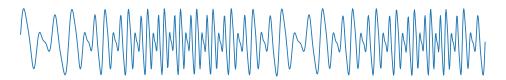
## Topological techniques for inference on periodic functions with phase variation

# Wojciech Reise Under the supervision of Frédéric Chazal and Bertrand Michel

Dec 6, 2023, Orsay







## Data with phase variation

# Signals with phase variation

A sample  $S_1, \ldots, S_N : [0,1] \to \mathbb{X}$  has **phase variation** if

$$S_n = f(\gamma_n) + W_n, \quad \text{for each } n \in \{1, \dots, N\},$$

where  $\gamma_1, \ldots, \gamma_N : [0,1] \to [0,1]$  are increasing homeomorphisms,  $f : [0,1] \to \mathbb{X}$  is continuous and  $W_n : [0,T] \to \mathbb{R}$  is a noise process.

#### Litterature

- Curve registration: estimating  $\gamma_n \circ \gamma_{n'}^{-1}$  (Tang and Muller 2008, Zhao and Itti 2018 )
- $\triangleright$  Computing a representative of f (Su et al. 2014)
- ▶ Clustering of  $S_1, ..., S_n$  (Srivastava et al. 2011)

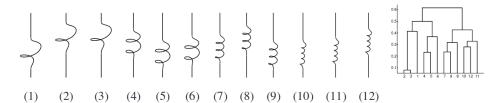
See Marron et al. 2015 for a review.

Fixed endpoints assumption

For all  $1 \le n \le N$ ,

$$\gamma_1(0) = \gamma_n(0), \tag{2}$$

$$\gamma_1(1) = \gamma_n(1). \tag{2}$$



Source: Srivastava et al. 2011.

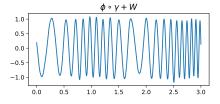
## Periodic data with phase variation

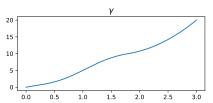
#### Definition

We call  $S:[0,1] \to \mathbb{X}$  a periodic function with phase variation if

$$S(t) = \phi(\gamma(t)) + W(t) \tag{3}$$

where  $\phi: \mathbb{R} \to \mathbb{X}$  is 1-periodic,  $\gamma: [0,1] \to [0,R]$  is an increasing homeomorphism and  $W: [0,1] \to \mathbb{X}$  is a noise process.





Example (Instantaneous phase estimation, Boashash, O'Shea, and Arnold 1990)

Decompose  $s(t) = a(t)\cos(\gamma_0(t))$  into an amplitude a(t), and a phase-variation component  $\gamma_0(t) = \arctan(H(s(t))/s(t))$ .

# Topological data analysis for periodic time series

We study S using persistent homology, a technique from Topological Data Analysis (TDA).

Contributions		
We describe the stru	octure of a topological descriptor of $\phi\circ\gamma$ .	(Chapter 3)
Let $S$ be a periodic	function with phase variation.	
1. We propose an	estimator of $\gamma$ from ${\cal S}$ ,	(Chapter 5)
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## Topological data analysis for periodic time series

We study S using persistent homology, a technique from Topological Data Analysis (TDA).

#### Contributions

We describe the structure of a topological descriptor of  $\phi \circ \gamma$ .

(Chapter 3)

Let S be a periodic function with phase variation.

1. We propose an estimator of  $\gamma$  from S.

(Chapter 5)

2. We construct a descriptor of  $\phi$  from S.

(Chapter 4)

#### TDA for time series

- Detecting periodicity in a time series (Perea 2019),
- Detecting financial crashes (Gidea and Katz 2018)
- Robust zero-crossings (Khasawneh and Munch 2018; Tanweer, Khasawneh, and Munch 2023),
- Analysis of gate signals for the study of multiple sclerosis (Bois et al. 2022).

## Outline

Additivity of persistence diagrams of periodic functions

Segmentation of periodic signals and phase estimation

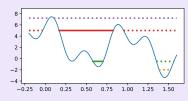
Signatures of periodic signals with phase variation

#### Intuition

The persistence diagram D(f) of a continuous function  $f:[0,T]\to\mathbb{R}$  is a multi-set of points in  $\mathbb{R}^2$ , which reflect when connected components appear and merge in  $(f^{-1}(]-\infty,t])_{t\in\mathbb{R}}$  as t increases.

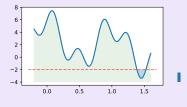


#### Definition (Chazal et al. 2016)

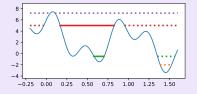


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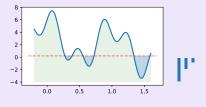
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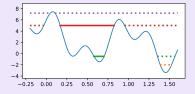


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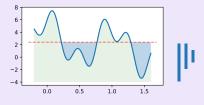
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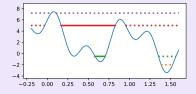


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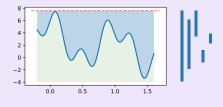
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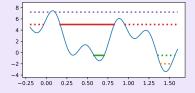


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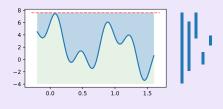
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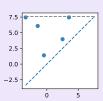




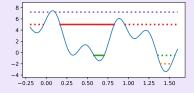
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#### Definition (Chazal et al. 2016)





## Stability: bottleneck distance

# Definition (Bottleneck distance)

We call a  $\epsilon$ -matching between two persistence diagrams D and D' a bijection  $\Gamma: A \to A'$  between some subsets of  $A \subset D$  and  $A' \subset D'$ , considered with multiplicity, if

$$d_{\infty}(a, \Gamma(a)) \le \epsilon,$$
 for any  $a \in A$ ,  
 $d_{\infty}(a, \Delta) \le \epsilon,$  for any  $a \in (D \setminus A) \cup (D' \setminus A').$ 

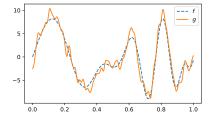
where  $\Delta = \{(x,x) \in \mathbb{R}^2\}$  denotes the diagonal.

