

IEOR 4703 - Monte Carlo Simulation

Solutions for Assignment 2

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1 Problem 1

The $1-\lambda$ confidence interval given by Chebyshev's Inequality is $\left[\hat{\theta}_n - \sqrt{\frac{\text{var}(\hat{\theta}_n)}{\lambda}}, \hat{\theta}_n + \sqrt{\frac{\text{var}(\hat{\theta}_n)}{\lambda}}\right]$.

The confidence interval by CLT is $\left[\hat{\theta}_n - z_{1-\frac{\lambda}{2}}\sqrt{\text{var}(\hat{\theta}_n)}, \hat{\theta}_n + z_{1-\frac{\lambda}{2}}\sqrt{\text{var}(\hat{\theta}_n)}\right]$. $z_{1-\frac{\lambda}{2}}$ is smaller than $\frac{1}{\sqrt{\lambda}}$. So Chebyshev's Inequality usually gives a conservative confidence interval. (Any sensible answer is OK.)

2 Problem 2

- (a) The exact solution given by BS formula is approximately 18.9586. The steps of simulation is following:

Set the parameters as the problem says

$n = 10000$

for $i = 1$ **to** n

Generate a standard normal variable R_i

$C_i = \exp(-rT) \max\{S_0 \exp((r - q - \frac{\sigma^2}{2})T + \sigma\sqrt{T}R_i) - K, 0\}$

$\hat{C}_n = \frac{1}{n} \sum_{i=1}^n C_i$

Delta hedging is also acceptable in this question. It just replaces the calculation of C_i in the algorithm. The details can be found in the answer to homework 1.

- (b) $\hat{\sigma} = \sqrt{\frac{\sum_{i=1}^n (C_i - \hat{C}_n)^2}{n-1}}$. The confidence interval is $\left[\hat{C}_n - z_{0.975} \frac{\hat{\sigma}}{\sqrt{n}}, \hat{C}_n + z_{0.975} \frac{\hat{\sigma}}{\sqrt{n}}\right]$.

- (c) Let $\hat{F}(x)$ be the empirical distribution function of $R_i, 1 \leq i \leq 10000$

$n = 10000$

$B = 1000$ (for example)

for $k = 1$ **to** B

for $i = 1$ **to** n

Generate $R_{i,k}$ from $\hat{F}(x)$

$C_{i,k} = \exp(-rT) \max\{S_0 \exp((r - q - \frac{\sigma^2}{2})T + \sigma\sqrt{T}R_{i,k}) - K, 0\}$

$\tilde{C}_k = \frac{1}{n} \sum_{i=1}^n C_{i,k}$

$\alpha = 0.05$

Let $\tilde{C}_{(\alpha/2)}$ and $\tilde{C}_{(1-\alpha/2)}$ be the $\alpha/2$ and $1 - \alpha/2$ percentiles of \tilde{C}_k .

The confidence interval is $\left[2\hat{C}_n - \tilde{C}_{(1-\alpha/2)}, 2\hat{C}_n - \tilde{C}_{(\alpha/2)}\right]$.

3 Problem 3

- (a) Set the parameters as the problem says
 $n = 50000$
for $i = 1$ **to** n
 $S_{i,0} = S_0$
for $j = 1$ **to** m
Generate a standard normal variable $R_{i,j}$
 $S_{i,j} = S_{i,j-1} \exp((r - q - \frac{\sigma^2}{2})\Delta t + \sigma\sqrt{\Delta t}R_{i,j})$
 $\bar{S}_i = \frac{1}{m} \sum_{j=1}^m S_{i,j}$
 $\theta_i = \exp(-rT) \max\{\bar{S}_i - K, 0\}$
 $\hat{\theta}_n = \frac{1}{n} \sum_{i=1}^n \theta_i$
 $\hat{\sigma} = \sqrt{\frac{1}{n-1} \sum_{i=1}^n (\theta_i - \hat{\theta}_n)^2}$

The confidence interval is $\left[\hat{\theta}_n - z_{0.995} \frac{\hat{\sigma}}{\sqrt{n}}, \hat{\theta}_n + z_{0.995} \frac{\hat{\sigma}}{\sqrt{n}} \right]$

- (b) Set the parameters as the problem says
 $n = 50000$
for $i = 1$ **to** $n/2$
 $S_{2i-1,0} = S_0$
 $S_{2i,0} = S_0$
for $j = 1$ **to** m
Generate a standard normal variable $R_{i,j}$
 $S_{2i-1,j} = S_{2i-1,j-1} \exp((r - q - \frac{\sigma^2}{2})\Delta t + \sigma\sqrt{\Delta t}R_{i,j})$
 $S_{2i,j} = S_{2i,j-1} \exp((r - q - \frac{\sigma^2}{2})\Delta t - \sigma\sqrt{\Delta t}R_{i,j})$
 $\bar{S}_{2i-1} = \frac{1}{m} \sum_{j=1}^m S_{2i-1,j}$
 $\bar{S}_{2i} = \frac{1}{m} \sum_{j=1}^m S_{2i,j}$
 $\theta_i = \frac{1}{2} \exp(-rT) (\max\{\bar{S}_{2i-1} - K, 0\} + \max\{\bar{S}_{2i} - K, 0\})$
 $\hat{\theta}_n = \frac{2}{n} \sum_{i=1}^{n/2} \theta_i$
 $\hat{\sigma}_A = \sqrt{\frac{1}{n/2-1} \sum_{i=1}^{n/2} (\theta_i - \hat{\theta}_n)^2}$

The confidence interval is $\left[\hat{\theta}_n - z_{0.995} \frac{\hat{\sigma}_A}{\sqrt{n/2}}, \hat{\theta}_n + z_{0.995} \frac{\hat{\sigma}_A}{\sqrt{n/2}} \right]$.

