



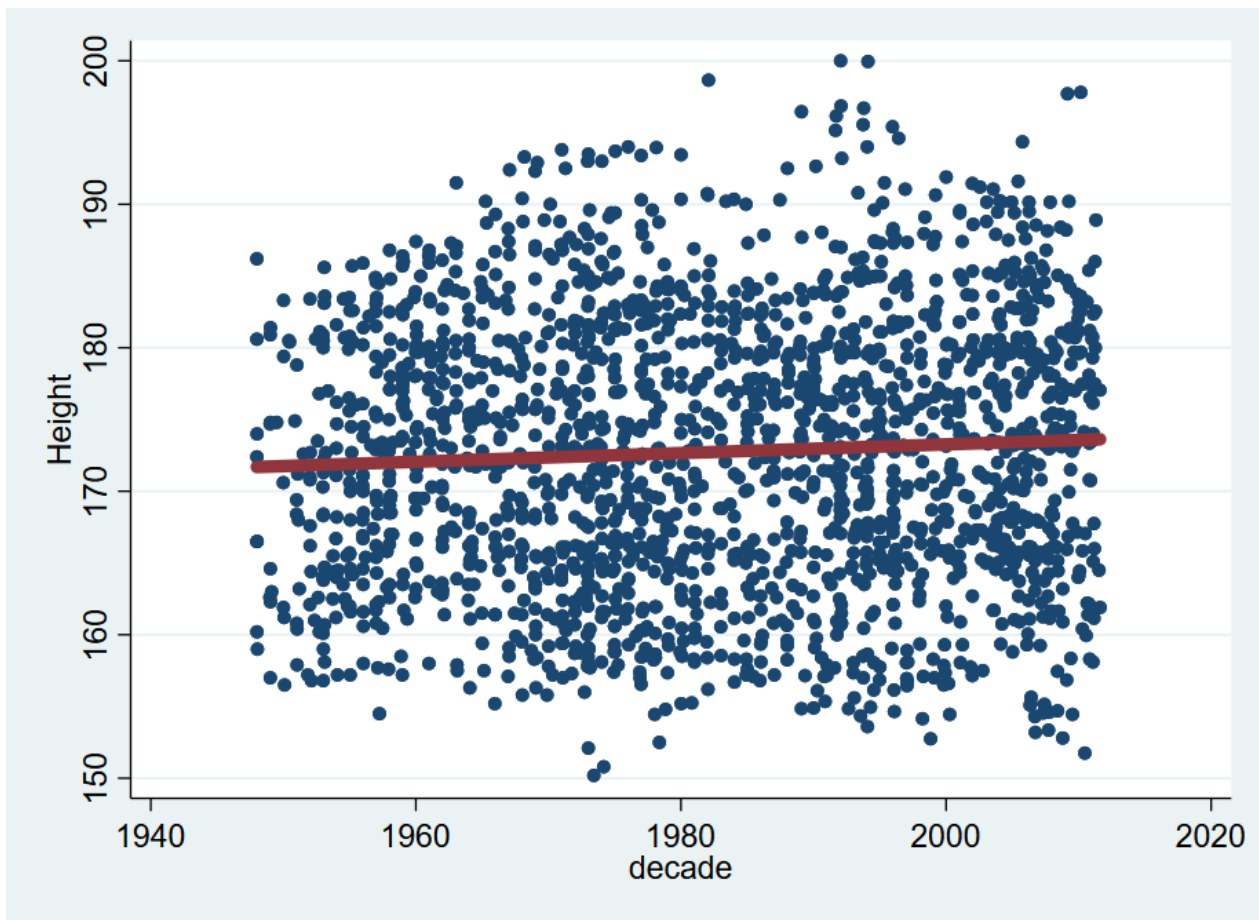
# A beginners guide to (growth trajectory analysis using) **multilevel models**

**Will Johnson**

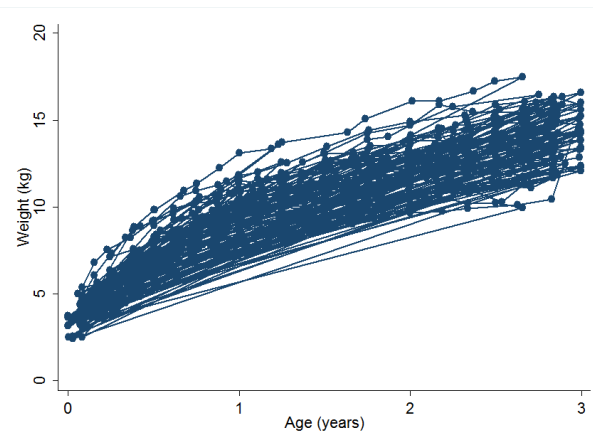
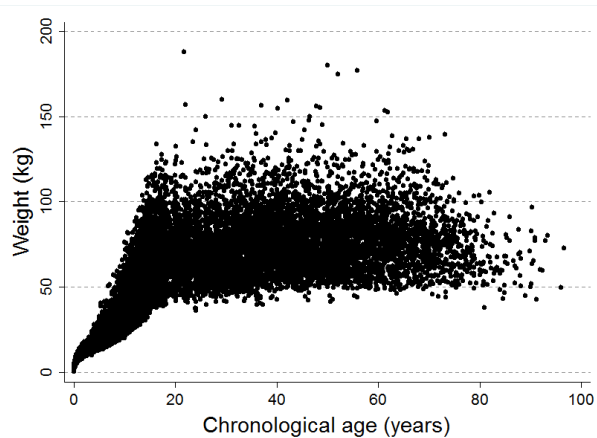
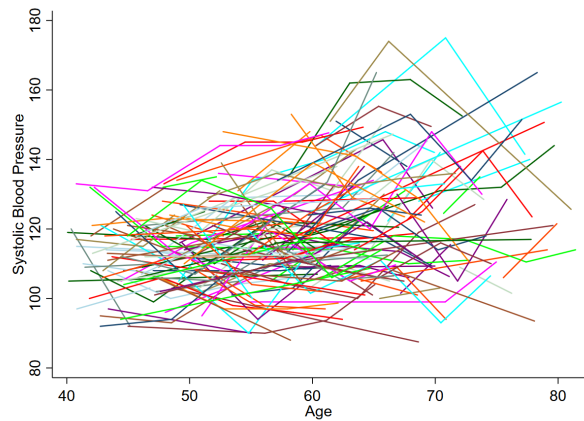
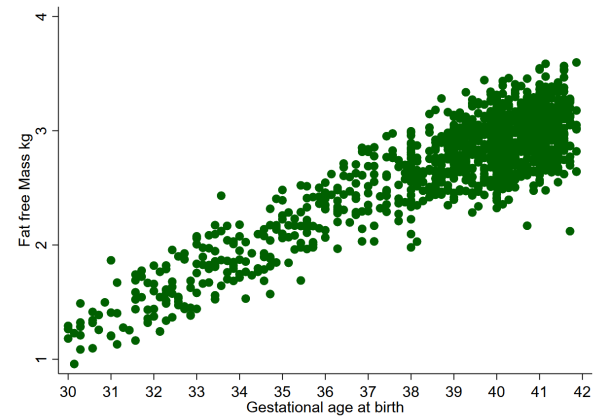
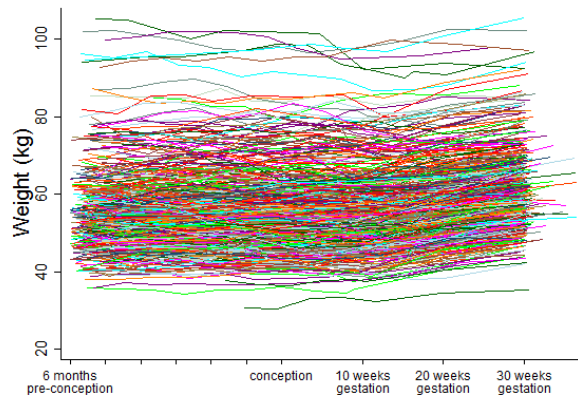
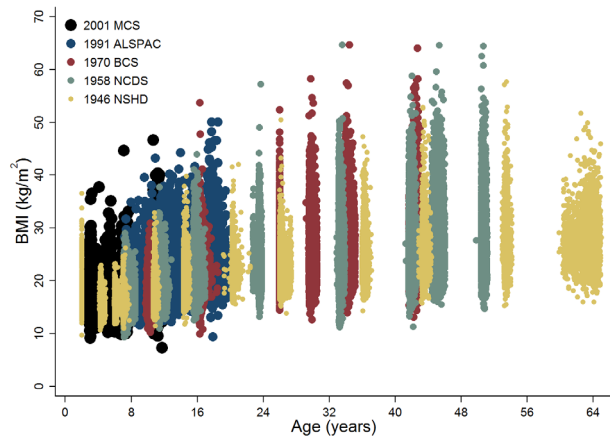
**W.O.Johnson@lboro.ac.uk**

School of Sport, Exercise and Health Sciences, Loughborough University

# What is a trajectory?



# Longitudinal data

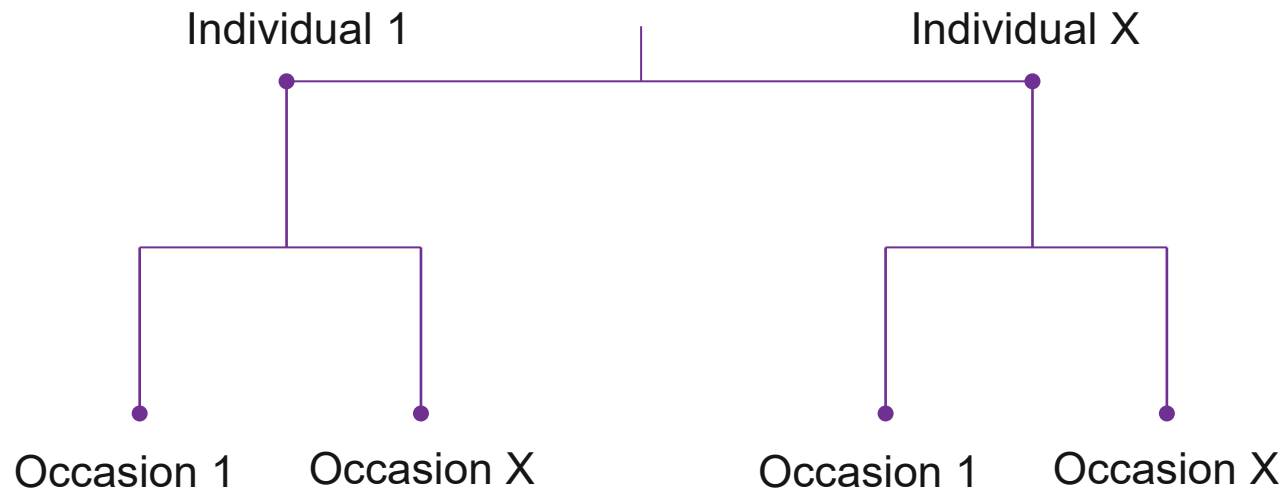




# Structure of longitudinal data

clustered within  
**individuals (level 2)**

serial measurement  
**occasions (level 1)**





# Structure of longitudinal data

Individual	Occasion			
	1	2	3	4
1	$y_{11}$	$y_{21}$		$y_{41}$
2	$y_{12}$	$y_{22}$	$y_{32}$	$y_{42}$
3	$y_{13}$		$y_{33}$	$y_{43}$



Individual	Occasion	
1	1	$y_{11}$
1	2	$y_{21}$
1	3	
1	4	$y_{41}$
2	1	$y_{12}$
2	2	$y_{22}$
2	2	$y_{32}$
2	2	$y_{42}$
3	1	$y_{13}$
3	2	
3	2	$y_{33}$
3	2	$y_{43}$

$y_{ij}$  is the response at  
occasion  $i$  for individual  $j$



# Other examples of multilevel data





# How many upper-level units?

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for Economics

Elff, Martin; Heisig, Jan Paul; Schaeffer, Merlin; Shikano, Susumu

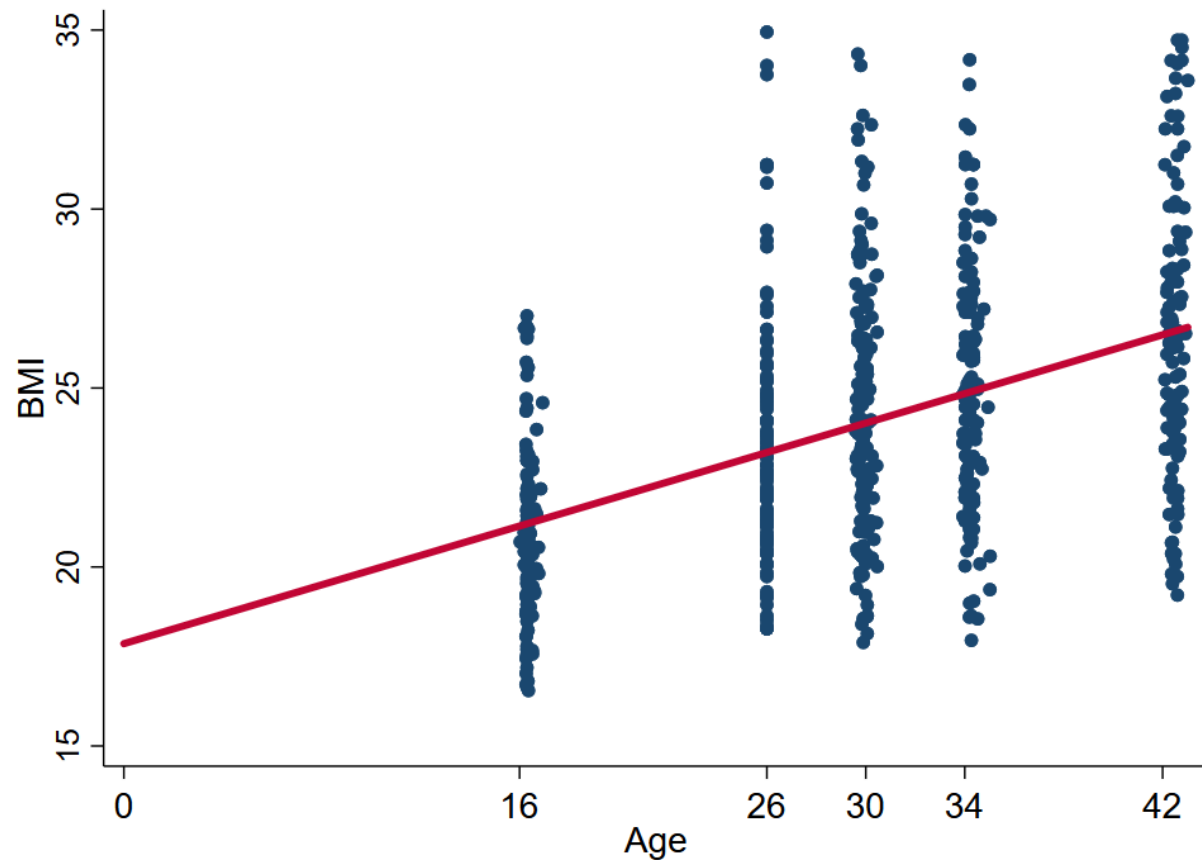
Article — Accepted Manuscript (Postprint)

