## FORMULARIO DE ESTADÍSTICA

MODELOS DE PROBABILIDAD					
Nombre	Símbolo	$p(x) = P(X = x); F(x) = P(X \le x); f(x) = \frac{\partial F(x)}{\partial x}$	Media $\mu$	Varianza $\sigma^2$	
Bernoulli	B(p)	$p(x) = p^x q^{1-x}; x = 0, 1$	p	pq	
Binomial	B(n,p)	$p(x) = \binom{n}{x} p^x q^{n-x}; x = 0, 1,, n$	np	npq	
Geométrica	G(p)	$p(x) = pq^{x-1}; x = 1, 2, \dots$	$\frac{1}{p}$	$\frac{q}{p^2}$	
Poisson	$P(\lambda)$	$p(x) = \frac{\lambda^x e^{-\lambda}}{x!}; x = 0, 1, \dots$	λ	λ	
Uniforme continua en $(a, b)$	U(a,b)	$f(x) = 1/(b-a); a \le x \le b$	$\frac{a+b}{2}$	$\frac{(b-a)^2}{12}$	
Exponencial	$\operatorname{Exp}(\lambda)$	$f(x) = \lambda e^{-\lambda x}; x > 0$	$1/\lambda$	$1/\lambda^2$	
		$F(x) = 1 - e^{-\lambda x}; x > 0$			
		$(x-\mu)^2$			
Normal	$N(\mu, \sigma^2)$	$f(x) = \frac{1}{\sqrt{2\pi\sigma^2}} e^{-\frac{(x-\mu)^2}{2\sigma^2}}; x \in \Re$	$\mu$	$\sigma^2$	

## REGRESIÓN SIMPLE

$$\hat{y} = a + bx$$

$$a = \bar{y} - b\bar{x}$$

$$b = \frac{s_{x,y}}{s_x^2} = \frac{\hat{s}_{x,y}}{\hat{s}_x^2}$$

INFERENCIA PARA UNA POBLACIÓN				
Población	Contraste de hipótesis	Estadístico de contraste	Región de rechazo ( $p$ -valor $< \alpha$ )	Intervalo de confianza $IC_{1-\alpha}$
Cualquier v.a. $X$ con $E[X] = \mu$ , $Var[X] = \sigma^2$ y $n \to \infty$	(1) $H_0: \mu = \mu_0; H_1: \mu \neq \mu_0$ (2) $H_0: \mu \geq \mu_0; H_1: \mu < \mu_0$ (3) $H_0: \mu \leq \mu_0; H_1: \mu > \mu_0$	(a) $z_0 = \frac{\bar{x} - \mu_0}{\sigma / \sqrt{n}}$ (b) $t_0 = \frac{\bar{x} - \mu_0}{\hat{s} / \sqrt{n}}$		$\mu \in (\bar{x} \mp z_{\alpha/2}\sigma/\sqrt{n})$ $\mu \in (\bar{x} \mp z_{\alpha/2}\hat{s}/\sqrt{n})$
$\begin{array}{c} \text{Normal} \\ N(\mu, \sigma^2) \end{array}$	(1) $H_0: \mu = \mu_0; H_1: \mu \neq \mu_0$ (2) $H_0: \mu \geq \mu_0; H_1: \mu < \mu_0$ (3) $H_0: \mu \leq \mu_0; H_1: \mu > \mu_0$	(a) $z_0 = \frac{\bar{x} - \mu_0}{\sigma / \sqrt{n}}$ (b) $t_0 = \frac{\bar{x} - \mu_0}{\hat{s} / \sqrt{n}}$	$ \begin{array}{c cccc} (1a) &  z_0  > z_{\alpha/2} & (1b) &  t_0  > t_{n-1;\alpha/2} \\ (2a) & z_0 < -z_{\alpha} & (2b) & t_0 < -t_{n-1;\alpha} \\ (3a) & z_0 > z_{\alpha} & (3b) & t_0 > t_{n-1;\alpha} \\ \end{array} $	$\mu \in (\bar{x} \mp z_{\alpha/2}\sigma/\sqrt{n})$ $\mu \in (\bar{x} \mp t_{n-1;\alpha/2}\hat{s}/\sqrt{n})$
