


Probability and Statistics

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Outline

- Probability Distributions 
- Joint and Conditional Probability Distributions
- Bayes' Rule
- Mean and Variance
- Properties of Gaussian Distribution
- Maximum Likelihood Estimation

Probability

- A **sample space S** is the set of all possible outcomes of a conceptual or physical, repeatable experiment. (S can be finite or infinite.)
 - E.g., S may be the set of all possible outcomes of a dice roll: S
(1 2 3 4 5 6)
 - E.g., S may be the set of all possible nucleotides of a DNA site: S
(A C G T)
- E.g., S may be the set of all possible time-space positions of an aircraft on a radar screen.
- An **Event A** is any subset of S
 - Seeing "1" or "6" in a dice roll; observing a "G" at a site; UA007 in space-time interval



Three Key Ingredients in Probability Theory

A **sample space** is a collection of all possible **outcomes**

Random variables X represents **outcomes** in sample space

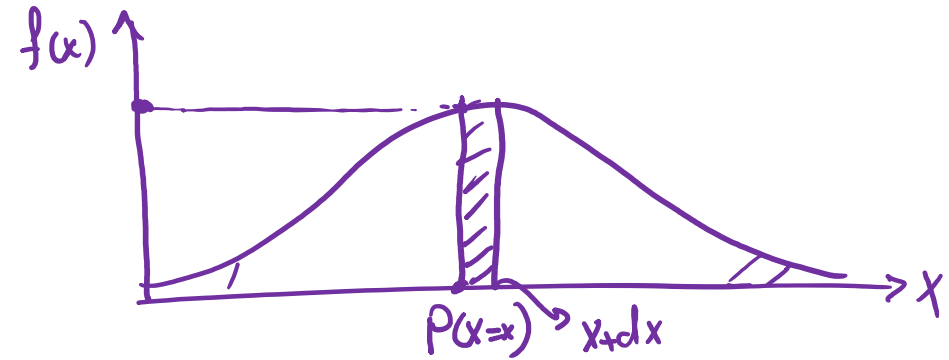
$$P(X=2) = \frac{1}{6}$$

Probability of a random variable to happen

$$p(x) = p(X = x)$$

$$p(x) \geq 0$$

density or
likelihood



Continuous variable

Continuous probability distribution

Probability density function

Density or likelihood value

Temperature (real number)

Gaussian Distribution

$$\int_x p(x) dx = 1$$

Discrete variable

Discrete probability distribution

Probability mass function

Probability value

Coin flip (integer)

Bernoulli distribution

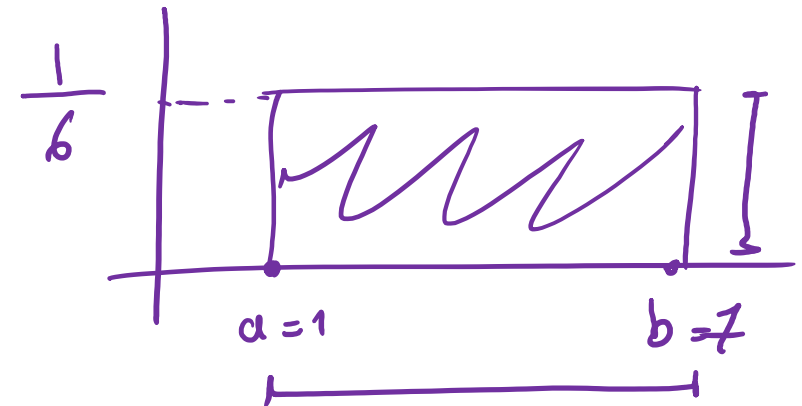
$$\sum_{x \in A} p(x) = 1$$

Continuous Probability Functions

- Examples:

- Uniform Density Function:

$$f_x(x) = \begin{cases} \frac{1}{b-a} & \text{for } a \leq x \leq b \\ 0 & \text{otherwise} \end{cases}$$

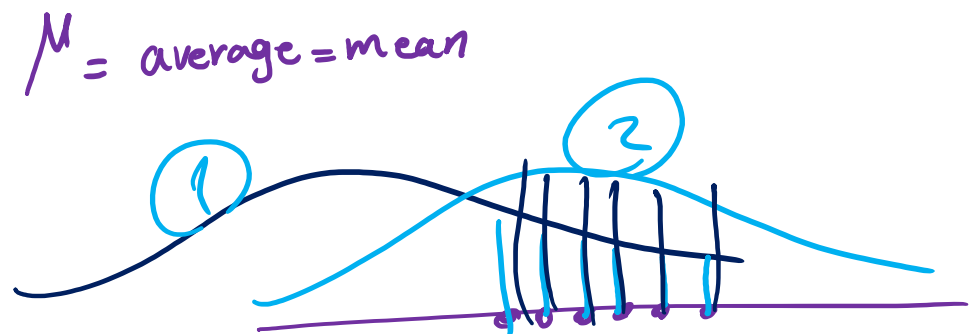


- Exponential Density Function:

