

Likelihood

1. Likelihood $L(\theta) = \prod_{i=1}^n f(y_i; \theta)$

Log likelihood $l(\theta; y) = \log L(\theta) = \sum_{i=1}^n \log f(y_i; \theta)$

Score $U(\theta) = l'(\theta)$, MLE $U(\hat{\theta}) = 0$

Observed information $J(\theta) = -\frac{dU}{d\theta} = -\frac{d^2l}{d\theta^2}$

Fisher's information $I(\theta) = E[J(\theta)] = -E\left[\frac{d^2l}{d\theta^2}\right]$

$E[U] = 0$, $Var(U) = E[U^2] - E[U]^2 = E[U^2] = -E[U'] = I$

$Var(\hat{\theta}) \approx \frac{1}{I(\hat{\theta})}$, $ESE(\hat{\theta}) = \sqrt{Var(\hat{\theta})}$

2. When θ is a p -dimensional parameter,

$\mathbf{U}(\theta)$ is a vector, $\mathbf{I}(\theta)$ and $\mathbf{J}(\theta)$ are matrices.

Let $p = 2$, $Var(\hat{\theta}) = I^{-1}(\theta_1, \theta_2) = \begin{pmatrix} Var(\hat{\theta}_1) & cov(\hat{\theta}_1, \hat{\theta}_2) \\ cov(\hat{\theta}_1, \hat{\theta}_2) & Var(\hat{\theta}_2) \end{pmatrix}$

3. Score test: $\frac{U(\theta)}{\sqrt{I(\theta)}} \sim N(0, 1)$ and $\frac{U^2(\theta)}{I(\theta)} \sim \chi_1^2$

Wald test: $\frac{\hat{\theta} - \theta_0}{\sqrt{I^{-1}(\theta_0)}} \sim N(0, 1)$ and $\frac{(\hat{\theta} - \theta_0)^2}{I^{-1}(\theta_0)} \sim \chi_1^2$

Likelihood Ratio test: $LR = \frac{L(\theta_0)}{L(\hat{\theta})}$, $-2 \log(LR) = -2(l(\theta_0) - l(\hat{\theta}))$, $2(l(\hat{\theta}) - l(\theta_0)) \sim \chi_1^2$

$qnorm(0.975, 0, 1) = 1.96$, $qchisq(0.95, 1) = 3.84$

4. Fisher's method of scoring $\theta_{r+1} = \theta_r + \frac{U(\theta_r)}{I(\theta_r)}$

Multiparameter case: $\theta_{r+1} = \theta_r + I^{-1}(\theta_r)U(\theta_r)$

5. Exponential family of distributions $f(y; \theta) = \exp\{a(y)b(\theta) + c(\theta) + d(y)\}$

$E\{a(Y)\} = -\frac{c'(\theta)}{b'(\theta)}$, $Var\{a(Y)\} = \frac{b''(\theta)c'(\theta) - c''(\theta)b'(\theta)}{\{b'(\theta)\}^3}$

6. Generalised Linear Models

Link function $\mu_i = E(Y_i)$, $g(\mu_i) = \mathbf{x}_i^T \beta$

Exponential family: $f(y; \theta) = \exp\{yb(\theta) + c(\theta) + d(y)\}$

Canonical link: $g(\mu_i) = b(\theta_i)$

Deviance: $D = -2 \left\{ l(\hat{\beta}) - l(\text{maximal model}, \Omega) \right\}$, where $\hat{\mu}_i = y_i$ in model Ω

Pearson residuals: $r_i = \frac{y_i - \hat{\mu}_i}{\sqrt{V(\hat{\mu}_i)}}$, where $V(\mu_i) = Var(Y_i)$