Multilevel Analysis with Few Clusters: Improving  
Likelihood-based Methods to Provide Unbiased  
Estimates and Accurate Inference

British Journal of Political Science



# Regression

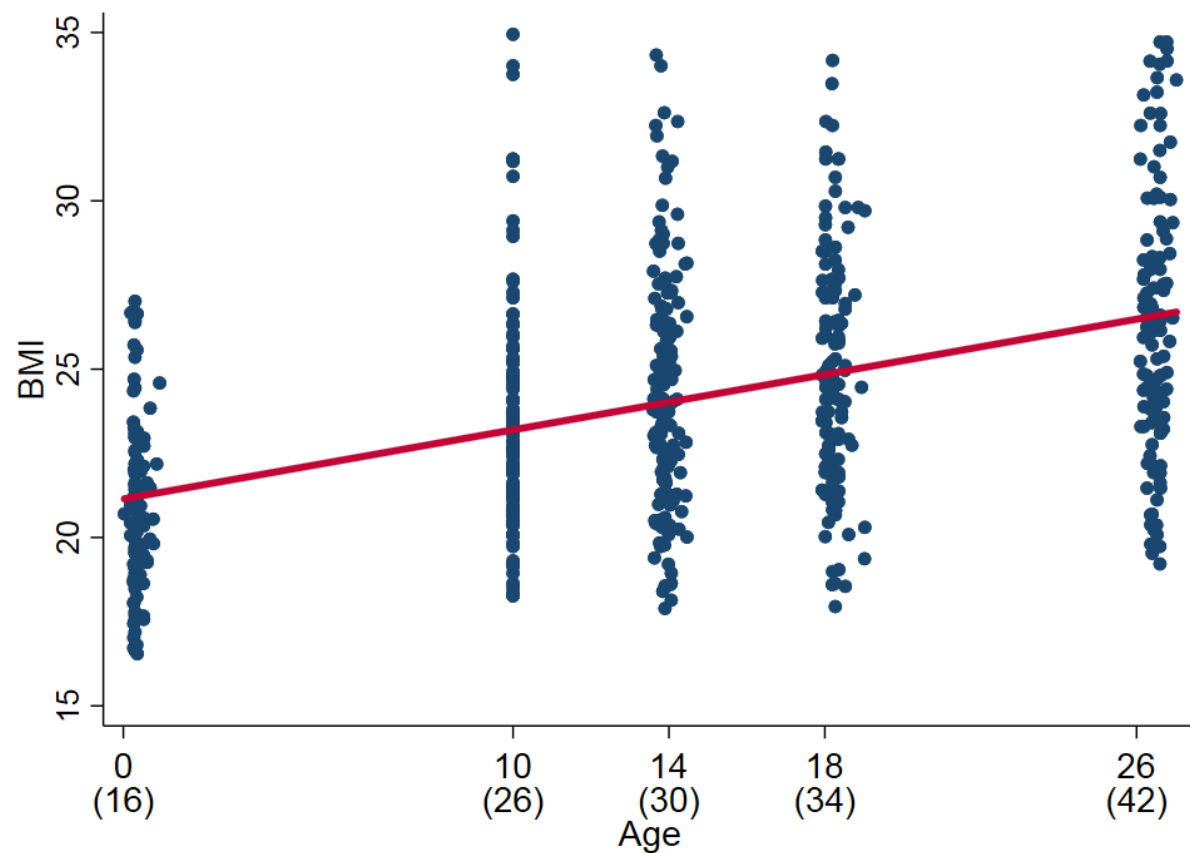


$$\text{bmi}_{ij} = \beta_{0i} \text{cons} + 0.205(0.016) \text{age}_{ij}$$
$$\beta_{0i} = 17.857(0.511) + e_{0ij}$$





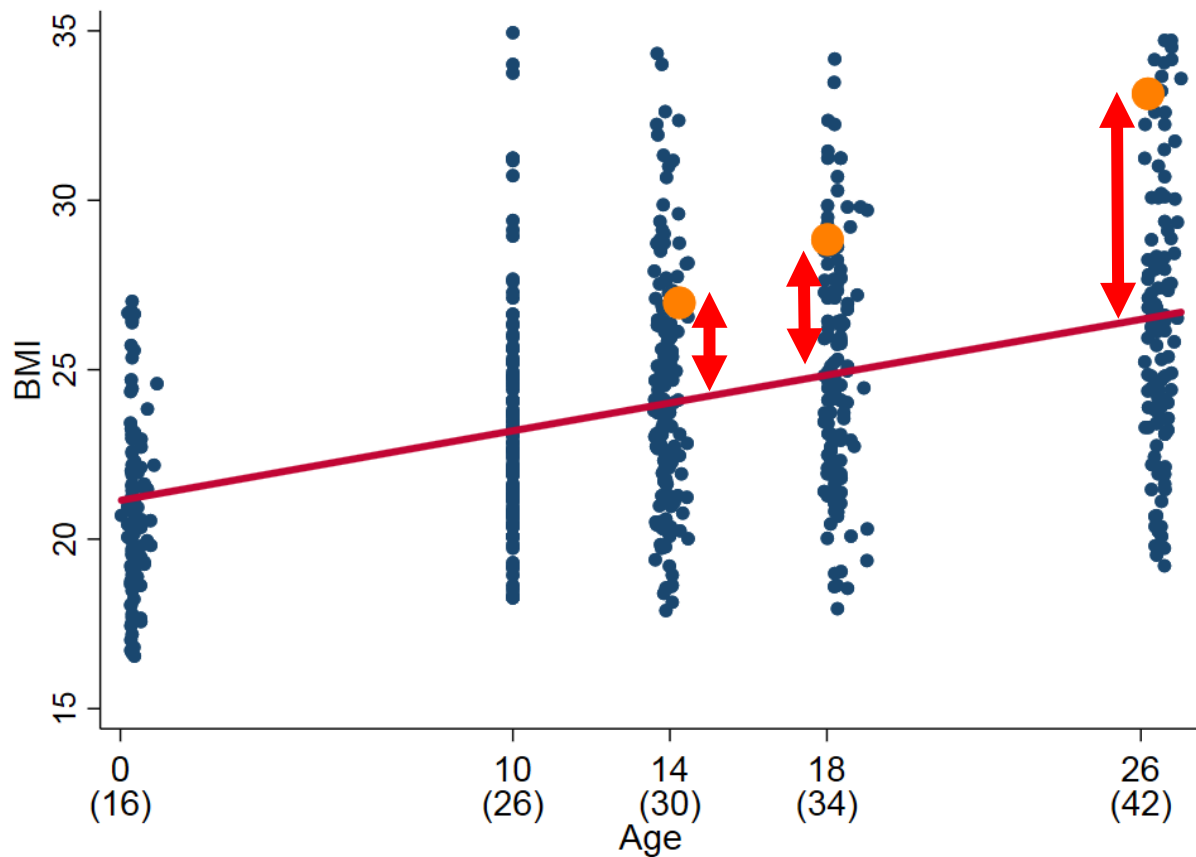
# Regression



$$\text{bmi}_{ij} = \beta_{0i}\text{cons} + 0.205(0.016)\text{agec}_{ij}$$
$$\beta_{0i} = 21.144(0.267) + e_{0ij}$$



# Regression



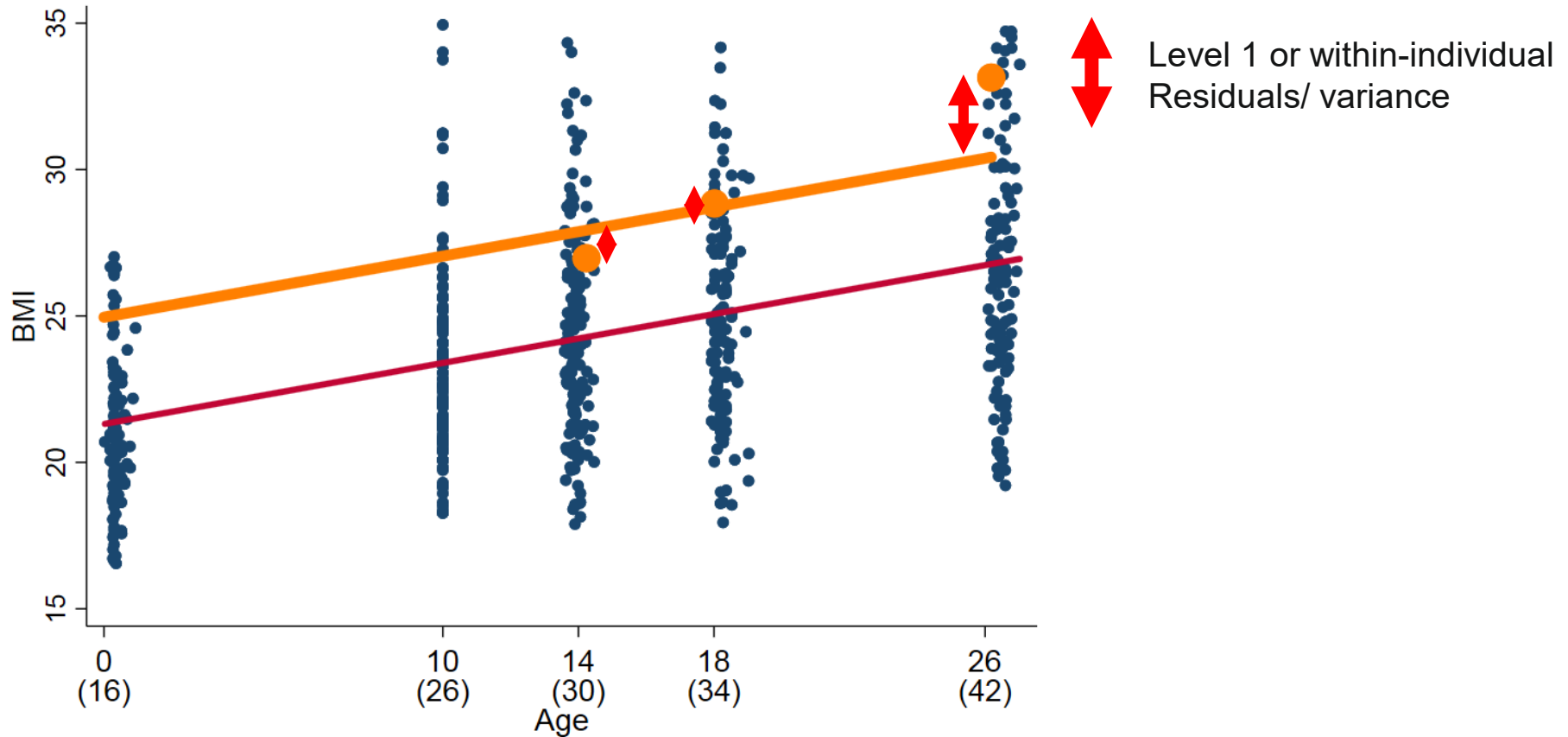
$$\text{bmi}_{ij} = \beta_{0i}\text{cons} + 0.205(0.016)\text{agec}_{ij}$$

