

A gene-based exchanged XGBoost method for detecting and ranking gene-gene interactions of qualitative trait

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Abstract

Among the large of number of statistical methods that have been proposed to identify gene-gene interaction in case-control genome-wide association studies (GWAS), gene-based methods have recently grown in popularity as they confer advantage in both statistical power and biological interpretability. All of the gene-based methods either jointly model the distribution of single nucleotide polymorphism (SNPs) sets prior to the statistical test, or specify an explicit function form of traits and SNPs, leading to a limited power to detect XXX. In this paper, we instead proposed a gene-based method that first applies XGBoost, a popular and highly effective boosted tree methods in machine learning for modeling additive models with non-linear terms, to model the function of qualitative trait with all the genes to be considered, and then form a subsample with an exchange strategy to evaluate the interaction of each pair of genes as a deviation from the multiplicative structure in the predicted probabilities. We use simulations to assess the capacity of geXGB. The benefits of our approach in terms of statistical power and robustness of genes are evaluated in pure and strict disease models with a wide range of heritability and MAF by comparing it to previous methods. We also apply our method to the gene pairs associated to rheumatoid arthritis(RA)in the Wellcome Trust Case Control Consortium(WTCCC)dataset according to the RA pathway hsa05323 in KEGG.

1 Introduction

Genome-wide association studies (GWAS) have identified over six thousand single-nucleotide polymorphisms (SNPs) associated with complex diseases or traits. Earlier GWAS analysis strategies were largely based on single locus models, which test the association between individual markers and a given phenotype independently. Although this type of approaches have successfully identified many regions of disease susceptibility, most of these SNPs identified have small effect sizes which failed to fully account for the heritability of complex traits. Genetic interaction has been hypothesized to play an important role in the genetic

basis of complex diseases and traits and to be one of the possible solutions to this problem of “missing heritability”. Even if genetic interaction explains only a tiny fraction of “missing heritability”, they can still provide some biological insight on the pathway level through by aiding the construction of novel gene pathway topologies.

The first investigations on genetic interactions have been at the SNP level, in which various statistical methods, including logic and logistic regression, odds-ratio, linkage disequilibrium(LD) and entropy-based statistic, are employed to detect SNP-SNP interactions (i.e. epistasis). Other techniques that have been used to study SNP-SNP interactions include multifactor dimensionality reduction, Tuning Relief, Random Jungle, BEAM, BOOST[1] and pRF[2]. These marker-based methods may encounter some common challenges, such as the complexity arising from the large number of pairwise or higher-order tests because all pairs or groups of SNPs have to be considered; and the extensive burden of correction they entail due to multiple testing. In this paper, we aim to improve the power of gene-gene interaction detection by moving beyond SNP level, and instead consider all potential pairs of SNPs from each of a pair of genes in a single gene-based interaction detection.

Gene-based approaches have been successful for regular GWAS tests of main (marginal) associations, and there are several potential advantages in extending this methodology to gene-gene interaction detections. Firstly, a gene-based approach can substantially reduce the number of tests needed. For example, for 20,000 genes, there are $\sim 2 \times 10^8$ possible pairwise gene-based interactions to be tested, while for 3 million SNPs there are over $\sim 5 \times 10^{12}$ possible marker-based interactions to be tested. Secondly, a gene-based interaction test may have greater power, because when there are multiple interactions between features in the targeted genes (or other kind of regions), the effect of these interactions may be aggregated by the algorithm. Such aggregation has already been seen in gene-based GWAS tests for main association effect. Thirdly, a gene-based approach may be better at leveraging prior biological knowledge, which is often on the level of genes. For example, one may test pairs of genes that exhibit protein-protein interactions (PPI) or that participate in the same pathways.

In the work of Peng et al [3], canonical correlation analysis between two genes is done on both the case and the control group, and a U-statistic, called CCU, is used to measure the difference of the correlation between these two genes, which is used to indicate the presence of interaction. A limitation of this method is that in the correlation analysis only linear relations are considered. To overcome this limitation, [4, 5] extended CCU to KCCU, where the canonical correlation analysis is kernelized to account for possible non-linearity. Li et al. [6] introduced another method called GBIGM which is entropy-based and non-parametric, which was based on an entropy-based non-parametric. More recently, Emily [7] developed a new method called AGGrGATOR which combines the p-values in marker-level interaction tests to measure the interaction between two genes. Earlier[8] this

strategy was successfully used for the interaction detection for quantitative phenotypes.

In this paper, rather than designing a new dedicated statistic, we apply a machine learning algorithm extreme gradient boost (XGBoost [9]) to propose a new approach, called gene-based exchanged extreme gradient boost (geXGB), to detect gene-gene interaction. The idea is to test the XGBoost model on a test dataset obtained from a exchange strategy in order to see how far it deviate from an additive form. Our idea has some similarity with [10]. Our method does not require explicit modeling of interacting terms and allow any kind of the functional form that interaction might take. An advantage of geXGB is that it is nonparametric, and hence may be more flexible for data-driven exploratory genome-wide association studies.

2 Materials and Methods

In this section we first detail the geXGB approach. Then we describe the various simulation studies conducted to assess the statistical power of our approach in gene-gene interaction detection. Finally, we apply our approach to the WTCCC dataset to evaluate our approach in a real-lift situation.

2.1 Overview of geXGB

Our method, gene-based exchange eXtreme Gradient Boost(geXGB), is a machine learning based procedure for detecting the interaction between two genes in susceptibility with a binary phenotype, typically a case/control disease status. Let random variable $y \in \{0, 1\}$ be the phenotype, where $y = 0$ stands for membership of the control group and $y = 1$ for membership of the case group. Let X_g , where $g = 1, \dots, G$ be the genotype of the G genes in our gene list, each a collection of m_g SNP markers, i.e. m_g discrete features that may take on a value of 0, 1 or 2 corresponds to the number of minor alleles at each locus for each observation.

