

第四次作业

1 1. EM算法理论练习

1.0.1 1.1 推导三硬币模型的EM算法中隐变量后验分布的计算公式以及参数更新公式

三硬币模型：

假设有三枚硬币，分别记为A、B、C，这些硬币正面向上的概率分别是 π, p, q 。

进行如下掷硬币试验：先掷硬币A，根据其结果选择硬币B或硬币C，正面选硬币B，反面选硬币C；然后掷选出的硬币，出现正面记为1，出现反面记为0；独立地重复n次试验。

假设只能观测到掷硬币的结果，不能观测到掷硬币的过程。问如何估计三硬币正面出现的概率，即三硬币模型的参数 $\theta = (\pi, p, q)$ 。

模型构建：

引入随机变量 $x \in \{0, 1\}$ 表示一次试验观测的结果是1或者0， x 是观测变量，可以观测。

引入随机变量 $z \in \{0, 1\}$ 表示未观测到的掷硬币A的结果， z 是隐变量，不可观测。其中 $z = 1$ 表示硬币A为正面， $z = 0$ 表示硬币A为反面。

写出似然函数的对数 $LL(\theta)$ ，并引入隐变量：

$$\begin{aligned} LL(\theta) &= \sum_{i=1}^n \log(P(x_i | \theta)) \\ &= \sum_{i=1}^n \log(P(x_i, z_i = 1 | \theta) + P(x_i, z_i = 0 | \theta)) \\ &= \sum_{i=1}^n \log\left(Q(z_i = 1) \frac{P(x_i, z_i = 1 | \theta)}{Q(z_i = 1)} + Q(z_i = 0) \frac{P(x_i, z_i = 0 | \theta)}{Q(z_i = 0)}\right) \\ &\geq \sum_{i=1}^n \left[Q(z_i = 1) \log\left(\frac{P(x_i, z_i = 1 | \theta)}{Q(z_i = 1)}\right) + Q(z_i = 0) \log\left(\frac{P(x_i, z_i = 0 | \theta)}{Q(z_i = 0)}\right)\right] \end{aligned} \quad (17)$$

E (expectation) 步：

由EM算法知，隐变量的后验分布计算公式为：

$$\begin{aligned} Q(z_i = 1) &= P(z_i = 1 | x_i, \theta^{(t)}) \\ &= \frac{P(z_i = 1 | \theta^{(t)})P(x_i | z_i = 1, \theta^{(t)})}{\sum_{z_i=0}^{z_i=1} P(z_i | \theta^{(t)})P(x_i | z_i, \theta^{(t)})} \\ &= \frac{\pi^{(t)}[p^{(t)}]^{x_i}[1-p^{(t)}]^{1-x_i}}{\pi^{(t)}[p^{(t)}]^{x_i}[1-p^{(t)}]^{1-x_i} + [1-\pi^{(t)}][q^{(t)}]^{x_i}[1-q^{(t)}]^{1-x_i}} \\ &\triangleq \mu_i^{(t)} \\ Q(z_i = 0) &= P(z_i = 0 | x_i, \theta^{(t)}) \\ &= \frac{P(z_i = 0 | \theta^{(t)})P(x_i | z_i = 0, \theta^{(t)})}{\sum_{z_i=0}^{z_i=1} P(z_i | \theta^{(t)})P(x_i | z_i, \theta^{(t)})} \\ &= 1 - \mu_i^{(t)} \end{aligned} \quad (18)$$

