

# STAT 510 HW#7

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Ex 16.7.7

1. (a) Under null,  $X$  and  $Z$  are conditionally independent given  $Y$ .

Hence,  $P(X=i, Z=k | Y=j) = P(X=i | Y=j) \cdot P(Z=k | Y=j)$

$$r_{ij} \cdot s_{kj} \cdot t_j = P(X=i, Z=k | Y=j) \cdot P(Y=j)$$

$$= P(X=i, Y=j, Z=k)$$

$$p_{ijk} = r_{ij} \cdot s_{kj} \cdot t_j$$

(b)  $N = 2812$ , Under alternative.

	$Y=0$		$Y=1$		$Y=2$	
$\hat{p}_{ijk}$	$Z=0$	$Z=1$	$Z=0$	$Z=1$	$Z=0$	$Z=1$
$X=0$	0.1742	0.0017	0.3261	0.0245	0.0263	0.0206
$X=1$	0.0992	0.0021	0.2140	0.0238	0.0515	0.0355

(c)  $t_0 = 0.27738$ ,  $t_1 = 0.58855$ ,  $t_2 = 0.13407$

$$r_{00} = 0.63462, r_{10} = 0.36538$$

$$r_{01} = 0.59577, r_{11} = 0.40423$$

$$r_{02} = 0.35013, r_{12} = 0.64987$$

$$s_{00} = 0.9859, s_{10} = 0.0141$$

$$s_{01} = 0.91782, s_{11} = 0.08218$$

$$s_{02} = 0.5809, s_{12} = 0.4190$$

Under Null

	$Y=0$		$Y=1$		$Y=2$	
$\hat{p}_{ijk}$	$Z=0$	$Z=1$	$Z=0$	$Z=1$	$Z=0$	$Z=1$
$X=0$	0.17355	0.00248	0.32182	0.02882	0.02727	0.01967
$X=1$	0.09992	0.00143	0.21836	0.01955	0.05056	0.03652



Under Null:  $l(P, n) = \sum n_{ijk} \log\left(\frac{n_{ijk}}{n}\right) = -5288.831$

Ex 16.7.7

1. (d) under Alternative,  $l(P, n) = \sum n_{ijk} \cdot \log(P_{ijk}) = -5292.116$

$2 \log(\Lambda(n)) = 2(-5288.831 + 5292.116) = 6.571$

(e). there are 11 free parameters for alternative

there are 3 free parameters for  $r_{ij}$

3 free parameters for  $s_{ki}$

2 free parameters for  $t_i$

$(3+3+2) = 8$  free parameters for Null

(f) the degree of freedom is  $(11-8) = 3$ , for  $\chi^2$

P-value = 0.0869 > 0.05

Hence, we accept the Null and

conclude that X and Z are conditionally independent given Y.

Ex 16.7.10

2. (a)  $f \sim N(0, 1)$

$l_i(\mu, x_i) = \log\left(\frac{1}{\sqrt{2\pi}} e^{-\frac{1}{2}(x_i - \mu)^2}\right)$

$l'_i(0, x_i) = x_i$ ,  $I_i(0) = 1$

$T_n(x_i) = \frac{1}{\sqrt{n}} \sum x_i$

(b).  $f \sim \text{Laplace}$

$l_i(\mu, x_i) = \log\left(\frac{1}{2} e^{-|x_i - \mu|}\right)$

$l'_i(0, x_i) = \frac{x_i}{|x_i|}$ ,  $I_i(0) = 1$

$T_n(x_i) = \frac{1}{\sqrt{n}} \sum \frac{x_i}{|x_i|}$

(c).  $f \sim \text{Logistic}$

$l_i(\mu, x_i) = \log\left(\frac{e^{-(x_i - \mu)}}{(1 + e^{-(x_i - \mu)})^2}\right)$

$l'_i(0, x_i) = \frac{e^{x_i} - 1}{e^{x_i} + 1}$ ,  $I_i(0) = \frac{1}{3}$

$T_n(x_i) = \frac{\sqrt{3}}{\sqrt{n}} \cdot \sum \frac{e^{x_i} - 1}{e^{x_i} + 1}$

(d).  $f \sim N(0, 1)$  is exactly  $N(0, 1)$

(e).  $f \sim \text{Logistic}$  has the same distribution under null.