Let $F^*(\cdot)$ be an unknown target function of interaction between genotype X and phenotype y , and Let $F(\cdot)$ be a highly accurate model of $F^*(\cdot)$ that can be learned from a given set of training data. We assume that the function of interaction should have an additive structure and we are interested in evaluating only whether a proper subset of features contribute additively to the response y . Suppose that our training set consists of X_g and X'_g genes and other genes $X_{\neg g, g'}$. We define X_g and X'_g to be *Non-interacting*, if and only if there are two functions F_1 and F_2 , so that

$$F^*(X_g, X'_g, X_{\neg g, g'}) = F_1(X_g, X_{\neg g, g'}) + F_2(X'_g, X_{\neg g, g'}) \quad (1)$$

And we use the size of deviation of $F^*(\cdot)$ from partial additivity as a means of identifying the interactions between X_g and X'_g required to reconstruct some percentage of the variation

in the values of $F^*(\cdot)$. Here, since the qualitative trait is binary that we can only get the probability of y to be 0 or 1, so we let the *log* of the probability of y being 1 be the target function above and hence get the zero hypothesis as follows:

$$P(y = 1|X_g, X'_g, X_{\neg g}, g') = F_1(X_g, X_{\neg g}, g') * F_2(X'_g, X_{\neg g}, g') \quad (2)$$

Any machine learning algorithms that deal with classification problems can be used for the above scheme. Here, We choose eXtreme Gradient Boost (XGBoost) as our classifier because gradient boosting decision tree (GBDT) is an effective and relatively model-agnostic way to approximate true target function which may have additive structure with *non-linear* terms. Chipman et al performed an extensive comparison of several algorithms on 42 data sets, in which GBDT showed performance similar to or better than Random Forests and a number of other types of models. XGBoost is an algorithm which improves upon GBDT for its computational efficiency with roughly the same error rate. The fact that XGBoost is generally accurate and fast makes it an excellent tool.

Our approach consists of two steps: 1) training 2) testing and ranking. Before training, We firstly start by training an XGBoost model with all the genes in the gene list and use cross-validation to choose the best parameter combination of the model. Then, with the selected parameter combination, for each selected pair of genes, we use our exchange strategies to generate a test dataset. Lastly, we calculate the predicted probability of our model on the test dataset and measure the strength of interaction by evaluating how much the prediction deviate from Equation (2). The various steps of the geXGB framework are illustrated in Figure 1.

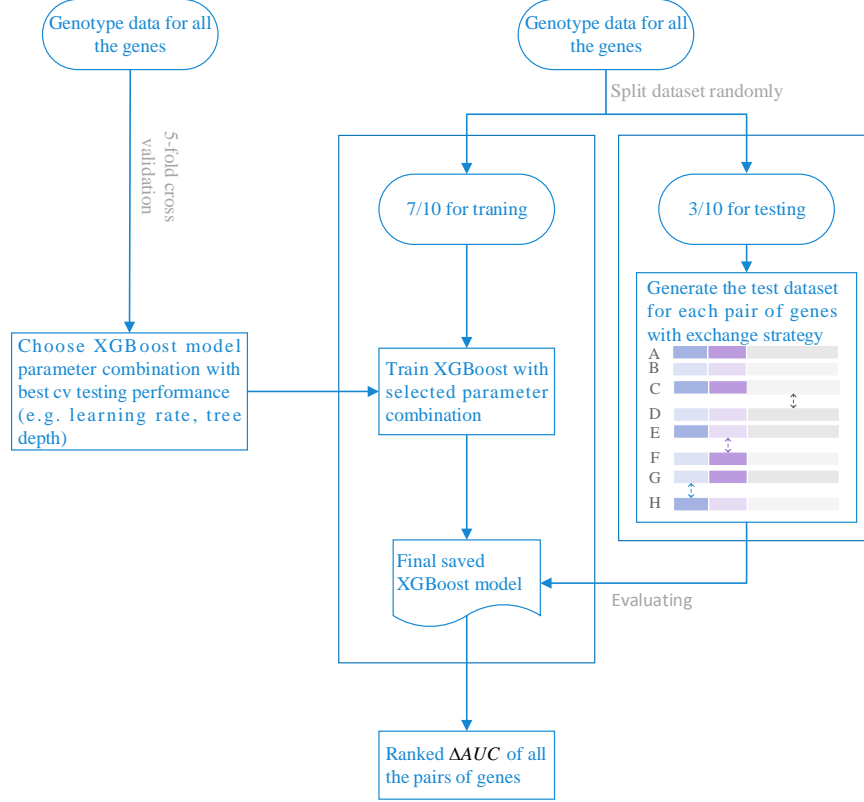


Figure 1: The framework of geXGB

2.1.1 Overview of XGBoost

XGBoost [9] is a scalable supervised machine learning system based on tree boosting, and recently has been dominating applied machine learning as well as in Kaggle competitions. It is an algorithm which improves GBDT for speed with the same performance. In this section, we use the standard Machine Learning notation and let x_i be features and y be a binary random variable which we attempt to predict.

2.1.1.1 Ensemble of CARTs

In this ensemble model, the base classifier is CART (Classifying And Regression Tree), which is similar to decision trees, but on each leaf, instead of a classification, a real-valued score is assigned. This makes ensemble training easier and may also provide more information beyond classification.

Let \mathcal{F} be the space of functions that can be represented by CARTs, the ensemble predictor is $\hat{y} = \sum_k f_k, f_k \in \mathcal{F}$. In our case we interpret it as in logistic regression, namely

$$p(y = 1|x) = \frac{1}{1 + e^{-\hat{y}(x)}} \quad (3)$$

Hence, the learning objective is

$$obj = \sum_i (l(y_i, \hat{y}(x_i)) + \sum_k \Omega(f_k) \quad (4)$$

Where $l(y, \hat{y}) = y \log(1 + e^{-\hat{y}}) + (1 - y) \log(1 + e^{\hat{y}})$ is the logistic regression loss function, and $\Omega(f_k)$ is the regularizer.

