# Logistic Regression and the ROC curve

Schwartz

July 22, 2017

#### Odd, even at best

Last year Leicester City was given 5000 to 1 odds to win the English Premier League. Actually, these are the longest odds *ever seen* for *any* top tier sporting league... *ever*. To put this in perspective, the current odds out of Vegas for "the most unlikely team to win the 2016/2017 NFL season" – woefully disastrous Cleveland\* Browns – are 200 to 1.

Since the clubs inception in 1890, Leicester City has only managed to appear in the Premier league 10 seasons. They had only been promoted the previous season and just barely escaped relegation in their final match that season. Only five teams – Arsenal, Chelsea, Liverpool, Man. City, and Man. U. – have held the trophy for the past 21 seasons.

Only a few stout souls put money down on Leicester City last year. And when Leicester City (literally against all odds) won the premiership last season in absolutely stunning, unbelievable, and unprecedented fashion, those stout souls got paid. Everyone, that is, except for John Micklethwait. John M has made the same bet – 20 pounds (\$29) that Leicester will win their division – every August for the past 20 years. Every year, that is, except this one. Last year he moved from London to New York and missed placing his bet. That's a pity for John M because if he had made his bet he would have won 100,000 pounds, or \$145,355.

Overall, \$3,000 was bet on Leicester City last season. The <u>unprecedented</u> \$15,000,000 payout nearly bankrupted the bookmakers. John M got \$0.

<sup>\*</sup>Cleveland's 52-year championship drought ended with the 2015/16 NBA season

#### Odds

$$\mbox{Odds} = \frac{p}{1-p} \Longrightarrow \mbox{p} = \frac{Odds}{1+Odds} = \frac{1}{1+Odds^{-1}}$$
 
$$1-p = \frac{1}{1+Odds}$$

# **Objectives**

# Morning

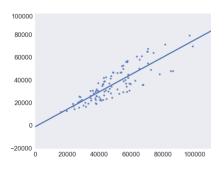
- Know why logistic regression is a thing:
  - Classification vs. Regression
  - Link functions
- Interpreting Logistic Regression
  - Fitted Values (probabilities)
  - Coefficients (log odds ratios)

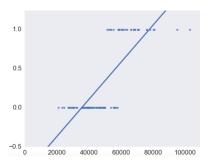
# <u>Afternoon</u>

- ightharpoonup T+, T-, F+, F- and other terminology
  - Confusion Matricies
- Thresholding Classification rules
  - ROC curves

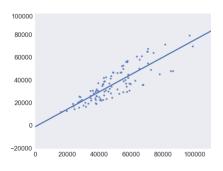


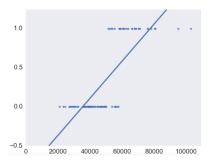
# Linear Regression





## Linear Regression

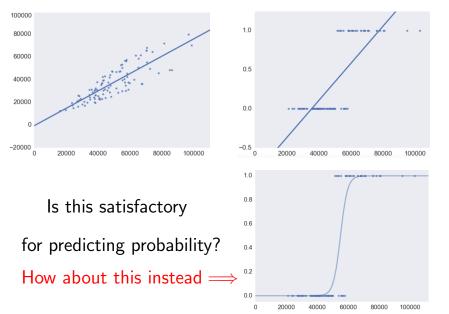




Is this satisfactory

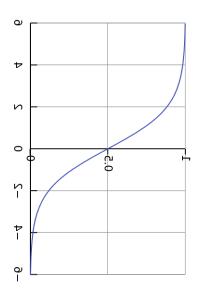
for predicting probability?

# Linear Regression



► The "logit"

$$g(p) = \log\left(\frac{p}{1-p}\right)$$

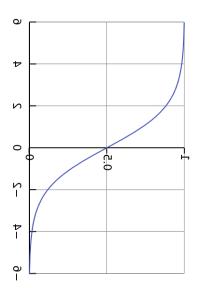


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$$p \in [0,1] \mapsto Z \in \mathbb{R}$$



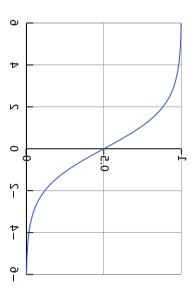
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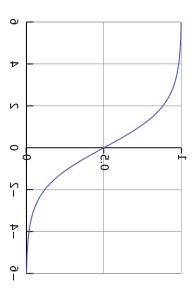
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Don't be at odds with odds!



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because how else can Z stay between 0 and 1??