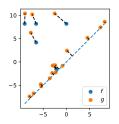
$$d_B(D,D') \coloneqq \inf\{\epsilon > 0 \mid \Gamma \text{ is an } \epsilon\text{-matching between } D' \text{ and } D'\}.$$

# Theorem (Bottleneck stability, Edelsbrunner and Harer 2010)

Let  $f,g:\mathbb{X}\to\mathbb{R}$  be two continuous functions on a compact space  $\mathbb{X}.$  Then,

$$d_B(D(f),D(g)) \leq ||f-g||_{\infty}.$$





#### Total *p*-persistence

## Definition (Total persistence)

The **persistence** of  $(b, d) \in D$  is d - b. The **total** *p*-**persistence** of a diagram *D* is

$$\operatorname{pers}_p(D) \coloneqq \left(\sum_{(b,d) \in D} (d-b)^p\right)^{1/p}.$$

Proposition (Plonka and Zheng 2016, Perez 2022)

For 
$$p = 1$$
,

$$pers_1(D(f)) + pers_1(D(-f)) = TV(f).$$

If f is  $\alpha$ -Hölder for  $p > 1 + 1/\alpha$ , then,  $\operatorname{pers}_p(D(f)) < \infty$ .

## Persistence diagrams of periodic functions

Let  $\phi : \mathbb{R} \to \mathbb{R}$  be a 1-periodic function and denote by  $\phi|_{[a,b]}$  the restriction of  $\phi$  to an interval [a,b].

Proposition (Invariance to reparametrisation)

Let  $\gamma:[0,1]\to [0,1]$  be an increasing homeomorphism. Then,  $D(\phi\circ\gamma)=D(\phi|_{[0,1]})$ .

Theorem (Additivity of persistence diagrams for periodic functions)

For  $R \in \mathbb{N}^*$ , there exists  $c \in [0,1]$  such that

$$D(\phi|_{[0,R]}) = RD(\phi|_{[c,c+1]}). \tag{4}$$

For any R > 1,

$$D(\phi|_{[0,R]}) = \lfloor R - 1 \rfloor D(\phi|_{[c,c+1]}) + D', \quad \text{with } \operatorname{pers}_{\rho}(D') \le 2\operatorname{pers}_{\rho}(D(\phi|_{[c,c+1]})). \tag{5}$$

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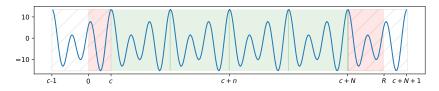
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#### Conclusion

The persistence diagram  $D(\phi \circ \gamma)$  contains information about

- $\triangleright$  extrema of  $\phi$ ,
- ▶ number of periods  $(\gamma(1) \gamma(0))$ .

# Proof of (5)



#### Proof.

Let  $c := \inf\{x \in [0,1[\mid \phi(x) = \max \phi\}, \ N = \max\{n \in \mathbb{N} \mid c+n \leq R\} \text{ and denote by } X_t := \phi^{-1}([-\infty,t[).$ 

**Step 1:** For any t < M,  $X_t \cap [0, c] \cap [c, c+1] = \emptyset$ , so

$$H_0(X_t \cap [0,R]) \simeq H_0(X_t \cap [0,c]) \oplus H_0(X_t \cap [c,c+N]) \oplus H_0(X_t \cap [c+N,R])),$$
 (6)

Step 2: similarly,

$$H_0(X_t \cap [c,c+N]) \simeq \bigoplus_{n=1}^N H_0(X_t \cap [c+(n-1),c+n])$$
 (7)

$$(x \mapsto x + n) \simeq \bigoplus_{n=1}^{N} H_0(X_t \cap [c, c+1])$$
 (8)

**Step 3:** The inclusion  $[0, c] \subset [c - 1, c]$  induces an injective morphism

$$H_0(X_t \cap [0,c]) \hookrightarrow H_0(X_t \cap [c-1,c]).$$

#### Outline

Additivity of persistence diagrams of periodic functions

Segmentation of periodic signals and phase estimation

Signatures of periodic signals with phase variation

#### Phase estimation

#### Setting

Consider S a periodic function with phase variation

$$S: [0,T] \rightarrow \mathbb{R}$$
 $t \mapsto \phi(\gamma(t)) + W(t),$ 

#### where

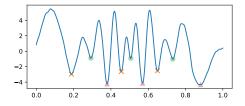
- 1.  $\phi: [0,1] \to \mathbb{R}$  is 1-periodic and unknown,
- 2.  $\gamma: [0, T] \rightarrow [0, N]$  with  $N \in \mathbb{N}$  unknown,
- 3.  $W:[0,T]\to\mathbb{R}$  is a continuous noise process.

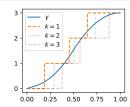
Goal: Given S, estimate  $\gamma$ .

# Proposed solution: segmenting the curve into periods

- 1. Estimate N using D(S).
- 2. Find  $t_1 < \ldots < t_N$  such that  $\gamma(t_n) \gamma(t_{n-1}) = 1$  for all  $n = 2, \ldots, N$ .

