

## Applied Probability 2019

1.) (a) We note that  $(T_n)_{n \geq 0}$  is the hold jump times,  
 $(V_n)_{n \geq 0}$  is the holding time of the  
 continuous markov chain

$\therefore$  At state  $i$ , holding time  $\sim \text{Exp}(g(i))$  Moves away at rate  $i$   
 Chain moves to state  $j$  ( $j \neq i$ ) at  $n$  with probability  
 $P_{ij}$

By applying thinning:  $i$  moves to  $j$  at rate  $g(i) P_{ij}$

$\therefore$  Process has Q-matrix:  $Q_{ij} = g(i) P_{ij} \mathbf{1}_{j \neq i} - g(i) \mathbf{1}_{j=i}$

(4/4)

$\therefore$  Markov Chain (continuous)

(0/17)

why is  $X_t$  defined for all  $t \geq 0$ ?

(b)  $X$  is irreducible if  $\forall x, y \in S, \exists$  sequence  $X = X_0, X_1, \dots = X_n = y$   
 s.t.  $Q(X_i, X_{i+1}) > 0$  for  $0 \leq i < n$

(2/2)

$x \in S$  is recurrent:  $\mathbb{P}(\{t: X(t) = x\} \text{ is unbounded}) = 1$

trc recurrent:  $\mathbb{E}_x [\inf \{t \geq 0: X_t = x\}] < \infty$ ,

(2/2)

$$\tau_1 = \inf \{t \geq 0: X_t \neq X_0\}$$

(c)  $X$  is irreducible iff  $\forall x, y \in S, \exists x_0, x_1, \dots, x_n$  s.t.  $x = x_0, y = x_n$   
 and  $Q(X_i, X_{i+1}) > 0$  ( $i \geq 0$ )

iff  $\forall x, y \in S, \exists x = x_0, \dots, x_n = y$  s.t.

$$P(X_i, X_{i+1}) > 0 \quad (i \geq 0) \quad \text{as } P = (X_i, X_{i+1}) =$$

$$Q(X_i, X_{i+1}) / g(X_i)$$

and  $X_i \neq X_{i+1}$  else  $Q(X_i, X_{i+1}) < 0$ .

(3/3)

iff  $Y$  is irreducible.

If  $Y$  is recurrent:  $\forall x \in S$ ,  $(Y_n)_{n \geq 0}$  visits  $x$  i.o. (A.S.)  
where  $Y_0 = x$ .

$\therefore$  If  $X_0 = Y_0 = 0$ ,  $X_{T_n} = x$  for infinite  $n$   
 $(T_n)_{n \geq 0}$  are jump times.

$Y$  recurrent  $\Rightarrow X$  non explosive  $\Rightarrow T_n \rightarrow \infty$  A.S.

$\therefore \{t: X_t \geq 0\}$  is unbounded (A.S.)

$\exists x \in S$ ,

If  $Y$  is transient:  $\exists N \in \mathbb{N}$  s.t.  $\mathbb{P}_x(Y_n \neq x \text{ for } n \geq N) > 0$

$$\hookrightarrow \{Y_n \neq x \text{ for } n \geq N\} \Rightarrow \{t: X_t = x\} \subseteq [0, T_{NN}]$$

$T_N$  bounded A.S.  $\Rightarrow \mathbb{P}_x(\{t: X_t = x\} \text{ bounded}) >$

$$\mathbb{P}_x(Y_n \neq x \text{ for } n \geq N) > 0$$

$\therefore X$  transient.

(313) (d) If  $\pi \cdot Q = 0$

$$\hookrightarrow \sum_{i: i \neq j} \pi_i Q_{i,j} = \pi_j g(j)$$

$$\text{LHS} = \sum_{i: i \neq j} \pi_i \cdot g_i \frac{Q_{i,j}}{g(j)} = \sum_{i: i \neq j} \pi_i g_i P_{i,j} = \sum_i \pi_i g_i P_{i,j} \quad (P_{i,i} = 0)$$

$\therefore (\pi_i g_i)_{i \in S}$  is an invariant measure of  $Y$ .

$Y$  +ve recurrent  $\Rightarrow$  Invariant measure unique (up to constant)

$$\Lambda = \sum_{i \in S} \pi_i g_i$$

$(\pi_i g_i / \Lambda)_{i \in S}$  is the invariant distribution.

You were asked to give invariant of  $X$   
using invariant of  $Y$ .

2.) (a)  $S = \mathbb{Z}_{\geq 0}$  is state space:

$$\begin{aligned} P(i, j) &= \lambda_i / (\lambda_i + \lambda_i P_i) & \text{if } j = i+1 \\ &= \lambda_i P_i / (\lambda_i + \lambda_i P_i) & \text{if } j = 0, i \neq 0 \\ &= 0 & \text{otherwise} \end{aligned}$$

WLOG:

$$(P_0 = 0)$$

(b)  $IP(N \rightarrow 0) > 0$  for  $N \geq 1$ . ("N → 0" represents chain moving from  $N$  to 0)

But  $IP(0 \rightarrow N) = 0$ .  $\therefore$  If we watch the reversed process, it cannot have the same distribution as the original one as  $N \rightarrow 0$  is not possible in the reversed version.

Not needed.

(c) Reversed  $Q$ -matrix:  $\hat{Q}(i, j) = \pi_j / \pi_i \cdot Q(i, j)$

$$\therefore \hat{Q}(n, 0) = \pi_0 / \pi_n \cdot Q(0, n) = 0 \neq \hat{Q}(0, n, 0) > 0$$

for  $n \geq 1$

(c1) Consider jump chain:  $P(i, j) = \frac{1}{1+P_i}$  if  $j = i+1$   
 $= \frac{P_i}{1+P_i}$  if  $j = 0$

Since chain is clearly irreducible:  $0 \rightarrow 1, \dots \rightarrow n \rightarrow 0 \Rightarrow 0$  communicates with all states.  $\Rightarrow$  irreducible.

$Y$  recurrent iff  $0$  is  $Y$ -recurrent ( $Y$  is jump chain)  
 iff  $IP_0(Y \text{ returns to } 0) = 1$   
 iff  $IP_0(Y \text{ never returns to } 0) = 0$

$$\text{iff } \prod_{k \geq 1} \frac{\lambda_k}{1+P_k} = 0 \quad \text{iff } \prod_k (1+P_k) = \infty$$

$\therefore X$  transient iff  $Y$  transient iff  $\prod_k (1+P_k) < \infty$

(c)  $\prod_k (1 + p_k) < \infty \Rightarrow$  Chain is transient. Let  $(T_n)_{n \geq 0}$  be jump times

Consider jump chain: (A.S.)  $\exists N$  s.t.  $\forall n \quad Y_{n+N} = N$ .

$\therefore \zeta = \sum_{k=1}^{N-1} T_N + \sum_{k \geq 0} E_{N+k} / \lambda_k (1 + p_k)$ ,  $(E_n)_{n \geq 1}$  is iid  $\text{Exp}(1)$ , independent of jump chain

\* Since  $T_N < \infty$  A.S.:  $\zeta = \infty$  iff  $\sum_{k \geq 0} E_{k+N} / \lambda_k (1 + p_k) = \infty$

Consider a birth chain,  $Q(n, n+1) = \lambda_n (1 + p_n)$

Chain is explosive A.S. if  $\sum_n 1/\lambda_n (1 + p_n) < \infty$

Chain is non-explosive if  $\sum_n 1/\lambda_n (1 + p_n) = \infty$

But chain explosive iff  $\zeta = \sum_{k \geq 0} E'_k / \lambda_k (1 + p_k) < \infty$

$(E'_k)_{k \geq 1} \sim \text{iid Exp}(1)$

$\therefore \sum_{k \geq 0} E_{k+N} / \lambda_k (1 + p_k) < \infty$  iff  $\sum_{k \geq 0} 1/\lambda_k (1 + p_k) < \infty$

Since this is independent of  $N$ : Chain is A.S. explosive

if  $\sum_{k \geq 0} 1/\lambda_k (1 + p_k) < \infty$ , non-explosive otherwise.