M (maximize) 步：

由于 $Q^{(t)}(z)\log Q^{(t)}(z)$ 是常数，所以参数 θ 更新公式为：

$$\begin{aligned} \theta^{(t+1)} &= \arg \max_{\theta} \sum_{i=1}^n \left[Q(z_i = 1) \log\left(\frac{P(x_i, z_i = 1 | \theta)}{Q(z_i = 1)}\right) + Q(z_i = 0) \log\left(\frac{P(x_i, z_i = 0 | \theta)}{Q(z_i = 0)}\right)\right] \\ &= \arg \max_{\theta} \sum_{i=1}^n \sum_{z_i} Q(z_i) \log(P(x_i, z_i | \theta)) \\ &= \arg \max_{\theta} \sum_{i=1}^n \left[\mu_i^{(t)} \log(\pi p^{x_i} (1-p)^{1-x_i}) + (1-\mu_i^{(t)}) \log((1-\pi) q^{x_i} (1-q)^{1-x_i})\right] \\ &= \arg \max_{\theta} \sum_{i=1}^n \left\{ \mu_i^{(t)} [\log(\pi) + x_i \log(p) + (1-x_i) \log(1-p)] + (1-\mu_i^{(t)}) [\log(1-\pi) + x_i \log(q) + (1-x_i) \log(1-q)] \right\} \\ &\triangleq \arg \max_{\pi, p, q} f(\pi, p, q) \end{aligned} \quad (19)$$

将 $f(x)$ 分别对 $\theta = (\pi, p, q)$ 求偏导，当 $\sum_{i=1}^n \mu_i^{(t)} \neq 0$ 且 $\sum_{i=1}^n (1-\mu_i^{(t)}) \neq 0$ 时：

$$\begin{cases} \frac{\partial f(\pi, p, q)}{\partial \pi} = \sum_{i=1}^n \left(\frac{\mu_i^{(t)}}{\pi} - \frac{1-\mu_i^{(t)}}{1-\pi} \right) \equiv 0 & \implies \pi \left(\sum_{i=1}^n (\mu_i^{(t)} - 1) \right) = (\pi - 1) \left(\sum_{i=1}^n \mu_i^{(t)} \right) & \implies \pi = \frac{1}{n} \\ \frac{\partial f(\pi, p, q)}{\partial p} = \sum_{i=1}^n \left(\mu_i^{(t)} \frac{x_i}{p} - \mu_i^{(t)} \frac{x_i - 1}{p - 1} \right) \equiv 0 & \implies (p - 1) \sum_{i=1}^n x_i \mu_i^{(t)} = p \sum_{i=1}^n (x_i - 1) \mu_i^{(t)} & \implies p = \frac{\sum_{i=1}^n x_i \mu_i^{(t)}}{\sum_{i=1}^n \mu_i^{(t)}} \\ \frac{\partial f(\pi, p, q)}{\partial q} = \sum_{i=1}^n \left((1-\mu_i^{(t)}) \frac{x_i}{q} - (1-\mu_i^{(t)}) \frac{x_i - 1}{q - 1} \right) \equiv 0 & \implies (q - 1) \sum_{i=1}^n x_i (1-\mu_i^{(t)}) = q \sum_{i=1}^n (x_i - 1) (1-\mu_i^{(t)}) & \implies q = \frac{\sum_{i=1}^n x_i (1-\mu_i^{(t)})}{\sum_{i=1}^n (1-\mu_i^{(t)})} \end{cases}$$

