

# 可靠度資料分析

# Reliability Data Analysis

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**Lecture 4 – Common Failure Distribution Models**

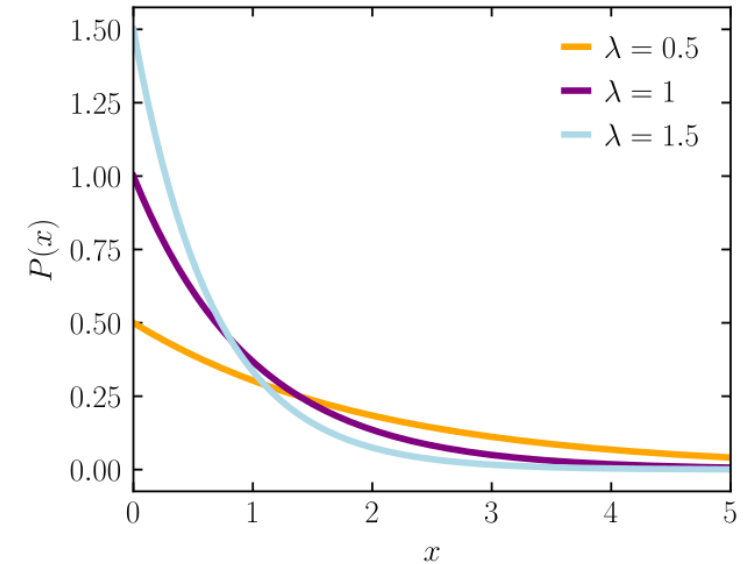
# Exponential Distribution

- Constant failure rate (CFR),  $\lambda$
- Assume that  $\lambda(t) = \lambda$ ,  $t \geq 0$ , and  $\lambda > 0$ .
- Thus, we have the following:

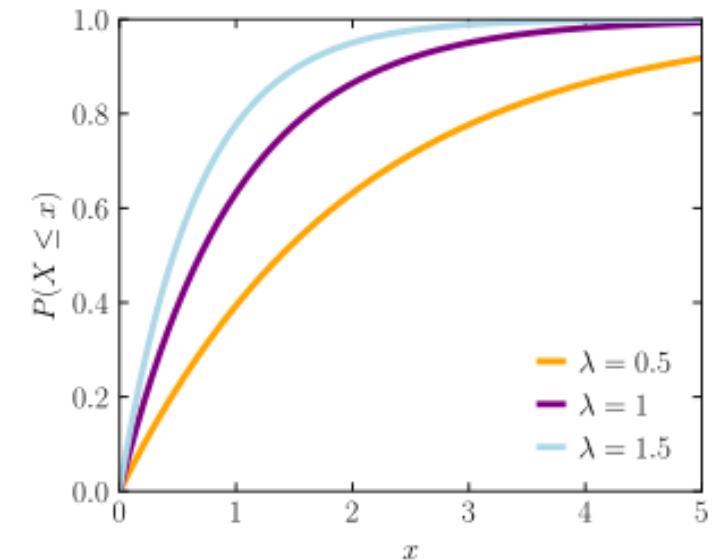
$$\begin{aligned}f(t) &= \lambda e^{-\lambda t} \\F(t) &= 1 - e^{-\lambda t} \\R(t) &= e^{-\lambda t}, \quad t \geq 0\end{aligned}$$

- $MTTF = \frac{1}{\lambda}$  and  $\sigma^2 = \frac{1}{\lambda^2}$

Probability density function



Cumulative distribution function

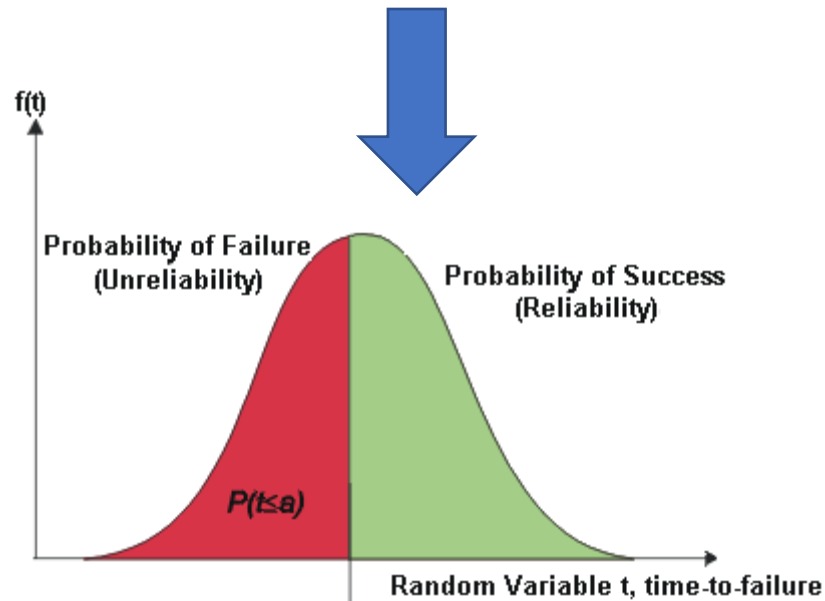


# Quick Note:

## Two ways of looking at Reliability Function

$$\bullet R(t) = \int_t^{\infty} f(t') dt' = 1 - \int_0^t f(t') dt' = \exp\left[-\int_0^t \lambda dt'\right], \quad t \geq 0$$

This is by definition of the PDF curve



This is by looking at hazard rate or instantaneous failure rate point of view

See next slide or “Hazard Rate” from Lecture 3

# FYI: Derivation of the Reliability Function, R(t)

$R(t)$  = probability of success  
 $N_S$  = number successful  
 $N_F$  = number failed  
 $N_O$  = total number =  $N_S + N_F$  = constant  
 $\lambda$  = failure rate

reliability is probability of success

$$R(t) = \frac{N_S}{N_O} = \frac{N_O - N_F}{N_O} = 1 - \frac{N_F}{N_O} \Rightarrow$$

$$\frac{dR}{dt} = \frac{d(1 - \frac{N_F}{N_O})}{dt} = -\frac{1}{N_O} \frac{dN_F}{dt} \Rightarrow -N_O \frac{dR}{dt} = \frac{dN_F}{dt} \Rightarrow -\frac{N_O}{N_S} \frac{dR}{dt} = \frac{1}{N_S} \frac{dN_F}{dt} = \lambda \Rightarrow$$

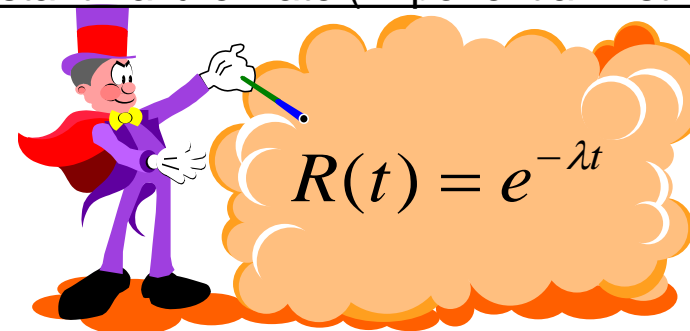
(up to here is  
derivation of  $\lambda$ )

$$-\frac{1}{R} \frac{dR}{dt} = \lambda \Rightarrow -\frac{dR}{R} = \lambda dt \Rightarrow -\int_1^R \frac{dR}{R} = -\ln R = \int_0^t \lambda dt \Rightarrow$$

General Case

$$R(t) = e^{-\int_0^t \lambda dt} \Rightarrow$$

Constant Failure Rate (Exponential Distribution) Case



# Memoryless Property

- Exponential distribution has memoryless property.
- The memoryless property means that the time to failure of a component is not dependent on how long the component has been operating
  - Probability of failure is constant over time, independent of past failures
  - Probability of failure is the same at any point in time
- To see this, consider the following conditional probability:

$$R(t|T_0) = \frac{R(t + T_0)}{R(T_0)} = \frac{e^{-\lambda(t+T_0)}}{e^{-\lambda T_0}} = \frac{e^{-\lambda t} \cdot e^{-\lambda T_0}}{e^{-\lambda T_0}} = e^{-\lambda t} = R(t)$$

# Exponential Distribution

- Useful for maintenance engineering
- To predict the amount of waiting time until the next event (i.e., success, failure, arrival, etc.)
  - The amount of time until the customer finishes browsing and actually purchases something in your store (success).
  - The amount of time until the hardware fails (failure).
  - The amount of time you need to wait until the bus arrives (arrival).

# Failure Modes

- Complex systems will fail through various means resulting from different physical characteristics.
- These failures can be separated according to the mechanisms or components causing the failures. These categories of failures are referred to as failure modes.
- If we define  $R_i(t)$  as the reliability of the  $i$ th failure mode, i.e., the probability that the  $i^{th}$  failure mode does not occur before time  $t$ . Then assuming independence among the failure modes, the system reliability denoted by  $R(t)$  can be expressed as follows:

$$R(t) = \prod_{i=1}^n R_i(t)$$

- That is, none of the  $n$  failure modes occurs before time  $t$ .

# System Hazard Rate for Exponential Reliability Function

- Let  $\lambda_i(t)$  denote the failure rate function for the  $i^{th}$  failure mode. Then we have  $R_i(t) = \exp[-\int_0^t \lambda_i(t')dt']$ , thus the system reliability is

$$\begin{aligned} R(t) &= \prod_{i=1}^n \exp \left[ - \int_0^t \lambda_i(t') dt' \right] \\ &= \exp \left[ - \int_0^t \sum_{i=1}^n \lambda_i(t') dt' \right] \\ &= \exp \left[ - \int_0^t \lambda(t') dt' \right] = \exp[-\lambda t] \end{aligned}$$

where  $\lambda(t') = \sum_{i=1}^n \lambda_i(t')$

- $MTTF = \frac{1}{\lambda} = \frac{1}{\sum_{i=1}^n \lambda_i} = \frac{1}{\sum_{i=1}^n 1/MTTF_i}$ ;  $MTTF_i = \frac{1}{\lambda_i}$



# Example

- An aircraft engine consists of three modules having constant failure rates of  $\lambda_1 = 0.002$ ,  $\lambda_2 = 0.015$ , and  $\lambda_3 = 0.0025$  failures per operating hour. Evaluate the reliability function for the engine and calculate the corresponding MTTF.

