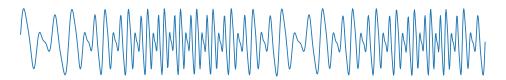
Topological techniques for inference on periodic functions with phase variation

Wojciech Reise Under the supervision of Frédéric Chazal and Bertrand Michel

Dec 6, 2023, Orsay







Data with phase variation

Signals with phase variation

A sample $S_1, \ldots, S_N : [0,1] \to \mathbb{X}$ has **phase variation** if

$$S_n = f(\gamma_n) + W_n, \quad \text{for each } n \in \{1, \dots, N\},$$

where $\gamma_1, \ldots, \gamma_N : [0,1] \to [0,1]$ are increasing homeomorphisms, $f : [0,1] \to \mathbb{X}$ is continuous and $W_n : [0,T] \to \mathbb{R}$ is a noise process.

Litterature

- Curve registration: estimating $\gamma_n \circ \gamma_{n'}^{-1}$ (Tang and Muller 2008, Zhao and Itti 2018)
- \triangleright Computing a representative of f (Su et al. 2014)
- ▶ Clustering of $S_1, ..., S_n$ (Srivastava et al. 2011)

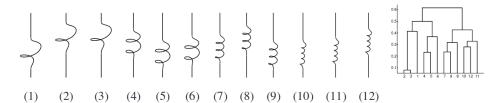
See Marron et al. 2015 for a review.

Fixed endpoints assumption

For all $1 \le n \le N$,

$$\gamma_1(0) = \gamma_n(0), \tag{2}$$

$$\gamma_1(1) = \gamma_n(1). \tag{2}$$



Source: Srivastava et al. 2011.

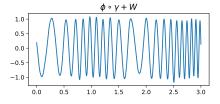
Periodic data with phase variation

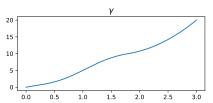
Definition

We call $S:[0,1] \to \mathbb{X}$ a periodic function with phase variation if

$$S(t) = \phi(\gamma(t)) + W(t) \tag{3}$$

where $\phi: \mathbb{R} \to \mathbb{X}$ is 1-periodic, $\gamma: [0,1] \to [0,R]$ is an increasing homeomorphism and $W: [0,1] \to \mathbb{X}$ is a noise process.





Example (Instantaneous phase estimation, Boashash, O'Shea, and Arnold 1990)

Decompose $s(t) = a(t)\cos(\gamma_0(t))$ into an amplitude a(t), and a phase-variation component $\gamma_0(t) = \arctan(H(s(t))/s(t))$.

Topological data analysis for periodic time series

We study S using persistent homology, a technique from Topological Data Analysis (TDA).

Contributions	
We describe the structure of a topological descriptor of $\phi \circ \gamma$. Let S be a periodic function with phase variation.	(Chapter 3)
1. We propose an estimator of γ from ${\cal S}$,	(Chapter 5)
2. We construct a descriptor of ϕ from S .	(Chapter 4)

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TDA for time series

- Detecting periodicity in a time series (Perea 2019),
- Detecting financial crashes (Gidea and Katz 2018)
- Robust zero-crossings (Khasawneh and Munch 2018; Tanweer, Khasawneh, and Munch 2023),
- ▶ Analysis of gate signals for the study of multiple sclerosis (Bois et al. 2022).

Outline

Additivity of persistence diagrams of periodic functions

Segmentation of periodic signals and phase estimation

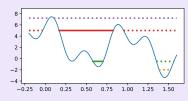
Signatures of periodic signals with phase variation

Intuition

The persistence diagram D(f) of a continuous function $f:[0,T]\to\mathbb{R}$ is a multi-set of points in \mathbb{R}^2 , which reflect when connected components appear and merge in $(f^{-1}(]-\infty,t])_{t\in\mathbb{R}}$ as t increases.

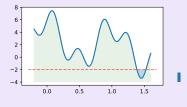


Definition (Chazal et al. 2016)

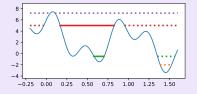


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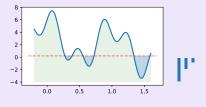
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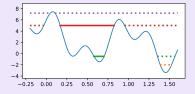


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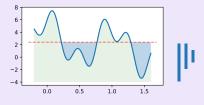
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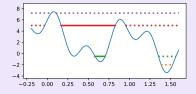


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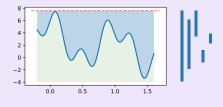
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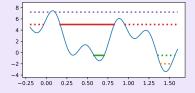


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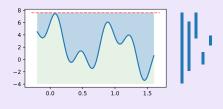
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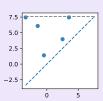




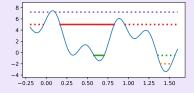
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Definition (Chazal et al. 2016)





Stability: bottleneck distance

Definition (Bottleneck distance)

We call a ϵ -matching between two persistence diagrams D and D' a bijection $\Gamma: A \to A'$ between some subsets of $A \subset D$ and $A' \subset D'$, considered with multiplicity, if

$$d_{\infty}(a, \Gamma(a)) \le \epsilon,$$
 for any $a \in A$,
 $d_{\infty}(a, \Delta) \le \epsilon,$ for any $a \in (D \setminus A) \cup (D' \setminus A').$

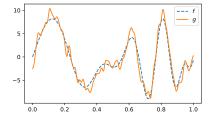
where $\Delta = \{(x,x) \in \mathbb{R}^2\}$ denotes the diagonal.