4 Problem 4

- (a) This is the same as the question (a) in Problem 3.

- (b) Set the parameters as the problem says
 $n = 50000$
 $p = 1000$ (for example)
// pilot simulation
for $i = 1$ **to** p
 $S_{i,0} = S_0$
for $j = 1$ **to** m
Generate a standard normal variable $R_{i,j}$
 $S_{i,j} = S_{i,j-1} \exp((r - q - \frac{\sigma^2}{2})\Delta t + \sigma\sqrt{\Delta t}R_{i,j})$
 $\bar{S}_i = \frac{1}{m} \sum_{j=1}^m S_{i,j}$
 $\theta_i = \exp(-rT) \max\{\bar{S}_i - K, 0\}$
 $Z_i = S_{i,m}$

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 $\hat{\theta}_p = \frac{1}{p} \sum_{i=1}^p \theta_i$ 
 $E(Z) = S_0 \exp((r - q)T)$ 
 $\text{var}(Z) = \frac{1}{p} \sum_{i=1}^p (Z_i - E(Z))^2$ 
 $\text{cov}(Z, \theta) = \frac{1}{p-1} \sum_{i=1}^p (Z_i - E(Z))(\theta_i - \hat{\theta}_p)$ 
 $c = -\frac{\text{cov}(Z, \theta)}{\text{var}(Z)}$ 
// main simulation
for  $i = 1$  to  $n$ 
     $S_{i,0} = S_0$ 
    for  $j = 1$  to  $m$ 
        Generate a standard normal variable  $R_{i,j}$ 
         $S_{i,j} = S_{i,j-1} \exp((r - q - \frac{\sigma^2}{2})\Delta t + \sigma\sqrt{\Delta t}R_{i,j})$ 
     $\bar{S}_i = \frac{1}{m} \sum_{j=1}^m S_{i,j}$ 
     $\theta_i = \exp(-rT) \max\{\bar{S}_i - K, 0\}$ 
     $Z_i = S_{i,m}$ 
 $\hat{\theta}_n = \frac{1}{n} \sum_{i=1}^n \theta_i$ 
 $\text{var}(\theta) = \frac{1}{n-1} \sum_{i=1}^n (\theta_i - \hat{\theta}_n)^2$ 
 $\tilde{\theta} = \hat{\theta}_n + \frac{c}{n} \sum_{i=1}^n (Z_i - E(Z))$ 
 $\hat{\sigma} = \sqrt{\text{var}(\theta) - \frac{\text{cov}(Z, \theta)^2}{\text{var}(Z)}}$ 
(Or define  $\tilde{\theta}_i = \theta_i + c(Z_i - E(Z))$  and  $\hat{\sigma} = \sqrt{\frac{\sum_{i=1}^n (\tilde{\theta}_i - \hat{\theta})^2}{n-1}}$ )

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The confidence interval is $\left[\tilde{\theta} - z_{0.995} \frac{\hat{\sigma}}{\sqrt{n}}, \tilde{\theta} + z_{0.995} \frac{\hat{\sigma}}{\sqrt{n}} \right]$.

- (c) Replace $Z_i = \exp(-rT) \max\{S_{i,m} - K, 0\}$ and $E(Z) = \text{BSFormula}(S_0, K, r, q, T, \sigma)$ in the former algorithm.
- (d) Replace $Z_i = \exp(-rT) \max\{\sqrt[m]{\prod_{j=1}^m S_{i,j}} - K, 0\}$ in the former algorithm. The key to calculate $E(Z)$ is that $\sqrt[m]{\prod_{j=1}^m S_{i,j}}$ is lognormal. Let

$$G = \log\left(\sqrt[m]{\prod_{j=1}^m S_{i,j}}\right) = \log S_0 + \sum_{j=1}^m \frac{m+1-j}{m} \left((r - q - \frac{\sigma^2}{2})\Delta t + \sigma\sqrt{\Delta t}R_j\right)$$

Its expectation is $\mu_G = \log S_0 + \frac{m+1}{2m}(r - q - \frac{\sigma^2}{2})T$ and its variance is $\sigma_G^2 = \frac{(m+1)(2m+1)}{6m^2}\sigma^2 T$. In the same way as calculating BS Formula, we get

$$E(Z) = e^{\mu_G + \frac{1}{2}\sigma_G^2 - rT} \Phi(d_1) - K e^{-rT} \Phi(d_2),$$

where $d_1 = \frac{\mu_G - \log K + \sigma_G^2}{\sigma_G}$ and $d_2 = \frac{\mu_G - \log K}{\sigma_G}$. There are other ways to calculate $E(Z)$.

- Conclusion: The control variate in (d) is the best.