

Chapter 1

MMFM Optimal Policy

In this section, the structure of the optimal policy for a machine which deteriorates according to the Markov Modulated Fluid Model will be determined. We will start with the simple model from earlier and will incrementally add complexity.

1.1 One CTMC-state, no jumps

The initial fluid level is given by $Q_0 \sim F(q)$ and the fluid rate in this single state equals r . The hazard rate then equals

$$\lim_{\delta \rightarrow 0} \frac{1}{\delta} \mathbb{P}(rx < Q_0 < r(x + \delta) | rx < Q_0) = \lim_{\delta \rightarrow 0} \frac{\mathbb{P}(rx < Q_0 < r(x + \delta))}{\delta \mathbb{P}(rx < Q_0)} \quad (1.1)$$
$$= rh(rx).$$

The rest of the problem can be solved in the same way as the simple problem. Hence, the optimal control limit μ^* is the first μ such that

$$rh(r\mu) = \beta \frac{c + V_\mu}{a} \quad (1.2)$$

(where V_μ is the expected total discounted cost when using control limit μ), or $\mu^* = \infty$ else.

1.2 HJB Equations and control limit for MMFM with constant jumps

In this section, the Hamilton-Jacobi-Bellman equations for the MMFM will be derived and these will be used to derive equations for the control limits. As earlier, we denote the state of the machine at time t by

$$X(t) = (S(t), L_0(t), L_c(t)).$$

The expected total discounted cost when starting in a state $X = (s_i, l_0, l_c)$ will be denoted by $V_i(l_0, l_c)$. The initial state is $X_0 = (s_0, 0, 0)$, transition intensities from state s_i to s_j are given by λ_{ij} and we will use $\lambda_i = \sum_j \lambda_{ij}$. When a jump

occurs between state s_i and s_j , the fluid level increases by $J_{ij} > 0$. When the machine is in state s_i , the fluid rate will decrease with rate $r_i > 0$.

From a state $X = (s_i, l_0, 0)$, in a small interval of length δ , a few things can happen:

- The machine can fail with probability $\delta r_i h(l_0) + o(\delta^2)$, in this case a cost $c + a$ is paid and the process is restarted.
- The machine can make a jump to a state s_j with probability $\delta \lambda_{ij} + o(\lambda^2)$, l_c is increased by J_{ij} in this case. As the probability of two jumps occurring in this time interval is $o(\delta^2)$, this possibility is neglected.
- The machine does not fail and no jump occurs with probability $1 - \delta(r_i h(l_0) + \lambda_{ij})$. In this case l_0 increases by δr_i .

Hence, for the expected cost of not repairing the machine, we have

$$\begin{aligned}
V_i(l_0, 0) &= \delta r_i h(l_0)(c + a + V_0(0, 0)) \\
&\quad + \delta \sum_{j \neq i} \lambda_{ij} V_j(l_0, J_{ij}) \\
&\quad + (1 - \delta(r_i h(l_0) + \lambda_i))e^{-\beta \delta} V_i(l_0 + \delta r_i, 0) + o(\delta^2) \\
&= \delta r_i h(l_0)(c + a + V_0(0, 0)) \\
&\quad + \delta \sum_{j \neq i} \lambda_{ij} V_j(l_0, J_{ij}) \\
&\quad + (1 - \delta(\beta + r_i h(l_0) + \lambda_i))V_i(l_0 + \delta r_i, 0) + o(\delta^2).
\end{aligned} \tag{1.3}$$

And preventive maintenance is chosen if this $V_i(l_0, 0)$ exceeds $c + V_0(0, 0)$. Suppose μ_i is the optimal control limit in this state. If $X = (s_i, \mu_i, 0)$ but we were to wait another time δ with maintenance, we would have

$$\begin{aligned}
c + V_0(0, 0) &\leq \delta r_i h(\mu_i)(c + a + V_0(0, 0)) \\
&\quad + \delta \sum_{j \neq i} \lambda_{ij} V_j(\mu_i, J_{ij}) \\
&\quad + (1 - \delta(\beta + r_i h(\mu_i) + \lambda_i))V_i(\mu_i + \delta r_i, 0) + o(\delta^2).
\end{aligned} \tag{1.4}$$