# Weibull Distribution

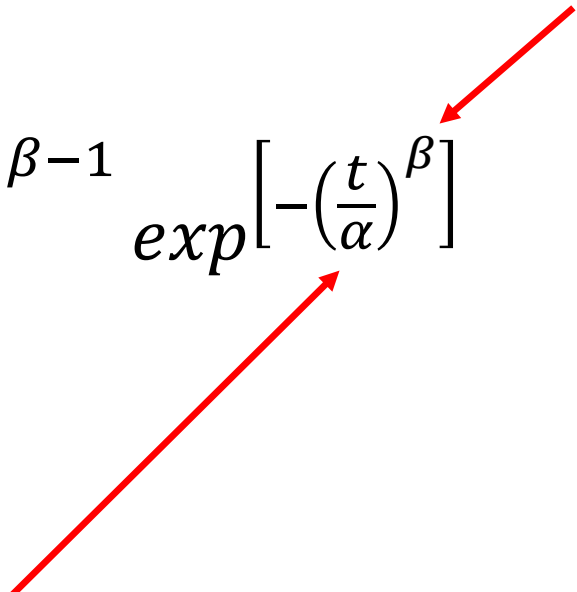
- Very flexible distribution
- Applicable to all three phases of life cycle
- Exponential is a special case of Weibull
- Shape parameter:  $\beta$
- Characteristic life or scale parameter
  - $\alpha$
  - $\eta$

# Weibull Distribution

$$f(t) = \left(\frac{\beta}{\alpha}\right) \left(\frac{t}{\alpha}\right)^{\beta-1} \exp\left[-\left(\frac{t}{\alpha}\right)^{\beta}\right]$$

$\beta$ : shape parameter

$\alpha$  or  $\eta$ : characteristic life ( $t_{63}$ )

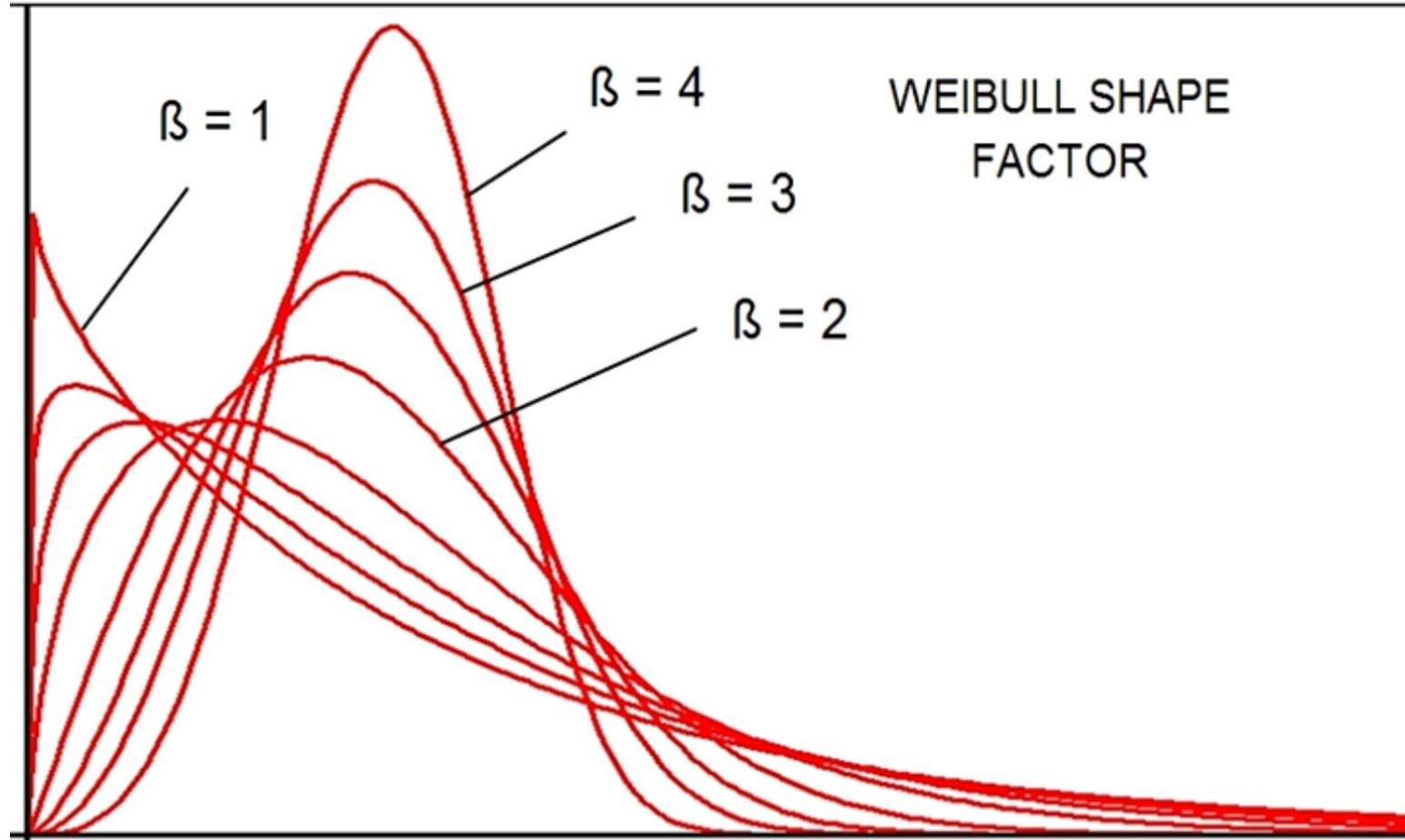
A diagram with two red arrows. One arrow points from the text 'beta: shape parameter' to the beta symbol in the exponent of the exponential term of the formula. The other arrow points from the text 'alpha or eta: characteristic life (t\_63)' to the alpha symbol in the denominator of the fraction (t/alpha) in the formula.

# Weibull PDF and CDF

$$f(t) = \left(\frac{\beta}{\alpha}\right) \left(\frac{t}{\alpha}\right)^{\beta-1} \exp\left[-\left(\frac{t}{\alpha}\right)^{\beta}\right]$$

$$F(t) = \int_0^t f(t) dt = 1 - \exp\left[-\left(\frac{t}{\alpha}\right)^{\beta}\right]$$

# Effect of Shape Parameter



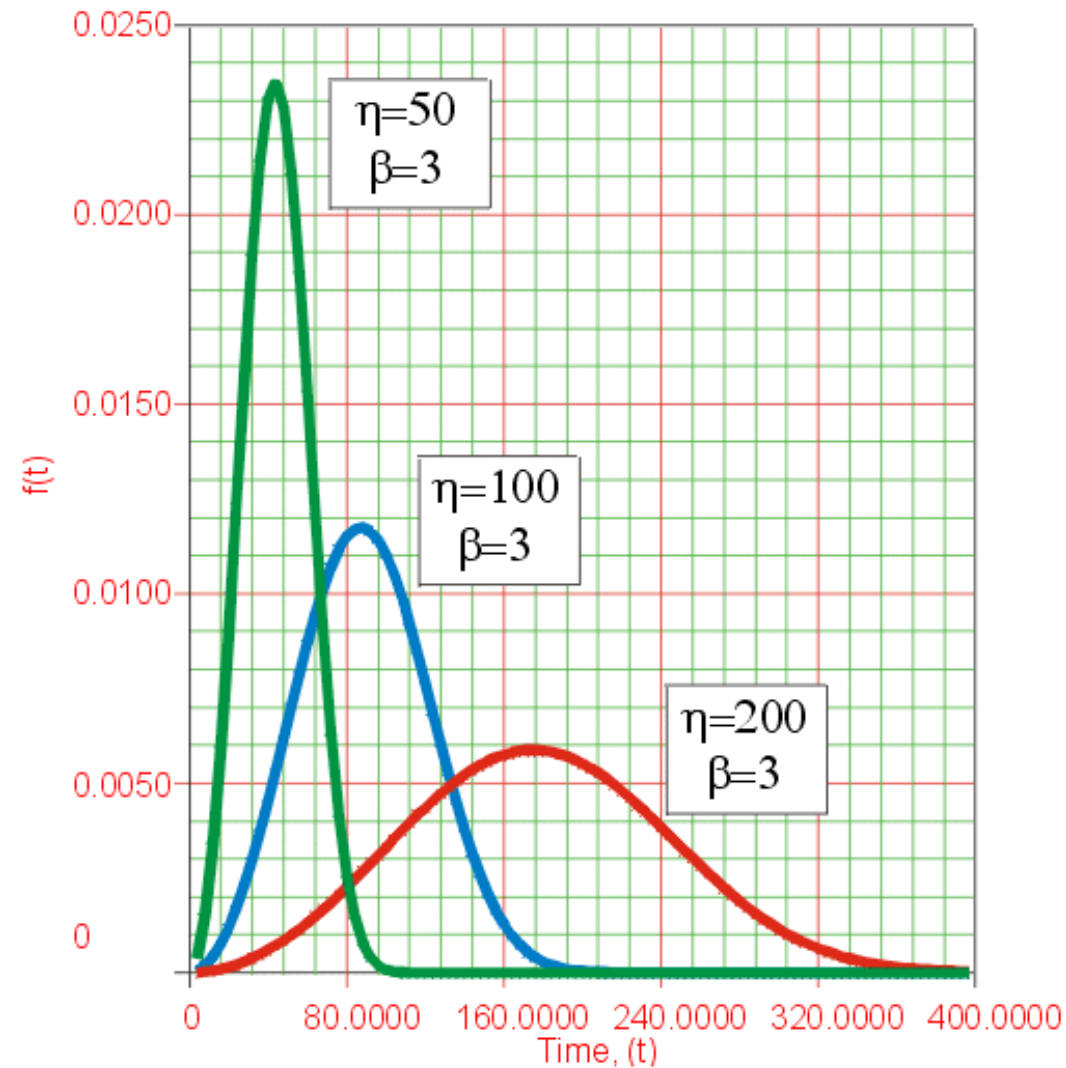
# Shape

- $\beta$  is referred to as the shape parameter
  - It affects the shape of the distribution in the sense that for  $\beta < 1$  the distribution looks like an exponential
  - For  $\beta = 1$  the distribution is exponential
  - For  $\beta > 3$ , the distribution is close to symmetrical
  - For  $1 < \beta < 3$  it is skewed
    - $\beta = 2.0$ : identical to the Rayleigh distribution
    - $\beta = 2.5$ : approximates the lognormal distribution (see later slides)

Weibull Distribution PDF

$$f(t) = \left(\frac{\beta}{\alpha}\right) \left(\frac{t}{\alpha}\right)^{\beta-1} \exp\left[-\left(\frac{t}{\alpha}\right)^\beta\right]$$

# Effect of Characteristic Life



# Scale Parameter

- $\alpha$  or  $\eta$  is the scale parameter
  - Influences both the mean and the spread of the distribution
  - Defines the age by which 63.2% of the units will fail, called characteristic lifetime

