

# AP

Def

1. Discrete time Markov chain
2. Continuous time random process
3. Jump time  $J_n$
4. Jump chain  $Y$
5. CTMC
6. Transition probability, time-homogeneous
7.  $P(t+s) = P(t)P(s)$
8. Holding time  $S_x$ : Memoryless property

1. Poisson Process [ $S_i \sim \exp(\lambda)$ ]
2. Markov property for Poisson process / strong markov  $\{X_{s+t} - X_s\}_{t \geq 0}$
3. Three equivalent def: Holding time, infinitesimal,  $X_t$
4. Superposition (Def 3)
5. Thinning (Def 2)
6. Conditional on the event ( $X_t = n$ ), we have  
 $f(t_1, \dots, t-n) = \frac{n!}{t^n} \mathbf{1}(0 \leq t_1 \leq \dots \leq t_n \leq t)$

1. Birth Process, [ $S_i \sim \exp(q_i)$ ], simple birth process
2.  $T = \inf_k T_k$ ,  $T \sim \text{Exp}(\sum_k q_k)$ ; The infimum is attained at a point  $T_K$  almost surely, and  
 $P(K = n) = q_n / \sum_k q_k$ ;  $T$  and  $K$  are independent
3.  $\zeta := \sum_n S_n < \infty$  explosion
4. Explosiveness equivalence in  $\sum \frac{1}{q_i}$
5. Equivalent three defs

1. Q-matrix
2. def of minimal CTMC with Q-matrix and initial distribution
3. Three constructions
4. Non-explosive if three conditions

(a)  $Y$  is ~  
(b) conditional on  $Y_0, \dots, Y_n$ ,  
 $S_1, \dots, S_{n+1}$  are indep.  $\sim \text{Exp}(q_{Y_{i-1}})$

1. Kolmogorov's forward & backward equations
2. Finite set  $P(t) = e^{tQ}$
3. Let  $I$  be a finite state space and  $Q$  be a matrix. Then it is a Q-matrix iff  $P(t) = \wedge tQ$  is a stochastic matrix for all  $t$ .
4. Let  $X$  be a right-continuous process with values in a finite set  $I$ , and let  $Q$  be a Q-matrix on  
I. Then the following are equivalent

1. Class structure
2. Recurrent and transient

3.  $x$  is recurrent for  $X$  if and only if  $\int_0^\infty p_{xx}(t) dt = \infty$ , and  $x$  is transient for  $X$  if and only if  $\int_0^\infty p_{xx}(t) dt < \infty$
4. If  $|I|$  is finite, then  $\lambda Q = 0$  if and only if  $\lambda P(s) = \lambda$  for all  $s \geq 0$
5. Positive recurrent equivalent to non-explosive + invariant distribution

inv for  $Y$   
 $\mu = q\pi$  → inv for  $X$   
 not necessarily distribution, just measure

## Queue

- $M/M/1$ :  $q(i, i+1) = \lambda, q(i, i-1) = \mu; M/M/\infty$ :  $q(i, i+1) = \lambda, q(i, i-1) = i\mu$
- Let  $\rho = \lambda/\mu$ . Then the queue length  $X$  (for a  $M/M/1$  process) is transient if and only if  $\rho > 1$ , recurrent if and only if  $\rho \leq 1$  and positive recurrent if and only if  $\rho < 1$ .  $X$  is non-explosive. In the positive recurrent case, the invariant distribution is  $\pi(n) = (1-\rho)\rho^n, n = 0, 1, \dots$ . And if  $\rho < 1$ , and  $X_0 \sim \pi$ , then the wait time (including service time) for a customer that arrives at time  $t$  is  $Exp(\mu - \lambda)$ .
- $M/M/\infty$ , The queue length  $X_t$  is positive recurrent for all  $\mu > 0, \lambda > 0$  with invariant distribution  $Poi(\rho)$  where  $\rho = \lambda/\mu$ .   
  $\forall \text{ DB, } \mu_i = \pi_i q_i, \text{ show } \mu \text{ is normalisable}$
- (Burke's Theorem). Consider an  $M/M/1$  queue with  $\mu > \lambda > 0$  or an  $M/M/\infty$  queue with  $\mu, \lambda > 0$ . At equilibrium (*i.e.*  $eX_0 \sim \pi$ ),  $D$  is a Poisson process of rate  $\lambda$  and  $X_t$  is independent of  $(D_s : s \leq t)$ .
- $(X, Y)$  is positive recurrent if and only if  $\lambda < \mu_1$  and  $\lambda < \mu_2$ . In this case, the invariant distribution is given by  $\pi(m, n) = (1-\rho_1)\rho_1^m(1-\rho_2)\rho_2^n$  where  $\rho_1 = \lambda/\mu_1, \rho_2 = \lambda/\mu_2$
- traffic equation:  $\bar{\lambda}_i = \bar{\lambda}_i + \sum_{j=1, j \neq i}^N \lambda_j p_{ji}$
- Jackson network positive recurrent
- $M/G/1$  queue: transition matrix, recurrent

## Renewal

- Definition  $\xi$  (holding time)  $T_n$  (jump time),  $N_t$ : renewal process
- If  $E\xi = 1/\lambda < \infty$  then as  $t \rightarrow \infty, N_t/t \rightarrow \lambda$  almost surely
- Size-biased picking:  $S_i, Y_i = S_i/S_n U, \hat{Y}_i$
- $f_{\hat{Y}}(y) \propto y f_{Y_1}(y), P(\hat{Y} \in dy) = ny P(Y_1 \in dy)$
- Let  $X$  be a non-negative random variable with distribution  $\mu$  and  $EX = m < \infty$ , Then the size-biased distribution of  $\mu$  is  $\hat{\mu}(dy) = y\mu(dy)/m$ . We write  $\hat{X}$  for a random variable with distribution  $\hat{\mu}$
- $A(t) = t - T_{N_t}, E(t) = T_{N_{t+1}} - t, L(t) = T_{N_{t+1}} - T_{N_t} = A(t) + E(t)$
- r.v. is arithmetic if  $P(\xi \in kZ) = 1$  for some  $k > 1, k \in \mathbb{Z}$
- $(L(t), E(t)) \rightarrow (\hat{\xi}, U\hat{\xi}), P(U\hat{\xi} \leq y) = \lambda \int_0^y P(\xi > z) dz$
- Renewal Reward
- $(\xi_i, R_i); R(t) = \sum_{i=1}^{N_t} R_i$
- $R(t)/t \rightarrow E[R]/E[\xi]$
- $\gamma(t) \rightarrow \lambda E(R\xi); \gamma(t) = E[R_{N_{t+1}}]$
- regenerative
- Little's formula:  $(\tau_n), N$ : arrival process,  $W_i$ : waiting time(including service); long-run queue  $L = \int_0^t X_s ds/t$ , waiting time  $W = (W_1 + \dots + W_n)/n$ , arrival rate  $\lambda = N_t/t$ . Then  $L = W\lambda$