

Intro Prob Lecture Notes

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Conditional Expectation

- Law of total expectation $E(X) = E(E(X|Y))$
 - Example: random variable N $p(N = n) = p_n$ for $n = 0, 1, 2, 3, \dots$ x_1, x_2, x_3, \dots i.i.d. and independent of N

$$* \rightarrow E(S) = \mu_x \mu_N$$

$$\begin{aligned} - \text{ } Var(S) &= E(S^2) - (E(S))^2 = E(S^2) - \mu_x^2 \mu_N^2 \\ * \left(\sum_{i=1}^N x_i \right) \left(\sum_{j=1}^N x_j \right) &= \sum_{i=1}^N \sum_{j=1}^N x_i x_j = \sum_{i=1}^N x_i^2 \sum_{i \neq j}^N x_i x_j \end{aligned}$$

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$$\begin{aligned} E(S^2) &= E(E(S^2|N)) \\ &= E\left(\sum_{i=1}^N x_i^2 \sum_{i \neq j}^N x_i x_j | N = n\right) \\ &= E\left(\sum_{i=1}^N x_i^2 \sum_{i \neq j}^N x_i x_j\right) \\ &= nE(x_1^2) + n(n-1)E(x_1 x_2) \\ &= n(\sigma_X^2 + \mu_X^2) + n(n-1)E(x_1)E(x_2) \\ &= n\sigma_X^2 + n\mu_X^2 + n(n-1)\mu_x^2 \\ &= n\sigma_X^2 + n^2\mu_X^2 \end{aligned}$$

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$$\begin{aligned}
E(S^2) &= E(E(S^2|N)) \\
&= E(n\sigma_X^2 + n^2\mu_X^2) \\
&= \sum_{n=0}^{\infty} (n\sigma_X^2 + n^2\mu_X^2)P_N(n) &= \sigma_X^2 \sum_{n=0}^{\infty} nP_N(n) + \mu_X^2 \sum_{n=0}^{\infty} n^2P_N(n) \\
&= \sigma_X^2\mu_N + \sigma_X^2E(N^2)
\end{aligned}$$

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$$Var(S) = \sigma_X^2\mu_N + \mu_X^2E(N^2) - \mu_X^2\mu_N^2 = \sigma^2\mu_N + \mu_X^2\sigma_N^2$$

Function with another variable

- Let X, Y r.v.'s and h is any function. Then

$$E(Xh(Y)|Y) = h(Y)E(X|Y)$$

or

$$E(Xh(Y)|Y = y) = h(y)E(X|Y = y)$$

- Proof:
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$$\begin{aligned}
E(Xh(Y)|Y = y) &= \int_{-\infty}^{\infty} xh(y)f_{X|Y}(x, y)dx \\
&= h(y) \int_{-\infty}^{\infty} xf_{X|Y}(x|y)dx \\
&= h(y)E(X|Y = y)
\end{aligned}$$

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$$\begin{aligned} E(g(X)h(Y)) &= E(E(g(X)h(Y)|Y)) \\ &= E(h(Y)E(g(X)|Y)) \end{aligned}$$

- Ex: $X \sim \text{geometric}(p)$

- Let Y be the outcome on the first toss. $f_Y(y) = \begin{cases} p & y = 1 \\ 1 - p & y = 0 \end{cases}$
- $E(X) = E(X|Y = 0) + E(X|Y = 1) = (1 + E(X))(1 - p) + 1(p) = 1 \cdot p + (1 - p)(1 + \mu) = \frac{1}{p}$
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$$\text{Var}(X) = E(X^2) - (E(X))^2$$

$$\begin{aligned} E(X^2) &= E(E(X^2|Y)) \\ &= E(X^2Y = 1)p + E(X^2|Y = 0)(1 - p) \\ &= 1 \cdot p + E((1 + X)^2)(1 - p) \\ &= p + E(1 + 2X + X^2)(1 - p) \\ &= p + (1 - p) + (1 - p)\frac{2}{p} + E(X^2)(1 - p) \\ pE(X^2) &= 1 + \frac{2(1 - p)}{p} \\ E(X^2) &= \frac{1}{p} + \frac{2(1 - p)}{p^2} \end{aligned}$$

* So

$$\text{Var}(X) = \frac{1}{p} + \frac{2(1 - p)}{p^2} - \left(\frac{1}{p}\right)^2 = \frac{1 - p}{p^2}$$

Conditional Variance

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$$\text{Var}(X|Y = y) = E([X - E(X|Y = y)]^2|Y = y)$$

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$$f(x, y) = xe^{-x(1+y)} \text{ for } x > 0, y > 0$$

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$$f_Y(y) = \frac{1}{(1+y)^2} \text{ for } y > 0$$

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$$f_{X|Y}(x|y) = (1+y)^2 xe^{-x(1+y)} \text{ for } x > 0$$

- $\text{Gamma}(2, \frac{1}{1+y})$

- $X|Y = y \sim \text{Gamma}(\alpha, \beta)$ with $\alpha = 2, \beta = \frac{1}{1+y}$