Parameter

$$f_x(x) = \frac{1}{\mu} e^{-\frac{x}{\mu}} \quad \text{for } x \geq 0$$

$$F_x(x) = 1 - e^{-\frac{x}{\mu}} \quad \text{for } x \geq 0$$



- Gaussian(Normal) Density Function

$$f_x(x) = \frac{1}{\sqrt{2\pi}b} e^{-\frac{(x-a)^2}{2b^2}}$$

μ → mean
 σ → Standard deviation
 σ^2 → Variance

Discrete Probability Functions

- Examples:

- Bernoulli Distribution:

- $$\begin{cases} 1 - p & \text{for } x = 0 \\ p & \text{for } x = 1 \end{cases}$$

In Bernoulli, just a **single** trial is conducted

- Binomial Distribution:


- $$P(X = k) = \binom{n}{k} p^k (1 - p)^{n-k}$$

k is number of successes

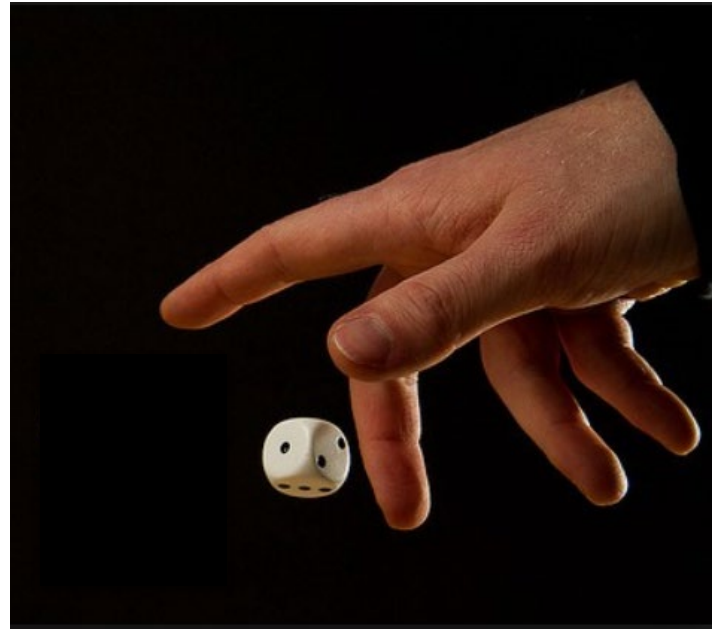
n-k is number of failures

$\binom{n}{k}$ The total number of ways of selection **k** distinct combinations of **n** trials, **irrespective of order**.

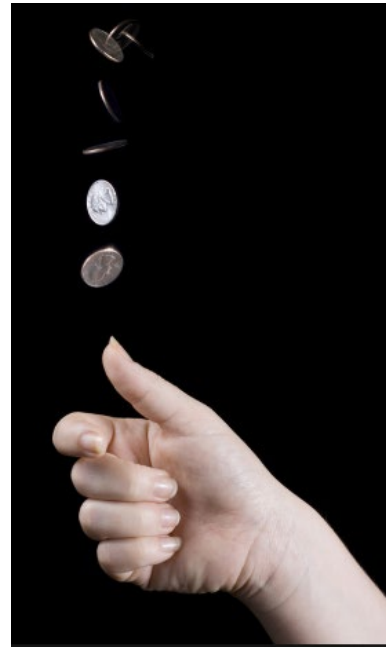
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Example



X = Throw a
dice



Y = Flip a coin

\mathbf{X} and \mathbf{Y} are random variables

\mathbf{N} = total number of trials

n_{ij} = Number of occurrence

		\mathbf{X}						C_j
		$x_{i=1} = 1$	$x_{i=2} = 2$	$x_{i=3} = 3$	$x_{i=4} = 4$	$x_{i=5} = 5$	$x_{i=6} = 6$	
\mathbf{Y}	$y_{j=2} = tail$	$n_{ij} = 3$	$n_{ij} = 4$	$n_{ij} = 2$	$n_{ij} = 5$	$n_{ij} = 1$	$n_{ij} = 5$	20
	$y_{j=1} = head$	$n_{ij} = 2$	$n_{ij} = 2$	$n_{ij} = 4$	$n_{ij} = 2$	$n_{ij} = 4$	$n_{ij} = 1$	15
	C_i	5	6	6	7	5	6	N=35

X**C_j**
 $x_{i=1} = 1 \quad x_{i=2} = 2 \quad x_{i=3} = 3 \quad x_{i=4} = 4 \quad x_{i=5} = 5 \quad x_{i=6} = 6$
Y $y_{j=2} = \text{tail}$
 $y_{j=1} = \text{head}$
C_i

$n_{ij} = 3$	$n_{ij} = 4$	$n_{ij} = 2$	$n_{ij} = 5$	$n_{ij} = 1$	$n_{ij} = 5$	20
$n_{ij} = 2$	$n_{ij} = 2$	$n_{ij} = 4$	$n_{ij} = 2$	$n_{ij} = 4$	$n_{ij} = 1$	15
5	6	6	7	5	6	N=35

$$P(X=2, Y=\text{tail}) = \frac{4}{35} = \frac{n_{ij}}{N}$$

$$P(Y=\text{head}) = \frac{15}{35} = \frac{C_j}{N} \quad P(X=3) = \frac{6}{35} = \frac{C_i}{N} = \sum_Y P(X=3, Y=y) \rightarrow \text{Sum Rule}$$