Bernoulli $B(p) \text{ con } n \to \infty$	(1) $H_0: p = p_0; H_1: p \neq p_0$ (2) $H_0: p \geq p_0; H_1: p < p_0$ (3) $H_0: p \leq p_0; H_1: p > p_0$	$z_0 = \frac{\hat{p} - p_0}{\sqrt{p_0 q_0 / n}}$	(1) $ z_0  > z_{\alpha/2}$ (2) $z_0 < -z_{\alpha}$ (3) $z_0 > z_{\alpha}$	$p \in \left(\hat{p} \mp z_{\alpha/2} \sqrt{\frac{\hat{p}\hat{q}}{n}}\right)$
$N  ext{ormal} \ N(\mu, \sigma^2)$	(1) $H_0: \sigma^2 = \sigma_0^2$ ; $H_1: \sigma^2 \neq \sigma_0^2$ (2) $H_0: \sigma^2 \geq \sigma_0^2$ ; $H_1: \sigma^2 < \sigma_0^2$ (3) $H_0: \sigma^2 \leq \sigma_0^2$ ; $H_1: \sigma^2 > \sigma_0^2$	$\chi_0^2 = \frac{(n-1)\hat{s}^2}{\sigma_0^2} = \frac{ns^2}{\sigma_0^2}$	(1) $\chi_0^2 > \chi_{n-1;\alpha/2}^2$ ó $\chi_0^2 < \chi_{n-1;1-\alpha/2}^2$ (2) $\chi_0^2 < \chi_{n-1;1-\alpha}^2$ (3) $\chi_0^2 > \chi_{n-1;\alpha}^2$	$\sigma^{2} \in \left(\frac{(n-1)\hat{s}^{2}}{\chi^{2}_{n-1;\alpha/2}}; \frac{(n-1)\hat{s}^{2}}{\chi^{2}_{n-1;1-\alpha/2}}\right)$
Cualquier v.a. $X$ con $\widehat{\theta}_{MV}$ y $n \to \infty$	(1) $H_0: \theta = \theta_0; H_1: \theta \neq \theta_0$ (2) $H_0: \theta \geq \theta_0; H_1: \theta < \theta_0$ (3) $H_0: \theta \leq \theta_0; H_1: \theta > \theta_0$	$t_0 = \frac{\hat{\theta}_{MV} - \theta_0}{\sqrt{\widehat{\text{Var}}(\hat{\theta}_{MV})}}$ $\widehat{\text{Var}}(\hat{\theta}_{MV}) = -\left(\frac{\partial^2 L(\hat{\theta}_{MV})}{\partial \theta^2}\right)^{-1}$	(1) $ t_0  > z_{\alpha/2}$ (2) $t_0 < -z_{\alpha}$ (3) $t_0 > z_{\alpha}$	$\theta \in \left(\hat{\theta}_{MV} \mp z_{\alpha/2} \sqrt{\widehat{\operatorname{Var}}(\hat{\theta}_{MV})}\right)$

INFERENCIA PARA DOS POBLACIONES				
Dos poblaciones o v.a. $X_1, X_2$	Contraste de hipótesis	Estadístico de contraste	Región de rechazo ( $p$ -valor $< \alpha$ )	