Let  $\hat{\gamma}:[0,T]\to\mathbb{R}^*$  be such that  $\hat{\gamma}(t_n)=n$  and interpolate.

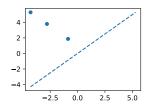




## Estimation of N: noiseless setting

We will denote by D(S)(A) the number of points from D(S) that are in  $A \subset \{(x,y) \in \mathbb{R}^2 \mid y-x>0\}$ .

$$\hat{N}(S) := \gcd\{D(S)(x) \mid x \text{ in } \operatorname{supp}(D(S))\}.$$



#### Proposition

Assume W = 0, so  $S = \phi \circ \gamma$  with  $\gamma : [0, T] \to [0, N]$ . For any  $A \subset \mathbb{R}^2$ ,

$$D(\phi \circ \gamma)(A) = ND(\phi|_{[0,1]})(A). \tag{9}$$

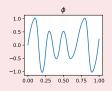
In particular,

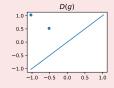
$$\hat{N}(\phi \circ \gamma) = N\hat{N}(\phi|_{[0,1]}).$$

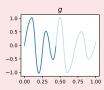
## Estimation of N: correctness in the noiseless setting

## Identifiability

There exists a 1-periodic function g such that  $D(g|_{[0,1]}) = D(\phi|_{[0,1]})/\hat{N}(\phi|_{[0,1]})!$ 







## Non-degeneracy

We say that  $\phi$  is **non-degenerate** if  $\hat{N}(\phi|_{[0,1]}) = 1$ .

## Example

If  $\phi$  has at least one unique critical value, it is non-degenerate.

#### Corollary

When  $\phi$  is non-degenerate,  $\hat{N}(\phi \circ \gamma) = N$ .

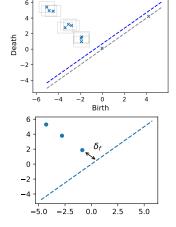
# Estimation of N: noisy signal

For the noisy signal  $S = \phi \circ \gamma + W$ , the points in D(S) have multiplicity 1.

#### Estimator

For  $\tau > 0$ , we define

$$\hat{N}_{\tau}(S) := \gcd\{|D(S)(B(x,\tau))| \mid x \in D(S), \text{pers}(x) > \tau\}.$$
 (10)



#### Separation constant

The separation constant is the smallest distance between points in  $D(\phi)$ ,

$$\delta_{\phi} := \min(d(x_1, x_2), d(x_1, \Delta) \mid x_1, x_2 \in D(\phi)).$$

# Proposition (Stability)

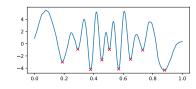
If  $\phi$  is non-degenerate, then for any au>0 such that  $2\|W\|_{\infty}< au<\delta/3$ , we have

$$\hat{N}_{\tau}(S) = N.$$

# Estimation of $\gamma$

#### Persistent minima

Let  $\tau > 0$  and  $\hat{C}_{\tau} = \{t_1, \dots, t_M\} \subset [0, T]$  be the set of local minima of S, corresponding to points in the diagram with persistence more than  $\tau$ .



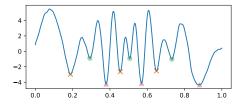
#### Proposition

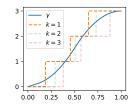
If  $\tau \in ]2||W||_{\infty}, \delta/3[$ , then, for some  $K \in \mathbb{N}$ ,

$$|\hat{\mathcal{C}}_{\tau}| = NK.$$

For each  $k \in \{1, \dots, K\}$ , we can define an estimator of  $\gamma$ 

$$\hat{\gamma}: \quad [0, T] \quad \to \quad \mathbb{R} \\
 t \quad \mapsto \quad \sum_{n=1}^{N} 1_{t_{(n-1)K+k} \le t}.$$
(11)



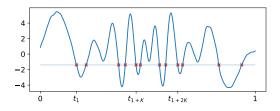


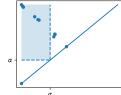
# Zero-crossings<sup>1</sup>

Let  $K := |\phi^{-1}(\alpha) \cap [0,1[]$  and assume that  $0 < K < \infty$ , for some  $\alpha \in \mathbb{R}$ .

#### Estimation of N

If K is known,  $N_{\alpha}(S) := \frac{|S^{-1}(\alpha)|}{K}$  is an estimator of N.





# Segmentation of the signal

If  $S^{-1}(\alpha) = \{t_1, \ldots, t_{NK}\}$ , then  $\gamma(t_{n+k}) - \gamma(t_n) = 1$  for  $1 \le k \le K$  and  $n \le NK - k$ .

#### Issues

- K is not known (and not necessarily finite),
- $N_{\alpha}$  is not stable.

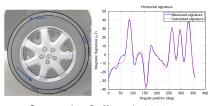
(Tanweer, Khasawneh, and Munch 2023)

<sup>&</sup>lt;sup>1</sup>Boualem Boashash, Peter O'Shea, and Morgan Arnold (1990). "Algorithms for Instantaneous Frequency Estimation: A Comparative Study". In: Advanced Signal Processing Algorithms, Architectures, and Implementations. Vol. 1348. SPIE, pp. 126–148. DOI: 10.1117/12.23471.