2.1.1.2 Gradient Boosting

It is not feasible to train all the trees in the ensemble together at once because it is hard to calculate the gradient as which is needed in traditional optimization methods. Instead, XGBoost use an additive training strategy: fix the trees have already learned, add new trees one at a time. Let $\hat{y}^{(t)}$ be the predictor at iteration t , then

$$\hat{y}^{(0)} = 0 \quad (5)$$

$$\hat{y}^{(t)} = \hat{y}^{(t-1)} + f_t \quad (6)$$

Where $f_t \in \mathcal{F}$ optimizes the following target function, which is obtained by the Taylor expansion of the lost function for logistic regression to the second order.

$$obj^t = \sum_i \left(g_i(\hat{y}^{(t-1)}(x_i)) f_t(x_i) + \frac{h_i^2(\hat{y}^{(t-1)}(x_i))}{2} \cdot f_t^2(x_i) \right) + \Omega(f_i) \quad (7)$$

Here $g_i(\hat{y}) = \frac{d}{d\hat{y}} l(y_i, \hat{y})$, $h_i(\hat{y}) = \frac{d^2}{d\hat{y}^2} l(y_i, \hat{y})$.

2.1.1.3 Regularizer and Training strategy for CARTs

For any $f \in \mathcal{F}$, let T be the number of leaves in the tree representing f , w_1, \dots, w_T be the scores on the leaves. Then the regularizer used in XGBoost is

$$\Omega(f) = \gamma T + \frac{1}{2} \lambda \sum_j w_j^2 \quad (8)$$

The purpose of the second term is that it can smoothen the leaf scores.

To optimize f_t , firstly note that given a tree structure, obj^t is a quadratic function of the scores w_j , and the minimum of obj^t as well as the w_j that minimizes obj^t can be easily calculated given the tree structure. Now the tree can be constructed by a greedy algorithm in which one starts with a tree with one single node, and repeatedly split its leaves in a way that maximizes the decrease in obj^t in each step.

2.1.2 Evaluating and Ranking

Our approach for gene-based gene-gene interaction detection is based on evaluating the extend a model trained with XGBoost (c.f. section 2.2) deviate from the “product form” in Equation (2). Our interaction estimation technique is based on the following observation: if Equation (2) is satisfied for X_g, X'_g , and let $\mathcal{P}(X_g, X'_g, X_{\neg g}, g') := P(y = 1 | X_g, X'_g, X_{\neg g}, g')$, $X_{\neg g}, g'$ be the vector consisting of genotypes that are in neither X_g nor X'_g , then, for two samples $(X^A), (X^B)$, we have

$$\begin{aligned} & \mathcal{P}(X_g^A, X_{g'}^A, X_{g,g'}^A) \mathcal{P}(X_g^B, X_{g'}^B, X_{g,g'}^B) \mathcal{P}(X_g^A, X_{g'}^A, X_{g,g'}^B) \mathcal{P}(X_g^B, X_{g'}^B, X_{g,g'}^A) \\ &= \mathcal{P}(X_g^A, X_{g'}^B, X_{g,g'}^A) \mathcal{P}(X_g^B, X_{g'}^A, X_{g,g'}^B) \mathcal{P}(X_g^A, X_{g'}^B, X_{g,g'}^B) \mathcal{P}(X_g^B, X_{g'}^A, X_{g,g'}^A) \end{aligned} \quad (9)$$

To verify that, note that if $\mathcal{P}(X_g, X_{g'}, X_{\neg g}, g') = F_1(X_g, X_{\neg g}, g') * F_2(X_{g'}, X_{\neg g}, g')$, then the left-hand-side of equation (9) above becomes:

$$\begin{aligned} & F_1(X_g^A, X_{g,g'}^A) F_2(X_{g'}^A, X_{g,g'}^A) F_1(X_g^B, X_{g,g'}^B) F_2(X_{g'}^B, X_{g,g'}^B) F_1(X_g^A, X_{g,g'}^B) F_2(X_{g'}^A, X_{g,g'}^B) \\ & F_1(X_g^B, X_{g,g'}^A) F_2(X_{g'}^B, X_{g,g'}^A) \end{aligned}$$

while the right-hand-side becomes:

$$\begin{aligned} & F_1(X_g^A, X_{g,g'}^A) F_2(X_{g'}^B, X_{g,g'}^A) F_1(X_g^B, X_{g,g'}^B) F_2(X_{g'}^A, X_{g,g'}^B) F_1(X_g^A, X_{g,g'}^B) F_2(X_{g'}^B, X_{g,g'}^B) \\ & F_1(X_g^B, X_{g,g'}^A) F_2(X_{g'}^A, X_{g,g'}^A). \end{aligned}$$

It is evident that these two are the same.

This observation is motivated by the well-known fact that a function $f(X, Y)$ is of the form $f(X, Y) = u(X)v(Y)$ if and only if for any a, b, c, d , $f(a, b)f(c, d) = f(a, d)f(c, b)$.

As the function \mathcal{P} is unknown, we use the predicted probability of XGBoost as an estimator of it. After model selection, we evaluated the gene interaction with equation (9). We want the distribution of the samples tested to be close to the samples we used to train the model to minimize error caused by interpolation. On the other hand, we don't want them to be completely identical because the learned probability may ill-behave at those places (c.f. [10]). Hence, we randomly split the available sample set into two, using the

7/10 of the samples to train our model with the selected model parameters and the rest 3/10 to generate the test dataset for evaluating equation (9). The process of using exchange strategy to get the test dataset was shown in Figure 2.