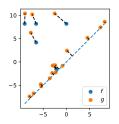
$$d_B(D,D') \coloneqq \inf\{\epsilon > 0 \mid \Gamma \text{ is an } \epsilon\text{-matching between } D' \text{ and } D'\}.$$

Theorem (Bottleneck stability, Edelsbrunner and Harer 2010)

Let $f,g:\mathbb{X}\to\mathbb{R}$ be two continuous functions on a compact space $\mathbb{X}.$ Then,

$$d_B(D(f),D(g)) \leq ||f-g||_{\infty}.$$





Total *p*-persistence

Definition (Total persistence)

The **persistence** of $(b, d) \in D$ is d - b. The **total** *p*-**persistence** of a diagram *D* is

$$\operatorname{pers}_p(D) \coloneqq \left(\sum_{(b,d) \in D} (d-b)^p\right)^{1/p}.$$

Proposition (Plonka and Zheng 2016, Perez 2022)

For
$$p = 1$$
,

$$pers_1(D(f)) + pers_1(D(-f)) = TV(f).$$

If f is α -Hölder for $p > 1 + 1/\alpha$, then, $\operatorname{pers}_p(D(f)) < \infty$.

Persistence diagrams of periodic functions

Let $\phi : \mathbb{R} \to \mathbb{R}$ be a 1-periodic function and denote by $\phi|_{[a,b]}$ the restriction of ϕ to an interval [a,b].

Proposition (Invariance to reparametrisation)

Let $\gamma:[0,1]\to [0,1]$ be an increasing homeomorphism. Then, $D(\phi\circ\gamma)=D(\phi|_{[0,1]})$.

Theorem (Additivity of persistence diagrams for periodic functions)

For $R \in \mathbb{N}^*$, there exists $c \in [0,1]$ such that

$$D(\phi|_{[0,R]}) = RD(\phi|_{[c,c+1]}). \tag{4}$$

For any R > 1,

$$D(\phi|_{[0,R]}) = \lfloor R - 1 \rfloor D(\phi|_{[c,c+1]}) + D', \quad \text{with } \operatorname{pers}_{\rho}(D') \le 2\operatorname{pers}_{\rho}(D(\phi|_{[c,c+1]})). \tag{5}$$

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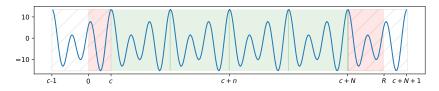
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Conclusion

The persistence diagram $D(\phi \circ \gamma)$ contains information about

- \triangleright extrema of ϕ ,
- ▶ number of periods $(\gamma(1) \gamma(0))$.

Proof of (5)



Proof.

Let $c := \inf\{x \in [0,1[\mid \phi(x) = \max \phi\}, \ N = \max\{n \in \mathbb{N} \mid c+n \leq R\} \text{ and denote by } X_t := \phi^{-1}([-\infty,t[).$

Step 1: For any t < M, $X_t \cap [0, c] \cap [c, c+1] = \emptyset$, so

$$H_0(X_t \cap [0,R]) \simeq H_0(X_t \cap [0,c]) \oplus H_0(X_t \cap [c,c+N]) \oplus H_0(X_t \cap [c+N,R])),$$
 (6)

Step 2: similarly,

$$H_0(X_t \cap [c,c+N]) \simeq \bigoplus_{n=1}^N H_0(X_t \cap [c+(n-1),c+n])$$
 (7)

$$(x \mapsto x + n) \simeq \bigoplus_{n=1}^{N} H_0(X_t \cap [c, c+1])$$
 (8)

Step 3: The inclusion $[0, c] \subset [c - 1, c]$ induces an injective morphism

$$H_0(X_t \cap [0,c]) \hookrightarrow H_0(X_t \cap [c-1,c]).$$

Outline

Additivity of persistence diagrams of periodic functions

Segmentation of periodic signals and phase estimation

Signatures of periodic signals with phase variation

Phase estimation

Setting

Consider S a periodic function with phase variation

$$S: [0,T] \rightarrow \mathbb{R}$$
 $t \mapsto \phi(\gamma(t)) + W(t),$

where

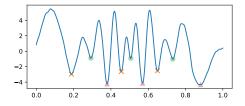
- 1. $\phi: [0,1] \to \mathbb{R}$ is 1-periodic and unknown,
- 2. $\gamma: [0, T] \rightarrow [0, N]$ with $N \in \mathbb{N}$ unknown,
- 3. $W:[0,T]\to\mathbb{R}$ is a continuous noise process.

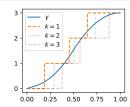
Goal: Given S, estimate γ .

Proposed solution: segmenting the curve into periods

- 1. Estimate N using D(S).
- 2. Find $t_1 < \ldots < t_N$ such that $\gamma(t_n) \gamma(t_{n-1}) = 1$ for all $n = 2, \ldots, N$.

Let $\hat{\gamma}:[0,T]\to\mathbb{R}^*$ be such that $\hat{\gamma}(t_n)=n$ and interpolate.