3. According to Bayes theorem:

$$\pi(H_1 | X) = f(X | H_1) \cdot P(H_1) / P(X)$$

$$\begin{aligned} \text{Odds}(H_1 | X=x) &= \frac{f(X | H_1) \cdot P(H_1) / P(X)}{f(X | H_0) \cdot P(H_0) / P(X)} \\ &= \frac{P(H_1)}{P(H_0)} \cdot \frac{\int_0^1 f(X | \theta) P_1(\theta) d\theta}{\int_0^1 f(X | \theta) P_0(\theta) d\theta} \end{aligned}$$

$$\text{Hence, Odds}(H_1 | X=x) = \text{Odds}(H_1) \cdot B_{10}(x)$$

Ex 15.7.4 4. (a)  $2 = P(X=0 | P=\frac{1}{2}) + P(X=5 | P=\frac{1}{2})$

$$= \left( \binom{5}{0} \frac{1}{2}^0 \cdot \frac{1}{2}^5 \right) \cdot 2$$

$$= 0.0625$$

(b)  $P(X=x | H_0)$

$$= \binom{5}{x} \frac{1}{2}^x \cdot \frac{1}{2}^{5-x}$$

$$P(X=x | H_1)$$

$$= \int_0^1 \binom{5}{x} p^x \cdot (1-p)^{5-x} \cdot 1 dp \quad \text{Since } P(H_1) \sim U(0,1)$$

$$= \frac{1}{720} \cdot \binom{5}{x} \cdot \Gamma(6-x) \Gamma(x+1)$$

$$B_{A0}(x) = \frac{\frac{1}{720} \binom{5}{x} \Gamma(6-x) \Gamma(x+1)}{\binom{5}{x} \frac{1}{2}^x \cdot \frac{1}{2}^{5-x}}$$

$$= \frac{2}{45} \cdot \Gamma(6-x) \Gamma(x+1)$$

(c)  $P(H_0 | X=x) = 1 / (1 + B_{A0}(x))$

When  $x=0$  or  $5$

$$P(H_0 | X=x) = 1 / (1 + \frac{2}{45} \cdot 5!)$$

$$= 0.3157$$

Since  $2 = 0.0625$

The Posterior Probability is not close to 2



Ex 15.7.11

5. (a)  $X|\theta \sim \text{Poisson}(\theta)$  $\theta \sim \text{Gamma}(\alpha, \lambda)$ 

$$\begin{aligned}
 f(x|\alpha, \lambda) &= \int e^{-\theta} \cdot \frac{(c\theta)^x}{x!} \frac{\lambda^\alpha}{\Gamma(\alpha)} e^{-\lambda\theta} \cdot \theta^{\alpha-1} d\theta \\
 &= \frac{c^x}{x!} \frac{\Gamma(x+\alpha)}{\Gamma(\alpha)} \cdot \frac{\lambda^\alpha}{(c+\lambda)^{x+\alpha}} \text{Pdf Gamma}(x+\alpha, (c+\lambda)) \\
 &= \frac{\Gamma(x+\alpha)}{x! \Gamma(\alpha)} \cdot \frac{c^x}{(c+\lambda)^{x+\alpha}}
 \end{aligned}$$

(b) Since  $X_v$  and  $X_c$  are independent, $\theta_c$  and  $\theta_v$  are independent under alternative.

$$f(X_v, X_c | H_A) = f(X_v | \alpha, \lambda) \cdot f(X_c | \alpha, \lambda)$$

(c).  $X_v \sim \text{Poisson}(c_v \theta_v)$ ,  $X_c \sim \text{Poisson}(c_c \theta_c)$ Under Null,  $\theta_v = \theta_c$ 

$$\begin{aligned}
 f(X_v, X_c | H_0) &= \int e^{-c\theta} \frac{(c\theta)^{x_c}}{x_c!} e^{-c_v\theta} \frac{(c_v\theta)^{x_v}}{x_v!} \frac{\lambda^\alpha}{\Gamma(\alpha)} e^{-\lambda\theta} \theta^{\alpha-1} d\theta \\
 &= \frac{\Gamma(x_v + x_c + \alpha)}{x_v! x_c! \Gamma(\alpha)} \frac{c_v^{x_v} c_c^{x_c} \lambda^\alpha}{(\lambda + c_v + c_c)^{x_v + x_c + \alpha}} \cdot \text{Gamma}(x_c + x_v + \alpha, (c + c_v + c_c)) \\
 &= \frac{\Gamma(x_v + x_c + \alpha)}{x_v! x_c! \Gamma(\alpha)} \frac{c_v^{x_v} c_c^{x_c} \cdot \lambda^\alpha}{(\lambda + (c_v + c_c))^{x_v + x_c + \alpha}}
 \end{aligned}$$

$$\begin{aligned}
 (d) \quad B_{A0}(X_v, X_c) &= \frac{f(X_v, X_c | H_A)}{f(X_v, X_c | H_0)} \\
 &= \frac{\Gamma(x_v + x_c + \alpha) \Gamma(\alpha)}{\Gamma(x_v + \alpha) \Gamma(x_c + \alpha)} \frac{(c_v + \lambda)^{x_v + \alpha} (c + \lambda)^{x_c + \alpha}}{(\lambda + c_v + c_c)^{x_v + x_c + \alpha} \cdot \lambda^\alpha} \quad \leftarrow \text{invert} \\
 &= (e^{16.251})
 \end{aligned}$$

(e). Posterior =  $1 / (1 + B_{A0}(X_v, X_c)) \approx 0$ .Hence, we reject the null and conclude that  $\theta_v \neq \theta_c$ 