

$$\hat{\theta}_{MLE} = s(x_1, \dots, x_n) = \bar{x} \in \mathbb{R} \text{ e.g. } 0.1984$$

some function of  
the data  $x_1, \dots, x_n$

$\mathcal{F}: \text{iid Bern}(\theta)$

by property 2

$$\hat{\theta}_{MLE} = s(x_1, x_2, \dots, x_n) = \bar{x} \sim N(\theta, SE[\hat{\theta}_{MLE}]^2)$$

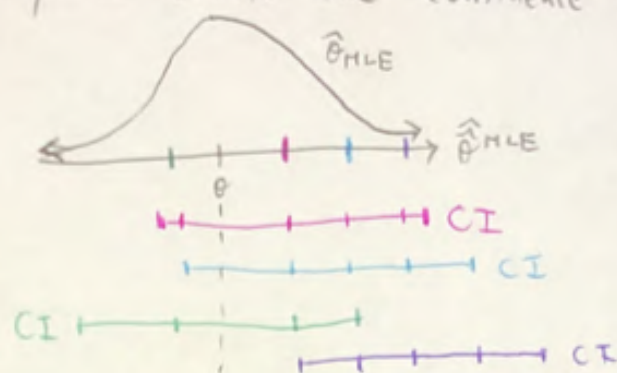
the same function  
except of the r.v.'s

A r.v., not a value,  
r.v.'s have distributions

We use the normality to create the confidence interval (CI). Why do CI's work?

$$\alpha = 5\% \quad 1 - \alpha = 95\%$$

$$z_{\frac{\alpha}{2}} = 1.96$$



CI's have 95% probability (before they're realized) of capturing (i.e. including) the real value of  $\theta$ .

Inference goal #3: "hypothesis testing" also called ~~theory testing~~  
"theory testing."

Consider a situation where \*I\* am trying to convince \*you\* of something.

Scenario I:

\*I\* declare

$H_a$ : Aliens and UFO's exist.

If they don't exist \*you\* need to provide \*me\* sufficient evidence that:

$H_0$ : Aliens and UFO's don't exist.

Scenario II:

\*I'll\* assume for the moment that

$H_0$ : Aliens and UFO's don't exist

and I will provide \*you\* sufficient evidence to the point that \*you're\* convinced that

$H_a$ : Aliens and UFO's exist.

Scenario II is more convincing and it is how science generally works. The theory I'm trying to demonstrate is called the "alternative hypothesis" ( $H_a$ ) since it's alternative to maybe business-as-usual. In Scenario II, you assume the opposite of the theory which is called the "null hypothesis" ( $H_0$ ). This is the "Hypothesis Testing" procedure.

In our context, theories are phrased as mathematical statements about  $\theta$ , the unknown parameter. We will study 3 types of  $H_a$ 's:

$$\begin{aligned} H_a: \theta \neq \theta_0 & \leftarrow \text{constant value} & \text{vs. } H_0: \theta = \theta_0 & \leftarrow \text{2-sided / 2 tailed test} \\ H_a: \theta > \theta_0 & & \text{vs. } H_0: \theta \leq \theta_0 & \leftarrow \text{Right-sided / Right tailed test} \\ H_a: \theta < \theta_0 & & \text{vs. } H_0: \theta \geq \theta_0 & \leftarrow \text{Left-sided / Left tailed test} \end{aligned}$$

There are two outcomes of a hypothesis test:

- (A) You were not shown sufficient evidence of  $H_a$ . Thus, you "fail to reject  $H_0$ " or "retain  $H_0$ "
- (B) You were indeed shown sufficient evidence of  $H_a$ . Thus, you "reject  $H_0$ " or "accept  $H_a$ "

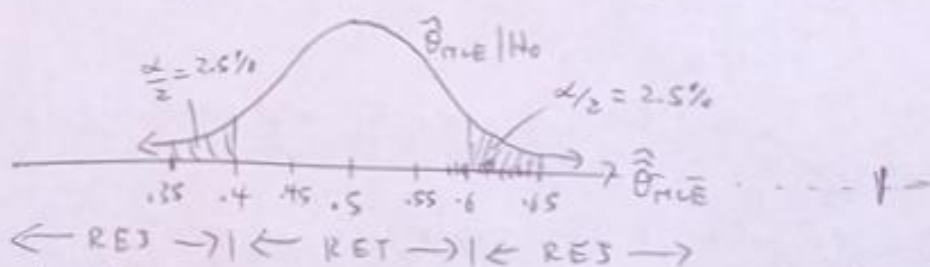
Imagine you're flipping a coin  $n=100$  times and you're counting the number of HEADS, then  $\hat{\theta} \sim \text{Bern}(\theta)$ . You want to prove the coin is unfair.

