

1-(a) From lecture, if $x_i | \theta \stackrel{iid}{\sim} N(\theta, 1)$, $i=1, \dots, n$,

\bar{x} is a sufficient statistic, & we know $\bar{x} | \theta \sim N(\theta, \frac{1}{n})$

Due to the sufficiency principle, we know $\pi(\theta | x_1, \dots, x_n) = \pi(\theta | \bar{x})$

$$\pi(\theta | \bar{x}) \propto f(\bar{x} | \theta) \pi(\theta)$$

$$\propto \exp \left\{ -\frac{n(\bar{x} - \theta)^2}{2} \right\} \cdot \exp \left\{ -\frac{(\theta - \mu)^2}{2\tau^2} \right\}$$

$$\propto \exp \left[-\frac{1}{2} \left\{ \left(n + \frac{1}{\tau^2} \right) \theta^2 - 2 \left(n\bar{x} + \frac{\mu}{\tau^2} \right) \theta \right\} \right]$$

$$\Rightarrow \text{a kernel for } N \left(\left(n + \frac{1}{\tau^2} \right)^{-1} \left(n\bar{x} + \frac{\mu}{\tau^2} \right), \left(n + \frac{1}{\tau^2} \right)^{-1} \right)$$

$$\Rightarrow \eta = \left(n + \frac{1}{\tau^2} \right)^{-1} \left(n\bar{x} + \frac{\mu}{\tau^2} \right) = \frac{n\tau^2 \bar{x} + \mu}{n\tau^2 + 1}$$

$$v^2 = \left(n + \frac{1}{\tau^2} \right)^{-1} = \frac{\tau^2}{n\tau^2 + 1}$$

(b)

(2)

$$p(\pi, a | \bar{x}) = E(L(\theta, a) | \bar{x})$$

$$= E(e^{c(a-\theta)} - c(a-\theta) - 1 | \bar{x})$$

$$= e^{ca} E(e^{-c\theta} | \bar{x}) - ca + cE(\theta | \bar{x}) - 1$$

$$\Rightarrow \delta^B(\bar{x}) = \delta^B(\bar{x}) = \underset{a \in \mathbb{R}}{\operatorname{argmin}} p(\pi, a | \bar{x})$$

$$\frac{\partial p(\pi, a | \bar{x})}{\partial a} = c \cdot E(e^{-c\theta} | \bar{x}) e^{ca} - c = 0$$

$$\Rightarrow e^{ca} = \frac{1}{E(e^{-c\theta} | \bar{x})}$$

$$\Rightarrow a = -\frac{1}{c} \log(E(e^{-c\theta} | \bar{x}))$$

$$\frac{\partial^2 p(\pi, a | \bar{x})}{\partial a^2} = c^2 E(e^{-c\theta} | \bar{x}) e^{ca} > 0$$

$$-c\theta | \bar{x} \sim N(-c\eta, c^2 v^2)$$

$$z = e^{-c\theta} | \bar{x} \sim \text{LogNormal}(-c\eta, c^2 v^2)$$

$$E(e^{-c\theta} | \bar{x}) = E(z | \bar{x}) = \exp\left(-c\eta + \frac{c^2 v^2}{2}\right)$$

$$\Rightarrow a = -\frac{1}{c} \cdot \left(-c\eta + \frac{c^2 v^2}{2}\right) = \eta - \frac{cv^2}{2} = \frac{n\tau^2 \bar{x} + \mu}{n\tau^2 + 1} - \frac{c\tau^2}{2(n\tau^2 + 1)}$$

(c) - i

$$\delta_1^B(\bar{x}) = \frac{n\bar{x}^2 + \mu}{n\bar{x}^2 + 1}$$

(3)

$$R(\theta, \delta_1^B) = E \left[\left(\theta - \frac{n\bar{x}^2 + \mu}{n\bar{x}^2 + 1} \right)^2 \right]$$

$$= E \left[\frac{(n\bar{x}^2 + \mu - (n\bar{x}^2 + 1)\theta)^2}{(n\bar{x}^2 + 1)^2} \right]$$

$$= \frac{1}{(n\bar{x}^2 + 1)^2} E \left[(n\bar{x}^2(\bar{x} - \theta) - (\mu - \theta))^2 \right]$$

$$\bar{x}|\theta \sim N(\theta, \frac{1}{n})$$

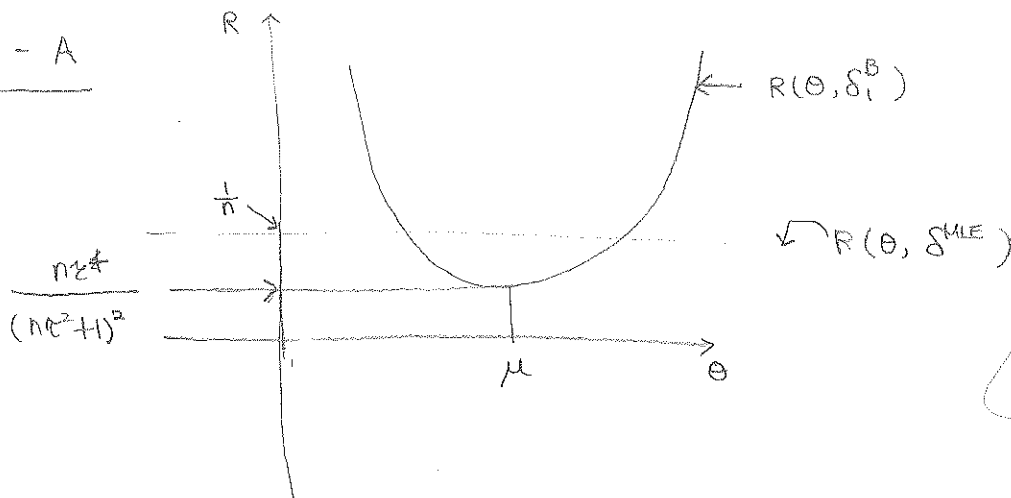
$$= \frac{1}{(n\bar{x}^2 + 1)^2} \left[\underbrace{n^2 \bar{x}^2 \frac{1}{n}}_{= n\bar{x}^2} + (\mu - \theta)^2 \right]$$

$$\frac{n\bar{x}^4}{(n\bar{x}^2 + 1)^2} = \frac{1}{\left(\sqrt{n} + \frac{1}{\sqrt{n}\bar{x}^2}\right)^2}$$

$$< \frac{1}{n}$$

$$\text{since } \frac{1}{\sqrt{n}\bar{x}^2} > 0 \text{ for } n \geq 1$$

ii - A



ii - B

From the graph, for some θ , $R(\theta, \delta_1^B) < R(\theta, \delta^{MLE})$

$\Rightarrow \delta_1^B$ is admissible

ii - C

$$\sup_{\theta \in \mathbb{R}} R(\theta, \delta_1^B) = \infty \quad \& \quad \sup_{\theta \in \mathbb{R}} R(\theta, \delta^{MLE}) = \frac{1}{n}$$

$\Rightarrow \delta^{MLE}$ is the minimax