(c)  $n \geq 1$ :  $\pi_n \cdot \lambda_n (1 + p_n) = \pi_{n-1} - \lambda_{n-1}$

$$\begin{aligned} \therefore \pi_n &= \frac{\lambda_{n-1}}{\lambda_n (1 + p_n)} \cdot \pi_{n-1} = \frac{\lambda_{n-1} \dots \lambda_0}{\lambda_n \dots \lambda_1} \cdot \frac{\pi_0}{(1 + p_n) \dots (1 + p_1)} \\ &= \frac{\lambda_0 \pi_0}{\lambda_n \prod_{k=1}^n (1 + p_k)} \end{aligned}$$

Check:

$$\pi_0 = 1, \quad \pi_0 \lambda_0 = \sum_{n \geq 1} \pi_n \lambda_n p_n$$

$$\text{RHS} = \lambda_0 \pi_0 \sum_{n \geq 1} P_n / \prod_{k=1}^n (1 + P_k)$$

$$1 - \sum_{n=1}^{\infty} P_n / \prod_{k=1}^n (1 + P_k) = 1 / \prod_{k=1}^{\infty} (1 + P_k)$$

$$\therefore \text{RHS} = \text{LHS} \quad \text{iff} \quad \prod_{k=1}^{\infty} (1 + P_k) = \infty$$

$$\therefore \text{Invariant measure exists} \quad \text{iff} \quad \prod_{k=1}^{\infty} (1 + P_k) = \infty$$

~~Invariant distribution exists  $\Rightarrow$  Invariant measure exists~~

Invariant measure  $\Leftarrow$

$$\text{Invariant distribution exists} \quad \text{iff} \quad \prod_{k=1}^{\infty} (1 + P_k) = \infty \quad \text{and}$$

$$\sum_{n \geq 1} \frac{1}{\lambda_n} \prod_{k=1}^n (1 + P_k) < \infty. \quad \checkmark$$

3) ca) Reversible in Equilibrium: For  $T \geq 0$ ,  $(X_t)_{t \geq 0}$  is reversible in equilibrium if  $\forall T \geq 0$ ,  $(X_t)_{t \in [0, T]}$ ,  $(X_{t+T-t})_{t \in [0, T]}$

(2/2) have the same distribution

Detailed Balance: Let  $Q$  be generator

(2/2)  $\forall x, y \in S = \text{stat space: } \pi_x Q(x, y) = \pi_y Q(y, x)$  (happy!)

If  $\pi$  is in detailed balance:  $\sum_x \pi_x Q_{xy} = \sum_x \pi_y Q_{yx} = \pi_y \sum_x Q_{yx} = 0$

(3/3)  $\therefore \pi$  is invariant

cb) Generator:  $Q$ ;  $S = \{0, \dots, N+s\}$ .

$$Q(n, n+1) = \lambda$$

$$Q(n+1, n) = \min\{n, s\} \cdot \mu$$

$$Q(n, n) = -(\lambda \mathbb{1}_{n \neq N+s} + \min\{n, s\} \mu)$$

$$Q = 0 \text{ otherwise!}$$

There are finite states; or Since  $Q(n, n \pm 1) > 0$  for  $n, n \pm 1 \in S$ :  $X_t$  is irreducible.

(1/2) Finite states  $\Rightarrow X$  is Jump chain recurrent  $\Rightarrow X$  recurrent

(1/2)  $\therefore$  Any invariant measure is unique up to constant

(as non-explosive) non-explosive?

Solve Detailed Balance:

$$0 \leq k \leq s-1: \pi_k \cdot \lambda = \pi_{k+1} (k+1) \mu \Rightarrow \pi_{k+1} = \pi_0 \cdot \left(\frac{\lambda}{\mu}\right)^{k+1} / (k+1)!$$

$$k \geq s: \pi_k \lambda = \pi_{k+1} (s\mu) \Rightarrow \pi_k = \pi_s \left(\frac{\lambda}{s\mu}\right)^{k-s}$$

$$\therefore \pi = \underline{\pi}: \pi_k = \pi_0 \left(\frac{\lambda}{\mu}\right)^k / k! \quad (0 \leq k \leq s)$$

$$= \pi_0 \left(\frac{\lambda}{\mu}\right)^s / s! \cdot \left(\frac{\lambda}{s\mu}\right)^{k-s} \quad (k \geq s)$$

After normalising: Distribution is unique invariant distribution

$$\Lambda = \sum_{k=0}^S (\lambda/\mu)^k / k! + \sum_{k=S+1}^{S+N} (\lambda/\mu)^S / S! (\lambda/S\mu)^{k-S} < \infty$$

$\tilde{\pi}_k = \pi_k / \Lambda$  is invariant distribution

Since we have solved detailed balance,  $X$  is reversible in equilibrium (\*)

(\*) :  $X$  is reversible in equilibrium iff Detailed Balance has a solution

4.) (a)  $\Pi$  is a Poisson process if:

$\forall$  Bounded, Measurable  $A \in \mathbb{R}^n$ :  $|\Pi \cap A| \sim \text{Poi}(\int_A \lambda)$

If  $A_m \in \mathcal{B}(\mathbb{R}^n)$ ,  $(A_m)_{m \geq 1}$  disjoint:  $(\Pi \cap A_m)_{m \geq 1}$  is independent sequence of R.V.

(b) If  $\forall x \in \mathbb{R}^S$ ,  $\int_{f^{-1}\{x\}} \lambda \, d\mu = 0$  and

$\forall$  Bounded, measurable  $B \in \mathbb{R}^S$ :  $\int_{f^{-1}(B)} \lambda < \infty$

Mean measure:  $\tilde{\Lambda}(B) = \int_{f^{-1}(B)} \lambda = \Lambda \cdot f^{-1}(B)$

$\therefore \tilde{\Lambda} = \Lambda \cdot f^{-1}$  (set function)



cc)

$f^{-1}\{r\} = \{x \in \mathbb{R}^2 : |x|^2 = r\}$  has 0 - Lebesgue measure

$\forall R \geq 0 : \mu(f^{-1}[0, R^2]) = \pi R^2 < \infty$  ( $\mu$  is Lebesgue measure)

$\therefore \forall$  Bounded  $B, \mu(f^{-1}(B)) < \infty$

$\therefore f(\Pi)$  is a Poisson process

$$\tilde{\Lambda}(B) = \Lambda \cdot f^{-1}(B)$$

$$\text{If } B = [0, R^2] : \Lambda \cdot f^{-1}([0, R^2]) = \mu(\{x : |x|^2 \leq R^2\}) \\ = \pi \cdot R^2$$

$$\therefore \Lambda \cdot f^{-1} = \pi \cdot \mu \text{ on } \mathcal{A} = \{[0, R^2] : R \geq 0\} \cup \{\emptyset\}$$

$\mathcal{A}$  is a  $\pi$ -system,  $\mu$   $\sigma$ -finite measure  $\Rightarrow \Lambda \cdot f^{-1} = \pi \cdot \mu$  on  $\sigma(\mathcal{A}) = \mathcal{B}(\mathbb{R}^2)$

$\therefore \tilde{\Lambda}(B) = \mu(B) \Rightarrow$  Homogenous  $\pi$ -intensity process

Since  $f(\Pi)$  is homogenous poisson process :

~~$\text{IP}(R_k \leq x)$~~   $\leftarrow R$  let  $S_k = R_k^2$ :

~~$S_0, S_1, S_2, \dots, S_{n-1}, S_n$~~

$S_0 = 0 : (S_{n+1} - S_n)_{n \geq 0} \sim \text{Exp}(\pi)$  R.V.

$\therefore S_n$  is the sum of  $n$  iid  $\text{Exp}(\pi)$  R.V.