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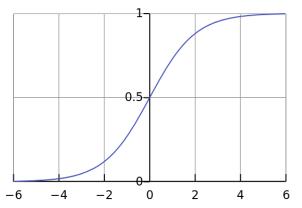
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So 
$$g(p) = Z = \log\left(\frac{p}{1-p}\right) \in \mathbb{R}$$
 (which is called the logit function) and  $Z = \beta_0 + \beta_1 x_1 + \dots + \beta_m x_m \in \mathbb{R}$  models the log odds

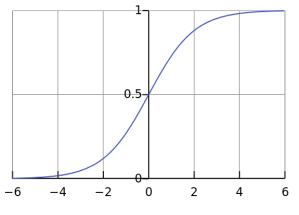
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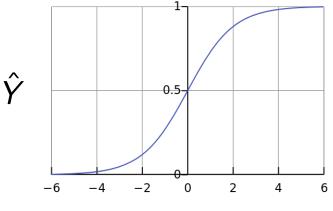
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- I.e., the odds are linear in x on a multiplicative, i.e., odds increase with x on a logorithmic scale with base exp(β<sub>i</sub>)
- ► The *log* odds  $log\left(\frac{Pr(Y=1|x)}{Pr(Y=0|x)}\right)$  are on a linear scale





# The Odds Ratio (OR)

▶ Equivalently,  $\exp(\beta_j)$  is the *odds ratio* (*OR*) between 1-unit differences in  $x_j$  (e.g., 0 versus 1) when other x's are constant

$$\exp(\beta_j) = \frac{Pr(Y = 1|x_j + 1, x_{-j})/Pr(Y = 0|x_j + 1, x_{-j})}{Pr(Y = 1|x)/Pr(Y = 0|x)}$$

since  $Pr(Y = 1|x_i + 1, x_{-i})$  $Pr(Y = 0|x_i + 1, x_{-i})$  $= \exp(\beta_0) \exp(\beta_1 x_1) \cdots \exp(\beta_i (x_i + 1)) \cdots \exp(\beta_m x_m)$  $= \exp(\beta_0) \exp(\beta_1 x_1) \cdots \exp(\beta_i x_i) \exp(\beta_i) \cdots \exp(\beta_m x_m)$ and Pr(Y=1|x)Pr(Y=0|x) $= \exp(\beta_0) \exp(\beta_1 x_1) \cdots \exp(\beta_i x_i) \cdots \exp(\beta_m x_m)$ 

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$$= \exp(\beta_0) \exp(\beta_1 x_1) \cdots \exp(\beta_j (x_j+1)) \cdots \exp(\beta_m x_m)$$

$$= \exp(\beta_0) \exp(\beta_1 x_1) \cdots \exp(\beta_j x_j) \exp(\beta_j) \cdots \exp(\beta_m x_m)$$
and 
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$$= \exp(\beta_0) \exp(\beta_1 x_1) \cdots \exp(\beta_i x_j) \cdots \exp(\beta_m x_m)$$

▶ So  $\beta_j$  is the change in log(OR) for one unit changes in  $x_j$ ...



# Logistic Regression Likelihood and Deviance

Likelihood

$$f(\mathbf{Y}|\boldsymbol{\beta}, \mathbf{x}) = \prod \left(\frac{1}{1 + e^{-\mathbf{x}_i^\mathsf{T}\boldsymbol{\beta}}}\right)^{Y_i} \left(\frac{1}{1 + e^{\mathbf{x}_i^\mathsf{T}\boldsymbol{\beta}}}\right)^{1 - Y_i}$$

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Deviance

$$D_{M} = -2 \left( \log f(\mathbf{Y}|\hat{\boldsymbol{\beta}}, \mathbf{x}) - \log f(\mathbf{Y}|\mathbf{Y}) \right)$$

$$\overset{\text{approx.}}{\sim} \chi_{n-p-1}^{2}$$

n =sample size

p = number of coefficients in model M

 $f(\mathbf{Y}|\mathbf{Y}) = \text{saturated model } (\mathbf{Y} \text{ perfectly predicted})$ 

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▶ In logistic regression

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▶ In linear regression

$$\begin{split} D_M &= \frac{RSS}{\sigma^2} & \text{[show this]} \\ &= \frac{\sum (Y_i - \hat{Y})^2}{\sigma^2} = (n - p - 1) \frac{s^2}{\sigma^2} \overset{\text{approx.}}{\sim} \chi^2_{n-p-1} \end{split}$$

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[what are residuals?] [what are residuals in logistic regression?]



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$$\hat{\boldsymbol{\beta}} = \operatorname*{argmax}_{\boldsymbol{\beta}} \prod \left( \frac{1}{1 + \mathrm{e}^{-\mathbf{x}^T \boldsymbol{\beta}}} \right)^{\mathbf{Y}_i} \left( \frac{1}{1 + \mathrm{e}^{\mathbf{x}^T \boldsymbol{\beta}}} \right)^{1 - \mathbf{Y}_i} \\ \iff \hat{\boldsymbol{\beta}} = \operatorname*{argmin}_{\boldsymbol{\beta}} D_{\boldsymbol{\beta}} = \operatorname*{argmax}_{\boldsymbol{\beta}} \left( \log f(\mathbf{Y}|\hat{\boldsymbol{\beta}}, \mathbf{x}) - \log f(\mathbf{Y}|\mathbf{Y}) \right)$$

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$$-\prod \left(\frac{1}{1+e^{-\mathbf{x}_{i}^{T}\boldsymbol{\beta}}}\right)^{Y_{i}} \left(\frac{1}{1+e^{\mathbf{x}_{i}^{T}\boldsymbol{\beta}}}\right)^{1-Y_{i}} + \lambda ||\boldsymbol{\beta}||^{2}?$$