$$P(Y=\text{tail} | X=4) = \frac{5}{7} = \frac{n_{ij}}{C_i}$$

$$P(X=4 | Y=\text{tail}) = \frac{5}{20} = \frac{n_{ij}}{C_j}$$

$$P(X, Y) = \frac{n_{ij}}{N} = \frac{n_{ij}}{C_i} \frac{C_i}{N} = \frac{n_{ij}}{C_j} \frac{C_j}{N}$$

$$P(X, Y) = P(Y|X) P(X) = P(X|Y) P(Y)$$

↳ product rule

Probability:

$$p(X = x_i) = \frac{c_i}{N}$$

Joint probability:

$$p(X = x_i, Y = y_j) = \frac{n_{ij}}{N}$$

Conditional probability:

$$p(Y = y_j | X = x_i) = \frac{n_{ij}}{c_i}$$

Sum rule

$$p(X = x_i) = \sum_{j=1}^L p(X = x_i, Y = y_j) \Rightarrow p(X) = \sum_Y P(X, Y)$$

Product rule

$$p(X = x_i, Y = y_j) = \frac{n_{ij}}{N} = \frac{n_{ij}}{c_i} \frac{c_i}{N} = p(Y = y_j | X = x_i) p(X = x_i)$$

$$p(X, Y) = p(Y|X)p(X)$$

$$P(a,b) = P(a|b)P(b) \quad P(a,b,c) = P(a|b,c)P(b,c)$$

Conditional Independence

- Examples:**

$$P(H,F,V,D) = P(H|F,V,D) \times P(F,V,D)$$

$$= P(H|F,D) \times P(F|V,D) \times P(V,D)$$

$P(\text{Virus} \mid \text{Drink Beer}) = P(\text{Virus})$
iff **Virus** is independent of **Drink Beer**

$$= P(H|F,D) \times P(F|V) \times P(V|D) \times P(D)$$

$P(\text{Flu} \mid \text{Virus}, \text{Drink Beer}) = P(\text{Flu} \mid \text{Virus})$


$$= P(H|F,D) \times P(F|V) \times P(V) \times P(D)$$
iff **Flu** is independent of **Drink Beer**, given **Virus**

$P(\text{Headache} \mid \text{Flu}, \text{Virus}, \text{Drink Beer}) =$
 $P(\text{Headache} \mid \text{Flu}, \text{Drink Beer})$
iff **Headache** is independent of **Virus**, given **Flu** and **Drink Beer**

Assume the above independence, we obtain:

$$\begin{aligned}
 &P(\text{Headache}; \text{Flu}; \text{Virus}; \text{Drink Beer}) \\
 &= P(\text{Headache} \mid \text{Flu}; \text{Virus}; \text{Drink Beer}) P(\text{Flu} \mid \text{Virus}; \text{Drink Beer}) \\
 &\quad P(\text{Virus} \mid \text{Drink Beer}) P(\text{Drink Beer}) \\
 &= P(\text{Headache} \mid \text{Flu}; \text{Drink Beer}) P(\text{Flu} \mid \text{Virus}) P(\text{Virus}) P(\text{Drink Beer})
 \end{aligned}$$

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Bayes' Rule

- $P(X|Y)$ = Fraction of the worlds in which X is true given that Y is also true.

$$P(X, Y) = P(X|Y)P(Y) \Rightarrow P(X|Y) = \frac{P(X, Y)}{P(Y)} = \frac{P(Y|X)P(X)}{P(Y)}$$

- For example:

- H = "Having a headache"
- F = "Coming down with flu"
- $P(\text{Headache}|\text{Flu})$ = fraction of flu-inflicted worlds in which you have a headache. How to calculate?

$$P(Y) = \sum_x P(Y, X=x) = \sum_x P(Y|X)P(X)$$

- Definition:

$$P(X|Y) = \frac{P(X, Y)}{P(Y)} = \frac{P(Y|X)P(X)}{P(Y)}$$

Corollary:

$$P(X, Y) = P(Y|X)P(X)$$

This is called **Bayes Rule**

Bayes' Rule

- $$P(\text{Headache}|\text{Flu}) = \frac{P(\text{Headache}, \text{Flu})}{P(\text{Flu})}$$

$$= \frac{P(\text{Flu}|\text{Headache})P(\text{Headache})}{P(\text{Flu})}$$

Other cases:

- $$P(Y|X) = \frac{P(X|Y)P(Y)}{P(X|Y)P(Y) + P(X|\neg Y)P(\neg Y)}$$


- $$P(Y = y_i|X) = \frac{P(X|Y)P(Y)}{\sum_{i \in S} P(X|Y = y_i)P(Y = y_i)}$$

- $$P(Y|X, Z) = \frac{P(X|Y, Z)P(Y, Z)}{P(X, Z)}$$

$$= \frac{P(X|Y, Z)P(Y, Z)}{P(X|Y, Z)P(Y, Z) + P(X|\neg Y, Z)P(\neg Y, Z)}$$

$P(X, Y, Z) = P(X|Y, Z)P(Y, Z)$

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Mean and Variance

$$E[g(x)] = \sum g(x) p(x)$$

- Expectation: The mean value, center of mass, first moment:

$$\underline{E_X[g(X)]} = \int_{-\infty}^{\infty} g(x) p_X(x) dx = \underline{\mu}$$

