Network modelling for high-dimensional data - II

Instructors:

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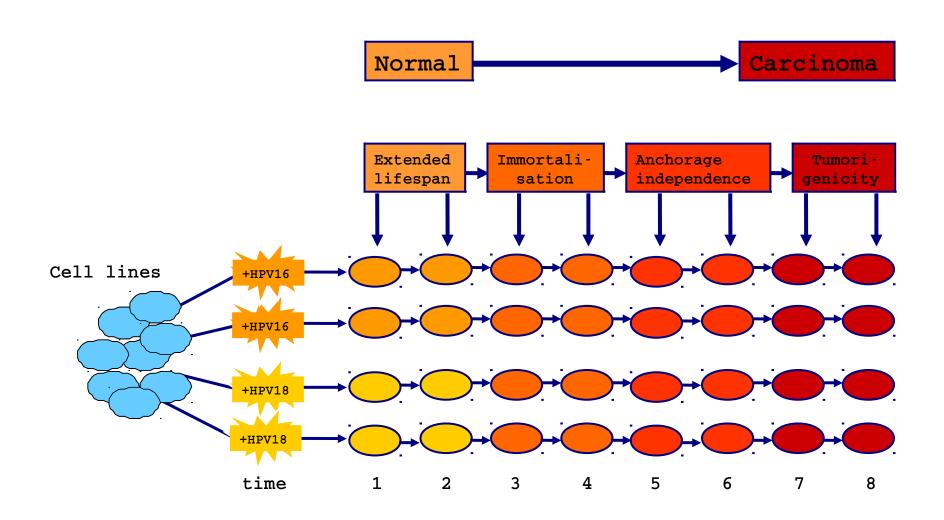
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Experiment



1.1 + 1.2: Libraries + experiment

```
# load libraries
library (Biobase)
library(lattice)
library(longitudinal)
library(rags2ridges)
library(ragt2ridges)
library(SparseTSCGM)
# load data
data(hpvP53)
# reformat data
Y <- longitudinal2array(t(exprs(hpvP53rna)))
# zero center data, variate- and cell line-wise
Y <- centerVAR1data(Y)
```

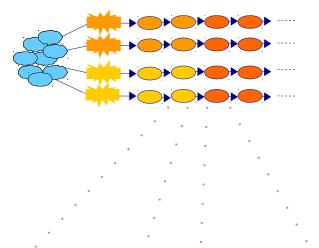




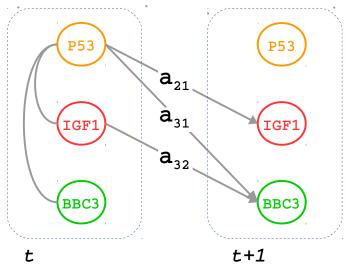
1.3: Data exploration

```
# plot time-courses of all variates
plotVAR1data(Y)
# plot time-courses of a single variate
plotVAR1data(Y[5, , , drop=FALSE])
# K-means clustering of the variates
cellLine <- 1
kClust <- kmeans(Y[, , cellLine],</pre>
                   centers=7, nstart=100)$cluster
# heatmap of reshuffled data
edgeHeat(Y[unlist(lapply(1:max(kClust),
         function(id, clusters) { which(clusters==id) },
         kClust)), , cellLine])
# plot time-courses of smallest cluster
plotVAR1data(Y[kClust==which.min(table(kClust)), , ,
               drop=FALSE])
```

Aim + model



Unrolled:



VAR(1) model:

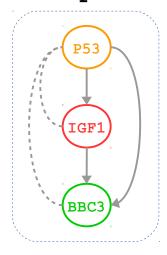
$$Y_{1,t+1} = + e_{1,t}$$

 $Y_{2,t+1} = a_{21} Y_{1,t} + e_{2,t}$
 $Y_{3,t+1} = a_{31} Y_{1,t} + a_{32} Y_{2,t} + e_{3,t}$

where:

$$Y_1 \leftrightarrow P53$$
; $Y_2 \leftrightarrow IGF1$; $Y_3 \leftrightarrow BBC3$

Curled up:



Model

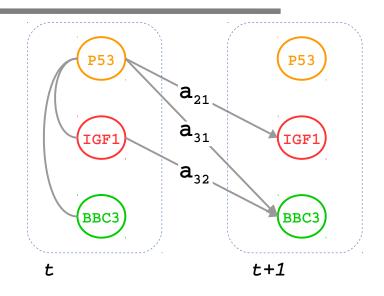
VAR(1) model:

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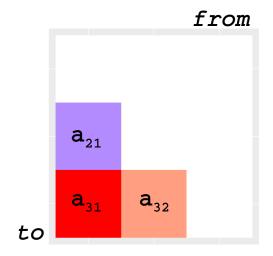


In matrix notation:

$$egin{array}{ll} \mathbf{Y}_t &= \mathbf{A}\mathbf{Y}_{t-1} + oldsymbol{arepsilon}_t \ & ext{with } oldsymbol{arepsilon}_t \sim \mathcal{N}(\mathbf{0}_{p imes 1}, oldsymbol{\Omega}_arepsilon^{-1}) \end{array}$$

A and ε: propagation of endo- and exogenous signal, resp.ε innovation of the system.

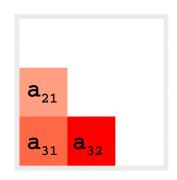
Parameters as heatmap:

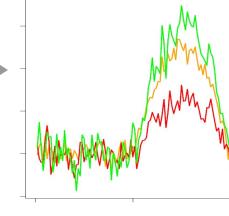


Model

VAR(1) model:

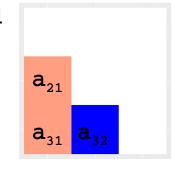
Coherent feed forward loop:

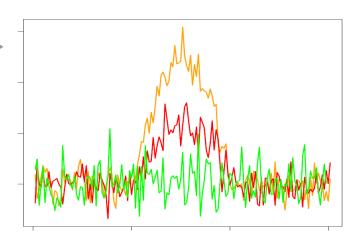




exogenous pulse

Incoherent feed forward loop:





TSCG

Time Series Chain Graph (TSCG)

→ conditional (in)dependencies

Temporal CI:

$$Y_{j_1,t} \perp \!\!\! \perp Y_{j_2,t+1} \mid$$
 other $Y_{j,t}$'s

Example:

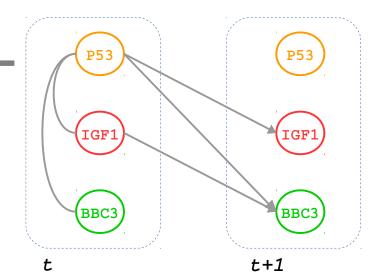
- → CI: BBC3 & P53 (no edge)
- → not CI: P53 & IGF1 (directed edge)

Contemporaneous CI:

$$Y_{j_1,t} \perp \!\!\! \perp Y_{j_2,t} \mid Y_{j,t-1}$$
's, other $Y_{j,t}$'s

Example:

- → CI: BBC3 & IGF1 (no edge)
- → not CI: P53 & IGF1 (undirected edge)

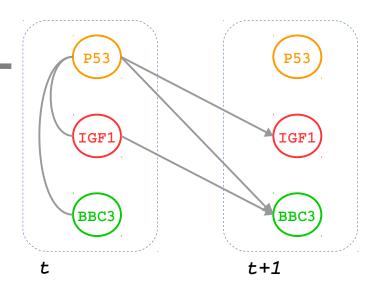


TSCG

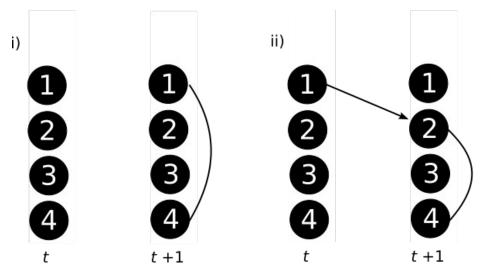
TSCG harbors global conditional (in)dependencies:

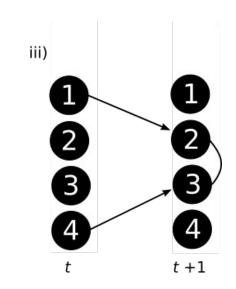
$$Y_{j_1,*} \perp \!\!\! \perp Y_{j_2,*} \mid$$
 other $Y_{j,*}$'s

Wermuth condition for CI: unconnected nodes may not exert influence on the same node.