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► McFadden's pseudo  $R^2 = 1 - D_M/D_0$ 

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- http://www.ats.ucla.edu/stat/mult\_pkg/faq/general/Psuedo\_RSquareds.htm

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where model R is nested in model F with k fewer parameters

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▶ For *non-nested* models, compare

$$AIC: -2\log f(Y|\hat{\theta}) + 2k$$

$$BIC: -2\log f(Y|\hat{\theta}) + k\log(n)$$

Remember,  $D_M = -2(logf(\mathbf{Y}|\hat{\theta}_M) - logf(\mathbf{Y}|\mathbf{Y}))$ , so

comparison of (reduced and full) models can be done using

$$D_R - D_F = -2 \left( \log f(Y|\hat{ heta}^R) - \log f(Y|\hat{ heta}^F) \right) \stackrel{\text{approx.}}{\sim} \chi_k^2$$

where model R is nested in model F with k fewer parameters

For non-nested models, compare

$$AIC: -2\log f(Y|\hat{\theta}) + 2k$$

$$BIC: -2\log f(Y|\hat{\theta}) + k\log(n)$$

How else could you compare nested or non-nested models?



Predict probabilities

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- Classify outcomes (based on probabilities)

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 Balancing observational comparison groups on propensity scores Pr(T|x) which controls bias from group covariate composition differences

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		Yes	No		
Actual Class	Yes	TP	FN		
	No	FP	TN		

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Which cells are Type I and Type II error?

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Which cells are Type I and Type II error? What is the power of a test?

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- Also called True Positive Rate
- ► Also called *Recall*

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- ▶ The "1 minus" related False Positive Rate is  $\left(\frac{\#FP}{\#TN+\#FP}\right)$

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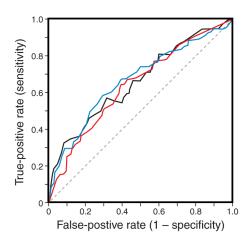
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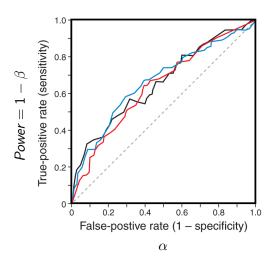
- ▶ Also called *True Positive Rate*
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- ► I.e, Type I & II error rates are 1-Specificity & 1-Sensitivity

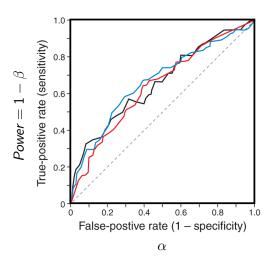
# ROC/AUC



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https://www.youtube.com/watch?v=JAQC59ArFJwhttps://www.youtube.com/watch?v=bhvvxNUbIpo

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- VERY useful concept of multiple testing contexts
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"Do we call hypotheses accurately?"

#### And just a one or two more...

			Predicted condition			
		Total population	Predicted Condition positive	Predicted Condition negative	$= \frac{\frac{\Sigma \text{ Condition positive}}{\Sigma \text{ Total population}}$	
True condition	True	condition positive	True positive	False Negative (Type II error)	True positive rate (TPR), Sensitivity, Recall = $\frac{\Sigma \text{ True positive}}{\Sigma \text{ Condition positive}}$	False negative rate (FNR), Miss rate $= \frac{\Sigma \text{ False negative}}{\Sigma \text{ Condition positive}}$
	condition	condition negative	False Positive (Type I error)	True negative	False positive rate (FPR), Fall-out $= \frac{\Sigma \text{ False positive}}{\Sigma \text{ Condition negative}}$	$\begin{aligned} & \text{True negative rate (TNR),} \\ & \text{Specificity (SPC)} \\ & = \frac{\Sigma \text{ True negative}}{\Sigma \text{ Condition negative}} \end{aligned}$
	Accuracy (ACC) =	$\begin{aligned} & \text{Positive predictive value} \\ & & \text{(PPV), Precision} \\ & = \frac{\Sigma \text{ True positive}}{\Sigma \text{ Test outcome positive}} \end{aligned}$	False omission rate (FOR) $= \frac{\Sigma \text{ False negative}}{\Sigma \text{ Test outcome negative}}$	Positive likelihood ratio $(LR+) = \frac{TPR}{FPR}$	Diagnostic odds ratio	
		$\frac{\Sigma \text{ True positive} + \Sigma \text{ True negative}}{\Sigma \text{ Total population}}$	False discovery rate (FDR) $= \frac{\Sigma \text{ False positive}}{\Sigma \text{ Test outcome positive}}$	Negative predictive value (NPV) $= \frac{\Sigma \text{ True negative}}{\Sigma \text{ Test outcome negative}}$	Negative likelihood ratio $(\text{LR}) = \frac{FNR}{TNR}$	$(DOR) = \frac{LR+}{LR-}$

Thanks, Wiki! (You're the besht!)