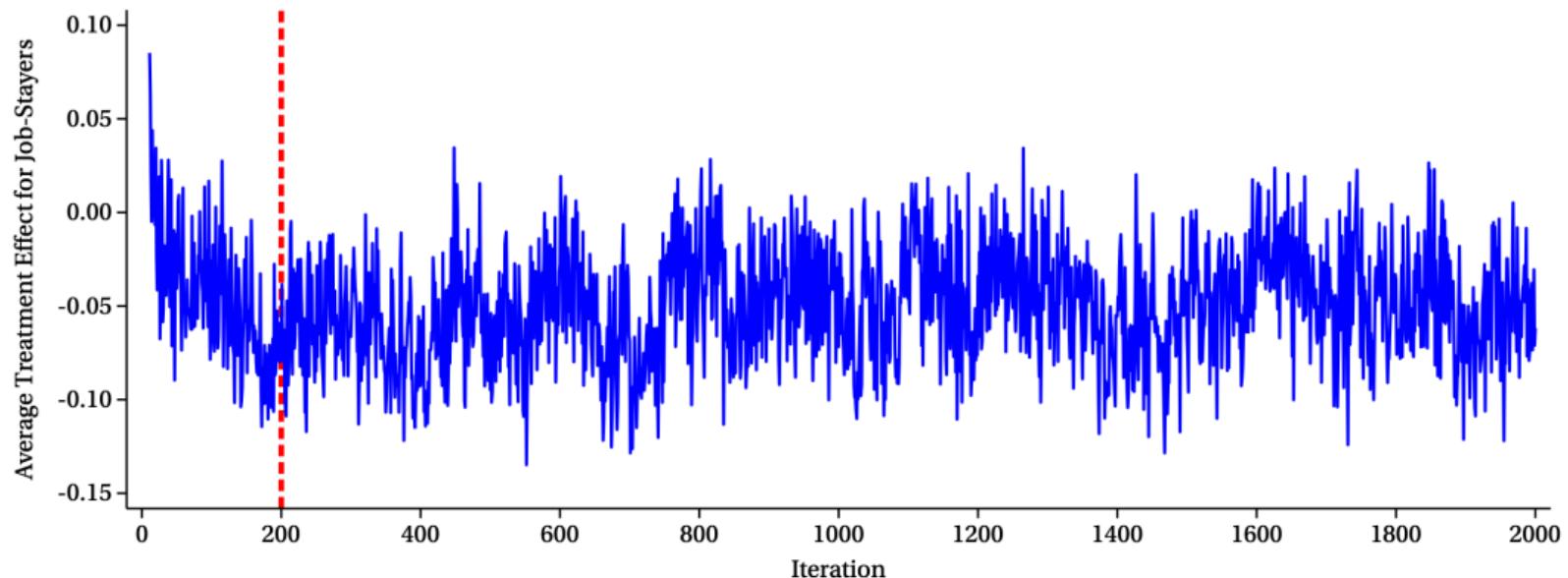
$$M_{it} = 1 \iff m_{it} \geq 0.$$

We impose uninformative conjugate priors on the hyper-parameters.

Estimating effects for job-stayers

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Note: the first ten iterations are omitted from the plot. [Back](#).

A novel algorithm for nested extreme value random variables

We are interested in drawing random variables with CDF:

$$\mathbb{P} \left[(X_j)_{j \in \mathbf{J}} \leq (x_j)_{j \in \mathbf{J}} \right] = \exp \left(\sum_{C \in \mathbf{C}} \left(\sum_{j \in C} \exp \left(\frac{-x_j}{\lambda} \right) \right)^{\lambda} \right),$$

where \mathbf{C} is a partition on \mathbf{J} . Both \mathbf{J} and \mathbf{C} may be large.

Per Cardell '97: $X_j = \zeta_{C(j)} + (1 - \sigma)\epsilon_j$, where $C(j)$ is the nest containing j , and both X_j and ϵ_j have marginal standard-Gumbel distributions.

By the Convolution Theorem, ζ_C has characteristic function

$$\phi_{\zeta}(t) = \frac{\Gamma(1 - i t)}{\Gamma(1 - i(1 - \sigma)t)}, \text{ where } \Gamma \text{ is the gamma function and } i \text{ is the imaginary unit.}$$

By the Fourier inversion theorem, the pdf of ζ is given by $f_{\zeta}(x) = \frac{1}{2\pi} \int_0^{\infty} e^{-tx} \phi_{\zeta}(t) dt$.

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where \mathbf{C} is a partition on \mathbf{J} . Both \mathbf{J} and \mathbf{C} may be large.

Algorithm:

1. Approximate $f_\zeta(x) = \frac{1}{2\pi} \int_0^\infty e^{-tx} \phi_\zeta(t) dt$ numerically.
2. Draw ζ_C using inverse transform sampling.
3. Draw ϵ_j from a standard Gumbel distribution.
4. Form $X_j = \zeta_{C(j)} + (1 - \sigma)\epsilon_j$.

Back.

Explicit forms for structural residuals

Supply elasticities

$$\eta_{f,o,x,t}^{status} = \tau \left(1 - (1 - \lambda) \sigma_{f,o,x,t|c_f}^{status} - \lambda \sigma_{f,o,x,t}^{status} \right).$$

$$\eta_{f,o,t} = \sum_{x \in \mathbf{X}} \left(\sigma_{x,resident|f,o,t} \eta_{f,o,x,t}^{resident} + \sigma_{x,migrant|f,o,t} \eta_{f,o,x,t}^{migrant} \right).$$

The amenity value of employment.

$$\tilde{\xi}_{f,o,x,t} = \log \sigma_{f,o,x,t|c_f}^{resident} - \tau \log(w_{f,o,t}) + D_{o,c_f,x,t}.$$

The amenity value of living in a location.

$$\bar{\xi}_{c,o,x,t} = \frac{1}{\lambda} \log \sigma_{c,o,x,t}^{resident} - \log \left(\sum_{f \in \mathbf{F}_{c,o,t}} \exp(\tau \log(w_{f,o,t}) + \xi_{f,o,x,t}) \right) + D_{o,x,t}.$$

Back.

