

Regression with Categorical Variables

- Back to SLR, where we have a response variable Y and one predictor X ,

$$Y \sim X.$$

- X : a categorical variable, e.g., gender, education level, etc.
- How to run regression?
- First, let's revisit the [corrosion data](#).

Corrosion:

Data consist of thirteen specimens of 90/10 Cu-Ni alloys with varying iron content in percent. The specimens were submerged in sea water for 60 days and the weight loss due to corrosion was recorded in units of milligrams per square decimeter per day.

Fe Iron content in percent

Loss Weight loss in mg per
square decimeter per day

Fe	loss
0.01	127.6
0.01	130.1
0.01	128.0
0.48	124.0
0.48	122.0
0.71	110.8
0.71	113.1
0.95	103.9
1.19	101.5
1.44	92.3
1.44	91.4
1.96	83.7
1.96	86.2

Fe	loss	fitted
0.01	127.6	128.5667
0.01	130.1	128.5667
0.01	128.0	128.5667
0.48	124.0	123.0000
0.48	122.0	123.0000
0.71	110.8	111.9500
0.71	113.1	111.9500
0.95	103.9	103.9000
1.19	101.5	101.5000
1.44	92.3	91.8500
1.44	91.4	91.8500
1.96	83.7	84.9500
1.96	86.2	84.9500

```
> (127.6+130.1+
    128) / 3
[1] 128.5667
```

```
> (110.8 + 113.1) / 2
[1] 111.95
```

```
> (92.3+91.4) / 2
[1] 91.85
```

```
> (83.7+86.2) / 2
[1] 84.95
```

```
> ga=lm(loss ~ factor(Fe), data=corrosion);
> summary(ga)
```

	Estimate	Std. Error	t value	Pr(> t)
(Intercept)	128.567	0.809	158.914	0.000
factor(Fe) 0.48	-5.567	1.279	-4.352	0.005
factor(Fe) 0.71	-16.617	1.279	-12.990	0.000
factor(Fe) 0.95	-24.667	1.618	-15.245	0.000
factor(Fe) 1.19	-27.067	1.618	-16.728	0.000
factor(Fe) 1.44	-36.717	1.279	-28.703	0.000
factor(Fe) 1.96	-43.617	1.279	-34.097	0.000

How to interpret those coefficients?

One-Way ANOVA Model

group 1	$y_{11},$	y_{12}	y_{13}	\cdots	y_{1n_1}
group 2	$y_{21},$	y_{22}	\cdots		y_{2n_2}
				\cdots	
group g	$y_{g1},$	$y_{g2},$	\cdots		y_{gn_g}

g is # of groups,

n_i denotes # of obs in the i -th group,

and the total sample size $n = \sum_{i=1}^g n_i$.

- The LS fit for y_{ij} is the corresponding group mean

$$\hat{y}_{ij} = \bar{y}_i.$$

- Residuals

$$r_{ij} = y_{ij} - \hat{y}_{ij} = y_{ij} - \bar{y}_i.$$

- RSS

$$\sum_{i=1}^g \sum_{j=1}^{n_i} (y_{ij} - \bar{y}_i)^2,$$

i.e., the within-group variation.

- The one-way ANOVA model is described as

$$y_{ij} = \mu + \alpha_i + e_{ij}, \quad e_{ij} \text{ iid } \sim N(0, \sigma^2).$$

- The unknown parameters are

$$(\mu, \alpha_1, \dots, \alpha_g).$$

- For simplicity, consider a simple case $g = 2, n_1 = 3$ and $n_2 = 2$, and write the one-way ANOVA model in matrix form

$$\begin{pmatrix} y_{11} \\ y_{12} \\ y_{13} \\ y_{21} \\ y_{22} \end{pmatrix} = \begin{pmatrix} 1 & 1 & 0 \\ 1 & 1 & 0 \\ 1 & 1 & 0 \\ 1 & 0 & 1 \\ 1 & 0 & 1 \end{pmatrix} \begin{pmatrix} \mu \\ \alpha_1 \\ \alpha_2 \end{pmatrix} + \text{err.}$$

- However, this model is over-parameterized. We need to put some constraint on μ or α_i 's.