$\infty$

$$\text{IP}(R_k \leq x) = \text{IP}(R_k^2 \leq x^2) = \text{IP}(\text{Poi}(x^2 \cdot \pi) \geq k)$$

$$= \sum_{n \geq k} e^{-x^2 \pi} \frac{(x^2 \pi)^n}{n!}$$

$$\therefore f_{R_k}(x) = \sum_{n \geq k} \frac{1}{n!} e^{-x^2 \pi} \left\{ -2x\pi (x^2 \pi)^{n-1} + \frac{n (x^2 \pi)^{n-1}}{(2x\pi)} \right\}$$

$$= \sum_{n \geq k} \frac{1}{n!} (2x\pi) (x^2 \pi)^{n-1} (n - 2x\pi)$$



$$= (2 \times \pi) \cdot \left\{ \sum_{n \geq k} \frac{1}{(n-1)!} (x' \pi)^{n-1} - \sum_{n \geq k} \frac{1}{n!} (x' \pi)^n \right\} e^{-x' \pi}$$

$$= (2 \times \pi) / (k-1)! (x' \pi)^{k-1} e^{-x' \pi}$$

(919)

(202)

2018

Paper 3

ca) If  $Q_n \neq 0$ :  $(n+1)^{\text{th}}$  service starts immediately.

$$Q_{n+1} = (Q_n - 1) \quad \left( \begin{array}{l} n^{\text{th}} \text{ customer serviced} \\ + A_n \quad \text{(arrivals during } n^{\text{th}} \text{ service time)} \end{array} \right)$$

Else: Wait until customer arrives.  $Q_n \rightarrow Q_n + 1$

(Same as above)

$$\therefore Q_{n+1} = Q_n + 1 - 1 + A_n = Q_n + A_n$$

$$\therefore Q_{n+1} = Q_n + A_n - 1_{Q_n \neq 0}$$

$$A_n \sim \text{Poi}(\lambda \cdot S)$$

$$\mathbb{E}[A_n] = \mathbb{E}[\text{Poi}(\lambda \cdot S)] = \mathbb{E}[\mathbb{E}[\text{Poi}(\lambda S) | S]] = \lambda \cdot \mathbb{E}[S] = \rho$$

$$\mathbb{E}[A_n'] = \mathbb{E}[\mathbb{E}[\text{Poi}(\lambda \cdot S') | S']] = \mathbb{E}[(\lambda \cdot S + \lambda' S')] = \rho + \lambda' \mathbb{E}[S']$$

(414)

cb) If  $Q_n$  in equilibrium:

$$\mathbb{E}[Q_{n+1}] = \mathbb{E}[A_n] + \mathbb{E}[Q_n] - \mathbb{E}[1_{Q_n \neq 0}] = \mathbb{E}[Q_n]$$

$$\therefore \mathbb{E}[1_{Q_n \neq 0}] = \rho$$

$$\mathbb{E}[Q_n^2] = \mathbb{E}[Q_{n+1}^2] = \mathbb{E}[A_n'^2] + \mathbb{E}[Q_n'^2] + \mathbb{E}[1_{Q_n \neq 0}] + 2\mathbb{E}[A_n \cdot Q_n] + 2\mathbb{E}[1_{Q_n \neq 0} \cdot A_n] - 2\mathbb{E}[Q_n \cdot 1_{Q_n \neq 0}]$$

$$\therefore p + \lambda' IE[S'] + p + 2 IE[A_n] IE[Q_n] - 2 IE[A_n] IE[Q_{n+1}] - 2 IE[Q_n] = 0$$

( $A_n, Q_n$  independent)

$$\therefore 2p + \lambda' IE[S'] + 2(p-1) IE[Q_n] - 2p' = 0$$

$$(12/12) \therefore IE[Q_n] = p + \frac{\lambda' IE[S']}{2(p+1-p)}$$

\* (root)

Highly sus part: Mean waiting time =:  $\frac{IE[Q_n]}{\lambda}$  = Average queue

$IE[Q_n]$  = Average # of people in queue (Not really)

$\lambda$  = arrival intensity.

Equilibrium exists  $\Rightarrow$  Regenerative

Little: Average wait time =  $IE[Q_n] / \lambda$

$$\therefore \text{Average wait (without service) time} = -p + \frac{IE[Q_n]}{\lambda}$$

$$= \frac{1}{2} \lambda IE[S'] / (1-p)$$

Seems to work, but  $Q_n$  is length immediately after someone leaves  $\Rightarrow$  "True" average queue should be longer.

\* Proper method: Consider waiting time (without service)

$$IE[\text{wait time}] = IE[\text{Queue length}] \cdot IE[S] + IE[R]$$

$R$  is remaining time of service.

Little:  $IE[\text{Queue length}] = IE[\text{Wait time}] \cdot \lambda$

$$\therefore IE[\text{wait time}] = \frac{IE[R]}{1 - \lambda IE[S]} = \frac{IE[R]}{1-p}$$

Consider  $\int_0^{T_n} R(t) dt$ ,  $T_n$  is time after  $n^{\text{th}}$  service

By adding triangle area:

$$\int_0^{T_n} R(t) dt = \sum_{k=1}^n S_k^2 / 2$$

In the long run: # of service completed at time  $t$   
 $\rightarrow$  # of arrivals up to  $t$ .

$$\therefore \frac{1}{n} \int_0^{T_n} R(t) dt = \frac{n}{T_n} \frac{1}{n} \int_0^{T_n} R(t) dt$$

$$\frac{n}{T_n} \rightarrow \lambda, \quad \frac{1}{n} \int_0^{T_n} R(t) dt$$

$$\therefore \mathbb{E} \left[ \frac{1}{T_n} \int_0^{T_n} R(t) dt \right] \rightarrow \lambda / 2 \mathbb{E}[S']$$

$$\therefore \mathbb{E}[R] \rightarrow \lambda / 2 \mathbb{E}[S']$$

$$\therefore \mathbb{E}[\text{Wait time}] = \lambda / 2 \mathbb{E}[S'] \times$$

No.: .....

Date: .....

## Applied Probability 2020

1.) ca)  $N$  is a birth process if & it is a cont. markov chain with:  
State space =  $\mathbb{Z}_{\geq 0}$

$Q$ -matrix generator:  $Q(n, n+1) = \lambda_n$

$Q(n, n) = -\lambda_n$

$Q(n, m) = 0$  (otherwise)

W Let  $(T_n)_{n \geq 0}$  be jump times,  $S_n = T_n - T_{n-1}$  ( $n \geq 1$ )  
 $(S_n)_{n \geq 1}$  are independent,  $S_n \sim \text{Exp}(\lambda_{n-1})$

(515)  $\xi = \sum_{n \geq 1} S_n$ :  $N$  is non-explosive iff  $\mathbb{P}(\xi = \infty) = 1$

cb) If  $\sum_{n=0}^{\infty} \frac{1}{\lambda_n} = \infty$ : Let  $A_n = \{S_{n+1} \geq \frac{1}{\lambda_n}\}$ ,  $\mathbb{P}(A_n) = e^{-\lambda_n \cdot (1/\lambda_n)} = e^{-1}$

$\therefore \sum_{n \geq 0} \mathbb{P}(A_n) = \infty \Rightarrow$  1st Borel Cancelli  $A_n$  occurs i.o.

$$\xi = \sum_{n \geq 1} S_n \geq \sum_{n \geq 1} S_n \cdot \mathbb{1}_{S_n \geq 1/\lambda_{n-1}}$$

$$\mathbb{E} \left[ e^{-\sum_{n=1}^{\infty} \lambda_n S_n} \right] = \prod_{n \geq 1} \mathbb{E} \left[ e^{-S_n} \right] = \prod_{n \geq 1} \left( \frac{\lambda_{n-1}}{1 + \lambda_{n-1}} \right)$$

$$= \left( \prod_{n \geq 1} \left( 1 + \frac{1}{\lambda_{n-1}} \right) \right)^{-1} \leq \left( \sum_{n \geq 1} \frac{1}{\lambda_{n-1}} \right)^{-1} = 0.$$

$$\therefore \sum_{n \geq 1} S_n = \infty \text{ (A.S.)} \Rightarrow \text{Non-explosive}$$

$$\text{If } \sum_{n \geq 0} \frac{1}{\lambda_n} < \infty: \mathbb{E}[\xi] = \sum_{n \geq 0} \frac{1}{\lambda_n} < \infty \Rightarrow \xi < \infty \text{ A.S.}$$