$$\beta_{0i} = 21.144(0.267) + e_{0ij}$$

$$[e_{0ij}] \sim N(0, \Omega_e) : \Omega_e = [11.674(0.661)]$$

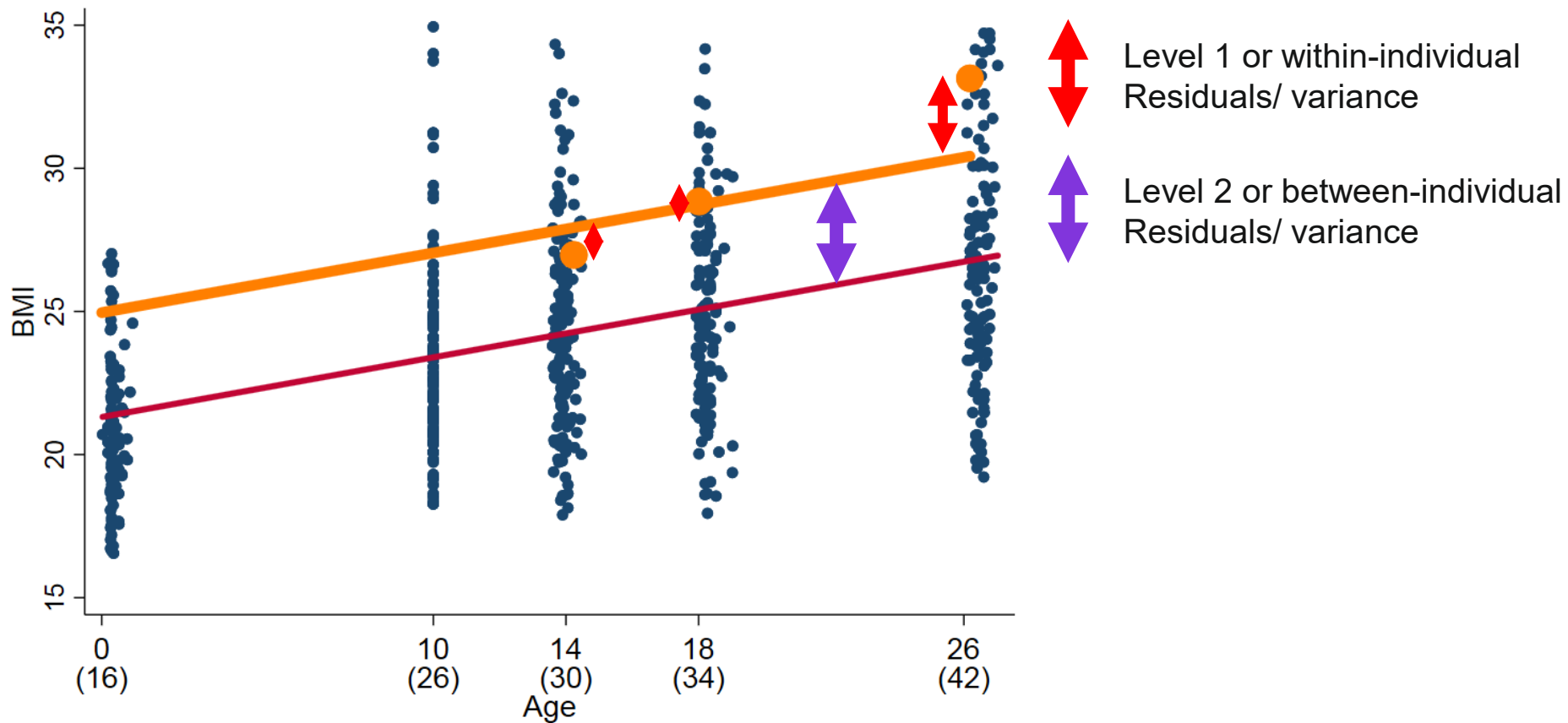


# Multilevel regression – random intercept



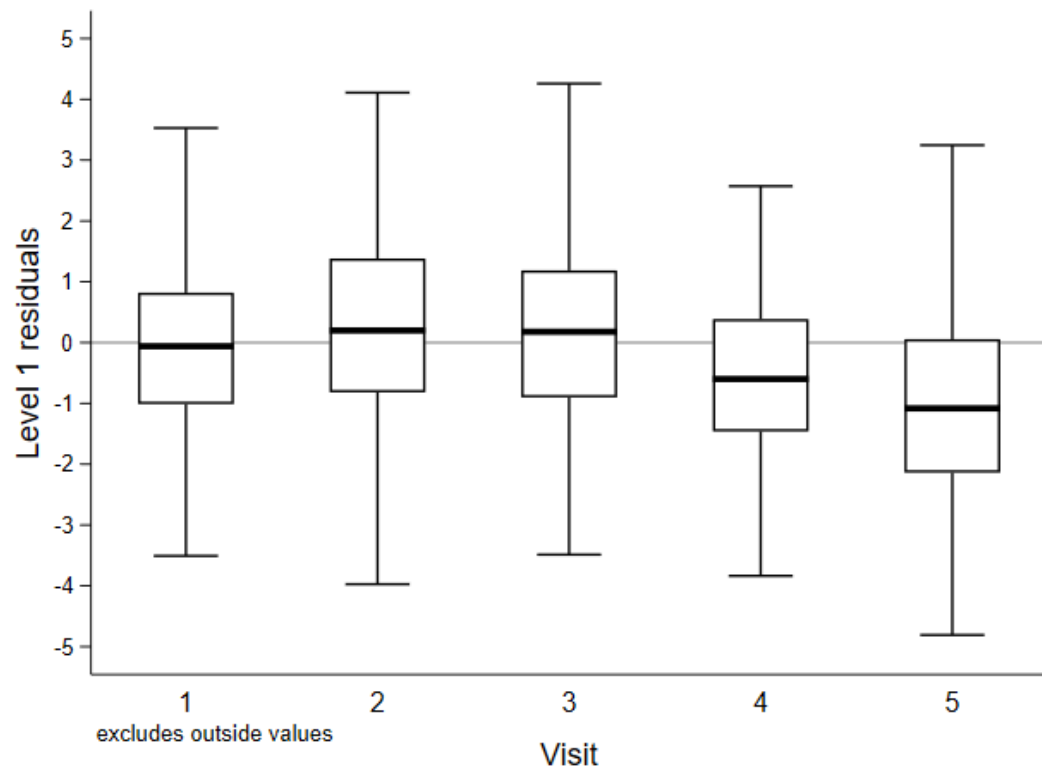


# Multilevel regression – random intercept





# Multilevel regression – random intercept



$$\text{bmi}_{ij} = \beta_{0ij} \text{cons} + \beta_1 \text{agec}_{ij}$$

$$\beta_{0ij} = \beta_0 + u_{0j} + e_{0ij}$$



$$[u_{0j}] \sim N(0, \Omega_u) : \Omega_u = [\sigma_{u0}^2]$$



$$[e_{0ij}] \sim N(0, \Omega_e) : \Omega_e = [\sigma_{e0}^2]$$



# Multilevel regression – random intercept

$$\text{bmi}_{ij} = \beta_{0ij} \text{cons} + 0.209(0.011) \text{agec}_{ij}$$

$$\beta_{0ij} = 21.311(0.263) + u_{0j} + e_{0ij}$$



$$[u_{0j}] \sim N(0, \Omega_u) : \Omega_u = [7.940(0.952)]$$

$$\text{sqrt}(7.940) = 2.8 \text{ kg/m}^2$$



$$[e_{0ij}] \sim N(0, \Omega_e) : \Omega_e = [4.104(0.282)]$$

$$\text{sqrt}(4.104) = 2.0 \text{ kg/m}^2$$



# Multilevel regression – random intercept

$$\text{bmi}_{ij} = \beta_{0ij} \text{cons} + 0.209(0.011) \text{agec}_{ij}$$

$$\beta_{0ij} = 21.311(0.263) + u_{0j} + e_{0ij}$$



$$[u_{0j}] \sim N(0, \Omega_u) : \Omega_u = [7.940(0.952)]$$

$$\text{sqrt}(7.940) = 2.8 \text{ kg/m}^2$$



$$[e_{0ij}] \sim N(0, \Omega_e) : \Omega_e = [4.104(0.282)]$$

$$\text{sqrt}(4.104) = 2.0 \text{ kg/m}^2$$

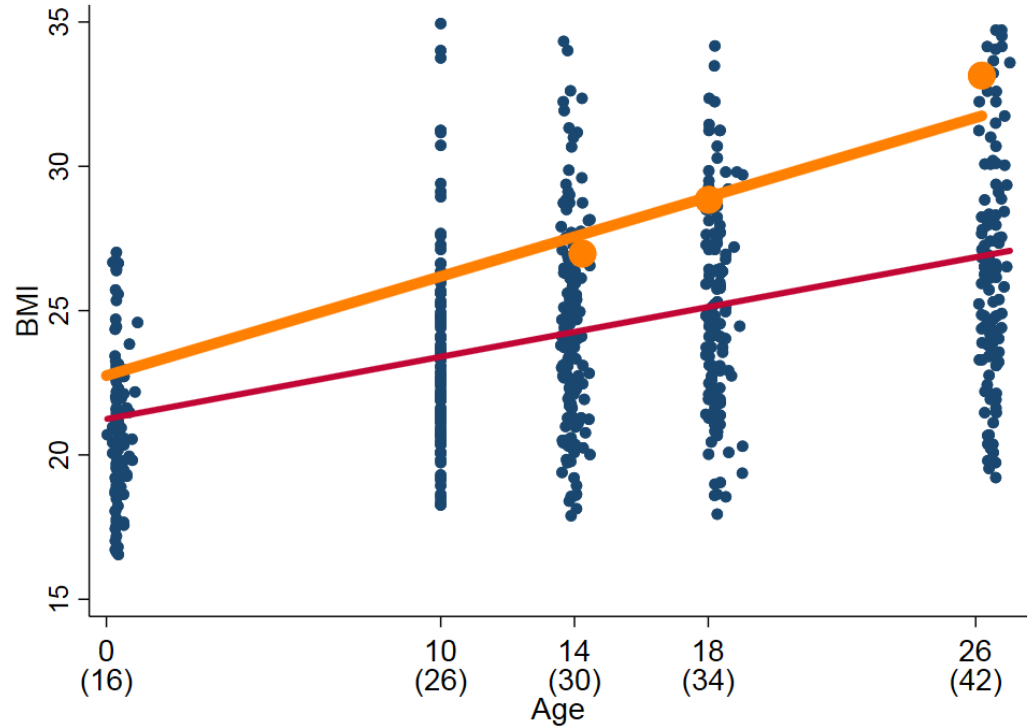
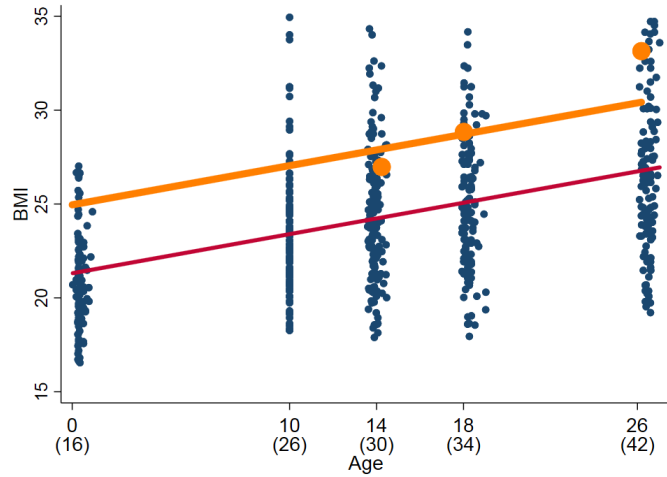
## Variance partitioning coefficient

Proportion of total variance that is due to between-individual differences (i.e., level 2 variance)

$$\begin{aligned} &= 7.940 / (4.103 + 7.940) \\ &= 0.659 \end{aligned}$$



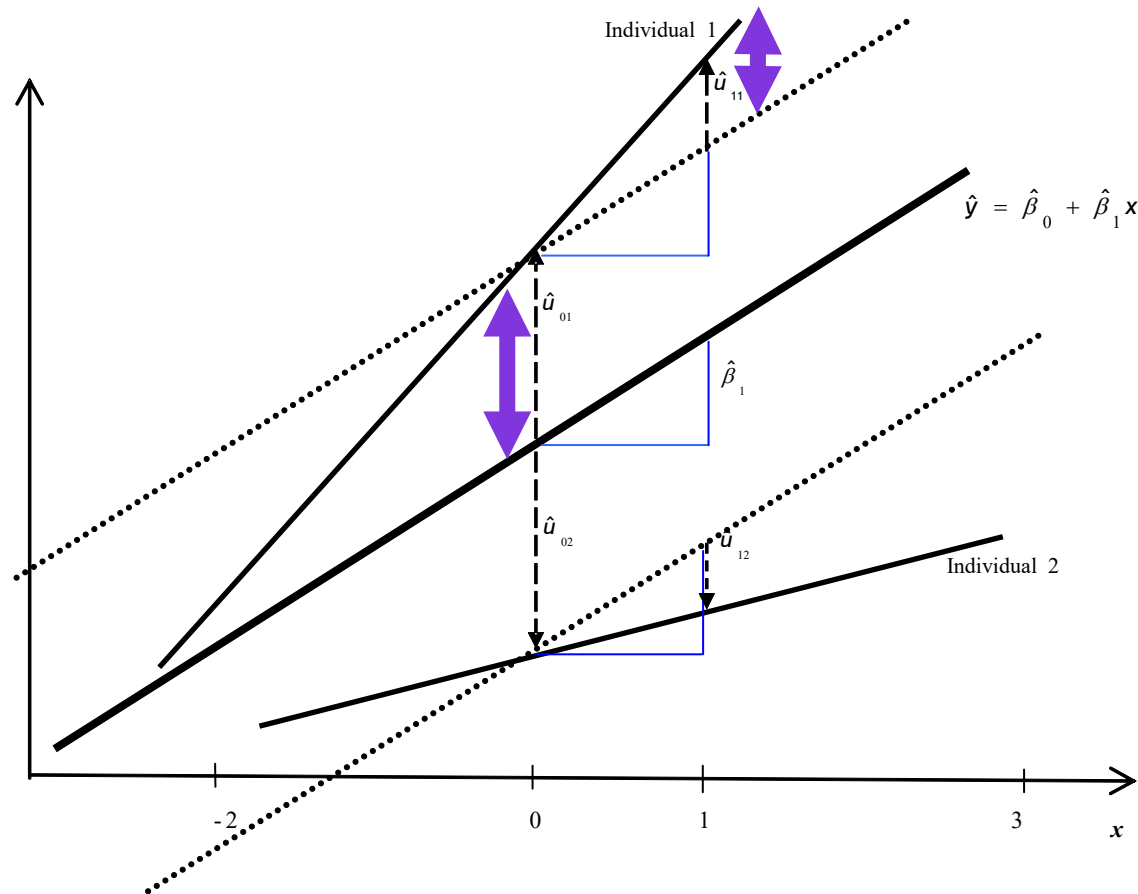
# Multilevel regression – random slope







# Multilevel regression – random slope





# Multilevel regression – random slope

$$\text{bmi}_{ij} = \beta_{0ij}\text{cons} + \beta_{1j}\text{agec}_{ij}$$

$$\beta_{0ij} = \beta_0 + u_{0j} + e_{0ij}$$

$$\beta_{1j} = \beta_1 + u_{1j}$$



$$\begin{bmatrix} u_{0j} \\ u_{1j} \end{bmatrix} \sim N(0, \Omega_u) : \Omega_u = \begin{bmatrix} \sigma_{u0}^2 & \\ \sigma_{u01} & \sigma_{u1}^2 \end{bmatrix}$$

$$\begin{bmatrix} 4.283(0.862) \\ 0.071(0.038) \quad 0.013(0.003) \end{bmatrix}$$

