

# 北京大学数学科学学院期末试题

2020–2021 年第一学期

考试科目 数理统计  
姓名 \_\_\_\_\_

考试时间 2020 年 1 月 11 日  
学号 \_\_\_\_\_

本试题共 8 道大题，满分 100 分。

1. (12 分) 设  $X_1, \dots, X_n$  是来自参数为  $\mu$  的正态分布  $N(\mu, \sigma_0^2)$  的简单随机样本， $\sigma_0^2$  已知。

- (1) 求  $\mu$  的最大似然估计；
- (2) 求  $\mu$  的矩估计；
- (3) 求 Fisher 信息量  $I(\mu)$ ；
- (4) 求  $\mu$  的无偏估计的方差下界；
- (5) (1) 中的最大似然估计是否是  $\mu$  的最小方差无偏估计 (需说明理由)？
- (6) 试找出  $\mu^2$  的一个无偏估计。

1. (12 points) Suppose  $X_1, \dots, X_n$  are i.i.d random samples from normal distribution  $N(\mu, \sigma_0^2)$ , with  $\sigma_0^2$  known.

- (1) Find the MLE of  $\mu$ ;
- (2) Find the moment estimate of  $\mu$ ;
- (3) Calculate the Fisher information  $I(\mu)$  of  $\mu$ ;
- (4) Find the variance lower bound of unbiased estimator for  $\mu$ ;
- (5) Is the MLE in (1) the UMVUE of  $\mu$  (Please give your reason)?
- (6) Try to find an unbiased estimator for  $\mu^2$ .

2. (12 分) 若随机变量  $X$  的分布密度可取下面的  $f_0(x)$  或  $f_1(x)$  :

$$f_0(x) = \begin{cases} 1 & \text{当 } 0 \leq x \leq 1 \\ 0 & \text{其他} \end{cases}; \quad f_1(x) = \begin{cases} \frac{C_1}{1+x^2} & \text{当 } 0 \leq x \leq 1 \\ 0 & \text{其他} \end{cases}$$

其中  $C_1$  是未知常数。基于  $X$  的一个观测值, 对检验问题  $H_0 : f(x) = f_0(x) \leftrightarrow H_1 : f(x) = f_1(x)$ , 利用 N-P 引理求检验水平为  $\alpha = 0.1$  的 UMP 检验  $\phi$ , 并求其第二类错误的概率。

2. (12 points) Suppose random variable  $X$  has density  $f_0(x)$  or  $f_1(x)$ :

$$f_0(x) = \begin{cases} 1 & \text{when } 0 \leq x \leq 1 \\ 0 & \text{otherwise} \end{cases}; \quad f_1(x) = \begin{cases} \frac{C_1}{1+x^2} & \text{when } 0 \leq x \leq 1 \\ 0 & \text{otherwise} \end{cases}$$

where  $C_1$  is unknown constant. Based on an observation of  $X$ , for the hypothesis testing  $H_0 : f(x) = f_0(x) \leftrightarrow H_1 : f(x) = f_1(x)$ , use N-P lemma to get a UMP test  $\phi$  of significance level  $\alpha = 0.1$ , and calculate the probability of Type 2 error.

3. (13 分) 设  $X_1, \dots, X_n$  和  $Y_1, \dots, Y_m$  是分别来自的参数为  $\lambda_1$  和  $\lambda_2$  的指数分布的简单随机样本,  $\lambda_1$  和  $\lambda_2$  未知, 即

$$p(X_i = x) = \lambda_1 e^{-\lambda_1 x}, \quad p(Y_j = y) = \lambda_2 e^{-\lambda_2 y}$$

试求  $\lambda_1/\lambda_2$  的置信水平为  $1 - \alpha$  的置信区间。

3. (13 points) Suppose  $X_1, \dots, X_n$  and  $Y_1, \dots, Y_m$  are random samples from exponential distribution with parameter  $\lambda_1$  and  $\lambda_2$  respectively, where  $\lambda_1$  and  $\lambda_2$  are unknown, that is:

$$p(X_i = x) = \lambda_1 e^{-\lambda_1 x}, \quad p(Y_j = y) = \lambda_2 e^{-\lambda_2 y}$$

