## Intro Prob Lecture Notes

William Sun

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## Law of the Unconscious Statistician (For Continuous Random Variables)

• If X is a continuous random variable with pdf f(x) and  $g: \mathbb{R} \to \mathbb{R}$  is any function such that

$$\int_{-\infty}^{\infty} |g(x)| f(x) dx < \infty$$

(This is a condition which will guarantee that the expected value exists and is finite.) then

$$\mathbb{E}(g(X)) = \int_{-\infty}^{\infty} g(x)f(x)dx$$

- Example: Suppose  $X \sim \exp(1)$ 
  - $-f_X(x) = e^{-x}$  for x > 0,  $f_X(x) = 0$  otherwise
  - Let's compute  $\mathbb{E}(X)$  and  $\mathbb{E}(X^2)$ .
  - $\operatorname{Supp}(X) := \{x : F_X(x) > 0\}$

$$\mathbb{E}(X) = \int_{-\infty}^{\infty} x f_X(x) dx$$

$$= \int_{0}^{\infty} x e^{-x} dx$$

$$= x(-e^{-x}) \Big|_{0}^{\infty} - \int_{0}^{\infty} -e^{-x} dx$$

$$= \int_{0}^{\infty} e^{-x} dx$$

$$= 1.$$

(Remember integration by parts!)

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$$\mathbb{E}(X^2) = \int_0^\infty x^2 e^{-x} dx$$

$$= x^2 (-e^{-x}) \Big|_0^\infty - \int_0^\infty -e^{-x} \cdot 2x dx$$

$$= 2 \int_0^\infty x e^{-x} dx$$

$$= 2$$

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$$Var(X) = \mathbb{E}(X^2) - (\mathbb{E}(X))^2$$
 
$$= 2 - 1^2$$
 
$$= 1$$

• Interesting and useful way to compute expectations when the random variable is non-negative: (Generalized)

– If  $X \ge 0$  then

$$\mathbb{E}(X) = \int_{0}^{\infty} P(X > u) du$$

(proof later)

– For example, if  $X \sim \exp(1)$  then first:

$$P(X > u) = \int_{u}^{\infty} e^{-x} dx$$
$$= e^{-x} \Big|_{u}^{\infty}$$
$$= e^{-u}$$

- Then:

$$\mathbb{E}(X) = \int_{0}^{\infty} P(X > u) du = \mathbf{1} - \mathbf{F}_{\mathbf{X}}(\mathbf{u})$$
$$= \int_{0}^{\infty} e^{u} du$$
$$= 1$$

- And

$$\mathbb{E}(X^2) = \int_0^\infty P(X^2 > u) du$$

$$= \int_0^\infty P(X^2 > \sqrt{u}) du$$

$$= \int_0^\infty e^{-\sqrt{u}} du$$

$$= \int_0^\infty e^{-w} \cdot 2w dw$$

$$= 2$$

## The Normal Distribution (The Gaussian Distribution (In France, the Laplace Distribution))

•  $X \sim \text{Normal } (\mu, \sigma^2) \text{ if the pdf (for } -\infty < x < \infty) \text{ is}$ 

$$f(x) = \frac{e^{-\frac{1}{2}(\frac{x-\mu}{\sigma})^2}}{\sigma\sqrt{2\pi}}$$

- (Trivially)  $\mu$  is the average/ $\mathbb{E}$ ,  $\sigma$  is the standard deviation /  $\sqrt{Var}$
- If p is fixed, then as  $n \to \infty$ , Binomial $(n, p) \to Normal$ 
  - Also Poisson( $\lambda$ ) as  $\lambda \to \infty$
  - Central Limit Theorem
- Totals 1

$$\int_{-\infty}^{\infty} \frac{e^{-\frac{1}{2}(\frac{x-\mu}{\sigma})^2}}{\sigma\sqrt{2\pi}} dx = 1 \to \frac{e^{-\frac{u^2}{2}}}{\sqrt{2\pi}} = 1$$

 $I^{2} = \int \frac{e^{-\frac{u^{2}}{2}}}{\sigma\sqrt{2\pi}} du \int \frac{e^{-\frac{v^{2}}{2}}}{\sigma\sqrt{2\pi}} dv = \int \int \frac{e^{-\frac{u^{2}+v^{2}}{2}}}{\sigma\sqrt{2\pi}} du dv = \dots = 1$ 

(See Text)

- Theorem: If  $X \sim \text{Normal}(\mu, \sigma^2)$  and a, b are any constants, then
  - $-Y = aX + b \sim \text{Normal}(a\mu + b, a^2\sigma^2)$
  - If you have a normal random variable, any linear transformation on it is also a normal random variable.
  - Consequence:  $X \sim \text{Normal}(\mu, \sigma^2)$ , then  $\frac{X-\mu}{\sigma} \sim \text{Normal}(0, 1)$ 
    - \* Any normal random variable can be converted into a *standard normal distribution* with a mean of 0 and a standard deviation of 1
  - Proof:

$$F_Y(y) = P(Y \le y) = P(aX + b \le y)$$

$$= P(X \le \frac{y - b}{a}) \text{(Assuming } a > 0)$$

$$= F_X(\frac{y - b}{a})$$

$$f_Y(y) = \frac{d}{dy} (F_X(\frac{y-b}{a}))$$

$$= f_X(\frac{y-b}{a}) \cdot \frac{1}{a}$$

$$= \frac{1}{a} (\frac{e^{-\frac{1}{2}(\frac{y-b}{a}-u})^2}}{\sigma\sqrt{2\pi}})$$

$$= \frac{e^{-\frac{1}{2}(\frac{y-b}{a}-u})^2}}{a\sigma\sqrt{2\pi}}$$