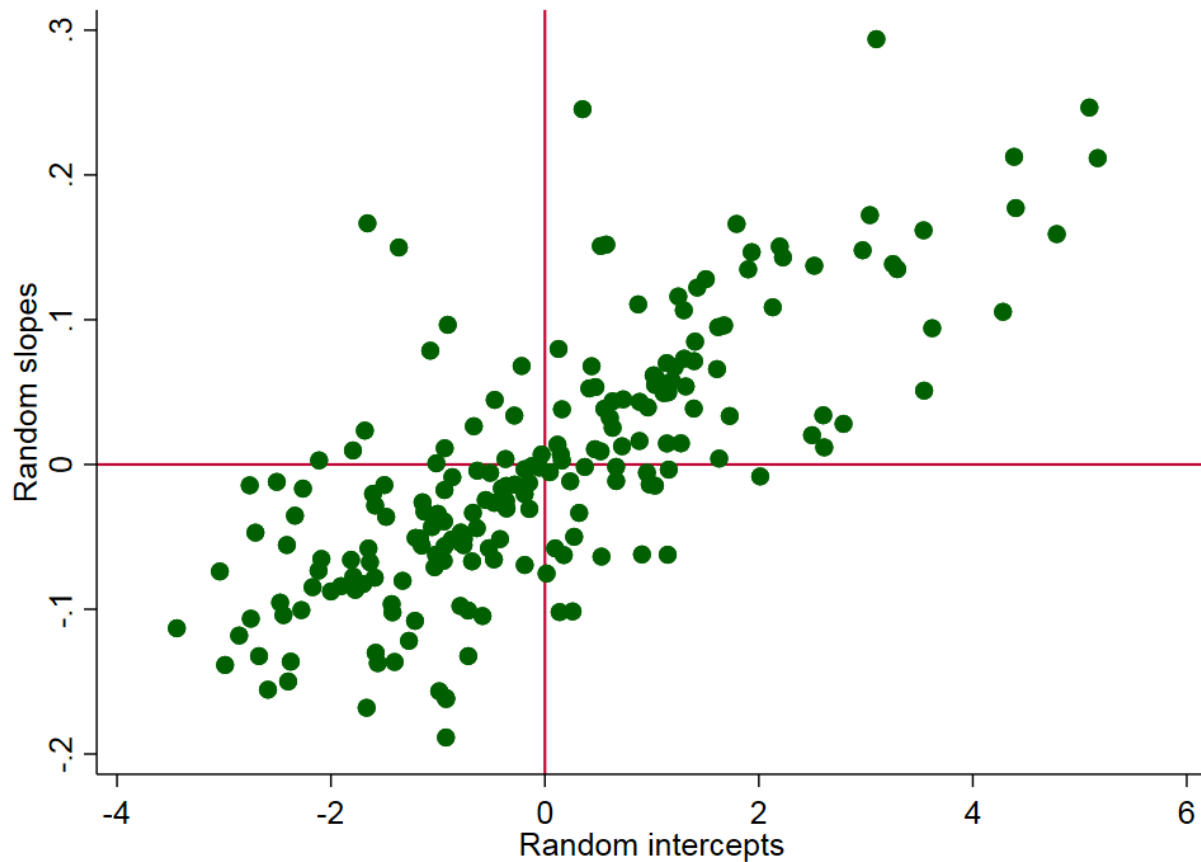


$$[e_{0ij}] \sim N(0, \Omega_e) : \Omega_e = [\sigma_{e0}^2]$$

$$[2.967(0.246)]$$



# Multilevel regression – random slope



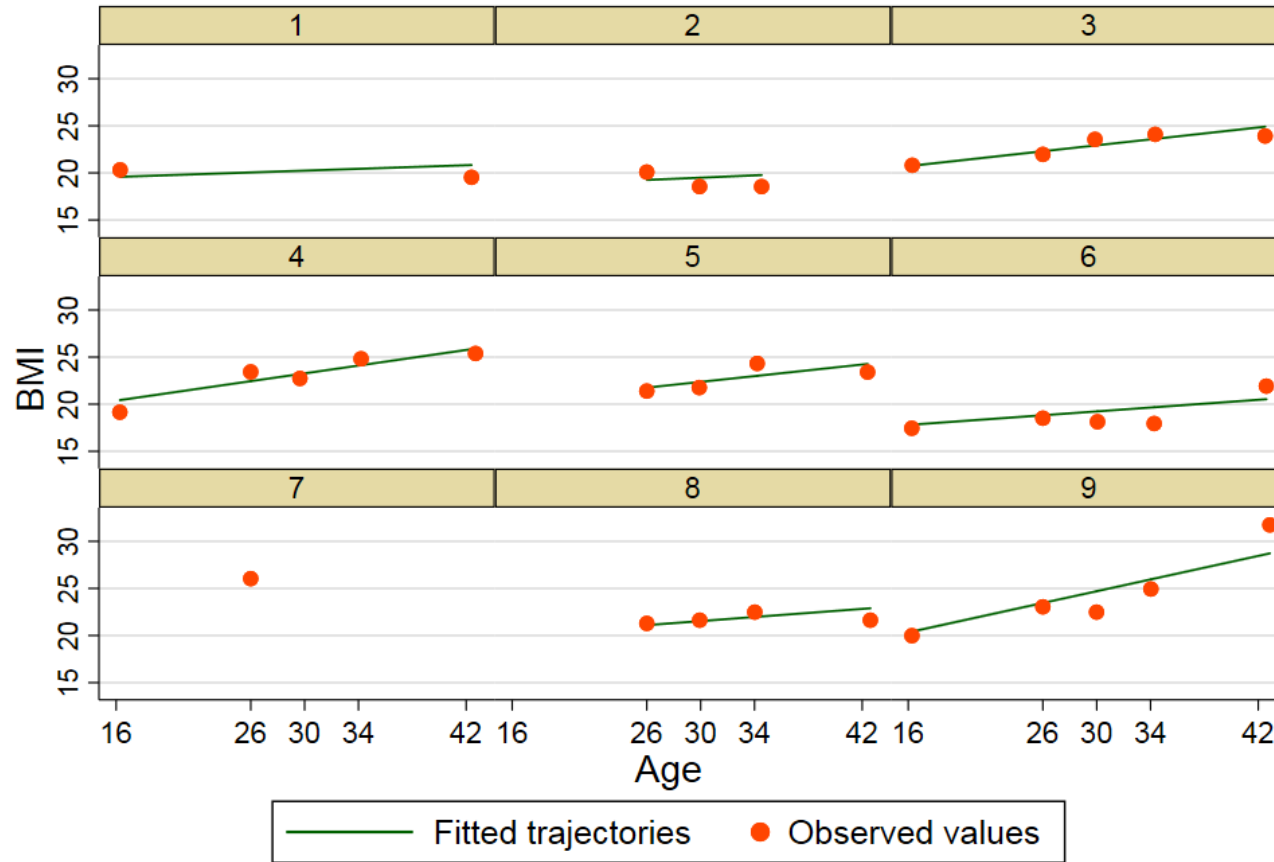
Covariance  
0.071

=

Correlation  
0.298



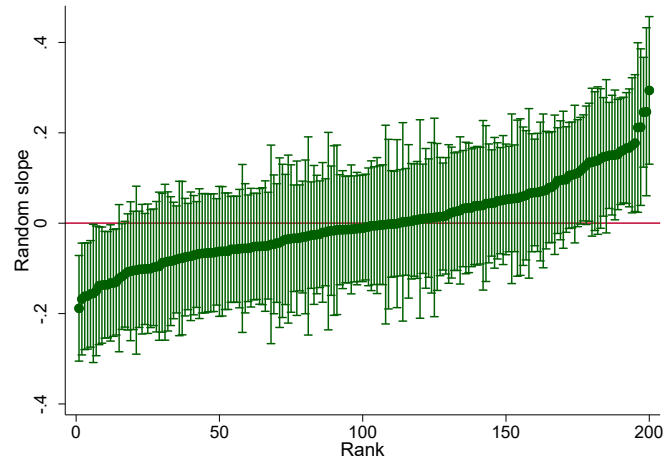
# Multilevel regression – random slope



Graphs by group(dum)

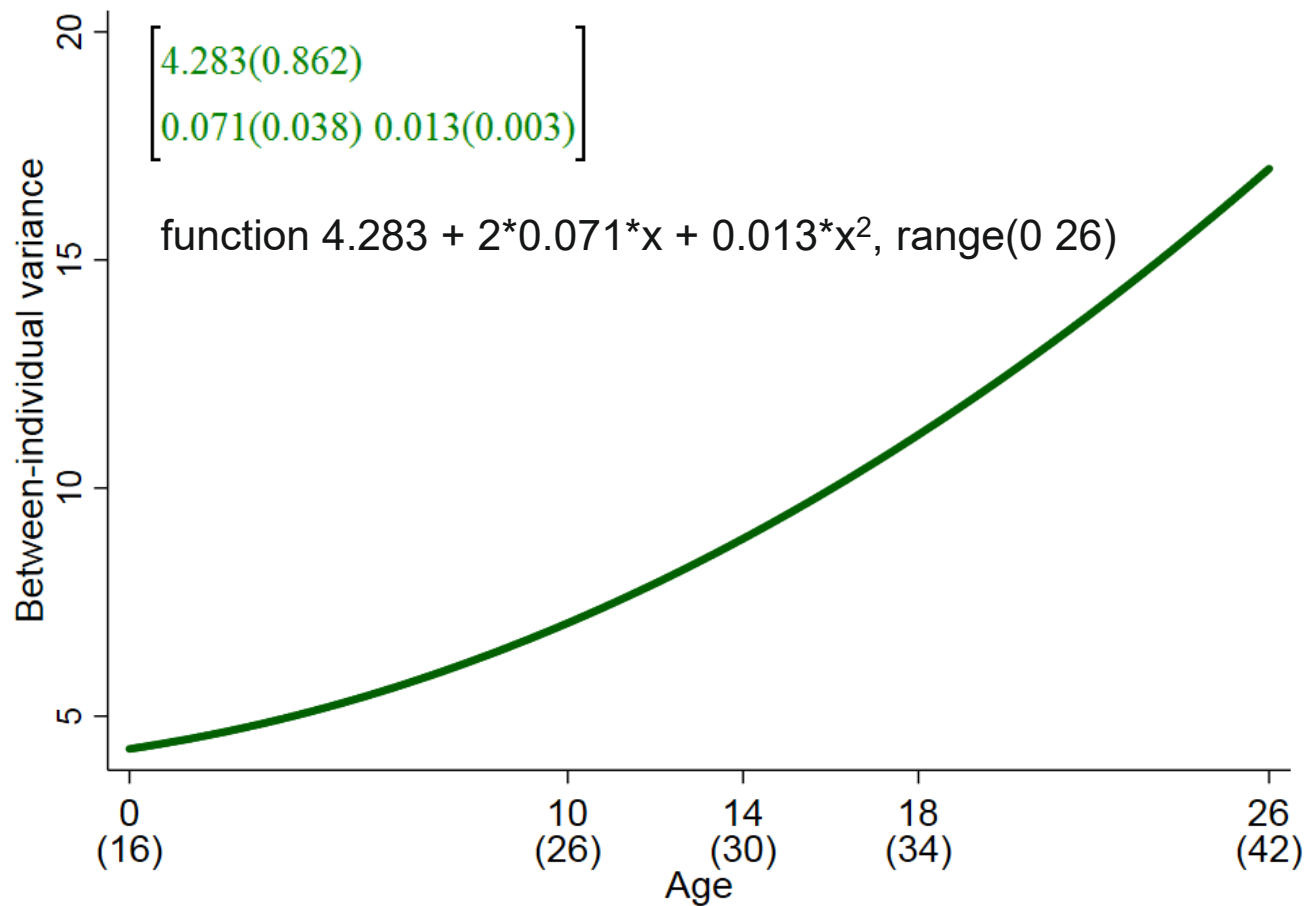


# Multilevel regression – random slope





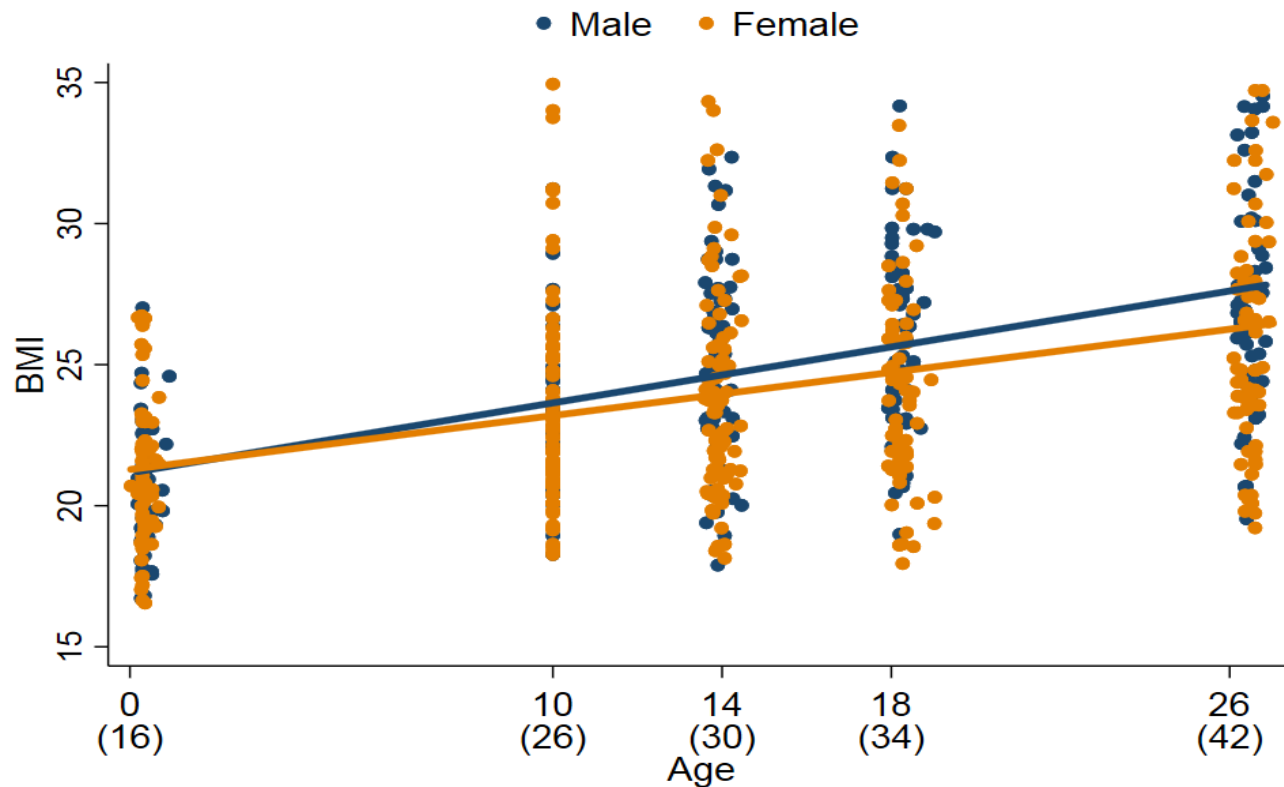
# Multilevel regression – random slope





## Including explanatory variables – level 2

$$\text{bmi}_{ij} = \beta_{0ij}\text{cons} + \beta_{1j}\text{agec}_{ij} + 0.118(0.427)\text{sex}_j + -0.057(0.026)\text{agec\_sex}_{ij}$$





## Including explanatory variables – level 1

$$\text{bmi}_{ij} = \beta_{0j}\text{cons} + \beta_{1j}\text{agec}_{ij} + 0.115(0.423)\text{sex}_j + -0.057(0.025)\text{agec\_sex}_{ij} + e_{4ij}\text{male}_j + e_{5ij}\text{female}_j$$

$$\beta_{0j} = 21.171(0.319) + u_{0j}$$

$$\beta_{1j} = 0.248(0.019) + u_{1j}$$

$$\begin{bmatrix} u_{0j} \\ u_{1j} \end{bmatrix} \sim N(0, \Omega_u) : \Omega_u = \begin{bmatrix} 4.189(0.851) \\ 0.074(0.037) \quad 0.012(0.003) \end{bmatrix}$$

