# Approach 2: Maximizing Likelihood

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### 1. Simple Linear Regression

### **Model Structure**

Using the maximum likelihood approach, we set up the regression model probabilistically. Since we are treating the target as a random variable, we will capitalize it. As before, we assume

$$Y_n = \beta_0 + \beta_1 x_n + \epsilon_n,$$

only now we give  $\epsilon_n$  a distribution (we don't do the same for  $x_n$  since its value is known). Typically, we assume the  $\epsilon_n$  are independently Normally distributed with mean 0 and an unknown variance. That is,

$$\epsilon_n \overset{ ext{i.i.d.}}{\sim} \mathcal{N}(0,\sigma^2).$$

The assumption that the variance is identical across observations is called *homoskedasticity*. This is required for the following derivations, though there are *heteroskedasticity-robust* estimates that do not make this assumption.

Since  $eta_0$  and  $eta_1$  are fixed parameters and  $x_n$  is known, the only source of randomness in  $Y_n$  is  $\epsilon_n$ . Therefore,

$$Y_n \overset{ ext{i.i.d.}}{\sim} \mathcal{N}(eta_0 + eta_1 x_n, \sigma^2),$$

since a Normal random variable plus a constant is another Normal random variable with a shifted mean.

#### Parameter Estimation

The task of fitting the linear regression model then consists of estimating the parameters with maximum likelihood. The joint likelihood and log-likelihood across observations are as follows.

$$egin{aligned} L(eta_0,eta_1;Y_1,\ldots,Y_N) &= \prod_{n=1}^N L(eta_0,eta_1;Y_n) \ &= \prod_{n=1}^N rac{1}{\sqrt{2\pi}\sigma} \mathrm{exp}igg(-rac{(Y_n-(eta_0+eta_1x_n))^2}{2\sigma^2}igg) \ &\propto \mathrm{exp}igg(-\sum_{n=1}^N rac{(Y_n-(eta_0+eta_1x_n))^2}{2\sigma^2}igg) \ \log L(eta_0,eta_1;Y_1,\ldots,Y_N) &= -rac{1}{2\sigma^2}\sum_{n=1}^N (Y_n-(eta_0+eta_1x_n))^2. \end{aligned}$$

Our  $\hat{\beta}_0$  and  $\hat{\beta}_1$  estimates are the values that maximize the log-likelihood given above. Notice that this is equivalent to finding the  $\hat{\beta}_0$  and  $\hat{\beta}_1$  that minimize the RSS, our loss function from the previous section:

$$ext{RSS} = rac{1}{2} \sum_{n=1}^N \left( y_n - \left( \hat{eta}_0 + \hat{eta}_1 x_n 
ight) 
ight)^2.$$

In other words, we are solving the same optimization problem we did in the <u>last section</u>. Since it's the same problem, it has the same solution! (This can also of course be checked by differentiating and optimizing for  $\hat{\beta}_0$  and  $\hat{\beta}_1$ ). Therefore, as with the loss minimization approach, the parameter estimates from the likelihood maximization approach are

$$egin{aligned} \hat{eta}_0 &= ar{Y} - \hat{eta}_1 ar{x} \ \hat{eta}_1 &= rac{\sum_{n=1}^N (x_n - ar{x}) (Y_n - ar{Y})}{\sum_{n=1}^N (x_n - ar{x})^2}. \end{aligned}$$

## 2. Multiple Regression

Still assuming Normally-distributed errors but adding more than one predictor, we have

$$Y_n \overset{ ext{i.i.d.}}{\sim} \mathcal{N}(oldsymbol{eta}^ op \mathbf{x}_n, \sigma^2).$$

We can then solve the same maximum likelihood problem. Calculating the log-likelihood as we did above for simple linear regression, we have

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$$egin{aligned} \log L(eta_0,eta_1;Y_1,\ldots,Y_N) &= -rac{1}{2\sigma^2}\sum_{n=1}^N \left(Y_n-oldsymbol{eta}^ op \mathbf{x}_n
ight)^2 \ &= -rac{1}{2\sigma^2}(\mathbf{y}-\mathbf{X}\hat{oldsymbol{eta}})^ op (\mathbf{y}-\mathbf{X}\hat{oldsymbol{eta}}). \end{aligned}$$

Again, maximizing this quantity is the same as minimizing the RSS, as we did under the loss minimization approach. We therefore obtain the same solution:

$$\hat{oldsymbol{eta}} = (\mathbf{X}^{ op}\mathbf{X})^{-1}\mathbf{X}^{ op}\mathbf{y}.$$

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