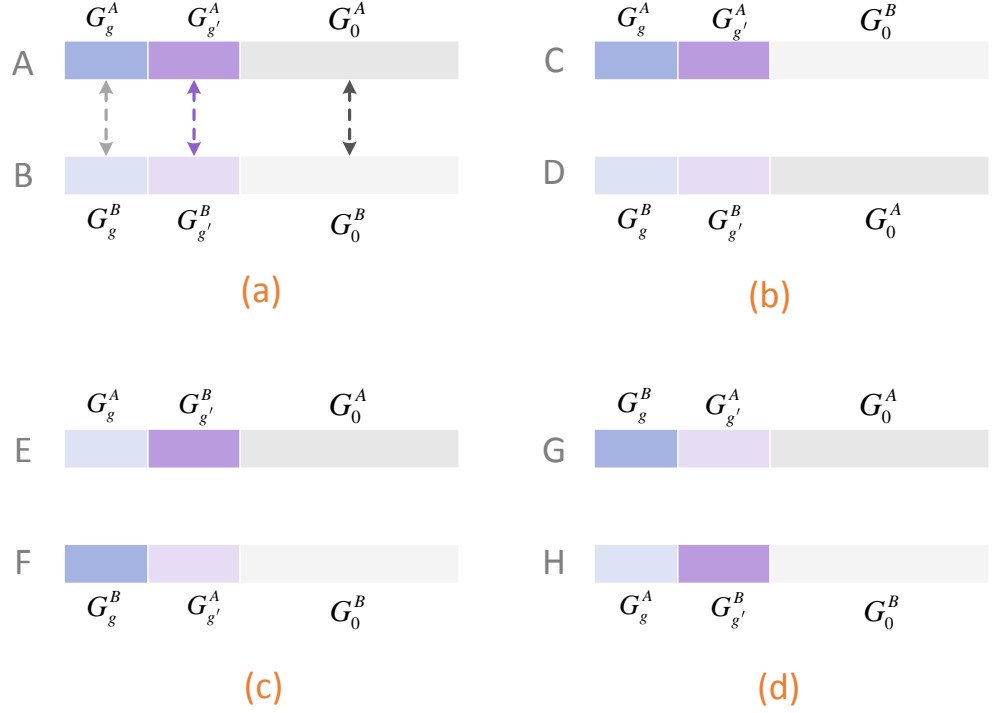


Figure 2: Illustration of exchange strategy of getting the test dataset

Let $C = \{X^{C^i}\}$ be a set of genotypes, which has the same genotype as A on all genes except the g and the g' -th, and the same genotype as B on the g and the g' -th.
Let $D = \{X^{D^i}\}$ be a set of genotypes, which has the same genotype as B on all genes except the g and the g' -th, and the same genotype as A on the g and the g' -th.
Let $E = \{X^{E^i}\}$ be a set of genotypes, which has the same genotype as A on all genes except the g -th, and the same genotype as B on the g -th.

Let $F = \{X^{F^i}\}$ be a set of genotypes, which has the same genotype as A on all genes except the g' -th, and the same genotype as B on the g' -th.

Let $G = \{X^{G^i}\}$ be a set of genotypes, which has the same genotype as B on all genes except the g -th, and the same genotype as A on the g -th.

Let $H = \{X^{H^i}\}$ be a set of genotypes, which has the same genotype as B on all genes except the g' -th, and the same genotype as A on the g' -th.

From Figure 2 we can see that A,B was two original samples from our test dataset. Based on A and B , we get C,D through exchanging the part $X_{\neg g,g'}$. A,B,C and D were used to calculate the left-hand of equation (9). While through exchange $X_{g'}$ and X_g from A,B respectively, we can get E,F and G,H , which were used to estimate the right-hand of equation(9). Finally, we calculated the difference called Δ_{inter} between predictions of A,B,C,D and E,F,G,H for each pairs of genes. Δ_{inter} can be treated as an indicator of the strength of gene-gene interaction that we ranked Δ_{inter} for all the gene pairs and return the potential gene pairs that may interact.

Hence, with the considerations above, we propose our ranking algorithm as follows:

Algorithm 1: geXGB

input : genotype dataset
 $S = \{(x_1, y_1), (x_2, y_2) \dots (x_n, y_n)\}, x_i \in \{0, 1, 2\}^{(m_1, \dots, m_G)}, y_i \in \{0, 1\}$; gene list file with position information for the G genes; buffer region size.

output: A list of all pairs of genes sorted by Δ_{inter} .

Train XGBoost model, using grid search to find the proper parameter combination of the XGBoost, using 5-fold cross validation for each parameter combination and select the best parameter combination that gives the best average predictive performance.

for $i = 1, \dots, N$ **do**
Divide dataset S randomly into a training set S_{train} with 7/10 samples and a testing set S_{test} with the rest.
XGBoostModel = trainXGBoost(S_{train} , parameter combination)
Sample the testing dataset to obtain two sets of genotype data $A = \{X^{A^i}\}$ and $B = \{X^{B^i}\}$ of equal size.
for $1 \leq g < g' \leq G$ **do**
Getting the
 $C = \{X^{C^i}\}, D = \{X^{D^i}\}, E = \{X^{E^i}\}, F = \{X^{F^i}\}, G = \{X^{G^i}\}, H = \{X^{H^i}\}$
according to the process shown in Figure2.
 $\Delta_{inter_{g,g'}} =$
 $\sum_i (Predict(XGBoostModel, X^{A^i})Predict(XGBoostModel, X^{B^i})$
 $Predict(XGBoostModel, X^{C^i})Predict(XGBoostModel, X^{D^i}) -$
 $Predict(XGBoostModel, X^{E^i})Predict(XGBoostModel, X^{F^i})$
 $Predict(XGBoostModel, X^{G^i})Predict(XGBoostModel, X^{H^i}))$
end
end
Return C_G^2 pairs of genes sorted by the total $\Delta_{inter_{g,g'}}$ in all N -iterations in decreasing order.

Because the first step is time consuming, in the simulated study, we did model selection for only one sample set and used the resulted parameters for the entire experiment.