$\therefore$  A.S. Explosive.

$$\text{ca) } f(t) = \mathbb{E}[N(t)] = \sum_{k \geq 0} P_{0,k}(t) \cdot k$$

$$f'(t) = \sum_{k \geq 0} k P_{0, \frac{k+1}{k}}'(t) = \sum_{k \geq 0} k \sum_{j \geq 0} Q_{jk} P_{0,j}(t) \quad (\text{Kolmogorov forward!})$$

$$= \sum_{j \geq 0} P_{0,j}(t) \underbrace{\sum_{k \geq 0} Q_{j,k}}_k$$

$$- Q_{j,j} j + (j+1) Q_{j,j+1} = \cancel{(j+1)} d j + \beta$$

$$= \beta \sum_{j \geq 0} P_{0,j}(t) + d \sum_{j \geq 0} j P_{0,j}(t) = \beta + d f(t)$$

$$f' - d f = \beta : \quad \text{A.} \quad f' - d f : \quad f(t) = C_1 e^{dt} \quad \left. \begin{array}{l} \\ \end{array} \right\} \begin{array}{l} \\ C_1 \text{ constant} \end{array}$$

$$\therefore f(t) = C_1 e^{dt} - \beta/d \text{ is a solution}$$

$$f(0) = 0 : C_1 = \beta/d ; \quad \text{C.} \quad f(t) = (\beta/d)(e^{dt} - 1)$$

2.) (a) Let  $J_0 = 0, \quad J_{n+1} = \inf \{ t \geq J_n : X_t \neq X_{J_n} \} \quad (n \geq 0)$

$$Y_n = X_{J_n}, \quad (Y_n)_{n \geq 0} \text{ is the jump chain}$$

X is irreducible if  $\forall x, y \in S: \exists \text{ sequence } x = x_0, \dots, x_N = y$   
s.t.  $Q(x_i, x_{i+1}) > 0$  for  $0 \leq i < N$

X is ~~irreducible~~ <sup>recurrent</sup> if:  $\forall \text{ stat } x, \quad \mathbb{P}_x \left( \left\{ t : X_t = x \right\} \text{ unbounded} \right) = 1$

If Y is recurrent:  $(X_t)_{t \geq 0}$  is non-explosive. ~~stat~~ state x.

" Since  $\exists$  subsequence  $n_k$  s.t.  $Y_{n_k} = x$

..  $f_t$

Fix state x: (A.S.)  $\exists n_k$  s.t.  $Y_{n_k} = x$ .

$$\therefore \{ J_{n_k} : k \geq 1 \} \subseteq \{ t : X_t = x \}$$

LHS is A.S. unbounded  $\Rightarrow \{ t : X_t = x \}$  is A.S. unbounded

$$\therefore \mathbb{P} \left( \left\{ t : X_t = x \right\} \text{ unbounded} \right) = 1$$

$\therefore$  X is recurrent.

If  $Y$  is transient:  $\exists$  state  $x$  s.t.  $IP(\bigcup_{n \geq 1} \bigcap_{m \geq n} \{Y_m \neq x\}) > 0$  ✓

$$\left\{ \left\{ t: X_t = x \right\} \text{ bounded} \right\} \stackrel{2}{=} \bigcup_{n \geq 1} \left\{ X_t \neq x \text{ for } t \geq T_n \right\} \cap \left\{ T_1, \dots, T_n \neq \infty \right\}$$

$$= \bigcup_{n \geq 1} \left\{ \bigcap_{m \geq n} \{Y_m \neq x\} \cap \left\{ T_n \neq \infty : n \in \mathbb{N} \right\} \right\}$$

$$\therefore IP(\{t: X_t = x\} \text{ bounded}) \geq IP(\bigcup_{n \geq 1} \bigcap_{m \geq n} \{Y_m \neq x\}) -$$

$$\sum_{n \geq 0} IP(T_n = \infty) = IP(\bigcup_{n \geq 1} \bigcap_{m \geq n} \{Y_m \neq x\})$$

$$= 0.$$

(818)  $\therefore X$  is also transient. ✓

ch) Jump chain: Random walk on  $\mathbb{Z}$ ,  $IP(+1) = 2/3$ ,  $IP(-1) = 1/3$   
 $\therefore$  Jump chain is transient ( $\rightarrow -\infty$ )  $\Rightarrow X$  is also transient ✓

$$(k \geq 0): (\text{Detailed Balance}): \pi_n \cdot 2 \cdot 3^n = \pi_{n+1} \cdot 3^{n+1} \Rightarrow \pi_{n+1} = \pi_n \left(\frac{2}{3}\right)$$

$$(n \leq 0): (\text{Detailed Balance}) \quad \pi_n \cdot 2 \cdot 3^{-n} = \pi_{n-1} \cdot 2 \cdot 3^{-n+1}$$

$$\therefore \pi_{n-1} = \pi_n / 6$$

$$\therefore \pi_n = \begin{cases} \left(\frac{2}{3}\right)^n \pi_0 & (n \geq 0) \\ \left(\frac{1}{6}\right)^{|n|} \pi_0 & (n \leq 0) \end{cases} \quad \left. \begin{array}{l} \text{Solve Detailed Balance} \\ \text{Invariant} \end{array} \right\}$$

$$\sum_n \pi_n = -1 + \frac{1}{1-2/3} + \frac{1}{1-1/6} = 2 + 6/5 < \infty$$

$\therefore$  Invariant distribution exists ✓

Theorem: If  $X$  is non-explosive, existence of invariant distribution  $\Rightarrow$   
 $X$  is +ve recurrent  $\Rightarrow X$  is recurrent ✓

(202)  $X$  transient  $\Rightarrow$  (Not recurrent)  $X$  is explosive. ✓



3.) (a) Renewal Process: Let  $(E_n)_{n \geq 1}$  be iid,  $\geq 0$  R.V.

$$S_k = \sum_{r=1}^k E_r$$

$$N(t) = \text{Max } \{ k \geq 0 : S_k \leq t \}$$

Let  $\{(E_n, R_n)_{n \geq 1}\}$  be iid:  $(E_n)_{n \geq 1}$  defines a renewal process.

Renewal reward theorem:  $\frac{1}{t} \sum_{n=1}^{N(t)} R_n \rightarrow \frac{E[R_1]}{E[E_1]} \quad (\text{A.S.})$

$$E\left[\frac{1}{t} \sum_{n=1}^{N(t)} R_n\right] \rightarrow E[R] / E[E]$$

(b) Lawful - Good Method.

Let  $E_n = m$ ,  $(E_n)_{n \geq 1}$  forms a renewal process.

By the memoryless property of Exponential distributions, if the machine is working after a check, the time until breakdown is still an exponential ( $\lambda$ ) R.V.

$R_n \sim \text{Min} \{ \text{Exp}(\lambda), m \}$  is the amount of time machine is working between  $(m-1)^{\text{th}}$ ,  $m^{\text{th}}$  check

(10/10)  $\therefore$  long run proportion of machine working:

(Renewal Reward theorem:)  $\sum_{n=1}^{N(t)} R_n \leq R(t) \leq \sum_{n=1}^{N(t)+1} R_n$

$$\therefore \frac{R(t)}{t} \rightarrow \frac{E[R]}{m} \quad (\text{A.S.})$$

$$E[R] = e^{-\lambda m} \cdot m + \int_0^m \lambda e^{-\lambda x} x \, dx$$

$$\therefore e^{-\lambda m} + \int_0^m \frac{1}{m} \lambda e^{-\lambda x} x dx \quad \text{proportion}$$

\* Evaluate:  $e^{-\lambda m} \int_0^m x \lambda e^{-\lambda x} dx = [x e^{-\lambda x}]_0^m + \int_0^m e^{-\lambda x}$

$$= -m e^{-\lambda m} + \frac{1}{\lambda} [-e^{-\lambda x}]_0^m$$

$$= -m e^{-\lambda m} + \frac{1}{\lambda} [1 - e^{-\lambda m}]$$

$$\therefore e^{-\lambda m} + \frac{1}{\lambda m} (1 - e^{-\lambda m}) = e^{-\lambda m}$$

$$= \frac{1 - e^{-\lambda m}}{\lambda m}$$

Chaotic Evil Method:

Let  $E_n$  be time between  $n-1^{\text{th}}$ ,  $n^{\text{th}}$  repair,  
 $d_n$  ( $d_n$ ) $_{n \geq 1}$  be working times of machine

$(E_n, d_n)_{n \geq 1}$  forms a renewal reward theorem

$$E_n = m [d_n/m]$$

$$E[d_n] = 1/\lambda$$

$$E[E_n] = m \sum_{k=1}^{\infty} k \int_{(k-1)m}^{km} \lambda e^{-\lambda x} dx$$

$$= m \sum_{k=1}^{\infty} k [e^{-\lambda x}]_{(k-1)m}^{km}$$

$$= m \sum_{k=1}^{\infty} k e^{-\lambda(k-1)m} - m \sum_{k=1}^{\infty} k e^{-\lambda k m}$$

$$= m \sum_{k=0}^{\infty} (k+1) e^{-\lambda k m} - m \sum_{k=0}^{\infty} k e^{-\lambda k m}$$

$$= m \sum_{k=0}^{\infty} (e^{-\lambda m})^k = \frac{m}{1 - e^{-\lambda m}} = \frac{m}{1 - e^{-\lambda m}}$$

$$\therefore \text{IP working} = \frac{1/\lambda}{m/1 - e^{-\lambda m}} = \frac{1 - e^{-\lambda m}}{\lambda m}$$

4.) (a) Customers arrive at times of a Poisson  $\lambda$  process, and there is 1 server; Service times are iid,  $\text{Exp}(\mu)$

State space =  $\mathbb{Z}_{\geq 0}$ ,  $Q(n, m) = \lambda \cdot 1_{m=n+1} + \mu \cdot 1_{m=n-1} + \cancel{\lambda \cdot 1_{m=n}} + (-\lambda - \mu \cdot 1_{n=0}) \cdot 1_{m=n}$

(b) Let  $(d_n)_{n \geq 1}$  be  $\text{Exp}(\mu)$  R.V.;  $K \sim \text{Geo}(\delta)$  be independent

$$S = \sum_{r=1}^K d_r : \mathbb{E}[e^{sS}] = \mathbb{E}[\mathbb{E}[e^{s \sum_{r=1}^K d_r} : K \sim \text{Geo}(\delta)]]$$

$$= \mathbb{E}\left[\left(\frac{\lambda}{\lambda - \theta}\right)^K\right] = \sum_{k=1}^{\infty} \delta (1-\delta)^{k-1} \left(\frac{\lambda}{\lambda - \theta}\right)^k$$

~~$$= \delta \left( \frac{\lambda}{\lambda - \theta} \right) \sum_{k=1}^{\infty} (1-\delta)^{k-1} \left( \frac{\lambda}{\lambda - \theta} \right)^{k-1}$$~~

$$= \frac{\lambda \delta}{\lambda - \theta} \frac{1}{1 - (1-\delta) \frac{\lambda}{\lambda - \theta}} = \frac{\lambda \delta}{\lambda - \theta - (1-\delta)\lambda} = \frac{\lambda \delta}{\lambda \delta - \theta}$$

(MGF of  $\text{Exp}(\lambda \delta)$  R.V.)

(7/7) Since the # of service / pax  $\sim \text{Geo}(\delta)$ , the total service time is a  $\text{Exp}(\lambda \cdot \delta)$  R.V.

If we consider only the length of queue, if we can let a customer remain at the front and not be serviced until he/she leaves the system.

$\therefore M(\lambda) / M(\delta \cdot \mu) / 1$  queue.

If  $\lambda < \delta \mu$ : Invariant distribution exists (part cii)

$$\pi_n = (\lambda / \delta \mu)^n$$

Since ~~Q matrix of M/M/1 queue~~ is bounded, chain is non-explosive.  $\therefore$

Q matrix bounded  $\Rightarrow$  Non-explosive;  $\therefore$  Chain is +ve recurrent

Since M/M/1 is a birth death chain, invariant  $\Rightarrow$  Detailed Balance.

∴ Equilibrium is possible (+ve recurrence), process reversible at equilibrium. ✓

(Using Customer remains at front until full completion):

Since departure corresponds to arrival of reversed process, reversed process has the same distribution as the original process: ✓

Departure process ~ Poisson  $\lambda$ -rate process ✓

(Relative position of customer do not affect departure process)

In equilibrium:  $IP(\text{queue empty}) = \pi_0$ ,  $\pi_0 = (1 - (\lambda/s\mu))$

$$IE[\# \text{ of arrivals in } [0, 1]] = \underbrace{IE[\text{Poi}(\lambda)]}_{\text{from outside}} + IE\left[\int_0^1 \mu \cdot \mathbb{1}_{\text{Queue not empty}} dt\right]$$

$$= \lambda + \mu \cdot (\lambda/s\mu) = \lambda(1 + 1/s)$$

If arrival process poissonian:  $\text{Poi}(\lambda(1 + 1/s))$  rate process.

$$\therefore \text{Invariant queue length: } \tilde{\pi}_n = c \left( \frac{\lambda(1 + 1/s)}{\mu} \right)^n = c \left( \frac{\lambda(1+s)}{s\mu} \right)^n$$

$$\therefore \tilde{\pi} \neq \pi \Rightarrow \text{Contradiction.}$$

$IE[\text{Time to first arrival}]$ :

$$\text{If queue empty (IP } 1 - \lambda/s\mu): IE[\text{Exp}(\lambda)] = 1/\lambda \quad \checkmark$$

$$\text{Else: } IE[\text{Min}\{\text{Exp}(\lambda), \text{Exp}(\mu)\}]$$

$$= 1/(\lambda + \mu) \quad \checkmark$$

$$\begin{aligned} \therefore IE &= (1 - \lambda/s\mu) \cdot 1/\lambda + \lambda/s\mu \cdot 1/(\lambda + \mu) \\ &= 1/\lambda - \lambda/s\mu (1/\lambda - 1/(\lambda + \mu)) = 1/\lambda - \lambda^2/s\mu\lambda(\lambda + \mu) \\ &= 1/\lambda - 1/s(\lambda + \mu) \quad \checkmark \end{aligned}$$

$\therefore$  If Poissonian:  $Poi \left( \underbrace{\frac{\lambda}{\mu} - \frac{1}{s(\lambda+\mu)}}_{\delta \lambda_1} \right)$  process

Invariant:  $\left( \frac{\lambda_1}{\mu} \right)^n \neq$

Invariant:  $Geo \left( 1 - \frac{\lambda_1}{\mu} \right) \neq Geo \left( 1 - \frac{\lambda}{s\mu} \right)$

$\therefore$  Not Poissonian. ✓

(3/3)  
(Prod)

## Applied Probability 2021

1) (a) Claim: Let  $A \subseteq \mathbb{R}^3$  be bounded, measurable.  $|\Pi_1 \cap A| + |\Pi_2 \cap A| = |A \cap (\Pi_1 \cup \Pi_2)|$  (A.S.)