## Application: estimating the speed of a moving vehicle

#### Context

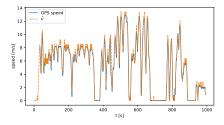
The magnetic signal measured in a car is  $\mathbf{B}(\theta, \theta_h) = Q(\theta_h)\mathbf{B}_E + \mathbf{B}_u(\theta) \in \mathbb{R}^3$ , where  $\theta_h$  is the orientation of the vehicle and  $\theta$  the angular position of a wheel<sup>2</sup>. As the car moves, we observe  $\mathbf{S}(t) = \mathbf{B}(\gamma(t), \gamma_h(t))$ .

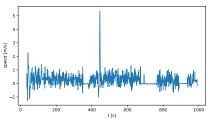


Source: Le Goff et al. 2012.

# Proposed solution

Choose a vector  $v \in \mathbb{S}^2$  and construct  $\hat{\gamma}$  for on  $S := \langle \mathbf{S}, v \rangle$ . Estimate the speed by  $(\hat{\gamma}(t) - \hat{\gamma}(t - t_0))/t_0$ , for some small delay  $t_0$ .

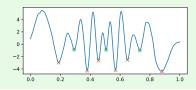


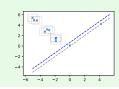


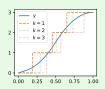
<sup>&</sup>lt;sup>2</sup> Pierre-Jean Bristeau (2012). "Techniques d'estimation du déplacement d'un véhicule sans GPS et autres exemples de conception de systèmes de navigation MEMS". PhD thesis. Ecole Nationale Superieure des Mines de Paris

#### Conclusion and future work

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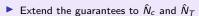




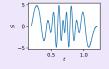


#### Limitations and future work

- 1. Identifiability
  - ▶ Use the order of local minima to lift the identifiability issue
- 2. More robust estimators



► Choose the sets to count multiplicity differently





- 3. The method is applicable only to  $N \in \mathbb{N}$ .
  - In practice, it is not a problem as  $\phi$  is often simple.
  - ▶ Use the approximate greatest common divisor.

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Additivity of persistence diagrams of periodic functions

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#### Problem statement

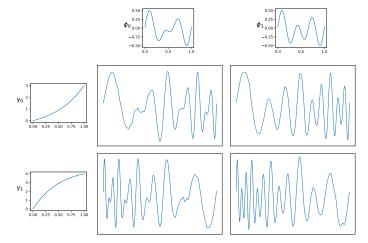
#### Data

Consider  $S = \phi \circ \gamma + W$ , where

- $\phi: \mathbb{R} \to \mathbb{R}$  is 1-periodic,
- $ightharpoonup \gamma: [0, T] \to \mathbb{R}$  an increasing bijection,
- ▶  $W : [0, T] \rightarrow \mathbb{R}$  is a cont. stoch. proc.

Aim

Given S, construct a signature of  $\phi$ .



Studied in Reise, Michel, and Chazal 2023.

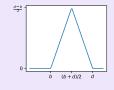
## Functional representations of persistence diagrams

Defining a mean of a collection of persistence diagrams is not necessarily easy, so it is common to compute statistics of diagrams in a functional space<sup>3</sup>.

#### Functional representation

Let  $\mathcal{H}$  be a functional Banach space.

$$\kappa: \quad \mathbb{R}^2 \quad \rightarrow \quad \mathcal{H} \ (b,d) \quad \mapsto \quad \kappa_{(b,d)}: \quad \mathbb{T} \quad \rightarrow \quad \mathbb{R} \ x \quad \mapsto \quad \kappa_{(b,d)}(x).$$





Persistence silhouette 4

Persistence image<sup>5</sup>

#### Definition

For  $p \ge 1$  and  $\epsilon > 0$ , the  $\epsilon$ -truncated p-persistence of (b,d) is  $w(d-b) = \max(d-b-\epsilon,0)^p$ . We define the **normalized functional** of D as persistence-weighted average of  $\kappa$ ,

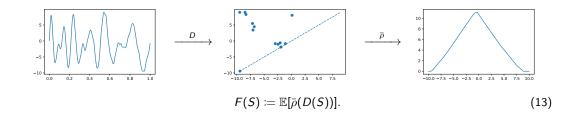
$$\bar{\rho}.(D): \quad \mathbb{T} \quad \to \quad \mathbb{R} \\ x \quad \mapsto \quad \frac{\sum_{(b,d)\in D} w(d-b)\kappa_{(b,d)}(x)}{\sum_{(b,d)\in D} w(d-b)}. \tag{12}$$

<sup>&</sup>lt;sup>3</sup>Frédéric Chazal and Bertrand Michel (2021). "An Introduction to Topological Data Analysis: Fundamental and Practical Aspects for Data Scientists". In: Frontiers in Artificial Intelligence 4. ISSN: 2624-8212.

<sup>&</sup>lt;sup>4</sup> Peter Bubenik (2015). "Statistical Topological Data Analysis Using Persistence Landscapes". In: Journal of Machine Learning Research 16.1, pp. 77–102

<sup>&</sup>lt;sup>5</sup>Henry Adams et al. (2017). "Persistence Images: A Stable Vector Representation of Persistent Homology". In: The Journal of Machine Learning Research 18.1, pp. 218–252

## Proposed approach: normalized functionals of persistence



## Properties of signatures

1. Consistency: thanks to the additivity of persistence,

$$\bar{\rho}(D(\phi|_{[0,R]})) \xrightarrow{\|\cdot\|_{\mathcal{H}}} \bar{\rho}(D(\phi|_{[c,c+1]})), \quad \text{as } R \to \infty.$$
 (14)

- 2. Stability: when  $\gamma$  and W are random and independent, how does F(S) depend on the law of  $\gamma$ ?
- 3. Estimation: how to estimate the signature from a sampled time series?