CI of nodes 1 and 4 forbids motifs:



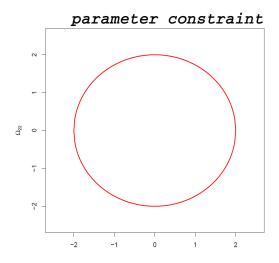


Estimation

Augment log-likelihood with ridge penalty:

$$|\lambda_a||\mathbf{A}||_2^2 + |\lambda_\omega||\mathbf{\Omega}_\varepsilon||_2^2$$

Include targets.



Analytic estimators:

$$\operatorname{vec}[\hat{\mathbf{A}}(\lambda_a)] = [\lambda_a \mathbf{I}_{p^2 \times p^2} + \hat{\mathbf{\Gamma}}(0) \otimes \mathbf{\Omega}_{\varepsilon}]^{-1} \operatorname{vec}[\mathbf{\Omega}_{\varepsilon} \hat{\mathbf{\Gamma}}(-1)],$$

$$\hat{\mathbf{\Omega}}_{\varepsilon}(\lambda_{\omega}) = \left\{ \left[\lambda_{\omega} \mathbf{I}_{p \times p} + \frac{1}{4} \mathbf{S}_{\varepsilon}^{2} \right]^{1/2} + \frac{1}{2} \mathbf{S}_{\varepsilon} \right\}^{-1}.$$

Initiate and iterate.

Rewrite vec(A) for computational efficiency

1.4 + 1.5: Penalty + estimation

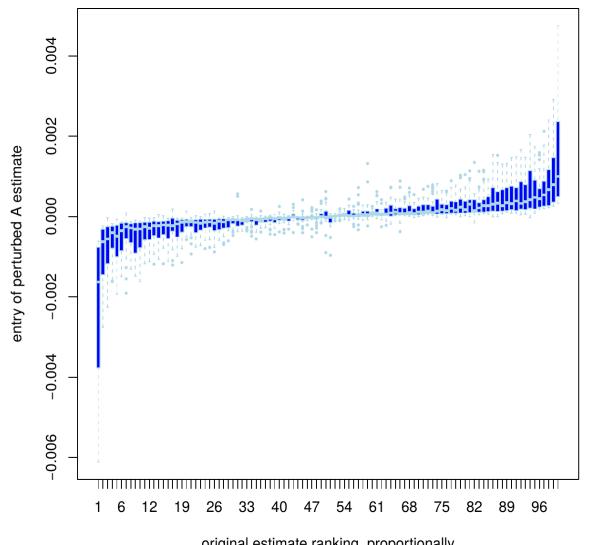


```
# find optimal penalty parameters
optLambdas <- optPenaltyVAR1(Y, lambdaMin =c(0.01, 0.00001),
                                 lambdaMax = c(1000, 1))
# fit the model
VAR1hat <- ridgeVAR1(Y=Y, lambdaA=optLambdas[1],</pre>
                           lambdaP=optLambdas[2])
# extract parameter estimates
Ahat <- VAR1hat$A; Phat <- VAR1hat$P
# add row and column names
rownames (Ahat) <- colnames (Ahat) <- rownames (Phat) <-
                  colnames(Phat) <- rownames(hpvP53rna)</pre>
# heatmap of estimate of A
edgeHeat(Ahat, main="ridge estimate of A")
# heatmap of partial correlation matrix estimate
edgeHeat(pcor(Phat), main="part. Cor. estimate", diag=FALSE)
```





Homework: code takes too long for practical



original estimate ranking, proportionally

Support



Known

Equality constraints for known absent temporal edges:

$$\max_{\{\mathbf{A}_c: \mathbf{C} \, \mathrm{VeC}(\mathbf{A}_c) = \mathbf{d}\}} \mathcal{L}^{\mathrm{pen}}(\mathbf{Y}; \mathbf{A}_c, \mathbf{\Omega}_{\varepsilon}; \lambda_a, \lambda_{\omega})$$

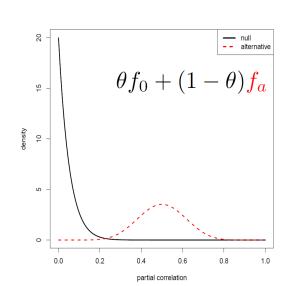
Analytic solution, efficiently evaluable.

