## 1 The Simulation Study

We compare our method to existing ones in a simulation study. We generate thirty datasets using simulation parameters calculated from real data and analyze the pseudo-data using our method and the popular R language packages edgeR, baySeq, and ShrinkBayes [6] [1]. Using ROC (receiver operating characteristic) curves, we rank the methods' abilities to identify heterosis genes.

#### 2 Simulated Data

We begin with a real heterosis RNA-seq dataset from a study by Paschold, Jia, Marcon, and others [5]. We select four libraries from each parent genotype and from the hybrid genotype, totaling twelve libraries for analysis. After library selection, we trim low-count features (genes): that is, we remove all the features with mean expression level below  $\exp(1)$  or with more than three zero counts, leaving 27888 features. Using the calcNormFactors(), estimateGLMTagwiseDisp(), and glmFit() functions in the edgeR package, we calculate normalization factors  $c_1, \ldots, c_{12}$ , dispersion parameters  $\psi_f$  for feature  $f = 1, \ldots, 27888$ , and main effects  $\mu_{f,t}$  for each f and treatment group t = 1 (parent 1), 2 (parent 2), 3 (hybrid). These estimates serve as simulation parameters for all of our thirty pseudo-datasets.

To simulate a dataset with N libraries per treatment group (3N total libraries), the count for feature f and library i is drawn from a NB(exp( $c_{\lceil 4i/N \rceil} + \mu_{f,\lceil i/N \rceil}$ ),  $\psi_f$ ) distribution independently of the other counts. Note that feature f is a heterosis feature if  $\mu_{f,3} > \max(\mu_{f,1}, \mu_{f,2})$  or if  $\mu_{f,3} < \min(\mu_{f,1}, \mu_{f,2})$ . Lastly, we apply the same trimming procedure as before and select a random subset of 25000 of the remaining features. For that dataset, we maintain a "truth vector"  $H = (h_1, \ldots, h_{25000})$ , where  $h_f = 1$  if feature f of the simulated dataset is a heterosis feature and  $h_f = 0$  otherwise.

We simulate 30 datasets total: 10 with N=4, 10 with N=8, and 10 with N=16.

# 3 edgeR

edgeR is one of the most popular R packages in RNA-sequencing data analysis. Its newest implementation applies a negative binomial loglinear model to the data. It uses a Cox-Reid adjusted profile likelihood to estimate dispersion parameters, and in the case of estimateGLMTagwiseDisp(), shrinks the final dispersion estimates towards those of neighboring features on a common trend. It then estimates main effects using a Fisher scoring algorithm [7] [4].

Using the calcNormFactors(), estimateGLMTagwiseDisp(), and glmFit() functions in the edgeR package, we calculate normalization factor estimates  $\hat{c}_i$  for  $i=1,\ldots,3N$ , dispersion parameter estimates  $\hat{\psi}_f$  for feature  $f=1,\ldots,25000$ , and main effects  $\hat{\mu}_{f,t}$  for each f and treatment group t=1 (parent 1), 2 (parent 2), 3 (hybrid). Using the glmLRT() function, we use likelihood ratio tests to perform the following hypothesis tests.

$$H_{0,f,1}: \mu_{f,3} - \mu_{f,1} = 0 \text{ vs } H_{a,f,1}: \mu_{f,3} - \mu_{f,1} \neq 0$$
  
 $H_{0,f,2}: \mu_{f,3} - \mu_{f,2} = 0 \text{ vs } H_{a,f,1}: \mu_{f,3} - \mu_{f,2} \neq 0$ 

We obtain p-values  $p_{f,1}$  and  $p_{f,2}$ , respectively, from each of the above tests. To translate the results into a test for heterosis for each feature, we compute the following p-values

$$p_{f, \text{edgeR}} = \begin{cases} p_{f,1}/2 & \widehat{\mu}_{f,3} < \widehat{\mu}_{f,1} \leq \widehat{\mu}_{f,2} \text{ or } \widehat{\mu}_{f,3} > \widehat{\mu}_{f,1} \geq \widehat{\mu}_{f,2} \\ p_{f,2}/2 & \widehat{\mu}_{f,3} < \widehat{\mu}_{f,2} \leq \widehat{\mu}_{f,1} \text{ or } \widehat{\mu}_{f,3} > \widehat{\mu}_{f,2} \geq \widehat{\mu}_{f,1} \\ 1 & \widehat{\mu}_{f,1} \leq \widehat{\mu}_{f,3} \leq \widehat{\mu}_{f,2} \text{ or } \widehat{\mu}_{f,2} \leq \widehat{\mu}_{f,3} \leq \widehat{\mu}_{f,1} \end{cases}$$