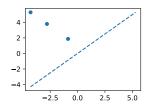




Estimation of N: noiseless setting

We will denote by D(S)(A) the number of points from D(S) that are in $A \subset \{(x,y) \in \mathbb{R}^2 \mid y-x>0\}$.

$$\hat{N}(S) := \gcd\{D(S)(x) \mid x \text{ in } \operatorname{supp}(D(S))\}.$$



Proposition

Assume W = 0, so $S = \phi \circ \gamma$ with $\gamma : [0, T] \to [0, N]$. For any $A \subset \mathbb{R}^2$,

$$D(\phi \circ \gamma)(A) = ND(\phi|_{[0,1]})(A). \tag{9}$$

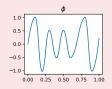
In particular,

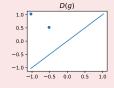
$$\hat{N}(\phi \circ \gamma) = N\hat{N}(\phi|_{[0,1]}).$$

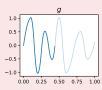
Estimation of N: correctness in the noiseless setting

Identifiability

There exists a 1-periodic function g such that $D(g|_{[0,1]}) = D(\phi|_{[0,1]})/\hat{N}(\phi|_{[0,1]})!$







Non-degeneracy

We say that $\phi|_{[0,1]}$ is non-degenerate if $\hat{N}(\phi|_{[0,1]})=1.$

Example

If ϕ has at least one unique critical value, it is non-degenerate.

Corollary

When ϕ is non-degenerate, $\hat{N}(\phi \circ \gamma) = N$.

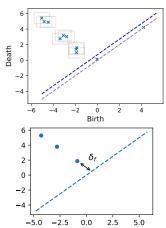
Estimation of N: noisy signal

For the noisy signal $S = \phi \circ \gamma + W$, the points in D(S) have multiplicity 1.

Estimator

For $\tau > 0$, we define

$$\hat{N}_{\tau}(S) := \gcd\{|D(S)(B(x,\tau))| \mid x \in D(S), \text{pers}(x) > \tau\}.$$
 (10)



Separation constant

The separation constant is the smallest distance between points in $D(\phi)$,

$$\delta_{\phi} := \min(d(x_1, x_2), d(x_1, \Delta) \mid x_1, x_2 \in D(\phi)).$$

Proposition (Stability)

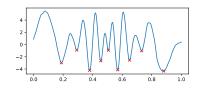
If ϕ is non-degenerate, then for any $\tau > 0$ such that $2\|W\|_{\infty} < \tau < \delta/3$, we have

$$\hat{N}_{\tau}(S) = N.$$

Estimation of γ

Persistent minima

Let $\tau > 0$ and $\hat{C}_{\tau} = \{t_1, \dots, t_M\} \subset [0, T]$ be the set of local minima of S, corresponding to points in the diagram with persistence more than τ .



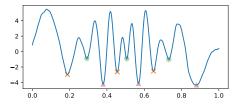
Proposition

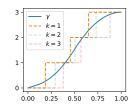
If
$$\tau \in]2||W||_{\infty}, \delta/3[$$
, then

$$|\hat{\mathcal{C}}_{\tau}| = NK.$$

For each $k \in \{1, \dots, K\}$, we can define an estimator of γ

$$\hat{\gamma}: \quad [0, T] \quad \to \quad \mathbb{R} \\
 t \quad \mapsto \quad \sum_{n=1}^{N} 1_{t_{(n-1)K+k} \le t}.$$
(11)



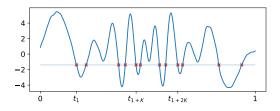


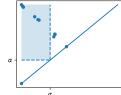
Zero-crossings¹

Let $K := |\phi^{-1}(\alpha) \cap [0,1[]$ and assume that $0 < K < \infty$, for some $\alpha \in \mathbb{R}$.

Estimation of N

If K is known, $N_{\alpha}(S) := \frac{|S^{-1}(\alpha)|}{K}$ is an estimator of N.





Segmentation of the signal

If $S^{-1}(\alpha) = \{t_1, \ldots, t_{NK}\}$, then $\gamma(t_{n+k}) - \gamma(t_n) = 1$ for $1 \le k \le K$ and $n \le NK - k$.

Issues

- K is not known (and not necessarily finite),
- N_{α} is not stable.

(Tanweer, Khasawneh, and Munch 2023)

¹Boualem Boashash, Peter O'Shea, and Morgan Arnold (1990). "Algorithms for Instantaneous Frequency Estimation: A Comparative Study". In: Advanced Signal Processing Algorithms, Architectures, and Implementations. Vol. 1348. SPIE, pp. 126–148. DOI: 10.1117/12.23471.

Application: estimating the speed of a moving vehicle

Context

The magnetic signal measured in a car is $\mathbf{B}(\theta, \theta_h) = Q(\theta_h)\mathbf{B}_E + \mathbf{B}_u(\theta) \in \mathbb{R}^3$, where θ_h is the orientation of the vehicle and θ the angular position of a wheel². As the car moves, we observe $\mathbf{S}(t) = \mathbf{B}(\gamma(t), \gamma_h(t))$.