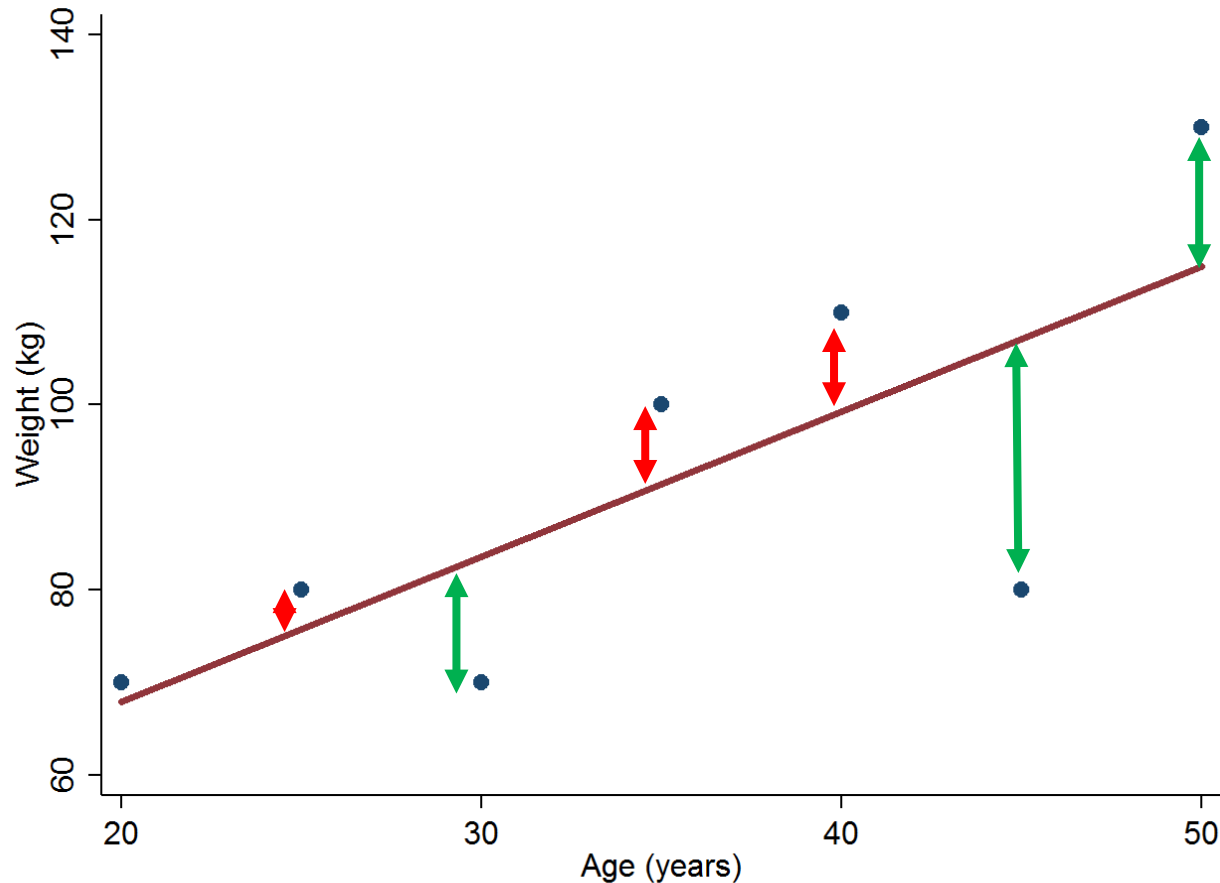
$$\begin{bmatrix} e_{4ij} \\ e_{5ij} \end{bmatrix} \sim N(0, \Omega_e) : \Omega_e = \begin{bmatrix} 2.699(0.343) \\ 0 \quad 3.202(0.322) \end{bmatrix}$$







# Including explanatory variables – level 1



**Measured**

**Self-reported**



# Recap

With longitudinal data:

- Multilevel models account for clustering by **partitioning variance** into within-individual (level 1) and between-individual (level 2)
- A **separate trajectory**, describing how  $y$  changes over age/time, is essentially estimated for each individual
  - Fixed intercept/slope = sample average trajectory
  - Fixed intercept/slope + random intercept/slope = individual trajectories
- Additional explanatory variables at level 1 (self-report) or 2 (sex) can be included to **model systematic differences between-individuals**
- Additional explanatory variables at level 1 (self-report) or 2 (sex) can be included to **model systematic differences within-individuals**



# Does accounting for clustering really matter?

**Fixed** intercept & slope

$$\beta_1 = 0.205(0.016)$$

Vs.

**Random** intercept & slope

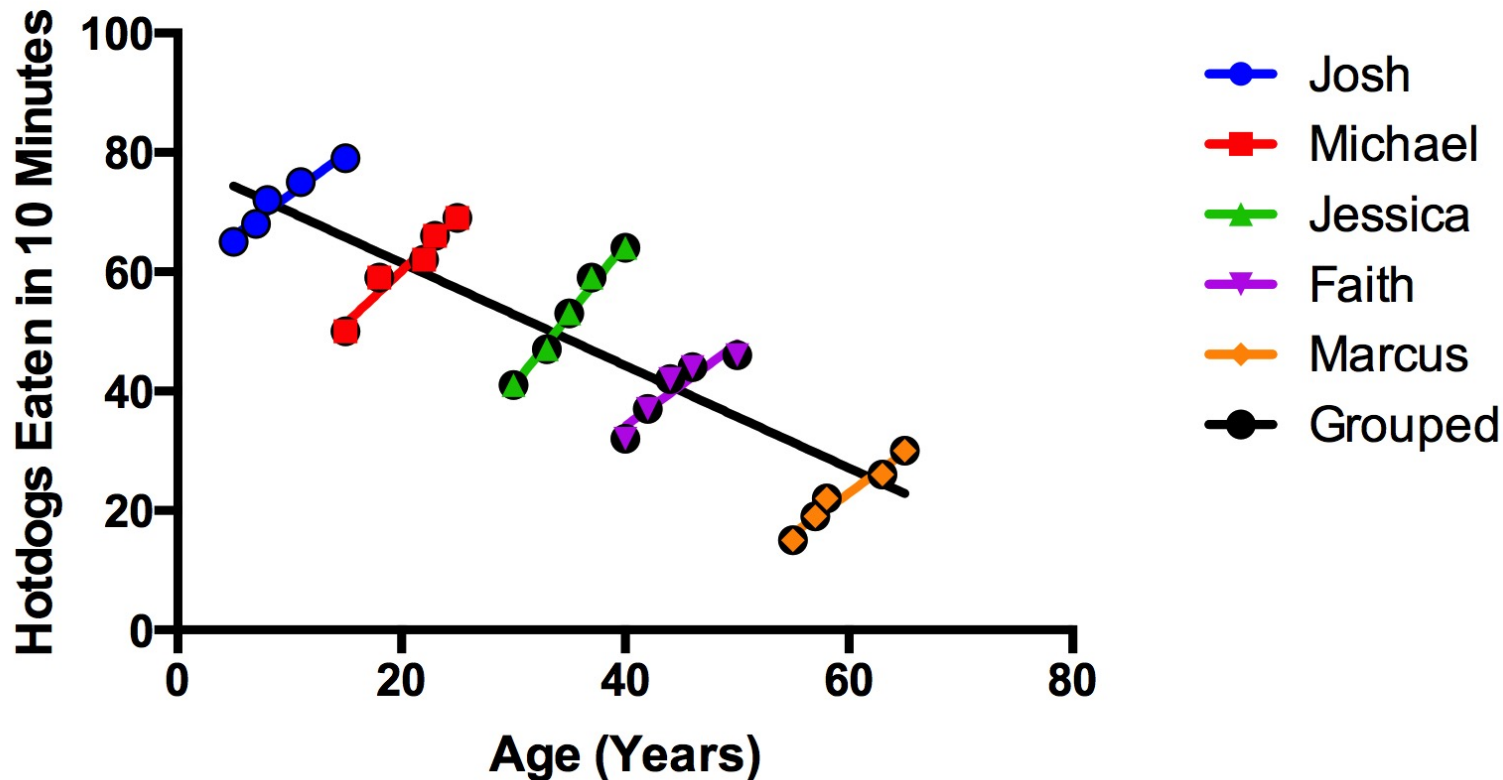
$$\beta_1 = 0.216(0.013)$$





# Does accounting for clustering really matter?

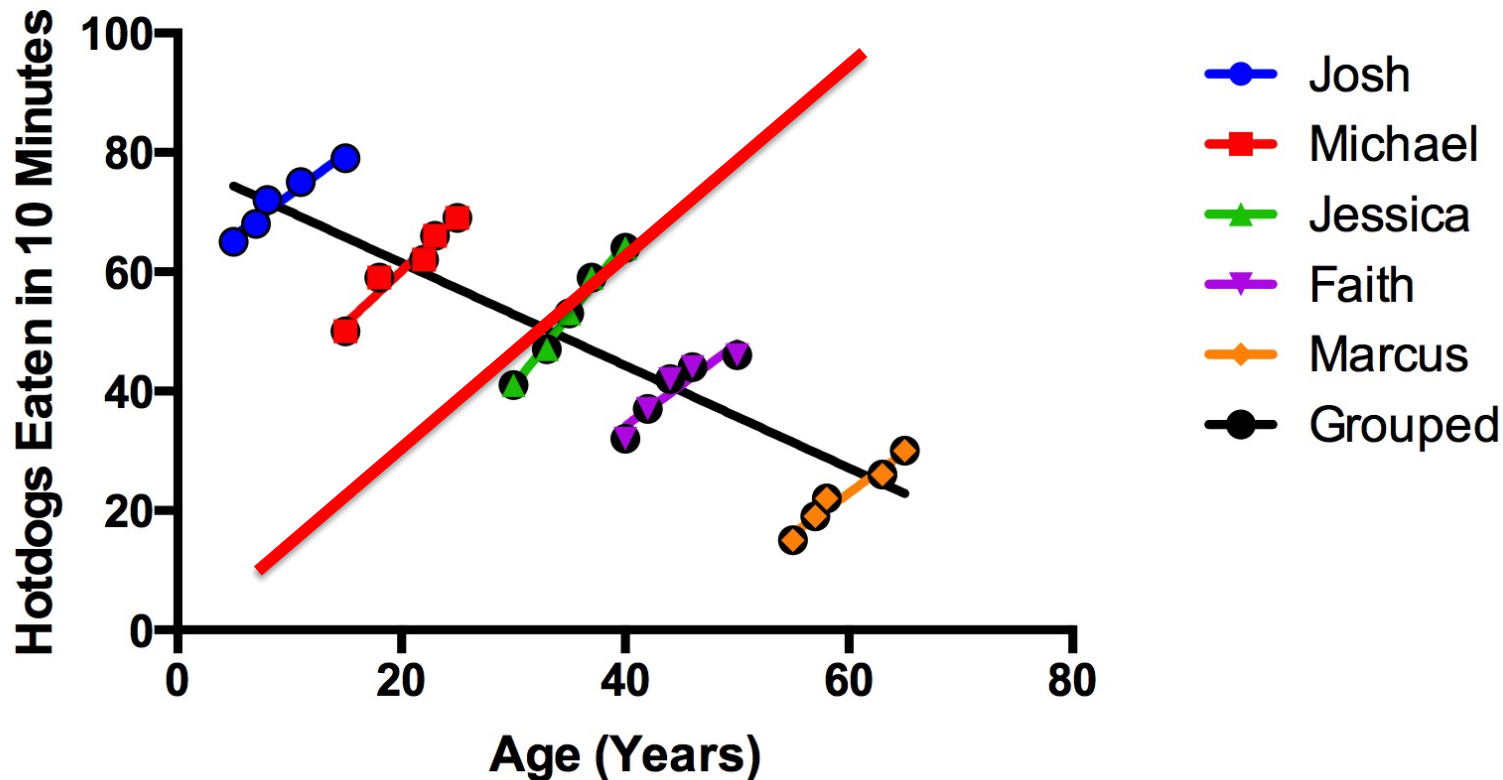
## Grouped Hotdogs Eaten in 10 Minutes vs. Age





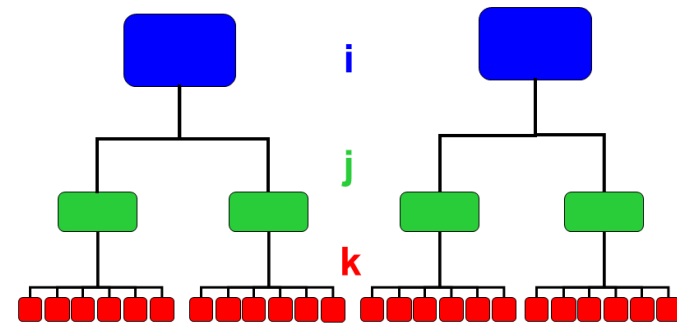
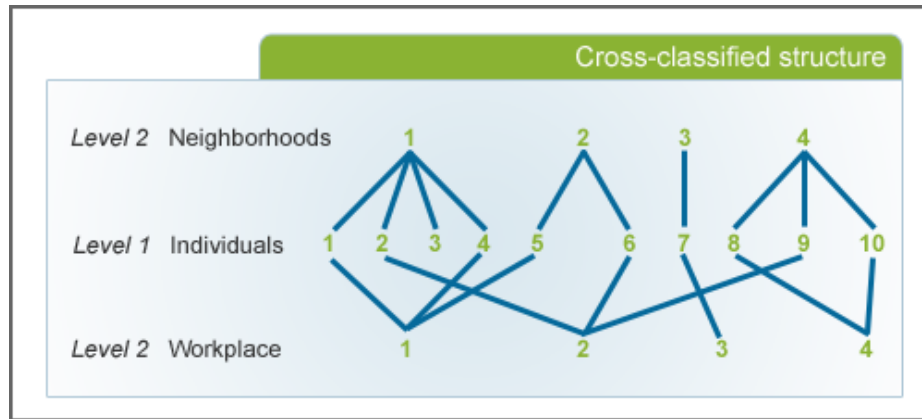
# Does accounting for clustering really matter?