How to the code categorical variables?

- $\mu = 0$: What's the design matrix \mathbf{X} ? How to interpret the parameters? (the default case)
- $\alpha_1 = 0$: What's the design matrix \mathbf{X} ? How to interpret the parameters? (`contr.treatment`)
- $\sum \alpha_i = 0$: What's the design matrix \mathbf{X} ? How to interpret the parameters? (`contr.sum`)
- Suffices to remember the default case; the interpretations are not important.

$$\begin{pmatrix} y_{11} \\ y_{12} \\ y_{13} \\ y_{21} \\ y_{22} \end{pmatrix} = \begin{pmatrix} 1 & 1 & 0 \\ 1 & 1 & 0 \\ 1 & 1 & 0 \\ 1 & 0 & 1 \\ 1 & 0 & 1 \end{pmatrix} \begin{pmatrix} \mu \\ \alpha_1 \\ \alpha_2 \end{pmatrix} + \text{err.}$$

$$\begin{pmatrix} y_{11} \\ y_{12} \\ y_{13} \\ y_{21} \\ y_{22} \end{pmatrix} = \begin{pmatrix} 1 & 0 \\ 1 & 0 \\ 1 & 0 \\ 0 & 1 \\ 0 & 1 \end{pmatrix} \begin{pmatrix} \alpha_1 \\ \alpha_2 \end{pmatrix} + \text{err}$$

```
> tmp= lm(loss ~ factor(Fe)-1, data=corrosion);  
> summary(tmp)
```

Coefficients:

	Estimate	Std. Error	t value	Pr(> t)	
factor(Fe) 0.01	128.5667	0.8090	158.91	4.19e-12	***
factor(Fe) 0.48	123.0000	0.9909	124.13	1.84e-11	***
factor(Fe) 0.71	111.9500	0.9909	112.98	3.24e-11	***
factor(Fe) 0.95	103.9000	1.4013	74.15	4.05e-10	***
factor(Fe) 1.19	101.5000	1.4013	72.43	4.66e-10	***
factor(Fe) 1.44	91.8500	0.9909	92.70	1.06e-10	***
factor(Fe) 1.96	84.9500	0.9909	85.73	1.70e-10	***

$$\begin{pmatrix} y_{11} \\ y_{12} \\ y_{13} \\ y_{21} \\ y_{22} \end{pmatrix} = \begin{pmatrix} 1 & 1 & 0 \\ 1 & 1 & 0 \\ 1 & 1 & 0 \\ 1 & 0 & 1 \\ 1 & 0 & 1 \end{pmatrix} \begin{pmatrix} \mu \\ \alpha_1 \\ \alpha_2 \end{pmatrix} + \text{err.}$$

$$\begin{pmatrix} y_{11} \\ y_{12} \\ y_{13} \\ y_{21} \\ y_{22} \end{pmatrix} = \begin{pmatrix} 1 & 0 \\ 1 & 0 \\ 1 & 0 \\ 1 & 1 \\ 1 & 1 \end{pmatrix} \begin{pmatrix} \mu \\ \alpha_2 \end{pmatrix} + \text{err} = \begin{pmatrix} \mu \\ \mu \\ \mu \\ \mu + \alpha_2 \\ \mu + \alpha_2 \end{pmatrix} + \text{err}$$

```
> ga=lm(loss ~ factor(Fe), data=corrosion);
> summary(ga)
```