$$H_a: \theta \neq 0.5$$

$$H_0: \theta = 0.5$$

$$\text{If } H_0 \text{ is true, } \theta = 0.5 \Rightarrow \hat{\theta}_{MLE} \sim N\left(\theta, \text{SE}[\hat{\theta}_{MLE}]^2\right)$$

$\sqrt{\frac{0.5(1-0.5)}{n}} = 0.05$



What constitutes "sufficient evidence." It's a probability of rejecting when  $H_0$  is true (denoted  $\alpha(\alpha)$ ). Everyone is different.

If  $\alpha = 5\%$ , in a 2-tailed test, we put  $\frac{1}{2}$  the probability in each tail. 5% is the most common scientific standard. Thus, we retain  $H_0$  for the most non-weird 95% of the  $\hat{\theta}$ 's and reject  $H_0$  for the 5% most weird  $\hat{\theta}$ 's.

$$\text{RET} = [\theta_0 \pm z_{\frac{\alpha}{2}} \text{SE}[\hat{\theta}_{MLE}]] = [0.5 \pm 1.96 \sqrt{\frac{0.5(1-0.5)}{100}}] = [0.402, 0.598]$$

e.g. if  $\bar{x} = \frac{61}{100} = 0.61$   $\alpha = 5\%$

$\hat{\theta}_{MLE} = 0.61 \notin RET \Rightarrow \text{Rej } H_0 \text{ and conclude coin is unfair.}$

e.g. if  $\bar{x} = \frac{59}{100} = 0.59$

$\hat{\theta}_{MLE} = 0.59 \in RET \Rightarrow \text{fail to reject } H_0 \text{ and conclude there's not enough evidence of coin being unfair.}$

We've covered the "frequentist" approach to statistical inference. But there are problems with it...

①  $T = \text{iid Bern}(\theta), x = \langle 0, 0, 0 \rangle$

$$\hat{\theta}_{MLE} = \bar{x} = 0$$

Is that a good point estimate? NO. You shouldn't be able to say something is absolutely impossible after  $n=3$  trials.

$CI_{\theta, 1-\alpha} = \left[ 0 \pm 1.96 \sqrt{\frac{0(1-0)}{3}} \right] = \{0\} \rightarrow$  Is this a good confidence set? NO. This is not a good set of "reasonable values".

② What if you had prior knowledge that  $\theta$  was restricted to e.g.  $[0.1, 0.2]$  and not the full  $(0, 1)$ . You can't "enter that into" your inference.

③ Consider the frequentist interpretation of a CI:

(1) Before you do the experiment, you have a 95% probability of capturing  $\theta$ . But this doesn't tell you anything about after your experiment.

After your experiment you have an interval e.g.  $[0.37, 0.43]$  and you

can't say:  $P(\theta \in [0.37, 0.43]) = 0.95$   
no randomness!!

(2) 95% of CIs will cover  $\theta$ . But again, I only make one!!! So interpretation doesn't help me!

In conclusion, any specific CI means NOTHING

④ Hypothesis tests result in a binary outcome: either you reject  $H_0$  or you fail to reject  $H_0$ . What if you want to know  $P(H_0|x)$  or  $P(H_1|x)$ ?

You cannot!!! One thing you can do is:

$$P_{val} := \text{~~P(seeing } \hat{\theta} \text{ or more extreme | } H_0 \text{)}~~ \neq P(H_0|x)$$

⑤  $T = \text{iid Bern}(\theta), x = \langle 0, 1, 0 \rangle \Rightarrow \hat{\theta}_{MLE} = 0.33$

$$CI_{\theta, 95\%} = \left[ 0.33 \pm 1.96 \sqrt{\frac{0.33 \cdot 0.67}{3}} \right] = [-0.20, 0.87]$$

Is this a reasonable confidence set? No. It's outside of the legal parameter space which is  $(0, 1)$ .



$$\alpha = 5\%$$

$$H_0: \theta = 0.5 \Rightarrow \text{RET} = \left[ 0.5 \pm 1.96 \sqrt{\frac{0.5 \cdot 0.5}{n}} \right] = [-0.066, 1.066]$$

Is this a good hypothesis? No, because you NEVER reject!!!

The problem in #5 is because the asymptotic normality of the MLE doesn't "kick in" until  $n$  is large (MLE property 2 is not true yet).

We will solve all these problems with Bayesian Inference. Unfortunately, we will get other problems instead. It's a tradeoff and a personal decision.