But we know that $V_i(\mu_i + \delta r_i, 0) = c + V_0(0, 0)$. So we can substitute that and subtract $c + V_0(0, 0)$ from both sides:

$$\begin{aligned}
0 &\leq \delta r_i h(\mu_i)(c + a + V_0(0, 0)) \\
&\quad + \delta \sum_{j \neq i} \lambda_{ij} V_j(\mu_i, J_{ij}) \\
&\quad + -\delta(\beta + r_i h(\mu_i) + \lambda_i)(c + V_0(0, 0)) + o(\delta^2).
\end{aligned} \tag{1.5}$$

Which can be rewritten to

$$r_i h(\mu_i) a + \sum_{j \neq i} \lambda_{ij} V_j(\mu_i, J_{ij}) \geq (\beta + \lambda_i)(c + V_0(0, 0)). \tag{1.6}$$

When we repeat this, but use $\mu_i - \delta$ as starting point instead of μ_i , we get

$$r_i h(\mu_i) a + \sum_{j \neq i} \lambda_{ij} V_j(\mu_i, J_{ij}) \leq (\beta + \lambda_i)(c + V_0(0, 0)). \tag{1.7}$$

Hence, μ_i should be chosen such that

$$r_i h(\mu_i) a + \sum_{j \neq i} \lambda_{ij} V_j(\mu_i, J_{ij}) = (\beta + \lambda_i)(c + V_0(0, 0)), \quad (1.8)$$

and maintenance will be chosen if $L_c(t) = 0$ and $L_0(t) \geq \mu_{S(t)}$. This intersection is difficult to find as the expected total discounted costs depend on the chosen control limits. Note that for $\lambda_{ij} = 0$, this equation simplifies to the earlier control limits. Also, the right hand side is a constant and the left hand side is increasing so that there cannot be multiple intersections. Before we are going to solve these equations, we are first going to solve them for a simpler problem.

1.3 Two equivalent CTMC-states with constant jumps at exponentially distributed time intervals

We simplify the model from the previous section by taking two CTMC-states with equal transition rates, fluid rates and jump sizes. The time intervals between jumps have rate λ and fluid quantity J . The fluid rate equals zero. As earlier, we denote the state of the machine at time t by

$$X(t) = (S(t), L_0(t), L_c(t)),$$

but in this simple two-state problem $S(t)$ isn't relevant since the two states are equivalent. When we fill in (1.8), we get

$$h(\mu)a + \lambda V(\mu, J) = (\beta + \lambda)(c + V(0, 0)) \quad (1.9)$$

Note that we will never preventively repair the machine when $L_c > 0$ since the machine cannot fail then. Hence, if we denote by $D(J)$ the expected discount over the time that the fluid has decreased by J , we can write $V(\mu, J) = D(J)V(\mu, 0)$. In the following section we will define this quantity and prove properties about it.

1.3.1 Time until control limit

Let $T_t(q)$ be the random variable denoting the time until the fluid level is q lower than it was at time t , i.e.

$$T_t(q) = \min\{\tau \geq 0 \mid Q(t + \tau) \leq Q(t) - q\}.$$

Note that, using this definition, $T_0(\mu)$ equals the time until the control limit is reached ($L_0(t) = \mu$) and $T(Q_0)$ equals the time until the machine fails.