## Stability of the signature

#### A model for S

Let  $\mu$  be a probability measure on  $(\Gamma_{0,R,\nu_{\min}},\mathcal{B}(\|\cdot\|_{\infty}))$  for some  $\nu_{\min}>0$ , where

$$\Gamma_{0,R,v_{\min}} = \{ \gamma \in C([0,T],\mathbb{R}) \mid \gamma(0) = 0, \ \gamma(T) = R, \ \gamma(s) - \gamma(t) \ge v_{\min}(s-t), \ \text{ for all } s \ge t \}, \quad (15)$$

Let  $\nu$  be a probability measure on  $(C([0,T],\mathbb{R}),\mathcal{B}(\|\cdot\|_{\infty}))$ , such that

$$||W||_{\infty} \le (\max \phi - \min \phi)/2 - \epsilon \text{ almost-surely,}$$
 (16)

$$t \mapsto W(t)$$
 has an  $\alpha$ -Hölder version. (17)

Let  $S := \phi \circ \gamma + W$ , where  $\gamma \sim \mu$  and  $W \sim \nu$  are independent.

#### Theorem

If  $\mu_1$ ,  $\mu_2$  are two probability measures on  $\Gamma_{0,R,\nu_{min}}$  and  $S_k = \phi \circ \gamma_k + W$ , then

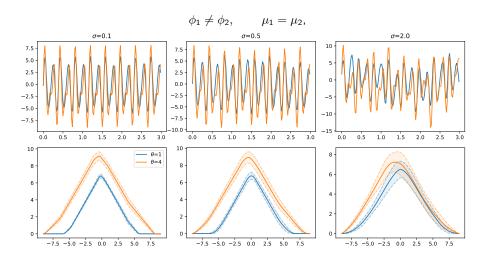
$$||F(S_1) - F(S_2)||_{\mathcal{H}} \le \frac{C}{v_{\min}^{\alpha}} \mathcal{W}_1(\mu_1, \mu_2)^{\alpha},$$
 (18)

where  $W_1$  is the Wasserstein distance, and C depends on the regularity of W,  $\|W\|_{\infty}$ ,  $\epsilon$ , p and  $\kappa$ .

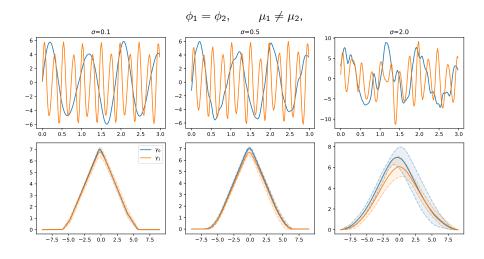
#### Comment

- $\checkmark$  As  $c \to 0$ ,  $||F(\phi \circ \gamma_1) F(\phi \circ \gamma_2 + cW)||_{\mathcal{H}} \to 0$ .
- X How to remove the fixed-endpoints assumption in (15)?

# Numerical examples: stability



# Numerical examples: stability



#### Estimation of signatures: introduction

Assume that only a single time series  $(S_n)_{n=1}^N \subset \mathbb{R}$  is given,

$$S_n = \phi(\gamma(t_n)) + W(t_n).$$

Can we estimate the signature?

Proposition (Chazal et al. 2014<sup>6</sup>, Berry et al. 2018<sup>7</sup>)

Let be  $D_1, \ldots, D_N$  i.i.d.persistence diagrams. When the (bracketing) entropy of  $(\bar{\rho}_x)_{x \in \mathbb{T}}$  is finite,

$$\sqrt{N}\left(\frac{1}{N}\sum_{n=1}^{N}\bar{\rho}(D_n)-\bar{\rho}^*\right)\stackrel{d}{\to}\mathbb{G},\tag{19}$$

for a zero-mean stochastic process G.

#### Procedure

We fix  $M \in \mathbb{N}$  and we generate  $S_1, \dots S_{N-M+1}$ , where

$$\mathbf{S}_n = (S_n, \ldots, S_{n+M-1}).$$

#### Challenge

 $S_1, \ldots, S_{N-M+1}$  are not independent! Under what assumptions on  $(\gamma(t_n))_{n\in\mathbb{N}}$  and  $(W_n)_{n\in\mathbb{N}}$  does an analogue of (19) hold?

<sup>&</sup>lt;sup>6</sup>Frédéric Chazal et al. (2014). "Stochastic Convergence of Persistence Landscapes and Silhouettes". In: Annual Symposium on Computational Geometry - SOCG'14. Kyoto, Japan: ACM Press, pp. 474–483. ISBN: 978-1-4503-2594-3. DOI: 10.1145/2582112.2582128

<sup>&</sup>lt;sup>7</sup> Eric Berry et al. (2018). Functional Summaries of Persistence Diagrams. arXiv: 1804.01618

## Quantifying dependence

#### Definition ( $\beta$ -mixing coefficients, Dedecker et al. 2007)

Let  $(X_n)_{n\in\mathbb{Z}}$  be a stationary sequence of random variables on a common measurable space. Then,

$$\beta_X(k) := \sup_{A,B} \sum_{A \in A,B \in B} |P(A \cap B) - P(A)P(B)|,$$

where  $\mathcal{A} \subset \sigma^X_{-\infty,0}$ ,  $\mathcal{B} \subset \sigma^X_{k,\infty}$  are finite partitions of the sample space and  $\sigma^X_{a,b} := \sigma((X_n)_{a \le n \le b})$ .