2.2 Simulation study

The goal of this simulation study is to evaluate the performance of geXGB procedure for gene-gene interaction detection. All simulated datasets were set to have 50 SNPs. Among them 2 SNPs were functional and the remaining 48 SNPs were non-functional. The 50 SNPs formed 5 genes, each had 10 SNPs. The 2 functional SNPs were put into the first and second gene, and the performance is measured by how likely our algorithm can rank the two interacting genes as the most significant. We chose the publicly available tool GAMETES [12] to generate the simulated genotype data. This tool is designed to generate

epistasis models that we refer to as pure and strict. Purely and strictly epistasis models constitute the most difficult type of disease association model to detect, as such associations may be observed only if all n -loci are included in the disease model. This requirement makes these types of models an attractive gold standard for simulation studies of complex multi-locus effects.

In this simulation study, to test the effects of heritability (which measures the strength of correlation between genotype and phenotype) and sample size, we performed experiments under two different scenarios. In the first scenario, we tested two-locus epistasis models with five different heritabilities (0.01, 0.025, 0.05, 0.1 and 0.2) and two different minor allele frequencies (MAF, 0.2 and 0.4) with prevalence set to be 0.2 and sample size to be 3000. Ten models for each of the 10 heritability-allele frequency combinations were generated, so that we had 100 models in total in accordance to Hardy-Weinberg proportions. For a specified genetic constrain combination, the 10 models were roughly sorted by ascending customized odds ratio (COR) used in GAMETES. COR was a metrics that calculated directly from the genetic model, and demonstrated a given model's delectability. From M1 to M10, it became much easier to detect the gene-gene interaction. The penetrance tables were generated for these 100 models in the absence of main effect. One hundred replicated data sets were generated from each model with balanced cases and controls, resulting in 10000 data sets in total in this scenario. In the second scenario, we set heritability to be 0.025 and MAF to be 0.2 and 0.4, prevalence to be 0.2 with sample size 10000. Then, 100 data sets were generated by random sampling from this large dataset for each of the 5 sample sizes 1000, 2000, 3000, 4000 and 5000. In this scenario, we have 1000 datasets in total.

2.3 Real data analysis

To assess the capacity of geXGB to deal with real case-control phenotype, we investigated the susceptibility of a set of pairs of genes to Rheumatoid Arthritis (RA), a chronic autoimmune joint disease where persistent inflammation affects bone remodeling leading to progressive bone destruction. We used the WTCCC(2007) dataset, which were genotyped in the British population using the Affymetrix GeneChip 500k. Quality control was performed in PLINK with several steps. First we removed samples with reported sex that did not match the heterozygosity rates observed on chromosome X [13]. We additionally filtered out SNPs with $> 10\%$ missingness, with a minor allele frequency (MAF) < 0.05 , or for which missingness was significantly correlated with phenotype ($p < 1 \times 10^{-4}$). We further filter out SNPs that are not in Hardy-Weinberg equilibrium in controls, as well as filter out samples with $> 10\%$ missing SNPs. After the QC steps, we have XX SNPs, XXX samples with XXX cases and XXX controls.

In this analysis, we aim to verify some gene-gene interaction in the RA pathway hsa05323 in KEGG pathway dataset. Genotyping coordinates are given in NCBI Build36/UCSC hg18 (National Center for Biotechnology Information, Bethesda, MD). There are XX genes in the pathway, and we can mapping XX based on Build36 annotation. For each gene, we add 10k to both the upstream and downstream. Gender was included as covariate in the analysis. Principal component analysis was conducted using GCAT[], and top 10 PCs were also included as covariates to account for potential population stratification.

2.4 Competitive methods

The performance of our procedure geXGB was compared to three previously published methods: Kernel Canonical Correlation-based U-statistic analysis (KCCU)[4, 5], the gene-based information gain method (GBIGM)[6] and A Gene-based Gene-Gene interaction test method (AGGrEGATOr)[7]. We adapted them to the task of ranking, by ranking the gene pairs by their p-values in ascending order.

3 Results

3.1 Results for the simulation study

To evaluate the statistical power of our geXGB and other three competitive methods, under each heritability-MAF combination, we measure the performance of geXGB as the percentage that the method ranks the interacting gene pair as the top one of 100 data sets. For all other methods, the number listed is the relative frequency for the single interacting pair to have the smallest p-value.