- N-th moment: $g(x) = x^n$
- N-th central moment: $g(x) = (x - \mu)^n$
- Mean: $E_X[X] = \int_{-\infty}^{\infty} x p_X(x) dx$

- $E[\alpha X] = \alpha E[X]$

- $E[\alpha + X] = \alpha + E[X]$

- Variance(Second central moment): $Var(x) =$

$$Var(x) = \underline{E_X[(X - E_X[X])^2]} = \underline{E_X[X^2] - E_X[X]^2}$$

$$E[X^2] = Var(x) + E[X]^2$$

- $Var(\alpha X) = \alpha^2 Var(X)$

- $Var(\alpha + X) = Var(X)$

$$g(x) = x \quad x = [1, 2, 3]$$

$$P(x=1) = \frac{1}{5} \quad P(x=2) = \frac{2}{5} \quad P(x=3) = \frac{2}{5}$$

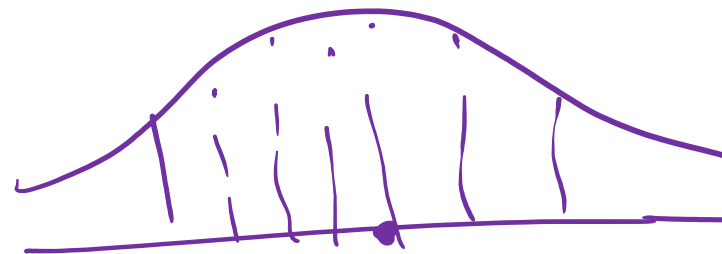
$$E[g(x)] = \sum_{i=1}^n g(x_i) P(x_i)$$

$$E[g(x)] = 1 \times \frac{1}{5} + 2 \times \frac{2}{5} + 3 \times \frac{2}{5} = \frac{11}{5}$$

$$\mu = \text{mean} = \frac{1+2+3}{3} = 2$$

$$x = [1, 2, 2, 3, 3]$$

$$\mu = \frac{1+2+2+3+3}{5} = \frac{11}{5} = E[g(x)]$$



$$X = \begin{bmatrix} 1 \\ 2 \\ 3 \end{bmatrix} \quad \text{height} = h$$

nxd

$$d=1 \quad n=3$$

$$\mu_h = \frac{1+2+3}{3} = 2$$

$$\bar{X} = \begin{bmatrix} \bar{h} \\ 1-\mu_h \\ 2-\mu_h \\ 3-\mu_h \end{bmatrix} = \begin{bmatrix} -1 \\ 0 \\ 1 \end{bmatrix}$$

$$\sigma_h^2 = \frac{1}{N} \sum_{i=1}^N (x_i - \mu_h)^2 = \left(\frac{1}{N} \sum_{i=1}^N \right) (x_i - E[h])^2$$

$$= E[(x_i - E[h])^2]$$

$$\sigma_h^2 = \frac{\bar{X}^T \bar{X}}{N} = \frac{1}{N} \begin{bmatrix} 1-\mu_h & 2-\mu_h & 3-\mu_h \end{bmatrix} \begin{bmatrix} 1-\mu_h \\ 2-\mu_h \\ 3-\mu_h \end{bmatrix}$$

$$= \frac{1}{N} \left[(1-\mu_h)^2 + (2-\mu_h)^2 + (3-\mu_h)^2 \right] = \frac{\sum_{i=1}^N (x_i - \mu_h)^2}{N}$$

$$X = \begin{matrix} & \begin{matrix} h & \text{weight} = w \end{matrix} \\ \begin{bmatrix} 1 \\ 2 \\ 3 \end{bmatrix} & \begin{bmatrix} 4 \\ 5 \\ 6 \end{bmatrix} \end{matrix}$$

3×2
 $n \times d$

$$\mu_h = 2 \quad \mu_w = 5$$

$$\bar{X} = \begin{bmatrix} \bar{h} & \bar{w} \\ 1 - \mu_h & 4 - \mu_w \\ 2 - \mu_h & 5 - \mu_w \\ 3 - \mu_h & 6 - \mu_w \end{bmatrix}$$

$$\text{Cov} = \frac{\bar{X}^T \bar{X}}{N} = \frac{1}{N} \begin{bmatrix} 1 - \mu_h & 2 - \mu_h & 3 - \mu_h \\ 4 - \mu_w & 5 - \mu_w & 6 - \mu_w \end{bmatrix} \begin{bmatrix} 1 - \mu_h & 4 - \mu_w \\ 2 - \mu_h & 5 - \mu_w \\ 3 - \mu_h & 6 - \mu_w \end{bmatrix}$$

$$\text{Cov} = \begin{bmatrix} \sigma_h^2 = \sigma_{hh} & \sigma_{hw} \\ \sigma_{wh} & \sigma_w^2 = \sigma_{ww} \end{bmatrix}$$

2×2
 $d \times d$

$$\begin{matrix} X - \mu_h & \perp & X - \mu_w \\ \begin{bmatrix} 1 - \mu_h \\ 2 - \mu_h \\ 3 - \mu_h \end{bmatrix} & & \begin{bmatrix} 4 - \mu_w \\ 5 - \mu_w \\ 6 - \mu_w \end{bmatrix} \end{matrix}$$

$$\bar{X} = \begin{matrix} & \begin{matrix} \bar{h} & \bar{w} \end{matrix} \\ \begin{bmatrix} 1 - \mu_h & 4 - \mu_w \\ 2 - \mu_h & 5 - \mu_w \\ 3 - \mu_h & 6 - \mu_w \end{bmatrix} \end{matrix}$$