Give the  $1 - \alpha$  confidence interval for  $\lambda_1/\lambda_2$

4. (13 分) 设  $X_i, i = 1, 2, \dots, n$  为参数为  $\lambda$  的泊松分布的独立同分布随机样本。 $\lambda$  的先验分布为  $Gamma(\alpha, \beta)$ , 即  $\pi_{\alpha, \beta}(\lambda) = \frac{\beta^\alpha}{\Gamma(\alpha)} \lambda^{\alpha-1} e^{-\beta \lambda}$ 。

(1) 求  $\lambda$  的后验分布

(2) 在平方误差下, 求  $\lambda$  的贝叶斯估计量。

4. (13 points) Suppose  $X_i, i = 1, 2, \dots, n$  are i.i.d samples from a Poisson distribution

with parameter  $\lambda$ . The prior of  $\lambda$  is  $\pi(\lambda) \sim \text{Gamma}(\alpha, \beta)$ , that is  $\pi_{\alpha, \beta}(\lambda) = \frac{\beta^\alpha}{\Gamma(\alpha)} \lambda^{\alpha-1} e^{-\beta\lambda}$ .

- (1) Find the posterior distribution of  $\lambda$ .
- (2) Find the Bayes estimator of  $\lambda$  using squared error loss.

5. (16 分) 设 ANOVA 模型  $Y_{ij} = \mu + \tau_i + \epsilon_{ij}$ ,  $i = 1, 2, \dots, I$ ,  $j = 1, \dots, n_i$ , 其中  $\sum_{i=1}^I \tau_i = 0$ , 且  $\epsilon_{ij}$  独立,  $E(\epsilon_{ij}) = 0$ ,  $Var(\epsilon_{ij}) = \sigma^2$ 。定义

$$\begin{aligned} SS_{TOT} &= \sum_{i=1}^I \sum_{j=1}^{n_i} (Y_{ij} - \bar{Y}_{..})^2; \\ SS_W &= \sum_{i=1}^I \sum_{j=1}^{n_i} (Y_{ij} - \bar{Y}_{i.})^2; \\ SS_B &= \sum_{i=1}^I n_i (\bar{Y}_{i.} - \bar{Y}_{..})^2. \end{aligned}$$

其中  $\bar{Y}_{i.} = \frac{1}{n_i} \sum_{j=1}^{n_i} Y_{ij}$ ,  $\bar{Y}_{..} = \frac{1}{\sum_{i=1}^I n_i} \sum_{i=1}^I \sum_{j=1}^{n_i} Y_{ij}$ .

(1) 证明  $SS_{TOT} = SS_W + SS_B$ .

(2) 若  $n_i = J$ , 证明  $E(SS_W) = I(J-1)\sigma^2$  以及  $E(SS_B) = J \sum_{i=1}^I \tau_i^2 + (I-1)\sigma^2$

5. (16 points) Suppose an ANOVA model:  $Y_{ij} = \mu + \tau_i + \epsilon_{ij}$ ,  $i = 1, 2, \dots, I$  and  $j = 1, \dots, n_i$ , where  $\sum_{i=1}^I \tau_i = 0$ . The  $\epsilon_{ij}$  are independent and  $E(\epsilon_{ij}) = 0$ ,  $Var(\epsilon_{ij}) = \sigma^2$ . Define

$$\begin{aligned} SS_{TOT} &= \sum_{i=1}^I \sum_{j=1}^{n_i} (Y_{ij} - \bar{Y}_{..})^2; \\ SS_W &= \sum_{i=1}^I \sum_{j=1}^{n_i} (Y_{ij} - \bar{Y}_{i.})^2; \\ SS_B &= \sum_{i=1}^I n_i (\bar{Y}_{i.} - \bar{Y}_{..})^2. \end{aligned}$$

where  $\bar{Y}_{i.} = \frac{1}{n_i} \sum_{j=1}^{n_i} Y_{ij}$  and  $\bar{Y}_{..} = \frac{1}{\sum_{i=1}^I n_i} \sum_{i=1}^I \sum_{j=1}^{n_i} Y_{ij}$ .

(1) Show that  $SS_{TOT} = SS_W + SS_B$ .