Similarly for known absent contemporaneous edges.

Infer

Sparsification by:

- → empirical Bayes,
- → largest elements.



1.7: Support determination + ..

```
# support determination of A
zerosA <- sparsifyVAR1(A=Ahat, SigmaE=symm(solve(Phat)),</pre>
                        threshold="top", top=25,
                        statistics=FALSE, verbose=FALSE) $zeros
# support determination of precision matrix
zerosP <- sparsify(Phat, threshold="top", top=10,</pre>
                    output="light", verbose=FALSE) $zeros
# format precision support
supportP <- support4ridgeP(zeros=zerosP, nNodes=nrow(Y))</pre>
# optimal penalty parameter determination
optLambdas <- optPenaltyVAR1(Y, lambdaMin=rep(10^(-5), 2),
                              lambdaMax=c(10, 0.1),
                              lambdaInit=c(5, 0.01),
                              zerosA=zerosA, zerosP=zerosP,
                              cliquesP=supportP$cliques,
                              separatorsP=supportP$separators,
                              zerosAfit="sparse")
```



1.8: Re-estimation

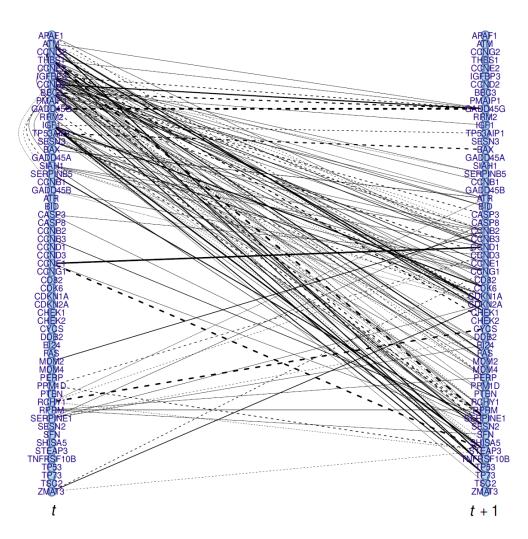
```
# re-fit the model
VAR1hat
               <- ridgeVAR1(Y=Y, lambdaA=optLambdas[1],</pre>
                             lambdaP=optLambdas[2],
                             zerosA=zerosA,
                             cliquesP=supportP$cliques,
                             separatorsP=supportP$separators,
                             zerosP=zerosP, zerosAfit="sparse")
# extract parameter estimates
Ahat <- VAR1hat$A; Phat <- VAR1hat$P
# add row and column names
rownames (Ahat) <- colnames (Ahat) <- rownames (Phat) <-
                   colnames(Phat) <- rownames(hpvP53rna)</pre>
# time-series chain graph
graphVAR1(Ahat, Phat, nNames=rownames(Ahat), type="TSCG",
          vertex.label.cex=0.5, vertex.label.font=1,
          vertex.size=4, edge.width=1.5)
```

HPV-induced oncogenesis

- → P53 pathway,
- \rightarrow 64 genes,
- → fit + LOOCV,
- \rightarrow sparsify A and Ω ,
- → re-fit + LOOCV.

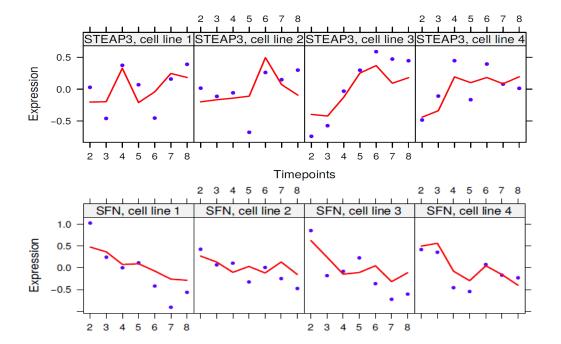
Biology

- → P53 knocked-out by HPV: zero connections
- → HPV-related IGF1/IGFBP3: many connections.