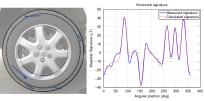
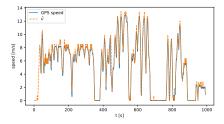
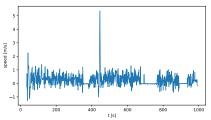


Fig. 2. Results of the optimization. Left: magnetic dipoles found after optimization. Right: measured and calculated signature

Proposed solution

Choose a vector $v \in \mathbb{S}^2$ and construct $\hat{\gamma}$ for on $S := \langle \mathbf{S}, v \rangle$. Define $\hat{v}(t) := \hat{\gamma}(t) - \hat{\gamma}(t - t_0)$, for some small delay t_0 .

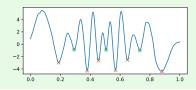


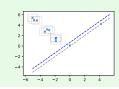


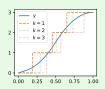
² Pierre-Jean Bristeau (2012). "Techniques d'estimation du déplacement d'un véhicule sans GPS et autres exemples de conception de systèmes de navigation MEMS". PhD thesis. Ecole Nationale Superieure des Mines de Paris

Conclusion and future work

Conclusion

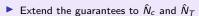




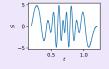


Limitations and future work

- 1. Identifiability
 - ▶ Use the order of local minima to lift the identifiability issue
- 2. More robust estimators



► Choose the sets to count multiplicity differently





- 3. The method is applicable only to $N \in \mathbb{N}$.
 - In practice, it is not a problem as ϕ is often simple.
 - ▶ Use the approximate greatest common divisor.

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Additivity of persistence diagrams of periodic functions

Segmentation of periodic signals and phase estimation

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Problem statement

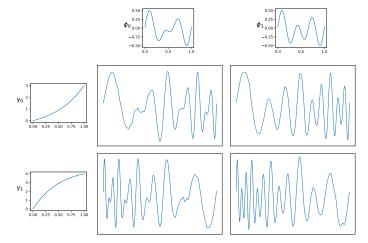
Data

Consider $S = \phi \circ \gamma + W$, where

- $\phi: \mathbb{R} \to \mathbb{R}$ is 1-periodic,
- $ightharpoonup \gamma: [0, T] \to \mathbb{R}$ an increasing bijection,
- ▶ $W : [0, T] \rightarrow \mathbb{R}$ is a cont. stoch. proc.

Aim

Given S, construct a signature of ϕ .



Studied in Reise, Michel, and Chazal 2023.

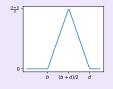
Functional representations of persistence diagrams

Defining a mean of a collection of persistence diagrams is not necessarily easy, so it is common to compute statistics of diagrams in a functional space³.

Functional representation

Let ${\cal H}$ be a functional Banach space.

$$\kappa: \quad \mathbb{R}^2 \quad \rightarrow \quad \mathcal{H} \ (b,d) \quad \mapsto \quad \kappa_{(b,d)}: \quad \mathbb{T} \quad \rightarrow \quad \mathbb{R} \ x \quad \mapsto \quad \kappa_{(b,d)}(x).$$





Persistence silhouette 4

Persistence image⁵

Definition

For $p \ge 1$ and $\epsilon > 0$, the ϵ -truncated p-persistence of (b,d) is $w(d-b) = \max(d-b-\epsilon,0)^p$. We define the **normalized functional** of D as persistence-weighted average of κ ,

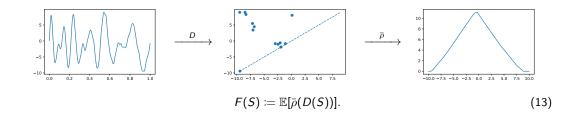
$$\bar{\rho}(D) := \frac{\sum_{(b,d)\in D} w(d-b)\kappa_{(b,d)}}{\sum_{(b,d)\in D} w(d-b)}.$$
(12)

³Frédéric Chazal and Bertrand Michel (2021). "An Introduction to Topological Data Analysis: Fundamental and Practical Aspects for Data Scientists". In: Frontiers in Artificial Intelligence 4. ISSN: 2624-8212.

⁴ Peter Bubenik (2015). "Statistical Topological Data Analysis Using Persistence Landscapes". In: Journal of Machine Learning Research 16.1, pp. 77–102

⁵Henry Adams et al. (2017). "Persistence Images: A Stable Vector Representation of Persistent Homology". In: The Journal of Machine Learning Research 18.1, pp. 218–252

Proposed approach: normalized functionals of persistence



Properties of signatures

1. Consistency: thanks to the additivity of persistence,

$$\bar{\rho}(D(\phi|_{[0,R]})) \xrightarrow{\|\cdot\|_{\mathcal{H}}} \bar{\rho}(D(\phi|_{[c,c+1]})), \quad \text{as } R \to \infty.$$
 (14)

- 2. Stability: when γ and W are random and independent, how does F(S) depend on the law of γ ?
- 3. Estimation: how to estimate the signature from a sampled time series?