当 $\sum_{i=1}^n \mu_i^{(t)} = 0$ 时，即 $\pi = 0$ 时，参数 p 的取值对结论无影响

当 $\sum_{i=1}^n (1-\mu_i^{(t)}) = 0$ 时，即 $\pi = 1$ 时，参数 q 的取值对结论无影响

1.0.1 1.2 假设硬币A、B、C正面向上的概率分别是0.7、0.3、0.6，按照三硬币模型的数据生成过程独立地重复100次试验并记录观测结果的序列

```
1 import numpy as np
2 def three_coins_model(pi,p,q): #分别为ABC硬币正面向上概率
3     z = np.random.choice([0,1],p=[1-pi,pi]) #投硬币A，其中z=1表示正面，z=0表示反面
4     if z == 1:
5         x = np.random.choice([0,1],p=[1-p,p]) #投硬币B，其中x=1表示正面，x=0表示反面
6     elif z == 0:
7         x = np.random.choice([0,1],p=[1-q,q]) #投硬币C，其中x=1表示正面，x=0表示反面
8     else:
9         return "error: something went wrong with z"
10    return x
11 x_list = []
12 for i in range(0,100):
13     x_list.append(three_coins_model(0.7,0.3,0.6))
14 print("100次独立重复试验结果: \n",x_list)
```

```
1 100次独立重复试验结果:
2 [0, 1, 0, 0, 0, 0, 0, 0, 0, 0, 0, 0, 0, 0, 1, 0, 0, 0, 0, 0, 1, 0, 1, 1, 0, 1, 0, 1, 1, 0, 0, 0, 0, 1, 1, 1, 0, 0,
0, 0, 1, 0, 0, 0, 0, 0, 0, 0, 0, 0, 0, 1, 0, 1, 1, 1, 0, 0, 0, 0, 0, 0, 0, 0, 1, 1, 1, 0, 0, 0, 0, 0, 1, 1, 1,
0, 1, 0, 0, 0, 0, 0, 0, 0, 0, 1, 0, 0, 1, 1, 0, 0, 1, 0, 1, 1, 0, 0, 1, 0, 0, 1, 1, 0, 1, 1]
```

1.0.1 1.3 利用EM算法根据上述观测序列估计各硬币正面向上的概率

```
1 import matplotlib.pyplot as plt
2 import pandas as pd
3 import seaborn as sns
4 import numpy as np
5 plt.rcParams['font.sans-serif'] = ['SimHei'] # 显示汉字
6 plt.rcParams['axes.unicode_minus']=False
7
8
9 def get_loglikelihood(theta_list, x): #获得指定参数下似然函数值
10     [pi, p, q] = theta_list
11     x = np.array(x)
12     likelihood = (pi*(p**x)*((1-p)**(1-x)))+(1-pi)*(q**x)*((1-q)**(1-x))
13     loglikelihood = np.log(likelihood).sum()
14     # for x_i in x:
15     #     loglikelihood += np.log(pi*(p**x_i)*((1-p)**(1-x_i))+(1-pi)*(q**x_i)*((1-q)**(1-x_i)))
16     return loglikelihood
17
18
19 # x表示结果序列; t表示循环迭代次数; initial_theta_list表示起始参数,初始值都为0.5
20 def EM_algorithm(x, t, initial_theta_list=[0.5, 0.5, 0.5]):
21     [pi, p, q] = initial_theta_list
22     loglikelihood = get_loglikelihood([pi, p, q], x)
23     result_list = [[pi, p, q, loglikelihood]]
24     for i in range(t):
25         mu = []
26         for x_i in x:
27             mu.append((pi*p**x_i*(1-p)**(1-x_i))/(pi*(p**x_i)*
28                 ((1-p)**(1-x_i))+(1-pi)*(q**x_i)*((1-q)**(1-x_i))))
29         pi = sum(mu)/len(mu)
30         p = sum([mu[i]*x[i] for i in range(len(x))]) / sum(mu)
31         q = sum([(1-mu[i])*x[i] for i in range(len(x))]) / \
32             sum([1-mu[i] for i in range(len(mu))])
33         loglikelihood = get_loglikelihood([pi, p, q], x)
34         result_list.append([pi, p, q, loglikelihood])
35         if result_list[i+1] == result_list[i]:
36             break
37     return result_list, pi, p, q, loglikelihood
38
39
40 def draw_plot(result_list):
41     data = pd.DataFrame(result_list, columns=["pi", "p", "q", "loglikelihood"])
42     fig = plt.figure(figsize=(12, 6))
43     ax1=fig.add_subplot(121)
44     plt.title('EM算法参数收敛情况')
45     plt.xlabel('迭代次数')
46     plt.ylabel('参数值')
47     plt.margins(x=0)
48     sns.lineplot(data=data.iloc[:,0:3])
49     ax2=fig.add_subplot(122)
50     plt.title('对数似然值变化情况')
51     plt.xlabel('迭代次数')
52     plt.ylabel('log(likelihood)')
53     plt.margins(x=0)
54     sns.lineplot(data=data["loglikelihood"])
55     plt.savefig("pic.png", dpi=300)
56
```

```

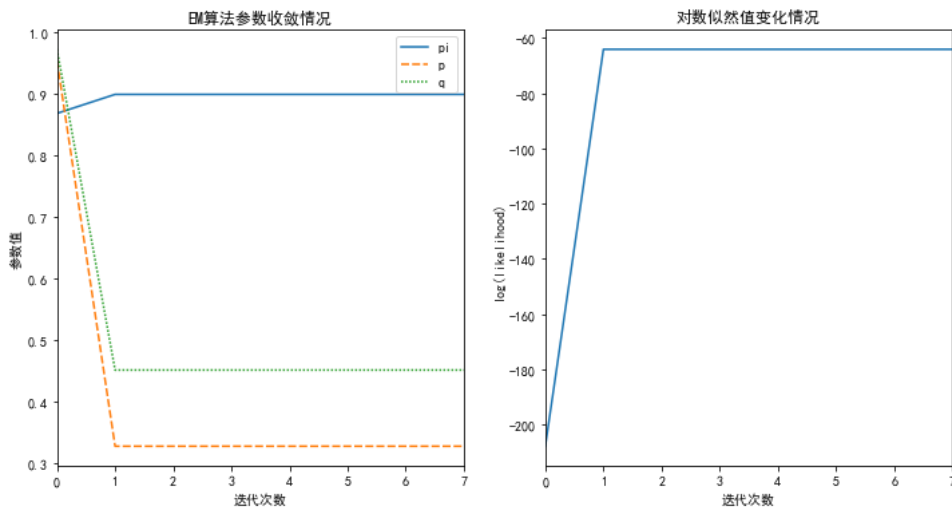
57
58 for i in range(10):
59     initial_theta_list = np.random.rand(3)
60     result_list, pi, p, q, loglikelihood = EM_algorithm(
61         x_list, 100, initial_theta_list=initial_theta_list)
62     print(pi, p, q, loglikelihood)
63 print("EM算法估计的各硬币正面向上的概率分别是: ", (pi, p, q))
64 draw_plot(result_list)
65