## Grouped Hotdogs Eaten in 10 Minutes vs. Age



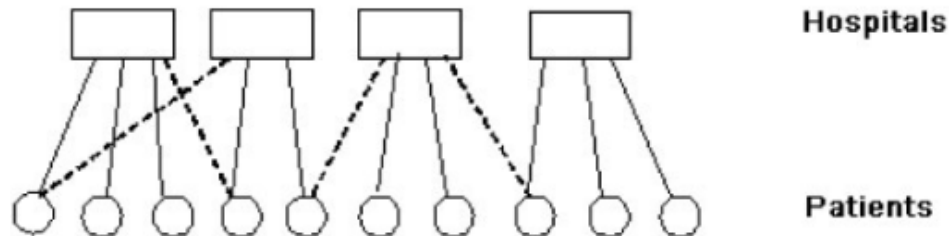


# Other structures



$$y_{ijk} = y_{\text{region}, \text{state}, \text{observation}}$$

**Multiple membership model: some patients attend more than one hospital**

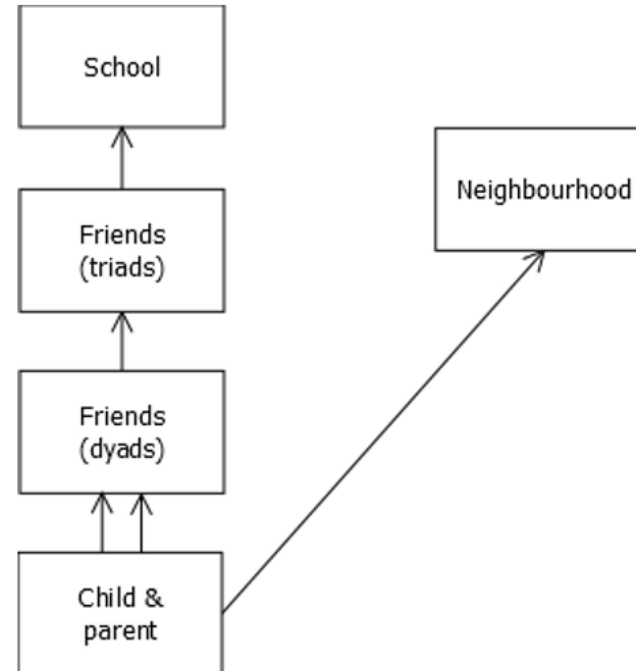




Article

## A Multilevel Analysis of Neighbourhood, School, Friend and Individual-Level Variation in Primary School Children's Physical Activity

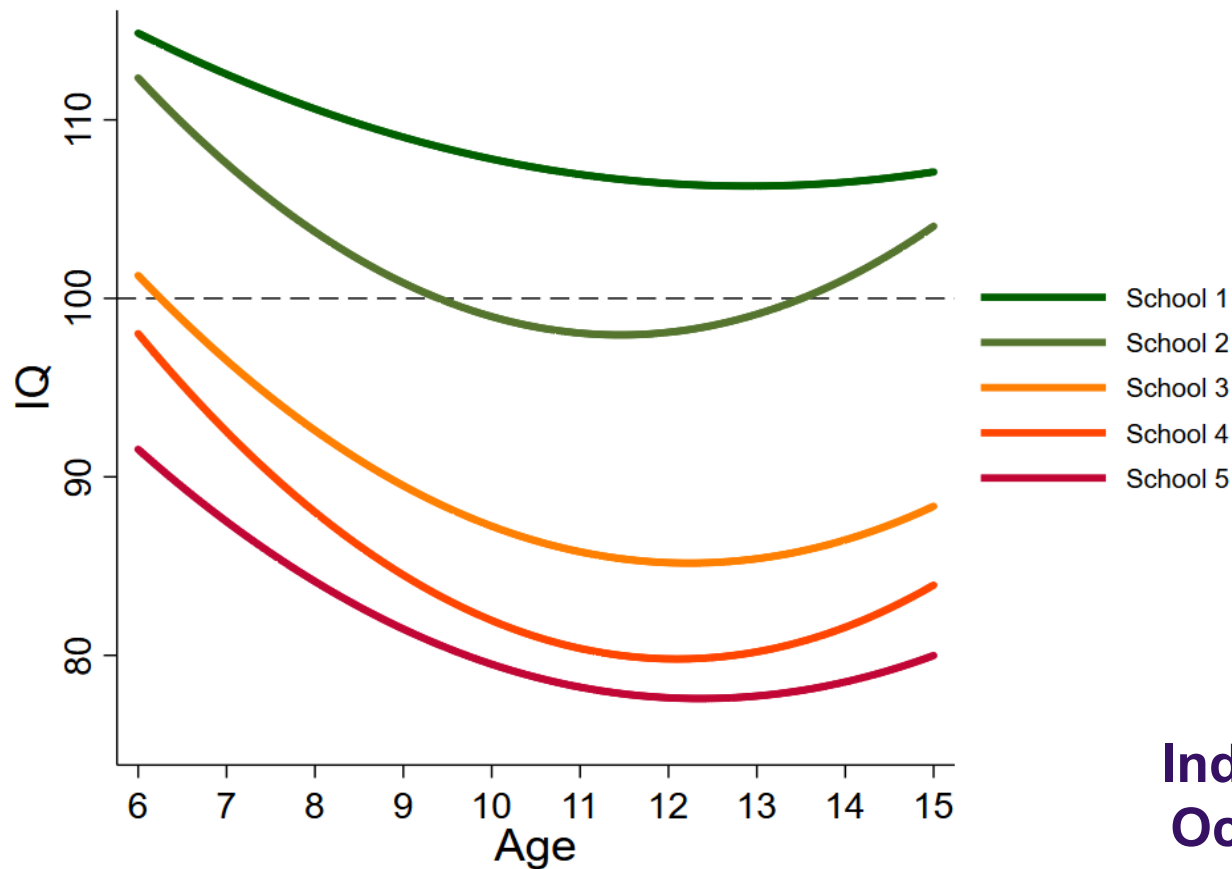
Ruth Salway <sup>1</sup>, Lydia Emm-Collison <sup>1</sup> , Simon J. Sebire <sup>1</sup>, Janice L. Thompson <sup>2</sup> ,  
Deborah A. Lawlor <sup>3,4</sup> and Russell Jago <sup>1,\*</sup>



**Figure 1.** Classification diagram for the multiple-membership multiple-classification model. An arrow indicates a nested relationship; a double arrow indicates multiple membership.



# Non-linear trajectories

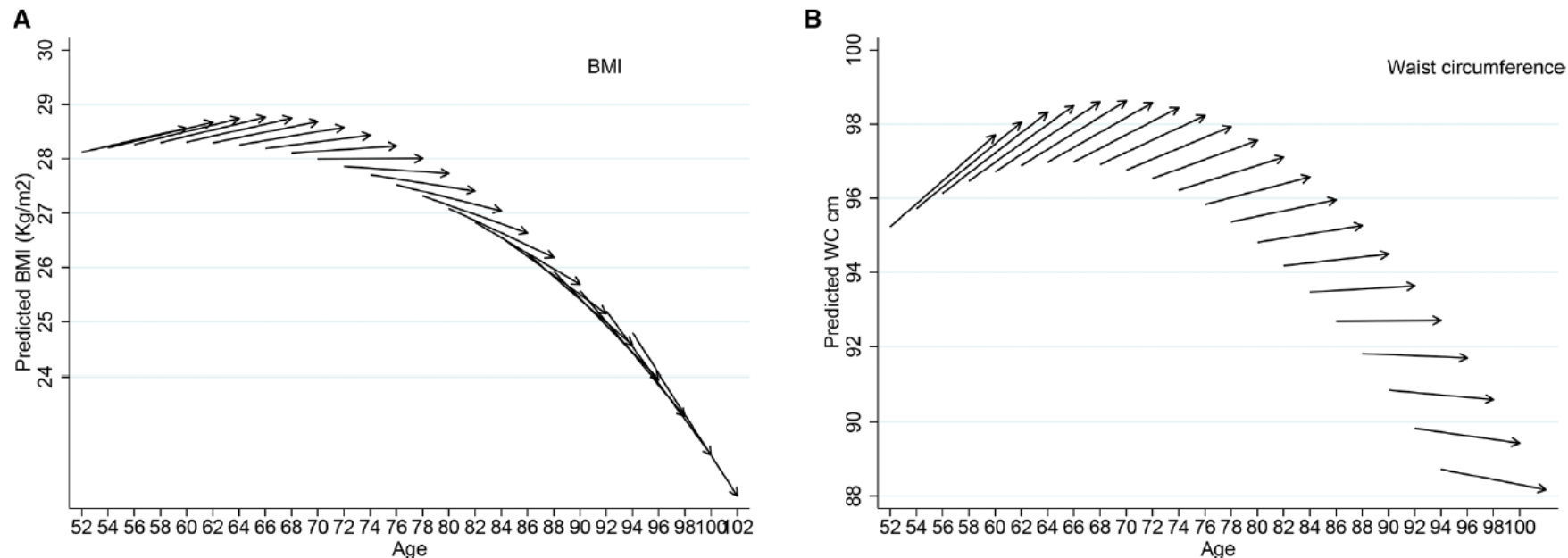


**Schools (level 3)**  
**Individuals (level 2)**  
**Occasions (level 1)**



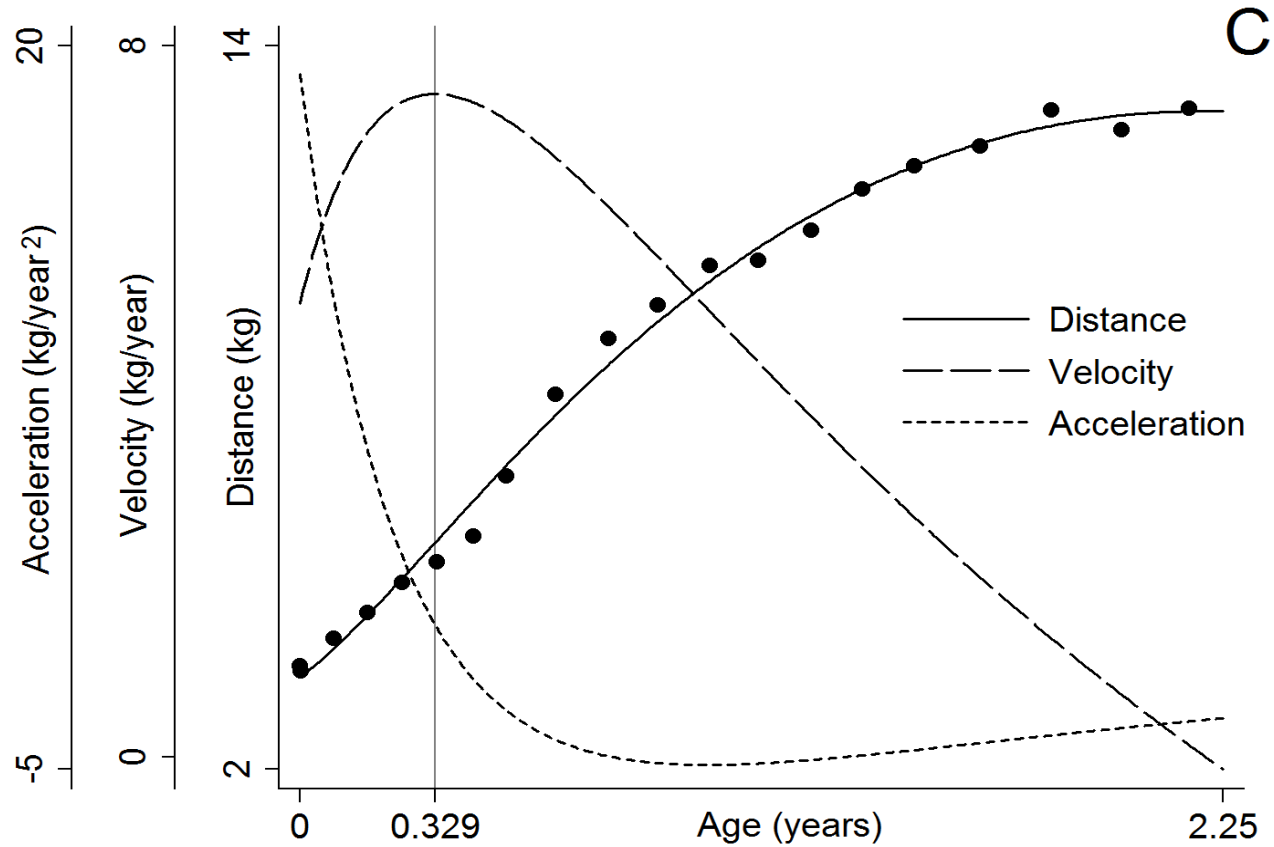


# Non-linear trajectories



**Figure 1** Vector graph showing 8-year ageing vectors of anthropometric markers, ELSA 2004–2005 to 2012–2013. BMI, body mass index; ELSA, English Longitudinal Study of Ageing; WC, waist circumference.

