

北京大学数学科学学院期末试题

2020-2021 年第一学期

考试科目 数理统计

考试时间 2020 年 1 月 11 日

姓名

学号

本试题共 8 道大题，满分 100 分。

1. (12 分) 设 X_1, \dots, X_n 是来自参数为 μ 的正态分布 $N(\mu, \sigma_0^2)$ 的简单随机样本， σ_0^2 已知。

(1) 求 μ 的最大似然估计；

(2) 求 μ 的矩估计；

(3) 求 Fisher 信息量 $I(\mu)$ ；

(4) 求 μ 的无偏估计的方差下界；

(5) (1) 中的最大似然估计是否是 μ 的最小方差无偏估计 (需说明理由)？

(6) 试找出 μ^2 的一个无偏估计。

1. (12 points) Suppose X_1, \dots, X_n are i.i.d random samples from normal distribution $N(\mu, \sigma_0^2)$, with σ_0^2 known.

(1) Find the MLE of μ ;

(2) Find the moment estimate of μ ;

(3) Calculate the Fisher information $I(\mu)$ of μ ;

(4) Find the variance lower bound of unbiased estimator for μ ;

(5) Is the MLE in (1) the UMVUE of μ (Please give your reason)?

(6) Try to find an unbiased estimator for μ^2 .

2. (12 分) 若随机变量 X 的分布密度可取下面的 $f_0(x)$ 或 $f_1(x)$:

$$f_0(x) = \begin{cases} 1 & \text{当 } 0 \leq x \leq 1 \\ 0 & \text{其他} \end{cases}; \quad f_1(x) = \begin{cases} \frac{C_1}{1+x^2} & \text{当 } 0 \leq x \leq 1 \\ 0 & \text{其他} \end{cases}$$

其中 C_1 是未知常数。基于 X 的一个观测值, 对检验问题 $H_0: f(x) = f_0(x) \leftrightarrow H_1: f(x) = f_1(x)$, 利用 N-P 引理求检验水平为 $\alpha = 0.1$ 的 UMP 检验 ϕ , 并求其第二类错误的概率。

2. (12 points) Suppose random variable X has density $f_0(x)$ or $f_1(x)$:

$$f_0(x) = \begin{cases} 1 & \text{when } 0 \leq x \leq 1 \\ 0 & \text{otherwise} \end{cases}; \quad f_1(x) = \begin{cases} \frac{C_1}{1+x^2} & \text{when } 0 \leq x \leq 1 \\ 0 & \text{otherwise} \end{cases}$$

where C_1 is unknown constant. Based on an observation of X , for the hypothesis testing $H_0: f(x) = f_0(x) \leftrightarrow H_1: f(x) = f_1(x)$, use N-P lemma to get a UMP test ϕ of significance level $\alpha = 0.1$, and calculate the probability of Type 2 error.

3. (13 分) 设 X_1, \dots, X_n 和 Y_1, \dots, Y_m 是分别来自的参数为 λ_1 和 λ_2 的指数分布的简单随机样本, λ_1 和 λ_2 未知, 即

$$p(X_i = x) = \lambda_1 e^{-\lambda_1 x}, \quad p(Y_j = y) = \lambda_2 e^{-\lambda_2 y}$$

试求 λ_1/λ_2 的置信水平为 $1 - \alpha$ 的置信区间。

3. (13 points) Suppose X_1, \dots, X_n and Y_1, \dots, Y_m are random samples from exponential distribution with parameter λ_1 and λ_2 respectively, where λ_1 and λ_2 are unknown, that is:

$$p(X_i = x) = \lambda_1 e^{-\lambda_1 x}, \quad p(Y_j = y) = \lambda_2 e^{-\lambda_2 y}$$

Give the $1 - \alpha$ confidence interval for λ_1/λ_2

4. (13 分) 设 $X_i, i = 1, 2, \dots, n$ 为参数为 λ 的泊松分布的独立同分布随机样本。 λ 的先验分布为 $Gamma(\alpha, \beta)$, 即 $\pi_{\alpha, \beta}(\lambda) = \frac{\beta^\alpha}{\Gamma(\alpha)} \lambda^{\alpha-1} e^{-\beta\lambda}$ 。