```

```

1  0.37791574858448046 0.741185877323549 0.09627985951218022 -64.10354778811555
2  0.8794275954353623 0.2739006813969091 0.8221133411822208 -64.10354778811556
3  0.5186876268611077 0.23051970499195382 0.45798174652242046 -64.10354778811556
4  0.533677915017902 0.03893590769298998 0.6845499628766888 -64.10354778811558
5  0.3676406814907094 0.09656613275127825 0.48152743580060614 -64.10354778811555
6  0.8785501870434601 0.26983609645585255 0.8475554180101406 -64.10354778811556
7  0.1395277206185737 0.07204563246756793 0.3834494672605473 -64.10354778811556
8  0.6373205893289423 0.48732703584085674 0.08110867464891316 -64.10354778811556
9  0.07238132728958667 0.6375894792064936 0.3167793360288392 -64.10354778811555
10 0.8988653112771975 0.3274905536772214 0.4511815095772829 -64.10354778811555
11 EM算法估计的各硬币正面向上的概率分别是: (0.8988653112771975, 0.3274905536772214, 0.4511815095772829)

```



2.2. 推导高斯混合模型的EM算法

多元高斯（正态）分布

$$P(\mathbf{x} | \Sigma, \mu) = \frac{1}{(2\pi)^{\frac{m}{2}} |\Sigma|^{\frac{1}{2}}} \exp \left\{ -\frac{1}{2} (\mathbf{x} - \mu)^{\top} \Sigma^{-1} (\mathbf{x} - \mu) \right\} \quad (21)$$

其中 \mathbf{x} 是 m 维随机变量, μ 是 m 维均值向量, Σ 是 $m \times m$ 的协方差矩阵

高斯混合模型对数似然函数

高斯混合模型是由一系列高斯分布按照参数 α_k 组成的混合模型, 其中 $\alpha_k \geq 0$ 且 $\sum_{k=1}^K \alpha_k = 1$ 。