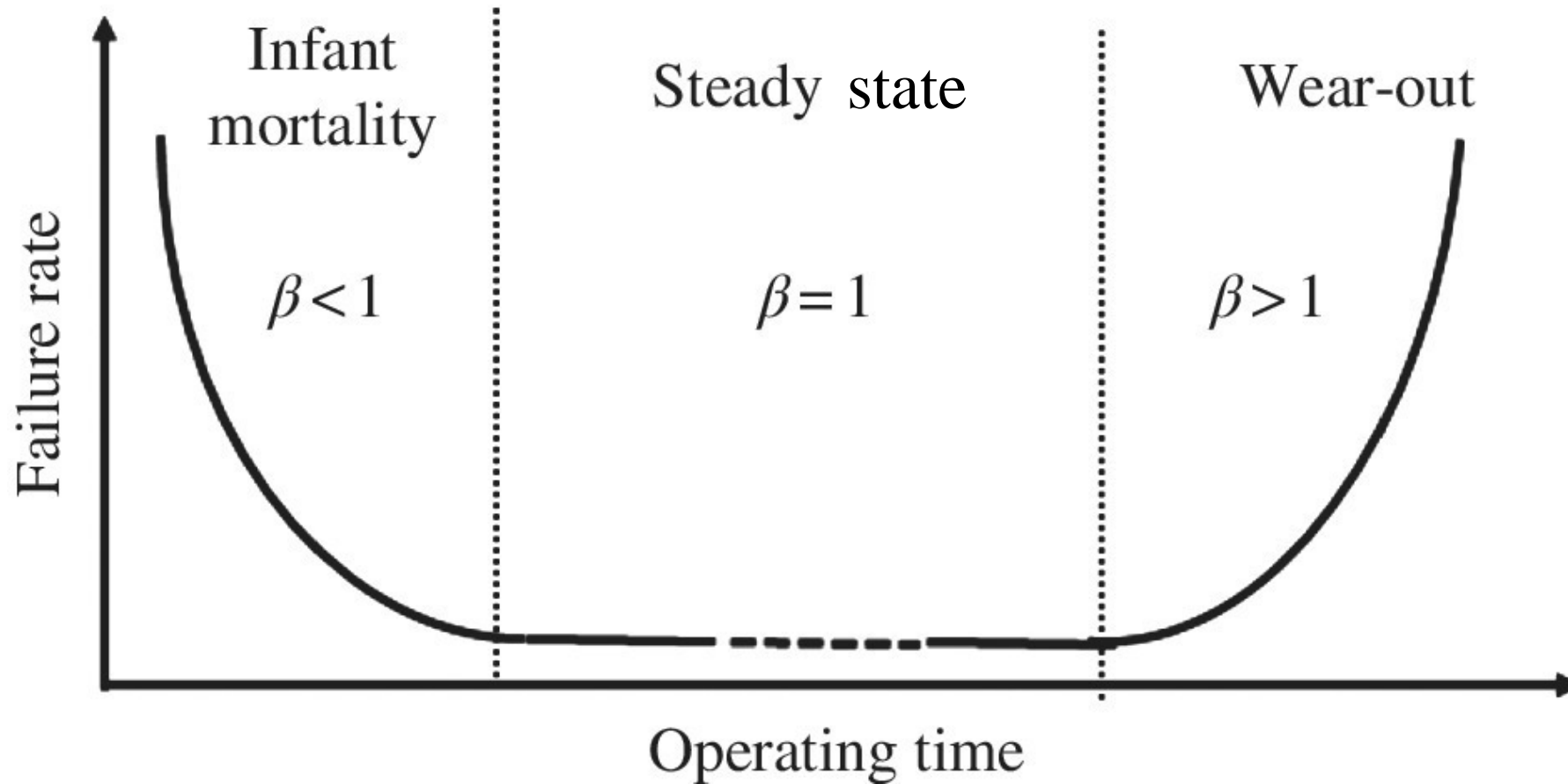
Weibull Distribution PDF

$$f(t) = \left(\frac{\beta}{\alpha}\right) \left(\frac{t}{\alpha}\right)^{\beta-1} \exp\left[-\left(\frac{t}{\alpha}\right)^\beta\right]$$

CDF at characteristic lifetime: 
$$F(\eta) = 1 - e^{\left(-\left(\frac{\eta}{\eta}\right)^\beta\right)} = 1 - e^{-(1)^\beta} = 1 - \frac{1}{e} \cong 0.632$$

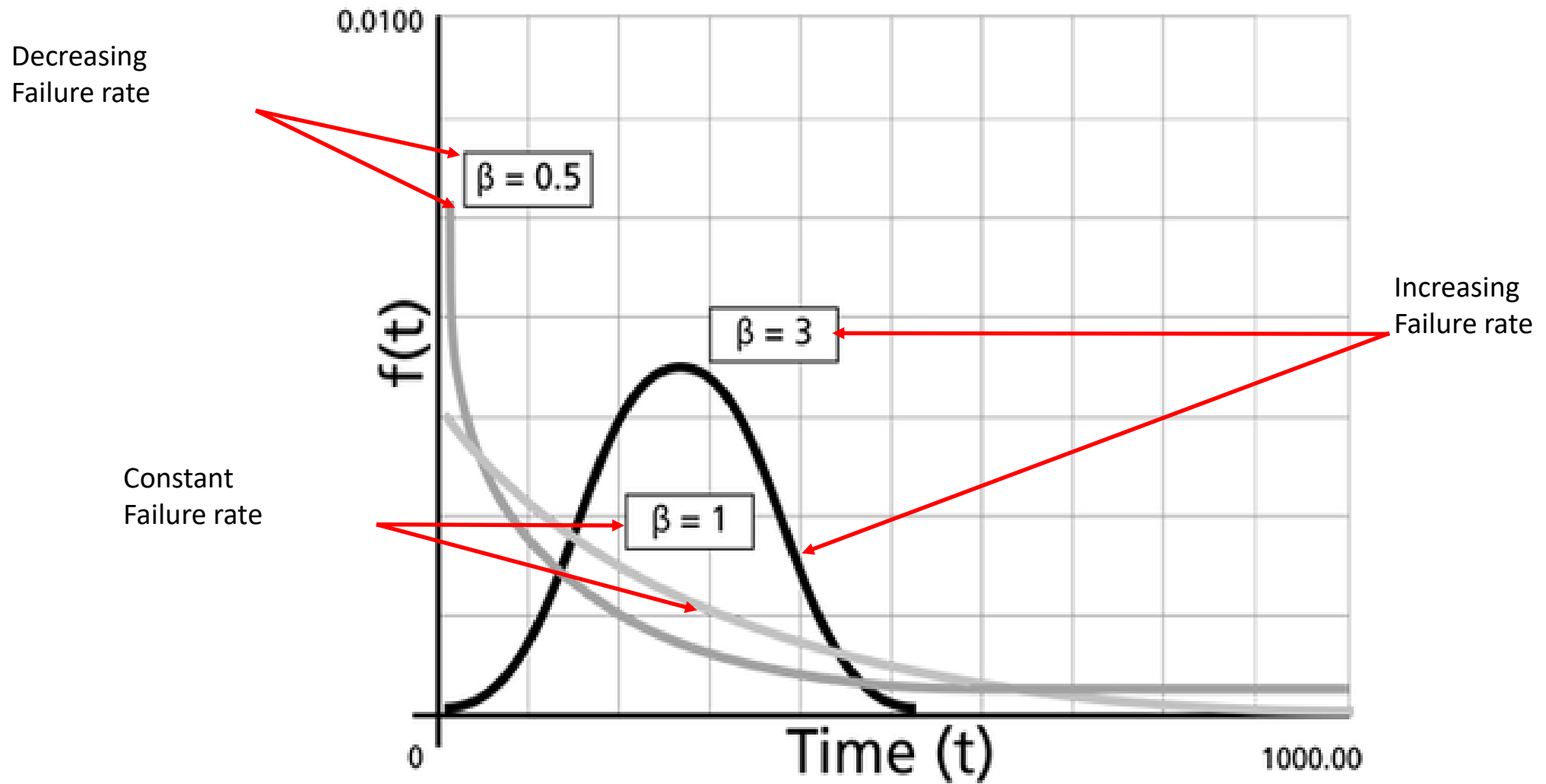


# Bathtub Curve & Weibull Distribution



<https://www.oreilly.com/library/view/thermodynamic-degradation-science/9781119276227/c09.xhtml>

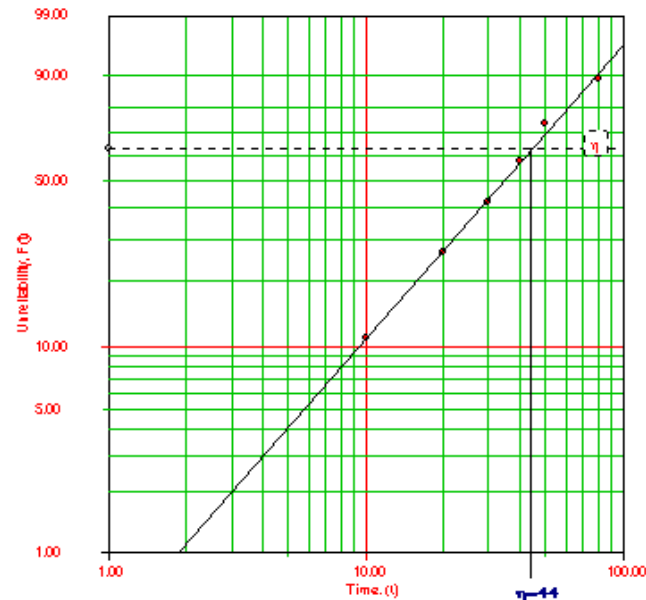
<http://www.applied-statistics.org/Glossary/BathTubCurve.html>



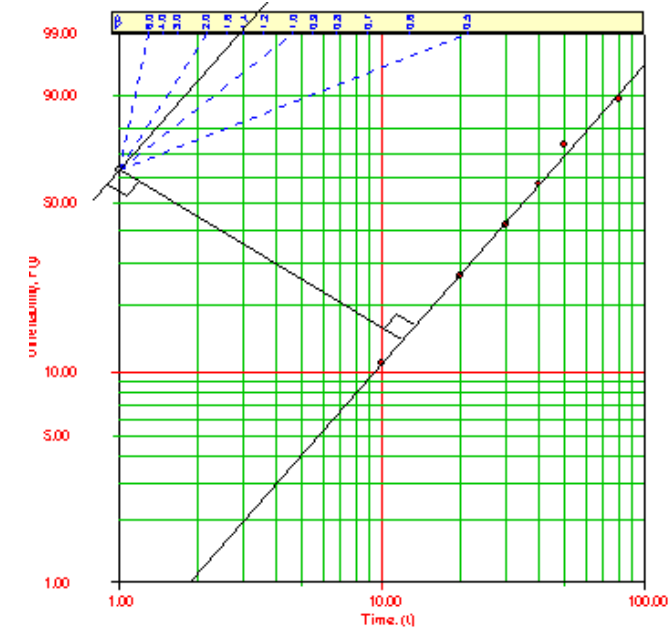
# Weibull Plotting

Failure Time (hours)	Unreliability Estimate
10	10.9%
20	26.6%
30	42.2%
40	57.8%
50	73.4%
80	89.1%