Stability of the signature

A model for S

Let μ be a probability measure on $(\Gamma_{0,R,\nu_{\min}},\mathcal{B}(\|\cdot\|_{\infty}))$ for some $\nu_{\min}>0$, where

$$\Gamma_{0,R,v_{\min}} = \{ \gamma \in C([0,T],\mathbb{R}) \mid \gamma(0) = 0, \ \gamma(T) = R, \ \gamma(s) - \gamma(t) \ge v_{\min}(s-t), \ \text{ for all } s \ge t \}, \quad (15)$$

Let ν be a probability measure on $(C([0,T],\mathbb{R}),\mathcal{B}(\|\cdot\|_{\infty}))$, such that

$$||W||_{\infty} \le (\max \phi - \min \phi)/2 - \epsilon \text{ almost-surely,}$$
 (16)

$$t \mapsto W(t)$$
 has an α -Hölder version. (17)

Let $S := \phi \circ \gamma + W$, where $\gamma \sim \mu$ and $W \sim \nu$ are independent.

Theorem

If μ_1 , μ_2 are two probability measures on $\Gamma_{0,R,\nu_{min}}$ and $S_k = \phi \circ \gamma_k + W$, then

$$||F(S_1) - F(S_2)||_{\mathcal{H}} \le \frac{C}{v_{\min}^{\alpha}} \mathcal{W}_1(\mu_1, \mu_2)^{\alpha},$$
 (18)

where W_1 is the Wasserstein distance, and C depends on the regularity of W, $\|W\|_{\infty}$, ϵ , p and κ .

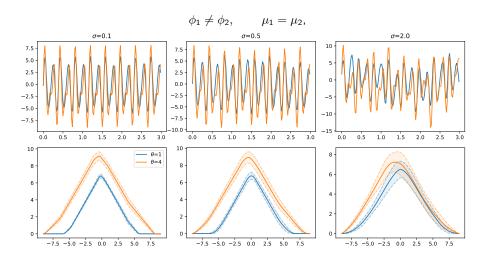
Comment

✓ As
$$c \to 0$$
, $||F(\phi \circ \gamma_1) - F(\phi \circ \gamma_2 + cW)||_{\mathcal{H}} \to 0$

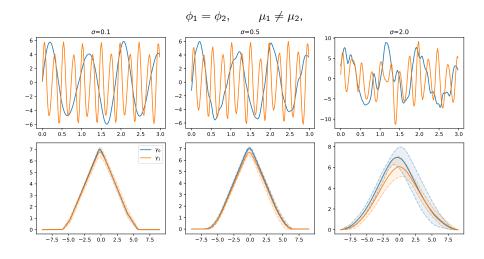
(continuity of $\bar{\rho}$).

X How to remove the fixed-endpoints assumption in (15)?

Numerical examples: stability



Numerical examples: stability



Estimation of signatures: introduction

Assume that only a single time series $(S_n)_{n=1}^N \subset \mathbb{R}$ is given,

$$S_n = \phi(\gamma(t_n)) + W(t_n).$$

Can we estimate the signature?

Proposition (Chazal et al. 2014⁶, Berry et al. 2018⁷)

Let be D_1, \ldots, D_N i.i.d.persistence diagrams. When the (bracketing) entropy of $(\bar{\rho}_x)_{x \in \mathbb{T}}$ is finite,

$$\sqrt{N}\left(\frac{1}{N}\sum_{n=1}^{N}\bar{\rho}(D_n)-\bar{\rho}^*\right)\stackrel{d}{\to}\mathbb{G},\tag{19}$$

for a zero-mean stochastic process G.

Procedure

We fix $M \in \mathbb{N}$ and we generate $S_1, \dots S_{N-M+1}$, where

$$\boldsymbol{S}_n = (S_n, \dots, S_{n+M-1}).$$

Challenge

 S_1, \ldots, S_{N-M+1} are not independent! Under what assumptions on $(\gamma(t_n))_{n\in\mathbb{N}}$ and $(W_n)_{n\in\mathbb{N}}$ does an analogue of (19) hold?

⁶Frédéric Chazal et al. (2014). "Stochastic Convergence of Persistence Landscapes and Silhouettes". In: Annual Symposium on Computational Geometry - SOCG'14. Kyoto, Japan: ACM Press, pp. 474–483. ISBN: 978-1-4503-2594-3. DOI: 10.1145/2582112.2582128

⁷Eric Berry et al. (2018). Functional Summaries of Persistence Diagrams. arXiv: 1804.01618

Quantifying dependence

Definition (β -mixing coefficients, Dedecker et al. 2007)

Let $(X_n)_{n\in\mathbb{Z}}$ be a stationary sequence of random variables on a common measurable space. Then,

$$\beta(k) := \sup_{A,B} \sum_{A \in A,B \in \mathcal{B}} |P(A \cap B) - P(A)P(B)|,$$

where $\mathcal{A} \subset \sigma^X_{-\infty,0}$, $\mathcal{B} \subset \sigma^X_{k,\infty}$ are finite partitions of the sample space and $\sigma^X_{a,b} \coloneqq \sigma((X_n)_{a \le n \le b})$.

Proposition (Kosorok 2008)

If $(\bar{\rho}_x)_{x\in\mathbb{T}}$ has finite bracketing entropy and $(S_n)_{n\in\mathbb{N}}$ is stationary with $\beta_S(k) = O(k^{-3})$, then

$$\sqrt{N}\left(\frac{1}{N}\sum_{n=1}^{N}\bar{\rho}(D_n)-\bar{\rho}^*\right) \xrightarrow{d} \mathbb{G}_{dep},\tag{20}$$

where \mathbb{G}_{dep} is a zero-mean stochastic process.