# Derivatives and traits

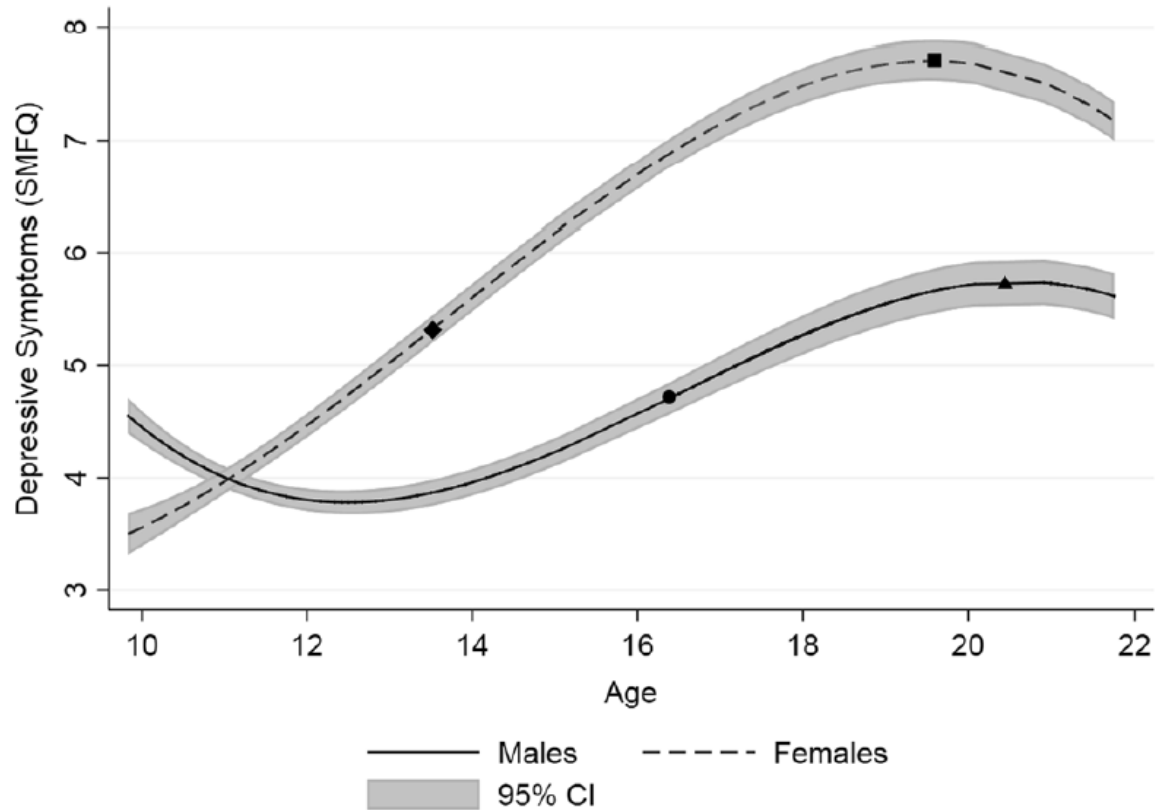




## Identifying Critical Points of Trajectories of Depressive Symptoms from Childhood to Young Adulthood

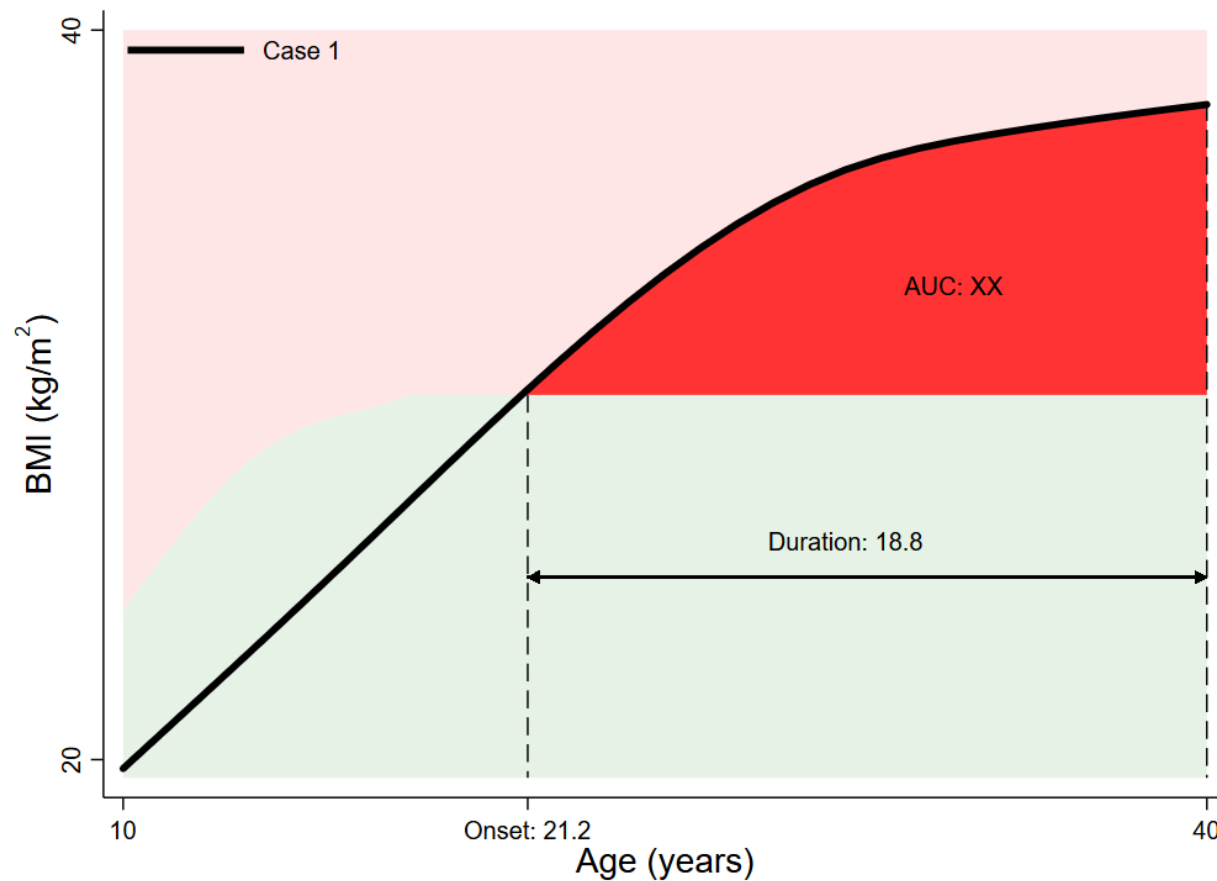
Alex S. F. Kwong<sup>1,2,3</sup> · David Manley<sup>1,2</sup> · Nicholas J. Timpson<sup>3,4</sup> · Rebecca M. Pearson<sup>3,4,5</sup> · Jon Heron<sup>3,4,5</sup> · Hannah Sallis<sup>3,4,5,6</sup> · Evie Stergiakouli<sup>3,4,7</sup> · Oliver S. P. Davis<sup>3,4</sup> · George Leckie<sup>2,8</sup>

Received: 7 December 2018 / Accepted: 8 December 2018 / Published online: 22 January 2019  
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# Derivatives and traits

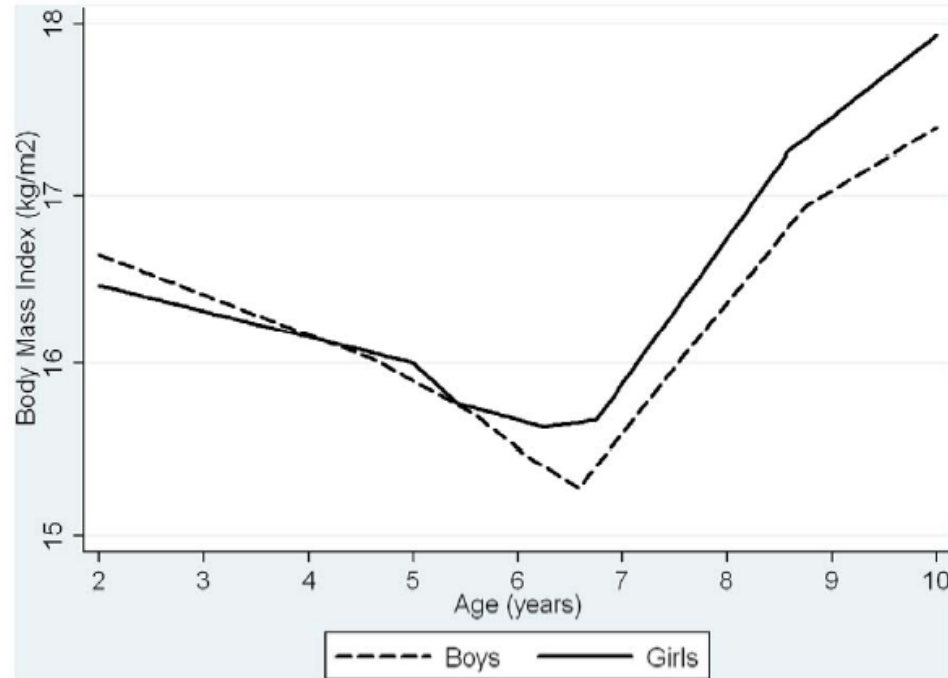




# Relating level 2 residuals to distal outcomes

## Changes in Ponderal Index and Body Mass Index across Childhood and Their Associations with Fat Mass and Cardiovascular Risk Factors at Age 15

Laura D. Howe<sup>1,2\*</sup>, Kate Tilling<sup>2</sup>, Li Benfield<sup>1,2</sup>, Jennifer Logue<sup>3</sup>, Naveed Sattar<sup>3</sup>, Andy R. Ness<sup>4</sup>, George Davey Smith<sup>1,2</sup>, Debbie A. Lawlor<sup>1,2</sup>



Article



Statistical Methods in Medical Research  
2016, Vol. 25(5) 1854–1874  
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DOI: 10.1177/0962280213503925  
smm.sagepub.com



### Linear spline multilevel models for summarising childhood growth trajectories: A guide to their application using examples from five birth cohorts

Laura D Howe,<sup>1</sup> Kate Tilling,<sup>2</sup> Alicia Matijasevich,<sup>3</sup> Emily S Petherick,<sup>4</sup> Ana Cristina Santos,<sup>5</sup> Lesley Fairley,<sup>4</sup> John Wright,<sup>4</sup> Iná S Santos,<sup>3</sup> Aluísio JD Barros,<sup>3</sup> Richard M Martin,<sup>2,6</sup> Michael S Kramer,<sup>7</sup> Natalia Bogdanovich,<sup>8</sup> Lidia Matush,<sup>8</sup> Henrique Barros<sup>5</sup> and Debbie A Lawlor<sup>1</sup>



Article



## **Joint modelling compared with two stage methods for analysing longitudinal data and prospective outcomes: A simulation study of childhood growth and BP**

**A Sayers,<sup>1</sup> J Heron,<sup>1</sup> ADAC Smith,<sup>1,2</sup> C Macdonald-Wallis,<sup>1,2</sup>  
MS Gilthorpe,<sup>3</sup> F Steele<sup>4</sup> and K Tilling<sup>1,2</sup>**

Statistical Methods in Medical Research

2017, Vol. 26(1) 437–452

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DOI: 10.1177/0962280214548822

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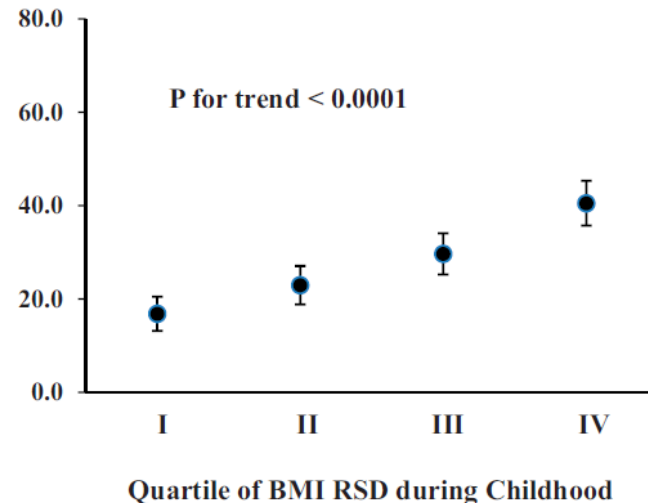
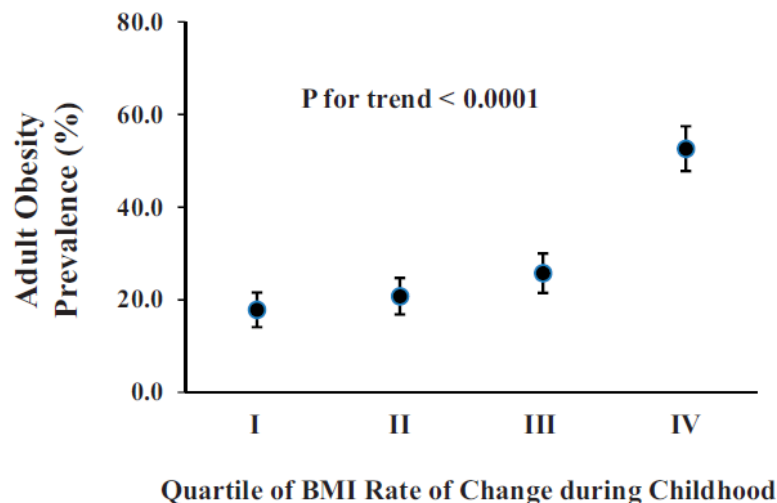




# Relating level 1 residuals to distal outcomes

## Variability and rapid increase in body mass index during childhood are associated with adult obesity

Shengxu Li,<sup>1\*</sup> Wei Chen,<sup>1</sup> Dianjianyi Sun,<sup>1,2</sup> Camilo Fernandez,<sup>1</sup> Jian Li,<sup>3</sup> Tanika Kelly,<sup>1</sup> Jiang He,<sup>1</sup> Marie Krousel-Wood<sup>1,4,5</sup> and Paul K Whelton<sup>1</sup>






# Relating level 1 residuals to distal outcomes

RESEARCH ARTICLE

WILEY **Statistics**  
in Medicine

## Estimating the association between blood pressure variability and cardiovascular disease: An application using the ARIC Study

Jessica K. Barrett<sup>1,2</sup>  | Raphael Huille<sup>2,3</sup> | Richard Parker<sup>4</sup> | Yuichiro Yano<sup>5</sup> | Michael Griswold<sup>6</sup>

<sup>1</sup>MRC Biostatistics Unit, University of Cambridge, Cambridge, UK

<sup>2</sup>Department of Public Health and Primary Care, University of Cambridge, Cambridge, UK

<sup>3</sup>École Nationale de la Statistique et de l'Administration Économique, Malakoff, France

<sup>4</sup>School of Social and Community Medicine, University of Bristol, Bristol, UK

<sup>5</sup>Department of Preventive Medicine, University of Mississippi Medical Center, Jackson, Mississippi

<sup>6</sup>Center of Biostatistics and Bioinformatics, University of Mississippi Medical Center, Jackson, Mississippi