	Estimate	Std. Error	t value	Pr(> t)
(Intercept)	128.567	0.809	158.914	0.000
factor(Fe) 0.48	-5.567	1.279	-4.352	0.005
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factor(Fe) 1.44	-36.717	1.279	-28.703	0.000
factor(Fe) 1.96	-43.617	1.279	-34.097	0.000

```
> 123-128.5667
[1] -5.5667
> 111.95-128.5667
[1] -16.6167
```

$$\begin{pmatrix} y_{11} \\ y_{12} \\ y_{13} \\ y_{21} \\ y_{22} \end{pmatrix} = \begin{pmatrix} 1 & 1 & 0 \\ 1 & 1 & 0 \\ 1 & 1 & 0 \\ 1 & 0 & 1 \\ 1 & 0 & 1 \end{pmatrix} \begin{pmatrix} \mu \\ \alpha_1 \\ \alpha_2 \end{pmatrix} + \text{err.}$$

$$\begin{pmatrix} y_{11} \\ y_{12} \\ y_{13} \\ y_{21} \\ y_{22} \end{pmatrix} = \begin{pmatrix} 1 & 1 \\ 1 & 1 \\ 1 & 1 \\ 1 & -1 \\ 1 & -1 \end{pmatrix} \begin{pmatrix} \mu \\ \alpha_1 \end{pmatrix} + \text{err} = \begin{pmatrix} \mu + \alpha_1 \\ \mu + \alpha_1 \\ \mu + \alpha_1 \\ \mu - \alpha_1 \\ \mu - \alpha_1 \end{pmatrix} + \text{err}$$

```
> newFe = factor(Fe)
> contrasts(newFe) = contr.sum(7)
> tmp= lm(loss ~ newFe);summary(tmp)
```

	Estimate	Std. Error	t value	Pr(> t)	
(Intercept)	106.5310	0.4167	255.644	2.42e-13	***
newFe1	22.0357	0.8007	27.519	1.52e-07	***
newFe2	16.4690	0.9354	17.607	2.15e-06	***
newFe3	5.4190	0.9354	5.793	0.00116	**
newFe4	-2.6310	1.2555	-2.096	0.08097	.
newFe5	-5.0310	1.2555	-4.007	0.00706	**
newFe6	-14.6810	0.9354	-15.695	4.24e-06	***

```
> tmp=round(ga$fitted, dig=5); tmp=unique(tmp)
[1] 128.5667 123.0000 111.9500 103.9000
101.5000 91.8500 84.9500
> mean(tmp)
[1] 106.531
> round(tmp-mean(tmp), dig=4)
[1] 22.0357 16.4690 5.4190 -2.6310 -5.0310
-14.6810 -21.5810
```

The F -test

- Are levels of the factor really different? State the hypothesis in terms of models

$$H_a : y_{ij} = \mu + \alpha_i + e_{ij}$$

$$H_0 : y_{ij} = \mu + e_{ij}$$

- They are two nested models, then we can use F -test.

$$\frac{(\text{RSS}_0 - \text{RSS}_a)/(g - 1)}{\text{RSS}_a/(n - g)} \sim F_{g-1, n-g},$$

under H_0 . The test statistic can also written as

$$\frac{\sum_{i=1}^g n_i (y_{i\cdot} - y_{\cdot\cdot})^2 / (g - 1)}{\sum_{i,j} (y_{ij} - y_{i\cdot})^2 / (n - g)} = \frac{\text{Between-group Variation} / (g - 1)}{\text{Within-group Variation} / (n - g)}.$$

Coagulation:

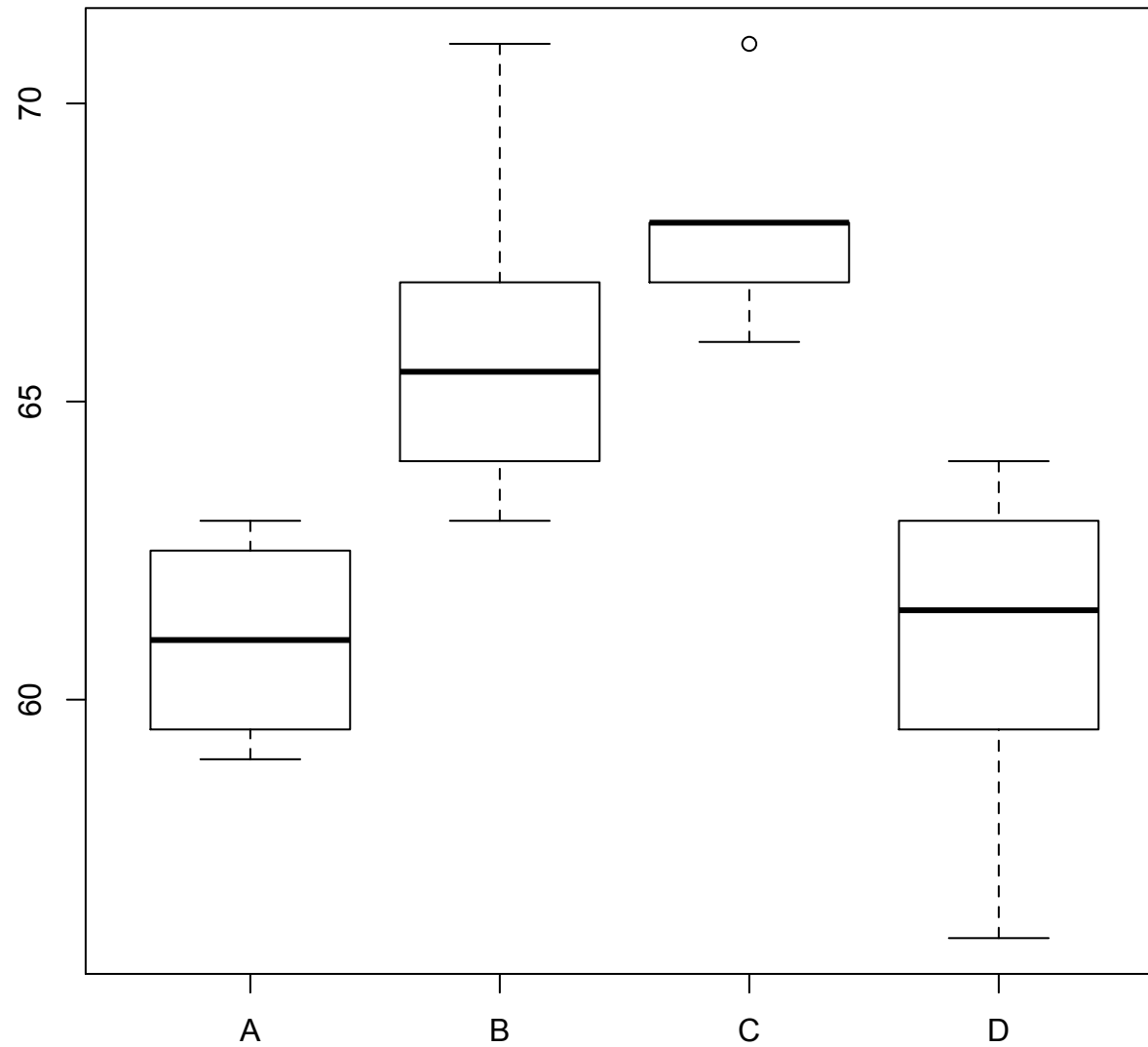
Dataset comes from a study of blood coagulation times. 24 animals were randomly assigned to four different diets and the samples were taken in a random order.