Standardization

$$\bar{X}^* = \begin{bmatrix} \bar{h}^* & \bar{w}^* \\ \frac{1 - \mu_h}{\sigma_h} & \frac{4 - \mu_w}{\sigma_w} \\ \frac{2 - \mu_h}{\sigma_h} & \frac{5 - \mu_w}{\sigma_w} \\ \frac{3 - \mu_h}{\sigma_h} & \frac{6 - \mu_w}{\sigma_w} \end{bmatrix}$$

$$\begin{aligned} \text{Correlation} &= \frac{\bar{X}^{*T} \bar{X}^*}{n} \\ \text{Cor} &= \frac{1}{N} \begin{bmatrix} \frac{1 - \mu_h}{\sigma_h} & \frac{2 - \mu_h}{\sigma_h} & \frac{3 - \mu_h}{\sigma_h} \\ . & . & . \end{bmatrix} \begin{bmatrix} \frac{1 - \mu_h}{\sigma_h} & . \\ \frac{2 - \mu_h}{\sigma_h} & . \\ \frac{3 - \mu_h}{\sigma_h} & . \end{bmatrix} \end{aligned}$$

$$\text{Cor} = \begin{bmatrix} \frac{\sigma_h^2}{\sigma_h^2} = 1 & -1 \leq \sigma_{hw} \leq 1 \\ -1 \leq \sigma_{wh} \leq 1 & \frac{\sigma_w^2}{\sigma_w^2} = 1 \end{bmatrix}$$

For Joint Distributions

- Expectation and Covariance:

- $E[X + Y] = E[X] + E[Y]$

- $cov(X, Y) = E[(X - E_X[X])(Y - E_Y[Y])] = E[XY] - E[X]E[Y]$

- $Var(X + Y) = Var(X) + 2cov(X, Y) + Var(Y)$

$$X = Z$$

$$\mu = 0 \quad \sigma = 1$$

$$Y = Z^2$$


$$\mu = .8 \quad \sigma = .8$$

$$E[X^2] = Var(X) + E[X]^2$$

$$\mu_Y = E[Z^2] = Var(Z) + \underbrace{E[Z]^2}_{0} = 1 + 0 = 1$$

$$Cov(X, Y) = E[XY] - \underbrace{E[X]}_0 E[Y] = \underbrace{E[Z^3]}_0 - 0 \times 1 = 0$$

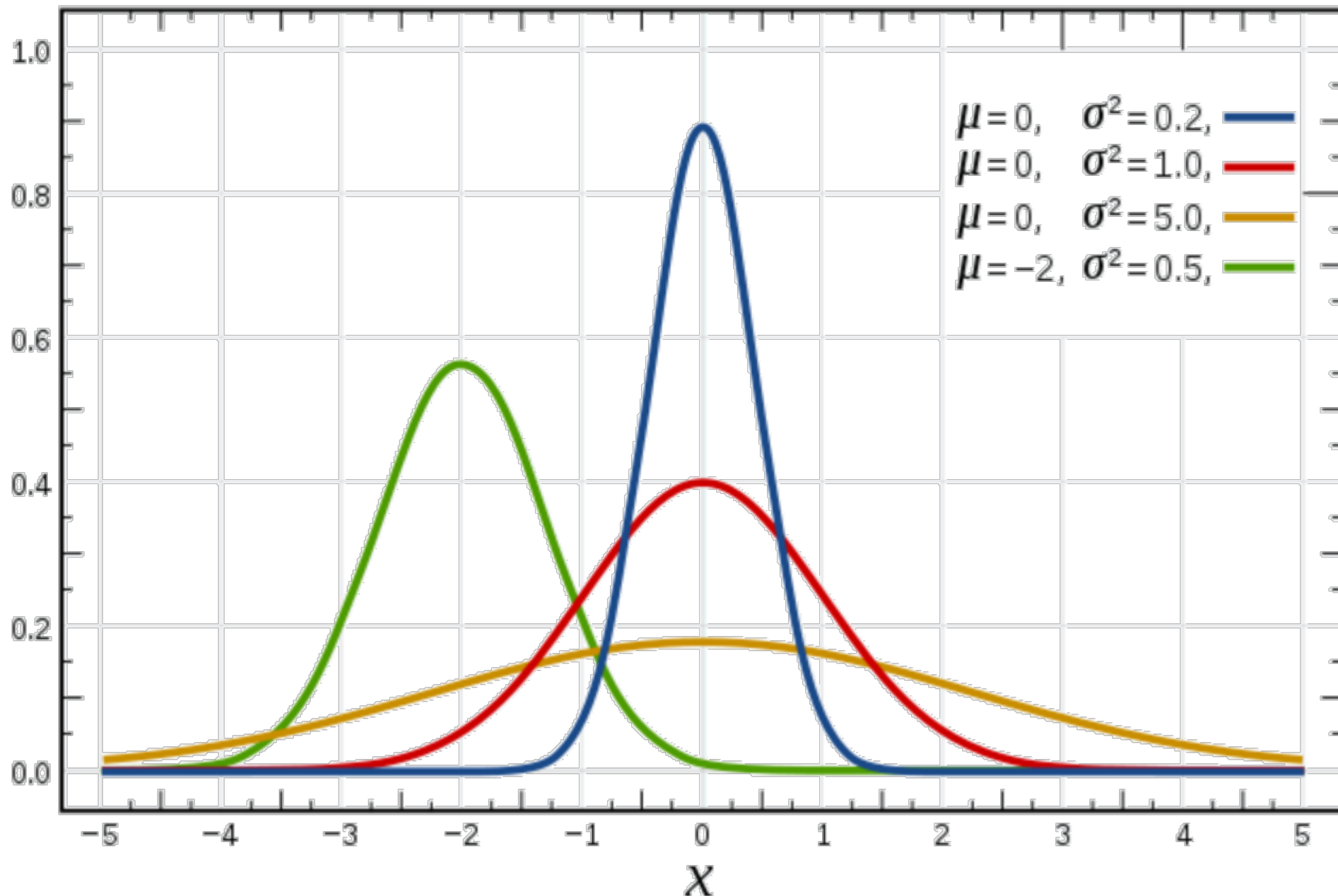
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Gaussian Distribution