#### Proposition (Kosorok 2008)

If  $(\bar{\rho}_x)_{x\in\mathbb{T}}$  has finite bracketing entropy and  $(\mathbf{S}_n)_{n\in\mathbb{N}}$  is stationary with  $\beta_{\mathbf{S}}(k) = O(k^{-3})$ , then

$$\sqrt{N}\left(\frac{1}{N}\sum_{n=1}^{N}\bar{\rho}(D(\mathbf{S}_n))-\bar{\rho}^*\right)\stackrel{d}{\to}\mathbb{G}_{dep},\tag{20}$$

where  $\mathbb{G}_{dep}$  is a zero-mean stochastic process.

## Proposition

For  $k \geq M + 1$ ,

$$\beta_{S}(k) \le \beta_{S}(k-M+1) \le \beta_{\phi(\gamma)}(k-M+1) + \beta_{W}(k-M+1),$$
 (21)

and

$$\beta_{\phi(\gamma)}(k) \leq \beta_{\operatorname{frac}(\gamma)}(k),$$

where  $frac(x) := x - \lfloor x \rfloor$ .

# Model for $(\gamma_n)_{n\in\mathbb{N}}$

#### Random walk model

For some h > 0, we set

$$\gamma_{n+1} = \gamma_n + hV_n,$$

for  $(V_n)_{n\in\mathbb{N}}\sim \mathbf{P}$  i.i.d. We assume that

- ▶  $supp(P) \subseteq [v_{min}, v_{max}] \subset ]0, \infty[$ ,
- ▶ for some c > 0 and a non-trivial interval  $I \subset [v_{\min}, v_{\max}]$ ,

$$\mathbf{P}(A) \ge c\lambda(A), \quad \text{for all } A \in \mathcal{B}(I).$$
 (22)

#### Proposition

If  $\gamma_0 \sim \mathcal{U}([0,1])$ , then  $(\operatorname{frac}(\gamma_n))_{n \in \mathbb{N}}$  is stationary and  $\beta_{\operatorname{frac}(\gamma)}(k) = O(e^{-ak})$  for some a > 0.

## Idea of the proof

1. By Thm 1 in Section 2.4 of Doukhan 1995, it suffices to show the **Doeblin condition**:

There is  $\mu_0$  and  $n_0 \in \mathbb{N}$ , such that for all  $n \geq n_0$  uniformly in  $x_0$ ,

$$P(\operatorname{frac}(\gamma_n) \in A \mid \gamma_0 = x_0) \ge \mu_0(A).$$

- 2.  $\sum_{k=1}^{n} V_k \sim \mathbf{P}^{\star n}$ , with  $\mathbf{P}^{\star n}$  lower-bounded by a uniform measure with growing support.
- 3. For  $n_0 \in \mathbb{N}$  big enough, the support is of length at least 1, and we obtain a lower-bound for the distribution of  $\operatorname{frac}(\sum_{k=1}^{n_0} V_k)$  on ]0,1[.

#### Conclusion and future work

## Conclusion

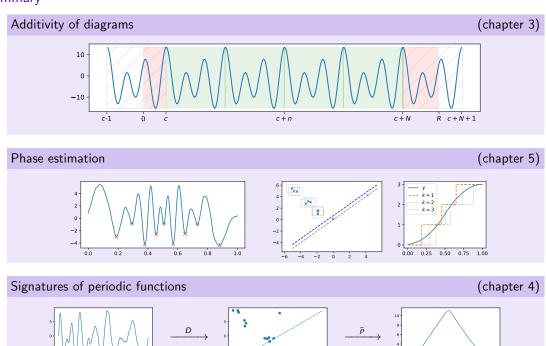
F(S) is a stable signature of  $\phi$  and can be estimated with standard techniques.

#### Limitations and future work

- Remove the assumption of fixed endpoints from the stability
  - ► Technical difficulties in defining the probability measures
  - ▶ Understand the distance between  $D(f|_{[0,T]})$  and  $D(f|_{[0,t]}) \cup D(f|_{[t,T]})$
- Numerical experiments to understand the discriminative power
  - Compare with registration-based methods.
  - Understand how the choice of the kernel

0.2 0.4 0.6

# Summary



-10.0 -7.5 -5.0 -2.5 0.0 2.5 5.0 7.5

-10.0 -7.5 -5.0 -2.5 0.0 2.5 5.0 7.5 10.0

#### Refernces I



Adams, Henry et al. (2017). "Persistence Images: A Stable Vector Representation of Persistent Homology". In: *The Journal of Machine Learning Research* 18.1, pp. 218–252.



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# Thank you!

## Persistence diagram of sub level sets: Definition

### 1. Persistence module

For each  $t \in \mathbb{R}$ , we calculate  $H_0(X_t)$ , where  $X_t := f^{-1}(]-\infty,t]$ ). For any  $s \leq t$ , the inclusion  $X_s \to X_t$  gives a map  $\iota_s^t : H_0(X_s) \to H_0(X_t)$ .

#### 2. Rectangle measure

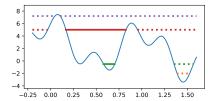
A measure m on rectangles of  $\mathbb{R}^2$ .