观测数据 $X = \{\mathbf{x}_1, \dots, \mathbf{x}_n\}$ 的似然函数为:

$$\begin{aligned}
 L(\theta) &= P(X) = \prod_{i=1}^n P(\mathbf{x}_i) \\
 &= \prod_{i=1}^n \left(\sum_{z_i \in \{1, \dots, K\}} P(\mathbf{x}_i, z_i) \right) \\
 &= \prod_{i=1}^n \left(\sum_{z_i \in \{1, \dots, K\}} P(\mathbf{x}_i | z_i) P(z_i) \right) \\
 &= \prod_{i=1}^n \left(\sum_{k=1}^K \alpha_k P(\mathbf{x}_i | \mu_k, \Sigma_k) \right)
 \end{aligned} \quad (22)$$

对数似然函数为

$$LL(\theta) = \sum_{i=1}^n \log \left(\sum_{k=1}^K \alpha_k P(\mathbf{x}_i | \mu_k, \Sigma_k) \right) \quad (23)$$

E(expectation)步:

计算隐变量 z_i 的分布 $Q(z_i)$, 即 z_i 的后验分布 $P(z_i = k | \mathbf{x}_i)$

$$\begin{aligned}
 Q(z_i = k) &= P(z_i = k | \mathbf{x}_i, \alpha^{(t)}) \\
 &= \frac{P(\mathbf{x}_i | z_i = k, \alpha^{(t)}) P(z_i = k | \alpha^{(t)})}{P(\mathbf{x}_i | \alpha^{(t)})}
 \end{aligned}$$

$$\begin{aligned}
&= \frac{P(\mathbf{x}_i | z_i = k, \alpha^{(t)}) P(z_i = k | \alpha^{(t)})}{\sum_{k=1}^K P(\mathbf{x}_i | z_i = k, \alpha^{(t)}) P(z_i = k | \alpha^{(t)})} \\
&= \frac{\alpha_k P(\mathbf{x}_i | \boldsymbol{\mu}_k, \boldsymbol{\Sigma}_k)}{\sum_{k=1}^K \alpha_k P(\mathbf{x}_i | \boldsymbol{\mu}_k, \boldsymbol{\Sigma}_k)} \\
&= \gamma_{ik}^{(t)}
\end{aligned} \tag{24}$$

且有 $\sum_{k=1}^K \gamma_{ik}^{(t)} = 1$

M(maximize)步:

参数 $\theta = (\alpha, \boldsymbol{\mu}, \boldsymbol{\Sigma})$ 的更新公式为:

$$\begin{aligned}
\theta^* &= \arg \max_{\theta} \sum_{i=1}^n \sum_{k=1}^K \gamma_{ik} \log (P(\mathbf{x}_i | z_i = k, \theta) P(z_i = k | \theta)) \\
&= \arg \max_{\theta} \sum_{i=1}^n \sum_{k=1}^K \gamma_{ik} \{ \log P(\mathbf{x}_i | \boldsymbol{\mu}_k, \boldsymbol{\Sigma}_k) + \log \alpha_k \} \\
&= \arg \max_{\alpha_k, \boldsymbol{\mu}_k, \boldsymbol{\Sigma}_k} \sum_{i=1}^n \sum_{k=1}^K \gamma_{ik} \left\{ \log \left\{ \frac{1}{(2\pi)^{\frac{n}{2}} |\boldsymbol{\Sigma}_k|^{\frac{1}{2}}} \exp \left\{ -\frac{1}{2} (\mathbf{x}_i - \boldsymbol{\mu}_k)^{\top} \boldsymbol{\Sigma}_k^{-1} (\mathbf{x}_i - \boldsymbol{\mu}_k) \right\} \right\} + \log \alpha_k \right\} \\
&= \arg \min_{\alpha_k, \boldsymbol{\mu}_k, \boldsymbol{\Sigma}_k} \sum_{i=1}^n \sum_{k=1}^K \gamma_{ik} \left\{ \frac{1}{2} \log |\boldsymbol{\Sigma}_k| + \frac{1}{2} (\mathbf{x}_i - \boldsymbol{\mu}_k)^{\top} \boldsymbol{\Sigma}_k^{-1} (\mathbf{x}_i - \boldsymbol{\mu}_k) - \log \alpha_k \right\} \\
&\triangleq \arg \min_{\theta} f(\alpha, \boldsymbol{\mu}, \boldsymbol{\Sigma})
\end{aligned} \tag{25}$$