- Gaussian Distribution:
$$f(x|\mu, \sigma^2) = \frac{1}{\sqrt{2\pi\sigma^2}} e^{-\frac{(x-\mu)^2}{2\sigma^2}}$$

Probability density function



Probability versus likelihood

Probability is about fact $P(X=H) = 0.5$

$$H, H, H, T \quad P(X=H) = \frac{3}{4}$$

$$P(X|\mu, \sigma) = f(x|\mu, \sigma) = \frac{1}{\sqrt{2\pi\sigma^2}} e^{-\frac{(x-\mu)^2}{2\sigma^2}}$$

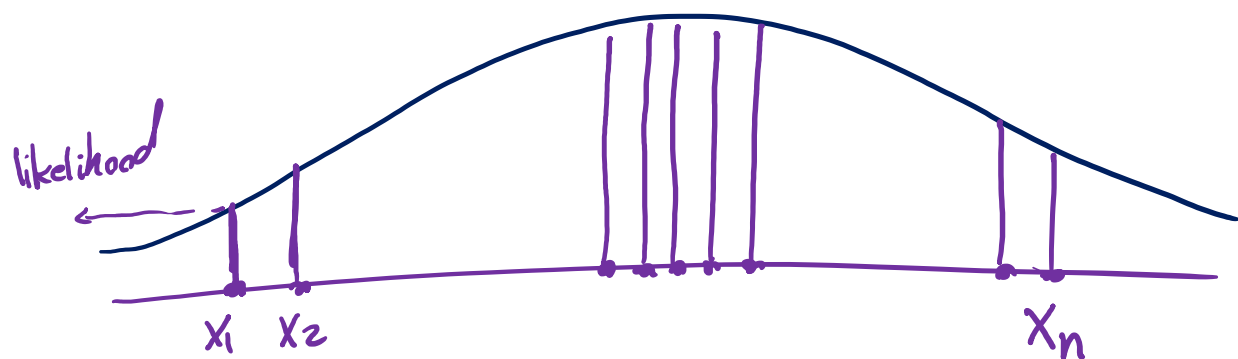
$$\mu = \frac{\sum_{i=1}^N x_i}{N}$$

$$\sigma^2 = \frac{\sum (x_i - \mu)^2}{N}$$

$$P(X|a, b) = f(x|a, b) = \frac{1}{\sqrt{2\pi b^2}} e^{-\frac{(x-a)^2}{2b^2}}$$

$$L(a, b | x) ?$$

①



$$f(x_1, x_2, \dots, x_n | \overset{\Theta}{a, b})$$

$$\text{iid } f(x_1 | a, b) f(x_2 | a, b) \dots f(x_n | a, b)$$

$$\log(f(x_1 | a, b) \dots f(x_n | a, b))$$

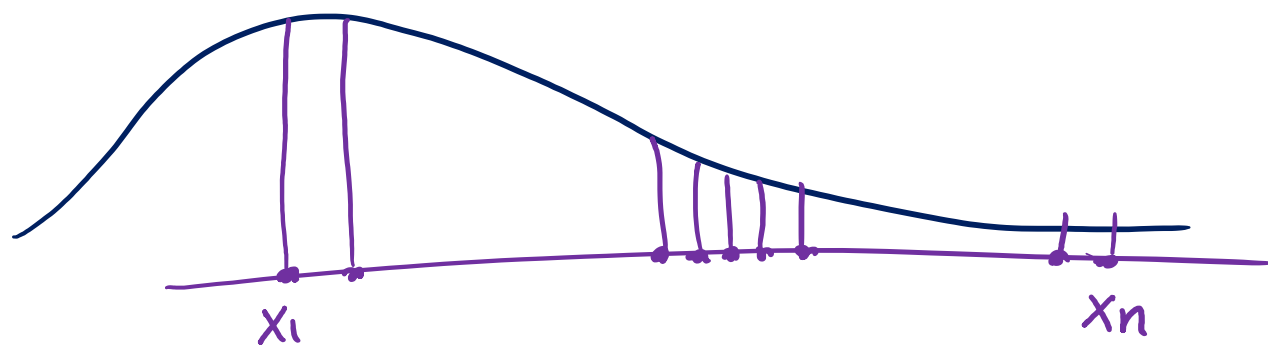
$$\log f(x_1 | a, b) + \dots + \log f(x_n | a, b)$$

