

Loss Reserving Curves Shape Analysis

Class Project – STAT 651

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Insurance Reserving: Annual Cost Prediction

Reserving is a huge financial burden

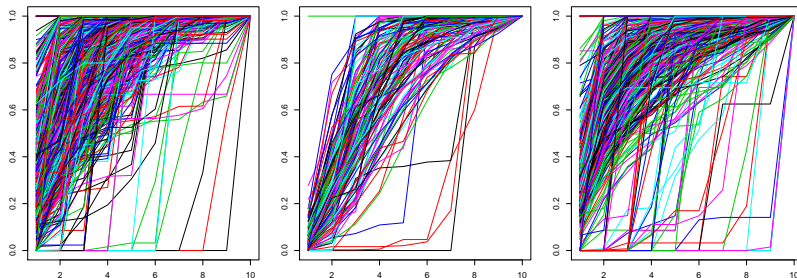
Sources of variation in annual cost:

- Standard business costs
- Claims made this year
- Payouts for claims made in past years

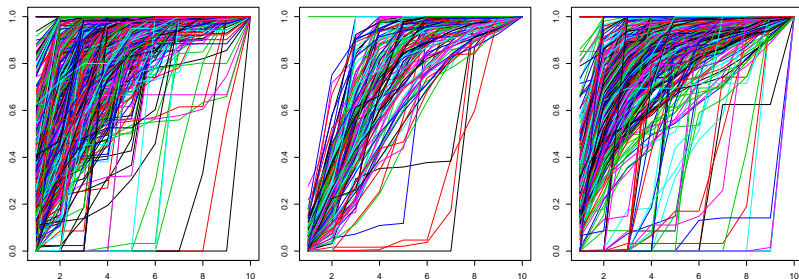
Terms:

- Accident year – Year the claim is made (incurred cost)
- Payment year – Year of claim payment (lagged or raw)
- Ultimate – Total payout for a given accident year

- Commercial Auto, Medical Malpractice, Workers Compensation
- 10 payment years per accident year (866, 130, 639)
- Cumulative percent “trajectory”



Trajectory Curve Shape Differences



- How do trajectories differ by business line?
- Reason: accurate annual predictions (cheaper)

$$y_{ik}(t) | \beta_i, \alpha, \gamma \sim N(\mathbf{H}(t)\beta_i, \alpha\gamma^{t-1})$$

$$\beta_i | \tau^2 \sim N_{10}(\mathbf{0}, \tau^2 \mathbf{P}^{-1})$$

$$\alpha \sim \text{InvGamma}(a_\alpha = 0.1, b_\alpha = 0.1)$$

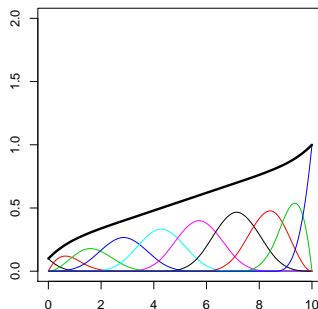
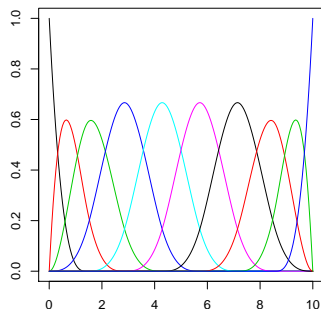
$$\gamma \sim \text{Unif}(m = 0, M = 1)$$

$$\tau^2 \sim \text{InvGamma}(a_\tau = 0.1, b_\tau = 0.1)$$

- $i \in (1, \dots, b = 3)$ denote line of business
- $k \in (1, \dots, n_i)$ denote an individual accident year
- $t \in (1, \dots, 10) \subset [1, 10]$ denotes payment year lag
- $y_{ik}(t)$ is percent of ultimate paid off at time t

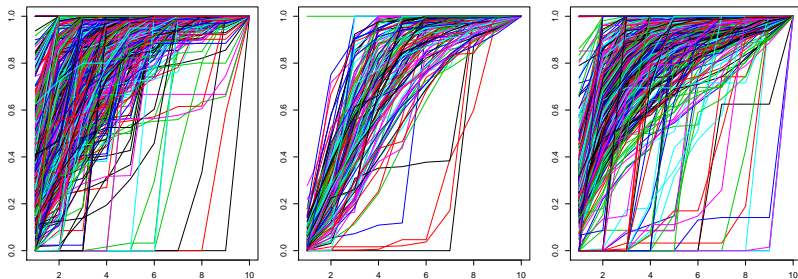
$$y_{ik}(t) | \beta_i, \alpha, \gamma \sim N(\mathbf{H}(t)\beta_i, \alpha\gamma^{t-1})$$

- $\mathbf{H}(t)$ evaluates to a row vector
- β_i weighted combination of basis functions forms curve
- Business line specific coefficients (one per year)



$$y_{ik}(t) | \beta_i, \alpha, \gamma \sim N(\mathbf{H}(t)\beta_i, \alpha\gamma^{t-1})$$

