

# Dynamic Bayesian Gaussian Graphical Models for Inferring Evolving Network Structure



Slav Kirov<sup>1</sup>, Micol Marchetti-Bowick<sup>2</sup>, Willie Neiswanger<sup>2</sup>

<sup>1</sup>Department of Mathematical Sciences and <sup>2</sup>Machine Learning Department, Carnegie Mellon University

# **ABSTRACT**

We propose a method for learning the structure of evolving Gaussian graphical models via Bayesian inference. Gaussian graphical models (GGMs) are often used to model the structure of a network, where they have found applications in computer vision (object pose), finance (relationships among stocks), biology (gene networks), and a number of related fields.

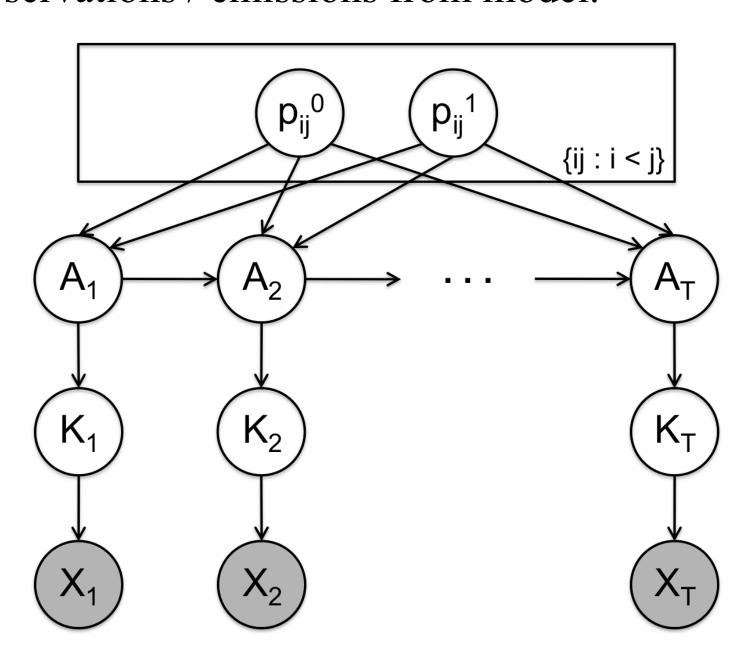
# INTRODUCTION

Learning the structure of a network is a problem of interest in a variety of fields. For a biologist, uncovering the structure of a gene network that characterizes functional associations between genes can provide insight into how genes interact during the regulation of a biological process. Similarly, a financial analyst may be interested in understanding how the prices of different stocks relate to one another. In both of these settings, learning a dynamic network that evolves over time can provide even more valuable information about how these relationships change in time.

## MODEL

Our generative model is a dynamic Bayesian network (DBN) that incorporates a Gaussian graphical model (GGM) at each time step, where

- $A_t = \{A_{ij}^t : i < j\}$ : adjacency matrix for the graph at time t.
- $K_t$ : precision matrix for multivariate Normal distribution.
- $X_t$ : observations / emissions from model.



$$p_{ij}^{0} \sim \text{Beta}(\alpha^{0}, \beta^{0}) \qquad p_{ij}^{1} \sim \text{Beta}(\alpha^{1}, \beta^{1})$$

$$A_{ij}^{t} \mid (A_{ij}^{t-1} = 0) \sim \text{Bernoulli}(p_{ij}^{0}) \qquad A_{ij}^{t} \mid (A_{ij}^{t-1} = 1) \sim \text{Bernoulli}(p_{ij}^{1})$$

$$K_t \mid A_t \sim \text{G-Wishart}_{A_t}(b, D)$$
  $X_t \mid K_t \sim \text{Normal}(0, K_t^{-1})$ 

Where:

$$p_{ij}^{0} = P(A_{ij}^{t} = 1 \mid A_{ij}^{t-1} = 0) \qquad 1 - p_{ij}^{0} = P(A_{ij}^{t} = 0 \mid A_{ij}^{t-1} = 0)$$

$$p_{ij}^{1} = P(A_{ij}^{t} = 0 \mid A_{ij}^{t-1} = 1) \qquad 1 - p_{ij}^{1} = P(A_{ij}^{t} = 1 \mid A_{ij}^{t-1} = 1)$$

#### INFERENCE

We experiment with two different inference methods for estimating the posterior distribution of our dynamic model.