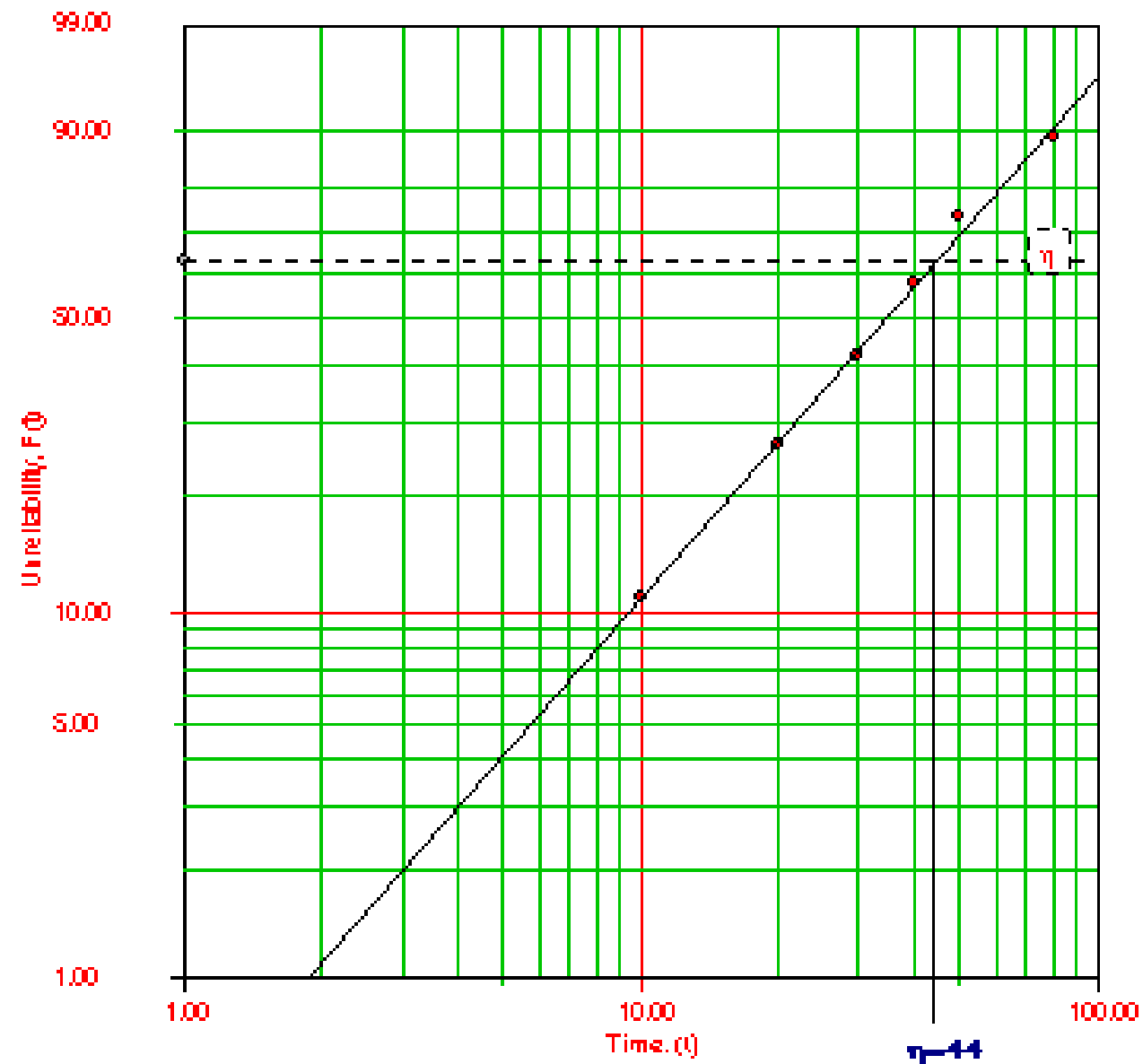
$\alpha$  or  $\eta$  is when 63% failed



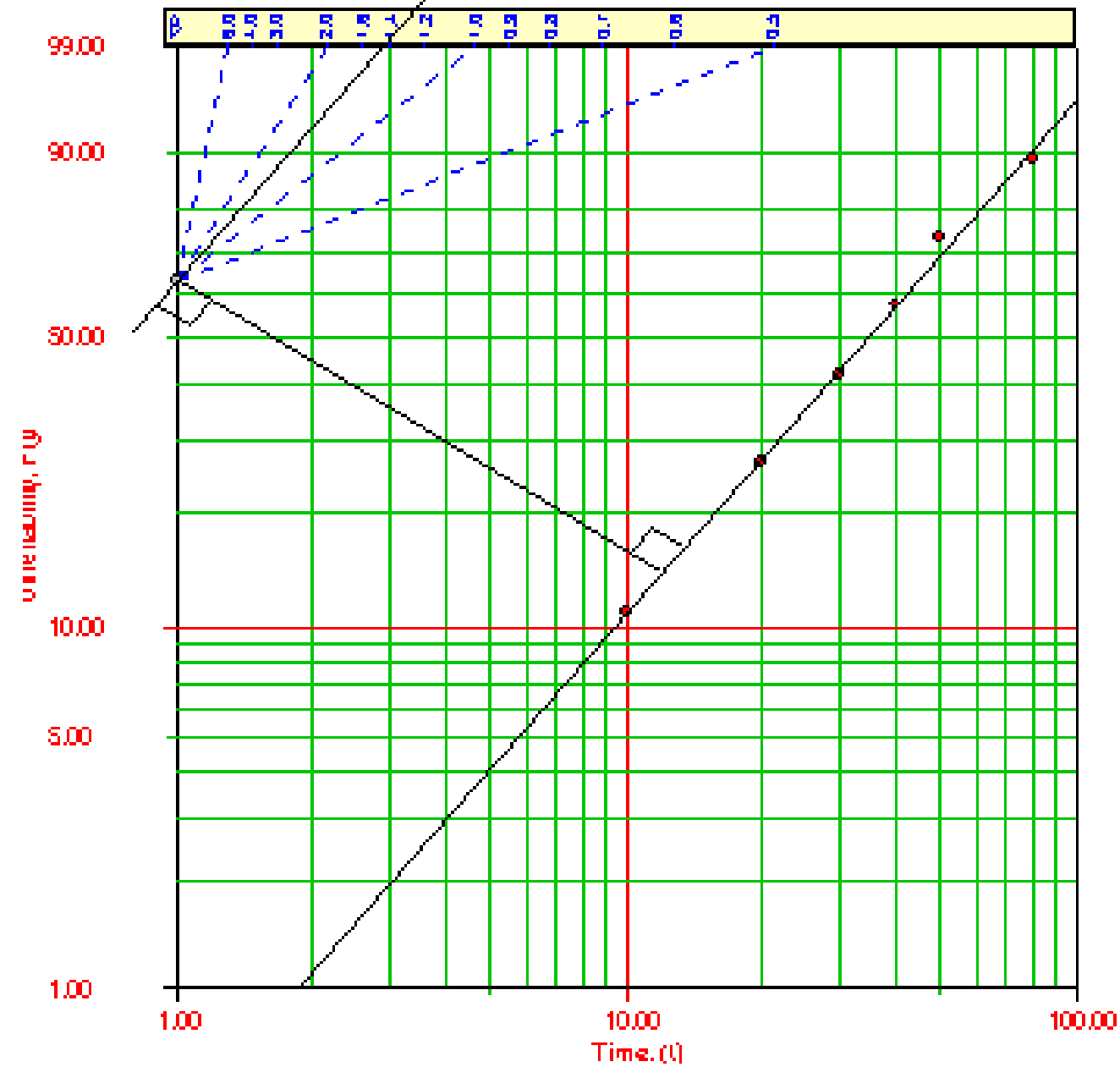
$\beta$ : is the slope



$\alpha$  or  $\eta$  is when 63% failed



$\beta$ : is the slope



# Example

- An item is randomly drawn from a two - parameter Weibull population having a shape parameter  $\beta = 1.5$  and a scale parameter  $\eta = 100.0$  hours. What is the probability that the item fails before achieving a life of  $x = 25$  hours?

# Why the Lognormal Distribution?

- Normal distribution
  - Most common
  - Represents **addition** of random variables
  - Application: dimension of a manufactured product
  - Used by quality engineers to reduce variability
- Lognormal distribution
  - Significant applications in science, finance, etc.
  - Represents **multiplication** of random variables

# Background of Lognormal

*So lets say that  $Y$  is the product  
of lots of independent, positive  
random variables ...*

$$Y = X_1 \times X_2 \times X_3 \times \dots \times X_i \times \dots$$

*independent random variables*



# Multiplication to Addition

$$Y = X_1 \times X_2 \times X_3 \times \dots \times X_i \times \dots$$

*how do we model Y?*

we can then apply the 'natural logarithm' to restate  
the above equation ...

$$\ln Y = \ln X_1 + \ln X_2 + \ln X_3 + \dots + \ln X_i + \dots$$

# Addition of Random Variables

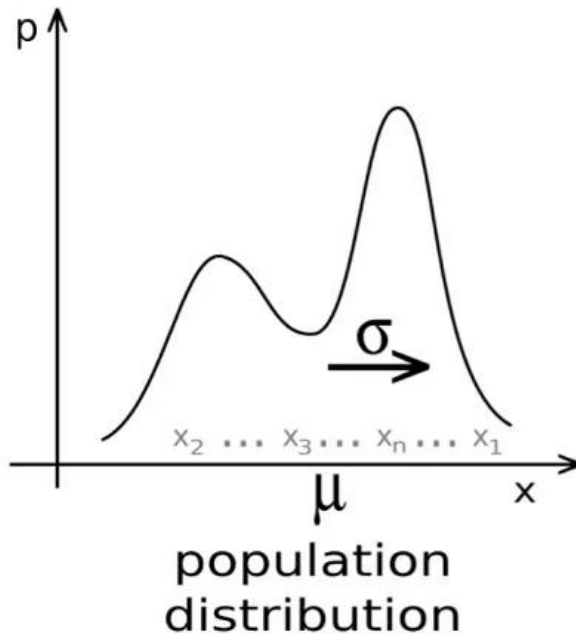
*each one of these logarithms  
is a random variable*

$$\ln Y = \ln X_1 + \ln X_2 + \ln X_3 + \dots + \ln X_i + \dots$$

*which means that  $\ln Y$  is a **normally distributed** random variable  
... due to the **Central Limit Theorem***

# Note: Central Limit Theorem

No matter the shape of the population of a distribution

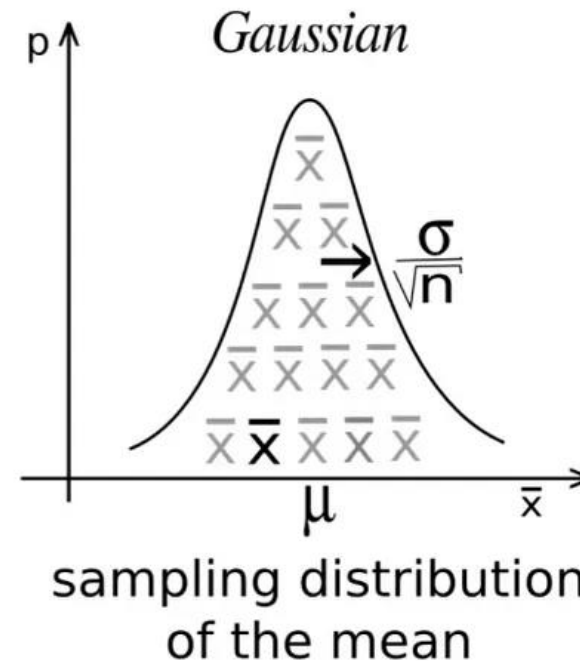


samples  
of size  $n$

$\bar{x}$

$\bar{x}$

If sample batches of data from that distribution (with replacement) and take the mean of each batch, the mean values from all those batches will be normally distributed.