Proposition

For k > M + 1,

$$\beta_{\mathbf{S}}(k) \le \beta_{\mathbf{S}}(k - M + 1) \le \beta_{\operatorname{frac}(\gamma)}(k - M + 1) + \beta_{W}(k - M + 1), \tag{21}$$

where frac(x) := x - |x|.

Model for $(\gamma_n)_{n\in\mathbb{N}}$

Random walk model

For some h > 0, we set

$$\gamma_{n+1} = \gamma_n + hV_n,$$

for $(V_n)_{n\in\mathbb{N}}\sim \mathbf{P}$ i.i.d. We assume that

- \triangleright supp(**P**) \subseteq [V_{min} , V_{max}] \subset $]0, \infty[$.
- for some c > 0 and a non-trivial interval $I \subset [v_{\min}, v_{\max}]$

$$\mathbf{P}(A) \ge c\lambda(A), \quad \text{for all } A \in \mathcal{B}(I).$$
 (22)

Proposition

If $\gamma_0 \sim \mathcal{U}([0,1])$, then $(\operatorname{frac}(\gamma_n))_{n \in \mathbb{N}}$ is stationary and $\beta_{\operatorname{frac}(\gamma_n)}(k) = O(e^{-ak})$ for some a > 0.

Idea of the proof

1. By Thm 1 in Section 2.4 of Doukhan 1995, it suffices to show the **Doeblin condition**:

There is μ_0 and $n_0 \in \mathbb{N}$, such that for all $n > n_0$ uniformly in x_0 ,

$$P(\operatorname{frac}(\gamma_n) \in A \mid \gamma_0 = x_0) \ge \mu_0(A).$$

- 2. $\sum_{k=1}^{n} V_k \sim \mathbf{P}^{\star n}$, with $\mathbf{P}^{\star n}$ lower-bounded by a uniform measure with growing support.
- 3. For $n_0 \in \mathbb{N}$ big enough, the support is of length at least 1, and we obtain a lower-bound for the distribution of frac($\sum_{k=1}^{n_0} V_k$) on]0,1[.

Conclusion and future work

Conclusion

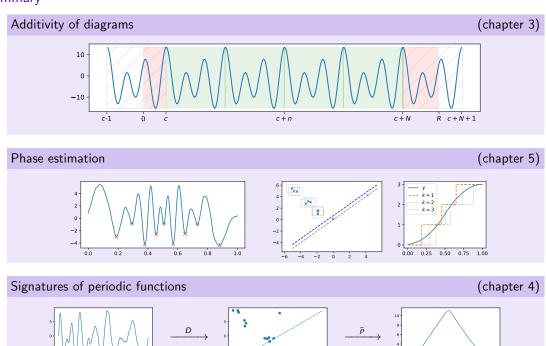
F(S) is a stable signature of ϕ and can be estimated with standard techniques.

Limitations and future work

- Remove the assumption of fixed endpoints from the stability
 - ► Technical difficulties in defining the probability measures
 - ▶ Understand the distance between $D(f|_{[0,T]})$ and $D(f|_{[0,t]}) \cup D(f|_{[t,T]})$
- Numerical experiments to understand the discriminative power
 - Compare with registration-based methods.
 - Understand how the choice of the kernel

0.2 0.4 0.6

Summary



-10.0 -7.5 -5.0 -2.5 0.0 2.5 5.0 7.5

-10.0 -7.5 -5.0 -2.5 0.0 2.5 5.0 7.5 10.0

Refernces I



Adams, Henry et al. (2017). "Persistence Images: A Stable Vector Representation of Persistent Homology". In: *The Journal of Machine Learning Research* 18.1, pp. 218–252.



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Thank you!

Persistence diagram of sub level sets: Definition

1. Persistence module

For each $t \in \mathbb{R}$, we calculate $H_0(X_t)$, where $X_t := f^{-1}(]-\infty,t]$). For any $s \leq t$, the inclusion $X_s \to X_t$ gives a map $\iota_s^t : H_0(X_s) \to H_0(X_t)$.

2. Rectangle measure

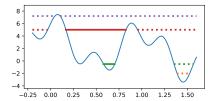
A measure m on rectangles of \mathbb{R}^2 .

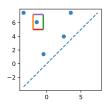
$$m(]a,b] \times [c,d[) = \dim \left(\frac{im(\iota_b^c) \cap \ker(\iota_c^d)}{im(\iota_o^c) \cap \ker(\iota_c^d)} \right),$$

3. Persistence diagram

The persistence diagram D(f) is a multi-set in \mathbb{R}^2 , where $(s,t) \in \mathbb{R}^2$ has multiplicity

$$m(s,t) = \lim_{s \to 0^+} m(]s - \delta, s + \delta] \times [t - \delta, t + \delta].$$





Proof of additivity of sub level sets: details

Proof.

