## Exercises for Lecture 11

## Statistical Computing with R, 2023-24

## Exercise 1

Consider the following for loop:

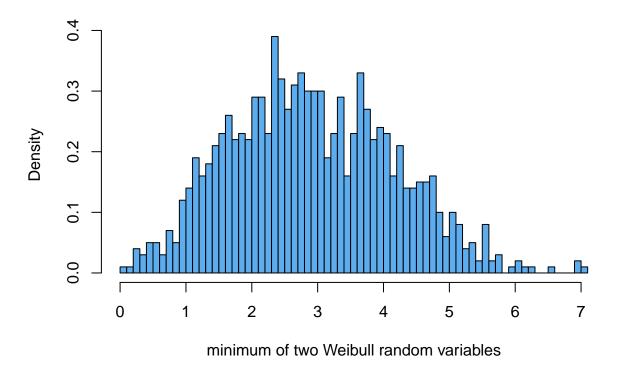
```
set.seed(13)
n.repl = 1000
weib.shape = c(2, 3)
weib.scale = c(5, 4)
min.weibs = rep(NA, n.repl)
for (i in 1:n.repl) {
    x1 = rweibull(1, shape = weib.shape[1], scale = weib.scale[1])
    x2 = rweibull(1, shape = weib.shape[2], scale = weib.scale[2])
    min.weibs[i] = min(x1, x2)
}
```

The code above implements a simulation / Monte Carlo experiment designed to evaluate the distribution of the random variable

$$Y = \min(X_1, X_2),$$

where  $X_1 \sim Weib(\lambda_1, \nu_1)$  and  $X_2 \sim Weib(\lambda_2, \nu_2)$  are two Weibull distributions with different scale  $(\lambda)$  and shape  $(\nu)$  parameters.

The estimated density of Y that can be obtained from the code above is



For the time being, you are not expected to know what a Monte Carlo experiment is, and how it can be used for estimation: we will cover this in the *Computational Statistics* course! What you need to understand, for now, is just what the code is doing so you can adjust it and reuse it to answer the questions given below.

- 1. Rewrite the for loop using replicate().
- 2. Compare the execution time of the for loop to that of the replicate() implementation using the benchmark() function. Set its argument replications = 100. Which of the two implementations is faster?
- 3. Use the code you wrote in (1) to perform two additional experiments:
- a. one with  $\lambda_1 = 0.5$ ,  $\lambda_2 = 2$ ,  $\nu_1 = 0.7$  and  $\nu_2 = 1$ ;
- b. one with  $\lambda_1 = \lambda_2 = 3$  and  $\nu_1 = \nu_2 = 1$ .

Create a histogram showing the estimated (empirical) distribution of  $Y = \min(X_1, X_2)$  in experiments (a) and (b).

## Exercise 2

During the lecture we implemented the EM algorithm for a mixture of two normal distributions whose variances were assumed to be known and equal to 1, i.e.  $\sigma_1 = \sigma_2 = 1$ . Now, let's consider the more general case of a mixture of two normals

$$f_X(x) = \pi_1 f_{X_1}(x; \mu_1, \sigma_1) + \pi_2 f_{X_2}(x; \mu_2, \sigma_2)$$

where  $X_1 \sim N(\mu_1, \sigma_1^2)$ ,  $X_2 \sim N(\mu_2, \sigma_2^2)$ , and all density parameters  $\mu_1, \mu_2, \sigma_1, \sigma_2$  and mixing proporton parameters  $\pi_1, \pi_2$  are unknown.

1. Adapt the negative log-likelihood function used in the lecture to this new problem where besides  $\mu_1$  and  $\mu_2$ , also  $\sigma_1$  and  $\sigma_2$  are unknown.

Now consider the following simulated dataset:

```
set.seed(13)
n = 2000; pi1 = 0.35
mu1 = 0.8; mu2 = 2.5
sigma1 = 0.8; sigma2 = 0.6
group = sample(1:2, n, replace = T, prob = c(pi1, 1-pi1))
table(group)

## group
## 1 2
## 698 1302

x = rep(NA, n)
x[group == 1] = rnorm(sum(group == 1), mu1, sd = sigma1)
x[group == 2] = rnorm(sum(group == 2), mu2, sd = sigma2)
```

- 2. Implement the EM algorithm to estimate  $\mu_1, \mu_2, \sigma_1, \sigma_2, \pi_1, \pi_2$  for this dataset. You may reuse and adjust suitably the code used during lecture 11 to solve this problem.
- 3. Use your algorithm implementation to estimate the parameters of interest using a single starting point. Make sure to check whether the algorithm has converged.
- 4. Consider 10 different starting points for the algorithm, and choose the solution that yields the highest loglikelihood. What are your maximum likelihood estimates of  $\mu_1, \mu_2, \sigma_1, \sigma_2, \pi_1, \pi_2$ ?
- 5. Compare your results to the true parameter values. Which inferred component corresponds to which group?
- 6. What is the percentage of misclassified observations?