Explicit forms for structural residuals

Labor-augmenting productivity.

$$\begin{aligned}\phi_{f,o,t} = & \log w_{f,o,t} + \log \left(1 + \frac{1}{\eta_{f,o,t}} \right) + (1 - \rho) \log L_{o,f,t} - \left(\frac{\rho}{\nu} \right) \log \nu \\ & + \left(\frac{\rho - \nu}{\nu} \right) \log \left(\sum_{o' \in \mathbf{O}_{f,t}} \exp \left(\log w_{f,o',t} + \log \left(1 + \frac{1}{\eta_{f,o',t}} \right) + \log L_{o',f,t} \right) \right).\end{aligned}\quad (1)$$

Back.

Explicit forms for moment conditions

Between split-samples:

- ▶ Amenity shocks are orthogonal to baseline residual productivity:

$$\Delta \hat{\xi}_{f,o,NZ,t}^s \perp \hat{\phi}_{f,o,t-1}^{-s} \mid o, c(f).$$

- ▶ Productivity shocks are orthogonal to firm-level amenity shocks:

$$\Delta \hat{\xi}_{f,o,NZ,t}^s \perp \Delta \hat{\phi}_{f,o,t}^{-s} \mid o, c(f).$$

- ▶ Productivity shocks are orthogonal to market-level amenity shocks:

$$\Delta \hat{\xi}_{c_f,o,NZ,t}^s \perp \Delta \hat{\phi}_{f,o,t}^{-s} \mid o.$$

Explicit forms for moment conditions

We construct a instrument to shift firm-level ‘aggregate labor’

$$\log\left(\sum_{o \in \mathbf{O}_{f,t}} e^{\phi_{f,o,t} + \rho \log l_{f,o,t}}\right).$$

- ▶ Form the firm-by-occupation amenity shocks $\tilde{\xi}_{f,o,NZ}^s$ by residualizing $\Delta\hat{\xi}_{f,o,NZ,t}^s$ on a location-by-occupation-by-sample fixed effect.
- ▶ Form the firm-level aggregate amenity shock

$$\check{\xi}_{f,NZ,t}^s \equiv \log\left(\sum_{o \in \mathbf{O}_{f,t}} e^{\hat{\phi}_{f,o,t}^s + \rho \tilde{\xi}_{f,o,NZ}^s}\right).$$

- ▶ This aggregate amenity shock is also orthogonal to productivity shocks (between split-samples):

$$\check{\xi}_{f,NZ,t}^s \perp \Delta\hat{\phi}_{f,o,t}^{-s}.$$

Back.

Measuring the amenity value of employment at a firm

We infer $\xi_{f,o,x,t}$ using the labor supply of residents:

$$\xi_{f,o,x,t} = \log \sigma_{f,o,x,t|c_f}^{\text{resident}} - \tau \log(w_{f,o,t}) + D_{c_f,o,x,t},$$

When no type- x residents are observed, amenities must be imputed. We assume that amenities are joint-normal across types, and impute using (an approximation of) the posterior mode.

1. Calculate the empirical covariance of type-*NZ* and type-*abroad* amenities across firm-by-occupations at which both are observed \Rightarrow a $N(\mu_{f,o,x,t}^\xi, s_{f,o,x,t}^\xi)$ prior.
2. The log posterior for $\tilde{\xi}_{f,o,x,t}$, given that $\sigma_{f,o,x,t|c_f}^{\text{resident}} = 0$ is

$$C + n_{x,c_f,t} \log \left(\frac{\sum_{f' \in \mathbf{F}_{c,o,t} \setminus f} \exp(\tau \log(w_{f',o,t}) + \tilde{\xi}_{f',o,x,t})}{\sum_{f' \in \mathbf{F}_{c,o,t}} \exp(\tau \log(w_{f',o,t}) + \tilde{\xi}_{f',o,x,t})} \right) + \frac{(\tilde{\xi}_{f,o,x,t} - \mu_{f,o,x,t}^\xi)^2}{2 s_{f,o,x,t}^{\xi^2}},$$

where $n_{x,c,t}$ is the number of type- x workers in location c .

Measuring the amenity value of employment at a firm

3. Taking a first-order condition for $\tilde{\xi}_{f,o,x,t}$ implies

$$-n_{x,c,t} \left(\frac{\exp(\tau \log(w_{f,o,t}) + \tilde{\xi}_{f,o,x,t})}{\sum_{f' \in \mathbf{F}_{c,o,t} \setminus f} \exp(\tau \log(w_{f',o,t}) + \tilde{\xi}_{f',o,x,t})} \right) - \frac{(\tilde{\xi}_{f,o,x,t} - \mu_{f,o,x,t}^\xi)^2}{s_{f,o,x,t}^{\xi}} = 0.$$

4. Evaluated at $\tilde{\xi}_{f',o,x,t} \approx \mu_{f',o,x,t}^\xi$, yields the approximate solution

$$\tilde{\xi}_{f,o,x,t} = \mu_{f,o,x,t}^\xi - s_{f,o,x,t}^{\xi} n_{x,c,t} \left(\frac{\exp(\tau \log(w_{f,o,t}) + \mu_{f,o,x,t}^\xi)}{\sum_{f' \in \mathbf{F}_{c,o,t} \setminus f} \exp(\tau \log(w_{f',o,t}) + \mu_{f',o,x,t}^\xi)} \right).$$

Back.