将 $f(\cdot)$ 分别对 $\boldsymbol{\mu}, \boldsymbol{\Sigma}$ 求偏导, 当 $\sum_{i=1}^n \gamma_{ik}^{(t)} \neq 0$ 对所有 $k = 1, 2, \dots, K$ 成立时:

$$\begin{cases} \frac{\partial f(\alpha, \boldsymbol{\mu}, \boldsymbol{\Sigma})}{\partial \boldsymbol{\mu}_k} = - \sum_{i=1}^n \gamma_{ik}^{(t)} (\mathbf{x}_i - \boldsymbol{\mu}_k) \boldsymbol{\Sigma}_k^{-1} \equiv 0 & \implies \boldsymbol{\Sigma}_k^{-1} \sum_{i=1}^n \gamma_{ik}^{(t)} \mathbf{x}_i = \boldsymbol{\mu}_k \boldsymbol{\Sigma}_k^{-1} \sum_{i=1}^n \gamma_{ik}^{(t)} & \implies \boldsymbol{\mu}_k = \\ \frac{\partial f(\alpha, \boldsymbol{\mu}, \boldsymbol{\Sigma})}{\partial \boldsymbol{\Sigma}_k} = \sum_{i=1}^n \left(\boldsymbol{\Sigma}_k^{-1} - (\mathbf{x}_i - \boldsymbol{\mu}_k)(\mathbf{x}_i - \boldsymbol{\mu}_k)^{\top} \boldsymbol{\Sigma}_k^{-2} \right) \gamma_{ik} \equiv 0 & \implies \boldsymbol{\Sigma}_k \sum_{i=1}^n \gamma_{ik}^{(t)} = \sum_{i=1}^n \gamma_{ik}^{(t)} (\mathbf{x}_i - \boldsymbol{\mu}_k)(\mathbf{x}_i - \boldsymbol{\mu}_k)^{\top} & \implies \boldsymbol{\Sigma}_k = \end{cases}$$

下面计算 α_k 的参数迭代过程:

$$\begin{aligned}
\alpha_k &= \arg \min_{\alpha} \sum_{i=1}^n \sum_{k=1}^K (-\log \alpha_k) \gamma_{ik} \\
&= \arg \max_{\alpha} \sum_{i=1}^n \sum_{k=1}^K \log \alpha_k \gamma_{ik}
\end{aligned} \tag{27}$$

由约束 $\sum_{k=1}^K \alpha_k = 1$ 构建拉格朗日对偶函数:

$$L(\alpha, \lambda) = \sum_{i=1}^n \sum_{k=1}^K \log \alpha_k \gamma_{ik} + \lambda \left(\sum_{k=1}^K \alpha_k - 1 \right) \tag{28}$$

将 $L(\alpha, \lambda)$ 对 α_k 求偏导:

$$\frac{\partial L(\alpha, \lambda)}{\partial \alpha_k} = \sum_{i=1}^n \frac{1}{\alpha_k} \gamma_{ik} + \lambda \equiv 0 \implies \alpha_k = \frac{\sum_{i=1}^n \gamma_{ik}}{\lambda} \tag{29}$$

由 $\sum_{k=1}^K \gamma_{ik}^{(t)} = 1$, 将上式代入得:

$$\sum_{i=1}^n \sum_{k=1}^K \gamma_{ik}^{(t)} = \lambda \sum_{k=1}^K \alpha_k = n \tag{30}$$