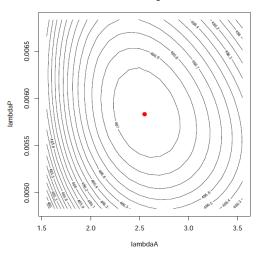


Diagnostics

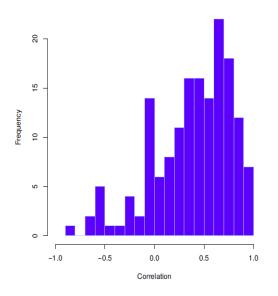
- → LOOCV contourplot,
- → fit vs. data.



cross-validated log-likelihood

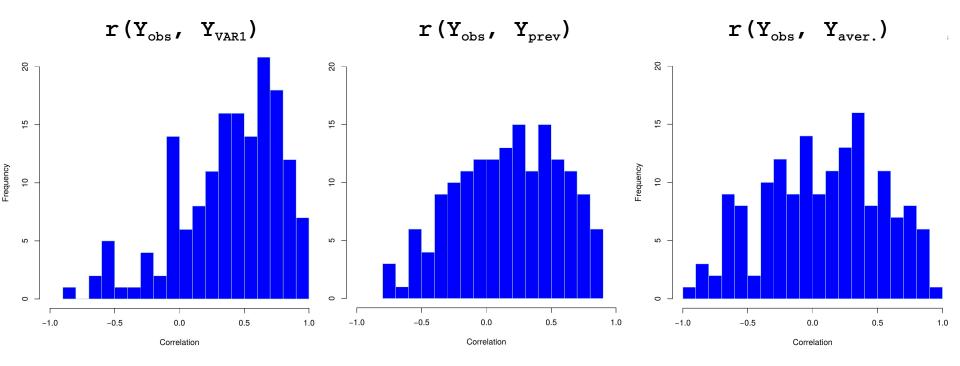


Histogram of the correlation fit vs. observation



Fit better than:

- → Previous observation,
- → Average other cell lines.





1.9: Evaluate fit

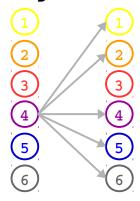
```
# calculate fits and fit-vs-observation correlations
Yhat <- array(dim=dim(Y))</pre>
for (i in 1:dim(Y)[3]) {
   Yhat[, -1, i] <- Ahat %*% Y[, -dim(Y)[2], i]
# calculate fits and fit-vs-observation correlations
for (i in 1:dim(Y)[3]) {
   Yhat[, -1, i] <- Ahat %*% Y[, -dim(Y)[2], i]</pre>
corFit <- numeric()</pre>
for (j in 1:dim(Y)[1]) {
    slHelper <- numeric()</pre>
    for (i in 1:dim(Y)[3]) {
      slHelper <- c(slHelper, cor(Yhat[j, -1, i],</pre>
                                     Y[j, -1, i], m="s"))
    corFit <- rbind(corFit, slHelper)</pre>
hist(corFit, xlab="Correlation", ylab="Frequency", n=20)
```

Node analysis

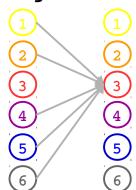
	$\deg^-(\mathbf{A})$	$\deg^+(\mathbf{A})$	between.	close.	eigen centr.
BBC3	0	17	17	0.00192	1.00000
CCND2	0	12	18	0.00187	0.68783
IGF1	1	14	0	0.00191	0.97635
IGFBP3	0	16	7	0.00190	0.87513
THBS1	0	11	0	0.00188	0.87717
CCNG1	6	0	0	0.00177	0.25154
CDKN2A	12	0	0	0.00181	0.49508
SERPINE1	8	4	0	0.00185	0.70869
SESN2	8	0	0	0.00180	0.26759
STEAP3	9	0	0	0.00179	0.36285

upper five ≈ 'regulators' lower five ≈ 'regulatees'

'regulator'



'regulatee'



Downstream

Mutual information.