coag coagulation time in
seconds

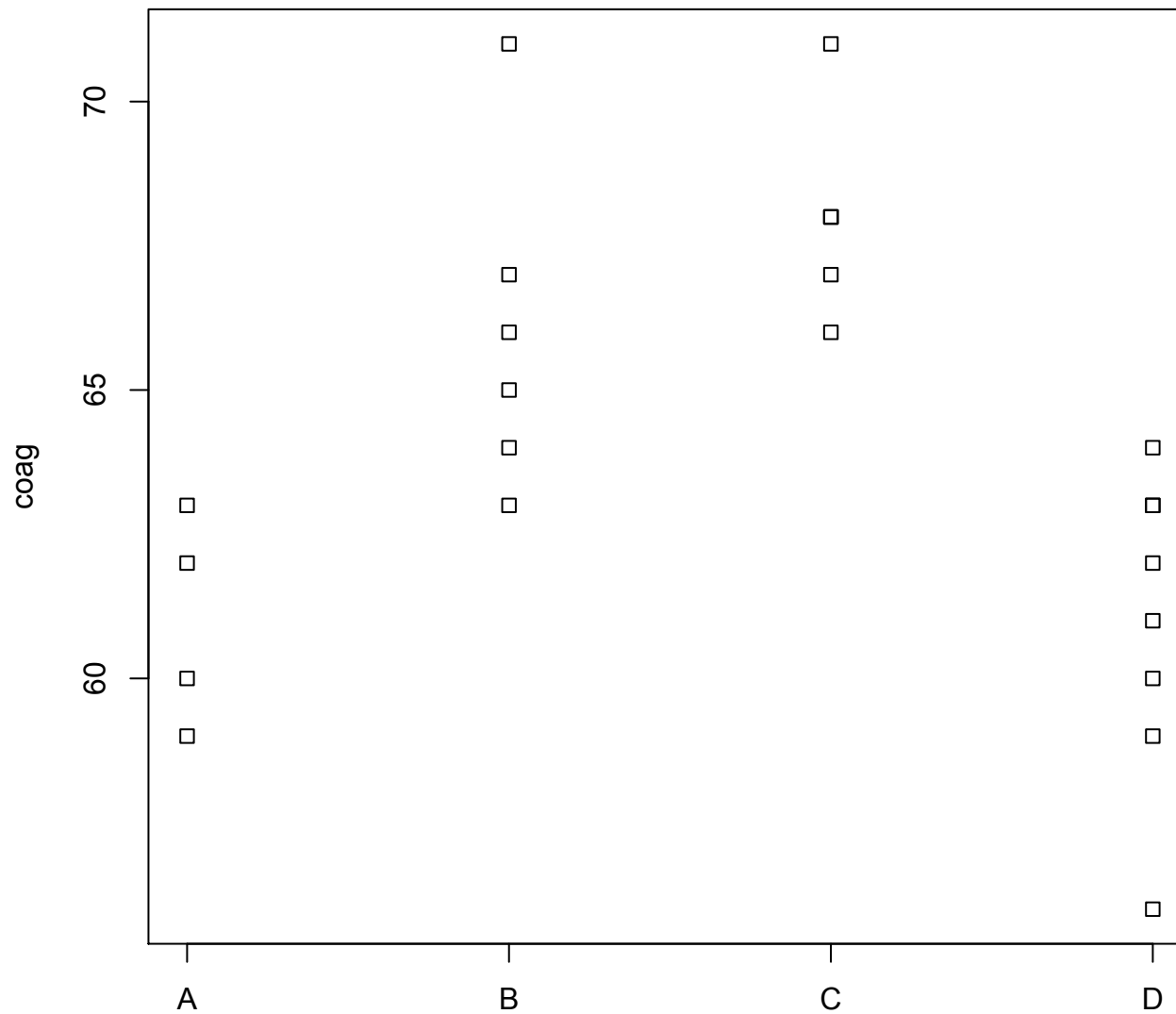
diet diet type - A,B,C or D

	coag	diet
1	62	A
2	60	A
3	63	A
4	59	A
5	63	B
6	67	B
7	71	B
8	64	B
9	65	B
21	63	D
22	64	D
23	63	D
24	59	D

