

# Intro Prob Lecture Notes

William Sun

April 3, 2017

## Independence of Random Variables

- If  $X_1, X_2, \dots, X_n$  are a collection of jointly distributed random variables, then we say that they are (mutually) independent if their joint distribution factors as the product of the individual marginal distributions

- For example, if  $X_1, X_2, \dots, X_n$  are jointly discrete then they will be independent if

\*

$$P_{X_1, X_2, \dots, X_n}(x_1, x_2, \dots, x_n) = P_{X_1}(x_1)P_{X_2}(x_2) \dots P_{X_n}(x_n)$$

\* for *all* possible values of  $x_1, x_2, \dots, x_n$

- Also, if  $X_1, X_2, \dots, X_n$  are jointly continuous then they will be independent if

\*

$$f(x_1, x_2, \dots, x_n) = f_{X_1}(x_1)f_{X_2}(x_2) \dots f_{X_n}(x_n)$$

\* for all  $x_1, x_2, \dots, x_n$

- Example: Dart board

- $f(x_1, x_2) = \frac{1}{\pi}$  if  $x_1^2 + x_2^2 \leq 1$ , 0 otherwise.

- Compute the marginals

\* If  $x < -1$  or  $x > 1$ , marginal is 0. Otherwise:

\*

$$f_{X_1}(x_1) = \int_{-\infty}^{\infty} f(x_1, x_2) dx_2 = \int_{-\sqrt{1-x_1^2}}^{+\sqrt{1-x_1^2}} \frac{1}{\pi} dx_2 = \frac{2\sqrt{1-x_1^2}}{\pi}$$

\* Similarly, when in range,

$$f_{X_2}(x_2) = \frac{2\sqrt{1-x_2^2}}{\pi}$$

– Multiply the marginals

$$* f\left(\frac{1}{2}, \frac{1}{2}\right) = \frac{1}{\pi}$$

\*

$$f_{X_1}\left(\frac{1}{2}\right)f_{X_2}\left(\frac{1}{2}\right) = \frac{2\sqrt{1-\frac{1}{4}}}{\pi} \cdot \frac{2\sqrt{1-\frac{1}{4}}}{\pi} = \frac{3}{\pi^2}$$

\*  $\frac{1}{\pi} \neq \frac{3}{\pi^2}$ , so  $X_1, X_2$  are *dependent*.

• Usually independence is an assumption that is made

–  $X_1, \dots, X_n$  are independent  $N(\mu, \sigma^2)$

–

$$\bar{X}_n = \frac{X_1 + \dots + X_n}{n}$$

–

$$S^2 = \frac{\sum_{i=1}^n (X_i - \bar{X}_n)^2}{n-1}$$

– are, in fact, independent (to be proved later)

## Sums of random variables

• If  $X_1, X_2$  are jointly distributed random variables, then what is the distribution (pmf) of  $X_1 + X_2$ ?

• Case 1: Suppose  $X_1, X_2$  jointly discrete,  $P_{X_1, X_2}(x_1, x_2) = P(X_1 = x_1, X_2 = x_2)$  given

$$- P_{X_1, X_2}(u) = P(X_1 + X_2 = u) = \sum_{x_1} P(X_1 + X_2, X_1 = x_1) = \sum_{x_1} P(X_1 = x_1, X_2 = u - x) = \sum_{x_1} P_{X_1, X_2}(x_1, u - x_1)$$

\* (Law of total probability)

– Formula:

$$P_{X_1+X_2}(u) = \sum_{x_1} P_{X_1, X_2}(x_1, u - x_1)$$

– A common assumption is that  $X_1, X_2$  independent. In this case:

\*

$$P_{X_1+X_2}(u) = \sum_{x_1} P_{X_1}(x_1)P_{X_2}(u - x_1)$$

\* Convolution  $(P_{X_1} * P_{X_2})(u)$

- Binomial theorem  $(a+b)^n = \sum_{k=0}^n \binom{n}{k} a^k b^{n-k}$

—

**Example:**  $X_1 \sim \text{Poisson}(\lambda_1)$ ,  $X_2 \sim \text{Poisson}(\lambda_2)$ , and they are independent

$$\begin{aligned} P_{X_1+X_2}(u) &= \sum_{x_1=0}^{\infty} \frac{e^{-\lambda_1} \lambda_1^{x_1}}{x_1!} \cdot \frac{e^{-\lambda_2} \lambda_2^{u-x_1}}{(u-x_1)!} \\ &= \frac{e^{-(\lambda_1+\lambda_2)}}{u!} \sum_{x_1=0}^u \frac{u!}{x_1!(u-x_1)!} \cdot \lambda_1^{x_1} \lambda_2^{u-x_1} \\ &= e^{-(\lambda_1+\lambda_2)} \frac{(\lambda_1 + \lambda_2)^u}{u!} \end{aligned}$$

\* The pdf of a  $\text{Poisson}(\lambda_1 + \lambda_2)$

• Note: For formula to be valid, upper limit must be  $\infty$

•

**Suppose**  $X_1 + X_2 = n$ .  $P(X_1 = k | X_1 + X_2 = n) = ?$

$$\begin{aligned} P(X_1 = k | X_1 + X_2 = n) &= \frac{P(X_1 = k, X_1 + X_2 = n)}{e^{-(\lambda_1+\lambda_2)} \frac{(\lambda_1+\lambda_2)^n}{n!}} \\ &= \frac{e^{-\lambda_1} \frac{\lambda_1^k}{k!} \cdot e^{-\lambda_2} \frac{\lambda_2^{n-k}}{(n-k)!}}{e^{-(\lambda_1+\lambda_2)} \frac{(\lambda_1+\lambda_2)^n}{n!}} \end{aligned}$$