### Correspondence

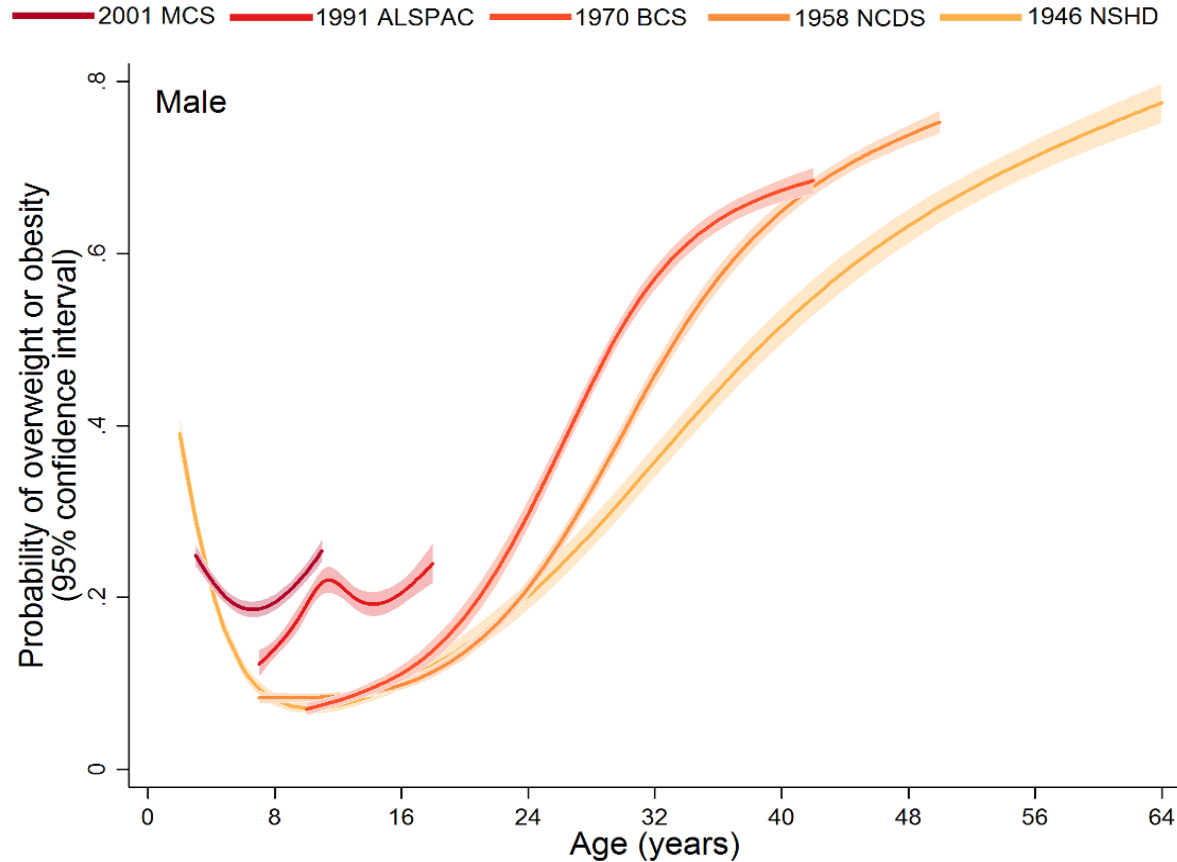
Jessica K. Barrett, MRC Biostatistics Unit,

The association between visit-to-visit systolic blood pressure variability and cardiovascular events has recently received a lot of attention in the cardiovascular literature. But, blood pressure variability is usually estimated on a person-by-person basis and is therefore subject to considerable measurement error. We demonstrate that hazard ratios estimated using this approach are subject to bias due to regression dilution, and we propose alternative methods to reduce this bias: a two-stage method and a joint model. For the two-stage method, in stage one, repeated measurements are modelled using a mixed effects model with a random component on the residual standard deviation (SD). The mixed effects model is used to estimate the blood pressure SD for each individual, which, in stage two, is used as a covariate in a time-to-event model. For the joint model, the mixed effects submodel and time-to-event submodel are fitted simultaneously using shared random effects. We illustrate the methods using data from the Atherosclerosis Risk in Communities study.





# Other outcome distributions



## Research Article

Received 8 March 2011,

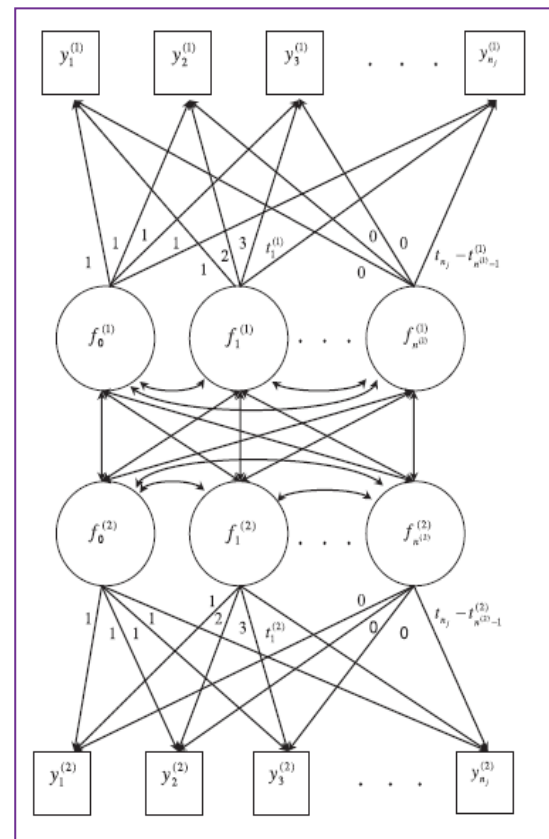
Accepted 6 March 2012

Published online 26 June 2012 in Wiley Online Library

(wileyonlinelibrary.com) DOI: 10.1002/sim.5385

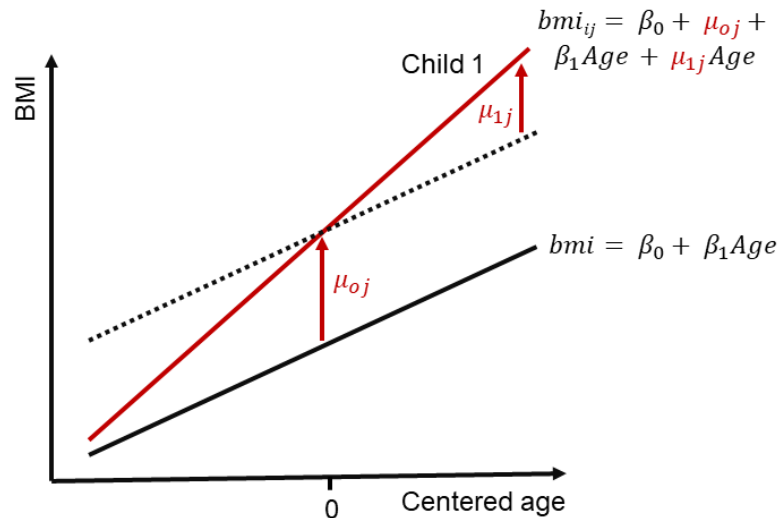
# Multivariate multilevel spline models for parallel growth processes: application to weight and mean arterial pressure in pregnancy

Corrie Macdonald-Wallis,<sup>a\*†</sup> Debbie A. Lawlor,<sup>a</sup> Tom Palmer<sup>a</sup>  
and Kate Tilling<sup>b</sup>



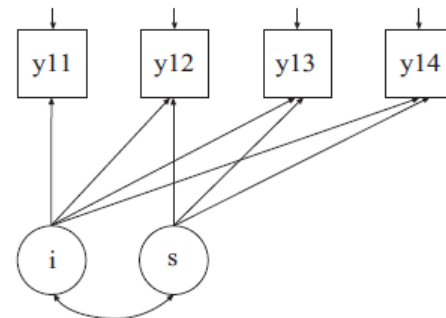


# Extend to SEM framework



Multilevel (growth curve) model

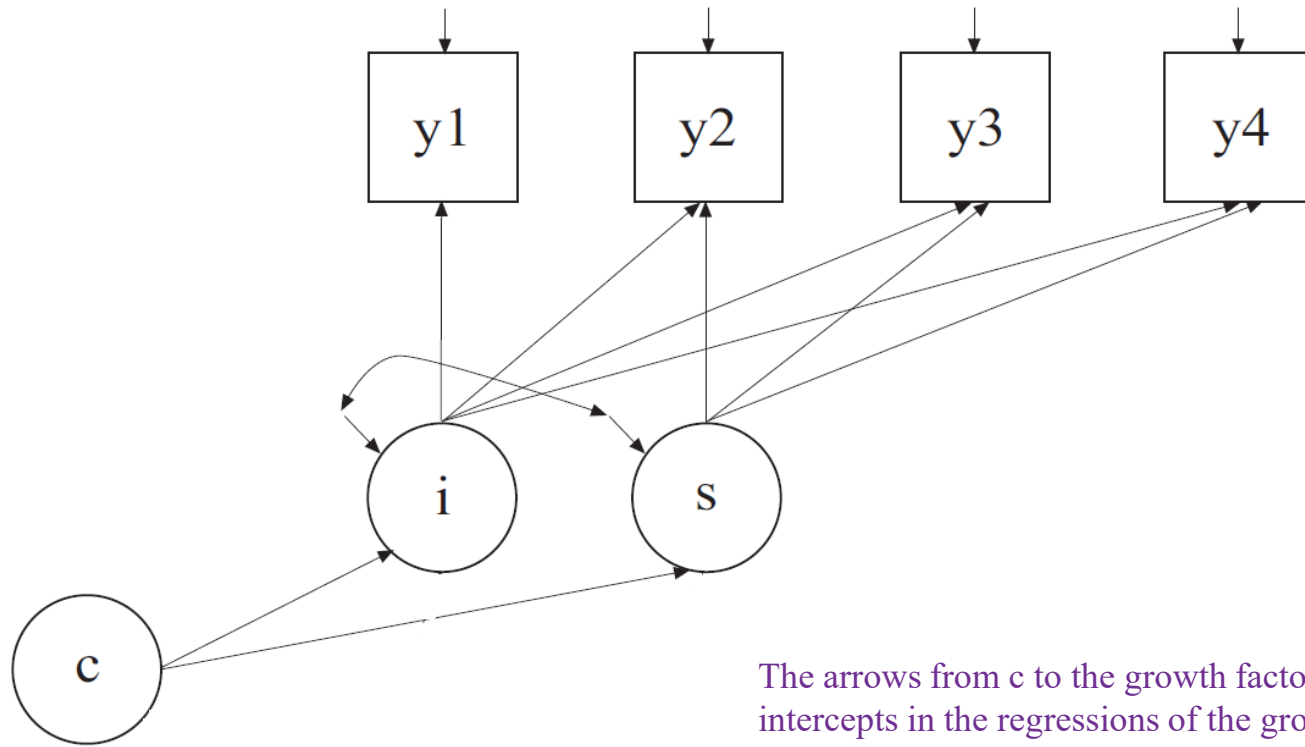
```
TITLE:    this is an example of a linear growth
          model for a continuous outcome
DATA:     FILE IS ex6.1.dat;
VARIABLE: NAMES ARE y11-y14 x1 x2 x31-x34;
          USEVARIABLES ARE y11-y14;
MODEL:    i s | y11@0 y12@1 y13@2 y14@3;
```



Latent (growth curve) model



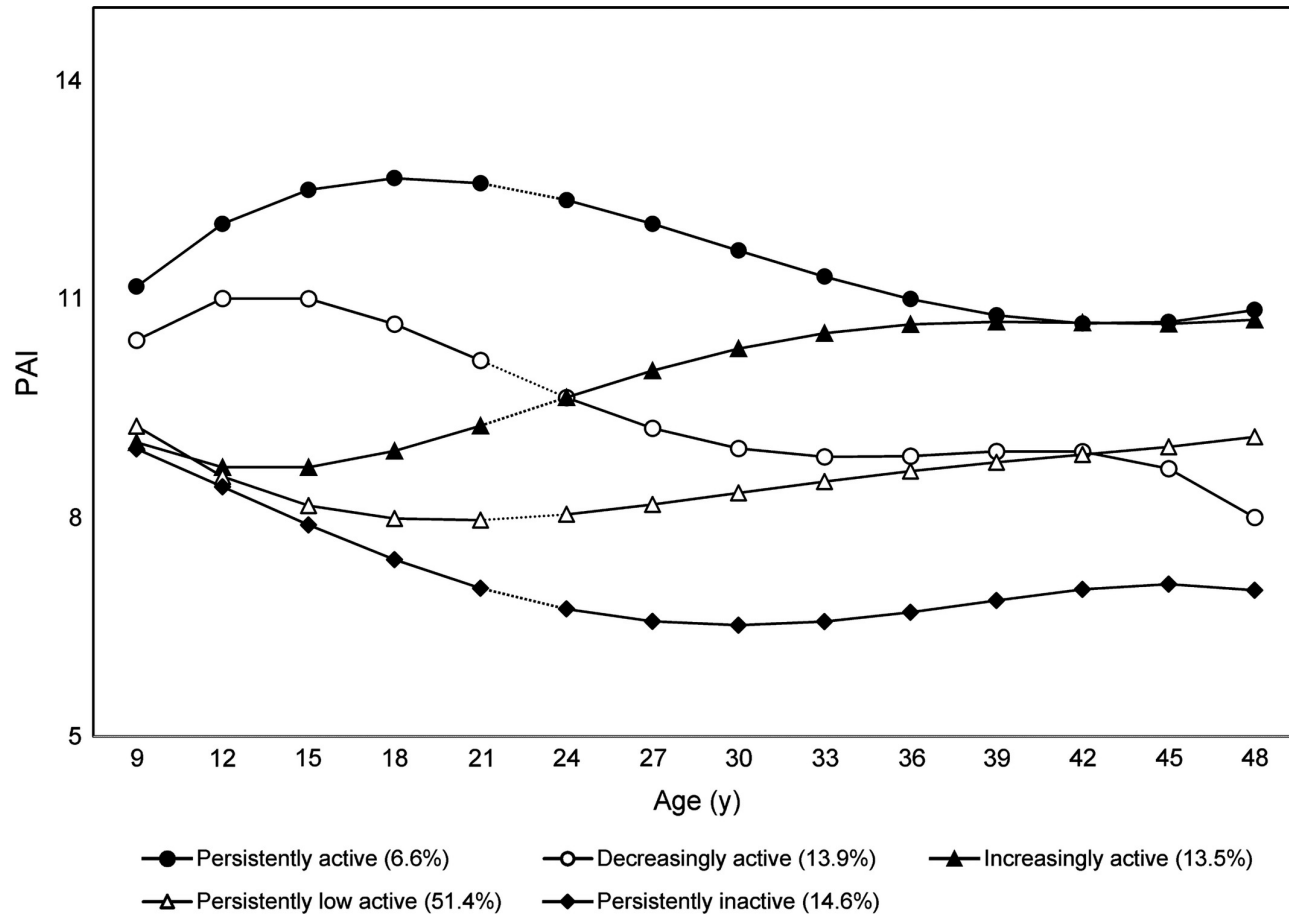
# Growth mixture model



The arrows from **c** to the growth factors **i** and **s** indicate that the intercepts in the regressions of the growth factors on **x** vary across the classes of **c**. This corresponds to the regressions of **i** and **s** on a set of dummy variables representing the categories of **c**.



# Growth mixture model



# Resources



## Centre for Multilevel Modelling

### Centre for Multilevel Modelling



People

The Centre for Multilevel Modelling (CMM) is a research centre based at the University of Bristol. Our researchers are drawn from the [School of Education](#) and [School of Geographical Sciences](#). We collaborate with a range of researchers working with [multilevel models](#).

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