Let $c := \inf\{x \in [0,1] | \phi(x) = \max \phi\}, \ N = \max\{n \in \mathbb{N} | c+n \le R\}$ and denote by $\mathbb{X}_t := \phi^{-1}([-\infty, t]).$

Step 1: For any t < M, $X_t \cap [0, c] \cap [c, c+1] = \emptyset$, so

$$H_0(\mathbb{X}_t \cap [0, R]) \simeq H_0(\mathbb{X}_t \cap [0, c]) \oplus H_0(\mathbb{X}_t \cap [c, c + N]) \oplus H_0(\mathbb{X}_t \cap [c + N, R]),$$
 (23)

Step 2: similarly,

$$H_0(\mathbb{X}_t \cap [c, c+N]) \simeq \bigsqcup_{n=1}^N H_0(\mathbb{X}_t \cap [c+(n-1), c+n])$$

$$(x \mapsto x+n) \quad \simeq \bigsqcup_{n=1}^N H_0(\mathbb{X}_t \cap [c, c+1])$$

$$(24)$$

$$(x \mapsto x + n) \simeq \bigsqcup_{n=1}^{N} H_0(\mathbb{X}_t \cap [c, c+1])$$
 (25)

Step 3: The inclusion $[0, c] \subset [c-1, c]$ induces an injective morphism

$$H_0(\mathbb{X}_t \cap [0,c]) \hookrightarrow H_0(\mathbb{X}_t \cap [c-1,c]).$$

Stability: bottleneck distance (detailed)

Definition (Edelsbrunner and Harer 2010, p. VIII.2)

We call a ϵ -matching between two persistence diagrams D and D' a bijection $\Gamma: A \to A'$ between some subsets of $A \subset D$ and $A' \subset D'$, considered with multiplicity, if

$$d_{\infty}(a, \Gamma(a)) \le \epsilon,$$
 for any $a \in A$,
 $d_{\infty}(a, \Delta) \le \epsilon,$ for any $a \in (D \setminus A) \cup (D' \setminus A').$

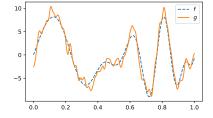
where $\Delta = \{(x, x) \in \mathbb{R}^2\}$ denotes the diagonal.

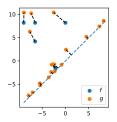
$$d_B(D, D') := \inf\{\epsilon > 0 \mid \Gamma \text{ is an } \epsilon\text{-matching between } D' \text{ and } D'\}.$$

Theorem (Bottleneck stability of diagrams)

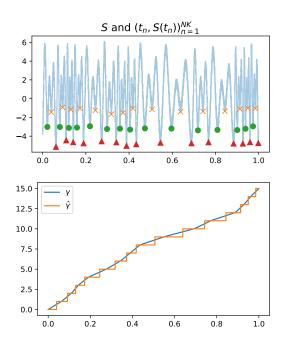
Let $f, g : \mathbb{X} \to \mathbb{R}$ be two continuous functions on a compact space \mathbb{X} . Then,

$$d_B(D(f),D(g)) \leq ||f-g||_{\infty}.$$



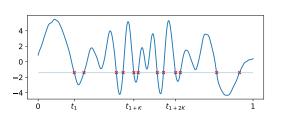


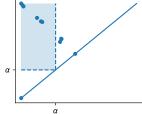
Landmarks for multiple periods



Zero-crossings from the persistence diagram

$$|S^{-1}(\alpha)| = 2 \lim_{\delta \to 0^+} |D(S) \cap (] - \infty, \alpha - \delta] \times [\alpha + \delta, \infty[)|.$$





Counting measure

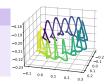
The persistence diagram D is also a counting measure on rectangles $A \subset \Delta_+ = \{(b,d) \in \mathbb{R}^2 \mid x < y\}$. By (4),

$$|D(\phi \circ \gamma) \cap A| = N |D(\phi|_{[0,1]}) \cap A|$$

Application: magnetic odometry and speed estimation

Problem

Using the magnetic signal **B**, recorded in a moving car, estimate the cars' trajectory. The angular position $t\mapsto \gamma(t)$ of a wheel in time is visible through $\mathbf{S}(t)=\mathbf{B}(\gamma(t),\gamma_h(t))$



Proposed solution

- 1. $S := \langle \mathbf{S}, v \rangle$, project **S** along a suitable direction $v \in \mathbb{S}^2$
- 2. $\hat{N}_{c,\tau}(S)$, for an appropriate scale τ ,
- 3. Derive an odometric sequence $t_1, \ldots, t_{\hat{N}_{\tau}(S)}$ from \mathcal{C}_{τ} .
- 4. Construct $\hat{\gamma}:[0,T]\to\mathbb{R}$.

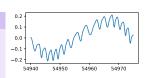
Results

	E_O		E_{l}	
Method	S_{ν_1}	$(abla {\sf S})_{ u_1}$	S_{v_1}	$(abla {\sf S})_{ u_1}$
$\hat{N}_{c, au}$	15.75	16.66	3.02	3.01
$\hat{N}_{0, au}$	15.75	16.66	3.02	3.01
ZC	9.91	5.62	6.35	16.51

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Results

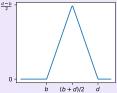
	E_O		E_{l}	
Method	S_{ν_1}	$(abla {\sf S})_{ u_1}$	S_{v_1}	$(abla {\sf S})_{ u_1}$
$\hat{ extsf{N}}_{c, au}$	15.75	16.66	3.02	3.01
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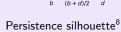
Normalized functionals of persistence

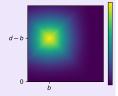
Functional representation

Let \mathcal{H} be a functional Banach space

- 1. $\operatorname{supp}(\kappa_{(b,d)}) \subset K$, K bounded,
- 2. $x \mapsto \kappa_{(b,d)}(x)$ (uniformly) Lipschitz,
- 3. $\|\kappa_{(b,d)} \kappa_{(b',d')}\|_{\mathcal{H}} \leq L_{\kappa} \|(b,d) (b',d')\|_{\kappa}$
- 4. $\|\kappa_{(b,b)}\|_{\mathcal{H}} \leq C$.