(Activity of  $\Pi_1, \Pi_2: \lambda, \nu$ )

IP: Let  $B_{N,k} = \prod_{i=1}^3 [k_i \cdot 2^{-N}, (k_i+1) 2^{-N})$ ,  $N \in \mathbb{N}$ ,  $k \in \mathbb{Z}^3$

$$\mathbb{P}(\text{~~not~~ } |\Pi_1 \cap \Pi_2 \cap A| \neq 0) \leq \sum_{k \in \mathbb{Z}^3} \mathbb{P}(|\Pi_1 \cap \Pi_2 \cap A \cap B_{N,k}| \neq 0)$$

$$\leq \sum_{k \in \mathbb{Z}^3} \mathbb{P}(|\Pi_i \cap A \cap B_{N,k}| > 0 : i=1,2) \leq \sum_{k \in \mathbb{Z}^3} (1 - e^{-\int_{A \cap B_{N,k}} \lambda dx} \cdot (1 - e^{-\int_{A \cap B_{N,k}} \nu dx}))$$

$$\leq \sum_{k \in \mathbb{Z}^3} \left( \int_{A \cap B_{N,k}} \lambda dx \right) \cdot \left( \int_{A \cap B_{N,k}} \nu dx \right)$$

$$\leq \lambda \nu \sum_{k \in \mathbb{Z}^3} \sup_{k \in \mathbb{Z}^3} \left[ \text{Vol}(A \cap B_{N,k}) \right] \text{Vol}(A \cap B_{N,k})$$

$$\leq \lambda \nu \text{~~Vol}(B_{N,k})~~ \sup_{k \in \mathbb{Z}^3} \left[ \text{Vol}(A \cap B_{N,k}) \right] \cdot \sum_{k \in \mathbb{Z}^3} \text{Vol}(A \cap B_{N,k})$$

$$= \lambda \nu \text{Vol}(A) \cdot 2^{-3N}$$

$$(\text{True } \forall N \in \mathbb{N}) : \mathbb{P}(|\Pi_1 \cap \Pi_2 \cap A| \neq 0) \leq \lim_{N \rightarrow \infty} \left\{ \lambda \nu |A| 2^{-3N} \right\} = 0$$

$\therefore \mathbb{P} = 0.$

$$\therefore |\Pi_1 \cap A| + |\Pi_2 \cap A| = |A \cap (\Pi_1 \cup \Pi_2)| \quad \text{W.H.T.}$$

For fixed bounded measurable  $A$ :  $|( \Pi_1 \cup \Pi_2 ) \cap A|$  is the sum of independent Poisson ~~Poisson~~  $\text{Poi} \left( \int_A z_i dx \right)$  P.V.

$$\therefore \sim \text{Poi} \left( \int_A (z_1 + z_2) dx \right)$$

disjoint

If  $(A_n)_{n \geq 1}$  independent:  $(|A_n \cap \Pi_i|)_{n \geq 1}$  are independent $\Pi_1, \Pi_2$  independent:  $\{|A_n \cap \Pi_i| : n \geq 1, i=1,2\}$  are all independent $\therefore (|A_n \cap (\Pi_1 \cup \Pi_2)|)_{n \geq 1}$  is independent

(7/7)

 $\therefore$  Ideal gas of 2 activity  $z_1 + z_2$ 

cb) Consider characteristic functions:

 $N(V_i) \sim \text{Poi}(z \cdot |V_i|)$ ; Let  $|V_i| = d_i$ 

$$\mathbb{E} [ e^{i\theta (N(V_i) - \mathbb{E}[N(V_i)] / |V_i|)} ] = \mathbb{E} [ e^{i\theta ( \text{Poi}(z d_i) - z d_i) / d_i} ]$$

$$= e^{-i\theta z d_i} \mathbb{E} [ e^{i\theta \text{Poi}(z d_i) / d_i} ]$$

$$= e^{-i\theta z d_i} \frac{(e^{i(\theta/d_i)} - 1) z d_i^2}{e}$$

$$\lim_{d_i \rightarrow \infty} \left\{ (e^{i(\theta/d_i)} - 1) z d_i^2 - i\theta z d_i \right\}$$

$$= \lim_{d_i \rightarrow \infty} \left\{ -\frac{\theta^2 z}{2} + o(1/d_i) \right\} = -\frac{z\theta^2}{2}$$

pointwise.

 $\therefore$  Converges to characteristic function of  $N(0, z)$  $\therefore \rightarrow N(0, z)$  in distribution

$$c_1) \text{ Let } K = \text{Supp}(g), \quad X = \sum_{x \in \Pi} g(x)$$

$$\mathbb{E} [ e^{\theta X} ] = \mathbb{E} [ \mathbb{E} [ e^{\theta \sum_{x \in \Pi \cap K} g(x)} \mid |\Pi \cap K| = n ] ]$$

$$\phi(\theta) = \mathbb{E} [ \mathbb{E} [ e^{\theta g(u)} ]^n : n \sim \text{Poi}(\int_K z) ]$$

$$= e^{(\int_K z dx) (-1 + \mathbb{E} [ e^{\theta g(u)} ] )} \quad u \sim \text{Uniform}(K)$$



$$= e^{\sum_k |k| \cdot \int_k \frac{1}{|k|} (e^{\theta g(x)} - 1) dx} = e^{\sum_k \int_k (e^{\theta g(x)} - 1) dx}$$

$$\phi'(\theta) \Big|_{\theta=0} : \sum_k \int_k g(x) dx = \mathbb{E}[X]$$

$$\phi''(\theta) \Big|_{\theta=0} : \sum_k \int_k g^2(x) dx + \left( \sum_k \int_k g(x) dx \right)^2 = \mathbb{E}[X^2]$$

$$\therefore \text{Var}(X) = \sum_k \int_k g^2(x) dx = \sum_k \int_{\mathbb{R}^3} g^2(x) dx$$

$$\mathbb{E}[X] = \sum_k \int_{\mathbb{R}^3} g(x) dx$$

(66)  
(200)

2.) (a) Let  $T_0 = 0$ ,  $T_{n+1} = \inf \{t \geq T_n : X_t \neq X_{T_n}\}$  ( $n \geq 1$ ):

$Y_n = X_{T_n}$ ;  $(Y_n)_{n \geq 0}$  is the jump chain.

Consider process  $\tilde{X}$  generated as follows:

$$P(x, y) = 1_{x \neq y} \cdot Q(x, y) / q(x) \quad q(x) = -Q(x, x)$$

$(\tilde{Y}_n)_{n \geq 0}$  is a discrete markov chain,  $\tilde{Y}_0 = X_0$  and transition matrix =  $P$ .

Let  $(\xi_n)_{n \geq 1}$  be iid  $\text{Exp}(1)$  R.V.;  $\tilde{T}_0 = 0$ ,  $\tilde{T}_n = \sum_{k=1}^n \xi_k / q(\tilde{Y}_{k-1})$

$$\tilde{X}_t = \tilde{Y}_k \quad \text{for } t \in [\tilde{T}_k, \tilde{T}_{k+1})$$

$X$ ,  $\tilde{X}$  have the same distribution;  $\tilde{Y}$  is the jump chain of  $\tilde{X} \Rightarrow \tilde{Y}, Y$  have the same distribution

$\tilde{Y}$  is a markov chain  $\Rightarrow Y$  is a markov chain

(4/4)

(b)  $X$  is recurrent if  $\forall$  state  $s$ :  $\mathbb{P}(\{t \geq 0 : X_t = s\} \text{ is unbounded}) = 1$

# Since  $X$  is irreducible: Recurrence is equivalent to 1 state being recurr

If  $(Y_n)_{n \geq 0}$  is recurrent:  $Y_n$  visits state  $s$  infinitely often  
 $\therefore \exists$  subsequence  $n_k$  s.t.  $J_{n_k} \in \{t \geq 0 : X_t = s\}$

$X$  non explosive:  $J_n \rightarrow \infty \Rightarrow J_{n_k} \rightarrow \infty$  (A.S.)

$\therefore \{t \geq 0 : X_t = s\}$  is unbounded (A.S.)

$\therefore$  Recurrent.

If  $(Y_n)_{n \geq 0}$  is transient:  $Y_n$  visits state  $s$  finitely often

If  $Y_n \neq s$  for  $n > N$ :  $\{t \geq 0 : X_t = s\} \subseteq [0, J_N]$ ,

$J_N < \infty$  (A.S.)