$$f(x_1 | \Theta)$$

$$\hookrightarrow \Theta \in \{a, b\}$$

$$L(a, b | x) ?$$


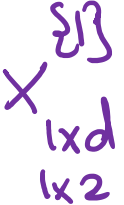
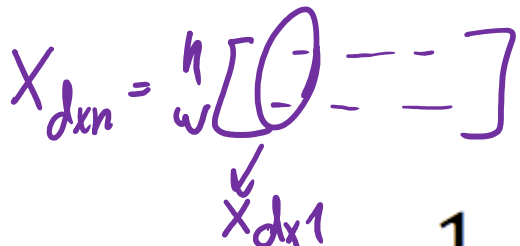
②



$$f(x_1, x_2, \dots, x_n | a, b)$$

$$f(x_1 | a, b) f(x_2 | a, b) \dots f(x_n | a, b)$$

Multivariate Gaussian Distribution

$X_{n \times d} =$ 
 $X_{1 \times d} =$ 
 $X_{d \times n} =$ 

$$p(x|\mu, \Sigma) = \frac{1}{(2\pi)^{n/2} |\Sigma|^{1/2}} \exp\left\{-\frac{1}{2} (x - \mu)^T \Sigma^{-1} (x - \mu)\right\}$$

$p(x|\mu, \Sigma)$ is a scalar (1x1).
 $(x - \mu)^T$ is a 1x1 scalar.
 Σ^{-1} is a dxd matrix.
 $(x - \mu)$ is a dxd matrix.
 $\mu = [\mu_h \ \mu_w]_{1 \times d}$ is a 1xd vector.
 x is a dxd matrix.
 Σ is a dxd matrix.

- Moment Parameterization $\mu = E(X)$

$$\Sigma = Cov(X) = E[(X - \mu)(X - \mu)^T]$$

- Mahalanobis Distance $\Delta^2 = (x - \mu)^T \Sigma^{-1} (x - \mu)$
- Tons of applications (MoG, FA, PPCA, Kalman filter,...)

Properties of Gaussian Distribution

- The **linear transform** of a Gaussian r.v. is a Gaussian. Remember that no matter how x is distributed

$$E(AX + b) = AE(X) + b$$

$$\text{Cov}(AX + b) = A\text{Cov}(X)A^T$$

this means that for Gaussian distributed quantities:

$$X \sim N(\mu, \Sigma) \rightarrow AX + b \sim N(A\mu + b, A\Sigma A^T)$$

- The **sum** of two independent Gaussian r.v. is a Gaussian

$$Y = X_1 + X_2, \quad X_1 \perp X_2 \rightarrow \mu_y = \mu_1 + \mu_2, \Sigma_y = \Sigma_1 + \Sigma_2$$

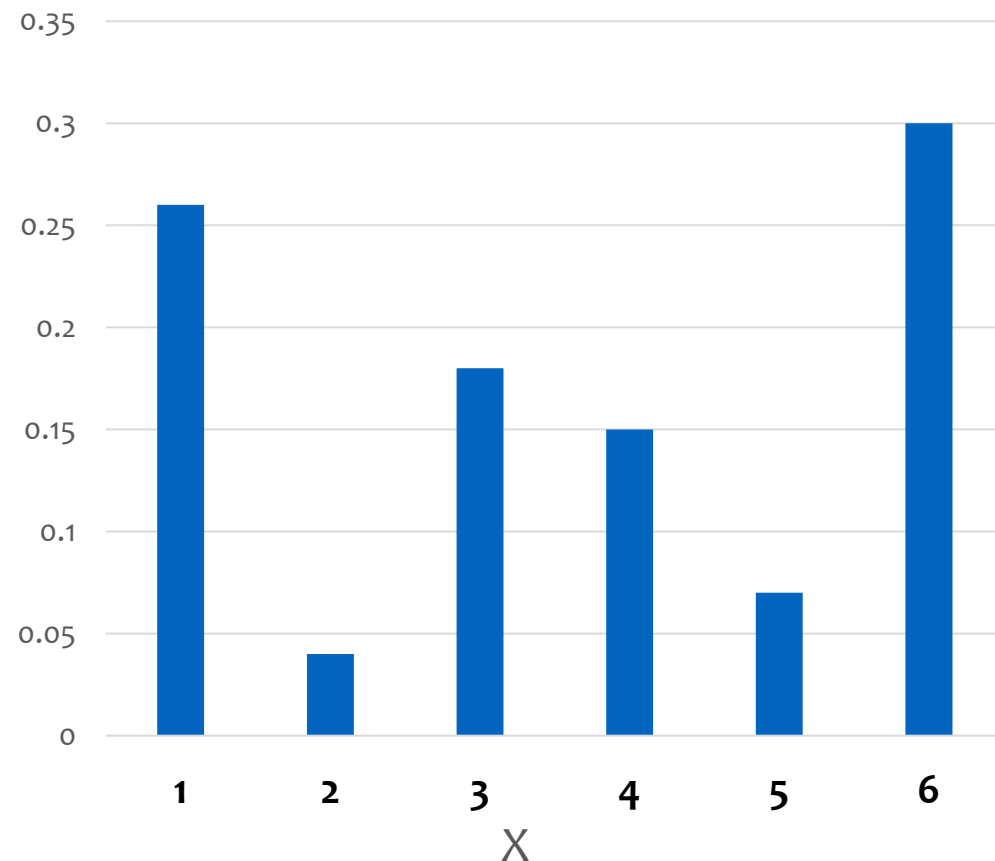
- The **multiplication** of two Gaussian functions is another Gaussian function (although no longer normalized)