# **Collapsed Gibbs Sampling**

First, we implemented a Gibbs sampler in which we integrate out the precision matrix variables  $K_{\{1:T\}}$  and resample each edge at each time point,  $A_{ii}^t$ , using the Gibbs proposal distribution.

$$P(A_{ij}^{t}|rest \setminus K_{1:T}) = \frac{1}{Z}P(X_{t}|A_{t})P(A_{ij}^{t}|A_{ij}^{t-1})P(A_{ij}^{t+1}|A_{ij}^{t})$$
(1)

$$= \frac{1}{Z} P(A_{ij}^t | A_{ij}^{t-1}) P(A_{ij}^{t+1} | A_{ij}^t) \int P(X_t, K_t | A_t) dK_t$$
 (2)

$$= \frac{1}{Z} P(A_{ij}^t | A_{ij}^{t-1}) P(A_{ij}^{t+1} | A_{ij}^t) \int P(X_t | K_t) P(K_t | A_t) dK_t$$
(3)

$$= \frac{1}{Z} P(A_{ij}^t | A_{ij}^{t-1}) P(A_{ij}^{t+1} | A_{ij}^t) (2\pi)^{-\frac{np}{2}} \frac{I_G(b+n, D+S)}{I_G(b, D)}$$
(4)

In (4),  $I_G(b, D)$  is the normalization constant of the prior G-Wishart distribution  $P(K_t)$  and  $I_G(b+n, D+S)$  is the normalization constant of the posterior G-Wishart distribution  $P(K_t|X_t)$ . These constants are computed using Monte Carlo simulation.

#### **Sequential Monte Carlo**

Second, we implemented a version of Sequential Monte Carlo (SMC) that builds on the Birth-Death MCMC inference method proposed by Mohammadi & Wit (2012). We chose SMC because it performs inference in an online fashion and is thus well-suited to time-varying models. At a high level, SMC proceeds with the following steps:

# **SMC Algorithm:**

- 1. At the first time step, run BD-MCMC using  $X_1$  as input and get N samples from the partial posterior  $P(A_1, K_1 | X_1)$  as output
- 2. At time t-l, we have N particles (samples of  $A_{1:t-l}, K_{1:t-l}$ ) that are approximately distributed according to  $P(A_{1:t-l}, K_{1:t-l} | X_{1:t-l})$
- 3.At time t, run a modified version of BD-MCMC to sample  $A_t$  and  $K_t$  from proposal distribution Q, where:

$$Q(A_t, K_t \mid A_{1:t-1}, X_{1:t}) = P(A_t, K_t \mid A_{t-1}, X_t)$$

- 4. Extend the particles with these new values
- 5. Compute the importance weight of each particle
- 6.Resample the *N* particles according to a categorical distribution made up of the normalized importance weights
- 7. Output N samples from the full posterior  $P(A_{1:T}, K_{1:T} | X_{1:T})$

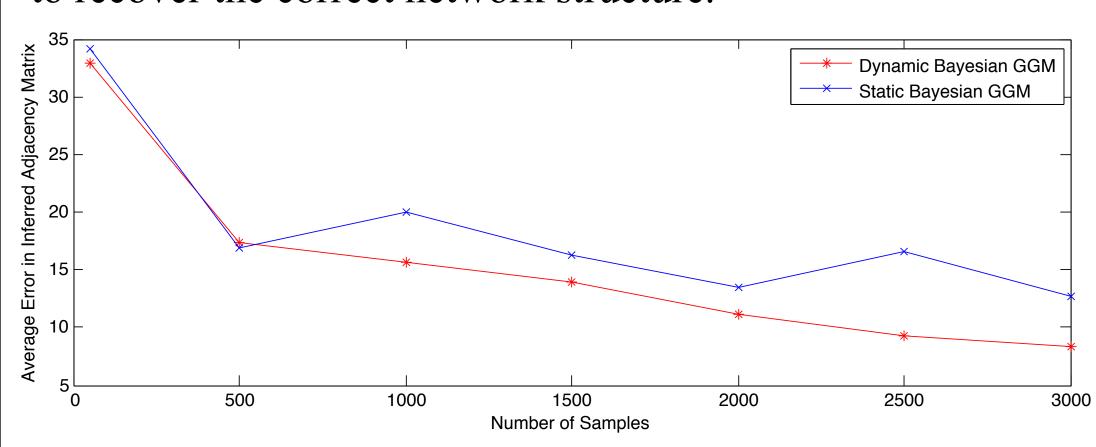
In step (3), we sample from the proposal distribution  $P(A_t, K_t | A_{t-1}, X_t)$  by modifying BD-MCMC to incorporate the prior over the graph structure  $P(A_t | A_{t-1})$ . BD-MCMC traverses the space of hidden variables by executing birth moves (adding an edge) and death moves (removing an edge), and re-sampling the precision matrix after each move.