Cualesquiera con $E(X_1) = \mu_1, E(X_2) = \mu_2$ $Var(X_1) = \sigma_1^2$ $Var(X_2) = \sigma_2^2$	(1) $H_0: \mu_1 = \mu_2$ ; $H_1: \mu_1 \neq \mu_2$ (2) $H_0: \mu_1 \geq \mu_2$ ; $H_1: \mu_1 < \mu_2$ (3) $H_0: \mu_1 \leq \mu_2$ ; $H_1: \mu_1 > \mu_2$	(a) $z_0 = \frac{\bar{x}_1 - \bar{x}_2}{\sqrt{\frac{\sigma_1^2}{n_1} + \frac{\sigma_2^2}{n_2}}}$ (b) $t_0 = \frac{\bar{x}_1 - \bar{x}_2}{\sqrt{\frac{\hat{s}_1^2}{n_1} + \frac{\hat{s}_2^2}{n_2}}}$	$\begin{array}{c cccc} (1{\rm a}) &  z_0  > z_{\alpha/2} & (1{\rm b}) &  t_0  > z_{\alpha/2} & (1{\rm b}) &  t_0  > t_{\nu;\alpha/2} \\ (2{\rm a}) & z_0 < -z_\alpha & (2{\rm b}) & t_0 < -z_\alpha & (2{\rm b}) & t_0 < -t_{\nu;\alpha} \\ (3{\rm a}) & z_0 > z_\alpha & (3{\rm b}) & t_0 > z_\alpha & (3{\rm b}) & t_0 > t_{\nu;\alpha} \\ & & & & & & & & & & & & & & \\ & & & & & & & & & & & \\ & & & & & & & & & & & \\ & & & & & & & & & & \\ & & & & & & & & & & \\ & & & & & & & & & & \\ & & & & & & & & & & \\ & & & & & & & & & & \\ & & & & & & & & & \\ & & & & & & & & & \\ & & & & & & & & & \\ & & & & & & & & & \\ & & & & & & & & & \\ & & & & & & & & \\ & & & & & & & & \\ & & & & & & & & \\ & & & & & & & & \\ & & & & & & \\ & & & & & & & \\ & & & & & & & \\ & & & & & & \\ & & & & & & \\ & & & & & & \\ & & & & & & \\ & & & & & & \\ & & & & & & \\ & & & & & & \\ & & & & & & \\ & & & & & & \\ & & & & & & \\ & & & & & & \\ & & & & & \\ & & & & & \\ & & & & & \\ & & & & & \\ & & & & & \\ & & & & & \\ & & & & & \\ & & & & & \\ & & & & \\ & & & & & \\ & & & & \\ & & & & & \\ & & & \\ & & & $	
Datos pareados $D = X_1 - X_2$	(1) $H_0: \mu_D = 0; H_1: \mu_D \neq 0$ (2) $H_0: \mu_D \geq 0; H_1: \mu_D < 0$ (3) $H_0: \mu_D \leq 0; H_1: \mu_D > 0$	(a) $z_0 = \frac{\bar{d}}{\sigma_D / \sqrt{n}}$ (b) $t_0 = \frac{\bar{d}}{\hat{s}_D / \sqrt{n}}$	$\begin{array}{ccccc} (1 {\bf a})   z_0  > z_{\alpha/2} & (1 {\bf b})   t_0  > z_{\alpha/2} & (1 {\bf b})   t_0  > t_{n-1;\alpha/2} \\ (2 {\bf a})  z_0 < -z_\alpha & (2 {\bf b})  t_0 < -z_\alpha & (2 {\bf b})  t_0 < -t_{n-1;\alpha} \\ (3 {\bf a})  z_0 > z_\alpha & (3 {\bf b})  t_0 > z_\alpha & (3 {\bf b})  t_0 > t_{n-1;\alpha} \\ \text{Bajo normalidad} & \text{Si } n \to \infty & \text{Bajo normalidad} \\ \text{o si } n \to \infty & & & & & & \\ \end{array}$	