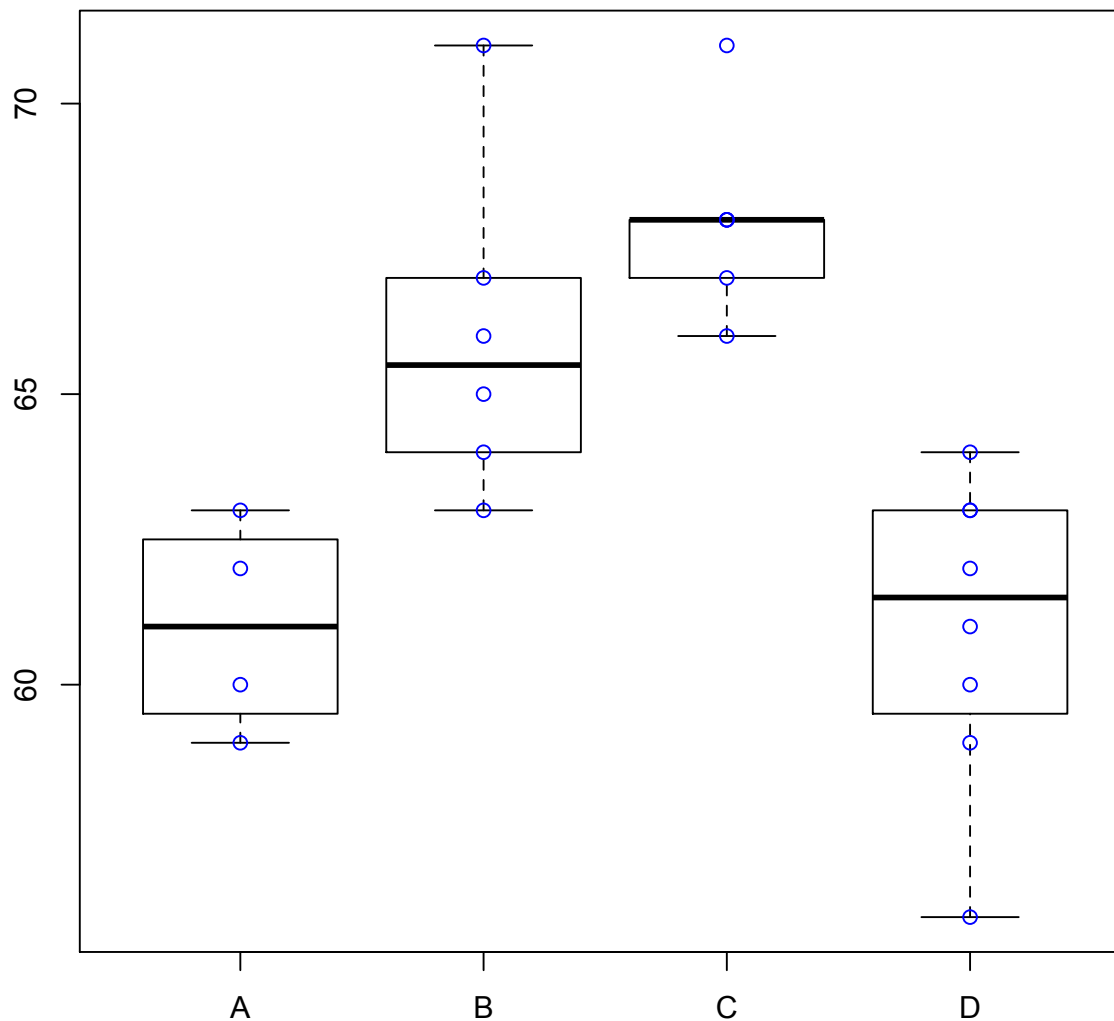
```
> attributes(diet)
$levels
[1] "A" "B" "C" "D"
```



```
> boxplot(coag ~ diet)
```



```
> stripchart(coag ~ diet, vertical=TRUE)
```



```
> boxplot(coag ~ diet, outline=FALSE)
> stripchart(coag ~ diet, vertical=TRUE,
  add=TRUE, col="blue", pch=1)
```

```
> g=lm(coag~diet)
```

```
> anova(g)
```

Analysis of Variance Table

Response: coag

	Df	Sum Sq	Mean Sq	F value	Pr(>F)
diet	3	228	76.0	13.571	4.658e-05
Residuals	20	112	5.6		