# **RESULTS**

We carry out several experiments on both synthetic and real data to evaluate the success of our method.

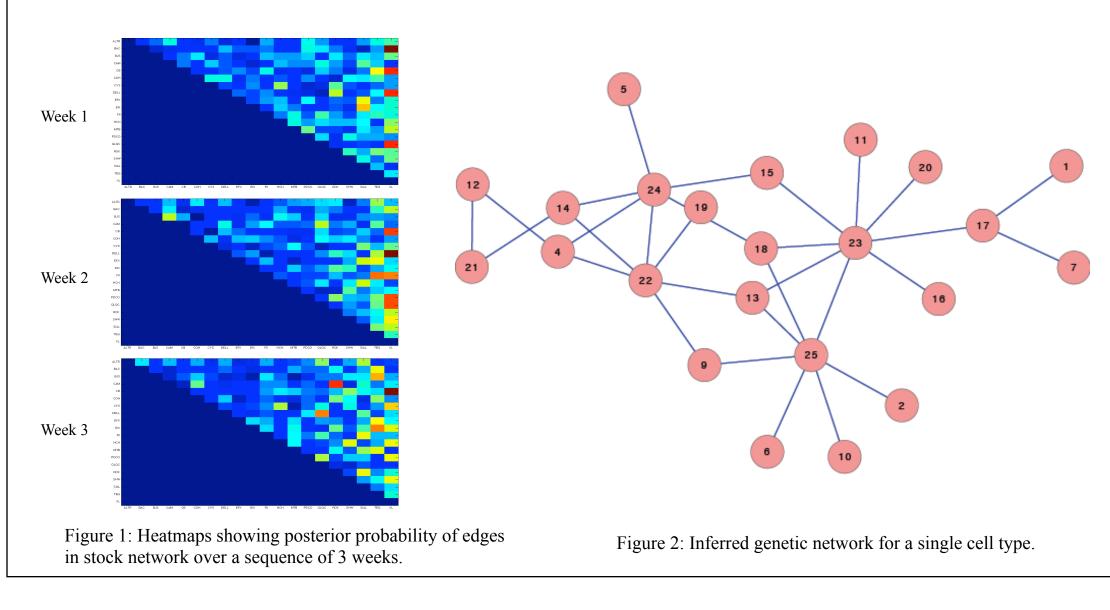
## **Synthetic Data Experiment**

We performed SMC inference on a synthetic time-varying network with 12 nodes over 10 time-steps, in order to judge performance. We compared its algorithm error with that of running BD-MCMC independently at each time step. The figure below shows that the dynamic model was better able to recover the correct network structure.



## Finance and Genetic Data Experiments

We perform inference on two real-world datasets. In the first, we infer the structure of a GGM for a set of 20 companies over a sequence of 10 weeks based on their stock prices, and in the second, we infer a network of 25 genes as the network varies over a sequence of 9 cell types.



## CONCLUSIONS

We have derived a time-varying, Bayesian model for inferring the structure of a sequence of Gaussian graphical models. By running this model on synthetic data (for which there exists ground-truth) we have shown that our model outperforms a static equivalent. We have also demonstrated our model on multiple real-life datasets in order to show its ability to infer network structure in the fields of both finance and genetics. We hope that pursuing Gaussian graphical model learning in a Bayesian setting will allow for flexible specification of dependencies, especially of the dynamics of time-varying graphs