# Note: Central Limit Theorem Demo Explanation

## Central Limit Theorem Premises

- Independent



- Identically distributed



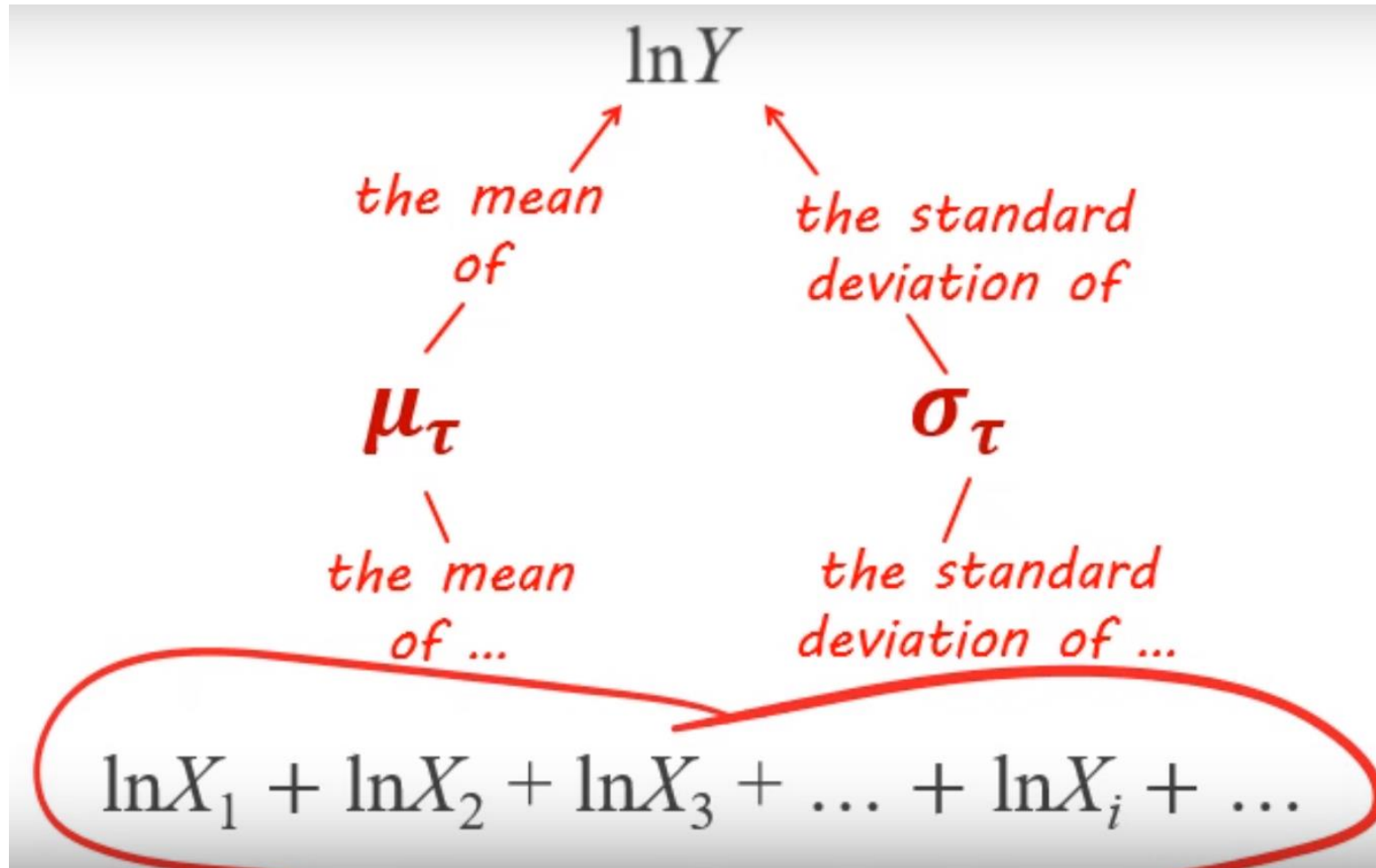
- Finite variance



## Candy Jar Demo

- Each trial (or individual estimate) does not know about other trials or estimates
- Looking at the same underlying problem
- Estimates are not so unreasonable that a single estimate will throw off the average

# Attributes of $\ln(Y)$

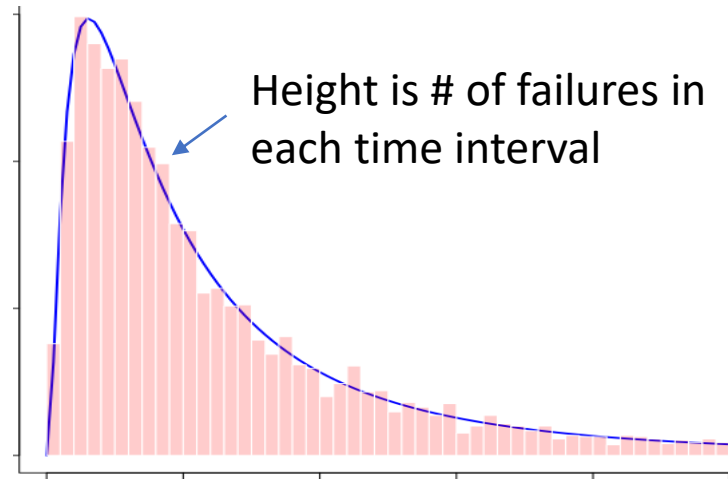


$\tau$ : referring to the  $X_i$  or normal domain (see following slides on "Conversion")