$$N(a, A)N(b, B) \propto N(c, C),$$

$$\text{where } C = (A^{-1} + B^{-1})^{-1}, c = CA^{-1}a + CB^{-1}b$$

Central Limit Theorem

Probability mass function of a **biased** dice



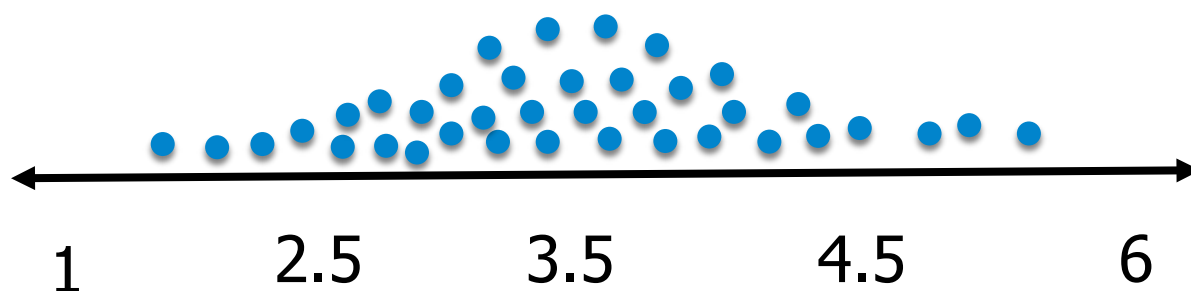
Let's say, I am going to get a sample from this pmf having a size of **$n = 4$**

$$S_1 = \{1,1,1,6\} \Rightarrow E(S_1) = 2.25$$

$$S_2 = \{1,1,3,6\} \Rightarrow E(S_2) = 2.75$$


\vdots

$$S_m = \{1,4,6,6\} \Rightarrow E(S_m) = 4.25$$



According to CLT, it will follow a bell curve distribution (normal distribution)

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Maximum Likelihood Estimation

- Probability: inferring probabilistic quantities for data given fixed models (e.g. prob. of events, marginals, conditionals, etc).
- Statistics: inferring a model given fixed data observations (e.g. clustering, classification, regression).

Main assumption:

Independent and identically distributed random variables
i.i.d

Maximum Likelihood Estimation

For Bernoulli (i.e. flip a coin):

Objective function: $P(x_i|\theta) = \theta^{x_i}(1-\theta)^{1-x_i}$ $x_i \in \{0,1\}$ or $\{head, tail\}$

$$L(\theta|X) = L(\theta|X = x_1, X = x_2, X = x_3, \dots, X = x_n)$$

i.i.d assumption

$$P(x_1, x_2, \dots, x_n | \theta) \\ P(x_1 | \theta) \dots P(x_n | \theta)$$

$$L(\theta|X) = \prod_{i=1}^n P(x_i|\theta)$$

$$L(\theta|X) = \prod_{i=1}^n P(x_i|\theta) = \prod_{i=1}^n \theta^{x_i} (1-\theta)^{1-x_i}$$

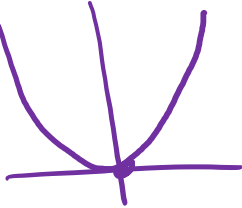
$$\begin{aligned} L(\theta|X) &= \theta^{x_1} (1-\theta)^{1-x_1} \times \theta^{x_2} (1-\theta)^{1-x_2} \dots \times \theta^{x_n} (1-\theta)^{1-x_n} = \\ &= \theta^{\sum x_i} (1-\theta)^{\sum (1-x_i)} \end{aligned}$$

We don't like multiplication, let's convert it into summation

What's the trick?

Take the log

$$L(\theta|X) = \theta^{\sum x_i} (1 - \theta)^{\sum (1 - x_i)}$$

$\log a^b$
 $y = x^2$ 

$$\log L(\theta|X) = l(\theta|X) = \log(\theta) \sum_{i=1}^n x_i + \log(1 - \theta) \sum_{i=1}^n (1 - x_i)$$

How to optimize θ ?

$$\frac{\partial l(\theta|X)}{\partial \theta} = 0 \quad \frac{\sum_{i=1}^n x_i}{\theta} - \frac{\sum_{i=1}^n (1 - x_i)}{1 - \theta} = 0$$

$x_i \in \{1, 0\}$

$$\theta = \frac{1}{n} \sum_{i=1}^n x_i$$

$\frac{\overbrace{1+1+1+\dots}^{8 \times 1} + \overbrace{0+0}^{2 \times 0}}{10} = 0.8$