Table 1: Results of the first scenario of the simulation study

MAF	Herita- bility	Model	M1	M2	M3	M4	M5	M6	M7	M8	M9	M10
		Method										
0.2	0.01	geXGB	0.14	0.17	0.58	0.75	0.48	0.38	0.71	0.91	0.93	0.49
		AGGrEGATOr	0.12	0.14	0.12	0.89	0.12	0.1	0.89	1	0.88	0.34
		KCCU	0.15	0.09	0.09	0.29	0.14	0.1	0.43	0.62	0.52	0.13
		GBIGM	0.09	0.08	0.11	0.13	0.12	0.17	0.11	0.08	0.1	0.09
	0.025	geXGB	0.98	0.97	0.94	1	1	0.99	0.99	1	0.9	0.94
		AGGrEGATOr	1	0.15	0.27	1	1	0.46	0.37	1	0.69	0.81
		KCCU	0.58	0.09	0.09	0.74	0.71	0.59	0.24	0.8	0.12	0.12
		GBIGM	0.08	0.11	0.07	0.11	0.1	0.12	0.13	0.2	0.14	0.1
	0.05	geXGB	1	1	1	1	1	1	1	1	1	1
		AGGrEGATOr	0.09	0.09	0.59	0.89	0.98	0.99	1	1	1	1
		KCCU	0.13	0.1	0.57	0.65	0.71	0.8	0.84	0.85	0.84	0.77
		GBIGM	0.18	0.16	0.08	0.22	0.23	0.12	0.17	0.19	0.12	0.12
	0.1	geXGB	1	1	1	1	1	1	1	1	1	1
		AGGrEGATOr	1	1	1	1	1	1	1	1	1	1
		KCCU	0.81	0.88	0.93	0.9	0.9	0.83	0.86	0.91	0.84	0.93
		GBIGM	0.15	0.14	0.14	0.23	0.19	0.17	0.16	0.16	0.15	0.19
	0.2	geXGB	1	1	1	1	1	1	1	1	1	1
		AGGrEGATOr	1	1	1	1	1	1	1	1	1	1
		KCCU	0.89	0.91	0.97	0.94	0.95	0.92	0.89	0.97	0.94	0.97
		GBIGM	0.19	0.21	0.31	0.18	0.29	0.23	0.22	0.21	0.23	0.22
0.4	0.01	geXGB	0.9	0.74	0.82	0.82	0.83	0.96	0.96	0.66	0.82	0.89
		AGGrEGATOr	0.71	0.73	0.09	0.1	0.77	0.96	0.94	0.17	0.9	0.81
		KCCU	0.34	0.34	0.05	0.08	0.4	0.29	0.77	0.16	0.73	0.61
		GBIGM	0.09	0.11	0.08	0.1	0.11	0.07	0.11	0.18	0.11	0.11
	0.025	geXGB	1	1	1	1	1	1	1	1	1	1
		AGGrEGATOr	0.99	1	0.56	0.12	0.06	0.26	0.91	0.51	0.12	1
		KCCU	0.58	1	0.24	0.08	0.06	0.11	0.24	0.32	0.12	1
		GBIGM	0.15	0.1	0.12	0.14	0.08	0.09	0.11	0.08	0.05	0.08
	0.05	geXGB	1	1	1	1	1	1	1	1	1	1
		AGGrEGATOr	1	0.68	0.97	0.91	0.15	0.42	0.35	0.19	0.91	1
		KCCU	0.86	0.15	0.9	0.95	0.1	0.37	0.41	0.14	0.74	1
		GBIGM	0.11	0.07	0.12	0.09	0.13	0.1	0.08	0.08	0.13	0.16
	0.1	geXGB	1	1	1	1	1	1	1	1	1	1
		AGGrEGATOr	0.98	0.06	1	0.96	1	1	1	1	1	1
		KCCU	0.62	0.17	1	0.95	1	1	1	1	1	1
		GBIGM	0.12	0.17	0.19	0.18	0.12	0.2	0.13	0.1	0.19	0.12
	0.2	geXGB	1	1	1	1	1	1	1	1	1	1
		AGGrEGATOr	0.93	1	1	0.99	1	0.8	0.27	0.09	0.14	0.12
		KCCU	0.28	1	1	0.83	1	0.76	0.41	0.15	0.28	0.12
		GBIGM	0.19	0.2	0.25	0.31	0.23	0.26	0.26	0.29	0.22	0.1

Table 2: Average power in the first scenario

MAF	0.2				0.4			
Methods	geXGB	AGGrE	KCCU	GBIGM	geXGB	AGGrE	KCCU	GBIGM
Heritability	GATOr				GATOr			
0.01	0.554	0.46	0.256	0.108	0.84	0.618	0.377	0.107
0.025	0.971	0.675	0.408	0.116	1	0.553	0.375	0.1
0.05	1	0.763	0.626	0.159	1	0.658	0.562	0.107
0.1	1	1	0.879	0.168	1	0.9	0.874	0.152
0.2	1	1	0.935	0.229	1	0.634	0.583	0.231

Table 3: Results of the second scenario

Heritability	Method	geXGB	AGGrGETOr	KCCU	GBIGM
	Sample Size				
0.2	1000	0.67	0.15	0.11	0.2
	2000	1	0.18	0.38	0.16
	3000	1	0.11	0.55	0.23
	4000	1	0.31	0.76	0.21
	5000	1	0.26	0.87	0.12
0.4	1000	0.68	0.18	0.13	0
	2000	0.97	0.16	0.11	0.04
	3000	1	0.35	0.2	0.11
	4000	1	0.54	0.37	0.11
	5000	1	0.65	0.58	0.05

Table 1 shows the empirical statistical power of all methods on the simulated data sets of the first scenario. Table 2 shows the average statistical power of Table 1. Table3 displays the average statistical power of all methods on the simulated data sets of the second scenario. Figure 3 is a boxplot to summary the Table1. Bold font shows the best performance in each model under different heritability-MAF combinations. Notice that a larger value indicates better performance. On average, geXGB is the best performing algorithm in this comparison, it always outperforms the other methods, but not for all the data sets: on some data sets it yields to AGGrEGATOr. However, its performance is remarkably consistent and to be the top performer for most data sets. AGGrEGATOr can achieve the same performance when MAF=0.2 and heritability is larger than 0.05. While with smaller heritability, geXGB gives much better performance. When heritability is the smallest 0.01, geXGB wins 6 models while AGGrEGATOr wins 3 for MAF=0.2 and 9 v.s.2 for MAF=0.4. AGGrEGATOr has the better average performance than KCCU. However, when it works poorly on a data set, its performance can be very low, even lower than KCCU, such as M7,M8,M9 of (h=0.2,MAF=0.4). The Figure 3 also shows that AGGrEGATOr has a large range of performance for the 10 models. But geXGB does not exhibit this poor behavior.

KCCU seems to have power to detect some disease models. Figure 4 shows the relationship between heritability and the average performance. From it, we can see that KCCU has the

similar performance pattern with AGGrEGATOr, although AGGrEGATOr is much more powerful in most of the simulated scenarios. GBIGM has almost no power in detecting gene-gene interaction with only one causal pair. This result replicated Emily [7]’s result of simulation.