所以得到 $\lambda = n$, 从而得到:

$$\alpha_k = \frac{\sum_{i=1}^n \gamma_{ik}}{n} \tag{31}$$

当 $\sum_{i=1}^n \gamma_{ik}^{(t)} = 0$ 对 k 成立时, $\boldsymbol{\mu}_k, \boldsymbol{\Sigma}_k$ 的取值对结果无影响。

3. 高斯混合模型EM算法代码实现

按下列参数生成高斯混合模型的数据, 共有三个高斯混合成分, 每个成分、生成 300 个数据点。

$$\begin{aligned}
\boldsymbol{\mu}_1 &= (3, 1) & \boldsymbol{\Sigma}_1 &= ((1, -0.5); (-0.5, 1)) \\
\boldsymbol{\mu}_2 &= (8, 10) & \boldsymbol{\Sigma}_2 &= ((2, 0.8); (0.8, 2)) \\
\boldsymbol{\mu}_3 &= (12, 2) & \boldsymbol{\Sigma}_3 &= ((1, 0); (0, 1))
\end{aligned} \tag{32}$$

使用 scikit-learn 库中的高斯混合模型实现上述数据的学习过程, 计算在不同个数的高斯混合成分下模型的 AIC 和 BIC 值, 并将学习得到的模型参数与真实模型参数进行对比。

1 3.1 根据参数生成数据点

```

1 import numpy as np
2 from numpy.random import multivariate_normal
3 import matplotlib.pyplot as plt
4 def make_data(len,mu,sigma): #alpha:隐变量权重(K);mu多元高斯分布均值(n*K);sigma多元高斯分布协方差矩阵
    (K*k),len表示生成点数量

```

```

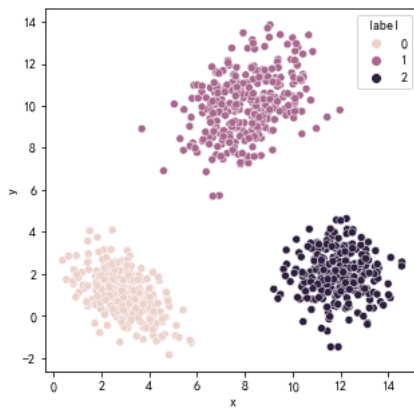
5     data = pd.DataFrame()
6     K = mu.shape[0]
7     for k in range(K):
8         x_k = multivariate_normal(mean = mu[k], cov = sigma[k], size=(len), check_valid="raise")
9         points = pd.DataFrame(x_k, columns = ["x", "y"])
10        points["label"] = k
11        data = data.append(points)
12        result = np.array(data.iloc[:,0:2])
13    return result, data
14 mu = np.array([[3,1],[8,10],[12,2]])
15 sigma = np.array([[1,-0.5],[-0.5,1]],[[2,0.8],[0.8,2]],[[1,0],[0,1]])
16 result, data = make_data(300, mu, sigma) #生成300个数据点
17 plt.figure(figsize=(5,5))
18 sns.scatterplot(x="x", y="y", data = data, hue = "label")

```

```

1 | <AxesSubplot: xlabel='x', ylabel='y'>

```



```

1 from sklearn.mixture import GaussianMixture as GMM
2 AIC_list = []
3 BIC_list = []
4 plt.figure(figsize=(18,12))
5 for k in range(1, 9):
6     gmm = GMM(n_components=k).fit(result)
7     labels = gmm.predict(result)
8     AIC_list.append(gmm.aic(result))
9     BIC_list.append(gmm.bic(result))
10    plt.subplot(2, 4, k)
11    plt.scatter(result[:, 0], result[:, 1], c=labels, s=4)
12    plt.title("K = {}".format(k))
13    plt.savefig("Gaussian_result.png", dpi=300)
14 plt.figure(figsize=(8,6))
15 plt.plot(range(1,9), AIC_list, label = "AIC")
16 plt.plot(range(1,9), BIC_list, label = "BIC")
17 plt.legend()
18 plt.savefig("AIC-BIC.png", dpi=100)

```