(2) If  $n_i = J$  for  $i = 1, 2, \dots, I$ . Prove that  $E(SS_W) = I(J-1)\sigma^2$  and  $E(SS_B) = J \sum_{i=1}^I \tau_i^2 + (I-1)\sigma^2$

6. (16 分) 设参数  $\theta \in [0, 1]$ ,  $X|\theta \sim p(x, \theta)$ , 数据为  $X_1, \dots, X_n$ , 损失函数为  $L(\theta, a) = (\theta - a)^2$ 。

(1) 试证明, 存在决策  $\delta_1$ , 对  $\theta$  的任意先验分布, 其贝叶斯风险  $\rho(\delta_1) \leq \frac{1}{4}$ ;

(2) 若  $X|\theta \sim B(1, \theta)$ , 即参数为  $\theta$  的两点分布, 先验分布  $\pi(\theta)$  为 Beta 分布  $Be(\alpha, \beta)$ , 即

$$\pi(\theta) = \frac{\Gamma(\alpha + \beta)}{\Gamma(\alpha)\Gamma(\beta)} x^{\alpha-1} (1-x)^{\beta-1}$$

求贝叶斯决策  $\delta^*$ 。

6. (16 points) Suppose parameter  $\theta \in [0, 1]$ ,  $X|\theta \sim p(x, \theta)$ , data are  $X_1, \dots, X_n$ , and the loss function is  $L(\theta, a) = (\theta - a)^2$ .

(1) Prove that there exists a decision rule  $\delta_1$  such that for any prior of  $\theta$ , the Bayes risk  $\rho(\delta_1) \leq \frac{1}{4}$ ;

(2) If  $X|\theta \sim B(1, \theta)$ , that is Binomial distribution with parameter  $\theta$ , prior  $\pi(\theta)$  is Beta distribution  $Be(\alpha, \beta)$ :

$$\pi(\theta) = \frac{\Gamma(\alpha + \beta)}{\Gamma(\alpha)\Gamma(\beta)} x^{\alpha-1} (1-x)^{\beta-1}$$

Find the Bayes decision rule  $\delta^*$ .

7. (18 分) 考虑测量误差的问题, 假设有下述数据生成机制:

$$Y_i = \beta_0 + \beta_1 \xi_i + \varepsilon_i, \varepsilon_i \sim N(0, \sigma_\varepsilon^2);$$

$$X_i = \xi_i + \delta_i, \delta_i \sim N(0, \sigma_\delta^2)$$

其中  $\varepsilon_i$  和  $\delta_i$  独立。在方程中,  $\xi_i$  为常数,  $\mathbb{E}X_i = \xi_i$ 。并且假设  $\sigma_\delta^2 = \lambda \sigma_\varepsilon^2$ , 其中  $\lambda$  已知。

(1) 用极大似然方法求出  $\beta_0$  和  $\beta_1$  的估计, 分别记作  $\hat{\beta}_0(\lambda)$  和  $\hat{\beta}_1(\lambda)$ 。

(2) 证明  $\lim_{\lambda \rightarrow 0} \hat{\beta}_1(\lambda) = S_{xy}/S_{xx}$ , 即  $y$  对  $x$  回归的斜率。

(3) 证明  $\lim_{\lambda \rightarrow +\infty} \hat{\beta}_1(\lambda) = S_{yy}/S_{xy}$ , 即  $x$  对  $y$  回归的斜率的倒数。

其中  $S_{uv} = \sum_{i=1}^n (u_i - \bar{u})(v_i - \bar{v})$ 。

7. (18 points) Consider a measurement error problem, suppose we have the following data generating process:

$$Y_i = \beta_0 + \beta_1 \xi_i + \varepsilon_i, \varepsilon_i \sim N(0, \sigma_\varepsilon^2);$$

$$X_i = \xi_i + \delta_i, \delta_i \sim N(0, \sigma_\delta^2)$$

where  $\varepsilon_i$  and  $\delta_i$  are independent. In the equation,  $\xi_i$  are constants and  $\mathbb{E}X_i = \xi_i$ . Also, suppose  $\sigma_\delta^2 = \lambda\sigma_\varepsilon^2$ , where  $\lambda$  is known.

- (1) Using MLE method to get the estimates for  $\beta_0$  and  $\beta_1$ , denoted as  $\hat{\beta}_0(\lambda)$  and  $\hat{\beta}_1(\lambda)$  respectively.
- (2) Show that  $\lim_{\lambda \rightarrow 0} \hat{\beta}_1(\lambda) = S_{xy}/S_{xx}$ , the slope of the ordinary regression of  $y$  on  $x$ .
- (3) Show that  $\lim_{\lambda \rightarrow +\infty} \hat{\beta}_1(\lambda) = S_{yy}/S_{xy}$ , the reciprocal of the slope of the ordinary regression of  $x$  on  $y$ .

where  $S_{uv} = \sum_{i=1}^n (u_i - \bar{u})(v_i - \bar{v})$ .