Cualesquiera con $E(X_1) = \mu_1, E(X_2) = \mu_2$ $Var(X_1) = Var(X_2) = \sigma^2$	(1) $H_0: \mu_1 = \mu_2; H_1: \mu_1 \neq \mu_2$ (2) $H_0: \mu_1 \geq \mu_2; H_1: \mu_1 < \mu_2$ (3) $H_0: \mu_1 \leq \mu_2; H_1: \mu_1 > \mu_2$	(a) $z_0 = \frac{\bar{x}_1 - \bar{x}_2}{\sigma \sqrt{\frac{1}{n_1} + \frac{1}{n_2}}}$ (b) $t_0 = \frac{\bar{x}_1 - \bar{x}_2}{\hat{s}_T \sqrt{\frac{1}{n_1} + \frac{1}{n_2}}}$ $\cos \hat{s}_T^2 = \frac{(n_1 - 1)\hat{s}_1^2 + (n_2 - 1)\hat{s}_2^2}{n_1 + n_2 - 2}$	$ \begin{array}{llll} (1 \mathrm{a}) \;  z_0  > z_{\alpha/2} & (1 \mathrm{b}) \;  t_0  > z_{\alpha/2} & (1 \mathrm{b})  t_0  > t_{n_1+n_2-2;\alpha/2} \\ (2 \mathrm{a}) \; z_0 < -z_\alpha & (2 \mathrm{b}) \; t_0 < -z_\alpha & (2 \mathrm{b}) \; t_0 < -t_{n_1+n_2-2;\alpha} \\ (3 \mathrm{a}) \; z_0 > z_\alpha & (3 \mathrm{b}) \; t_0 > z_\alpha & (3 \mathrm{b}) t_0 > t_{n_1+n_2-2;\alpha} \\ \mathrm{Bajo \; normalidad} & \mathrm{Si} \; n_1, n_2 \to \infty & \mathrm{Bajo \; normalidad} \\ \mathrm{o \; si} \; n_1, n_2 \to \infty & & & \end{array} $	
v.a. de Bernoulli $X_1{\sim}\;B(p_1),X_2{\sim}\;B(p_2)$	(1) $H_0: p_1 = p_2; H_1: p_1 \neq p_2$ (2) $H_0: p_1 \geq p_2; H_1: p_1 < p_2$ (3) $H_0: p_1 \leq p_2; H_1: p_1 > p_2$	$z_0 = \frac{\hat{p}_1 - \hat{p}_2}{\sqrt{\hat{p}_0 \hat{q}_0 \left(\frac{1}{n_1} + \frac{1}{n_2}\right)}} \operatorname{con} \hat{p}_0 = \frac{n_1 \hat{p}_1 + n_2 \hat{p}_2}{n_1 + n_2}$	(1) $ z_0  > z_{\alpha/2}$ (2) $z_0 < -z_{\alpha}$ (3) $z_0 > z_{\alpha}$ Si $n_1, n_2 \to \infty$	
v.a. Normales $X_1 \sim N\left(\mu_1, \sigma_1^2\right)$ $X_2 \sim N\left(\mu_2, \sigma_2^2\right)$	(1) $H_0: \sigma_1^2 = \sigma_2^2$ ; $H_1: \sigma_1^2 \neq \sigma_2^2$ (2) $H_0: \sigma_1^2 \geq \sigma_2^2$ ; $H_1: \sigma_1^2 < \sigma_2^2$ (3) $H_0: \sigma_1^2 \leq \sigma_2^2$ ; $H_1: \sigma_1^2 > \sigma_2^2$	$F_0 = \frac{\hat{s}_1^2}{\hat{s}_2^2}$	$ \begin{array}{c} F_0 > F_{n_1-1;n_2-1;\alpha/2} \circ F_0 < F_{n_1-1;n_2-1;1-\alpha/2} \\ \text{donde } F_{n_1-1;n_2-1;1-\alpha/2} = 1 \left/ F_{n_2-1;n_1-1;\alpha/2} \right. \\ (2) \ F_0 < F_{n_1-1;n_2-1;1-\alpha} \\ (3) \ F_0 > F_{n_1-1;n_2-1;\alpha} \end{array} $	