Persistence image⁹

Normalized functionals of persistence diagrams

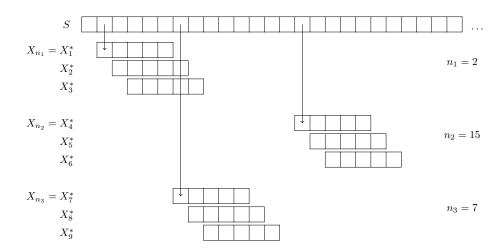
For some $p \ge 1$ and $\epsilon > 0$,

$$\bar{\rho}(D) := \frac{\sum_{(b,d)\in D} w(d-b)\kappa_{(b,d)}}{\sum_{(b,d)\in D} w(d-b)}, \quad \text{where } w(d-b) = \max(d-b-\epsilon,0)^p. \tag{26}$$

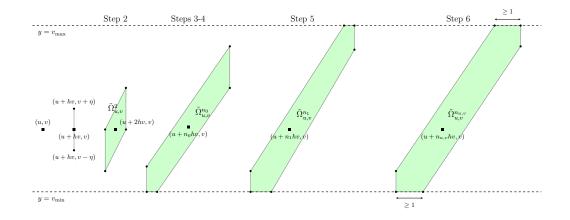
⁸Peter Bubenik (2015). "Statistical Topological Data Analysis Using Persistence Landscapes". In: Journal of Machine Learning Research 16.1, pp. 77–102

⁹Henry Adams et al. (2017). "Persistence Images: A Stable Vector Representation of Persistent Homology". In: The Journal of Machine Learning Research 18.1, pp. 218–252

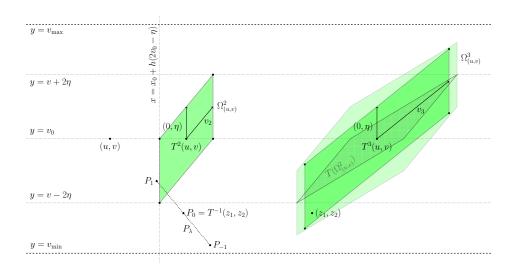
Bootstrap



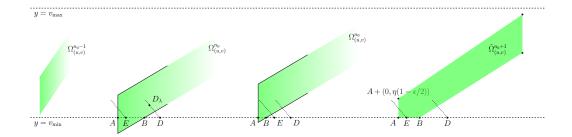
Mixing: overview



Mixing: initial



Mixing: boundary



Measures of dependence

Types of dependence

There are different ways to measure dependence in a time series $(X_n)_{n\in\mathbb{N}}\subset\mathbb{X}$:

- ► *m*-dependence.
- strong-mixing,
- weak-dependence,

Strong mixing

The β -mixing coefficient of a time series $(X_n)_{n\in\mathbb{N}}\subset\mathbb{X}$ is

$$\beta_X(k) = \sup_{A,B} \sum_{A \in A, B \in B} |P(A \cap B) - P(A)P(B)|,$$

where $A \subset \sigma^X_{-\infty,0}$, $B \subset \sigma^X_{k,\infty}$ are finite partitions of the sample space and $\sigma^X_{a,b} \coloneqq \sigma((X_n)_{a \le n \le b})$.

Example

- 1. If $(X_n)_n$ is *m*-dependent, then $\beta_X(k) = 0$ for $k \ge m$.
- 2. Markov chains: irreducible and aperiodic.

Proposition

For any measurable function $f : \mathbb{X} \to \mathbb{Y}$, $\beta_X(k) \leq \beta_{f(X)}(k)$.

Measures of dependence

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There are different ways to measure dependence in a time series $(X_n)_{n\in\mathbb{N}}\subset\mathbb{X}$:

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Strong mixing

The β -mixing coefficient of a time series $(X_n)_{n\in\mathbb{N}}\subset\mathbb{X}$ is

$$\beta_X(k) = \sup_{A,B} \sum_{A \in A, B \in B} |P(A \cap B) - P(A)P(B)|,$$

---> preserved by measurable functions!

where $A \subset \sigma^X_{-\infty,0}$, $B \subset \sigma^X_{k,\infty}$ are finite partitions of the sample space and $\sigma^X_{a,b} \coloneqq \sigma((X_n)_{a \le n \le b})$.

Example

- 1. If $(X_n)_n$ is *m*-dependent, then $\beta_X(k) = 0$ for $k \ge m$.
- 2. Markov chains: irreducible and aperiodic.

Proposition

For any measurable function $f : \mathbb{X} \to \mathbb{Y}$, $\beta_X(k) \leq \beta_{f(X)}(k)$.