Considering about the relationship between performance of geXGB and genetic constrains, as expected, power of all the simulation is affected greatly by the heritability, i.e. the effect size of the interaction. With a difference, between heritability of 0.01 to 0.025, geXGB almost doubling the power for a given sample size of $n = 3000$ with $MAF = 0.2$. Other methods are also shown an steady upward trend (Table 1). It also depends on the minor allele frequencies (MAF) of the interacting SNPs, e.g. for the cases of $h=0.01$, the power of the different simulations ranges between 0.14-0.93 for $MAF = 0.2$, while it ranges between 0.66-0.96 for $MAF = 0.4$ (Table 1). The average power is 0.554 for $MAF = 0.2$, which is much lower than 0.84 for $MAF = 0.4$ (Table 2). Sample size of the data set has a considerable effect on power as well, geXGB, AGGrGETOr and KCCU have lower power estimated for $n = 1000$ compared to $n = 5000$ (Table 3).

Thus, our results demonstrate that geXGB is particularly efficient in detecting gene-gene interaction where one SNP pair is assumed to be causal by the purely and strictly epistasis without main effect. Compared to other methods, geXGB has the capacity to accurately identify a wide range of epistatic signals.

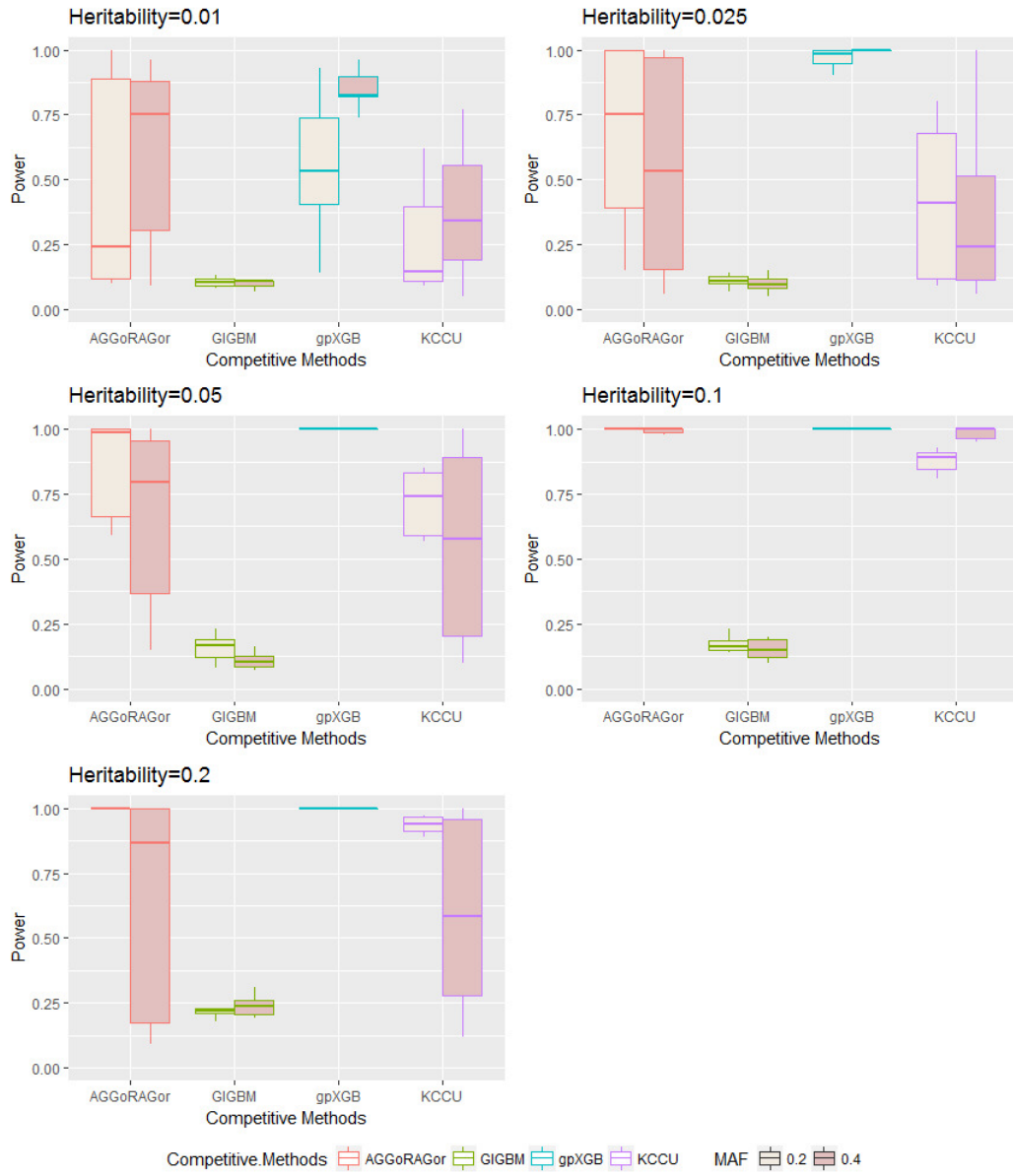


Figure 3: Performance on the simulated datasets

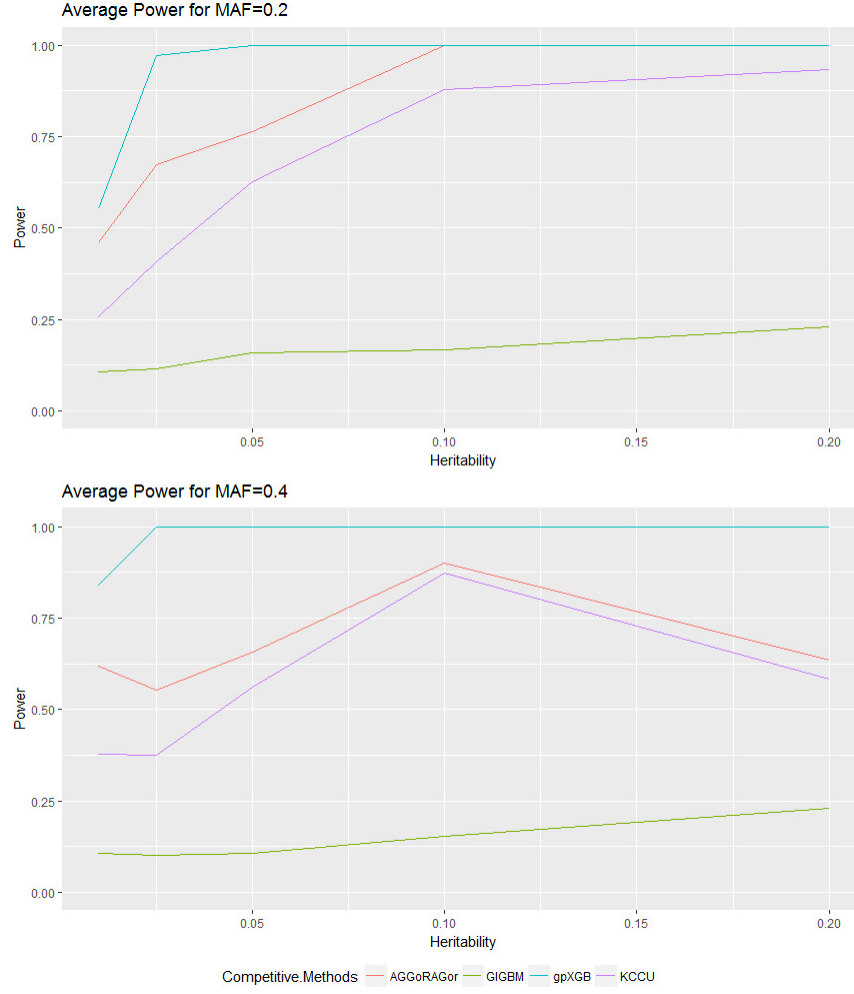


Figure 4: Average performance on the simulated datasets

3.2 Real data sets analysis

4 Discussion

In this paper, we proposed a new machine learning algorithm based procedure, called geXGB, to detect interaction at the gene level by applying the XGBoost and permutation strategies. We proved the capacity of geXGB to detect gene interaction, especially for the pure and strict type interactions. According to a power study based on a wide range of disease models with different genetic constrain combinations, geXGB is shown to be the

most powerful method when one causal pair is associated with the case-control phenotypes.

In the geXGB procedure, we consider the interaction between two genes as an additive model, each additive term represent some kind of non-linear interaction between SNPs from the two genes. We did not model the interactions form between SNPs explicitly. Such assumption made the method more powerful compared to AGGrEGATOr or KCCU. Our Method also can include the covariates matrix, which can correct confounding information such as population stratification with no more computational resources. Necessary for GWAS. We have applied a version with parallelize calculation for the permutation.

The application of geXGB procedure to the association between RA and 43 genes from RA pathway confirmed that the capacity for geXGB to be a robust and valid method compared to competitive methods and also gives promising new insights in the etiology of RA.