# Visual Interpretation from Histogram Perspective

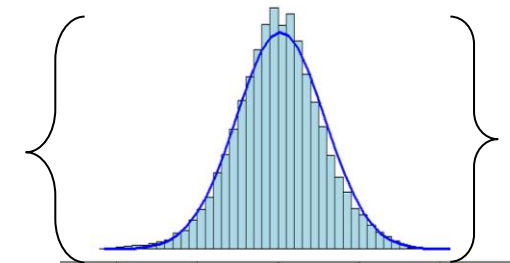
$$\textit{Time to failure} = A_0 e^x$$

Bucketizing failure times in a test



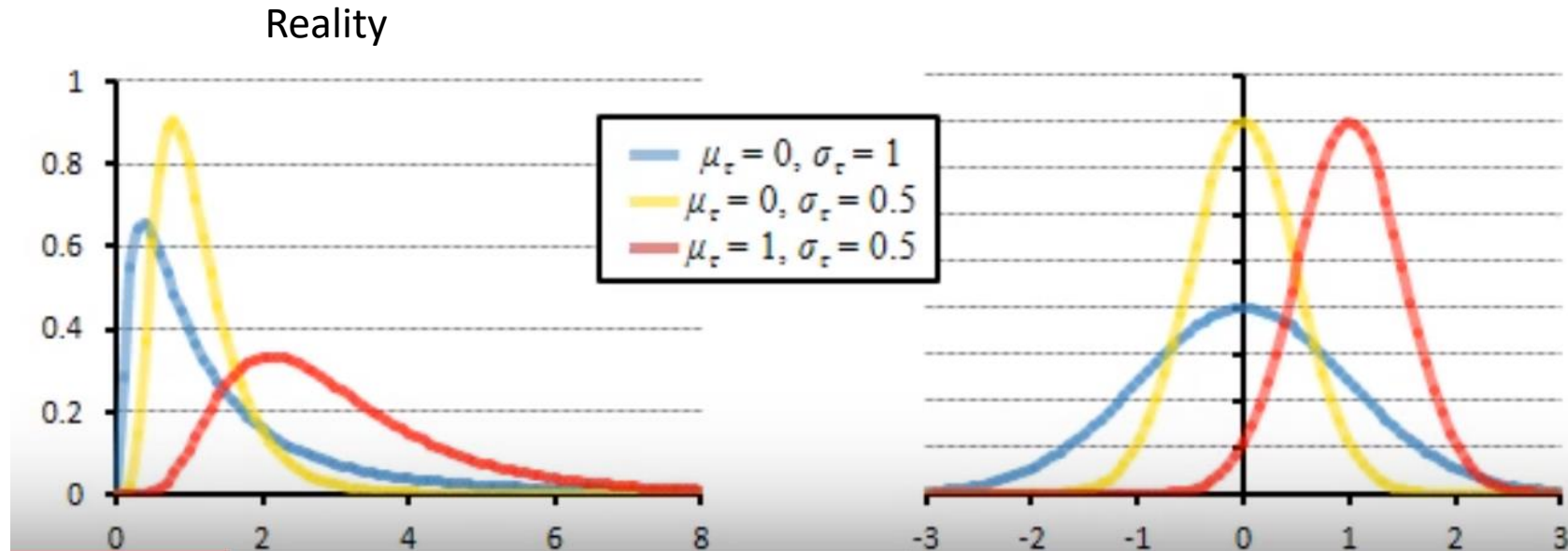
Reality

$$= A_0 e$$



Underlying factor x is normally distributed that influences overall distribution

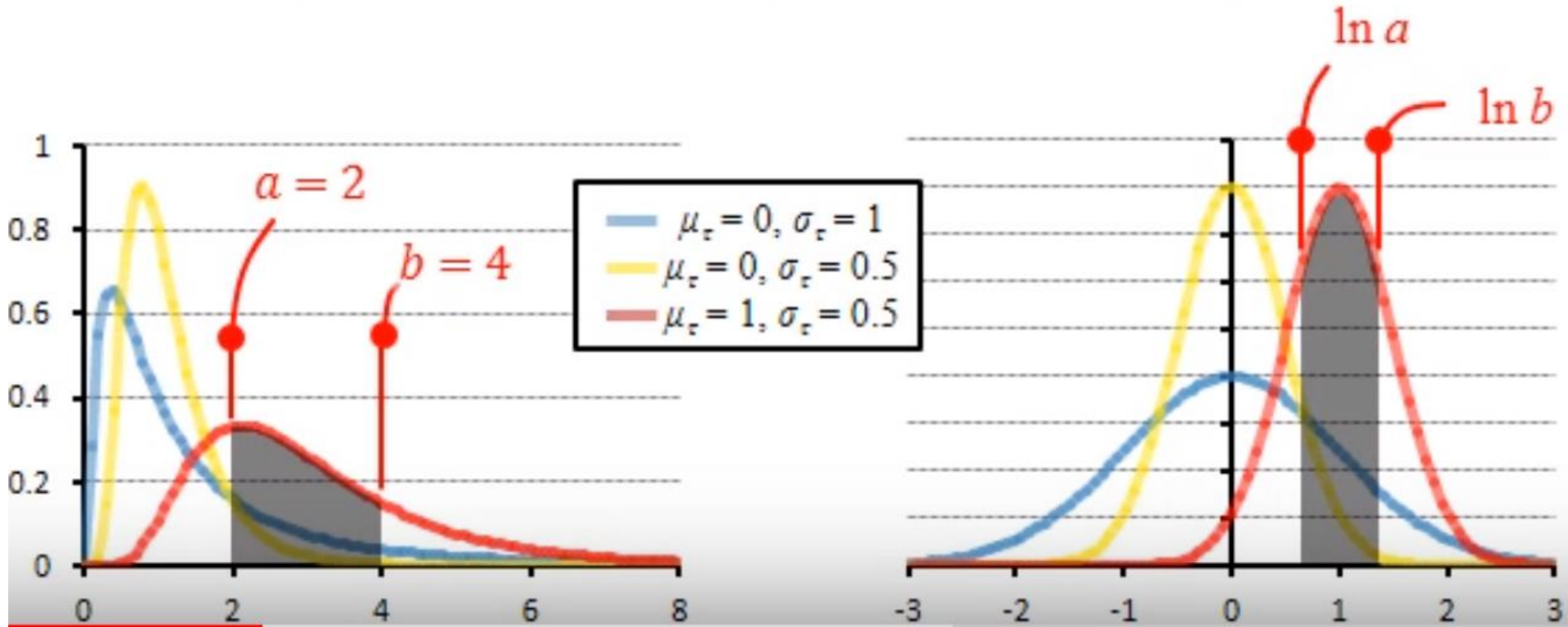
# Conversion



$t$   $\tau = \ln(t)$   $\tau$

**Note:** we are now using  $t$  instead of  $Y$

# Equivalent Areas





# Comparison to Normal

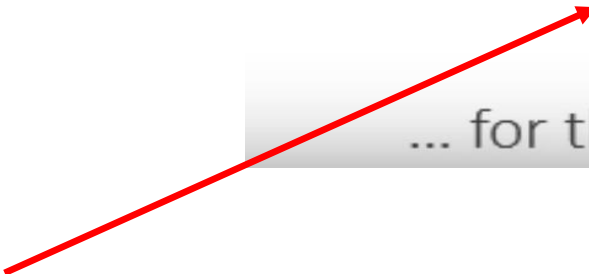
so the probability density function for the normal distribution ...

$$f(t) = \frac{1}{\sigma\sqrt{2\pi}} e^{-\frac{(t-\mu)^2}{2\sigma^2}}$$

... becomes ...

$$f(y) = \frac{1}{\mathbf{y\sigma_\tau}\sqrt{2\pi}} e^{-\frac{(\mathbf{\ln y} - \mu_\tau)^2}{2\sigma_\tau^2}}$$

... for the lognormal distribution

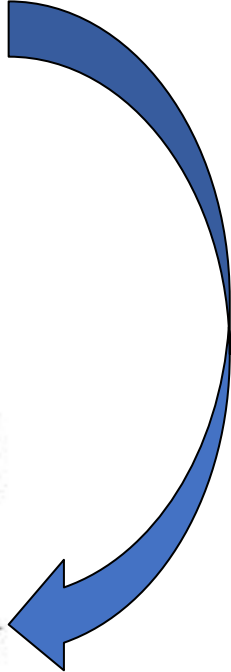


1/y pre-factor is needed to ensure that the integral of  $f(y) \times dy$  will continue to represent the probability of failure. Due to the fact that  $d\ln(y) = (1/y) dy$

# Lognormal Probability Density Function

$$f(t) = \frac{1}{t\sigma_\tau\sqrt{2\pi}} e^{-\frac{(\ln t - \mu_\tau)^2}{2\sigma_\tau^2}}$$

$\mu_\tau$  also written  
as  $\ln(t_{50})$

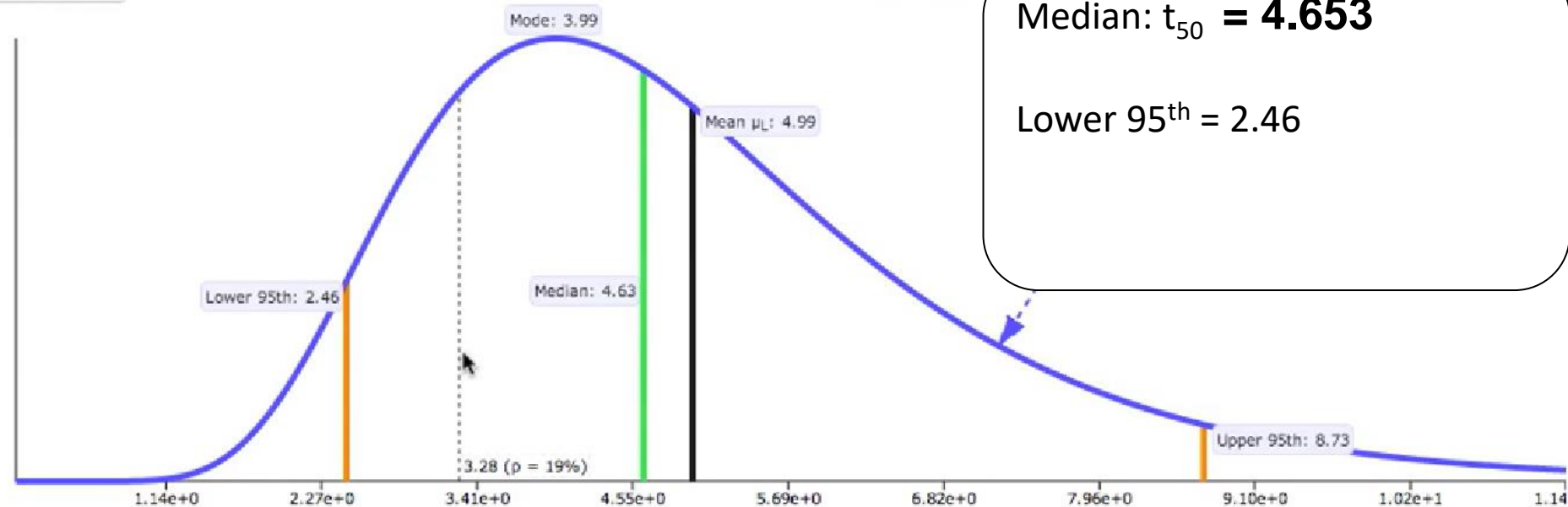


The mean is  $\mu = e^{\mu_\tau + \frac{\sigma_\tau^2}{2}}$

The median is  $t_{50} = e^{\mu_\tau}$

The mode is  $\hat{t} = e^{\mu_\tau - \sigma_\tau^2}$

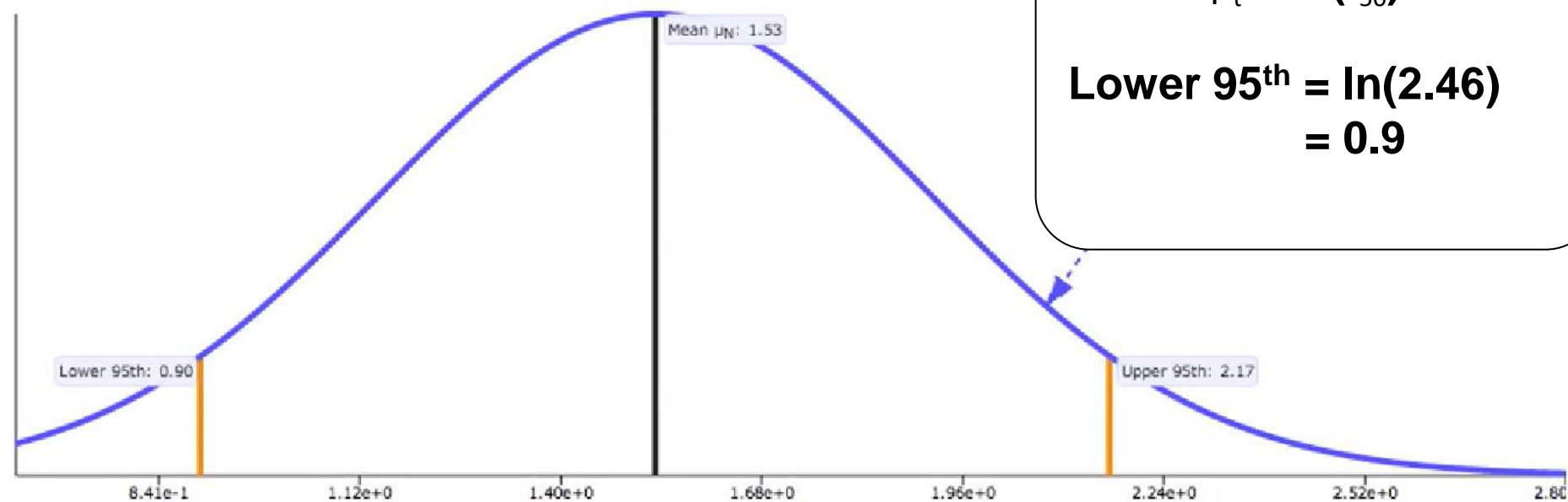
$\tau$ : referring to the  $X_i$  or normal domain