(1) 求 λ 的后验分布

(2) 在平方误差下, 求 λ 的贝叶斯估计量。

4. (13 points) Suppose $X_i, i = 1, 2, \dots, n$ are i.i.d samples from a Poisson distribution

with parameter λ . The prior of λ is $\pi(\lambda) \sim \text{Gamma}(\alpha, \beta)$, that is $\pi_{\alpha, \beta}(\lambda) = \frac{\beta^\alpha}{\Gamma(\alpha)} \lambda^{\alpha-1} e^{-\beta\lambda}$.

(1) Find the posterior distribution of λ .

(2) Find the Bayes estimator of λ using squared error loss.

5. (16 分) 设 ANOVA 模型 $Y_{ij} = \mu + \tau_i + \epsilon_{ij}$, $i = 1, 2, \dots, I$, $j = 1, \dots, n_i$, 其中 $\sum_{i=1}^I \tau_i = 0$, 且 ϵ_{ij} 独立, $E(\epsilon_{ij}) = 0$, $\text{Var}(\epsilon_{ij}) = \sigma^2$. 定义

$$SS_{TOT} = \sum_{i=1}^I \sum_{j=1}^{n_i} (Y_{ij} - \bar{Y}_{..})^2;$$

$$SS_W = \sum_{i=1}^I \sum_{j=1}^{n_i} (Y_{ij} - \bar{Y}_{i.})^2;$$

$$SS_B = \sum_{i=1}^I n_i (\bar{Y}_{i.} - \bar{Y}_{..})^2.$$

其中 $\bar{Y}_{i.} = \frac{1}{n_i} \sum_{j=1}^{n_i} Y_{ij}$, $\bar{Y}_{..} = \frac{1}{\sum_{i=1}^I n_i} \sum_{i=1}^I \sum_{j=1}^{n_i} Y_{ij}$.

(1) 证明 $SS_{TOT} = SS_W + SS_B$.

(2) 若 $n_i = J$, 证明 $E(SS_W) = I(J-1)\sigma^2$ 以及 $E(SS_B) = J \sum_{i=1}^I \tau_i^2 + (I-1)\sigma^2$

5. (16 points) Suppose an ANOVA model: $Y_{ij} = \mu + \tau_i + \epsilon_{ij}$, $i = 1, 2, \dots, I$ and $j = 1, \dots, n_i$, where $\sum_{i=1}^I \tau_i = 0$. The ϵ_{ij} are independent and $E(\epsilon_{ij}) = 0$, $\text{Var}(\epsilon_{ij}) = \sigma^2$. Define

$$SS_{TOT} = \sum_{i=1}^I \sum_{j=1}^{n_i} (Y_{ij} - \bar{Y}_{..})^2;$$

$$SS_W = \sum_{i=1}^I \sum_{j=1}^{n_i} (Y_{ij} - \bar{Y}_{i.})^2;$$

$$SS_B = \sum_{i=1}^I n_i (\bar{Y}_{i.} - \bar{Y}_{..})^2.$$

where $\bar{Y}_{i.} = \frac{1}{n_i} \sum_{j=1}^{n_i} Y_{ij}$ and $\bar{Y}_{..} = \frac{1}{\sum_{i=1}^I n_i} \sum_{i=1}^I \sum_{j=1}^{n_i} Y_{ij}$.

(1) Show that $SS_{TOT} = SS_W + SS_B$.

(2) If $n_i = J$ for $i = 1, 2, \dots, I$. Prove that $E(SS_W) = I(J-1)\sigma^2$ and $E(SS_B) = J \sum_{i=1}^I \tau_i^2 + (I-1)\sigma^2$

6. (16 分) 设参数 $\theta \in [0, 1]$, $X|\theta \sim p(x, \theta)$, 数据为 X_1, \dots, X_n , 损失函数为 $L(\theta, a) = (\theta - a)^2$.