Genome-wide implementation of geXGB is hardly feasible since the input of the model is all the candidate genes, too many unrelated genes may be harmful for the performance of the XGBoost model. To avoid the computational all the gene pairs, another strategy consists in using prior biological knowledge to reduce the number of genes need to be considered. Indeed, geXGB procedure can easily be used in a network based approach since it is widely assumed that protein-protein interaction network or pathway-based approach can be successfully combined to GWAS. Furthermore, the term “gene” refers to a collection of SNPs and can be any locus, even some non-function unit such as non-coding RNA. For all the reason we believe that the geXGB procedure can help detecting a part of the missing heritability.

5 Conflict of interest

6 Acknowledgements

References

- [1] Wan, X., et al., *BOOST: A fast approach to detecting gene-gene interactions in genome-wide case-control studies*. Am J Hum Genet, 2010. 87(3): p. 325-40.
- [2] Li, J., et al., *Detecting gene-gene interactions using a permutation-based random forest method*. BioData Min, 2016. 9: p. 14.
- [3] Peng, Q., J. Zhao, and F. Xue, *A gene-based method for detecting gene-gene co-association in a case-control association study*. Eur J Hum Genet, 2010. 18(5): p. 582-7.

- [4] Yuan, Z., et al., *Detection for gene-gene co-association via kernel canonical correlation analysis*. BMC Genet, 2012. 13: p. 83.
- [5] Larson, N.B., et al., *Kernel canonical correlation analysis for assessing gene-gene interactions and application to ovarian cancer*. Eur J Hum Genet, 2014. 22(1): p. 126-31.
- [6] Li, J., et al., *A gene-based information gain method for detecting gene-gene interactions in case-control studies*. Eur J Hum Genet, 2015. 23(11): p. 1566-72.
- [7] Emily, M., *AGGrEGATOr: A Gene-based GENE-Gene interActTiOn test for case-control association studies*. Stat Appl Genet Mol Biol, 2016. 15(avg2): p. 151-71.
- [8] Ma, L., A.G. Clark, and A. Keinan, *Gene-based testing of interactions in association studies of quantitative traits*. PLoS Genet, 2013. 9(2): p. e1003321.
- [9] Chen, T. and C. Guestrin. *XGBoost: A Scalable Tree Boosting System*. in *ACM SIGKDD International Conference on Knowledge Discovery and Data Mining*. 2016.
- [10] Mentch, Lucas, and Giles Hooker, *Formal hypothesis tests for additive structure in random forests*. Journal of Computational and Graphical Statistics, 2016.
- [11] Greene, C.S., et al., *Enabling personal genomics with an explicit test of epistasis*. Pac Symp Biocomput, 2010: p. 327-36.
- [12] Urbanowicz, R.J., et al., *GAMETES: a fast, direct algorithm for generating pure, strict, epistatic models with random architectures*. BioData Min, 2012. 5(1): p. 16.
- [13] Laurie, C.C., et al., *Quality control and quality assurance in genotypic data for genome-wide association studies*. Genet Epidemiol, 2010. 34(6): p. 591-602.