A generalized 'correlation' measure, between a variate at some time and all variates at a future time point (given the past):

$$\mathcal{I}(\mathbf{Y}_{*,t+\tau}, Y_{j,t} \mid \mathbf{Y}_{*,t-1}) = \mathcal{H}(\mathbf{Y}_{*,t+\tau} \mid \mathbf{Y}_{*,t-1}) - \mathcal{H}(\mathbf{Y}_{*,t+\tau} \mid Y_{j,t}, \mathbf{Y}_{*,t-1})$$

where:

$$\mathcal{H}(\mathbf{Y}_{*,t+\tau} \mid \mathbf{Y}_{*,t-1}) = \log(|\mathbb{V}(\mathbf{Y}_{*,t+\tau} \mid \mathbf{Y}_{*,t-1})|)$$

```
E.g.:

'regulator' (deg^{+}=17):

\rightarrow I(all_{t+1}, BBC3_{t} | ...) = 0.05605

'regulatee' (deg^{-}=12):

\rightarrow I(all_{t+1}, CDKN2A_{t} | ...) = 0.00000
```

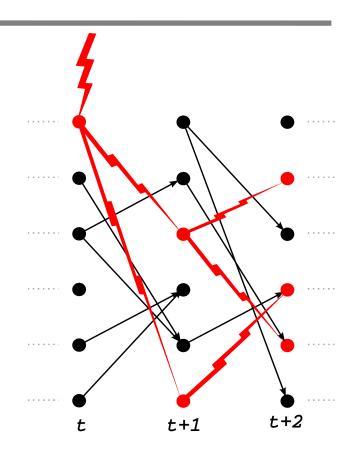
Downstream

Impulse response.

Assess the change in variates at future time points due to a change in a variate at the current time:

$$\frac{\partial \mathbf{Y}_{*,t+\tau}}{\partial \boldsymbol{\varepsilon}_{*,t}} = \mathbf{A}^{\tau}$$

Facilitates prediction of knock-out effect.



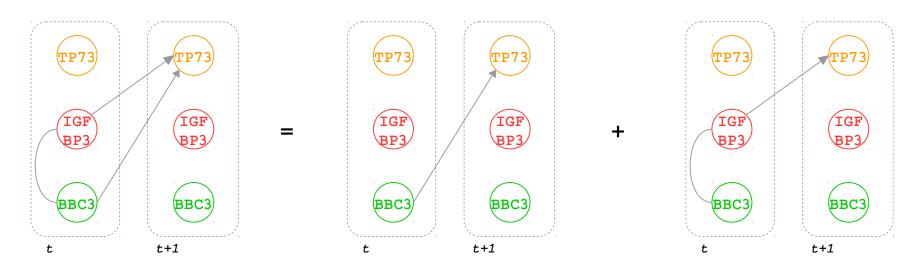
E.g.:

- \rightarrow 'regulator' (deg⁺=17): effect of BBC3_t on all_{t+1}= 0.01497
- \rightarrow 'regulatee' (deg=12): effect of CDKN2A_t on all_{t+1} = 0.00000

Downstream

Covariance decomposition.

Given TSCG, decompose conditional covariance in terms of paths:



-0.003168 = -0.002483 + -0.0006845



1.10: Down-stream analysis

```
# calculate node-wise network stats
nodeStats <- nodeStatsVAR1(Ahat, Phat, as.table=TRUE)</pre>
# show node degree table
print(nodeStats[, 1:7])
# specify time lag
lag <- 1
# evaluate mutual informations with specified lag
MIs <- mutualInfoVAR1(Ahat, solve(Phat), lag)
# evaluate impulse response with specified lag
IRs <- impulseResponseVAR1(Ahat, lag)</pre>
```

References + contributors

- → Viktorian Miok
- → Renske Steenbergen
- → Saskia Wilting



Based on:

Miok, V., Wilting, S.M., Van Wieringen, W.N. (2017), "Ridge estimation of the VAR(1) model and its timeseries chain graph model from multivariate timecourse omics data", Biometrical Journal, 59(1), 172-191.





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