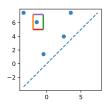
$$m(]a,b] \times [c,d[) = \dim \left( \frac{im(\iota_b^c) \cap \ker(\iota_c^d)}{im(\iota_o^c) \cap \ker(\iota_c^d)} \right),$$

#### 3. Persistence diagram

The persistence diagram D(f) is a multi-set in  $\mathbb{R}^2$ , where  $(s,t) \in \mathbb{R}^2$  has multiplicity

$$m(s,t) = \lim_{s \to 0^+} m(]s - \delta, s + \delta] \times [t - \delta, t + \delta].$$





#### Proof of additivity of sub level sets: details

#### Proof.

Let  $c := \inf\{x \in [0,1] | \phi(x) = \max \phi\}, \ N = \max\{n \in \mathbb{N} | c+n \le R\}$  and denote by  $\mathbb{X}_t := \phi^{-1}([-\infty, t]).$ 

Step 1: For any t < M,  $X_t \cap [0, c] \cap [c, c+1] = \emptyset$ , so

$$H_0(\mathbb{X}_t \cap [0, R]) \simeq H_0(\mathbb{X}_t \cap [0, c]) \oplus H_0(\mathbb{X}_t \cap [c, c + N]) \oplus H_0(\mathbb{X}_t \cap [c + N, R]),$$
 (23)

Step 2: similarly,

$$H_0(\mathbb{X}_t \cap [c, c+N]) \simeq \bigsqcup_{n=1}^N H_0(\mathbb{X}_t \cap [c+(n-1), c+n])$$

$$(x \mapsto x+n) \quad \simeq \bigsqcup_{n=1}^N H_0(\mathbb{X}_t \cap [c, c+1])$$

$$(24)$$

$$(x \mapsto x + n) \simeq \bigsqcup_{n=1}^{N} H_0(\mathbb{X}_t \cap [c, c+1])$$
 (25)

Step 3: The inclusion  $[0, c] \subset [c-1, c]$  induces an injective morphism

$$H_0(\mathbb{X}_t \cap [0,c]) \hookrightarrow H_0(\mathbb{X}_t \cap [c-1,c]).$$

## Stability: bottleneck distance (detailed)

# Definition (Edelsbrunner and Harer 2010, p. VIII.2)

We call a  $\epsilon$ -matching between two persistence diagrams D and D' a bijection  $\Gamma: A \to A'$  between some subsets of  $A \subset D$  and  $A' \subset D'$ , considered with multiplicity, if

$$d_{\infty}(a, \Gamma(a)) \le \epsilon,$$
 for any  $a \in A$ ,  
 $d_{\infty}(a, \Delta) \le \epsilon,$  for any  $a \in (D \setminus A) \cup (D' \setminus A').$ 

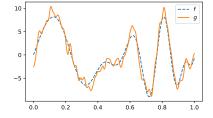
where  $\Delta = \{(x, x) \in \mathbb{R}^2\}$  denotes the diagonal.

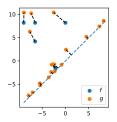
$$d_B(D, D') := \inf\{\epsilon > 0 \mid \Gamma \text{ is an } \epsilon\text{-matching between } D' \text{ and } D'\}.$$

# Theorem (Bottleneck stability of diagrams)

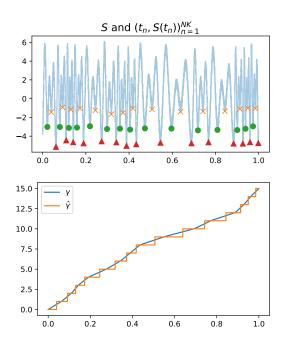
Let  $f, g : \mathbb{X} \to \mathbb{R}$  be two continuous functions on a compact space  $\mathbb{X}$ . Then,

$$d_B(D(f),D(g)) \leq ||f-g||_{\infty}.$$



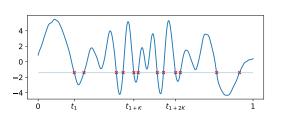


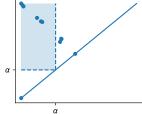
# Landmarks for multiple periods



## Zero-crossings from the persistence diagram

$$|S^{-1}(\alpha)| = 2 \lim_{\delta \to 0^+} |D(S) \cap (] - \infty, \alpha - \delta] \times [\alpha + \delta, \infty[)|.$$





## Counting measure

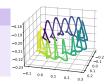
The persistence diagram D is also a counting measure on rectangles  $A \subset \Delta_+ = \{(b,d) \in \mathbb{R}^2 \mid x < y\}$ . By (4),

$$|D(\phi \circ \gamma) \cap A| = N |D(\phi|_{[0,1]}) \cap A|$$

## Application: magnetic odometry and speed estimation

#### Problem

Using the magnetic signal **B**, recorded in a moving car, estimate the cars' trajectory. The angular position  $t\mapsto \gamma(t)$  of a wheel in time is visible through  $\mathbf{S}(t)=\mathbf{B}(\gamma(t),\gamma_h(t))$ 



#### Proposed solution

- 1.  $S := \langle \mathbf{S}, v \rangle$ , project **S** along a suitable direction  $v \in \mathbb{S}^2$
- 2.  $\hat{N}_{c,\tau}(S)$ , for an appropriate scale  $\tau$ ,
- 3. Derive an odometric sequence  $t_1, \ldots, t_{\hat{N}_{\tau}(S)}$  from  $\mathcal{C}_{\tau}$ .
- 4. Construct  $\hat{\gamma}:[0,T]\to\mathbb{R}$ .