Since  $\exists N$  s.t.  $Y_n \neq s$  for  $n > N$  A.S.,  $\{t \geq 0 : X_t = s\}$  is bounded

has  $\mathbb{P} = 0$ .  $\therefore \mathbb{P}(\{X_t\}_{t \geq 0} \text{ is unbounded}) = 0$ .  $\therefore$  Not recurrent.

$$\int_0^\infty P_{00}(t) dt = \int_0^\infty \mathbb{E}_0[1_{X_t=0}] dt = \mathbb{E}_0\left[\int_0^\infty 1_{X_t=0} dt\right]$$

$$= \mathbb{E}_{x_0}\left[\sum_{n \geq 1} J_n \cdot 1_{Y_n=0}\right] = \frac{1}{q(0)} \cdot \mathbb{E}_0[\text{\# of visits to } 0 \text{ by } (Y_n)_{n \geq 0}]$$

$$= \infty \Rightarrow (0)_n \text{ is } (Y_n)_{n \geq 0} \text{ recurrent}$$

$$\Rightarrow X \text{ - recurrent.}$$

(6/6)

(c) By considering jump chain probabilities: Equivalent to consider.

$$q_{x,y} = e^{-|x-y|^2}$$

$$q_{x,y} = \frac{1}{(1+|x-y|)^2}$$

$$\int_0^{\infty} P_{0,0}(t) dt = \frac{1}{2\pi} \int_0^{\infty} \int_{-\pi}^{\pi} e^{-t\lambda(k)} dk dt$$

$$= \frac{1}{2\pi} \int_{-\pi}^{\pi} \frac{1}{\lambda(k)} dk$$

$$\lambda(k) = \sum_{z \in \mathbb{Z}} (1 - e^{ikz}) e^{-z^2} = R + iI, \quad R, I \in \mathbb{R}$$

$$0 \leq R(k) = \sum_{z \in \mathbb{Z}} (1 - \cos kz) e^{-z^2} \quad (\text{all terms } \geq 0)$$

$$\leq \sum_{z \in \mathbb{Z}} R(k) k^2 z^2 e^{-z^2} = A \cdot k^2, \quad A \text{ is constant in } k$$

$$|I(k)| \leq \sum_{z \in \mathbb{Z}} | \sin(kz) | \cdot e^{-z^2} \leq \sum_{z \in \mathbb{Z}} |k| |z| e^{-z^2} = B \cdot |k|$$

$$I(k) = -I(-k)$$

$$\therefore \int_0^{\infty} P_{0,0}(t) dt = \frac{1}{2\pi} \int_{-\pi}^{\pi} \frac{1}{R+iI} dk = \frac{1}{2\pi} \int_{-\pi}^{\pi} \frac{R-iI}{R^2+I^2} dk$$

$$= \frac{1}{\pi} \int_0^{\pi} \frac{R(k)}{R^2(k)+I^2(k)} dk \geq \frac{1}{\pi} \int_0^{\pi} \frac{R(k)}{A^2 k^4 + B^2 k^2} dk$$

\* Sus part:  $I_R = \sum_{z \in \mathbb{Z}} (1 - \cos kz) e^{-z^2}$

$$= \frac{1}{\pi} \sum_{z \in \mathbb{Z}} e^{-z^2} \int_0^{\pi} \frac{1 - \cos kz}{A^2 k^4 + B^2 k^2} dk$$

$$I(k) = \sum_{z \in \mathbb{Z}} \sin(kz) e^{-z^2}$$

$$= 0 + \sum_{n \in \mathbb{N}} \{ \sin(kn) + \sin(-kn) \} e^{-n^2} = 0$$

$$\therefore \int_0^\infty P_{00}(t) dt = \frac{1}{2\pi} \int_{-\pi}^{\pi} \frac{1}{R(k)} dk \geq \frac{1}{2\pi} \int_{-\pi}^{\pi} \frac{1}{Ak^2} dk = \infty$$

(10/10)

 $\therefore 1^{\text{st}}$  chain is recurrent.

(20/2)

3.) (a) State space =  $\mathbb{Z}_{\geq 0}$ :

$$Q \text{ matrix: } Q(n, n+1) = \lambda r^n$$

$$Q(n, n-1) = \min\{n, 2\} \cdot \mu \quad (n \geq 1)$$

$$Q(n, n) = \lambda \cdot r^n + \min\{n, 2\} \cdot \mu$$

Since this is a birth-death cycle:  $\pi$  is invariant iff  $\pi$  satisfies Detailed Balance.

$$\pi_0 \cdot \lambda = \pi_1 \cdot \mu$$

$$\pi_n \lambda r^n = \pi_{n+1} (2\mu) \quad (n \geq 1)$$

$$\therefore \pi_{n+1} = \left(\frac{\lambda}{2\mu}\right)^n r^{1+\dots+n} = \left(\frac{\lambda}{2\mu}\right)^n r^{(n+1) \cdot n/2}$$

If  $r=1$ :  $\pi$  is an invariant distribution iff

$$\sum_{n \geq 1} \left(\frac{\lambda}{2\mu}\right)^n r^{(n+1)n/2} = \sum_{n \geq 1} \left(\frac{\lambda}{2\mu}\right)^n < \infty$$

$$\text{iff } \lambda < 2\mu$$

If  $r < 1$ : Given  $\lambda, \mu$ : Pick  $N \in \mathbb{N}$  s.t.  $n \geq N \Rightarrow$

$$r^{(n+1)/2} \cdot \left(\frac{\lambda}{2\mu}\right) < 1/2$$

$$\begin{aligned} \therefore \sum_{n \geq 1} \left(\frac{\lambda}{2\mu}\right)^n r^{(n+1) \cdot n/2} &= \sum_{n < N} \left(\frac{\lambda}{2\mu}\right)^n r^{(n+1)n/2} \\ &\quad + \sum_{n \geq N} \left(\frac{\lambda}{2\mu}\right)^n r^{(n+1)n/2} \\ &\leq \sum_{n < N} \left(\frac{\lambda}{2\mu}\right)^n r^{(n+1)n/2} + \sum_{n \geq N} \left(\frac{1}{2}\right)^n < \infty. \end{aligned}$$

$\therefore$  Invariant distribution exists  $\forall \lambda, \mu$ .

$$\therefore r < 1 \text{ or } (r=1, \lambda/2\mu < 1)$$

(b) State space:  $S = \{0, \dots, N\}$  (Finite)

$$Q\text{-matrix: } Q(n, n+1) = \lambda \quad (0 \leq n < N)$$

$$Q(n, n-1) = \mu \quad (0 < n \leq N)$$

$$Q(n, n) = \lambda \cdot 1_{n < N} + \mu \cdot 1_{n > 0}$$

$$\lambda = 1$$

Since chain is irreducible: ~~Jump~~ chain is recurrent  $\Rightarrow$  Chain is recurrent.  $(X_t)_{t \geq 0}$

$\therefore$  Any invariant distribution is unique

$Q$  / finite dim. matrix  $\Rightarrow$  solution to  $A \cdot Q = 0$   
 $\therefore r(Q) < \dim(Q)$

$$\text{Detailed Balance: } \pi_n \cdot \lambda = \mu \pi_{n+1}$$

$$\therefore \pi_n = \left(\frac{\lambda}{\mu}\right)^n \pi_0 \quad (0 \leq n \leq N)$$

$$\pi_n = \left(\frac{\lambda}{\mu}\right)^n / \sum_{k=0}^N \left(\frac{\lambda}{\mu}\right)^k \text{ is an invariant distribution}$$

(6/6) cc1. Since we have solved detailed balance, process is reversible at equilibrium. Let  $\tilde{X}$  be reversed process.

Since departure of  $(X_t)_{t \geq 0}$  corresponds to arrivals in  $(\tilde{X}_t)_{t \geq 0}$ , and  $\tilde{X}$  has the same distribution as  $X$ :

(7/7) De Arrival process at equilibrium: Customers arrive at rate  $\lambda$ ; Each customer will be turned away with probability  $\mu$  (if queue store is full), independently (chain in equilibrium) why?