- Means $= \alpha\beta = \frac{2}{1+y}$

- Variance $= \alpha\beta^2 = \frac{2}{(1+y)^2}$

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$$\begin{aligned} \text{Var}(X|Y = y) &= E([X - E(X|Y = y)]^2 | Y = y) \\ &= E\left(\left[X - \frac{2}{1+y}\right]^2 | Y = y\right) \\ &= E\left(X^2 - \frac{4}{1+y}X + \frac{4}{(1+y)^2} | Y = y\right) \\ &= \int_0^\infty x^2 f_{X|Y}(x, y) dx - \frac{4}{1+y} \int_0^\infty x f_{X|Y}(x, y) dx + \frac{4}{(1+y)^2} \\ &= (1+y)^2 \int_0^\infty x^2 x e^{-x(1+y)} dx - \frac{4}{1+y} \frac{2}{1+y} + \frac{4}{(1+y)^2} \\ &= \frac{(1+y)^2 \cdot 3!}{(1+y)^4} - \frac{8}{(1+y)^2} + \frac{4}{(1+y)^2} \\ &= \frac{2}{(1+y)^2} \end{aligned}$$

- Example: Z_1, Z_2 independent standard normals.

- $X = \mu_X + \sigma_X Z_1$

- $Y = \mu_Y + \sigma_Y \rho Z_1 + \sigma_X(?) \sqrt{1 - \rho^2} Z_2$

- $Z_1 = \frac{X - \mu_X}{\sigma_X}$

$$\begin{aligned}
& - Y = \mu_Y + \sigma_Y \rho \left(\frac{X - \mu_X}{\sigma_X} \right) + \sigma_Y \sqrt{1 - \rho^2} Z_2 \\
& - Y|X \sim \text{Normal}(\mu_Y + \frac{\sigma_Y}{\sigma_X} \rho (x - \mu_X), \sigma_Y^2 (1 - \rho^2)) \\
& - \text{Then:} \\
& - E(Y|X) = \mu_Y + \frac{\sigma_Y \rho (X - \mu_X)}{\sigma_X^2 (1 - \rho^2)} \\
& - \text{Var}(Y|X) = \sigma_Y^2 (1 - \rho^2) \\
& - E(Y|X) = \mu_Y - \frac{\sigma_Y \mu_X \rho}{\sigma_X} + \frac{\sigma_Y (?) \rho}{\sigma_X} X \\
& - \text{Var}(E(Y|X)) = \frac{\sigma_Y^2 \rho^2}{\sigma_X^2} \text{Var}(X) = \sigma_Y^2 \rho^2 \\
& * \implies \rho^2 = \frac{\text{Var}(E(Y|X))}{\text{Var}(Y)}
\end{aligned}$$

Moment Generating Functions

- For a random variable X , $M_X(s) = E(e^{sX})$ (when this function is finite/exists in an open neighborhood of $s = 0$)

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$$\begin{aligned}
e^y &= 1 + y + \frac{y^2}{2!} + \frac{y^3}{3!} + \dots \\
e^{sX} &= 1 + sX + \frac{s^2 X^2}{2!} + \frac{s^3 X^3}{3!} + \dots \\
E(e^{sX}) &= 1 + E(X)s + E(X^2) \frac{s^2}{2!} + \dots \\
&= \sum_{n=0}^{\infty} E(X^n) \frac{s^n}{n!} = M_X(s)
\end{aligned}$$

- Suppose $X \sim \text{uniform}(0, 1)$. Compute $M_X(s)$

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$$\begin{aligned}
 M_X(s) &= E(e^{sX}) \\
 &= \int_0^1 e^{sX} 1 dx \\
 &= \left. \frac{e^{sX}}{s} \right|_{x=0}^1 \\
 &= \frac{e^s - 1}{s}
 \end{aligned}$$

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$$\begin{aligned}
 e^s &= 1 + s + \frac{s^2}{2!} + \frac{s^3}{3!} + \dots \\
 e^s - 1 &= s + \frac{s^2}{2!} + \frac{s^3}{3!} + \dots \\
 \frac{e^s - 1}{s} &= 1 + \frac{s}{2!} + \frac{s^2}{3!} + \dots + \frac{s^{n-1}}{n!} + \frac{s^{n(?)}}{(n+1)!} \\
 &= 1 + \left(\frac{1}{2}\right)s + \left(\frac{1}{3}\right)\frac{s^2}{2} + \dots + \left(\frac{1}{n}\right)\frac{s^{n-1}}{(n-1)!} + \left(\frac{1}{n+1}\right)\frac{s^n}{n!} + \dots \\
 E(X^n) &= \frac{1}{n+1}
 \end{aligned}$$

• Application of mgf:

$$\begin{aligned}
 - M'_X(s) &= \frac{d}{ds}(M_X(s)) = \frac{d}{ds}E(e^{sX}) = E\left(\frac{d}{ds}(e^{sX})\right) = E(Xe^{sX}) \\
 - \mathbf{M}'_{\mathbf{X}}(\mathbf{0}) &= \mathbf{E}(\mathbf{X}) \\
 - \frac{d^n}{ds^n} M'_X(s) &= E(X^n e^{sX}) \\
 - \frac{d^n}{ds^n} (\mathbf{M}_{\mathbf{X}}(\mathbf{0})) &= \mathbf{E}(\mathbf{X}^n)
 \end{aligned}$$