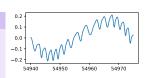
#### Results

	$E_O$		$E_{l}$	
Method	$S_{\nu_1}$	$( abla {\sf S})_{ u_1}$	$S_{v_1}$	$( abla {\sf S})_{ u_1}$
$\hat{N}_{c, au}$	15.75	16.66	3.02	3.01
$\hat{N}_{0, au}$	15.75	16.66	3.02	3.01
ZC	9.91	5.62	6.35	16.51

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Using the magnetic signal **B**, recorded in a moving car, estimate the cars' trajectory. The angular position  $t \mapsto \gamma(t)$  of a wheel in time is visible through  $\mathbf{S}(t) = \mathbf{B}(\gamma(t), \gamma_h(t))$ 



#### Proposed solution

- 1.  $S := \langle \mathbf{S}, v \rangle$ , project **S** along a suitable direction  $v \in \mathbb{S}^2$
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- 3. Derive an odometric sequence  $t_1, \ldots, t_{\hat{N}_{\tau}(S)}$  from  $\mathcal{C}_{\tau}$ .
- 4. Construct  $\hat{\gamma}:[0,T]\to\mathbb{R}$ .

#### Results

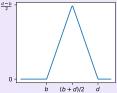
	$E_O$		$E_{l}$	
Method	$S_{\nu_1}$	$( abla {\sf S})_{ u_1}$	$S_{v_1}$	$( abla {\sf S})_{ u_1}$
$\hat{ extsf{N}}_{c, au}$	15.75	16.66	3.02	3.01
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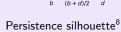
## Normalized functionals of persistence

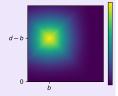
#### Functional representation

Let  $\mathcal{H}$  be a functional Banach space

- 1.  $\operatorname{supp}(\kappa_{(b,d)}) \subset K$ , K bounded,
- 2.  $x \mapsto \kappa_{(b,d)}(x)$  (uniformly) Lipschitz,
- 3.  $\|\kappa_{(b,d)} \kappa_{(b',d')}\|_{\mathcal{H}} \leq L_{\kappa} \|(b,d) (b',d')\|_{\kappa}$
- 4.  $\|\kappa_{(b,b)}\|_{\mathcal{H}} \leq C$ .







Persistence image<sup>9</sup>

# Normalized functionals of persistence diagrams

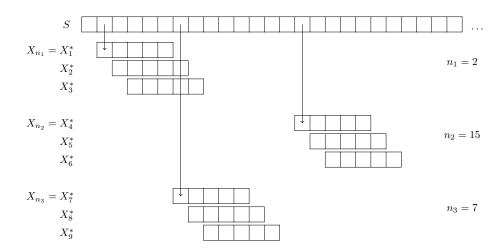
For some  $p \ge 1$  and  $\epsilon > 0$ ,

$$\bar{\rho}(D) := \frac{\sum_{(b,d)\in D} w(d-b)\kappa_{(b,d)}}{\sum_{(b,d)\in D} w(d-b)}, \quad \text{where } w(d-b) = \max(d-b-\epsilon,0)^p. \tag{26}$$

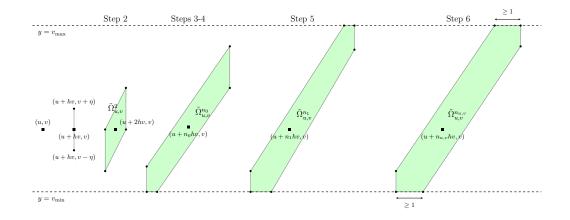
<sup>&</sup>lt;sup>8</sup>Peter Bubenik (2015). "Statistical Topological Data Analysis Using Persistence Landscapes". In: Journal of Machine Learning Research 16.1, pp. 77–102

<sup>&</sup>lt;sup>9</sup>Henry Adams et al. (2017). "Persistence Images: A Stable Vector Representation of Persistent Homology". In: The Journal of Machine Learning Research 18.1, pp. 218–252

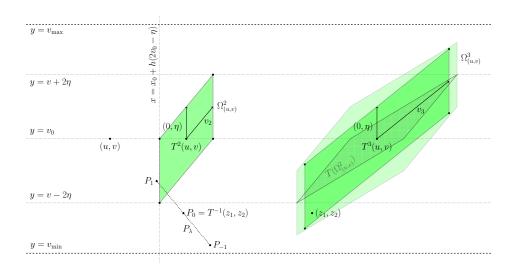
# Bootstrap



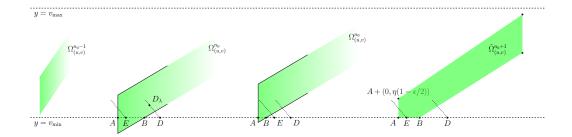
# Mixing: overview



# Mixing: initial



# Mixing: boundary



## Measures of dependence

# Types of dependence

There are different ways to measure dependence in a time series  $(X_n)_{n\in\mathbb{N}}\subset\mathbb{X}$ :

- ► *m*-dependence.
- strong-mixing,
- weak-dependence,

#### Strong mixing

The  $\beta$ -mixing coefficient of a time series  $(X_n)_{n\in\mathbb{N}}\subset\mathbb{X}$  is

$$\beta_X(k) = \sup_{A,B} \sum_{A \in A, B \in B} |P(A \cap B) - P(A)P(B)|,$$

where  $A \subset \sigma^X_{-\infty,0}$ ,  $B \subset \sigma^X_{k,\infty}$  are finite partitions of the sample space and  $\sigma^X_{a,b} \coloneqq \sigma((X_n)_{a \le n \le b})$ .

#### Example

- 1. If  $(X_n)_n$  is *m*-dependent, then  $\beta_X(k) = 0$  for  $k \ge m$ .
- 2. Markov chains: irreducible and aperiodic.

#### Proposition

For any measurable function  $f : \mathbb{X} \to \mathbb{Y}$ ,  $\beta_X(k) \leq \beta_{f(X)}(k)$ .

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---> preserved by measurable functions!

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