Dos poblaciones o v.a. $X_1, X_2$	Parámetro	Intervalo de confianza ${\rm IC}_{1-lpha}$
Cualesquiera con $E(X_1) = \mu_1, \ E(X_2) = \mu_2$ $Var(X_1) = \sigma_1^2$ $Var(X_2) = \sigma_2^2$	$\mu_1 - \mu_2$	$\mu_1 - \mu_2 \in \left(\bar{x}_1 - \bar{x}_2 \mp z_{\alpha/2} \sqrt{\frac{\sigma_1^2}{n_1} + \frac{\sigma_2^2}{n_2}}\right) \text{ Bajo normalidad o si } n_1, n_2 \to \infty$
		$\mu_1 - \mu_2 \in \left(\bar{x}_1 - \bar{x}_2 \mp z_{\alpha/2} \sqrt{\frac{\hat{s}_1^2}{n_1} + \frac{\hat{s}_2^2}{n_2}}\right) \text{ Si } n_1, n_2 \to \infty$
		$\mu_1 - \mu_2 \in \left(\bar{x}_1 - \bar{x}_2 \mp t_{\nu;\alpha/2} \sqrt{\frac{\hat{s}_1^2}{n_1} + \frac{\hat{s}_2^2}{n_2}}\right) \text{ con } \nu \approx \frac{\left(\frac{\hat{s}_1^2}{n_1} + \frac{\hat{s}_2^2}{n_2}\right)^2}{\frac{1}{n_1 - 1} \left(\frac{\hat{s}_1^2}{n_1}\right)^2 + \frac{1}{n_2 - 1} \left(\frac{\hat{s}_2^2}{n_2}\right)^2} \text{ Bajo normalidad}$
Cualesquiera con $E(X_1) = \mu_1, \ E(X_2) = \mu_2$ $Var(X_1) = Var(X_2) = \sigma^2$	$\mu_1 - \mu_2$	$\mu_1 - \mu_2 \in \left(\bar{x}_1 - \bar{x}_2 \mp z_{\alpha/2}\sigma\sqrt{\frac{1}{n_1} + \frac{1}{n_2}}\right)$ Bajo normalidad o si $n_1, n_2 \to \infty$
		$\mu_1 - \mu_2 \in \left(\bar{x}_1 - \bar{x}_2 \mp z_{\alpha/2} \hat{s}_T \sqrt{\frac{1}{n_1} + \frac{1}{n_2}}\right) \text{ Si } n_1, n_2 \to \infty$
		$\mu_1-\mu_2\in\left(ar{x}_1-ar{x}_2\mp t_{n_1+n_2-2;lpha/2}\hat{s}_T\sqrt{rac{1}{n_1}+rac{1}{n_2}} ight)$ Bajo normalidad
v.a. de Bernoulli $X_1{\sim} B(p_1), X_2{\sim} B(p_2)$	$p_1 - p_2$	$p_1 - p_2 \in \left(\hat{p}_1 - \hat{p}_2 \mp z_{\alpha/2} \sqrt{\frac{\hat{p}_1 \hat{q}_1}{n_1} + \frac{\hat{p}_2 \hat{q}_2}{n_2}}\right) \text{ Si } n_1, n_2 \to \infty$
v.a. Normales $X_1 \sim N \left(\mu_1, \sigma_1^2\right)$ $X_2 \sim N \left(\mu_2, \sigma_2^2\right)$	$\frac{\sigma_1^2}{\sigma_2^2}$	$\frac{\sigma_1^2}{\sigma_2^2} \in \left(\frac{\hat{s}_1^2}{\hat{s}_2^2} F_{n_2-1;n_1-1;1-\alpha/2}; \frac{\hat{s}_1^2}{\hat{s}_2^2} F_{n_2-1;n_1-1;\alpha/2}\right) \text{ donde } F_{n_2-1;n_1-1;1-\alpha/2} = 1 / F_{n_1-1;n_2-1;\alpha/2}$