Lemma 1. $Q_0 \leq \mu \Leftrightarrow T_t(Q_0) \leq T_t(\mu)$

Proof. \Rightarrow :

$$\begin{aligned} Q_0 \leq \mu &\Rightarrow Q(t) - \mu \leq Q(t) - Q_0 \\ &\Rightarrow (Q(t + \tau) \leq Q(t) - \mu \Rightarrow Q(t + \tau) \leq Q(t) - Q_0) \\ &\Rightarrow T_t(Q_0) \leq T_t(\mu) \end{aligned} \quad (1.10)$$

\Leftarrow : We will prove that $Q_0 > \mu \Rightarrow T_t(Q_0) > T_t(\mu)$:
We know that

$$Q_0 > \mu \Rightarrow Q(t) - \mu > Q(t) - Q_0.$$

Since $Q(t)$ is piecewise continuous and does not decrease at the discontinuities, we know that

$$Q(t + T_t(\mu)) = Q(t) - \mu > Q(t) - Q_0 \Rightarrow T_t(Q_0) > T_t(\mu).$$

□

To find the distribution of $T(q)$, we will condition on the number of jumps. Let $N_t(q)$ be the random variable denoting the number of jumps that occur in the interval $(t, t + T_t(q)]$. We will now compute its distribution. For simplicity, we will assume that $r = 1$. The probability that zero jumps occur equals the probability that the exponentially distributed time interval is larger than q :

$$\mathbb{P}(N_t(q) = 0) = e^{-\lambda q}.$$

The probability that exactly one jump occurs equals the probability that exactly one Poisson event happens in the interval $(t, t + q]$ while none happen in $(t + q, t + q + J]$. Resulting in

$$\mathbb{P}(N_t(q) = 1) = \lambda q e^{-\lambda q} e^{-\lambda J} = \lambda q e^{-\lambda(q+J)}.$$

For each $k \geq 0$, by conditioning on the time until the first jump, we get the following recursion

$$\begin{aligned} \mathbb{P}(N_t(q) = k + 1) &= \int_0^q \lambda e^{-\lambda x} \mathbb{P}(N_t(q - x + J) = k) dx \\ &= \int_0^q \lambda e^{-\lambda(q-y)} \mathbb{P}(N_t(y + J) = k) dy, \end{aligned} \tag{1.11}$$

since after this first jump, the fluid level equals $q - x + J$ and k jumps should occur. The second equality follows after the substitution $y = q - x$. The solution of this recursion, is given by

$$\mathbb{P}(N_t(q) = k) = \lambda q \frac{(\lambda(q + kJ))^{k-1}}{k!} e^{-\lambda(q+kJ)}.$$

Which can be seen by substituting it into (1.11) and setting $k = 1$ to see that it also satisfies the calculated expression for $\mathbb{P}(N_t(q) = 1)$.

Now we can define the quantity $D(q)$ in the following way

$$D(q) = \mathbb{E}[e^{-\beta T(q)}] = \sum_{k=0}^{\infty} e^{-\beta(q+kJ)} \mathbb{P}(N_t(q) = k).$$

This quantity has the following properties

Lemma 2. $D(A)D(B) = D(A + B)$

Proof.

$$\begin{aligned}
D(A)D(B) &= \left[\sum_{k=0}^{\infty} e^{-\beta(A+kJ)} \mathbb{P}(N_t(A) = k) \right] \left[\sum_{m=0}^{\infty} e^{-\beta(B+mJ)} \mathbb{P}(N_t(B) = m) \right] \\
&= \sum_{n=0}^{\infty} e^{-\beta(A+B+nJ)} \sum_{k=0}^n \mathbb{P}(N_t(A) = k) \mathbb{P}(N_t(B) = n-k) \\
&= \sum_{n=0}^{\infty} e^{-\beta(A+B+nJ)} \mathbb{P}(N_t(A+B) = n).
\end{aligned} \tag{1.12}$$

The last step holds since on the second last line, the second sum equals the probability that n jumps occur before a fluid quantity $A+B$ is drained, conditioned on the number of jumps that occur before a quantity A is drained. \square

Lemma 3.