$\therefore$  Arrival process is a thinned Poisson process  $\Rightarrow$  Poissonian

$\therefore$  Departure process is also Poissonian

\* Comment: Last part seems a little suspicious. I am really appealing to the "Pasta" property.

$$\begin{aligned}
 4.) \quad c_1) \quad f_t(c, \lambda) &= \mathbb{E}_i \left[ f(X(t), \lambda + L(t)) \right] \\
 &= e^{-q_i t} f(c, \lambda + t e_i) + \int_0^t \sum_{j \neq i} \Omega_{ij} e^{-q_i s} \mathbb{E}_i \left[ f(X(t-s), \lambda + s e_i + L(t-s)) \right] ds \\
 e^{q_i t} f_t(c, \lambda) &= f(c, \lambda + t e_i) + \sum_{j \neq i} \int_0^t \Omega_{ij} e^{q_i s} \mathbb{E}_i \left[ f(X(s), \lambda + (t-s)e_i + L(s)) \right] ds \\
 \therefore e^{q_i t} \left\{ q_i f_t(c, \lambda) + \frac{\partial f_t}{\partial t}(c, \lambda) \right\} &= \frac{\partial f}{\partial \lambda_i}(c, \lambda + t e_i) + \sum_{j \neq i} \int_0^t \Omega_{ij} e^{q_i s} \frac{\partial f}{\partial \lambda_i}(j, \lambda + (t-s)e_i) ds \\
 + \sum_{j \neq i} \left\{ \Omega_{ij} e^{q_i t} f_t(j, \lambda) + \int_0^t \Omega_{ij} e^{q_i s} \frac{\partial f}{\partial t}(j, \lambda + (t-s)e_i) ds \right\} \\
 \uparrow \quad e^{q_i t} \cdot \sum_j \Omega_{ij} f_t(j, \lambda) &= \sum_{j \neq i} \int_0^t \Omega_{ij} e^{q_i s} \frac{\partial f}{\partial \lambda_i}(j, \lambda + (t-s)e_i) ds \\
 e^{q_i t} \frac{\partial f_t}{\partial t}(c, \lambda) + \frac{\partial f_t}{\partial t}(c, \lambda) &= \sum_{j \neq i} \Omega_{ij} f_t(j, \lambda) \\
 \frac{\partial f_t}{\partial t} &= \underline{\Omega} \cdot f_t(c, \lambda)
 \end{aligned}$$



4.) (a)  $\mathbb{E}_i [f(X(t), \lambda + L(t))] = e^{-q_i t} f(i, \lambda + t e_i) +$   
 $\int_0^t \sum_{j \in I} \omega_{ij} e^{-q_i s} f_{t-s}(j, \lambda + s e_i) ds$

$$\therefore e^{q_i t} \cdot f_t(i, \lambda) = f(i, \lambda + t e_i) + \sum_{j \in I} \int_0^t \omega_{ij} e^{q_i s} f_s(j, \lambda + (t-s)e_i) ds$$

Diff. wrt  $t$ :

$$e^{q_i t} \cdot \left\{ q_i f_t(i, \lambda) + \frac{\partial}{\partial t} f_t(i, \lambda) \right\} = \frac{\partial}{\partial \lambda_i} f(i, \lambda + t e_i) + \sum_{j \in I} \omega_{ij} e^{q_i t} f_t(j, \lambda) + \sum_{j \in I} \int_0^t \omega_{ij} e^{q_i s} \frac{\partial}{\partial \lambda_i} f_s(j, \lambda + (t-s)e_i) ds$$

$$\therefore \frac{\partial f_t}{\partial t}(i, \lambda) = \sum_{j \in I} \omega_{ij} f_t(j, \lambda) + \frac{\partial f}{\partial \lambda_i}(i, \lambda + t e_i) + \int_0^t \sum_{j \in I} \omega_{ij} e^{q_i s} \frac{\partial f_s}{\partial \lambda_i}(j, \lambda + (t-s)e_i) ds$$

$$= \mathbb{E}_i \left[ \frac{\partial f}{\partial \lambda_i}(X(t), \lambda) \right] = \frac{\partial}{\partial \lambda_i} f_t(i, \lambda)$$

(7/7)  $\therefore \frac{\partial f_t}{\partial t} = M \cdot f_t(i, \lambda)$

(b)  $\int_{\mathbb{R}^n} e^{-\frac{1}{2} y^T \Omega y} \sum_{j \in I} y_j \sum_{k \in I} \omega_{jk} f(k, y^{\frac{1}{2}}) dy$

Let  $f$  be supported in  $K$ :

Consider the vector  $\underline{A} \in \mathbb{R}^n$ ,  $A_j = y_j f(j, y^{\frac{1}{2}})$

$$\nabla \cdot \underline{A} = \sum_{j=1}^n y_j \frac{\partial f}{\partial \lambda_j}(j, y^{\frac{1}{2}}) \cdot y_j + f(i, y^{\frac{1}{2}})$$

$$\nabla \cdot (e^{\frac{1}{2} y^T \Omega y} \cdot \underline{A}) = e^{\frac{1}{2} y^T \Omega y} \nabla \cdot \underline{A} + e^{\frac{1}{2} y^T \Omega y} (\Omega \cdot y)^T \cdot \underline{A}$$

Apply Divergence theorem to the region  $K$ :  $A_j = 0$  as  $f = 0$  on  $\partial K$

$$0 = \int_K \nabla \cdot (e^{\frac{1}{2} y^T \omega y} A) dy = \int_{\mathbb{R}^n} \nabla \cdot (e^{\frac{1}{2} y^T \omega y} A) dy \quad \text{as } A = 0 \text{ outside } K.$$

$$= \int_{\mathbb{R}^n} e^{\frac{1}{2} y^T \omega y} \cdot \left\{ y_i \sum_{j=1}^n \frac{\partial f(j, y'/2)}{\partial \lambda_j} y_j + f(i, y'/2) + \sum_{j=1}^n \sum_{k=1}^n y_{jk} \omega_{jk} y_i f(k, y'/2) \right\} dy$$

$$= \int_{\mathbb{R}^n} e^{\frac{1}{2} y^T \omega y} f(i, y'/2) + y_i \sum_{j=1}^n \left( \sum_{k=1}^n y_{jk} \omega_{jk} f(k, y'/2) \right) - \frac{\partial f(j, y'/2)}{\partial \lambda_j} dy$$

$$\therefore \int_{\mathbb{R}^n} e^{\frac{1}{2} y^T \omega y} f(i, y'/2) dy = - \int_{\mathbb{R}^n} e^{\frac{1}{2} y^T \omega y} y_i \sum_{j=1}^n y_j M f(j, y'/2) dy$$

(4/6)

$$(c) \int_{\mathbb{R}^n} e^{\frac{1}{2} y^T \omega y} \left\{ \mathbb{E}_i [g(\lambda + L(t)) | X(t) = j] dy + y_i \sum_{j=1}^n y_j \frac{\partial}{\partial t} f(j, y'/2) \right\} dy = 0$$

as for  $\|\lambda\| \geq t + T$ ,  $\|\lambda + L(t)\| \geq T \Rightarrow g = 0$  (c.b) valid

~~$$\frac{\partial}{\partial t} \mathbb{E}_i [g(\lambda + L(t)) | X(t) = j] = 0$$~~

~~$$\lim_{t \rightarrow \infty} \mathbb{E}_i [g(\lambda + L(t)) | X(t) = j] = 0$$~~

For fixed  $y$ :  $\mathbb{E}_i [g(y'/2 + L(t)) | X(t) = j] = 0$  for  $t \geq y'/2 + T$   
 $(\|y'/2 + L(t)\| \geq T)$

$$\therefore \int_0^\infty \frac{\partial}{\partial t} f_i(j, y'/2) dt = - \mathbb{E}_i [g(\lambda + y'/2)]$$

$$\therefore \int_{\mathbb{R}^n} e^{\frac{1}{2} y^T \omega y} \mathbb{E}_i [g(y'/2)] dy = \int_{\mathbb{R}^n} e^{\frac{1}{2} y^T \omega y} g(y'/2) \cdot y_i y_j dy$$

$$= \int_0^\infty \int_{\mathbb{R}^n} e^{\frac{1}{2} y^T \omega y} \mathbb{E}_i [g(y'/2 + L(t)) | X(t) = j] dy dt =$$

$$\int_{\mathbb{R}^n} e^{\frac{1}{2} y^T \omega y} \int_0^\infty \mathbb{E}_i [g(y'/2 + L(t)) | X(t) = j] dt dy$$

(6/6)

(20d)

No.: .....

Date: .....

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Hom Need some help on:

2022 Q2 part (a):