## 4 ShrinkBayes

ShrinkBayes is based on the inla package, which applies an integrated nested Laplace approximation to fit models in empirical Bayes fashion. ShrinkBayes applies a zero-inflated negative binomial model with normal distributions as priors [8]. In our usage, we make the following reparameterization

$$\begin{split} \phi_f &= \frac{\mu_{f,1} + \mu_{f,2}}{2} \qquad \text{(parental mean)} \\ \alpha_f &= \frac{\mu_{f,2} - \mu_{f,1}}{2} \qquad \text{(half parental difference)} \\ \delta_f &= \mu_{f,3} - \frac{\mu_{f,1} + \mu_{f,2}}{2} \qquad \text{(hybrid effect)} \end{split}$$

We use the ShrinkSeq() and FitAllShrink() functions to fit the model and use inla.make.lincombs(), BFUpdatePosterior(), and SummaryWrap() to calculate posterior probabilities  $P(\delta_f + \alpha_f > 0 \mid \text{data})$ ,  $P(\delta_f - \alpha_f > 0 \mid \text{data})$ ,  $P(\delta_f - \alpha_f < 0 \mid \text{data})$ , and  $P(\delta_f + \alpha_f < 0 \mid \text{data})$ , along with estimates of  $\phi_f$ ,  $\alpha_f$ , and  $\delta_f$  for  $f = 1, \ldots, 25000$ . Using this information, we calculate the posterior probability that each feature f is not a heterosis feature,

$$p_{f,\mathtt{ShrinkBayes}} = \begin{cases} 1 & |\widehat{\delta}_f| < |\widehat{\alpha}_f|. \ \mathtt{Otherwise}, \\ P(\delta_f + \alpha_f > 0 \mid \mathtt{data}) & \widehat{\delta}_f > -\widehat{\alpha}_f \\ P(\delta_f - \alpha_f > 0 \mid \mathtt{data}) & \widehat{\delta}_f > \widehat{\alpha}_f \\ P(\delta_f - \alpha_f < 0 \mid \mathtt{data}) & \widehat{\delta}_f < \widehat{\alpha}_f \\ P(\delta_f + \alpha_f < 0 \mid \mathtt{data}) & \widehat{\delta}_f < -\widehat{\alpha}_f \end{cases}$$

## 5 baySeq

baySeq uses an empirical Bayes procedure to calculate the posterior probabilities that each feature follows each of the multiple models supplied by the user [2]. In the baySeq framework, a user-supplied model is an an assignment of libraries to treatment groups. In the case of heterosis experiments, it is appropriate to consider the following five models.

$$\begin{split} M_1: \mu_{f,1} &= \mu_{f,2} = \mu_{f,3} \\ M_2: \mu_{f,1} &= \mu_{f,2} \\ M_3: \mu_{f,1} &= \mu_{f,3} \\ M_4: \mu_{f,2} &= \mu_{f,3} \\ M_5: \text{All } \mu_{f,t}\text{'s are distinct.} \end{split}$$

Now, let  $p_{f,baySeq}$  be the posterior probability that feature f of a given simulated dataset is not a heterosis feature. We can calculate

$$p_{f,\mathtt{baySeq}} = \begin{cases} 1 & \widehat{\mu}_{f,1} \leq \widehat{\mu}_{f,3} \leq \widehat{\mu}_{f,2} \text{ or } \widehat{\mu}_{f,2} \leq \widehat{\mu}_{f,3} \leq \widehat{\mu}_{f,3} \leq \widehat{\mu}_{f,1} \\ P(M_1|\mathtt{data}) + P(M_3|\mathtt{data}) + P(M_4|\mathtt{data}) & \mathtt{otherwise} \end{cases}$$

We calculate estimates  $\hat{\mu}_{f,t}$  for  $f=1,\ldots,25000$  and t=1,2,3 using edgeR as described previously.

#### 6 ROC curves

We use receiver operating characteristic (ROC) curves to compare the effectiveness of our method versus edgeR, ShrinkBayes, and baySeq. A ROC curve is a tool for measuring the effectiveness of a binary

classifier. It is a graph of the true positive rate (TPR) of detection against the false positive rate (FPR), so a high area under the curve (AUC) is favorable. Landau and Liu [3] describe most of the details of calculation. However, note that in this study, posterior probabilities sometimes replace p-values, and we test for heterosis, not differential expression.

### References

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