Median:  $t_{50} = 4.653$

Lower 95<sup>th</sup> = 2.46

Underlying Normal Distribution with mean  $\mu_N =$   and standard



Mean:  $\mu_\tau = \ln(t_{50}) = 1.53$

Lower 95<sup>th</sup> =  $\ln(2.46)$   
= 0.9

# Lognormal PDF and CDF

$$f(t) = \frac{1}{\sigma t \sqrt{2\pi}} \exp \left\{ - \left[ \frac{\ln(t) - \ln(t_{50})}{\sigma \sqrt{2}} \right]^2 \right\}$$

$$F(t) = \int_0^t f(t) dt$$

The integral needs to be numerically evaluated

# Lognormal CDF Evaluation

$$F(t) = \int_0^t f(t)dt$$

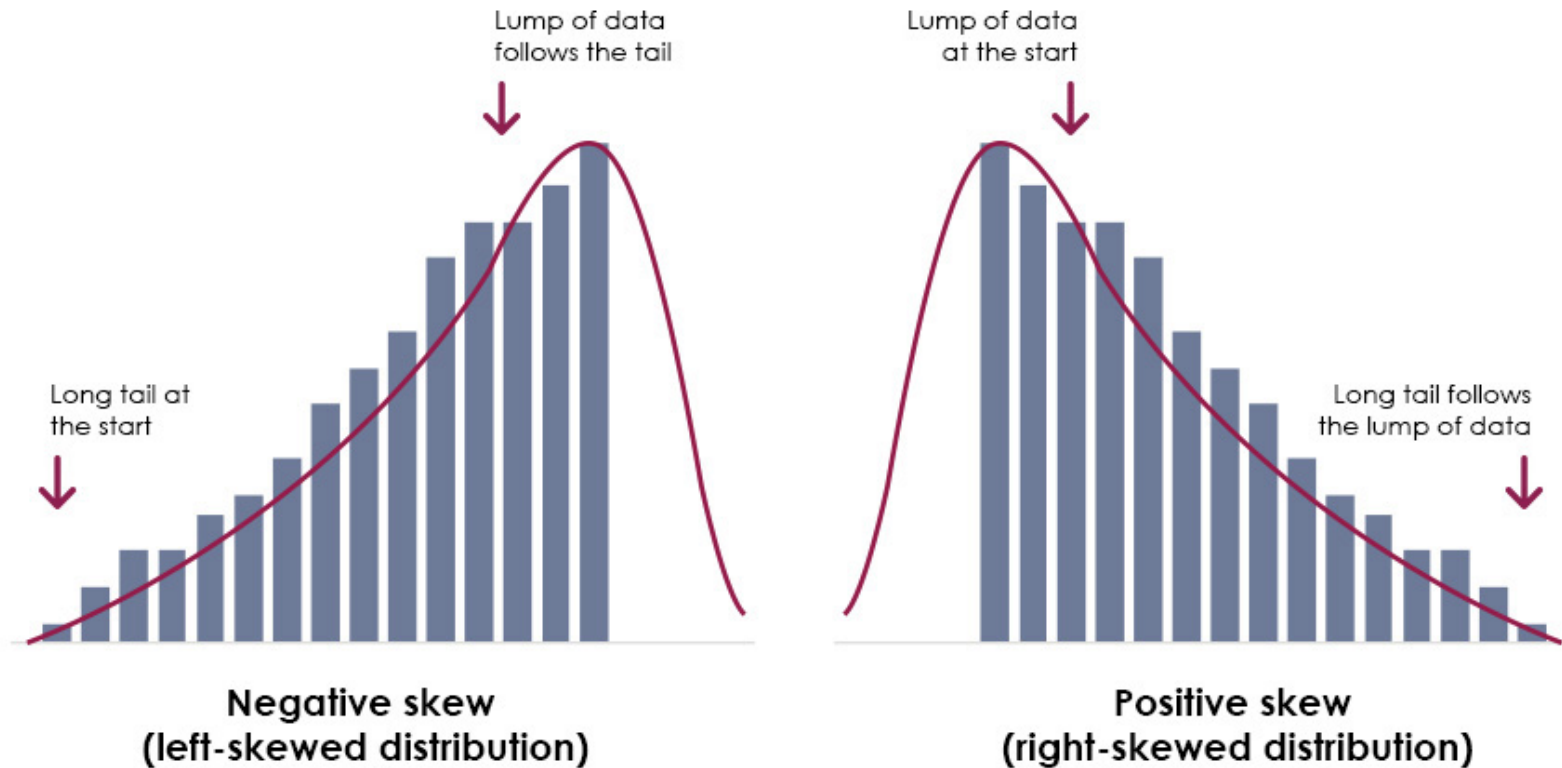
Numerical evaluation of integral by error function complement

$$F(t) = \frac{1}{2} \operatorname{erfc} \left( \frac{\ln(t_{50}) - \ln(t)}{\sigma\sqrt{2}} \right) \quad \text{for } t \leq t_{50}$$

$$F(t) = 1 - \frac{1}{2} \operatorname{erfc} \left( \frac{\ln(t) - \ln(t_{50})}{\sigma\sqrt{2}} \right) \quad \text{for } t \geq t_{50}$$

# Skewed Data

- Skewness is a way to describe the symmetry of a distribution

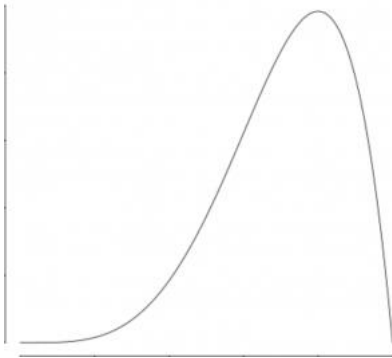


# Skewed Data

- Skewness is a way to describe the symmetry of a distribution

## Left Skewed

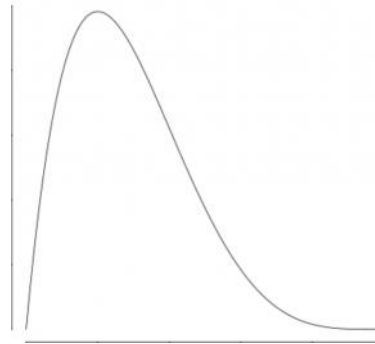
Left Skewed Distribution



- “tail” on the left side of the distribution
- negatively-skewed

## Right Skewed

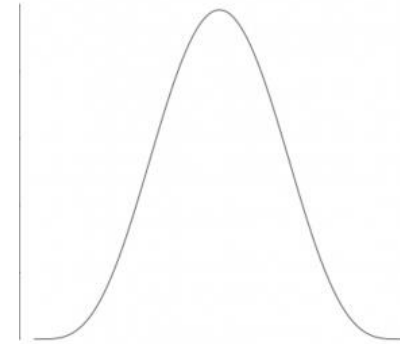
Right Skewed Distribution



- “tail” on the right side of the distribution
- positively-skewed

## No Skew

No Skew

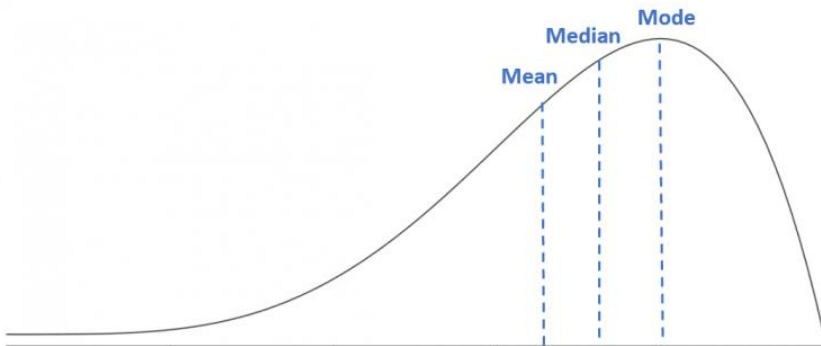


- symmetrical on both sides

# Skewed Data

- Skewness is a way to describe the symmetry of a distribution

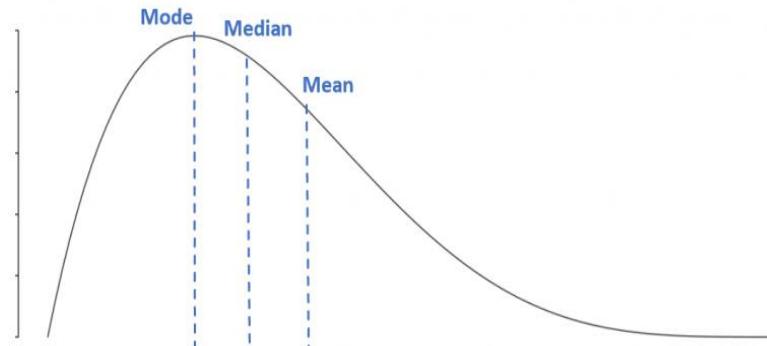
## Left Skewed



Left Skewed Distribution

- $\text{Mean} < \text{Median} < \text{Mode}$
- Most of the data points falls on the **right side** of the mean

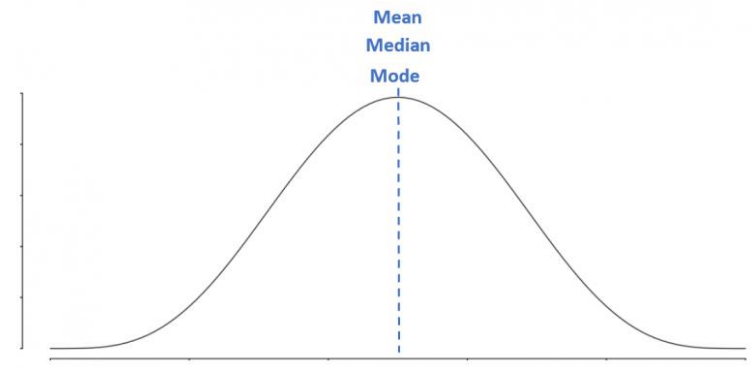
## Right Skewed



Right Skewed Distribution

- $\text{Mode} < \text{Median} < \text{Mean}$
- Most of the data points falls on the **left side** of the mean

## No Skew



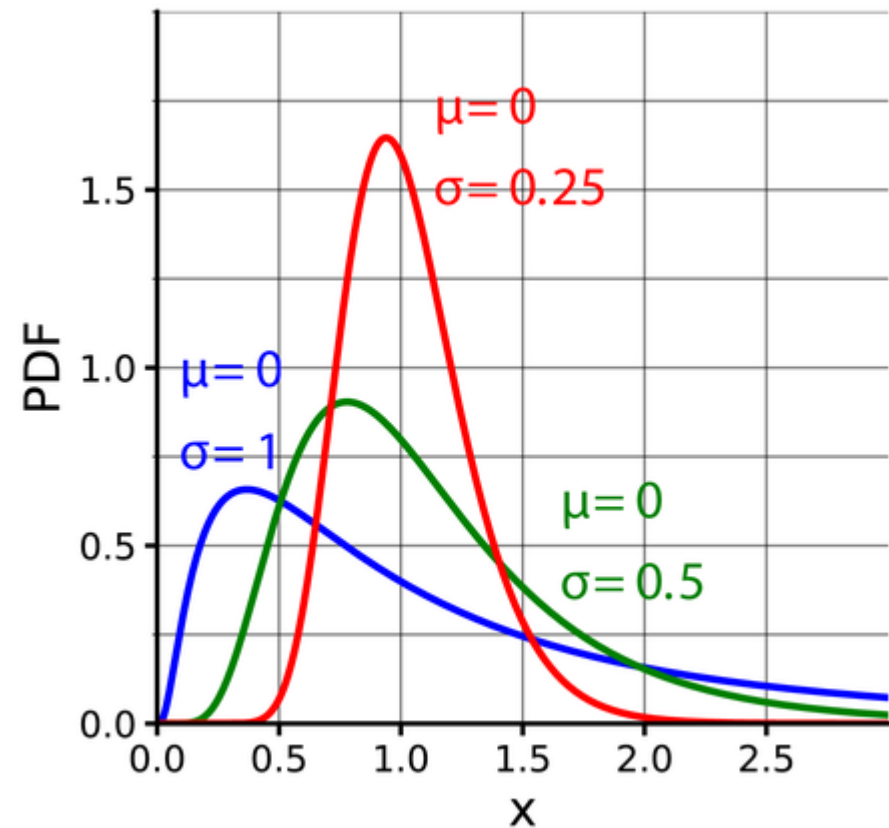
No Skew

- $\text{Mean} = \text{Median} = \text{Mode}$



# Lognormal: Right-skewed data

- Values are positively skewed
- Better at predicting longer lifetimes
- Gradual degradation over time