(1) 试证明, 存在决策 δ_1 , 对 θ 的任意先验分布, 其贝叶斯风险 $\rho(\delta_1) \leq \frac{1}{4}$;

(2) 若 $X|\theta \sim B(1, \theta)$, 即参数为 θ 的两点分布, 先验分布 $\pi(\theta)$ 为 Beta 分布 $Be(\alpha, \beta)$, 即

$$\pi(\theta) = \frac{\Gamma(\alpha + \beta)}{\Gamma(\alpha)\Gamma(\beta)} x^{\alpha-1} (1-x)^{\beta-1}$$

求贝叶斯决策 δ^* .

6. (16 points) Suppose parameter $\theta \in [0, 1]$, $X|\theta \sim p(x, \theta)$, data are X_1, \dots, X_n , and the loss function is $L(\theta, a) = (\theta - a)^2$.

(1) Prove that there exists a decision rule δ_1 such that for any prior of θ , the Bayes risk $\rho(\delta_1) \leq \frac{1}{4}$;

(2) If $X|\theta \sim B(1, \theta)$, that is Binomial distribution with parameter θ , prior $\pi(\theta)$ is Beta distribution $Be(\alpha, \beta)$:

$$\pi(\theta) = \frac{\Gamma(\alpha + \beta)}{\Gamma(\alpha)\Gamma(\beta)} x^{\alpha-1} (1-x)^{\beta-1}$$

Find the Bayes decision rule δ^* .

7. (18 分) 考虑测量误差的问题, 假设有下列数据生成机制:

$$Y_i = \beta_0 + \beta_1 \xi_i + \varepsilon_i, \varepsilon_i \sim N(0, \sigma_\varepsilon^2);$$

$$X_i = \xi_i + \delta_i, \delta_i \sim N(0, \sigma_\delta^2)$$

其中 ε_i 和 δ_i 独立。在方程中, ξ_i 为常数, $\mathbb{E}X_i = \xi_i$ 。并且假设 $\sigma_\delta^2 = \lambda \sigma_\varepsilon^2$, 其中 λ 已知。

(1) 用极大似然方法求出 β_0 和 β_1 的估计, 分别记作 $\hat{\beta}_0(\lambda)$ 和 $\hat{\beta}_1(\lambda)$ 。

(2) 证明 $\lim_{\lambda \rightarrow 0} \hat{\beta}_1(\lambda) = S_{xy}/S_{xx}$, 即 y 对 x 回归的斜率。

(3) 证明 $\lim_{\lambda \rightarrow +\infty} \hat{\beta}_1(\lambda) = S_{yy}/S_{xy}$, 即 x 对 y 回归的斜率的倒数。

其中 $S_{uv} = \sum_{i=1}^n (u_i - \bar{u})(v_i - \bar{v})$ 。

7. (18 points) Consider a measurement error problem, suppose we have the following data generating process:

$$Y_i = \beta_0 + \beta_1 \xi_i + \varepsilon_i, \varepsilon_i \sim N(0, \sigma_\varepsilon^2);$$

$$X_i = \xi_i + \delta_i, \delta_i \sim N(0, \sigma_\delta^2)$$

where ε_i and δ_i are independent. In the equation, ξ_i are constants and $\mathbb{E}X_i = \xi_i$. Also, suppose $\sigma_\delta^2 = \lambda\sigma_\varepsilon^2$, where λ is known.

- (1) Using MLE method to get the estimates for β_0 and β_1 , denoted as $\hat{\beta}_0(\lambda)$ and $\hat{\beta}_1(\lambda)$ respectively.
- (2) Show that $\lim_{\lambda \rightarrow 0} \hat{\beta}_1(\lambda) = S_{xy}/S_{xx}$, the slope of the ordinary regression of y on x .
- (3) Show that $\lim_{\lambda \rightarrow +\infty} \hat{\beta}_1(\lambda) = S_{yy}/S_{xy}$, the reciprocal of the slope of the ordinary regression of x on y .

where $S_{uv} = \sum_{i=1}^n (u_i - \bar{u})(v_i - \bar{v})$.