$$\frac{d}{dA} D(A) = -(\beta + \lambda)D(A) + \lambda D(A+J)$$

Proof. TODO \square

1.3.2 Control limit

We can now rewrite (1.9) as

$$h(\mu)a + \lambda D(J)V(\mu, 0) = (\beta + \lambda)(c + V(0, 0)). \tag{1.13}$$

And since we will repair when control limit μ is reached, we get

$$\begin{aligned}
h(\mu)a + \lambda D(J)(c + V(0, 0)) &= (\beta + \lambda)(c + V(0, 0)) \\
\Rightarrow h(\mu) &= \frac{c + V(0, 0)}{a} [\beta + \lambda(1 - D(J))].
\end{aligned} \tag{1.14}$$

Note that for $\lambda = 0$ or $J = 0$ ($D(0) = 1$) this reduces to the control limit of the simple problem. Another nice thing to mention is that this control limit is equivalent to the simple problem with discount $\beta^* = \beta + \lambda(1 - D(J))$ and hence we could use the same iteration methods. However, given a control limit μ , $V(0, 0)$ is still difficult to compute.

1.3.3 Calculation of the expected total discounted cost

The presence of jumps complicates the calculation of the time between repairs. We can derive the distribution of the time length of a run with initial fluid level Q_0 random with density function $f_{Q_0}(q)$. This distribution $f(x)$, is given by

$$\begin{aligned}
f(x) &= \lim_{\delta \rightarrow 0} \frac{1}{\delta} \mathbb{P}(x < T_0(Q_0) < x + \delta) \\
&= \sum_{k=0}^{\lfloor \frac{x}{J} \rfloor} \mathbb{P}(N_0(q - kJ) = k) \lim_{\delta \rightarrow 0} \frac{1}{\delta} \mathbb{P}(x - kJ < Q_0 < x - kJ + \delta) \\
&= \sum_{k=0}^{\lfloor \frac{x}{J} \rfloor} \mathbb{P}(N_0(q - kJ) = k) f_{Q_0}(x - kJ)
\end{aligned} \tag{1.15}$$

The total discounted cost V^μ when using control limit μ can, similarly as with the simple problem, be calculated in the following way

$$V^\mu = \frac{(c+a)\mathbb{P}(Q_0 < \mu)\mathbb{E}[e^{-\beta Q_0} | Q_0 < \mu] + c\mathbb{P}(Q_0 \geq \mu)D(\mu)}{1 - \mathbb{E}e^{-\beta T(\mu \wedge Q_0)}}. \quad (1.16)$$

Alternatively, we could also find the optimal value for μ by differentiating the above equation by μ .

TODO

1.4 Multiple CTMC-states with different fluid rates

When there are states with different jump quantities, transition rates and fluid rates, a few complications arise:

1. The computation of $D(q)$ is more difficult as we need to take multiple paths with different probabilities and jump sizes into account. Moreover, it is also relevant from which state you start and where you end.
2. There are different control limits μ_i for different states, all depending on each other.
3. A jump can cause the machine to be repaired after the control limit.

We define

$$D_{ij}^t(q) := \mathbb{E}[e^{-\beta T(q)} \mathbb{1}\{S(t+T_t(q)) = s_j\} | S(t) = s_i]$$

and

$$D_i^t(q) := \sum_j D_{ij}^t(q).$$

Since these quantities don't really depend on t , this will be suppressed in the notation.