- α baseline or initial variance
- γ geometric variance decay
- $t \in (1, \dots, 10) \subset [1, 10]$ denotes payment year lag



$$y_{ik}(t) | \beta_i, \alpha, \gamma \sim N(\mathbf{H}(t)\beta_i, \alpha\gamma^{t-1})$$
$$\beta_i | \tau^2 \sim N_{10}(\mathbf{0}, \tau^2 \mathbf{P}^{-1})$$

$$\mathbf{P} = \begin{bmatrix} 2 & -1 & 0 & \dots & \dots & 0 \\ -1 & 2 & -1 & \ddots & & \vdots \\ 0 & -1 & \ddots & \ddots & \ddots & \vdots \\ \vdots & \ddots & \ddots & \ddots & -1 & 0 \\ \vdots & & \ddots & -1 & 2 & -1 \\ 0 & \dots & \dots & 0 & -1 & 1 \end{bmatrix}$$

\mathbf{P} imposes smoothness and shrinkage.

Model Summary

$$y_{ik}(t) | \beta_i, \alpha, \gamma \sim N(\mathbf{H}(t)\beta_i, \alpha\gamma^{t-1})$$

$$\beta_i | \tau^2 \sim N_{10}(\mathbf{0}, \tau^2 \mathbf{P}^{-1})$$

$$\alpha \sim \text{InvGamma}(a_\alpha = 0.1, b_\alpha = 0.1)$$

$$\gamma \sim \text{Unif}(m = 0, M = 1)$$

$$\tau^2 \sim \text{InvGamma}(a_\tau = 0.1, b_\tau = 0.1)$$

- Model trajectory curves as noisy versions of business line averages
- Variance decays since all curves approach ultimate
- Fairly diffuse priors

Model Fitting

Alternate Formulation:

$$\mathbf{y}_{ik} | \beta_i, \alpha, \gamma \sim N \left(\mathbf{H}\beta_i, \alpha \begin{bmatrix} \gamma^0 & 0 & \dots & 0 \\ 0 & \gamma^1 & \ddots & \vdots \\ \vdots & \ddots & \ddots & 0 \\ 0 & \dots & 0 & \gamma^9 \end{bmatrix} \right)$$

- Gibbs sampler, 33 total parameters
- Conditional conjugacy with β_i, α, τ^2
- Metropolis–Hastings for γ
- Utilize vectorization and stacking (Matrix package)
- Identifiability brought on by $\gamma^0 = 1$

Conjugacy Notes

Normal-Normal Case

$$\mathbf{y}|\boldsymbol{\beta}, \boldsymbol{\Sigma} \sim N_n(\mathbf{X}\boldsymbol{\beta}, \boldsymbol{\Sigma})$$

$$\boldsymbol{\beta} \sim N_p(\boldsymbol{\mu}, \mathbf{V})$$

$$\boldsymbol{\beta}|- \sim N_p((\mathbf{X}'\boldsymbol{\Sigma}^{-1}\mathbf{X} + \mathbf{V}^{-1})^{-1}(\mathbf{X}'\boldsymbol{\Sigma}^{-1}\mathbf{y} + \mathbf{V}^{-1}\boldsymbol{\mu}), (\mathbf{X}'\boldsymbol{\Sigma}^{-1}\mathbf{X} + \mathbf{V}^{-1})^{-1})$$

Normal-Inverse Gamma Case

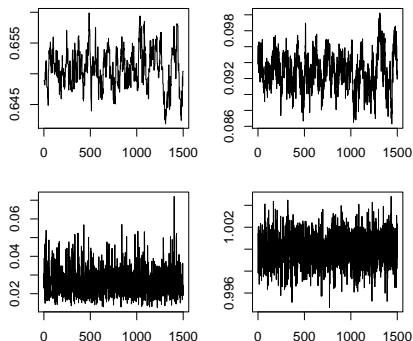
$$\mathbf{y}|\boldsymbol{\mu}, \sigma^2 \sim N_n(\boldsymbol{\mu}, \sigma^2\mathbf{R})$$

$$\sigma^2 \sim IG(a, b)$$

$$\sigma^2|- \sim IG\left(\frac{n}{2} + a, \frac{1}{2}(\mathbf{y} - \boldsymbol{\mu})'\mathbf{R}^{-1}(\mathbf{y} - \boldsymbol{\mu}) + b\right)$$

Convergence

- \hat{R} : minimum = 0.9996, maximum = 1.054
- 10 chains, 2000 iterations, 500 burnin
- 2 iterations per second, 16x speed increase with Matrix package
- Lowest effective sample size: 800 for γ , 1000 for α (out of 15000)



Sensitivity and Model Fit

Considerations:

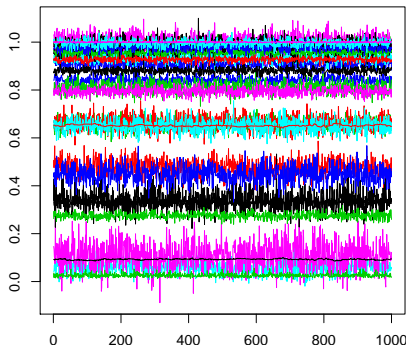
- $IG(2, 20)$ variance priors impose large expected value
- Remove γ
- Use DIC with alternate pD computation to compare fits

	Original Model	Inflated Variance Prior	Removed γ
DIC	-24041.3	-24008.2	-16271.5

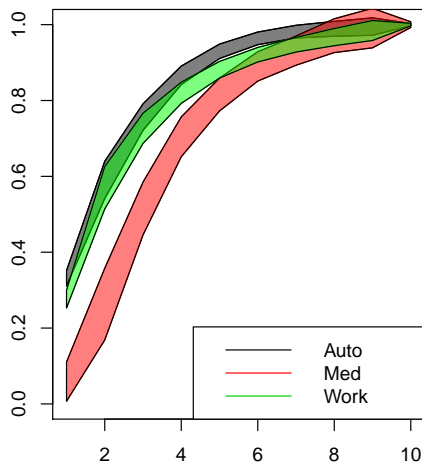
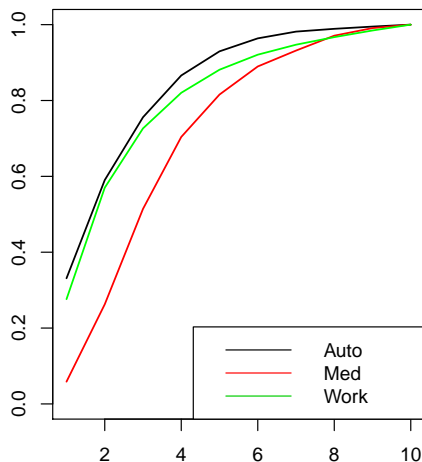
Results

Auto	$\beta_{1,1}$	$\beta_{1,2}$	$\beta_{1,3}$	$\beta_{1,4}$	$\beta_{1,5}$	$\beta_{1,6}$	$\beta_{1,7}$	$\beta_{1,8}$	$\beta_{1,9}$	$\beta_{1,10}$	α
Bayes	0.332	0.475	0.659	0.841	0.931	0.971	0.988	0.99	1.001	1	0.098
Freq	0.332	0.475	0.659	0.841	0.931	0.971	0.989	0.99	1.001	1	0.022
Med	$\beta_{2,1}$	$\beta_{2,2}$	$\beta_{2,3}$	$\beta_{2,4}$	$\beta_{2,5}$	$\beta_{2,6}$	$\beta_{2,7}$	$\beta_{2,8}$	$\beta_{2,9}$	$\beta_{2,10}$	γ
Bayes	0.059	0.096	0.337	0.67	0.807	0.913	0.936	1.017	0.971	1	0.645
Freq	0.059	0.094	0.338	0.67	0.807	0.913	0.936	1.018	0.971	1	
Work	$\beta_{3,1}$	$\beta_{3,2}$	$\beta_{3,3}$	$\beta_{3,4}$	$\beta_{3,5}$	$\beta_{3,6}$	$\beta_{3,7}$	$\beta_{3,8}$	$\beta_{3,9}$	$\beta_{3,10}$	τ^2
Bayes	0.277	0.448	0.651	0.795	0.88	0.928	0.957	0.98	0.994	1	0.646
Freq	0.277	0.447	0.651	0.795	0.88	0.928	0.957	0.98	0.994	1	

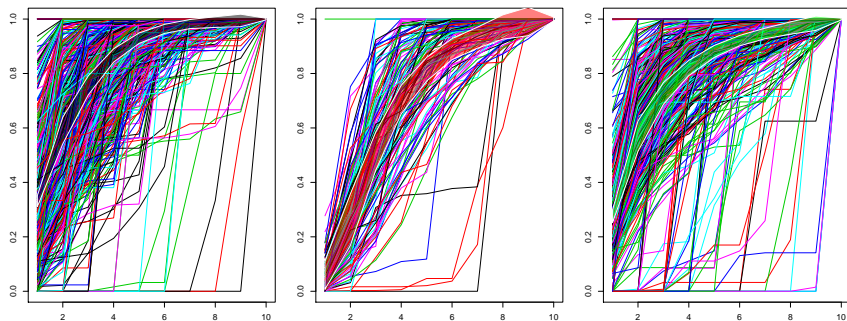
- All variables significant
- Frequentist method ignores hierarchy
- Methods match coefficients
- α reflects heteroskedasticity



Posterior Average Trajectories by Insurance Type



Posterior Average Trajectory Bands by Insurance Type



Conclusions

- Medical malpractice takes longer to reach ultimate
- Variance decay over payment years exists
- Matrix library in R
- Improvement: model differences instead of raw confidence regions
- Improvement: Experiment with restricted likelihood support
- Future: Utilize ultimate magnitude