# Normal vs. Lognormal

Distribution	$Y = \ln(X)$	$X = e^Y$
Notation	$N(\mu_y, \sigma_y^2)$	$lognormal(\mu_x, \sigma_x^2)$
PDF	$\frac{1}{\sigma_y \sqrt{2\pi}} \exp \left\{ - \left[ \frac{x - \mu_y}{\sigma_y \sqrt{2}} \right]^2 \right\}$	$\frac{1}{\sigma_y x \sqrt{2\pi}} \exp \left\{ - \left[ \frac{\ln x - \mu_y}{\sigma_y \sqrt{2}} \right]^2 \right\}$
Mean	$\mu_y$	$\mu_x = \exp(\mu_y + \frac{1}{2} \sigma_y^2)$
Variance	$\sigma_y^2$	$\sigma_x^2 = \mu_x^2 (e^{\sigma_y^2} - 1)$ $= \exp(2\mu_y + \sigma_y^2) (e^{\sigma_y^2} - 1)$

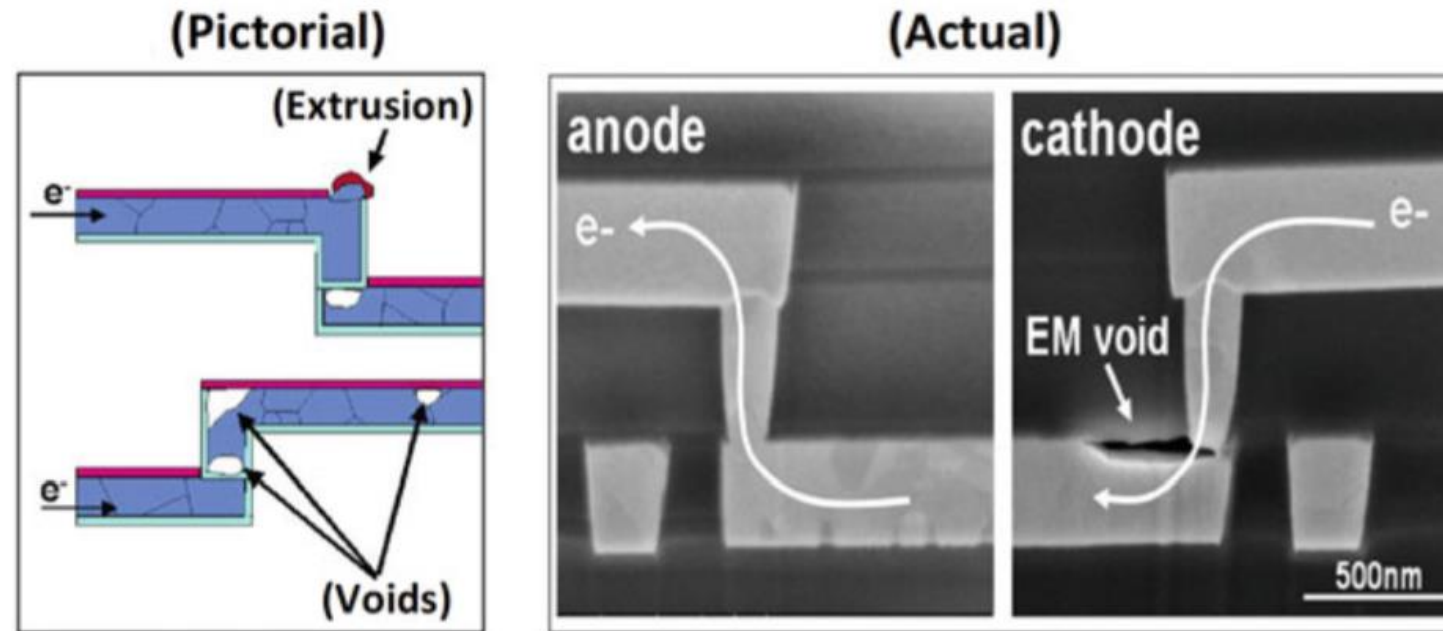
Thomopoulos, Nick T. "Statistical distributions." Applications and Parameter Estimates. Cham, Switzerland: Springer International Publishing (2017).

# Example

- In a fracture failure experiment, one of the variables,  $x$ , is detected as lognormal with  $LN(2.5, 1.1^2)$ . Compute the corresponding mean and variance of the  $x$ .

# Lognormal Applications

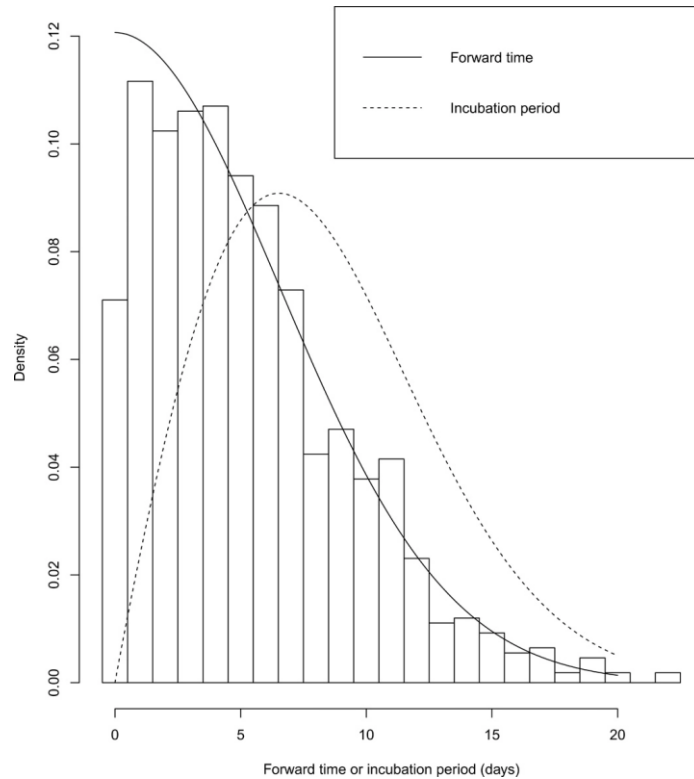
- Electromigration



McPherson, Joe W. Reliability physics and engineering: time-to-failure modeling. Springer, 2018.

# Lognormal Applications

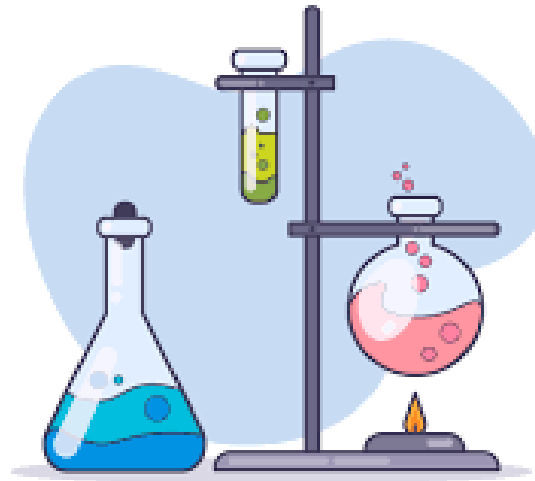
## Medical Science



### Modeling onset of symptoms from COVID

Qin, Jing, et al. "Estimation of incubation period distribution of COVID-19 using disease onset forward time: a novel cross-sectional and forward follow-up study." *Science advances* 6.33 (2020)

## Chemical Reactions



### Modeling chemical kinetics

Schnoerr, David, Guido Sanguinetti, and Ramon Grima. "Comparison of different moment-closure approximations for stochastic chemical kinetics." *The Journal of Chemical Physics* 143.18 (2015)

## Corrosion & Fatigue Cracking



Pic from <https://www.thoughtco.com/how-rust-works-608461>

Fu, Guoyang, et al. "Prediction of fracture failure of steel pipes with sharp corrosion pits using time-dependent reliability method with lognormal process." *Journal of Pressure Vessel Technology* 141.3 (2019).

# Summary

- Times of failures should lead your selection of appropriate probability distribution

Model	Description	Characteristic	Advantages	Disadvantages
Exponential	The failure rate is constant over time	Memoryless (the age of the item has no effect on its future failure rate)	Simple and easy to understand (only one parameter to estimate)	May not reflect real-life scenarios where the failure rate changes over time
Log-normal	The failure rate changes over time	Model is appropriate when underlying process has a large number of causes of failure and the failure times are spread out over a large range	Can reflect real-life scenarios where the failure rate changes over time	Requires more parameters to estimate
Weibull	The shape parameter determines the shape of the failure rate curve, which can be constant, increasing or decreasing over time	Flexibility in modeling	Can model a wide range of failure rate patterns	Estimation process can be more complex

# Summary

	Exponential	Log-normal	Weibull
PDF	$f(t) = \lambda e^{-\lambda t}$	$f(t) = \frac{1}{\sigma t \sqrt{2\pi}} \exp \left\{ - \left[ \frac{\ln(t) - \ln(t_{50})}{\sigma \sqrt{2}} \right]^2 \right\}$	$f(t) = \left( \frac{\beta}{\alpha} \right) \left( \frac{t}{\alpha} \right)^{\beta-1} \exp \left[ - \left( \frac{t}{\alpha} \right)^\beta \right]$
CDF	$F(t) = 1 - e^{-\lambda t}$	$F(t) = \frac{1}{2} \operatorname{erfc} \left( \frac{\ln(t_{50}) - \ln(t)}{\sigma \sqrt{2}} \right)$	$F(t) = 1 - \exp \left[ - \left( \frac{t}{\alpha} \right)^\beta \right]$
Parameter	$\lambda$	$\mu = \ln(t_{50}), \sigma$ (*note: these refer to the parameters in normal form, i.e. $\mu = \mu_y, \sigma = \sigma_y$ on slide "Normal vs. Lognormal")	$\beta = \text{shape parameter}$ (affects shape of distribution)  $\eta = \text{scale parameter}$ or characteristic lifetime (affects spread of distribution)