Now we are going to assume that the states are ordered such that

$$\mu_1 \leq \dots \leq \mu_N.$$

Ways to find such an ordering are discussed later. When $L_0(t) = \mu_N$, we know that in each state, preventive repair would be chosen such that we can rewrite the control limit (1.8) from

$$r_N h(\mu_N) a + \sum_{j \neq N} \lambda_{Nj} V_j(\mu_N, J_{Nj}) = (\beta + \lambda_N)(c + V_0(0, 0)) \quad (1.17)$$

to

$$\begin{aligned} r_N h(\mu_N) a + \sum_{j \neq N} \lambda_{Nj} D_j(J_{Nj})(c + V_0(0, 0)) &= (\beta + \lambda_N)(c + V_0(0, 0)) \\ \Rightarrow r_N h(\mu_N) &= \frac{c + V_0(0, 0)}{a} \left[\beta + \sum_{j \neq N} \lambda_{Nj} (1 - D_j(J_{Nj})) \right]. \end{aligned} \quad (1.18)$$

So that given the total discounted cost $V_0(0,0)$, we can find μ_N . This also provides a way to determine which state might be the state s_N with the highest control limit: simply compute for which state this equation results in the highest control limit.

When we want to repeat this for μ_{N-k} , the equation becomes more complex because it matters in which state you end after the jump quantity is drained:

$$\begin{aligned}
& r_{N-k}h(\mu_{N-k})a \\
& + \sum_{j \neq N-k} \lambda_{N-k,i} \left[\sum_{j \leq N-k} D_{ij}(J_{N-k,i})(c + V_0(0,0)) \right. \\
& \quad \left. + \sum_{j > N-k} D_{ij}(J_{N-k,i})V_j(\mu_{N-k},0) \right] \\
& = (\beta + \lambda_{N-k})(c + V_0(0,0)).
\end{aligned} \tag{1.19}$$

The most difficult part of solving the equations above, is that $V_j(\mu_{N-k},0)$ is not known. But we might find an expression by solving the following differential equations. $V_m(x,0)$ adhere to the following ODE for $x \in [\mu_{m-1}, \mu_m]$:

$$\begin{aligned}
r_m \frac{d}{dx} V_m(x,0) &= (\beta + r_m h(x) + \sum_{i \neq m} \lambda_{mi}(1 - D_{im}(J_{mi})))V_m(x,0) \\
&\quad - r_m h(x)(c + a + v_0(0,0)) \\
&\quad - \sum_{i \neq m} \lambda_{mi} \sum_{j > m} D_{ij}(J_{mi})V_j(x,0) \\
&\quad - \sum_{i \neq m} \lambda_{mi} \sum_{j < m} D_{ij}(J_{mi})(c + V_0(0,0))
\end{aligned} \tag{1.20}$$

Using this ODE, $V_N(x,0)$ can be computed when μ_N is known. When $V_N(x,0)$ is computed, μ_{N-1} can be computed and after that $V_{N-1}(x,0)$ can be computed. This can be repeated until $V_1(x,0)$. This suggests the following method for finding the policy $\pi = (\mu_1, \dots, \mu_N)$ for a given ordering s_1, \dots, s_N and an estimate of $V_0(0,0)$:

1. Compute μ_N using (1.18)
2. Compute $V_N(x,0)$ using the ODE (1.20).
3. Compute μ_{N-1} (1.19)
4. Compute $V_{N-1}(x,0)$ using the ODE (1.20).
5. ... repeat the above two steps for $N-2, \dots, 1$

Some complications might arise when a control limit exceeds a previously computed control limit (e.g. when $\mu_{N-1} > \mu_N$). Also, there is currently no method yet to compute $V_0^\pi(0,0)$ for some policy π . If we were to have a way to calculate $V_0^\pi(0,0)$, we could iterate the above method until the policy converges.

Before computing the control limit of the next state in the ordering, we could also determine the ordering ourselves by seeing for which of the remaining states, the control limit would be the highest and picking that state as the next state in the ordering.

Note that we do not yet have an easy way to compute $D_{jk}(J_{ij})$. Currently, the simplest way seems to start with all paths from s_j to s_k of length one and calculate the expected discount and probability of this path and repeat this for paths of longer length. Although the number of these paths grows exponential with their length, the discount might still converge rather quickly because of the exponential in the expectation. Furthermore these discounts can be pre-computed.